#### Lecture 2

- **Read:** Chapter 1.6-1.9
- Independence and Disjointness
- Pairwise Independence and Mutual Independence
- Counting
- Independent Trials
- Reliability Problems
- Multiple Outcomes

## Randomness and Probability

- Recall: We call a phenomenon random if individual outcomes are uncertain but there is a regular distribution of outcomes in a large number of repetitions
- The probability of any outcome of a random phenomenon is the proportion of times the outcome would occur in a very long series of repetitions

## Probabilities in a Finite Sample Space

- If the sample space is finite, each distinct event is assigned a probability
- The probability of an event is the sum of the probabilities of the distinct outcomes making up the event
- If a random phenomenon has k equally likely outcomes, each individual outcome has probability 1/k
- For any event A,

$$P[A] = \frac{\text{number of outcomes in } A}{\text{number of outcomes in } S}$$

## Example

- Roll a fair die and look at the face value
- Sample space:  $S = \{1, 2, 3, 4, 5, 6\}$
- This is a finite sample space, and each outcome is equally likely
- That is,

$$P[X=j]=1/6$$
 ,  $\forall j \in S$ 

where X is the face value of the die after rolling

- $P[X \ge 5] = P[X = 5] + P[X = 6] = 1/6 + 1/6 = 1/3$
- $P[X \le 2] = ?$

## Independent Events

• Events A and B are independent if and only if

$$P[AB] = P[A]P[B]$$

Always check this condition if you are asked about independence!

## Are Disjoint Events Independent?

**Example:** Suppose  $A \cap B = \emptyset$ 



- Then,  $P[A \cap B] = P[\emptyset] = 0$
- For the independence condition  $(P[A \cap B] = P[A] \cdot P[B])$  to hold, one of the events has to have probability 0

## Definition: Two Independent Events

**Definition 1.6:** Events A and B are independent if and only if

$$P[AB] = P[A]P[B]$$

.....

Equivalent definitions:

$$P[A|B] = P[A] \qquad \qquad P[B|A] = P[B]$$

## Definition: Three Independent Events

**Definition 1.7:**  $A_1$ ,  $A_2$ , and  $A_3$  are independent if and only if

- 1.  $A_1$  and  $A_2$  are independent
- 2.  $A_2$  and  $A_3$  are independent
- 3.  $A_1$  and  $A_3$  are independent
- 4.  $P[A_1 \cap A_2 \cap A_3] = P[A_1]P[A_2]P[A_3]$

### Independence and Disjointness

- Two events A and B are independent if knowing that one occurs does not change the probability that the other occurs. If A and B are independent, then  $P[A \cap B] = P[A]P[B]$
- Events A and B are disjoint if they have no outcomes in common  $(A \cap B = \emptyset)$ . If A and B are disjoint, then  $P[A \cup B] = P[A] + P[B]$
- If A and B are disjoint, then the fact A occurs tells us that B cannot occur. So, disjoint events are not independent
- Independence cannot be shown in a Venn diagram because it involves the probabilities of the events rather than just the outcomes that make up the events
- If A and B are independent, so are A and B<sup>c</sup> (Clearly, if B does not provide any information about A occurring, then neither does B<sup>c</sup>)



### Pairwise Independent Events

**Definition:** Two events A and B are independent if

$$P[A \cap B] = P[A] \cdot P[B]$$

• Recall: When  $P[A] \neq 0$  and  $P[B] \neq 0$ ,  $P[A \cap B] = P[A|B] \cdot P[B] = P[B|A] \cdot P[A]$ 

• So, equivalently, independence means:

$$P[A|B] = P[A]$$
 and  $P[B|A] = P[B]$ 

 $\Longrightarrow$  conditioning on an independent event does not change probability of the other

#### Remarks:

- $A \cap B = \emptyset$  does **not** mean A and B are independent (e.g., consider A and  $A^c$ )
- Independence is a property of events and of P



## Pairwise Independent Events: Example

**Example:** Consider an experiment involving two rolls of a four-sided die, each outcome is equally likely with probability 1/16

- 1.  $A_i = \{1 \text{st roll is i}\}, B_j = \{2 \text{nd roll is j}\} \text{ independent? (Looks right. Yes!)}$
- 2.  $A = \{1 \text{st roll is } 1\}, B = \{\text{sum is } 5\} \text{ independent? (Does not look obvious, but calculate. Yes!)}$
- A= {max. of two rolls is 2}, B= {min. of two rolls is 2} independent? (Looks unlikely. No!)

### Independence of a Collection of Events

• A collection  $A_1, A_2, ..., A_n$  of events are mutually independent if for any finite subset S of  $\{1, 2, ..., n\}$ ,

$$P[\cap_{i\in S}A_i]=\prod_{i\in S}P[A_i]$$

 Write out these conditions for three events: Three events are mutually independent if <u>all</u> these conditions hold:

$$P[A_1 \cap A_2 \cap A_3] = P[A_1]P[A_2]P[A_3]$$

$$P[A_1 \cap A_2] = P[A_1]P[A_2]$$

$$P[A_2 \cap A_3] = P[A_2]P[A_3]$$

$$P[A_1 \cap A_3] = P[A_1]P[A_3]$$

Warning: Pairwise independence 

→ mutual independence



## Independence of a Collection of Events: Example

#### Example: Two independent fair coin tosses

- $A = \{HT, HH\}$ : First toss is H
- $B = \{TH, HH\}$ : Second toss is H
- $C = \{TT, HH\}$ : First toss = second toss
- Verify they are pairwise independent, e.g., P[C|A] = P[C]
- Verify they are not mutually independent,  $P[A \cap B \cap C] \neq P[A] \cdot P[B] \cdot P[C]$

### Pairwise Independence and Mutual Independence

 Are there sets of random events which are pairwise independent but not mutually independent?

Suppose a box contains 4 tickets labeled by: 112 121 211 222 Choose one ticket at random and consider random events:

$$A_1 = \{1 \text{ occurs at the first place}\}$$

$$A_2 = \{1 \text{ occurs at the second place}\}$$

$$A_3 = \{1 \text{ occurs at the third place}\}$$

$$P[A_1] = 1/2, P[A_2] = 1/2, P[A_3] = 1/2$$

$$A_1A_2 = \{112\}, A_1A_3 = \{121\}, A_2A_3 = \{211\}$$

$$P[A_1A_2] = P[A_1A_3] = P[A_2A_3] = 1/4$$

So, we conclude that the three events  $A_1, A_2, A_3$  are pairwise independent

However, 
$$A_1A_2A_3 = \emptyset$$
  
 $P[A_1A_2A_3] = 0 \neq P[A_1]P[A_2]P[A_3] = (1/2)^3$ 

... Pairwise independence of a given set of random events does not imply that these events are mutually independent.



## Pairwise Independence and Mutual Independence (cont.)

- - Yes! The definition of mutual independence implies it!

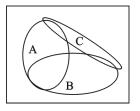
# Pairwise Independence and Mutual Independence (cont.)

• Suppose A, B, C are random events satisfying just the relation P[ABC] = P[A]P[B]P[C]. Does it follow that  $\overline{A}, \overline{B}, C$  are pairwise independent?

Toss two different standard dice: white and black. The sample space S of the outcomes consists of all ordered pairs  $(i, j), i, j = 1, ..., 6, S = \{(1, 1), (1, 2), ..., (6, 6)\}$ Each point in S has probability 1/36 $A_1 = \{ \text{first die} = 1, 2, \text{ or } 3 \}, A_2 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_3 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_4 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4, \text{ or } 5 \}, A_5 = \{ \text{first die} = 3, 4,$  $A_3 = \{\text{sum of faces is 9}\}\$  $A_1A_2 = \{31, 32, 33, 34, 35, 36\}, A_1A_3 = \{36\},$  $A_2A_3 = \{36, 45, 54\}, A_1A_2A_3 = \{36\}$  $P[A_1] = 1/2, P[A_2] = 1/2, P[A_3] = 1/9$  $P[A_1A_2A_3] = 1/36 = (1/2)(1/2)(1/9) = P[A_1]P[A_2]P[A_3]$  $P[A_1A_2] = 1/6 \neq 1/4 = P[A_1]P[A_2]$  $P[A_1A_3] = 1/36 \neq 1/18 = P[A_1]P[A_3]$  $P[A_2A_3] = 1/12 \neq 1/18 = P[A_2]P[A_3]$  $\therefore$  Just P[ABC] = P[A]P[B]P[C] does not mean mutual

## Conditioning May Affect Independence

• Assume A and B are independent:

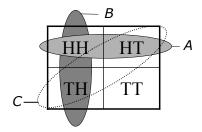


• If we are told that C occurred, are A and B independent?

## Conditioning May Affect Independence

#### Example 1:

- Two independent fair (p=1/2) coin tosses
- Event A: First toss is H
- Event B: Second toss is H
- P[A] = P[B] = 1/2



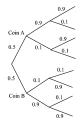
- Event C: The two outcomes are different.
- Conditioned on C, are A and B independent?



## Conditioning May Affect Independence

Example 2: Choose at random among two unfair coins A and B

- P[H|coin A] = 0.9 and P[H|coin B] = 0.1
- Keep tossing the chosen coin



- Given coin A was selected, are future tosses independent? (Yes!)
- If we do not know which coin it is, are future tosses independent?
  - Compare P[toss 11 = H] and  $P[\text{toss } 11 = H \mid \text{first } 10 \text{ tosses are } H]$  ( $\neq !$ )

## Conditionally Independent Events

**Definition:** Two events A and B are conditionally independent given an event C with probability  $P[C] \ge 0$  if

$$P[A \cap B|C] = P[A|C] \cdot P[B|C]$$

- We can show this is equivalent to  $P[A|B \cap C] = P[A|C]$ , i.e., if C is known, the additional information that B has occurred does not change the probability of A
- Remark: Independence does not imply conditional independence and vice versa!

## The King's Sibling

- The king comes from a family of two children
- What is the probability that his sibling is female?

## Using Independence in Modeling

**Reliability:** Complex systems are often simply modeled by assuming that they consists of several independent components

- Consider a network with n links
- For link i, event  $W_i = \text{link } i$  is operational, and is independent of other components
- Probability that a link i is working well:  $P[W_i] = p_i$

## Using Independence in Modeling (cont.)

- What is probability that the network is reliable for communications?
  - Series connection:  $P[\cap_i W_i] = p_1 p_2 \cdots p_n$ Overall reliability is worse than any of the components



■ Parallel connection:

 $P[\text{system is up}] = 1 - P[\cap_i W_i^c] = 1 - (1 - p_1)(1 - p_2) \cdots (1 - p_n)$ Reliability is better than any of the components, for small  $p_i$ , roughly sum



■ Mixture: Same story, can decompose the problem and reduce the series and parallel combinations

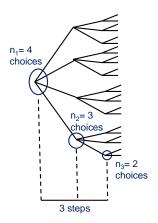


## Counting

**Motivation:** Discrete Uniform Law: Recall that for such a law, all sample points are equally likely, and computing probabilities reduces to just counting, i.e.,

$$P[A] = \frac{\text{number of elements of } A}{\text{total number of sample points}}$$

## Fundamental Principle of Counting



- r steps with  $n_i$  choices at each step i
- total number of choices =  $n_1 \times n_2 \times ... \times n_r$

### Fundamental Principle of Counting: Examples

- number of license plates with 3 letters followed by 4 digits =  $26 \cdot 26 \cdot 26 \cdot 10 \cdot 10 \cdot 10 \cdot 10$ 
  - if repetition is prohibited: 26 · 25 · 24 · 10 · 9 · 8 · 7
- number of subsets of a set with n elements =  $2^n$

#### Permutations

- k-permutation: an ordered sequence of k distinguishable objects
- $(n)_k$  = no. of possible k-permutations of n distinguishable objects:

$$(n)_k = n(n-1)(n-2)\cdots(n-k+1) = \frac{n!}{(n-k)!}$$

This follows from the fundamental counting principle

• Example: number of words with 4 distinct letters =

$$\frac{26!}{22!} = 26 \times 25 \times 24 \times 23$$

### Combinations

- k-combination:
  - Pick a subset of k out of n objects
  - Order of selection does not matter
  - Each subset is a k-combination
- Remark: In a combination, there is no ordering involved, e.g., 2-permutations of {A, B, C} are AB, AC, BA, CA, BC, CB, while the combinations of 2 out of the 3 letters would be AB, AC, BC

## How Many Combinations?

- $\binom{n}{k}$  = "n choose k" = no. of possible k-element subsets (i.e., order is not important) that can be obtained out of a set of n distinguishable objects
- To find  $\binom{n}{k}$ , we perform the following two subexperiments to assemble a k-permutation of n distinguishable objects:
  - 1. Choose a k-combination out of n objects.
  - 2. Choose a k-permutation of the k objects in the k-combination.

$$(n)_k = \binom{n}{k} k!$$
  $\binom{n}{k} = \frac{n!}{k!(n-k)!}$ 

## Example: Independent Trials and Binomial Probabilities

- n independent coin tosses, P[H] = p
- What is the probability of obtaining an *n*-sequence with *k* heads?
- $P[HTTHHH] = p^4(1-p)^2$
- P[sequence] =  $p^{\#heads} \cdot (1-p)^{\#tails}$

$$egin{aligned} P[k ext{ heads}] &= \sum_{k- ext{head seq.}} P[seq.] \ &= p^k \cdot (1-p)^{n-k} \cdot (\# ext{ of k-head seqs.}) \ &= inom{n}{k} p^k \cdot (1-p)^{n-k} \end{aligned}$$

#### **Partitions**

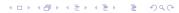
How many ways are there to divide a set of n distinct elements into r disjoint sets of  $n_1, n_2, ..., n_r$  elements each, with  $n_1, n_2, ..., n_r \leq n$ ?

- A combination of k elements out of n breaks up into two disjoint sets of elements of size k and n - k
- Note:  $\binom{n}{k} = \#$  of ways of breaking n elements into subsets of size k and n-k each
- Using the counting principle, we can answer the above question as

$$\binom{n}{n_1} \binom{n-n_1}{n_2} \binom{n-n_1-n_2}{n_3} \dots \binom{n-n_1-n_2-\dots-n_{r-1}}{n_r}$$

$$= \frac{n!}{n_1! n_2! \dots n_r!} = \binom{n}{n_1, n_2, \dots n_r}$$

This is called the multinomial coefficient



### Partitions: Example

- Anagrams: How many letter sequences can be obtained by rearranging the word TATTOO?
- Think of an anagram for this word as including 6 slots, which need to be filled with the letters T, A, or O
- Each solution corresponds to selecting 3 slots for T, 2 slots for O, and 1 slot for A
- How many ways are there to divide the 6 slots into such a partition?
- **Answer:**  $\frac{6!}{1!2!3!} = 60$
- Note: The letters are not distinct, but the slots are!



#### Problem 1.8.6

- A basketball team has
  - 3 pure centers, 4 pure forwards, 4 pure guards
  - one swingman who can play either guard or forward
- A pure player can play only the designated position
- How many lineups are there (1 center, 2 forwards, 2 guards)?

### Problem 1.8.6 Solution

#### Three possibilities:

- 1. swingman plays guard:  $N_1$  lineups
- 2. swingman plays forward:  $N_2$  lineups
- 3. swingman does not play:  $N_3$  lineups

$$\textit{N} = \textit{N}_1 + \textit{N}_2 + \textit{N}_3$$

# Problem 1.8.6 Solution (cont.)

$$N_1 = {3 \choose 1} {4 \choose 2} {4 \choose 1} = 72$$

$$N_2 = {3 \choose 1} {4 \choose 1} {4 \choose 2} = 72$$

$$N_3 = {3 \choose 1} {4 \choose 2} {4 \choose 2} = 108$$

### Multiple Outcomes

- *n* independent trials
- r possible trial outcomes  $(s_1, ..., s_r)$
- $\bullet \ P[\{s_k\}] = p_k$

## Multiple Outcomes (2)

- Outcome is a sequence:
  - Example:  $s_3 s_4 s_3 s_1$

$$P[s_3s_4s_3s_1] = p_3p_4p_3p_1 = p_1p_3^2p_4$$
  
=  $p_1^{n_1}p_2^{n_2}p_3^{n_3}p_4^{n_4}$ 

Probability depends on how many times each outcome occurred

# Multiple Outcomes (3)

 $N_i$  = no. of times  $s_i$  occurs

$$P[N_1 = n_1, ..., N_r = n_r] = M p_1^{n_1} p_2^{n_2} p_3^{n_3} p_4^{n_4}$$

$$M = Multinomial Coefficient$$

$$= \frac{n!}{n_1! n_2! ... n_r!}$$

## Service Facility Design

- c: service capacity of a facility
- n: # of customers it is assigned
- p: probability customer requires service
- *N* = number of customers requiring service
- Given n and p, choose c
- Criterion?
  - lacktriangle We establish a probability p such that P[N > c] < p
  - We choose the smallest c such that P[N > c] is still < p

## Card Play

- 52-card deck, dealt to 4 players, i.e., 13 cards each
- Find P[each gets an ace]
- Count size of the sample space: Partition the 52 card deck into 4 sets of 13 cards each

$$\binom{52}{13, 13, 13, 13} = \frac{52!}{13!13!13!13!}$$

 Count number of ways of distributing the four aces (One ace in each player's hand)

$$4 \times 3 \times 2$$

 Count number of ways of dealing the remaining 48 cards (Each hand has an additional 12 cards)

$$\binom{48}{12, 12, 12, 12} = \frac{48!}{12!12!12!12!}$$

• So,  $P[\text{each gets an ace}] = \frac{\# \text{ hands with one ace each}}{\# \text{ hands in sample space}} = \frac{4 \times 3 \times 2 \cdot \frac{48!}{12!12!12!12!}}{\frac{52!}{13!13!13!13!}}$ 

## N People

- *N* people in the class
- Each person can pick heads or tails
- A person will win if only  $\underline{1}$  person selects heads
  - $\rightarrow$  Ethernet network  $\rightarrow$  what is the optimal strategy?

#### Three Doors

- Three doors, one has a prize behind it
- You pick one, what is the chance you win?
- After you pick your door, I open one of the others and show you there is no prize behind it
- Should you change your choice? (Yes, you should change!)
- If so, how? What is the probability you win?