

# On twisted sums of random multiplicative functions (Warwick)

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## 1 Motivation, history

Let  $\alpha_S: \mathbb{N} \rightarrow S^1$  be Steinhaus (completely multiplicative),  $\alpha_S: \mathbb{N} \rightarrow \{-1, 0, 1\}$  be Rademacher.

Original motivation:  $\alpha_R$  models  $\mu$ . RH is  $\sum_{n \leq x} \mu(n) \ll x^{1/2+\varepsilon}$ . Wintner (1944) proved  $\sum_{n \leq x} \alpha_R(n) \ll x^{1/2+\varepsilon}$  holds almost surely **for all**  $x$ .

(Non-proof: Since  $\mathbb{E} \alpha_R(n) \alpha_R(m) = \delta_{n,m}$  for squarefree  $n, m$ , we have  $\mathbb{E}[(\sum_{n \leq x} \alpha_R(n))^2] \ll x$ , so  $\mathbb{P}(|\sum_{n \leq x} \alpha_R(n)| > Tx^{1/2}) \ll 1/T^2$ .)

It is expected that  $\alpha_R$  misses some properties of  $\mu$ . E.g.  $\sum_{n \leq x} \mu(n) = O(\sqrt{x}(\log \log \log x)^{5/4})$  is expected (Gonek) while Harper proved that if  $V(x) \rightarrow \infty$  then  $\sum_{n \leq x} \alpha_R(n) \geq \sqrt{x}(\log \log x)^{1/4}/V(x)$  for infinitely many  $x$  a.s. On the other hand,  $\alpha_S$  is the large- $T$  limit of  $(n^{it})_{t \leq T}$  and the large- $q$  limit of  $(\chi \bmod q)$ , e.g.

$$\sum_{n \leq x} n^{it} \xrightarrow{d} \sum_{n \leq x} \alpha_S(n)$$

as  $T \rightarrow \infty$  (since  $(p^{it})_{p \leq x} \xrightarrow{d} (\alpha_S(p))_{p \leq x}$ ; here  $t$  is uniform in  $[0, T]$ ). And we certainly want to understand such sums – they build up  $\zeta$ .

Significant work on moments of  $\sum_{n \leq x} \alpha_S(n)$  and law of iterated logarithm (Halász, Erdős, Lau-Tenenbaum-Wu, Basquin, Harper, Caich). Focus of this talk: distribution.

The distribution of  $\sum_{n \leq x} \alpha_S(n)$ , appropriately normalized is still not known, but very recently a precise conjecture was stated (2024). Harper's work (2017) shows that normalizing by standard deviation gives trivial limiting distribution (Helson), and that 'true' normalization is by  $\sqrt{x}/(\log \log x)^{1/4}$ . This relates to critical multiplicative chaos. Lots of activity on variants:

$$\sum_{n \leq x} \alpha_S(n) f(n).$$

E.g. very recently (2025), Hardy proved a version of the conjecture where  $f(n) = \mathbf{1}_{P(n) > \sqrt{n}}$ . Most works, until Hardy's, focused on much 'sparser'  $f$ , where correct normalization turns out to be **standard deviation**, and limiting distribution was **standard Gaussian**:

1.  $f(n) = \mathbf{1}_{\omega(n)=k}$ : Hough, Harper.  $k = o(\log \log x)$ . The sum has length  $x/\log^{1+o(1)} x$ .
2.  $f(n) = \mathbf{1}_{[x, x+H]}$ : Chatterjee-Sound ( $H = o(x/\log x)$ ), Sound-Xu ( $H \leq x/(\log x)^{2 \log 2 - 1 + \varepsilon}$ ), Pandey-Wang-Xu ( $H \ll_A x/\log^A x$  for every  $A > 0$ ), Harper-Sound-Xu.
3.  $f(n) = \mathbf{1}_{Q(\mathbb{Z})}$ : Najnudel, Klurman-Shkredov-Xu, Wang-Xu, Chinis-Shala.

None of these  $f$  are multiplicative.

## 2 Work with Wong

Motivated by works of Najnudel-Paquette-Sim, Mo Dick Wong and I considered

$$S_x := \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1/2} \sum_{n \leq x} \alpha_S(n) f(n)$$

for multiplicative  $f$ , with  $|f(p)|^2$  equal to  $\theta \in (0, 1/2)$  on average; recently extended to  $|f(p)|^2$  equal to  $\theta \in (0, 1)$  on average, inspired by ideas in Najnudel-Paquette-Simm-Vu. So cannot cover  $f \equiv 1$ , but can take e.g. indicator

of sums of two squares. In fact,  $f$  does not have to be bounded,  $f(p)$  can infinitely often be as large as  $p^{1/2+o(1)}$  ( $\sum_p |f(p)|^3 \log^C p / p^{3/2} < \infty$ ). If  $f$  takes the values 0 and 1 only, its support is  $\asymp x(\log x)^{\theta-1}$

What is the limiting distribution we found?  $G \cdot \sqrt{V}$  for standard complex Gaussian  $G$  independent of  $V$ , and  $V = 1/(2\pi) \int_{\mathbb{R}} |1/2 + it|^{-2} m_{\infty}(dt)$ . Here  $m_{\infty}$  is a random measure, translation-invariant, constructed by a limit from a sequence of measures involving  $\alpha$  and  $f$ . We use the notation  $G$  for standard complex Gaussian throughout. This talk will not be about  $m_{\infty}$  (see Mo Dick's talk for that), but about **how one connects the random sum with the random measure**.

One approach for this is through moments. It is not relevant here because the moments of  $V$  explode:  $\mathbb{E}V^p$  is finite iff  $p < 1/\theta$ , which forces  $\mathbb{E}|S_x|^{2p}$  to diverge for  $p \geq 1/\theta$  (recall that if  $X_n \xrightarrow{d} X$  in distribution then  $\liminf \mathbb{E}|X_n|^{2p} \geq \mathbb{E}|X|^{2p}$  by Fatou's lemma.)

Next approach is martingale CLT, introduced into this area by Harper (2010). Why martingales?

$$S_x = \sum_{p \leq x} Z_{p,x}$$

where

$$Z_{p,x} := \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1/2} \sum_{n \leq x, P(n)=p} \alpha_S(n) f(n).$$

If we define filtrations  $\mathcal{F}_{p-} = \sigma((\alpha(q) : q < p))$  then  $\mathbb{E}[Z_{p,x} | \mathcal{F}_{p-}] = 0$  (essentially conditioning of values of  $\alpha(q)$  for  $q < p$ .) Let

$$V_x = \sum_{p \leq x} |Z_{p,x}|^2.$$

Informally,  $S_x \approx G\sqrt{V_x}$ . Formally, **McLeish CLT** makes this formal in a special case: if  $\mathbb{E}V_x \rightarrow 1$ ,  $\mathbb{E}V_x^2 \rightarrow 1$ , and  $\sum_{p \leq x} \mathbb{E}|Z_{p,x}|^4 \rightarrow 0$ , then  $S_x \xrightarrow{d} G$ . Note that these conditions imply  $V_x \xrightarrow{p} 1$ . First condition is trivial:  $\mathbb{E}V_x = 1$  regardless of  $f$ . (Recent restatement by Sound–Xu: suffices to check  $\max_p \mathbb{E}|Z_{p,x}|^2 \rightarrow 0$  and  $\mathbb{E}|S_x|^4 \rightarrow 2$ , if  $f$  takes the values 0 and 1.)

This CLT is not relevant if limiting distribution is not Gaussian, so we turn to the general form of martingale CLT. We define

$$BP_x := \sum_{p \leq x} \mathbb{E}[|Z_{p,x}|^2 | \mathcal{F}_{p-}].$$

If  $BP_x \xrightarrow{p} V$ ,  $\sum_{p \leq x} \mathbb{E}|Z_{p,x}|^4 \rightarrow 0$  and  $\sum_{p \leq x} \mathbb{E}[Z_{p,x}^2 | \mathcal{F}_{p-}] \xrightarrow{d} 0$ , then  $S_x \xrightarrow{d} G\sqrt{V}$ . Rest of the talk will be about  $BP_x \xrightarrow{p} V$  – other conditions are ‘trivial’ to verify, *even in critical case* (e.g.  $\sum_{p \leq x} \mathbb{E}[Z_{p,x}^2 | \mathcal{F}_{p-}] \equiv 0$ ).

### 3 Two approaches to bracket process

It is convenient to assume that  $f$  is supported on squarefrees. Moreover, we shall only talk about Steinhaus (we put  $\alpha = \alpha_S$ ). We suppose  $\sum_{p \leq x} |f(p)|^2 \sim \theta \text{Li}(x)$ . First step is to understand how  $BP_x$  looks like. Note that

$$Z_{p,x} = \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1/2} \alpha(p) f(p) \sum_{m \leq x/p, P(m) < p} \alpha(m) f(m)$$

so

$$\mathbb{E}[|Z_{p,x}|^2 | \mathcal{F}_{p-}] = \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1} |f(p)|^2 \sum_{m \leq x/p, P(m) < p} \alpha(m) f(m)^2.$$

This shows

$$BP_x = \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1} \sum_{p \leq x} |f(p)|^2 \sum_{m \leq x/p, P(m) < p} \alpha(m) f(m)^2.$$

We introduce

$$s_{t,x} := t^{-1/2} \sum_{n \leq t, P(n) < x} \alpha(n) f(n)$$

so that

$$BP_x = \frac{x}{\sum_{n \leq x} |f(n)|^2} \sum_{p \leq x} \frac{|f(p)|^2}{p} |s_{x/p,p}|^2.$$

It is useful to massage  $BP_x$  a little: since  $F(x) = \sum_{p \leq x} |f(p)|^2/p \sim \theta \log \log x$  then, informally,

$$BP_x \approx \frac{x}{\sum_{n \leq x} |f(n)|^2} \int_{2^-}^x |s_{x/t,t}|^2 dF(t) \sim \theta \frac{x}{\sum_{n \leq x} |f(n)|^2} \int_{2^-}^x \frac{|s_{x/t,t}|^2}{t \log t} dt. \quad (1)$$

There are now (at least) two rather different approaches relating this to the measure  $m_\infty$ . Both approaches employ the following form of Plancherel's identity:

$$\int_{\mathbb{R}} |\sum_{n \leq t} g(n)|^2 t^{-2-2r} dt = \frac{1}{2\pi} \int_{\mathbb{R}} |\sum_n g(n) n^{-1/2-r-it}|^2 |1/2 + r + it|^{-2} dt \quad (2)$$

holds for  $r > 0$ , a.s. Proof: Through Perron.

**First argument.** We introduce

$$U_x := (\sum_{n \leq x} |f(n)|^2/n)^{-1} \int_1^x |s_{t,\infty}|^2 \frac{dt}{t}.$$

We combine two lemmas:

$$U_x \xrightarrow{p} \frac{1}{2\pi} \int_{\mathbb{R}} |1/2 + it|^{-2} m_\infty(dt) \quad (3)$$

and

$$\mathbb{E}|U_x - BP_x|^2 \rightarrow 0 \quad (4)$$

The statement (4) is proved by a straightforward technical computation, but isn't deep. The statement (3) is also not too deep (given the construction(s) of  $m_\infty$ ). Perhaps the difficult part is *guessing* that  $BP_x \xrightarrow{p} \frac{1}{2\pi} \int_{\mathbb{R}} |1/2 + it|^{-2} m_\infty(dt)$ , which is motivated by Plancherel.

To prove (4) one expands the square and computes three expectation using  $\mathbb{E}\alpha(n)\overline{\alpha(m)} = \delta_{n,m}$ . This leads to summing solutions to  $ab = cd$  with certain weights and constraints. Here it is useful to mention some facts:

- Solutions to  $ab = cd$  are given by  $a = n_1 n_2$ ,  $b = n_3 n_4$ ,  $c = n_1 n_3$ ,  $d = n_2 n_4$ . Proof: let  $g_1 = (a, c)$   $g_2 = (b, d)$  and define  $a', b', c', d'$  accordingly to get  $a'b' = c'd'$  with  $(a', c') = 1$  and  $(b', d') = 1$ . Then  $a' = d'$  and  $b' = c'$  forcing  $a = g_1 a'$ ,  $b = g_2 b'$ ,  $c = g_1 b'$ ,  $d = g_2 a'$ .
- De Bruijn-van Lint:  $\sum_{n \leq x, P(n) \leq y} |f(n)|^2 \sim \sum_{n \leq x} |f(n)|^2 \cdot \rho_\theta(\frac{\log x}{\log y})$  holds in the regime  $\log y \asymp \log x$ .
- Wirsing:  $\sum_{n \leq x} |f(n)|^2 \sim C_f x (\log x)^{\theta-1}$  if  $|f(p)|^2 \sim \theta$  on average.

Let us expand on (3). Notation:

$$m_x(dt) = |A(1/2 + 1/(2 \log x) + it)|^2 dt / \mathbb{E}|A(1/2 + 1/(2 \log x) + it)|^2$$

where  $A(s) = \sum_n f(n)/n^s$ . The measure  $m_\infty(dt)$  is constructed by taking  $\lim_x m_x(dt)$ . Note that

$$\mathbb{E}|A(1/2 + r/2 + it)|^2 = \mathbb{E}|\sum_n \alpha(n) f(n)/n^{1/2+r/2+it}|^2 = \sum_n |f(n)|^2/n^{1+r}$$

is independent of  $t$ . So, for  $r = 1/\log x$ , Plancherel implies that

$$\begin{aligned} \int_{\mathbb{R}} |s_{t,\infty}|^2 t^{-1-1/\log x} dt &= \frac{1}{2\pi} \int_{\mathbb{R}} |A(1/2 + 1/(2 \log x) + it)|^2 |1/2 + 1/2(1/\log x) + it|^{-2} dt \\ &= \frac{1}{2\pi} (\sum_n |f(n)|^2/n^{1+1/\log x}) \int_{\mathbb{R}} |1/2 + 1/(2 \log x) + it|^{-2} m_x(dt). \end{aligned}$$

From Wirsing,  $\sum_n |f(n)|^2/n^{1+1/\log x} \sim C_f \Gamma(\theta)(\log x)^\theta$ . From consequences of Plancherel and Wirsing,

$$(C_f \Gamma(\theta)(\log x)^\theta)^{-1} \int_{\mathbb{R}} |s_{t,\infty}|^2 t^{-1-1/\log x} dt \xrightarrow{p} \frac{1}{2\pi} \int_{\mathbb{R}} |1/2 + it|^{-2} m_\infty(dt).$$

(Here we relied on  $m_x(I) \xrightarrow{p} m_\infty(I)$ .) By Tauberian theorem (basically, approximating  $\mathbf{1}_{t \leq x}$  by  $P(t^{-1/\log x})$ ), this implies that

$$(C_f \Gamma(\theta)(\log x)^\theta)^{-1} \int_1^x |s_{t,\infty}|^2 \frac{dt}{t} \xrightarrow{p} \frac{1}{\Gamma(1+\theta)} \frac{1}{2\pi} \int_{\mathbb{R}} |1/2 + it|^{-2} m_\infty(dt)$$

as needed. (In  $U_x$  we divide by  $\sum_{n \leq x} |f(n)|^2/n$ , which is asymptotic to  $C_f(\log x)^\theta/\theta$  by Wirsing.)

**Second argument.** Some notation:

$$m_{y,x}(dt) = |A_y(1/2 + it + 1/(2 \log x))|^2 dt / \mathbb{E} |A_y(1/2 + it + 1/(2 \log x))|^2$$

where  $A_y(s) = \sum_{P(n) < y} f(n)/n^s$ . Recall (1). From Plancherel's identity, we may obtain the limit of a somewhat similar expression, namely

$$x \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1} \int_0^\infty q(t^{1/\log x}) \frac{|s_{x/t, x^a}|^2 dt}{t \log x} \quad (5)$$

for any fixed  $a > 0$  and any 'nice' function  $q$ . Details: Plancherel implies that

$$\begin{aligned} \int_{\mathbb{R}} |s_{t,y}|^2 t^{-1-\frac{r}{\log y}} dt &= \frac{1}{2\pi} \int_{\mathbb{R}} |A_y(\frac{1}{2} + \frac{r}{2 \log y} + it)|^2 |\frac{1}{2} + \frac{r}{2 \log y} + it|^{-2} dt \\ &= \mathbb{E} |A_y(\frac{1}{2} + \frac{r}{2 \log y})|^2 \cdot \frac{1}{2\pi} \int_{\mathbb{R}} |\frac{1}{2} + \frac{r}{2 \log y} + it|^{-2} m_{y,y^{1/r}}(dt) \end{aligned}$$

holds for  $r > 0$  so

$$\left( \sum_{P(n) < y} \frac{|f(n)|^2}{n^{1+r/\log y}} \right)^{-1} \int_0^\infty \frac{|s_{t,y}|^2 dt}{t^{1+r/\log y}} \xrightarrow{y \rightarrow \infty} V;$$

now substitute  $t = x/t$  and  $y = x^a$ . This gives (5) with  $q(z) = z^{r/a}$ . Here we rely on  $m_{x,y}(I) \xrightarrow{p} m_\infty(I)$  if  $x, y \rightarrow \infty$  together. The main difference between (1) and (5), however, is that in (5) the 'smoothness parameter' in  $s_{x/t, x^a}$ , namely  $x^a$ , does not depend on the integration variable  $t$ , while in (1) the smoothness parameter in  $s_{x/t, t}$  is  $t$ , the integration variable itself.

To circumvent this issue, we modify  $S_x$ . We divide the primes in  $[2, x]$  into finitely many disjoint intervals  $(I_k)_k$ , and if  $n \leq x$  has  $P(n) \in I_k$ , we 'keep' this  $n$  in the modified version of  $S_x$  only if  $P(n/P(n))$  (the second largest prime factor of  $n$ ) is smaller than  $\min I_k$ . In this way, the new  $S_x$  is

$$S'_x = \sum_{p \leq x} Z'_{p,x}$$

where, if  $p \in I_k$ , then

$$Z'_{p,x} := \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1/2} \sum_{n \leq x, P(n)=p, P(n/P(n)) < \min I_k} \alpha(n) f(n)$$

and the new bracket process takes the shape

$$BP'_x \approx x \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1} \sum_k \sum_{p \in I_k} \frac{|f(p)|^2}{p} |s_{x/p, \min I_k}|^2 \approx x \left( \sum_{n \leq x} |f(n)|^2 \right)^{-1} \sum_k \int_2^x \frac{|s_{x/t, \min I_k}|^2 dt}{t \log t}.$$

For fixed  $k$ , the smoothness parameter is now fixed within the integral, namely it is  $\min I_k$ . This allows us to handle the  $k$ th integral using (5). One has to justify working with this modified  $S_x$ . The idea is that the second moment of the discarded terms is

$$\sum_k \frac{\sum_{n \leq x: P(n), P(n/P(n)) \in I_k} |f(n)|^2}{\sum_{n \leq x} |f(n)|^2} \ll \sum_k \sum_{p, q \in I_k} \frac{1}{pq} \ll \sum_k (\log(\log \max I_k / \log \min I_k))^2.$$

We discard also  $P(n) \leq x^\varepsilon$  (we lose  $O(\varepsilon)$ ) and then take  $I_k = [x^{\varepsilon+\delta k}, x^{\varepsilon+\delta(k+1)}]$ , so that the summand is  $\ll \log^2((\varepsilon + \delta(k+1))/(\varepsilon + \delta k)) \ll \min\{\delta/\varepsilon, 1/k\}^2$  and this is good enough; the total loss is  $O_\varepsilon(\delta) + O(\varepsilon)$  and this is manageable if we take  $\delta$  to 0 and then  $\varepsilon$  to 0.