

# Midterm 2012 Questions & Answers

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## Part I

1. Prove that  $P(A|B) = P(B|A) \frac{P(A)}{P(B)}$ .

Answer:	$P(A \cap B) = P(B \cap A)$	Commutative
	$P(B A) = \frac{P(B \cap A)}{P(A)}$	Defn Cond Prob
	$\therefore, P(B A) = \frac{P(A \cap B)}{P(A)}$	Substitution
	$\Rightarrow P(B A)P(A) = P(A \cap B)$	Algebra
	$\Rightarrow P(A \cap B) = P(B A)P(A)$	Reflective
	$P(A B) = \frac{P(A \cap B)}{P(B)}$	Defn Cond Prob
	$\therefore, P(A B) = P(B A) \frac{P(A)}{P(B)}$	Substitution
		Q.E.D.

2. Prove that if sets  $A$  and  $B$  in sample space  $S$  are mutually exclusive (disjoint),  $P(A) > 0$ , and  $P(B) > 0$ , then  $A$  and  $B$  cannot be independent.

Answer:	$A$ and $B$ are mutually exclusive	Given
	$\Rightarrow P(A \cap B) = 0$	Defn of ME
	$P(A) > 0$ and $P(B) > 0$	Given
	$\therefore, P(A)P(B) > 0$	Algebra
	Assume $A$ and $B$ are independent	Assumption
	$\Rightarrow P(A \cap B) = P(A)P(B)$	Defn of $\perp$
	Since $P(A)P(B) > 0$ , $P(A \cap B) > 0$	Reflective Prop
	But $P(A \cap B) = 0$	Contradiction
	$\therefore$ , assumption is false	Observation
	$\Rightarrow A$ and $B$ are not independent	Q.E.D.

3.  $X$  and  $Y$  are jointly distributed with pdf  $f(x, y) = 9x^2y^2$ , where  $0 < x < 1$  and  $0 < y < 1$ . Are  $X$  and  $Y$  independent? Justify your answer.

Answer:	Let $g(x) = \begin{cases} 3x^2 & \text{if } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$	
	Is $g(x)$ a pdf?	
	(a) $g(x) \geq 0 \forall x \in (0, 1)$	
	(b) $\int_{-\infty}^{\infty} g(x)dx = \int_0^1 3x^2dx = 1$	
	$\therefore, g(x) = g_X(x)$ is a pdf	Theorem 1.6.5
	Let $h(y) = \begin{cases} 3y^2 & \text{if } 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$	
	Is $h(y)$ a pdf?	
	(a) $h(y) \geq 0 \forall y \in (0, 1)$	
	(b) $\int_{-\infty}^{\infty} h(y)dy = \int_0^1 3y^2dy = 1$	

$$\begin{aligned}\therefore, h(y) &= h_Y(y) \text{ is a pdf} \\ \therefore, g(x)h(y) &= g_X(x)h_Y(y) = \begin{cases} 9x^2y^2 & \text{if } 0 < x < 1 \wedge 0 < y < 1 \\ 0 & \text{otherwise} \end{cases} \\ \therefore, f_{X,Y}(x,y) &= g_X(x)h_Y(y) \\ \Rightarrow X \text{ and } Y &\text{ are independent}\end{aligned}$$

Theorem 1.6.5

Definition 4.2.5

OR

$$\begin{aligned}\text{Let } g(x) &= \begin{cases} 9x^2 & \text{if } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases} \\ \text{Let } h(y) &= \begin{cases} y^2 & \text{if } 0 < y < 1 \\ 0 & \text{otherwise} \end{cases} \\ \therefore, g(x)h(y) &= \begin{cases} 9x^2y^2 & \text{if } 0 < x < 1 \wedge 0 < y < 1 \\ 0 & \text{otherwise} \end{cases} \\ \therefore, f_{X,Y}(x,y) &= g(x)h(y) \\ \Rightarrow X \text{ and } Y &\text{ are independent}\end{aligned}$$

Lemma 4.2.7

$$\text{This method also works for letting } \begin{matrix} g(x) \\ h(y) \end{matrix} = \begin{matrix} 3x^2 \\ 3y^2 \end{matrix} \text{ or } \begin{matrix} g(x) \\ h(y) \end{matrix} = \begin{matrix} x^2 \\ 9y^2 \end{matrix}.$$

**4.  $X$  and  $Y$  are independent and  $\mathbb{E}(Z|X, Y) = X + Y$ . Show that  $\text{cov}(X, Z) = \text{var}(X)$ .**

Answer:

$$\begin{aligned}\text{cov}(X, Z) &= \mathbb{E}(XZ) - \mathbb{E}(X)\mathbb{E}(Z) \\ \Rightarrow \text{cov}(X, Z) &= \mathbb{E}[\mathbb{E}(XZ|X, Y)] - \mathbb{E}(X)\mathbb{E}[\mathbb{E}(Z|X, Y)] \\ \Rightarrow \text{cov}(X, Z) &= \mathbb{E}[X\mathbb{E}(Z|X, Y)] - \mathbb{E}(X)\mathbb{E}[\mathbb{E}(Z|X, Y)] \\ \mathbb{E}(Z|X, Y) &= X + Y \\ \therefore, \text{cov}(X, Z) &= \mathbb{E}[X(X + Y)] - \mathbb{E}(X)\mathbb{E}(X + Y) \\ \Rightarrow \text{cov}(X, Z) &= \mathbb{E}(X^2 + XY) - \mathbb{E}(X)\mathbb{E}(X + Y) \\ \Rightarrow \text{cov}(X, Z) &= \mathbb{E}(X^2) + \mathbb{E}(XY) - \mathbb{E}(X)[\mathbb{E}(X) + \mathbb{E}(Y)] \\ \Rightarrow \text{cov}(X, Z) &= \mathbb{E}(X^2) + \mathbb{E}(XY) - \mathbb{E}(X)^2 - \mathbb{E}(X)\mathbb{E}(Y) \\ \Rightarrow \text{cov}(X, Z) &= [\mathbb{E}(X^2) - \mathbb{E}(X)^2] + [\mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y)] \\ \text{var}(X) &= \mathbb{E}(X^2) - \mathbb{E}(X)^2 \\ \text{cov}(X, Y) &= \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) \\ \therefore, \text{cov}(X, Z) &= \text{var}(X) + \text{cov}(X, Y) \\ X \text{ and } Y &\text{ are independent} \\ \Rightarrow \text{cov}(X, Y) &= 0 \\ \therefore, \text{cov}(X, Z) &= \text{var}(X)\end{aligned}$$

Defn of Cov

LIE

Prop of  $\mathbb{E}$

Given

Substitution

Algebra

Prop of  $\mathbb{E}$

Algebra

Commutative

Defn of Var

Defn of Cov

Substitution

Given

Theorem 4.5.5

Substitution

*Q.E.D.*

## Part II

5. The following table provides the joint *pmf* for the bivariate vector  $(X, Y)$ ,  $f_{X,Y}(x, y)$ . Empty cells correspond to  $(x, y)$  combinations outside the support of  $(X, Y)$ . Answer questions (a), (b), (c), (d), (e) and (f) using this table. Hint: You can get full credit without any lengthy calculations as long as you explain your reasoning carefully.

		X													
		1	2	3	4	5	6	7	8	9	10	11	12	13	14
Y	1	0	0	0	0	0	0	0	0	0	0	0	0	0	0
	2	0	.0075	.0075	.0075	.0075	.0075	0	0	.0075	.0075	.0075	.0075	.0075	0
	3	0	.0075	.02	.01	.01	.0075	0	0	.0075	.01	.01	.02	.0075	0
	4	0	.0075	.02	.02	.01	.0075	0	0	.0075	.01	.02	.02	.0075	0
	5	0	.0075	.01	.02	.01	.0075	0	0	.0075	.01	.02	.01	.0075	0
	6	0	.0075	.0075	.0075	.0075	.0075	0	0	.0075	.0075	.0075	.0075	.0075	0
	7	0	0	0	0	0	0	0	0	0	0	0	0	0	0
	8	0	0	0	0	0	0	0	0	0	0	0	0	0	0
	9	0	.0075	.0075	.0075	.0075	.0075	0	0	.0075	.0075	.0075	.0075	.0075	0
	10	0	.0075	.01	.02	.01	.0075	0	0	.0075	.01	.02	.01	.0075	0
	11	0	.0075	.01	.02	.02	.0075	0	0	.0075	.02	.02	.01	.0075	0
	12	0	.0075	.01	.01	.02	.0075	0	0	.0075	.02	.01	.01	.0075	0
	13	0	.0075	.0075	.0075	.0075	.0075	0	0	.0075	.0075	.0075	.0075	.0075	0
	14	0	0	0	0	0	0	0	0	0	0	0	0	0	0

- (a) Consider events  $A = 1 \leq X \leq 7$ ,  $B = 7 \leq Y \leq 14$ . Are  $A$  and  $B$  independent? Provide careful argument for your answers.

Answer:  $P(A \cap B) = P(1 \leq X \leq 7 \cap 7 \leq Y \leq 14) = \sum_{x=1}^7 \sum_{y=7}^{14} P(X = x \wedge Y = y) = .25$   
 $P(A) = P(1 \leq X \leq 7) = \sum_{x=1}^7 \sum_{y=1}^{14} P(X = x \wedge Y = y) = .5$   
 $P(B) = P(7 \leq Y \leq 14) = \sum_{y=7}^{14} \sum_{x=1}^{14} P(X = x \wedge Y = y) = .5$   
 $\therefore, P(A \cap B) = P(A)P(B)$   
 $\Rightarrow A$  and  $B$  are independent

- (b) Are random variables  $X$  and  $Y$  independent? Provide a careful argument for your answer.

Answer:  $P(X \cap Y) = P(X)P(Y)$   
 $\Leftrightarrow P(X = x \wedge Y = y | x, y \in [1, 14]) = P(X = x | x \in [1, 14])P(Y = y | y \in [1, 14])$   
 $\Leftrightarrow X$  and  $Y$  are independent  
Let  $x = 2$  and  $y = 6$ :  
 $P(X = 2 \wedge Y = 6) = .0075$   
 $P(X = 2) = .0075 * 10 = .075$   
 $P(Y = 6) = .0075 * 10 = .075$   
 $\therefore, P(X = 2)P(Y = 6) = .075 * .075 \approx .0056$   
 $\therefore, P(X = 2 \wedge Y = 6) \neq P(X = 2)P(Y = 6)$

$$\Rightarrow P(X \cap Y) \neq P(X)P(Y)$$

$$\Rightarrow X \text{ and } Y \text{ are NOT independent}$$

**(c) Are events  $A^c$  and  $B$  independent?**

Answer:  $A^c = X < 1 \cup X > 7$

$$P(A^c \cap B) = P[(X < 1 \cup X > 7) \cap 7 \leq Y \leq 14]$$

$$= P[(X < 1 \cap 7 \leq Y \leq 14) \cup (X > 7 \cap 7 \leq Y \leq 14)]$$

$$= P(X < 1 \cap 7 \leq Y \leq 14) + P(X > 7 \cap 7 \leq Y \leq 14)$$

$$= 0 + \sum_{x=8}^{14} \sum_{y=7}^{14} P(X = x \wedge Y = y)$$

$$= .25$$

$$P(A^c) = P(X < 1 \cup X > 7) = P(X < 1) + P(X > 7) = 0 + .5 = .5$$

From (a),  $P(B) = .5$

$$\therefore P(A^c \cap B) = P(A^c)P(B)$$

$$\Rightarrow A^c \text{ and } B \text{ are independent}$$

**OR**

$$A \perp B \Rightarrow A^c \perp B$$

Theorem 1.3.9

**(d) Prove the following theorem: Let  $X$  and  $Y$  be independent random variables. For any  $C \subset \mathbb{R}$  and  $D \subset \mathbb{R}$ ,  $P(X \in C, Y \in D) = P(X \in C)P(Y \in D)$ .**

Answer: Let  $X \in C$  and  $Y \in D$  for some  $C \subset \mathbb{R}$  and  $D \subset \mathbb{R}$

$$\therefore P(X \in C | Y \in D) = \frac{P(X \in C \cap Y \in D)}{P(Y \in D)}$$

$X$  and  $Y$  are independent

$$\Rightarrow P(X \in C | Y \in D) = P(X \in C)$$

$$\therefore P(X \in C) = \frac{P(X \in C \cap Y \in D)}{P(Y \in D)}$$

$$\Rightarrow P(X \in C)P(Y \in D) = P(X \in C \cap Y \in D)$$

$$P(X \in C \cap Y \in D) = P(X \in C)P(Y \in D)$$

Given  
Defn Cond Prob  
Given  
Prop of  $\perp$   
Substitution  
Algebra  
Reflective Prop  
Q.E.D.

**(e) Is  $Y$  mean independent of  $X$ ?**

Answer:  $\mathbb{E}(Y|X) = \mathbb{E}(Y)$

$$\Leftrightarrow \mathbb{E}(Y|X = x \in [1, 14]) = \mathbb{E}(Y)$$

$$\Leftrightarrow Y \text{ is mean independent of } X$$

Let  $x = 5$

$$\mathbb{E}(Y|x = 5) = \sum_{y=1}^{14} yP(Y = y|x = 5)$$

$$= \sum_{y=1}^{14} y \frac{P(Y=y \wedge X=5)}{P(X=5)}$$

$$= \sum_{y=1}^{14} y \frac{P(Y=y \wedge X=5)}{.11}$$

$$\approx 8.23$$

Since  $Y$  is symmetrical around its midpoint,  $\mathbb{E}(Y) = 7.5$

$$\therefore \mathbb{E}(Y|x = 5) \neq \mathbb{E}(Y)$$

$$\Rightarrow \mathbb{E}(Y|X) \neq \mathbb{E}(Y)$$

$$\Rightarrow Y \text{ is NOT mean independent of } X$$

(f) Define the following random variable:  $Z = 1_{X=Y}$ , where  $1_A$  is an indicator function that takes the value of one if statement  $A$  is true and the value 0 if  $A$  is false. Write the *pmf* of random variable  $Z$ .

Answer:  $P(Z = 1) = P(X = Y) = .12$   
 $\therefore, P(Z = 0) = 1 - P(X = Y) = 1 - .12 = .88$   
 $\therefore, f_Z(z) = \begin{cases} .12 & \text{if } z = 1 \\ .88 & \text{if } z = 0 \\ 0 & \text{otherwise} \end{cases}$

6. Suppose that random variable  $X$  is distributed logistic with parameters  $\mu = 0$  and  $\beta = 1$ . Define the variable  $Y = X^2$ . Also, note that  $\mathbb{E}(X^4) = \frac{26}{45}\pi^4$ .

(a) Is  $\mathbb{E}(Y)$  larger, smaller or equal to  $[\mathbb{E}(X)]^2$ ? Explain your answer.

Answer:  $var(X) = \frac{\pi^2\beta^2}{3} = \frac{\pi^2}{3}$  and  $\mathbb{E}(X) = \mu = 0$  Logistic Dist

$$var(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2$$

$$\therefore, \frac{\pi^2}{3} = \mathbb{E}(X^2) - 0^2$$

$$\Rightarrow \mathbb{E}(X^2) = \frac{\pi^2}{3}$$

$$\therefore, \mathbb{E}(Y) = \mathbb{E}(X^2) = \frac{\pi^2}{3} > 0 = \mathbb{E}(X)^2$$

$$\Rightarrow \mathbb{E}(Y) > \mathbb{E}(X)^2$$

(b) Is the variance of  $Y$  larger, smaller or equal to the variance of  $X$ ? Explain your answer.

Answer:  $var(X) = \frac{\pi^2}{3}$  and  $\mathbb{E}(X) = 0$  Logistic Dist

$$\Rightarrow \mathbb{E}(X^2) = \frac{\pi^2}{3}$$

$$var(Y) = var(X^2) = \mathbb{E}(X^4) - \mathbb{E}(X^2)^2$$

$$\therefore, var(Y) = \frac{26\pi^4}{45} - \frac{\pi^4}{9} = \frac{21\pi^4}{45}$$

$$\therefore, var(Y) = \frac{21\pi^4}{45} > \frac{\pi^2}{3} = var(X)$$

$$\Rightarrow var(Y) > var(X)$$

(c) Characterize the support of  $Y$ .

Answer:  $Y = X^2$  and  $x \in (-\infty, \infty)$

$$\therefore, x \in (-\infty, 0] \Rightarrow y \in [0, \infty) \text{ and } x \in [0, \infty) \Rightarrow y \in [0, \infty)$$

$$\therefore, y \in [0, \infty)$$

(d) What is the *pdf* of  $Y$ ? Hint: Note that because of symmetry about 0,  $f_X(x) = f_X(-x)$ .

Answer: Is  $f_X(x)$  continuous? Yes.

Can you partition the range of  $X$  into continuous pieces? Yes,  $A_1 = \{x|x \in (-\infty, 0)\}$ ,  $A_2 = \{x|x \in (0, \infty)\}$ , and  $A_0 = \{x|x = 0\}$

Do  $A_1$  and  $A_2$  correspond to the same range of  $Y$ ? Yes,  $y \in (0, \infty)$  for both  $A_1$  and  $A_2$

Are  $A_1$  and  $A_2$  monotonic? Yes, both always increase

Is the transformation identical across partitions? Yes,  $Y = X^2$

Is  $\frac{d}{dy}g^{-1}(y)$  continuous on  $A_1$  and  $A_2$ ? Yes,  $\frac{d}{dy}(-\sqrt{y})$  for  $A_1$  and  $\frac{d}{dy}\sqrt{y}$  for  $A_2$ .

$$\therefore, f_Y(y) = \begin{cases} \sum_{i=1}^k f_X(g_i^{-1}(y)) \left| \frac{d}{dy}g_i^{-1}(y) \right|, & \forall y \in [0, \infty) \\ 0, & \text{otherwise} \end{cases} \quad \text{Theorem 2.1.8}$$

$$= \begin{cases} f_X(g_1^{-1}(y)) \left| \frac{d}{dy}g_1^{-1}(y) \right| + f_X(g_2^{-1}(y)) \left| \frac{d}{dy}g_2^{-1}(y) \right|, & \forall y \in [0, \infty) \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned}
&= \begin{cases} f_X(-\sqrt{y}) \left| \frac{d}{dy}(-\sqrt{y}) \right| + f_X(\sqrt{y}) \left| \frac{d}{dy}\sqrt{y} \right|, & \forall y \in [0, \infty) \\ 0, & \text{otherwise} \end{cases} \\
f_X(-\sqrt{y}) &= f_X(\sqrt{y}) \quad \text{Symmetry of } f_X \\
\therefore, f_Y(y) &= \begin{cases} f_X(\sqrt{y}) \left| \frac{d}{dy}(-\sqrt{y}) \right| + f_X(\sqrt{y}) \left| \frac{d}{dy}\sqrt{y} \right|, & \forall y \in [0, \infty) \\ 0, & \text{otherwise} \end{cases} \\
&= \begin{cases} \frac{e^{-\sqrt{y}}}{(1+e^{-\sqrt{y}})^2} \left| -\frac{1}{2\sqrt{y}} \right| + \frac{e^{-\sqrt{y}}}{(1+e^{-\sqrt{y}})^2} \left| \frac{1}{2\sqrt{y}} \right|, & \forall y \in [0, \infty) \\ 0, & \text{otherwise} \end{cases} \\
&= \begin{cases} \frac{2e^{-\sqrt{y}}}{(1+e^{-\sqrt{y}})^2} \frac{1}{2\sqrt{y}}, & \forall y \in [0, \infty) \\ 0, & \text{otherwise} \end{cases} \\
&= \begin{cases} \frac{1}{\sqrt{y}e^{\sqrt{y}}(1+e^{-\sqrt{y}})^2}, & \forall y \in [0, \infty) \\ 0, & \text{otherwise} \end{cases}
\end{aligned}$$

(e) Define the random variable  $Z = F(x)$  where  $F(\cdot)$  is the cdf of  $X$ ,  $F_X(x)$ . What is the pdf of  $Z$ ?

Answer:  $Z = F_X(x) = \int \frac{e^{-x}}{(1+e^{-x})^2} dx$

Let  $u = 1 + e^{-x}$

$\Rightarrow du = -e^{-x} dx$

$\therefore, \int \frac{e^{-x}}{(1+e^{-x})^2} dx = \int -\frac{1}{u^2} du = \frac{1}{u} + C = \frac{1}{1+e^{-x}} + C$

pdf must integrate to 1

$\therefore, \frac{1}{1+e^{-x}} \Big|_{-\infty}^{\infty} + C = 1$

$\Rightarrow 1 - 0 + C = 1$

$\Rightarrow C = 0$

$\therefore, Z = \frac{1}{1+e^{-X}}$ , where  $x \in (-\infty, \infty)$

$\Rightarrow \begin{cases} Z \in (0,1) \\ X = g^{-1}(Z) = \ln\left(\frac{Z}{1-Z}\right) \end{cases}$

Is  $f_X(x)$  continuous on  $X$ ? Yes

Is  $g(x)$  monotone across the range of  $X$ ? Yes, it is constantly decreasing

Does  $g^{-1}(z)$  have a continuous derivative? Yes, it is  $\frac{d}{dz} g^{-1}(z) = \frac{1}{z(1-z)}$

$\therefore, f_Z(z) = \begin{cases} f_X(g^{-1}(z)) \left| \frac{d}{dz} g^{-1}(z) \right|, & \forall z \in (0,1) \\ 0, & \text{otherwise} \end{cases}$

$= \begin{cases} \frac{e^{-\ln\left(\frac{z}{1-z}\right)}}{\left[1+e^{-\ln\left(\frac{z}{1-z}\right)}\right]^2} \left| \frac{1}{z(1-z)} \right|, & \forall z \in (0,1) \\ 0, & \text{otherwise} \end{cases}$

$= \begin{cases} \frac{\frac{1-z}{z}}{\left(1+\frac{1-z}{z}\right)^2} \frac{1}{z(1-z)}, & \forall z \in (0,1) \\ 0, & \text{otherwise} \end{cases}$

Theorem 2.1.5

$$= \begin{cases} \frac{\frac{1-z}{z}}{\frac{1}{z^2}} \frac{1}{z(1-z)}, & \forall z \in (0,1) \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} 1, & \forall z \in (0,1) \\ 0, & \text{otherwise} \end{cases}$$

$$\therefore, Z \sim U(0,1)$$

**OR**

$$Z = \frac{1}{1+e^{-X}}, \text{ where } x \in (-\infty, \infty)$$

$$\Rightarrow z \in (0,1)$$

Is  $F_X(x)$  monotonically increasing in  $x$ ? Yes, it's increasing.

$$\therefore, F_Z(z) = F_X(F_X^{-1}(z)) = z, \forall z \in (0,1)$$

Theorem 2.1.3

$$\therefore, f_Z(z) = \frac{dF_X(F_X^{-1}(z))}{dz} = \frac{dz}{dz} = 1, \forall z \in (0,1)$$

$$\therefore, f_Z(z) = \begin{cases} 1, & \forall z \in (0,1) \\ 0, & \text{otherwise} \end{cases}$$

$$\therefore, Z \sim U(0,1)$$