

# Conditional Expectation Functions

## Econometrics II

Douglas G. Steigerwald

UC Santa Barbara

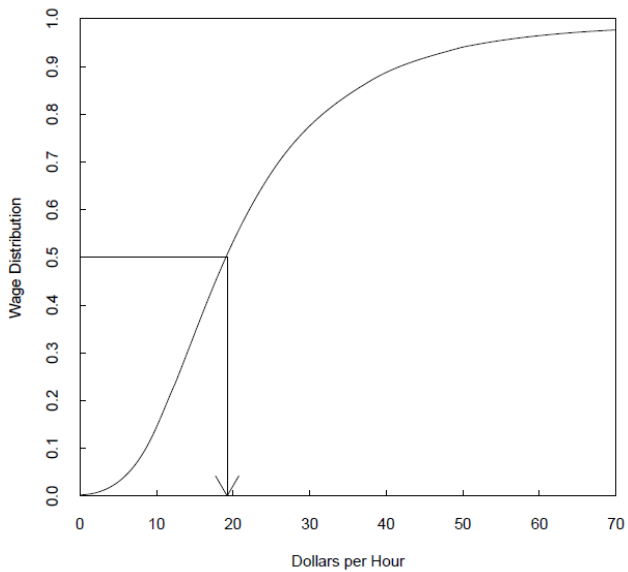
# Overview

Reference: B. Hansen Econometrics Chapters 1 and 2.0 - 2.8

- most commonly applied econometrics tool
  - ▶ least-squares estimation (regression)
- tool to estimate
  - ▶ approximate conditional mean of dependent variable
  - ▶ as a function of covariates (regressors)
  - ▶  $(y, x_1, \dots, x_K) := (y, x^T)$
- data is observational *not* experimental
  - ▶ causality is difficult to infer
- example - wages
  - ▶ random variable before measurement
  - ▶ observed wages are outcomes of the random variable
  - ▶ underpins the application of statistics to economics

# Distribution of Wages

- probability distribution
  - ▶  $F(u) = \mathbb{P}(\text{wage} \leq u)$
- median - measure of location (central tendency)
  - ▶ If  $F$  is continuous,  $m$  uniquely solves  $F(m) = \frac{1}{2}$
  - ▶ Otherwise,  $m = \inf \left\{ u : F(u) \geq \frac{1}{2} \right\}$
  - ▶ not a linear operator, some calculations are tricky
  - ▶ robust to tail perturbations
- nonparametric distribution estimate (following slide)
  - ▶ 50,742 full-time non-military wage earners March 2009 CPS
  - ▶  $\hat{m} = \$19.23$



# Quantiles

a useful way to summarize a probability distribution

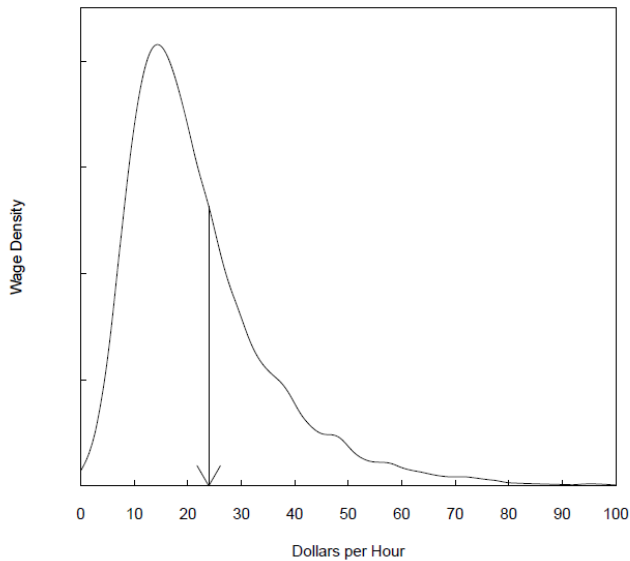
- for any  $\alpha \in (0, 1)$ , the  $\alpha^{th}$  quantile is
  - ▶ If  $F$  is continuous,  $q_\alpha$  uniquely solves  $F(q_\alpha) = \alpha$
  - ▶ Otherwise,  $q_\alpha = \inf \{u : F(u) \geq \alpha\}$
  - ▶  $q_{0.5} = m$
- quantile function,  $q_\alpha$ , viewed as a function of  $\alpha$  is the inverse of  $F$
- if  $\alpha$  is represented in percentage terms (10% instead of .1), quantiles are referred to as percentiles
  - ▶  $q_{0.5} = m$  is called the 50th percentile
  - ▶  $q_{0.9}$  is called the 90th percentile

# Density of Wages

If  $F$  is differentiable, density function exists

$$f(u) = \frac{d}{du} F(u)$$

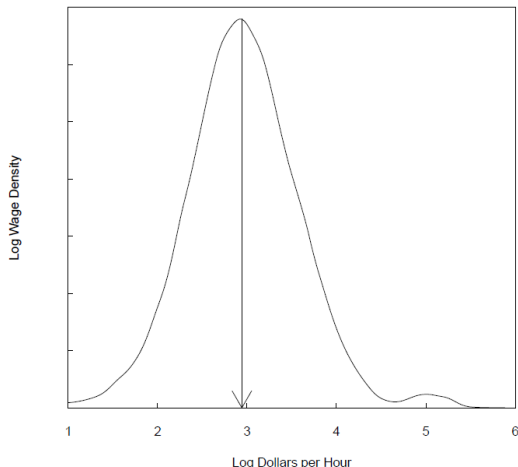
- $F(u)$  and  $f(u)$  contain the same information
- density is easier to interpret visually
- mean - measure of location
  - ▶ if  $F$  is continuous,  $\mu := \mathbb{E}(u) = \int_{-\infty}^{\infty} uf(u) du$
  - ▶ formal definition, 240A Lecture on Random Variables and Distributions
  - ▶ linear operator, not robust
- nonparametric density estimate for wages (following slide)
- $\hat{\mu} = \$23.90$
- data are skew, 64% of workers earn less than  $\hat{\mu}$



## Density of Log Wage

gains can be made by taking the natural logarithm of wages

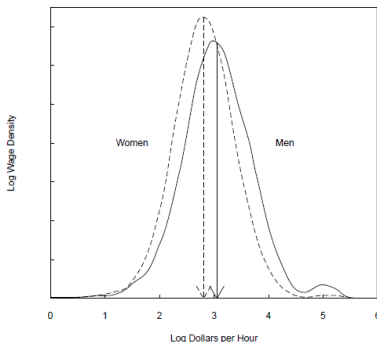
- skewness and thick tails can be reduced
- $\hat{\mu}$  for log wage is a much better measure of central tendency





# Conditional Expectations

*Is the wage distribution the same for all workers?*



- Men versus women (43% of workers are women)
  - ▶  $\mathbb{E}(\log(\text{wage}) \mid \text{gender} = \text{man}) = 3.05$
  - ▶  $\mathbb{E}(\log(\text{wage}) \mid \text{gender} = \text{woman}) = 2.81$
  - ▶ 24% difference in average wages between men and women

# Toolkit: Log Differences

If  $y^*$  is  $c\%$  greater than  $y$

$$\begin{aligned}y^* &= \left(1 + \frac{c}{100}\right) y \\ \log y^* - \log y &= \log \left(1 + \frac{c}{100}\right) \approx \frac{c}{100}\end{aligned}$$

*key logic*  $\log(1 + x) \approx x$

- example:  $100 * (\log w - \log z) = c$ 
  - ▶ then  $w$  is approximately  $c\%$  larger than  $z$
  - ▶ approximation is quite good for  $|c| \leq 10$

Approximation Accuracy

# Gender and Race

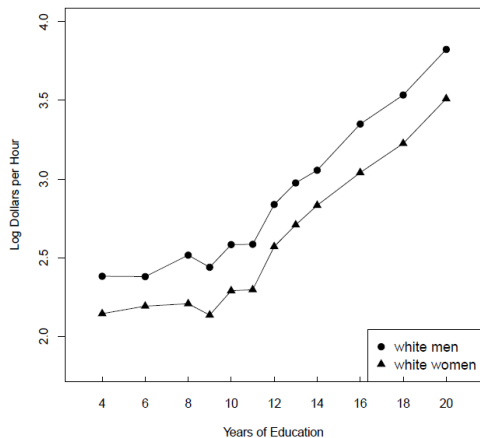
- conditional means reduce distributions to a single summary measure
  - ▶ primary focus of regression analysis
  - ▶ major focus of econometrics

## Conditional Means

	White	Black	Other
Men	3.07	2.86	3.03
Women	2.82	2.73	2.86

- male-female wage gap
  - ▶ 25% for whites 13% for blacks
- black-white wage gap
  - ▶ 21% for men 9% for women

# Education



- after 9 years, conditional mean increases at a different rate
- male-female gap is constant across education levels
  - ▶ constant percentage difference in wages

# Conditional Expectation Function

## Discrete Conditioning Variables

### CEF

$$\mathbb{E}(\log(\text{wage}) | \text{gender}, \text{race}, \text{education})$$

simplify notation

$$\mathbb{E}(y | x_1, x_2, \dots, x_k) = m(x_1, x_2, \dots, x_k)$$

for  $x = (x_1, x_2, \dots, x_k)^T$

$$\mathbb{E}(y | x) = m(x)$$

- CEF  $\mathbb{E}(y | x)$  is a random variable because it is a function of  $x$
- given  $x$ , it is not random

$$\mathbb{E}(\log(\text{wage}) | \text{gender} = \text{man}, \text{race} = \text{white}, \text{education} = 12) = 2.84$$

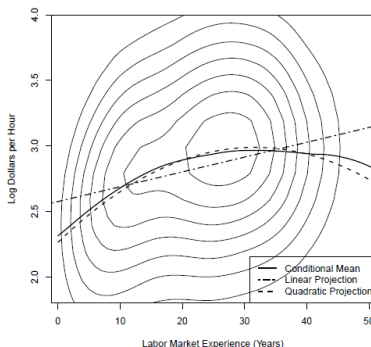
# Conditional Expectation Function

## Continuous Variables with Joint Density Function

$f(y, x)$  is the joint density function for  $(y, x)$

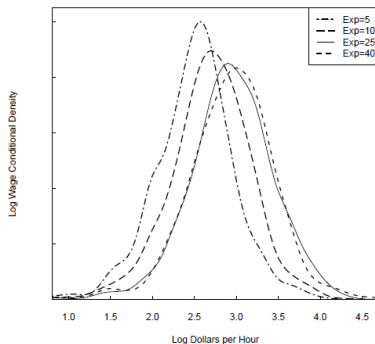
- $y = \log(\text{wage})$      $x = \text{experience}$

for white men with 12 years of education contours of  $f(y, x)$  are



# Conditional Density

a "slice" of the joint density contours yields the conditional density

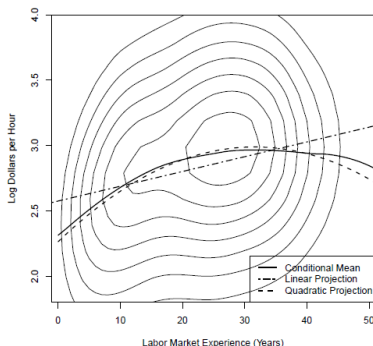


- shifts right and becomes more diffuse as experience increases
  - ▶ little change as experience increases past 25 years
- the conditional density is denoted  $f_{y|x}(y|x)$  **Conditional Density**

# Conditional Expectation Function

$$m(x) := \mathbb{E}(y|x) = \int_{\mathbb{R}} y f_{y|x}(y|x) dy$$

- mean of idealized subpopulation with value  $x$ 
  - ▶  $x$  continuous implies this subpopulation is infinitely small



- conditional mean (CEF) is nonlinear



# Error

define: CEF error  $e = y - m(x)$

$$y = m(x) + e$$
$$\mathbb{E}(e|x) = 0$$

note,  $\mathbb{E}(e|x) = 0$  is not a restriction, these equations hold by definition

Error Properties Theorem: (derived from  $f(y, x)$ )

- ①  $\mathbb{E}(e|x) = 0 \Rightarrow \mathbb{E}(e) = 0$
- ②  $\mathbb{E}(h(x)e) = 0$  if  $\mathbb{E}|h(x)e| < \infty$
- ③  $\mathbb{E}|y|^r < \infty \Rightarrow \mathbb{E}|e|^r < \infty$  ( $r \geq 1$ )

# Error Properties

1.  $\mathbb{E}(e|x) = 0$

- *not* a restriction, but a definition
- called mean independence
  - ▶ mean independence  $\nRightarrow$  independence
  - ▶  $e = x\epsilon$  with  $\epsilon \sim \mathcal{N}(0, 1)$  independent of  $x \Rightarrow e|x \sim \mathcal{N}(0, x^2)$
  - ▶ empirics :  $e$  and  $x$  are rarely assumed independent

2.  $\mathbb{E}(h(x)e) = 0$

- $e$  is uncorrelated with any function of the covariates

3.  $\mathbb{E}|y|^r < \infty \Rightarrow \mathbb{E}|e|^r < \infty$

- $\mathbb{E}y^2 < \infty \Rightarrow \text{Var}(e) < \infty$

# Toolkit: Law of Iterated Expectations

To show property 1

## Simple Law

$$\mathbb{E}(\mathbb{E}(y|x)) = \mathbb{E}(y) \quad \text{if } \mathbb{E}(y) < \infty$$

note  $\mathbb{E}(\mathbb{E}(y|x)) = \int_{\mathbb{R}} \mathbb{E}(y|x) f_x(x) dx$

General Law (allows for 2 sets of conditioning variables)

$$\mathbb{E}(\mathbb{E}(y|x_1, x_2) | x_1) = \mathbb{E}(y|x_1) \quad \text{if } \mathbb{E}(y) < \infty$$

- "smaller information set wins"

# Toolkit: Conditioning Theorem

To show property 2

condition on  $x \rightarrow$  effectively treat  $x$  as constant

Simple Property

$$\mathbb{E}(g(x)|x) = g(x) \quad \text{for any function } g(\cdot)$$

example  $\mathbb{E}(x|x) = x$

General Property (allows for an additional random variable)

$$\mathbb{E}(g(x)y|x) = g(x)\mathbb{E}(y|x)$$

$$\mathbb{E}(g(x)y) = \mathbb{E}(g(x)\mathbb{E}(y|x))$$

# Proofs

- 1 Proof of Simple LIE
- 2 Proof of General LIE
- 3 Proof of Conditioning Theorems
- 4 Proof of Error Properties Theorem

# Review

- Implication of observational data?
- causality is difficult to infer

Should we model  $\mathbb{E}(y|x)$  as linear in  $x$ ?

- no

What are the key properties of  $e = y - \mathbb{E}(y|x)$ ?

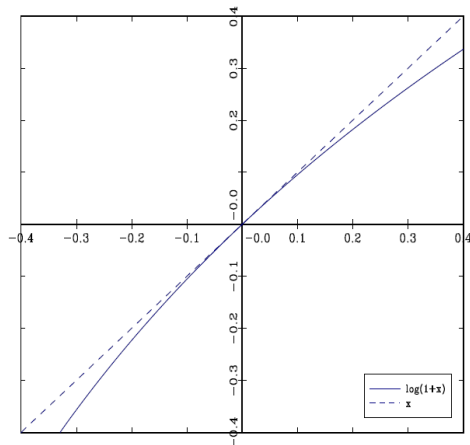
- $\mathbb{E}(e|x) = 0$  (by construction)
- uncorrelated with any function of  $x$

# Approximation Accuracy

Taylor Series expansion

$$\log(1+x) = x - \frac{1}{2}x^2 + \frac{1}{3}x^3 - \dots = x + O(x^2) \approx x$$

highly accurate if  $|x| \leq .1$



# Landau Notation (Big O)

we say  $f(x) = O(g(x))$  as  $x \rightarrow 0$  if

$$|f(x)| \leq M |g(x)| \text{ for all } x \leq x_0$$

Consider  $f(x) = -\frac{1}{2}x^2 + \frac{1}{3}x^3$

$$\left| -\frac{1}{2}x^2 + \frac{1}{3}x^3 \right| \leq \frac{1}{2}|x^2| + \frac{1}{3}|x^3|$$

for suitably chosen  $x_0$ , if  $x \leq x_0$

$$\frac{1}{2}|x^2| + \frac{1}{3}|x^3| \leq \frac{1}{2}|x^2| + \frac{1}{3}|x^2| = \frac{5}{6}x^2$$

so  $f(x) = O(x^2)$  as  $x \rightarrow 0$

Return to Log Wage



## Definition of Conditional Density

- if  $(y, x)$  have joint density  $f(y, x)$  then
  - ▶  $x$  has marginal density

$$f_x(x) = \int_{\mathbb{R}} f(y, x) dy$$

- for any  $x$  such that  $f_x(x) > 0$ , the conditional density of  $y$  given  $x$  is defined as

$$f_{y|x}(y|x) = \frac{f(y, x)}{f_x(x)}$$

- consider  $f(\log(\text{wage})|\text{experience} = 5)$

$$\frac{f(y, x = 5)}{\mathbb{P}(x = 5)} \quad \begin{array}{l} \leftarrow \text{the "slice"} \\ \leftarrow \text{the scale factor} \end{array}$$

- ▶ if there are fewer individuals with 5 years of experience than with 10 years of experience, the higher conditional density could correspond to workers with 5 years of experience, even if the joint density is higher for workers with 10 years of experience

## Return to Conditional Density

# Proof of Simple Law of Iterated Expectations

*Simple LIE:*  $\mathbb{E}(\mathbb{E}(y|x)) = \mathbb{E}(y)$

- assume  $(y, x)$  have joint density  $f(y, x)$  (for convenience)
  - ▶  $\mathbb{E}(y|x)$  is a function of the random variable  $x$  only
  - ▶ to calculate its expectation, integrate with respect to the density  $f_x(x)$  of  $x$

$$\mathbb{E}(\mathbb{E}(y|x)) = \int_{\mathbb{R}^k} \mathbb{E}(y|x) f_x(x) dx$$

- ▶ which equals (by substitution and by noting that  $f_{y|x}(y|x) f_x(x) = f(y, x)$ )

$$\begin{aligned} \int_{\mathbb{R}^k} \left( \int_{\mathbb{R}} y f_{y|x}(y|x) dy \right) f_x(x) dx &= \int_{\mathbb{R}^k} \int_{\mathbb{R}} y f_{y,x}(y, x) dy dx \\ &= \mathbb{E}(y), \end{aligned}$$

because  $\int_{\mathbb{R}^k} f_{y,x}(y, x) dx = f_y(y)$ . ■

Return to Proofs

# Proof of General Law of Iterated Expectations

*General LIE:*  $\mathbb{E}(\mathbb{E}(y|x_1, x_2) | x_1) = \mathbb{E}(y|x_1)$

- assume  $(y, x_1, x_2)$  have joint density  $f(y, x_1, x_2)$  (for convenience)
  - ▶  $\mathbb{E}(y|x_1, x_2)$  is a function of the random variables  $x_1$  and  $x_2$
  - ▶ integrate with respect to the density of  $x_2$  given  $x_1$

$$\begin{aligned}\mathbb{E}(\mathbb{E}(y|x_1, x_2) | x_1) &= \int_{\mathbb{R}^{k_2}} \mathbb{E}(y|x_1, x_2) f(x_2|x_1) dx_2 \\ &= \int_{\mathbb{R}^{k_2}} \left( \int_{\mathbb{R}} y f(y|x_1, x_2) dy \right) f(x_2|x_1) dx_2\end{aligned}$$

- ▶ note that  $f(y|x_1, x_2) f(x_2|x_1) = \frac{f(y, x_1, x_2)}{f(x_1, x_2)} \frac{f(x_1, x_2)}{f(x_1)} = f(y, x_2|x_1)$ , so

$$\begin{aligned}&= \int_{\mathbb{R}^{k_2}} \int_{\mathbb{R}} y f(y, x_2|x_1) dy dx_2 \\ &= \mathbb{E}(y|x_1),\end{aligned}$$

the mean of  $y$  given the value of  $x_1$ . ■

Return to Proofs

# Proof of Conditioning Theorems

*General CT 1:*  $\mathbb{E}(g(x)y|x) = g(x)\mathbb{E}(y|x)$

- assume  $(y, x_1, x_2)$  have joint density  $f(y, x_1, x_2)$  (for convenience)

$$\begin{aligned}\mathbb{E}(g(x)y|x) &= \int_{\mathbb{R}} g(x)y f_{y|x}(y|x) dy \\ &= g(x) \int_{\mathbb{R}} y f_{y|x}(y|x) dy \\ &= g(x) \mathbb{E}(y|x),\end{aligned}$$

where  $\mathbb{E}|g(x)y| < \infty$  is needed to ensure the first equality is well defined. ■

*General CT 2:*  $\mathbb{E}(g(x)y) = \mathbb{E}(g(x)\mathbb{E}(y|x))$

- Proof: application of simple LIE. ■

Return to Proofs

# Proof of Error Properties Theorem

Parts 1 and 2 follow trivially

$$\text{Part 3: } \mathbb{E} |y|^r < \infty \Rightarrow \mathbb{E} |e|^r < \infty \quad (r \geq 1)$$

- $e = y - m(x)$
- $(\mathbb{E} |e|^r)^{1/r} = (\mathbb{E} |y - m(x)|^r)^{1/r}$ 
  - ▶  $(\mathbb{E} |y - m(x)|^r)^{1/r} \leq (\mathbb{E} |y|^r)^{1/r} + (\mathbb{E} |m(x)|^r)^{1/r}$  (Minkowski's Inequality - generalizes Triangle Inequality)
    - ★  $\mathbb{E} |\mathbb{E}(y|x)|^r \leq \mathbb{E} |y|^r$  for any  $r \geq 1$  (Conditional Expectation Inequality)
  - ▶ the two parts on the right are finite by  $\mathbb{E} |y|^r < \infty$
- $(\mathbb{E} |e|^r)^{1/r} < \infty$  implies  $\mathbb{E} |e|^r < \infty$ . ■

## Background : Triangle Inequality

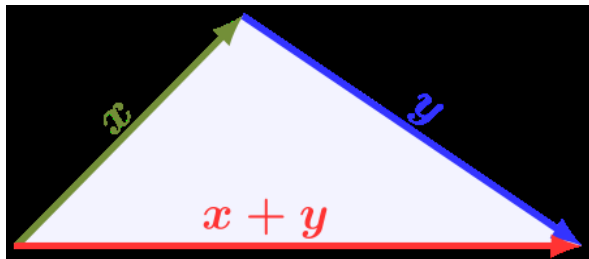
Let  $x$  and  $y$  be real numbers. The triangle inequality is

$$|x + y| \leq |x| + |y|.$$

- Proof

- ▶  $x + y \leq |x| + |y|$
- ▶  $-(x + y) = (-x) + (-y) \leq |x| + |y|$
- ▶  $|x + y| \leq \max\{-(x + y), x + y\}$  ■

To understand why it is called the triangle inequality, let  $x$  and  $y$  be vectors



# Background : Triangle Inequality for a Random Variable

- Let  $x$  be a random variable
  - ▶ for convenience,  $x$  is a discrete random variable
  - ▶ set of possible values  $\{x_1, x_2, \dots, x_n\}$
  - ▶ with probabilities  $\{p_1, p_2, \dots, p_n\}$

The triangle inequality is

$$|\mathbb{E}x| \leq \mathbb{E}|x|.$$

- Proof

- ▶  $|\sum_{i=1}^n x_i p_i| \leq \sum_{i=1}^n |x_i p_i| = \sum_{i=1}^n |x_i| p_i$  ■

Return to Proofs