

# Solutions Midterm - Econ 241A Probability, Statistics and Econometrics

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1. Let  $Y = \exp(X)$ , where  $X \sim N(\mu, \sigma^2)$ . Show that

$$\frac{\sigma_Y}{E[Y]} = \sqrt{e^{\sigma^2} - 1}$$

where  $\sigma_Y$  is the standard deviation of  $Y$ .

Answer.

Given that  $Y \sim \ln N(\mu, \sigma^2)$  then

$$\text{Var}(Y) = \sigma_Y^2 = (e^{\sigma^2} - 1)e^{2\mu + \sigma^2}$$

Given that  $E[Y]^2 = e^{2\mu + \sigma^2}$

$$\frac{\sigma_Y}{E[Y]} = \sqrt{e^{\sigma^2} - 1}$$

2. Show that  $E[\epsilon|X] = 0$  implies that  $\text{Cov}(\epsilon, X) = 0$ .

Answer.

$$E[\epsilon X] = E[E[\epsilon X|X]] = E[XE[\epsilon|X]] = E[X \cdot 0] = 0$$

$$E[\epsilon] = E[E[\epsilon|X]] = E[0] = 0$$

$$\text{Cov}(\epsilon, X) = E[\epsilon X] - E[\epsilon]E[X] = 0 - 0 = 0$$

3. Assume  $w$  is a random variable and  $u(w)$  has a convergent Taylor expansion around  $E[w] = \mu_w$ , i.e.

$$u(w) = u(\mu_w) + u'(\mu_w)(w - \mu_w) + u''(\mu_w)(w - \mu_w)^2 + \sum_{n=3}^{\infty} \frac{1}{n!} u^{(n)}(\mu_w)(w - \mu_w)^n$$

- (a) Give an exact expression for  $E[u(w)]$  (state or not if any more assumptions are required).

Answer.

$$E[u(w)] = u(\mu_w) + u''(\mu_w)E[(w - \mu_w)^2] + \sum_{n=3}^{\infty} \frac{1}{n!} u^{(n)}(\mu_w)E[(w - \mu_w)^n]$$

It is required that expectations goes through the limit of the Taylor expansion.

- (b) Assume that  $w \sim N(0, \sigma_w^2)$ . Give an even more detailed expression for  $E[u(w)]$ .

Answer.

$$E[u(w)] = u(0) + u''(0)\sigma^2 + \frac{3}{4!}u^{(4)}(0)\sigma^4 + \frac{3 \cdot 5}{6!}u^{(6)}(0)\sigma^6 + \frac{3 \cdot 5 \cdot 7}{8!}u^{(8)}(0)\sigma^8 + \dots$$

- (c) Assume that  $u(w) = \exp(w)$  and  $w \sim N(\mu_w, \sigma_w^2)$ . Give an *alternative* expression for  $E[u(w)]$ .

Answer.

$$E[u(w)] = E[e^w] = e^{\mu + 0.5\sigma^2}$$

- (d) Let  $w_0 \sim N(\mu, \sigma_0^2)$  and  $w_1 \sim N(\mu, \sigma_1^2)$  maintaining  $u(w) = \exp(w)$ . Establish a condition over  $\sigma_0^2$  and  $\sigma_1^2$  such as  $E[u(w_0)] \leq E[u(w_1)]$ .

Answer.

Based on (c) it is sufficient and necessary that  $\sigma_0^2 < \sigma_1^2$

4. Let  $X$  denote the math score on the ACT college entrance exam of a randomly selected student. Let  $Y$  denote the verbal score on the ACT college entrance exam of a randomly selected student. If  $X$  and  $Y$  are distributed jointly normal such as  $X \sim N(\mu_X, \sigma_X^2)$ ,  $Y \sim N(\mu_Y, \sigma_Y^2)$  and  $\text{corr}(X, Y) = \rho$ . State the following in terms of the given parameters  $(\mu_X, \sigma_X^2, \mu_Y, \sigma_Y^2, \rho)$  and the standard normal cdf  $\Phi(z)$ .

- (a) What is the probability that a randomly selected student's verbal ACT score is between 10 and 20 points?

Answer.

$$P(10 \leq Y \leq 20) = P\left(\frac{10 - \mu_Y}{\sigma_Y} \leq \frac{Y - \mu_Y}{\sigma_Y} \leq \frac{20 - \mu_Y}{\sigma_Y}\right) = \Phi\left(\frac{20 - \mu_Y}{\sigma_Y}\right) - \Phi\left(\frac{10 - \mu_Y}{\sigma_Y}\right)$$

- (b) What is the probability that a randomly selected student's verbal ACT score is between 10 and 20 points given that  $X = 20$ ?

Answer.

$$P(10 \leq Y \leq 20|X = 20) = \Phi\left(\frac{20 - \mu_{Y|X}}{\sigma_{Y|X}}\right) - \Phi\left(\frac{10 - \mu_{Y|X}}{\sigma_{Y|X}}\right)$$

$$\text{Where } \mu_{Y|X=20} = \mu_Y + \frac{\text{cov}(X,Y)}{\sigma_X^2}(20 - \mu_X) \text{ and } \sigma_{Y|X} = \sigma_Y \sqrt{1 - \rho^2}.$$

5. The Gini coefficient is commonly used to measure inequality of income. If income is represented by  $Y$ , a continuous random variable with cdf  $F(y)$  with mean  $E[Y] = \mu$ , then the Gini coefficient is given by

$$G = \frac{1}{\mu} \int_0^\infty F(y)(1 - F(y))dy$$

- (a) Assume that  $Y \sim U[a, b]$  (clearly  $a > 0$ ). Compute  $G$ .

Answer.

We know that

$$F(y) = \frac{y - a}{b - a}$$

$$\mu = \frac{b + a}{2}$$

So using the formula for the Gini coefficient

$$G = \frac{2}{a + b} \int_a^b \frac{y - a}{b - a} \left(1 - \frac{y - a}{b - a}\right) dy$$

If we make a change in variables such as  $z = \frac{y-a}{b-a}$  and  $dy = (b - a)dz$ , then

$$G = \frac{2(b-a)}{a+b} \int_0^1 z(1-z)dz$$

$$= \frac{2(b-a)}{a+b} \left[ \frac{z^2}{2} - \frac{z^3}{3} \right]_0^1$$

$$= \frac{2(b-a)}{a+b} \left[ \frac{1}{2} - \frac{1}{3} \right]$$

$$G = \frac{b-a}{3(a+b)}$$

- (b) Assume that  $Y \sim \exp(\lambda)$  (i.e.  $f_Y(y) = \lambda \exp(-\lambda y)$ ,  $y > 0$ ). Compute  $G$ .

Answer.

We know that

$$F(y) = 1 - e^{-\lambda y}$$

$$\mu = \frac{1}{\lambda}$$

So using the formula for the Gini coefficient

$$G = \lambda \int_0^\infty (1 - e^{-\lambda y})(e^{-\lambda y})dy$$

$$G = \lambda \int_0^\infty (e^{-\lambda y} - e^{-2\lambda y})dy$$

Given that  $f_Y(y) = \lambda e^{-\lambda y}$  is a pdf and integrates to 1, we have

$$G = \lambda \left[ \frac{1}{\lambda} - \frac{1}{2\lambda} \right]$$

$$G = \frac{1}{2}$$

6. A household has preferences represented by

$$u(w) = -\frac{1}{2}(w - a)^2$$

where  $w$  is random variable which represents wealth and we assume that  $a$  is high enough such as  $0 \leq w < a$ . The household maximizes expected utility  $E[u(w)]$ .

- (a) Show that maximizing expected utility is equivalent to maximizing

$$aE[w] - \frac{1}{2}E[w]^2 - \frac{1}{2}Var(w)$$

Answer.

$$\begin{aligned} E[u(w)] &= E\left[-\frac{1}{2}(w-a)^2\right] \\ E[u(w)] &= -\frac{1}{2}(E[w]^2 - 2aE[w] + a^2) \\ E[u(w)] &= -\frac{1}{2}E[w]^2 - \frac{1}{2}Var(w) + aE[w] - \frac{1}{2}a^2 \end{aligned}$$

Adding  $\frac{1}{2}a$  (an affine transformation) provides the answer.

- (b) The household has an endowment  $w_0 > 0$ , and decides to invest nonnegative amounts  $\phi$  in a risky asset and  $\phi_f$  in a risk-free asset such as  $w_0 = \phi_f + \phi$  (before knowing the risky asset's return). The household consumes wealth  $w$  after the return of the risky asset  $R$  is determined, i.e.  $w = R_f\phi_f + R\phi$ , where  $R_f$  is the return to the risk-free asset. What is the demand for the risky asset  $\phi_f$  that maximizes expected utility?

$$\max_{(\phi, \phi_f) \geq 0} \quad aE[w] - \frac{1}{2}E[w]^2 - \frac{1}{2}Var(w) \quad (1)$$

$$\text{s. t.} \quad w_0 = \phi_f + \phi \quad (2)$$

$$w = R_f\phi_f + R\phi \quad (3)$$

Answer.

Solving for  $\phi_f$  in Equation 2 and plugging into Equation 3, we get

$$w = (R - R_f)\phi + R_fw_0$$

Taking expected value and variance of  $w$

$$\begin{aligned} E[w] &= (\mu - R_f)\phi + R_fw_0 \\ Var[w] &= \phi^2\sigma^2 \end{aligned}$$

Where  $\mu = E[R]$  and  $\sigma = Var(R)$ . Plugging back into the utility function (Equation 1)

$$\max_{\phi \geq 0} \quad a[(\mu - R_f)\phi + R_fw_0] - \frac{1}{2}[(\mu - R_f)\phi + R_fw_0]^2 - \frac{1}{2}\phi^2\sigma^2$$

The FOC wrt to  $\phi$  is

$$\begin{aligned} a(\mu - R_f) - [(\mu - R_f)\phi + R_fw_0](\mu - R_f) - \phi\sigma^2 &= 0 \\ \phi &= \frac{(\mu - R_f)(a - R_fw_0)}{(\mu - R_f)^2 + \sigma^2} \end{aligned}$$

if  $\mu > R_f$  and  $\phi = 0$  otherwise.

- (c) What is the effect of an increase in the endowment on the final demand for the risky asset (conditional on positive demand of the risky asset  $\phi > 0$ )? Discuss.

Answer.

The effect is negative since  $\frac{\partial \phi}{\partial w_0} < 0$ . The risky asset is an inferior good. Assuming these preferences is not appealing.