

Bootstrap method

Woojoo Lee

We want to assess the uncertainty of a statistic under the nonparametric framework.

→ The bootstrap makes us do it ! i.e. we can estimate variance of a statistic and compute the confidence interval with bootstrap.

Let $T_n = \bar{X}_n$.

Q) How can you get the variance estimate for \bar{X}_n ?

Remark) The key problem is that we do not know how to generate random samples from F .

Suppose that we have x_1, \dots, x_n from an unknown distribution F . The parameter of interest is

$$\theta = E(h(X)) = \int h(x)dF(x)$$

We note that

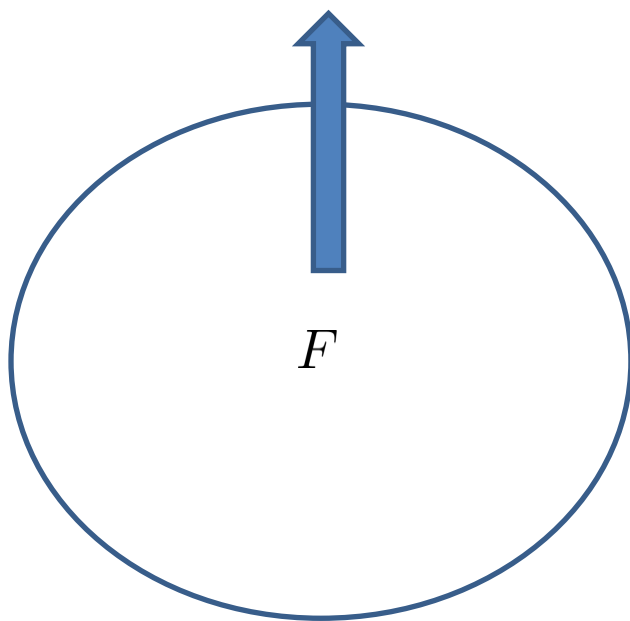
- $F \approx \hat{F}_n$
- $\hat{\theta} \approx \int h(x)d\hat{F}_n(x)$
- Often, the difficulty lies in computing $\text{var}(\hat{\theta})$.

$\text{Var}(\hat{\theta})$ requires "many samples" ! but we have only one sample in practice.

→ Pretend to know F by using \hat{F}_n and generate new samples from \hat{F}_n !

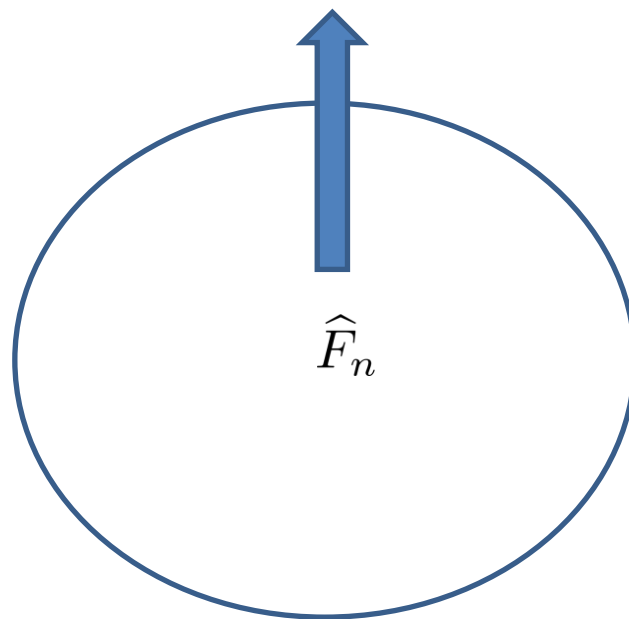
Real world

$$\{X_1, \dots, X_n\}$$



Bootstrap world

$$\{X_1^*, \dots, X_n^*\}$$



Key procedure: Sampling n observations from \hat{F}_n with replacement !

Bootstrap algorithm for computing the variance of a statistic

1. Compute the empirical CDF \hat{F}_n
2. Sample n observations from \hat{F}_n with replacement
3. Compute the statistic of interest with the bootstrap samples.
4. Repeat the above steps 2-3 B times.

Denote the statistic from b th bootstrap sample by $\hat{\theta}_b$.

Compute

$$var_B(\hat{\theta}) = \frac{1}{B-1} \sum_{b=1}^B (\hat{\theta}_b - \bar{\theta}_B)^2$$

where

$$\bar{\theta}_B = \frac{1}{B} \sum_{b=1}^B \hat{\theta}_b.$$

Sometimes, we are interested in the bias of an estimator. The bias is defined as

$$E(\hat{\theta}) - \theta.$$

We can estimate this bias from bootstrap:

$$\frac{1}{B} \sum_{b=1}^B \hat{\theta}_b - \hat{\theta}$$

.

R-example

```
### Example: calculating se for median
```

```
set.seed(1201)
```

```
data<-rnorm(100,5,3)
```

```
B<-1000
```

```
b.samples<-lapply(1:B,function(i) sample(data,replace=T))
```

```
b.median<-sapply(b.samples,median)
```

```
hist(b.median)
```

```
sqrt(var(b.median))
```

```
## in theory,
```

```
##  $p*(1-p)/(n*f(5)^2)=1/(100*4*f(5)^2)=0.1414$  ( $f(5)=dnorm(0,0,3)$ )
```

```
##  $\sqrt{0.1414}=0.376$ 
```

R-example

```
### Example: calculating se for skewness
```

```
library(moments)
```

```
set.seed(1201)
```

```
data<-rnorm(100,5,3)
```

```
B<-1000
```

```
b.samples<-lapply(1:B,function(i) sample(data,replace=T))
```

```
b.skewness<-sapply(b.samples,skewness)
```

```
hist(b.skewness)
```

```
sqrt(var(b.skewness))
```


R-example

```
library(boot)  
data(bigcity)
```

```
## we want to know the mean ratio of the populations,  
## i.e. pop 1930/pop 1920
```

```
row.bigcity<-dim(bigcity)[1]
```

```
boots.bigcity<-function(index){  
  b.bigcity<-bigcity[index,]  
  b.ratio<-sum(b.bigcity$x)/sum(b.bigcity$u)  
  return(b.ratio)  
}
```

```
B<-1000  
b.samples<-lapply(1:B,function(i) sample(c(1:row.bigcity),replace=T))  
b.ratio<-sapply(b.samples,boots.bigcity)
```

```
hist(b.ratio)
```

```
sqrt(var(b.ratio))
```

R-example

```
##### As an alternative, you may use "boot".  
##### Before calling boot,  
##### you need to define a function that will return the statistic  
##### that you want to bootstrap.
```

```
library(boot)  
ratio <- function(d, indices) sum(d$x[indices])/sum(d$u[indices])  
RES.city<-boot(bigcity, ratio, R = 999)  
  
boot.ci(RES.city,type=c("norm","basic","perc","bca"))
```

Bootstrap confidence intervals

Bootstrap confidence intervals

We will study three different bootstrap confidence intervals.

The normal confidence interval based on bootstrap:

Suppose that $\hat{\theta}$ is the observed statistic. Then, the $1 - \alpha$ normal interval is given by

$$(\hat{\theta} - z_{\alpha/2} \hat{se}_{boot}, \hat{\theta} + z_{1-\alpha/2} \hat{se}_{boot})$$

where \hat{se}_{boot} is the bootstrap estimate of standard error.

Remark) This interval is not accurate if the sampling distribution of $\hat{\theta}$ is not close to normal.

Suppose that $\widehat{\theta}^*_{\alpha}$ is α -percentile of $\widehat{\theta}^*_i$. The $1 - \alpha$ bootstrap percentile interval is given by

$$(\widehat{\theta}^*_{\alpha/2}, \widehat{\theta}^*_{1-\alpha/2})$$

Some properties

- Easy to use !
- This works well when the bootstrap distribution is symmetric and centered on the observed statistic.
- Many literatures report that this can be narrow in small samples.

The $1 - \alpha$ basic bootstrap confidence interval is given by

$$(2\hat{\theta} - \hat{\theta}^*_{1-\alpha/2}, 2\hat{\theta} - \hat{\theta}^*_{\alpha/2})$$

Q) Derive the basic bootstrap CI.

Remark) The key idea is that the distribution of $\hat{\theta}^* - \hat{\theta}$ is approximately the same as that of $\hat{\theta} - \theta$.

Algorithm for bootstrap-t (studentized) confidence interval

- Generate B bootstrap samples ($*$ denotes the statistic from the bootstrap samples.)
- Compute $t_i^* = \frac{\hat{\theta}_i^* - \hat{\theta}}{\hat{\sigma}_i^*}$ where $\hat{\theta}_i^*$ and $\hat{\sigma}_i^*$ denote the statistic and s.e. from i th bootstrap sample, respectively.
- Compute α -percentile of t_i^* : find t_α^* satisfying $\#(t_i^* \leq t_\alpha^*)/B = \alpha$.
- The bootstrap-t $1 - \alpha$ confidence interval is given by

$$(\hat{\theta} - t_{\alpha/2}^* \hat{\sigma}, \hat{\theta} + t_{1-\alpha/2}^* \hat{\sigma})$$

Some properties

- The bootstrap t-confidence interval reflects the skewness of the data.
- More accurate than the percentile and the basic confidence intervals.

R-example for various bootstrap CIs

Example: calculating se for median

```
set.seed(1201)
```

```
data<-rnorm(100,5,3)
```

```
B<-1000
```

```
b.samples<-lapply(1:B,function(i) sample(data,replace=T))
```

```
b.median<-sapply(b.samples,median)
```

```
hist(b.median)
```

```
sqrt(var(b.median))
```

normal interval

```
c(mean(b.median)-2*sqrt(var(b.median)),mean(b.median)+2*sqrt(var(b.median)))
```

percentile interval

```
c(quantile(b.median,0.025),quantile(b.median,0.975))
```

basic interval

```
c(2*median(data)-quantile(b.median,0.975),2*median(data)-quantile(b.median,0.025))
```


Bootstrap in regression

Consider a simple linear regression model:

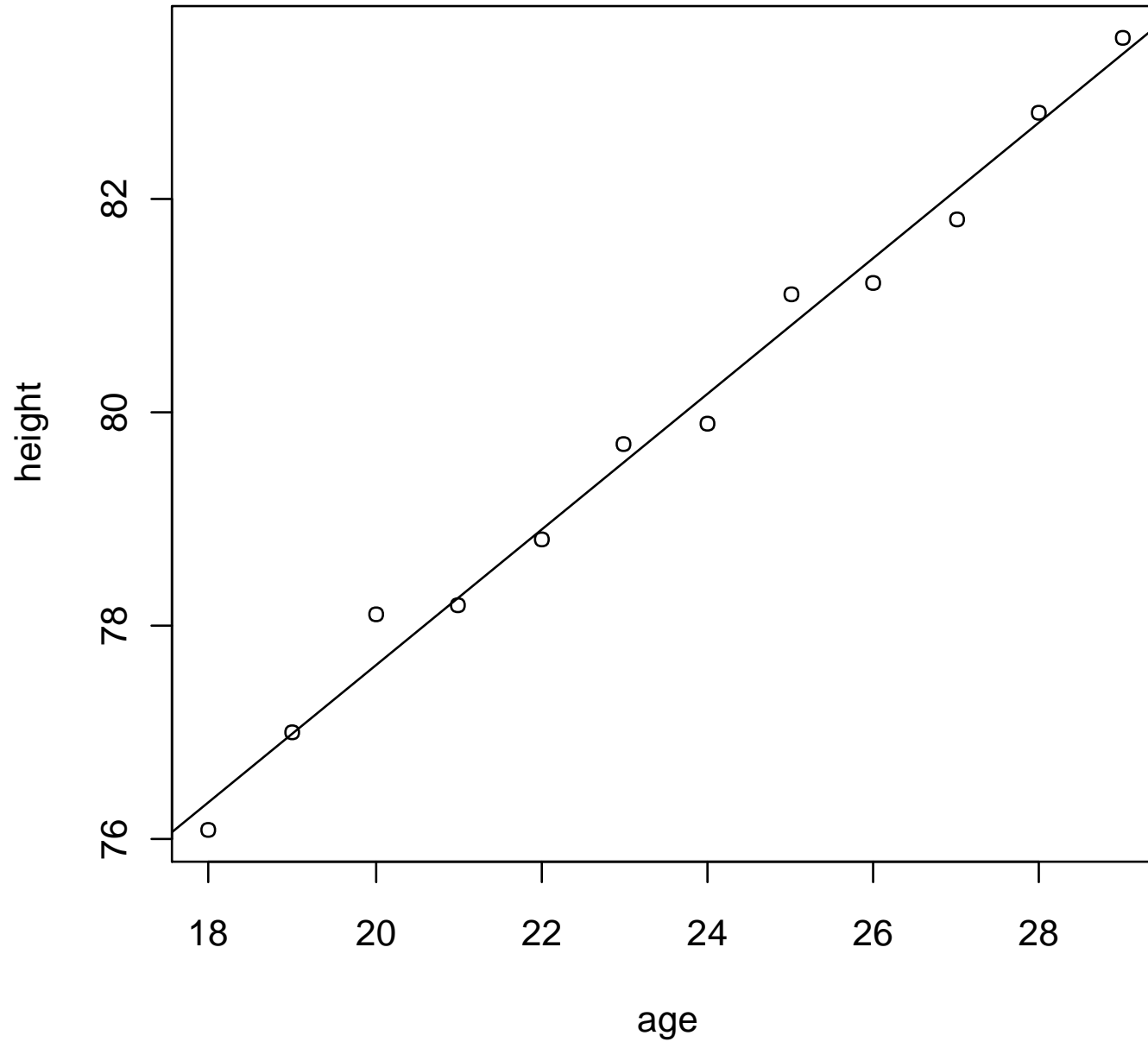
$$y_i = \beta_0 + x_i\beta_1 + e_i$$

where e_i is i.i.d. with mean 0 and variance σ^2 . Note that there is no parametric assumption on the distribution for e_i .

Two alternatives bootstrapping methods are available.

- case 1) bootstrapping pairs (x_i, y_i)
- case 2) bootstrapping residuals

R-example for bootstrapping in regression



Q) Explain the bootstrap method for case 1 (case sampling).

```
boots.pair<-function(index){  
  b.age<-age[index]  
  b.height<-height[index]  
  b.coeff<-coef(lm(b.height~b.age))[2]  
  return(b.coeff)  
}
```

```
B<-1000  
set.seed(1210)  
b.samples<-lapply(1:B,function(i)  
  sample(c(1:length(height)),replace=T))  
b.PWD<-sapply(b.samples,boots.pair)
```

```
mean(b.PWD)  
sd(b.PWD)
```

Q) Explain the bootstrap method for case 2 (model-based sampling).

```
lm.res<- lm(height ~ age)
residual<-resid(lm.res)
fit<-fitted(lm.res)
```

```
boots.resid<-function(index){
  newy<-fit+residual[index]
  b.coef<-coef(lm(newy~age))[2]
  return(b.coef)
}
```

```
B<-1000
set.seed(1210)
b.samples<-lapply(1:B,function(i)
  sample(c(1:length(height)),replace=T))
b.PWD<-sapply(b.samples,boots.resid)
```

```
mean(b.PWD)
sd(b.PWD)
```

Q) Which one is better ?