Midterm Fall 2015

Please answer all questions. Show your work.

The exam is open book/open note; closed any devices that can communicate. (No laptops, cell phones, Morse code keys.)

1. Suppose $x \sim U(0,1)$. Find the covariance between x and x^2 .

Answer:

First we need the expected values. E(x)=.5, obviously. $E(x^2)=\int_0^1 x^2\times 1dx=\frac{1}{3}x^3|_0^1=\frac{1}{3}$. So the covariance is $E(x\times x^2)-E(x)\,E(x^2)$. The first term is $E(x\times x^2)=\int_0^1 x^3\times 1dx=\frac{1}{4}x^4|_0^1=\frac{1}{4}$. So the answer is $\frac{1}{4}-\frac{1}{2}\times\frac{1}{3}=\frac{1}{12}$.

2. There were two kinds of Sneetches in the world, the Star-Belly Sneetches had bellies with stars. The Plain-Belly Sneetches had none upon thars. 90 percent of Sneetches had a star. The Star-Belly Sneetches believed that when they saw a bad thing, two-thirds of the time it was due to a Plain-Belly Sneetch. In other words, people believed that 2/3rd of bad Sneetches were Plain-Bellies. (The Star-Belly Sneetches were wrong about this, but for the purpose of the problem pretend they were right.) Among all Sneetches, only 1 percent were really bad.

If a Star-Belly comes upon a Plain-Belly, what is the probability that the Plain-Belly is bad?

Answer:

Bayes law tells us that

$$p(bad|PlainBelly) = p(PlainBelly|bad) \times \frac{P(bad)}{P(PlainBelly)} = \frac{2}{3} \times \frac{.01}{.1} = \frac{2}{30} = \frac{1}{15}$$

3. The probability of dying is distributed exponentially with expected number of years $1/\lambda$. Consider an annuity that pays out continuously a rate p per year and stops payment at death. If the continuously compounded interest rate is r, (so a dollar at time t is worth e^{-rt} dollars now, then the present value of payments through year τ is

$$\frac{p}{r}[1-e^{-r\tau}]$$

What is the expected net present value of the annuity?

Answer:

The pdf of the exponential is $f(t) = \lambda \exp(-\lambda t)$. So the expected present value of payments is

$$A = \int_{0}^{\infty} \left(\frac{p}{r} [1 - e^{-r\tau}] \right) \lambda \exp(-\lambda t) dt = \frac{p\lambda}{r} \left[\int_{0}^{\infty} e^{-\lambda t} dt - \int_{0}^{\infty} e^{-(r+\lambda)t} dt \right]$$

With appropriate constants, both integrals are integrals of an exponential pdf equaling 1.

$$A = \frac{p\lambda}{r} \left[\frac{1}{\lambda} \int_{0}^{\infty} \lambda e^{-\lambda t} dt - \frac{1}{r+\lambda} \int_{0}^{\infty} (r+\lambda) e^{-(r+\lambda)t} dt \right] = \frac{p\lambda}{r} \left[\frac{1}{\lambda} - \frac{1}{r+\lambda} \right] = \frac{p}{r+\lambda}$$

4. Suppose that x, ε , and v are jointly normally distributed

$$\begin{bmatrix} x \\ \varepsilon \\ v \end{bmatrix} \sim N \begin{pmatrix} \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma_x^2 & 0 & 0 \\ 0 & \sigma_\varepsilon^2 & 0 \\ 0 & 0 & \sigma_v^2 \end{bmatrix}$$

Further, $y = \beta x + \varepsilon$ and z = x + v.

Find

$$\frac{\operatorname{cov}(y,z)}{\operatorname{var}(z)}$$

Answer:

$$var(z) = \sigma_x^2 + \sigma_v^2 + 2 \times 0$$

$$cov(y, z) = cov(\beta x + \varepsilon, x + v) = \beta \sigma_x^2 + \beta \sigma_{x\varepsilon} + \sigma_{\varepsilon x} + \sigma_{\varepsilon v} = \beta \sigma_x^2 + 0 + 0 + 0$$

$$\frac{cov(y, z)}{var(z)} = \beta \frac{\sigma_x^2}{\sigma_x^2 + \sigma_v^2}$$

5. Consider a simulation that produces a yes/no answer where the probability of "yes" is p. The total number of independent Monte Carlo trials is n. If we observe k yeses, we estimate

$$\hat{p} = \frac{k}{n}$$

- (a) Find mean, μ , and the variance, V, of \hat{p} in terms of p, k, and n.
- (b) In a large number of trials, \hat{p} is approximately normally distributed. Taking $\hat{p} \sim N(\mu, V)$, then it can be shown $P(|\hat{p} \mu| > 1.96\sqrt{V}) = .05$. If we think p = .1, how many observations do we need do that the probability \hat{p} is off by 0.01 is five percent?

Answer:

The expected value of $\hat{p} = p$ since we are looking at the sum of n trials each with expectation p and then we are dividing by n. The variance of a single draw is p(1-p). We

have n draws so the variance of the sum is np(1-p). We divide by n which divides the variance by n^2 . So V=p(1-p)/n.

We have

$$P\big(|\hat{p}-p|>1.96\sqrt{p(1-p)/n}\big)=.05$$
 So we want $1.96\sqrt{p(1-p)/n}=0.01$. For $p\approx .9$ that's $1.96\sqrt{.9\times.1/n}=0.01$.
$$n=\left(\frac{1.96\sqrt{.9\times.1}}{.01}\right)^2=n=1.96\times 900\approx 3,457$$