

Ve406 Lecture 22

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- In factor analysis, we have implicitly assumed the latent factor is continuous

$$\mathbf{X}_{n \times k} = \mathbf{F}_{n \times 1} \mathbf{W}_{k \times 1}^T + \boldsymbol{\epsilon}_{n \times k}$$

- What if the latent variables are not continuous but categorical? e.g.

$$Z \sim \text{Binomial}(1, 0.2)$$

$$X \mid z = 0 \sim \text{Normal}(0, 9)$$

$$X \mid z = 1 \sim \text{Normal}(1, 9)$$

which corresponds to

$$\mathbf{X}_{n \times 1} = \mathbf{Z}_{n \times 1} \mathbf{W}_{1 \times 1}^T + \boldsymbol{\epsilon}_{n \times 1}$$

- More generally, the following is known as a **Gaussian mixture model**.

$$Z \sim \text{Multinomial}(1, \mathbf{p})$$

$$X \mid z = k \sim \text{Normal}(\mu_k, \sigma_k^2)$$

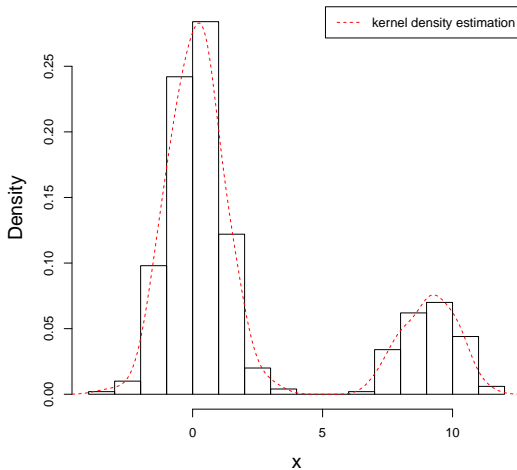
where \mathbf{p} is a vector probabilities, $\mathbf{1}^T \mathbf{p} = 1$, known as the **mixing proportions**.

- In general, a mixture model assumes x_i are generated in the following fashion

```
> set.seed(2)
> n = 500
> z.vec = rbinom(n, size = 1, prob = 0.2)
> # z is latent, thus not observed in practice

> # x is observed data
> x.vec = double(n)
>
> for (i in 1:n){
+   if (z.vec[i] == 0) {
+     x.vec[i] = rnorm(1, mean = 0, sd = 1)
+   } else {
+     x.vec[i] = rnorm(1, mean = 9, sd = 1)
+   }
+ }
```

Distribution of observed x



- In terms of probability density function, it means

$$f_{X,Z}(x, z) = f_Z(z)f_{X|Z}(x | z)$$

- Consider the marginal distributions,

$$f_X(x) = \sum_z f_Z(z)f_{X|Z}(x | z) = \sum_z \Pr(Z = z)f_{X|Z}(x | z)$$

- Since the following formula is also true for the joint density

$$f_{X,Z}(x, z) = f_Z(z)f_{X|Z}(x | z) = f_X(x)f_{Z|X}(z | x)$$

we have

$$f_{Z|X}(z | x) = \frac{f_Z(z)f_{X|Z}(x | z)}{f_X(x)} = \frac{f_Z(z)f_{X|Z}(x | z)}{\sum_z \Pr(Z = z)f_{X|Z}(x | z)}$$

Q: What is the significance of this formula?

- Of course, the parameters are unknown and need to be estimated in practice,

$$Z \sim \text{Binomial}(1, p)$$

$$X \mid z = 0 \sim \text{Normal}(\mu_0, \sigma_0^2)$$

$$X \mid z = 1 \sim \text{Normal}(\mu_1, \sigma_1^2)$$

- The MLE of $\theta^T = [p, \mu_0, \mu_1, \sigma_0, \sigma_1]$ cannot be computed in a closed form.

$$\begin{aligned} f_X(x; p, \mu_0, \mu_1, \sigma_0, \sigma_1) &= \sum_z \Pr(Z = z) f_{X|Z}(x \mid z) \\ &= (1 - p) \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp\left(-\frac{(x_i - \mu_0)^2}{2\sigma_0^2}\right) \\ &\quad + p \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left(-\frac{(x_i - \mu_1)^2}{2\sigma_1^2}\right) \end{aligned}$$

Q: What is the likelihood function if we assume observations are independent?

- We can numerically find MLE,

```
> obj_func = function(theta, x) {  
+   p       = theta[1]  
+   p0      = 1 - p  
+   mu0     = theta[2]  
+   mu1     = theta[3]  
+   sigma0  = theta[4]  
+   sigma1  = theta[5]  
+  
+   pdf      = p0 * dnorm(x, mu0, sigma0) +  
+     p * dnorm(x, mu1, sigma1)  
+  
+   res      = -sum(log(pdf))  
+   return(res)  
+  
+ }  
  
> res.nlm = nlm( # Newton based numerical methods  
+   obj_func, c(.25, 10, 10, 10, 10), x.vec)
```

```
> res.nlm # True parameters [0.2, 0, 9, 1, 1]
```

```
$estimate
```

```
[1] 0.2159996 0.1487675 8.9709822 1.0059598 0.9431263
```

```
> p = res.nlm$estimate[1]; p0 = 1 - p
```

```
> mu0 = res.nlm$estimate[2]
```

```
> mu1 = res.nlm$estimate[3]
```

```
> sigma0 = res.nlm$estimate[4]
```

```
> sigma1 = res.nlm$estimate[5]
```

```
> x.p = seq(min(x.vec), max(x.vec), length = 200)
```

```
>
```

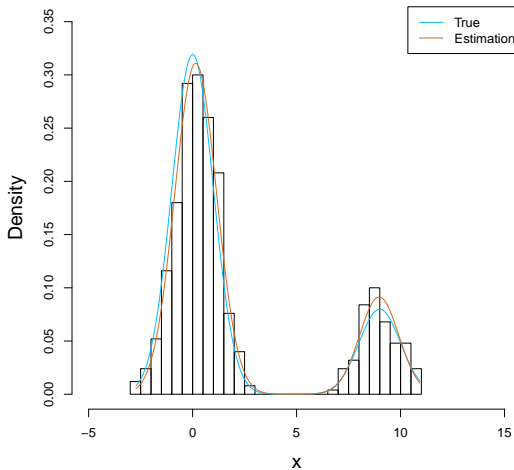
```
> est.pdf = p0 * dnorm(x.p, mu0, sigma0) +
```

```
+ p * dnorm(x.p, mu1, sigma1)
```

```
> true.pdf = 0.8 * dnorm(x.p, 0, 1) +
```

```
+ 0.2 * dnorm(x.p, 9, 1)
```


Distribution of observed x



- Of course, R has a package and a function for mixture models

```
> library(mixtools)
> res = normalmixEM(x.vec)
> res
```

```
$lambda
[1] 0.784 0.216

$mu
[1] 0.148769 8.970986

$sigma
[1] 1.0059601 0.9431269
```

which is more stable and faster than general purpose minimisation routines.

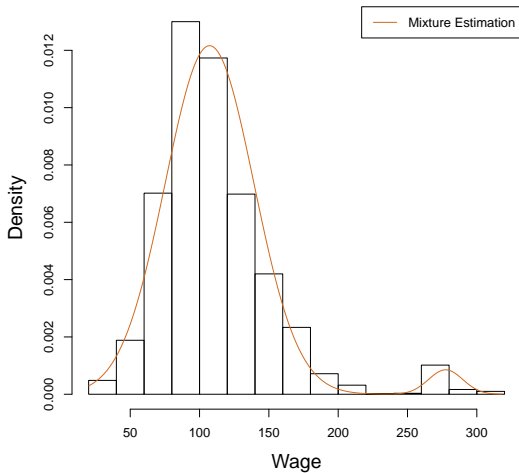
Q: Now we know how to estimate parameters in a mixture model, what is the significance of it in terms of regression?

- Let us finish lectures for this semester by looking at my archenemy again!

wage	Raw wage in the Mid-Atlantic region
age	Age of the worker
year	The year that wage information was recorded
education	A factor with levels: <ol style="list-style-type: none">1. < HS Grad2. HS Grad3. Some College4. College Grad5. Advanced Degree

```
> library(ISLR);  
> wage.df =  
+   Wage[, c("year", "age", "education", "wage")]  
>  
> res = normalmixEM(wage.df$wage)
```

Distribution of Wage



```
> str(wage.df)
```

```
'data.frame': 3000 obs. of 4 variables:  
 $ year      : int  2006 2004 2003 2003 2005 2008 2009 2008 2006 2004 ...  
 $ age       : int  18 24 45 43 50 54 44 30 41 52 ...  
 $ education: Factor w/ 5 levels "1. < HS Grad",...: 1 4 3 4 2 4 3 3 3 2 ...  
 $ wage      : num  75 70.5 131 154.7 75 ...
```

```
> X = as.matrix(wage.df[, 1:2])
```

```
>
```

```
> X = cbind(X, as.integer(wage.df$education))
```

```
>
```

```
> wage.gauss.MM = regmixEM(wage.df[,4], X, k=2)
```

```
> summary(wage.gauss.MM)
```

```
summary of regmixEM object:  
      comp 1      comp 2  
lambda 7.48706e-02 0.925129  
sigma  5.42237e+01 25.622410  
beta1 -2.99259e+03 -2120.478107  
beta2  1.48038e+00  1.081827  
beta3  1.72743e+00  0.484611  
beta4  3.92691e+01 12.019221  
loglik at estimate: -14510.38
```

The end of Ve406!