

## Day 2 – Joint modelling and Non-stationary processes

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## Day 2 goals

Today we focus on two ideas that come up constantly in disease mapping:

### 1. Joint modelling

- ▶ multiple related malaria outcomes
- ▶ shared and outcome-specific spatial structure
- ▶ *multiple likelihoods* (Binomial + Poisson)

### 2. Non-stationary spatial processes

- ▶ when “one Matérn everywhere” is not realistic
- ▶ two practical constructions:
  - ▶ mixture of SPDEs (smooth/rough)
  - ▶ **structured range** (range changes with a covariate)

## Why do we need joint models?

You often have **multiple imperfect views** of the same latent risk:

- ▶ prevalence surveys (direct infection measurements)
- ▶ facility case counts (incidence proxy)

## Part A — Joint modelling (same likelihood, two groups)

Example: two prevalence processes

Suppose we have:

- ▶ children <5 prevalence data
- ▶ pregnant women prevalence data

At each location  $s_i$ , group  $g \in \{c, p\}$ :

$$Y_{ig} \mid p_{ig} \sim \text{Binomial}(N_{ig}, p_{ig})$$

$$\text{logit}(p_{ig}) = \eta_{ig}$$

Decomposition: shared + group-specific spatial fields

A common and interpretable joint structure:

## Identifiability and scaling

Potential issue:  $S_0$  and  $S_g$  can compete.

Typical solutions:

- ▶ set priors so that group-specific fields are “smaller” than the shared field
- ▶ consider constraints or priors on variance:
  - ▶  $\sigma_{S_g}^2 < \sigma_{S_0}^2$  (probabilistically)

Conceptually: shared structure should capture the dominant signal.

What do we assume about  $S_0(s)$  and  $S_g(s)$ ?

A standard choice: **Matérn Gaussian random fields**

$$S(s) \sim \text{GRF}(0, \text{Matérn}(\rho, \sigma^2))$$

- ▶  $\rho$ : range (how quickly correlation decays)
- ▶  $\sigma^2$ : marginal variance

**Stationarity** here means  $\rho, \sigma^2$  are constant over space.

## Mesh and SPDE

- ▶ Mesh nodes: basis functions
- ▶ Field represented as:

$$S(s) \approx \sum_{k=1}^K w_k \phi_k(s)$$

where  $w_k$  are GMRF weights.

Joint model uses the same mesh for  $S_0$  and  $S_g$ , but:

- ▶  $S_0$ : one field
- ▶  $S_g$ : replicated by group

## Part B — Joint modelling with multiple likelihoods

Motivation: prevalence + case counts

Often you have:

- ▶ **Binomial** prevalence surveys:
  - ▶  $Y_i$  positives out of  $N_i$  tested
- ▶ **Poisson** facility case counts:
  - ▶  $C_i$  cases with exposure  $E_i$  (population, person-time, etc.)

Each dataset is incomplete on its own.

Joint modelling combines them while respecting their data-generating mechanisms.

## The two likelihoods

### Prevalence

$$Y_i \mid p_i \sim \text{Binomial}(N_i, p_i), \quad \text{logit}(p_i) = \eta_i^{(B)}$$

### Counts

$$C_i \mid \lambda_i \sim \text{Poisson}(E_i \lambda_i), \quad \log(\lambda_i) = \eta_i^{(P)}$$

Different link functions:

- ▶ logit for probabilities
- ▶ log for rates

## Linking them through shared latent structure

A flexible shared-structure joint model:

$$\eta_i^{(B)} = \beta_0^{(B)} + \beta^\top x_i + S_0(s_i) + S_B(s_i)$$



## Why include outcome-specific fields $S_B, S_P$ ?

Because the two processes are not identical:

- ▶ prevalence and incidence are related but not the same
- ▶ facility data can reflect access and reporting biases
- ▶ survey prevalence can reflect age structure and diagnostics

Outcome-specific fields mop up these systematic differences.

## What can go wrong?

Two common practical pitfalls:

### 1. **Confounding:**

- ▶ covariates and spatial fields compete

### 2. **Mis-specified exposure $E_i$ :**

- ▶ Poisson component can dominate if  $E_i$  is wrong scale

### 3. **Different spatial supports:**

- ▶ facility counts might be aggregated (administrative units)

## Part C — Non-stationarity

### Why stationarity can be unrealistic

Stationary Matérn assumes:

- ▶ same smoothness everywhere
- ▶ same correlation length everywhere

But malaria transmission may be:

- ▶ smooth in the north (broad ecological gradients)
- ▶ rough in the south (heterogeneous land use, urbanisation)
- ▶ different across ecozones

Non-stationarity: the covariance structure changes over space.

### Two practical nonstationary constructions

#### 1. Mixture of SPDEs (smooth + rough)

- ▶ simple, and robust

#### 2. Structured range in the SPDE

## Part C1 — Mixture of two SPDEs

Idea: blend a smooth and a rough field

Let:

- ▶  $S_{\text{smooth}}(s)$ : long-range Matérn
- ▶  $S_{\text{rough}}(s)$ : short-range Matérn
- ▶  $\omega(s) \in [0, 1]$ : spatial weight (e.g., function of latitude)

Define:

$$S(s) = \omega(s)S_{\text{smooth}}(s) + (1 - \omega(s))S_{\text{rough}}(s)$$

Then:

$$\eta(s) = \beta_0 + \beta^\top x(s) + S(s)$$

## What does $\omega(s)$ do?

- ▶ If  $\omega(s) \approx 1$ : field behaves like smooth long-range
- ▶ If  $\omega(s) \approx 0$ : field behaves like rough short-range
- ▶ If  $\omega(s)$  changes with location: correlation structure changes

**Interpretation** -  $\omega(s)$  is a *nonstationarity driver*.

## Choosing $\omega(s)$

Simple choices:

- ▶ monotone latitude function (north–south structure)
- ▶ ecozone indicator (piecewise structure)
- ▶ logistic function of a covariate:

$$\omega(s) = \text{logit}^{-1}(a + bz(s))$$

In the practical session:

- ▶ start with latitude-based  $\omega(s)$

## Part C2 — Structured range in the SPDE

Goal: let the Matérn range vary with a covariate

Recall: for Matérn SPDE models, range is approximately:

$$\rho(s) \approx \frac{\sqrt{8}}{\kappa(s)}$$

So if we model:

$$\log \kappa(s) = \theta_0 + \theta_1 z(s)$$

then

$\rho(s)$  varies smoothly with  $z(s)$ .

What does the sign of  $\theta_1$  mean?

## Why use mesh-vertex covariates?

In the SPDE approximation:

- ▶ the field is defined through weights on mesh vertices
- ▶ so nonstationary parameters need to be specified at the same support (vertices)

Thus  $z(s)$  is evaluated at mesh nodes:

- ▶  $z_k = z(\text{vertex}_k)$

This makes the range a spatially varying parameter.

## What changes compared to stationary SPDE?

Stationary SPDE:

- ▶ one  $\kappa$ , one  $\tau$  for all space

Structured-range SPDE:

- ▶  $\kappa(s)$  and  $\tau(s)$  can depend on covariates
- ▶ introduces additional parameters  $\theta$  controlling how they vary

Conceptual result:

## Practical guidance for modelling

- ▶ Start with a simple covariate  $z(s)$  (e.g. northing or ecozone index)
- ▶ Scale it to have mean 0, sd 1
- ▶ Use weak-to-moderate priors on the slope parameter to avoid extreme range variation

In interpretation:

- ▶ show the covariate map
- ▶ show the predicted prevalence map
- ▶ (optionally) show a derived “implied range” surface

## Part D — Examples and discussion

### Example 1: Joint model outputs to discuss

For the joint prevalence model: - How similar are children vs pregnant spatial patterns?  
- Is the shared field dominant? - Where do groups diverge?

Discussion prompts: - Are differences biological, behavioural, or sampling-related? -  
Could this be explained by covariates?



## Example 2: Multi-likelihood model outputs to discuss

For binomial + poisson: - Does adding case counts sharpen predictions? - Do we see evidence of strong coupling (( ))? - Where do outcome-specific fields absorb discrepancies?

Discussion prompts: - What might cause facility counts to disagree with survey prevalence?

### Example 3: Nonstationary mixture outputs to discuss

- ▶ Where does the model behave smoothly vs roughly?
- ▶ Does (s) align with your ecological intuition?
- ▶ How sensitive is the result to the choice of (s)?

## Example 4: Structured range outputs to discuss

### Summary

**Joint models** - share spatial information across related outcomes - can handle: - multiple groups (same likelihood) - multiple likelihoods (binomial + poisson)

**Nonstationarity** - mixture SPDE: simple and robust - structured range SPDE: mechanistic, covariate-driven

**Take-away** - choose the simplest model that answers your scientific question - use nonstationarity when stationary assumptions are visibly violated

