# Probability & Statistics for EECS: Homework #6 Solutions

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The Cauchy distribution has PDF

$$f(x) = \frac{1}{\pi \left(1 + x^2\right)}$$

for all x. Find the CDF of a random variable with the Cauchy PDF. Hint: Recall that the derivative of the inverse tangent function  $\tan^{-1}(x)$  is  $\frac{1}{1+x^2}$ .

## Solution

Given that the PDF of the Cauchy distribution is

$$f(x) = \frac{1}{\pi(1+x^2)},$$

and the hint that the derivative of the inverse tangent function  $\tan^{-1}(x)$  is  $\frac{1}{1+x^2}$ , we can calculate the CDF of the Cauchy distribution by definition, *i.e.*, integrating the PDF over range  $(-\infty, x]$ . Therefore, the CDF of the Cauchy distribution F(x) is as follows:

$$F(x) = \int_{-\infty}^{x} f(t)dt = \frac{1}{\pi} \tan^{-1}(t) \Big|_{-\infty}^{x} = \frac{1}{\pi} \tan^{-1}(x) - \frac{1}{\pi} (-\frac{\pi}{2}) = \frac{1}{2} + \frac{1}{\pi} \tan^{-1}(x),$$

where  $x \in (-\infty, \infty)$ .

The Pareto distribution with parameter a > 0 has PDF

$$f(x) = \frac{a}{x^{a+1}}$$

for  $x \ge 1$  (and 0 otherwise). This distribution is often used in statistical modeling. Find the CDF of a Pareto r.v. with parameter a; check that it is a valid CDF.

## Solution

Given that the PDF of the Pareto distribution with parameter  $\alpha > 0$  is

$$f(x) = \begin{cases} \frac{a}{x^{a+1}}, & x \ge 1\\ 0, & \text{Otherwise} \end{cases},$$

we can calculate the CDF of the Pareto distribution by definition, *i.e.*, integrating the PDF over range  $(-\infty, x]$ .

Therefore, the CDF of the Pareto distribution F(x) is as follows:

$$F(x) = \int_{-\infty}^{x} f(t)dt = \int_{1}^{x} \frac{a}{t^{a+1}}dt = -\frac{1}{t^{a}}\Big|_{1}^{x} = 1 - \frac{1}{x^{a}},$$

where  $x \in [1, \infty)$ . When  $x \in (-\infty, 1)$ , by definition, F(x) = 0. We then check if F(x) is a valid CDF as follows:

- Increasing: Due to the fact that  $\frac{1}{x^a}$ , a > 0 is decreasing over  $[1, \infty)$ , CDF  $F(x) = 1 \frac{1}{x^a}$  is increasing over the corresponding support  $[1, \infty)$ .
- Right-continuous: Due to the fact that  $1 \frac{1}{x^a}$ , a > 0 is continuous over  $[1, \infty)$ , CDF F(x) is right-continuous over the corresponding support  $[1, \infty)$ .
- Convergence to 0 and 1 in the limits: Due to the fact that F(x) = 0, x < 1 and  $\lim_{x \to \infty} \frac{1}{x^a} = 0$  when a > 0, CDF F(x) have its limits as follows:

$$\lim_{x \to -\infty} F(x) = 0, \lim_{x \to \infty} F(x) = 1 - 0 = 1.$$

In summary, the CDF F(x) is valid.

The Beta distribution with parameters  $a=3,\,b=2$  has PDF

$$f(x) = 12x^2(1-x)$$
, for  $0 < x < 1$ .

Let X have this distribution.

- (a) Find the CDF of X.
- (b) Find P(0 < X < 1/2).
- (c) Find the mean and variance of X (without quoting results about the Beta distribution).

#### Solution

(a) The CDF of X is

$$\begin{split} F(X) &= \int_0^x f(t)dt = \int_0^x 12t^2(1-t)dt \\ &= \int_0^x 12t^2dt - \int_0^x 12t^3dt \\ &= 4t^3|_0^x - 3t^4|_0^x \\ &= x^3(4-3x), \quad \text{for } 0 < x < 1 \end{split}$$

- (b) According to CDF F(x),  $P(0 < x < 1/2) = F(1/2) = \frac{5}{16}$ .
- (c) According to PDF, the mean of X is

$$E(X) = \int_0^1 x f(x) dx = \int_0^1 12x^2 (1 - x) dx$$
$$= \int_0^1 12x^3 dx - \int_0^1 12x^4 dx$$
$$= \frac{3}{5}$$

We have

$$E(X^{2}) = \int_{0}^{1} x^{2} f(x) dx = \int_{0}^{1} 12x^{4} (1 - x) dx$$
$$= \int_{0}^{1} 12x^{4} dx - \int_{0}^{1} 12x^{5} dx$$
$$= \frac{2}{5}$$

Thus, we have

$$Var(X) = E(X^2) - EX^2 = \frac{1}{25}$$

The Exponential is the analog of the Geometric in continuous time. This problem explores the connection between Exponential and Geometric in more detail, asking what happens to a Geometric in a limit where the Bernoulli trials are performed faster and faster but with smaller and smaller success probabilities. Suppose that Bernoulli trials are being performed in continuous time; rather than only thinking about first trial, second trial, etc., imagine that the trials take place at points on a timeline. Assume that the trials are at regularly spaced times  $0, \Delta t, 2\Delta t, \ldots$ , where  $\Delta t$  is a small positive number. Let the probability of success of each trial be  $\lambda \Delta t$ , where  $\lambda$  is a positive constant. Let G be the number of failures before the first success (in discrete time), and T be the time of the first success (in continuous time).

- (a) Find a simple equation relating G to T. Hint: Draw a timeline and try out a simple example.
- (b) Find the CDF of T. Hint: First find P(T > t).
- (c) Show that as  $\Delta t \to 0$ , the CDF of T converges to the Expo( $\lambda$ ) CDF, evaluating all the CDFs at a fixed  $t \ge 0$ .

#### Solution

- (a)  $T = G\Delta t$ .
- (b) For  $t \ge 0$ ,  $P(T > t) = P(G > \frac{t}{\Delta t}) = P$  (no success in the first  $\lfloor \frac{t}{\Delta t} \rfloor$  trials)  $= (1 \lambda \Delta t)^{\lfloor \frac{t}{\Delta t} \rfloor + 1}$ . Thus The CDF of T is  $P(T < t) = 1 P(T > t) = 1 (1 \lambda \Delta t)^{\lfloor \frac{t}{\Delta t} \rfloor + 1}.$
- (c) As  $\Delta t \to 0$ ,

$$\begin{split} \lim_{\Delta t \to 0} P\left(T \le t\right) &= \lim_{\Delta t \to 0} \left[1 - \left(1 - \lambda \Delta t\right)^{\left\lfloor \frac{t}{\Delta t} \right\rfloor + 1}\right] = 1 - \lim_{\Delta t \to 0} \left(1 - \lambda \Delta t\right)^{\frac{t}{\Delta t}} \\ &= 1 - \lim_{\Delta t \to 0} \left[\left(1 - \lambda \Delta t\right)^{\frac{1}{\lambda \Delta t}}\right]^{\lambda t} = 1 - e^{-\lambda t}. \end{split}$$

Thus for  $t \geq 0$ , the CDF of T converges to the Expo( $\lambda$ ) CDF as  $\Delta t \rightarrow 0$ .

Let  $Z \sim \mathcal{N}(0,1)$ , and c be a nonnegative constant. Find  $E(\max(Z-c,0))$ , in terms of the standard Normal CDF  $\Phi$  and PDF  $\varphi$ .

Let  $\varphi$  be the PDF of  $\mathcal{N}(0,1)$ , then we have

$$E(\max(Z - c, 0)) = \int_{-\infty}^{\infty} \max(z - c, 0)\varphi(z) dz$$

$$= \int_{c}^{\infty} (z - c)\varphi(z) dz$$

$$= \int_{c}^{\infty} z\varphi(z) dz - c \int_{c}^{\infty} \varphi(z) dz$$

$$= \frac{-1}{\sqrt{2\pi}} e^{-z^{2}/2} \Big|_{c}^{\infty} - c(1 - \Phi(c))$$

$$= \frac{1}{\sqrt{2\pi}} e^{-c^{2}/2} - c(1 - \Phi(c))$$

$$= \varphi(c) + c\Phi(c) - c$$
(1)

(Optional Challenging Problem) Let  $X \sim \mathcal{N}(0,1)$ , its corresponding CDF is denoted as  $\Phi$  and the corresponding PDF is denoted as  $\varphi$ .

(a) If x > 0, show the following inequality holds:

$$\frac{x}{x^2 + 1}\varphi(x) \le 1 - \Phi(x) \le \frac{1}{x}\varphi(x).$$

(b) Define the function g(x) as follows:

$$g(x) = \frac{2}{\sqrt{\pi}} \int_{x}^{\infty} e^{-t^2} dt, \forall x \ge 0.$$

Show the following inequality holds:

$$g(x) \le e^{-x^2}, \forall x \ge 0.$$

#### Solution

(a) Let X be a standard normal random variable. These notes present upper and lower bounds for the complementary cumulative distribution function

$$\Phi^{c}(x) = P(X > x) = \frac{1}{\sqrt{2\pi}} \int_{x}^{\infty} e^{-t^{2}/2} dt.$$

An upper bound is easy to obtain. Since t/x > 1 for t in  $(x, \infty)$ , we have

$$\Phi^{c}(x) = \frac{1}{\sqrt{2\pi}} \int_{x}^{\infty} e^{-t^{2}/2} dt$$

$$< \frac{1}{\sqrt{2\pi}} \int_{t}^{\infty} \frac{t}{x} e^{-t^{2}/2} dt$$

$$= \frac{1}{\sqrt{2\pi}} \frac{1}{x} e^{-x^{2}/2}.$$

We can also show there is a lower bound

$$\Phi^c(x) > \frac{1}{\sqrt{2\pi}} \frac{x}{x^2 + 1} e^{-x^2/2}.$$

To prove this lower bound, define

$$g(x) = \Phi^{c}(x) - \frac{1}{\sqrt{2\pi}} \frac{x}{x^2 + 1} e^{-x^2/2}.$$

We will show that g(x) is always positive. Clearly g(x) > 0. From the derivative

$$g'(x) = -\frac{2}{\sqrt{2\pi}} \frac{e^{-x^2/2}}{(x^2+1)^2}$$

we see that g is strictly decreasing. Since the limit of g(x) as x goes infinity vanishes, g must always be positive.

Combining the inequalities above we have

$$\frac{x}{x^2+1} < \sqrt{2\pi}e^{x^2/2}\Phi^c(x) < \frac{1}{x}.$$

(b) The function g(x) as defined is related to the complementary error function, which is used in statistics to describe the tail distribution of the normal curve. To prove that  $g(x) \le e^{-x^2}$  for all  $x \ge 0$ , we can start by expressing g(x) in terms of the error function and then use known inequalities to establish the desired result.

Let's first write down g(x):

$$g(x) = \frac{2}{\sqrt{\pi}} \int_{x}^{\infty} e^{-t^2} dt$$

We know that the complementary error function, denoted as  $\operatorname{erfc}(x)$ , is defined as:

$$\operatorname{erfc}(x) = \frac{2}{\sqrt{\pi}} \int_{x}^{\infty} e^{-t^{2}} dt$$

So we can say that  $g(x) = \operatorname{erfc}(x)$ .

The inequality to prove is:

$$\operatorname{erfc}(x) \le e^{-x^2}$$
, for  $x \ge 0$ 

To prove this, we'll use a standard approach that involves comparing the rate of decrease of both sides as x increases. Specifically, we'll look at the derivatives of both sides with respect to x and show that the derivative of  $\operatorname{erfc}(x)$  is always less than or equal to the derivative of  $e^{-x^2}$ , and that  $\operatorname{erfc}(x)$  and  $e^{-x^2}$  are equal at  $x = \infty$ , from which the result will follow.

The derivative of  $e^{-x^2}$  with respect to x is:

$$\frac{d}{dx}e^{-x^2} = -2xe^{-x^2}$$

And the derivative of erfc(x) is:

$$\frac{d}{dx}\operatorname{erfc}(x) = -\frac{2}{\sqrt{\pi}}e^{-x^2}$$

We can now compare the absolute values of the derivatives for  $x \geq 0$ :

$$2xe^{-x^2}$$
 versus  $\frac{2}{\sqrt{\pi}}e^{-x^2}$ 

Since  $x \ge 0$  and  $\sqrt{\pi} > 1$ , it is clear that:

$$-2xe^{-x^2} \le -\frac{2}{\sqrt{\pi}}e^{-x^2}$$

This shows that the rate at which  $e^{-x^2}$  decreases is always greater than or equal to the rate at which  $\operatorname{erfc}(x)$  decreases for  $x \geq 0$ . Since both functions tend to 0 as x approaches infinity, and  $e^{-x^2}$  decreases faster, it follows that for all  $x \geq 0$ :

$$\operatorname{erfc}(x) \le e^{-x^2}$$
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