1. Suppose Y is a discrete random variable with probability function $p(y) = ky(1/4)^y$, $y = 0, 1, 2, 3, \ldots$ Find

(a)
$$k$$
 and (b) $E(Y)$ and $V(Y)$.

Solution. Let p = 1/4.

(a) We have that

$$1 = \sum_{y=0}^{\infty} kyp^y$$

$$= \sum_{y=0}^{\infty} kp \left(\frac{d}{dp}p^y\right)$$

$$= kp \frac{d}{dp} \sum_{y=0}^{\infty} p^y$$

$$= kp \frac{d}{dp} \left(\frac{1}{1-p}\right)$$

$$= \frac{kp}{(1-p)^2}.$$

It follows that $k = \frac{(1-p)^2}{p} = \frac{9}{4}$.

(b) We have that

$$E(Y) = \sum_{y=0}^{\infty} ky^{2}p^{y}$$

$$= \sum_{y=0}^{\infty} ky^{2}p^{y} - \sum_{y=0}^{\infty} kyp^{y} + \sum_{y=0}^{\infty} kyp^{y}$$

$$= \sum_{y=0}^{\infty} k(y^{2} - y)p^{y} + \sum_{y=0}^{\infty} kyp^{y}$$

$$= \sum_{y=0}^{\infty} kp^{2} \left(\frac{d^{2}}{dp^{2}}p^{y}\right) + \sum_{y=0}^{\infty} kyp^{y}$$

$$= kp^{2} \frac{d^{2}}{dp^{2}} \sum_{y=0}^{\infty} p^{y} + 1$$

$$= kp^{2} \frac{d^{2}}{dp^{2}} \left(\frac{1}{1-p}\right) + 1$$

$$= \frac{2kp^{2}}{(1-p)^{3}} + 1$$

$$= \frac{5}{3},$$

and

$$E(Y^{2}) = \sum_{y=0}^{\infty} ky^{3}p^{y}$$

$$= \sum_{y=0}^{\infty} ky^{3}p^{y} - 3\sum_{y=0}^{\infty} ky^{2}p^{y} + 2\sum_{y=0}^{\infty} kyp^{y} + 3\sum_{y=0}^{\infty} ky^{2}p^{y} - 2\sum_{y=0}^{\infty} kyp^{y}$$

$$= \sum_{y=0}^{\infty} k(y^{3} - 3y^{2} + 2y)p^{y} + 3E(Y) - 2$$

$$= \sum_{y=0}^{\infty} ky(y - 1)(y - 2)p^{y} + 3$$

$$= \sum_{y=0}^{\infty} kp^{3} \left(\frac{d^{3}}{dp^{3}}p^{y}\right) + 3$$

$$= kp^{3} \frac{d^{3}}{dp^{3}} \sum_{y=0}^{\infty} p^{y} + 3$$

$$= \frac{6kp^{3}}{(1-p)^{4}} + 3.$$

Since $V(Y) = E(Y^2) - E(Y)^2$, it follows that

$$V(Y) = E(Y^{2}) - E(Y)^{2}$$

$$= \frac{6kp^{3}}{(1-p)^{4}} + 3 - \frac{25}{9}$$

$$= \frac{8}{9}.$$

2. Verify the identity $\binom{n}{k} = \frac{n}{k} \binom{n-1}{k-1}$ and use it to show that $E[Y^k] = npE[(X+1)^{k-1}]$ where Y is a binomial random variable with parameters n and p and X is a binomial random variable with parameters n-1 and p.

Proof. We have that

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

$$= \frac{n(n-1)!}{k(k-1)!(n-k)!}$$

$$= \frac{n}{k} \frac{(n-1)!}{(k-1)!(n-k)!}$$

$$= \frac{n}{k} \frac{(n-1)!}{(k-1)!((n-1)-(k-1))!}$$

$$= \frac{n}{k} \binom{n-1}{k-1}.$$

Now

$$E[Y^k] = \sum_{y=0}^n y^k p(y)$$
 [Definition]

$$= \sum_{y=0}^n y^k \binom{n}{y} p^y (1-p)^{n-y}$$

$$= \sum_{y=1}^n y^k \binom{n}{y} p^y (1-p)^{n-y}$$

$$= \sum_{y=1}^n y^{k-1} n \binom{n-1}{y-1} p^y (1-p)^{n-y}$$

$$= np \sum_{y=1}^n y^{k-1} \binom{n-1}{y-1} p^{y-1} (1-p)^{n-y}$$

$$= np \sum_{x=0}^{n-1} (x+1)^{k-1} \binom{n-1}{x} p^x (1-p)^{(n-1)-x}$$
 [Let $y = x+1$]

$$= np \sum_{x=0}^{n-1} (x+1)^{k-1} p(x)$$

$$= np E[(X+1)^{k-1}],$$

which is what we wanted to show.

3. Using the recursion relation found in problem 2 for the binomial random variable with parameters n and p, find $E[Y^2]$ and then V(Y).

Solution. If we set k=2 in the formula in problem 2, we get

$$E[Y^{2}] = npE[X + 1]$$

$$= np(E[X] + E[1])$$

$$= np((n - 1)p + 1)$$

$$= (np)^{2} - np^{2} + np.$$

Thus

$$\begin{split} V(Y) &= E[Y^2] - E[Y]^2 \\ &= (np)^2 - np^2 + np - (np)^2 \\ &= np - np^2 \\ &= np(1-p). \end{split}$$

- 4. Using the identity from problem 2, show that
 - (a) if Y is a negative binomial random variable with parameters r and p, then

$$E[Y^k] = \frac{r}{p}E[(X-1)^{k-1}],$$

where X is a negative binomial random variable with parameters r + 1 and p.

(b) Use the relation in (a) to find E[Y] and V(Y).

Solution.

(a) From problem 2, we know that

$$\frac{1}{r} \binom{y-1}{r-1} = \frac{1}{y} \binom{y}{r};$$

thus

$$\begin{split} E[Y^k] &= \sum_{y=r}^{\infty} y^k p(y) \\ &= \sum_{y=r}^{\infty} y^k \binom{y-1}{r-1} p^r (1-p)^{y-r} \\ &= \frac{r}{p} \sum_{y=r}^{\infty} \sum_{y=r}^{\infty} y^k \binom{y-1}{r-1} p^r (1-p)^{y-r} \\ &= \frac{r}{p} \sum_{y=r}^{\infty} y^k \frac{1}{r} \binom{y-1}{r-1} p^{r+1} (1-p)^{y-r} \\ &= \frac{r}{p} \sum_{y=r}^{\infty} y^k \frac{1}{y} \binom{y}{r} p^{r+1} (1-p)^{y-r} \\ &= \frac{r}{p} \sum_{y=r}^{\infty} y^{k-1} \binom{y}{r} p^{r+1} (1-p)^{y-r} \\ &= \frac{r}{p} \sum_{x=r+1}^{\infty} (x-1)^{k-1} \binom{x-1}{r} p^{r+1} (1-p)^{(x-1)-r} \\ &= \frac{r}{p} \sum_{x=r+1}^{\infty} (x-1)^{k-1} p(x) \\ &= \frac{r}{p} E[(X-1)^{k-1}]. \end{split}$$

(b) If we set k=1 in (a), we immediately get that $E[Y]=\frac{r}{p}$. Now

$$\begin{split} V(Y) &= E[Y^2] - E[Y]^2 \\ &= \frac{r}{p} E[X-1] - \frac{r^2}{p^2} \\ &= \frac{r}{p} (E[X] - E[1]) - \frac{r^2}{p^2} \\ &= \frac{r}{p} \left(\frac{r+1}{p} - 1 \right) - \frac{r^2}{p^2} \\ &= \frac{r^2 + r}{p^2} - \frac{r}{p} - \frac{r^2}{p^2} \\ &= \frac{r - pr}{p^2} \\ &= \frac{r(1-p)}{p^2}. \end{split}$$

5. Using the identity from problem 2, show that if Y is a hypergeometric random variable with parameters N, r, and n, then

$$E[Y^k] = \frac{nr}{N}E[(X+1)^{k-1}],$$

where X is a hypergeometric random variable with parameters N-1, r-1, and n-1.

Proof. We have that

$$\begin{split} E[Y^k] &= \sum_{y=0}^n y^k p(y) \\ &= \sum_{y=1}^n y^k \frac{\binom{r}{y} \binom{N-r}{n-y}}{\binom{N}{n}} \\ &= \sum_{y=1}^n y^k \frac{\binom{r}{y} \binom{N-r}{n-y}}{\binom{N}{n}} \\ &= \sum_{y=1}^n y^k \frac{\frac{r}{y} \binom{r-1}{y-1} \binom{N-r}{n-y}}{\binom{N-1}{n-1}} \\ &= \sum_{y=1}^n y^k \frac{\frac{r}{y} \binom{r-1}{y-1} \binom{N-r}{n-y}}{\binom{N-1}{n-1}} \\ &= \frac{nr}{N} \sum_{y=1}^n y^{k-1} \frac{\binom{r-1}{y-1} \binom{N-r}{n-y}}{\binom{N-1}{n-1}} \\ &= \frac{nr}{N} \sum_{x=0}^{n-1} (x+1)^{k-1} \frac{\binom{r-1}{x} \binom{(N-1)-(r-1)}{(n-1)-x}}{\binom{N-1}{n-1}} \\ &= \frac{nr}{N} \sum_{x=0}^{n-1} (x+1)^{k-1} p(x) \\ &= \frac{nr}{N} E[(X+1)^{k-1}], \end{split}$$
 [Let $y = x+1$]

which is what we wanted to prove.

6. If Y is a hypergeometric random variable with parameters N, r, and n, use the recursion relation found in problem 5 to find E[Y] and V(Y).

Solution. Plugging in k=1 in the formula from problem 5 immediately shows us that

 $E[Y] = \frac{nr}{N}$. Using this same formula, we have that

$$\begin{split} V(Y) &= E[Y^2] - E[Y]^2 \\ &= \frac{nr}{N} E[X+1] - \left(\frac{nr}{N}\right)^2 \\ &= \frac{nr}{N} (E[X] + E[1]) - \left(\frac{nr}{N}\right)^2 \\ &= \frac{nr}{N} \left(\frac{(n-1)(r-1)}{N-1} + 1\right) - \left(\frac{nr}{N}\right)^2 \\ &= \frac{nr}{N} \left(\frac{(n-1)(r-1) + N - 1}{N-1}\right) - \left(\frac{nr}{N}\right)^2 \\ &= \frac{nr}{N} \left(\frac{(n-1)(r-1) + N - 1}{N-1} - \frac{nr}{N}\right) \\ &= \frac{nr}{N} \left(\frac{N(n-1)(r-1) + N^2 - N - nr(N-1)}{N(N-1)}\right) \\ &= \frac{nr}{N} \left(\frac{Nnr - Nn - Nr + N + N^2 - N - Nnr + nr}{N(N-1)}\right) \\ &= \frac{nr}{N} \left(\frac{-Nn - Nr + N^2 + nr}{N(N-1)}\right) \\ &= \frac{nr}{N} \left(\frac{N^2 - Nr - Nn + nr}{N(N-1)}\right) \\ &= \frac{nr}{N} \left(\frac{N(N-r) - n(N-r)}{N(N-1)}\right) \\ &= \frac{nr}{N} \left(\frac{N(N-r) - n(N-r)}{N(N-1)}\right). \end{split}$$

- 7. The number of chocolate chips in 1 cup of chocolate chip ice cream has a Poisson distribution with a mean of 10 chips per cup.
 - (a) What is the probability that a cup of chocolate chip ice cream has 9 chocolate chips.
 - (b) What is the probability that a half cup of chocolate chip ice cream has at least 3 chocolate chips?

Solution.

(a) We have that $\lambda = 10$, so that

$$P(Y=9) = p(9) = \frac{10^9}{9!}e^{-10} \approx 0.12511.$$

(b) Now $\lambda = 5$, so that

$$\begin{split} P(Y \ge 3) &= 1 - P(Y < 3) \\ &= 1 - p(0) - p(1) - p(2) \\ &= 1 - e^{-5} - 5e^{-5} - \frac{5^2}{2}e^{-5} \\ &\approx 0.875348. \end{split}$$