1.

(i)a.利用 Quadrature integrtion,求 f_z(z) 積分從 0 到 0.7,

用 R 算的真實值為 0.872516089908

	真實值-I.hat <0.0001,所需要到樣本數	
Rectangle(矩形法)	1670	
Trapezoidal(梯型法)	44	
Simpson(辛普森)	6	

b.利用 Monte Carlo method with U(0,0.7), 求 f_z(z) 積分從 0 到 0.7 .

由於 Monte Carlo 估計量本身為隨機變數,故每次|真實值-I.hat|<0.0001,所需要到樣本數都不一樣。

對於不同的重複生成次數,建構 95%的信賴區間:

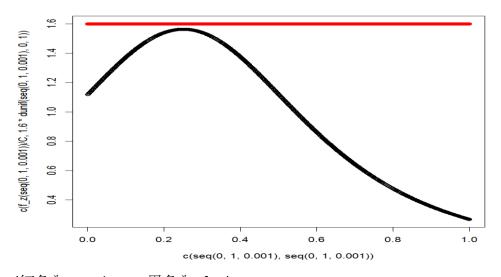
重複生成次數	樣本平均	95%信賴區間
1670	0.8726629	[0.8714788 0.8738471]
44	0.8762858	[0.8688807
6	0.8668569	[0.8501588 0.8835549]

(ii)利用 empirical CDF 去估計 F_z(0.7)=P(Z < 0.7) 的值,

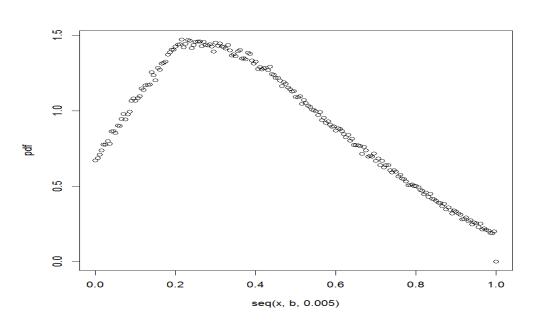
1. sample z1, z2, ..., zn iid from f_z

2.
$$F_n(0.7) = 1/n \sum_{i=0}^n I(Z_i \le 0.7)$$

首先利用 Rejection sampling 生成z1, z2, ..., zn 選擇 g~U(0,1)作為 envolope, e(x)=c*g(x),取 c=1.6 畫圖如下



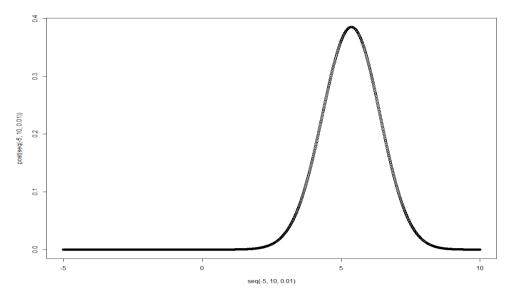
(紅色為 envolope; 黑色為 f_z)



(此為生成後z1, z2, ..., zn 利用 empirical pdf 估計 f_z)

由於 $F_n(0.7) = 1/n \sum_{i=0}^n I(Z_i \le 0.7)$ 為一個估計量,可以得到在樣本數為 10^5 時估計值為 0.87254 , bias 為 2.391009e-05.

2.估計 mean and variance of posterior density , 由於此處無法理論推得 posterior density 屬於某個特定分配 , 模擬 theta=5 ,X ~N(5,1) , 先畫出 posterior density ,



(posterior density 給定 theta=5,X~N(5,1)的畫圖)

$$E(\theta|X) = \int_{-\infty}^{\infty} \theta p(\theta|X) d\theta$$
$$= \int_{-\infty}^{\infty} \theta \frac{p(X|\theta)p(\theta)}{p(X)} d\theta$$

我們生成 x1,x2,...,xn $\sim N(5,1)$,取 $n=10^4$,利用 Simpson method 算出積分值 作為一個估計值 ,得到對於 mean 及 variance 的樣本估計為

I_bar=4.528903 , I_var=1.213188 , bias=0.471097

發現利用樣本平均和樣本方差估計 posterior mean and variance ,不是一個不偏的估計量.

(附上程式碼部分)

1.

```
f z<-function(z){ (1+( (z-0.25)/0.5 )^2 )^-1.5}
C<-integrate(f_z,0,1)$value
I true<-integrate(f z,0,0.7)$value /C
plot(seq(0,1,0.001),f_z(seq(0,1,0.001))/C)
#Quadrature integration
a<-0; b<-0.7; n<-100
h < -seq(a,b,(b-a)/n)
#Rectangle rule
I_upper<-0;I_lower<-0
for(i in 1:n){
  I_upper < -I_upper + f_z(h[i+1])*(b-a)/n/C
  I_lower < -I_lower + f_z(h[i])*(b-a)/n/C
}
a<-0; b<-0.7; n<-0
I upper<-0
while(abs(I_upper-I_true)>=10^-4){
  cat("\n","Now is",n,"-th iteration ,bias is",abs(I_upper-I_true),"\n")
  n<-n+1
  h < -seq(a,b,(b-a)/n)
  I upper<-0
  for(i in 1:n){ I_upper < I_upper + f_z(h[i+1])*(b-a)/n/C}
```

```
}
cat("\n", "The sample size is", n, ", bias is", abs(I upper-I true), "\n")
a<-0; b<-0.7; n<-0
I lower<-0
while(abs(I lower-I true)>=10^-4){
  cat("\n","Now is",n,"-th iteration ,bias is",abs(I_lower-I_true),"\n")
  n<-n+1
  h < -seq(a,b,(b-a)/n)
  I_lower<-0
  for(i in 1:n){ I_lower < I_lower + f_z(h[i])*(b-a)/n/C}
}
cat("\n","The sample size is",n,",bias is",abs(I_lower-I_true),"\n")
#Trapezoidal rule
I_T<-0
for(i in 1:n){ I_T<-I_T+(f_z(h[i])+f_z(h[i+1]))/2 }
I_T<-I_T*(b-a)/n/C
a<-0; b<-0.7; n<-0
I_T<-0
while(abs(I_T-I_true)>=10^-4){
  cat("\n","Now is",n,"-th iteration ,bias is",abs(I_T-I_true),"\n")
  n<-n+1
  h < -seq(a,b,(b-a)/n)
  I T<-0
  for(i in 1:n){ I_T<-I_T+(f_z(h[i])+f_z(h[i+1]))/2 }
  I_T<-I_T*(b-a)/n/C
}
cat("\n","The sample size is",n,",bias is",abs(I_T-I_true),"\n")
#Simpsons rule
if(n%%2==0){
  z1<-0;z2<-0
  for(i in 2:(n/2)){ z1<-z1+2*f_z(h[2*i-1]) }
  for(i in 1:(n/2)){ z2<-z2+4*f_z(h[2*i]) }
  I_S<-(f_z(h[1])+f_z(h[n+1])+z1+z2)*(b-a)/n/3/C
  rm(z1,z2)
```

```
} else { print("n need to be even") }
a<-0; b<-0.7; n<-0
I S<-0
while(abs(I S-I true)>=10^-4){
  cat("\n","Now is",n,"-th iteration ,bias is",abs(I S-I true),"\n")
  n<-n+2
  h < -seq(a,b,(b-a)/n)
  I S<-0
  z1<-0;z2<-0
  for(i in 2:(n/2)){ z1<-z1+2*f_z(h[2*i-1]) }
  for(i in 1:(n/2)){ z2<-z2+4*f z(h[2*i]) }
  I_S<-(f_z(h[1])+f_z(h[n+1])+z_1+z_2)*(b-a)/n/3/C
};rm(z1,z2)
cat("\n","The sample size is",n,",bias is",abs(I S-I true),"\n")
#Monte Carlo method with U(0,0.7)
G<-function(z,n){
  z < -c(z,rep(0,n))
  for(i in 1:n){z[i+1]<-(16807*z[i])\%\%(2^31-1)}
  u < -z[-1]/(2^31-1)
  return(u)
}
confi.of.I_M<-function(n,N,alpha){
  #n=100 ;seed<-1;N=50 ;alpha=0.05
  x < -rep(0,N); seed < -1
  for(i in 1:N){
     u < -rep(0,n)
     for(j in 1:n){
       u[j] < f_z(0.7*G(seed,1))/(10/7)
       seed<-G(seed,1)*(2^31-1)}
     x[i]<-x[i]+sum(u)/n/C
                                            #I M<-sum(u)/n/C
  }
  I.bar<-sum(x)/N
  I.sd < -sd(x)
  cat("The 95% confidence interval for estimating I is [",
       I.bar-I.sd/sqrt(N)*qnorm(1-alpha),"",
```

```
I.bar+I.sd/sqrt(N)*qnorm(1-alpha),"]","\n",I.bar)
}
confi.of.I M(6,50,0.05)
seed<-2
I M<-0;n<-0
while(abs(I M-I true)>=10^-4){
  cat("\n","Now is",n,"-th iteration ,bias is",abs(I_M-I_true),"\n")
  n<-n+1
  u < -rep(0,n)
  for(j in 1:n){
    u[j] < -f_z(0.7*G(seed,1))/(10/7)
    seed<-G(seed,1)*(2^31-1)}
  I_M<-sum(u)/n/C
}
cat("\n","The sample size is",n,",bias is",abs(I_M-I_true),"\n")
#Acce. Rejection
windows();plot(x=c(seq(0,1,0.001),seq(0,1,0.001)),y=c(f_z(seq(0,1,0.001))/
C,1.6*dunif(seq(0,1,0.001),0,1)),
                  col = c(rep(1,1001), rep(2,1001)))
c<-1.6
envo<-function(x,c){ f_z(x)/C/c}
acc.rej.exp<-function(n,c=1.6){
  u1 < -runif(n,0,1)
  Y<-u1
  u2 < -runif(n,0,1)
  X < -rep(0,n)
  N<-length(which(u2 <= envo(Y,c)))
  for(i in 1:N){ X[i] < Y[which(u2 <= envo(Y,c))][i] }
                                                             #exp(-(Y-
1)^2/2)
  while(N<n){
    uu1 < -runif(n-N,0,1)
    Y<-uu1
    uu2 < -runif(n-N,0,1)
    if(length(which(uu2<=envo(Y,c)))>0){
       for(i in 1:length(which(uu2<=envo(Y,c)))){ X[N+i]<-
Y[which(uu2<=envo(Y,c))][i] }}
    N<-N+length(which(uu2 <= envo(Y,c)))
```

```
return(X)
}
acc.rej.exp(1)
PDF<-function(a,b,n,band){
  #a=0;b=1;n=1000;band=0.2
  x<-a
  pdf < -rep(0, length(seq(x, b, 0.005)))
  i<-1
  while(a<=b){
     I<-0
     for(i in 1:n){ if( abs(a-acc.rej.exp(1)) \le band){ I \le I+1 }
     pdf[j] < -I/(2*n*band)
     a<-a+0.005
    j<-j+1
  }
  windows();plot(seq(x,b,0.005),pdf)
  return(pdf) }
PDF(0,1,5000,0.2)
# F(0.7)
cdf<-0;n<-1
while(abs(cdf-I_true)>=10^-4){
  cat("\n","bias is",abs(cdf-I_true),"\n")
  if(n<10^5){n<-n*10}
  for(i in 1:n){ if(acc.rej.exp(1)<=0.7){ I<-I+1 }
  cdf<-I/n
}
cat("\n","The sample size is",n,",bias is",abs(cdf-I true),"\n")
```

2.

```
px.theta<-function(x,theta){ 1/sqrt(2*pi)*exp(-(x-theta)^2/2) }
p.theta<-function(theta){ 1/(pi*(1+theta^2)) }
px<-function(theta){ px.theta(x,theta)*p.theta(theta) }
#plot(seq(-10,10,0.01),px(seq(-10,10,0.01))) #plot p(x)
post<-function(theta){ px.theta(x,theta)*p.theta(theta)/I_Spx }
f<-function(theta){ theta*post(theta) } #theta*p(theta|x)
#plot p(theta|x)
theta<-5;x<-rnorm(1,theta,1)</pre>
```

```
a<-theta-10; b<-theta+10; n<-100
h < -seq(a,b,(b-a)/n)
if(n%%2==0){
  z1<-0;z2<-0
  for(i in 2:(n/2)){ z1<-z1+2*px(h[2*i-1]) }
  for(i in 1:(n/2)){ z2<-z2+4*px(h[2*i]) }
  I Spx<-(px(h[1])+px(h[n+1])+z1+z2)*(b-a)/n/3
  rm(z1,z2)
} else { print("n need to be even") }
plot(seq(-5,10,0.01),post(seq(-5,10,0.01)))
#estmator with theta=5
N=10000; theta=5
y < -rep(0,N)
for(j in 1:N){
  x<-rnorm(1,theta,1)
  a<-theta-10; b<-theta+10; n<-100
  h < -seq(a,b,(b-a)/n)
  if(n%%2==0){
    z1<-0;z2<-0
    for(i in 2:(n/2)){ z1<-z1+2*px(h[2*i-1]) }
    for(i in 1:(n/2)){ z2<-z2+4*px(h[2*i]) }
    I_Spx<-(px(h[1])+px(h[n+1])+z1+z2)*(b-a)/n/3
    rm(z1,z2)
  } else { print("n need to be even") }
  if(n%%2==0){
    z1<-0;z2<-0
    for(i in 2:(n/2)){ z1<-z1+2*f(h[2*i-1]) }
    for(i in 1:(n/2)){ z2<-z2+4*f(h[2*i]) }
    I_Sposterior<-(f(h[1])+f(h[n+1])+z1+z2)*(b-a)/n/3
    rm(z1,z2)
  } else { print("n need to be even") }
  y[j]<-I_Sposterior
};rm(i,I_Spx,I_Sposterior)
I_bar<-sum(y)/N
I_var<-(sum(y^2)-N*I_bar^2)/(N-1)
list(theta=theta,I_bar=I_bar,I_var=I_var)
```

```
#use simpsons to est px's integral
a<--5; b<-10; n<-100
h < -seq(a,b,(b-a)/n)
if(n%%2==0){
  z1<-0;z2<-0
  for(i in 2:(n/2)){ z1<-z1+2*px(h[2*i-1]) }
  for(i in 1:(n/2)){ z2<-z2+4*px(h[2*i]) }
  I_Spx<-(px(h[1])+px(h[n+1])+z1+z2)*(b-a)/n/3
  rm(z1,z2)
} else { print("n need to be even") }
                                               #integrate(px,-Inf,Inf)
#use simpsons to est [theta*p(theta|x)]'s integral
a<--5; b<-10; n<-1000
h < -seq(a,b,(b-a)/n)
if(n%%2==0){
  z1<-0;z2<-0
  for(i in 2:(n/2)){ z1<-z1+2*f(h[2*i-1]) }
  for(i in 1:(n/2)){ z2<-z2+4*f(h[2*i]) }
  I_Sposterior < -(f(h[1]) + f(h[n+1]) + z1 + z2)*(b-a)/n/3
  rm(z1,z2)
 } else { print("n need to be even") }
                                              #integrate(f,-Inf,Inf)
```