High-Throughput Sequencing Course Statistical Inference: Part II

Biostatistics and Bioinformatics



Summer 2018





TWO-SAMPLE MODEL: INFERENCE

- ► The mRNA abundance level in the untreated population is μ_0
- ▶ The mRNA abundance level in the untreated population is μ_1
- ► Assumed model:
 - ▶ Untreated Population: $Y = \mu_0 + \epsilon$
 - ▶ Treated Population: $X = \mu_1 + \epsilon'$
- ► Statistical Hypotheses
 - $H_0: \mu_0 = \mu_1$ (no treatment effect)
 - $H_0: \mu_0 \neq \mu_1$ (treatment effect)

TWO-SAMPLE MODEL: ESTIMATION

- ► What is often of interested is estimate the unknown parameters or quantities
- ► Examples
 - ▶ Mean level for the untreated group μ_0
 - ▶ Mean level for the treated group μ_1
 - ► Fold-change $\rho = \frac{\mu_1}{\mu_0}$
 - Standardized difference $\Delta = |\mu_1 \mu_0|/\sigma$
- ► Two types of estimates
 - ► Point estimate
 - ► Interval estimate

CONFIDENCE INTERVALS

- ► Example: The sample mean (the average of the observations) is a point estimate of the population (true) mean
- ► It is either equal to the true value of the parameter or is not
- ► As it is a single number it does not provide any direct measure of accuracy
- ► An interval estimate incorporates some measure of accuracy
- ► Thus it is generally more appropriate to present an interval estimate
- ► A common example of an interval estimate is the confidence interval

ESTIMATION EXAMPLE (ONE-SAMPLE MODEL)

- ► Truth: The RNA abundance follows a normal distribution with mean $\mu = 0$ and standard deviation $\sigma = 1$
- ▶ Assumption: The RNA abundance follows a normal distribution with unknown mean μ and unknown standard deviation σ
- ▶ Goal: The population mean μ is to be estimated on the basis of sample of size n=7
- ► Objectives:
 - Produce point estimate of μ
 - ▶ Produce a 95% confidence interval of μ

ESTIMATION EXAMPLE (SIMULATE DATA)

```
mu <- 0
sigma <- 1
n <- 7
set.seed(12131)
x <- rnorm(n, mu, sigma)
x

## [1] 1.5227356 -2.7829224 0.3571897 0.5478351 1.2733071 0.5166791
## [7] -1.3890287
```

POINT ESTIMATOR

- \blacktriangleright A point estimator of μ is the so called sample mean
- ▶ The sample mean \bar{x}_n is obtained by simply averaging all the observations
- ► Note that an alternative is to used the sample median (rather than sample mean)
- ► The sample median is obtained by first sorting the observations (in say ascending order)
- ► The median is the middle observation (among the sorted observation)
- ► The median is more robust against outliers

POINT ESTIMATES

► The data

```
x

## [1] 1.5227356 -2.7829224 0.3571897 0.5478351 1.2733071 0.5166791

## [7] -1.3890287
```

► The sample mean

```
mean(x)
## [1] 0.006542226
```

► The data sorted in ascending order

```
sort(x)
## [1] -2.7829224 -1.3890287 0.3571897 0.5166791 0.5478351 1.2733071
## [7] 1.5227356
```

► Sample median

```
median(x)
## [1] 0.5166791
```

CONFIDENCE INTERVAL ESTIMATORS

- ▶ To construct a confidence interval for μ we need to deal with the nuisance parameter σ
- We can estimate it using the sample standard deviation s_n (details omitted)
- ▶ A 95% confidence interval for μ is obtained as

$$[\bar{x}_n - \frac{s_n}{\sqrt{n}}t(0.975, n-1), \bar{x}_n + \frac{s_n}{\sqrt{n}}t(0.975, n-1)]$$

- ▶ t(0.975, n-1) is the 0.975 quantile of a t distribution with n-1=6 degrees of freedom
- $\frac{s_n}{\sqrt{n}}$ is called the standard error
- ▶ $\frac{s_n}{\sqrt{n}}t(0.975, n-1)$ is called the margin of error
- ► The confidence interval is obtained as the point estimate plus or minus the margin of error

SIMULATE EXPERIMENT 1

► Calculate the sample mean

```
xbar <- mean(x)
xbar
## [1] 0.006542226
```

► Calculate standard deviation

```
s <- sd(x)
s
## [1] 1.544261
```

► Calculate standard error

```
se <- s/sqrt(n)
se
## [1] 0.5836759
```

► Calculate margin of error

```
me <- qt(0.975, df = n - 1) * se me ## [1] 1.428204
```

► Calculate 95% CI

```
c(xbar - me, xbar + me)
## [1] -1.421661 1.434746
```

COVERED OR NOT COVERED

- ▶ The goal is to estimate μ
- ▶ If μ (the true but unknown parameter) is contained in the confidence interval, we say that it is covered
- ► Otherwise, it is not covered
- ▶ Note that when doing a simulation study, we can ascertain if μ is covered or not.
- ► Why?
- ▶ In real data analysis, we cannot ascertain if μ is covered by the confidence interval
- ► Why?
- ▶ We can only state that we are 95% confident that μ is covered by the interval estimate based on the data from our experiment
- ► More on "confidence" later

REPEAT THE EXPERIMENT

Repeat the Experiment 10 times

exp	n	mu	sigma	xbar	S	lcl	ucl	cover	len
1	7	0	1	0.48	0.42	0.09	0.87	FALSE	0.78
2	7	0	1	0.34	0.88	-0.47	1.15	TRUE	1.63
3	7	0	1	-0.51	1.18	-1.60	0.58	TRUE	2.18
4	7	0	1	-0.87	0.67	-1.49	-0.25	FALSE	1.24
5	7	0	1	-0.09	0.95	-0.97	0.78	TRUE	1.76
6	7	0	1	0.30	1.62	-1.20	1.80	TRUE	3.00
7	7	0	1	-0.68	0.52	-1.15	-0.20	FALSE	0.96
8	7	0	1	0.06	1.30	-1.15	1.26	TRUE	2.41
9	7	0	1	0.28	1.02	-0.66	1.23	TRUE	1.89
10	7	0	1	-0.31	0.48	-0.76	0.14	TRUE	0.89

CONFIDENCE INTERVAL: COMMON

MISUNDERSTANDING

- ► A (not the) 95% CI for the mean based on the first experiment was (0.09, 0.87)
- ▶ A (not the) 95% CI for the mean based on the second experiment was (-0.47, 1.15)
- ▶ It is wrong to say that the probability that the first CI does not contain the true value $\mu = 0$ is 95%
- ▶ It is also wrong to say that the probability that the second CI contains the true value $\mu = 0$ is 95%
- ▶ We conduct one and only one experiment
- ▶ Based on the first experiment, we can say that we are 95% confident that it contains the true value
- \blacktriangleright Note that μ is not covered by the first experiment
- \blacktriangleright If we repeated the experiment a large number of times, 95% of the CIs would cover the true value
- ► We are 95% confident that the first experiment is among these (which it is not)

RECAP: ASSUMPTIONS

- ► We do not need to make distributional assumptions (e.g., normality) on the sample mean for the purpose of point estimation
- ► The sample mean, however, is not robust against outliers
- ► Why did 1984 UNC geography graduates have high average salary?
- We made distributional assumptions for using the confidence interval
- \blacktriangleright The margin of error was based on a t distribution

A MORE COMPLICATED EXAMPLE: OUTLINE

- ▶ Suppose that you are measuring a quantity that is between 0 and θ
- ▶ How would you estimate θ ?
- ► Would you take the sample average?
- ► How about the sample mean?
- ► If the measurements are uniformly distributed, it turns out that the maximum observation is an "optimal" estimator
- ▶ It is also intuitively speaking a "reasonable" estimator
- ► Why?

A MORE COMPLICATED EXAMPLE: SIMULATION

► Simulate data from a uniform distribution on [0, 1]

```
n <- 10
theta <- 1
set.seed(2313)
x <- runif(n, 0, theta)
x
## [1] 0.34807917 0.12084940 0.11035999 0.03917718 0.79590237 0.72536724
## [7] 0.80347454 0.95498314 0.62601926 0.19549397</pre>
```

► Sample mean

```
mean(x)
## [1] 0.4719706
```

► Sample median

```
median(x)
## [1] 0.4870492
```

► Maximum observation

```
max(x)
## [1] 0.9549831
```

A MORE COMPLICATED EXAMPLE: RECAP

- ► An estimator is "valid" if it depends only on the data and no unknown quantities (including the parameter to be estimated)
- ► Why?
- ▶ Both the sample mean and median are valid estimators of θ
- ► There are, however, not good estimators
- ► In fact, in this case, the sample mean and median should be close to 0.5
- ► Why?
- ► The maximum observation is not only a valid estimator but also intuitively reasonable estimator
- ► This example has a rich history

QUICK NOTE: ESTIMATE VERSUS ESTIMATOR

- ▶ We use the terms estimate and estimators interchangeably
- ► There is a subtle but important distinction
- ► Suppose that you decide to estimate the population mean using the sample mean (once you get the data)
- ► The sample mean is the estimator
- ► Its outcome is random before you collect the data
- ► Once you collect the data and plug them into the estimator you get an (not the) estimate