

Reproduire les résultats de Harvest models of small populations of a large carnivore using Bayesian forecasting par Andrén et al. 2020

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02/09/2020

Motivation

Reproduire pour comprendre les résultats de H., Andrén, Hobbs, N. T., Aronsson, M., Brøseth, H., Chapron, G., Linnell, J. D. C., Odden, J., Persson, J., and Nilsen, E. B.. 2020. Harvest models of small populations of a large carnivore using Bayesian forecasting. *Ecological Applications* 30(3):02063. 10.1002/eap.2063.

Les données sont disponibles, mais pas le code. Haha, Erlend le dernier auteur est un fervent défenseur de la science reproductible, c'est loupé sur ce coup-là. Je suppose que les analyses ont été faites par Hobbs, qui a fait plusieurs papiers avec un modèle approchant. Voir par exemple :

- Raiho AM, Hooten MB, Bates S, Hobbs NT (2015) Forecasting the Effects of Fertility Control on Overabundant Ungulates: White-Tailed Deer in the National Capital Region. PLoS ONE 10(12): e0143122. doi:10.1371/journal.pone.0143122
- Hobbs, N.T., Andrén, H., Persson, J., Aronsson, M. and Chapron, G. (2012), Native predators reduce harvest of reindeer by Sámi pastoralists. *Ecological Applications*, 22: 1640-1654. doi:10.1890/11-1309.1
- Ketz, A. C., T. L. Johnson, R. J. Monello, and N. T. Hobbs. 2016. Informing management with monitoring data: the value of Bayesian forecasting. *Ecosphere* 7(11):e01587. <10.1002/ecs2.1587>

Données

On récupère les données de monitoring et harvest pour le lynx. Les colonnes sont : * year – the year of census (February) * run – the run in the data * country – code for country; S = Sweden and N = Norway * region – code for management region; Z = Jämtland, Y = Västerbotten, AC = Västerbotten, BD = Norrbotten, 2 – 8 = the different large carnivore management regions in Norway (2 – 8) * census – number of lynx family groups censused in that year in that region * harvest – total number of lynx harvested in that year in that region * harvest_F_>1yr – number of females older than one year harvested in that year in that region * harvest_F_kitten – number of female kittens (10 months old) harvested in that year in that region

```
dat <- read.csv("eap2063-sup-0003-datas1.csv")
dat
```

##	year	run	country	region	census	harvest	harvest_F_.1yr	harvest_F_kitten
## 1	1998	1	S	Z	82	44	15	1
## 2	1999	2	S	Z	84	30	14	2
## 3	2000	3	S	Z	63	52	14	7
## 4	2001	4	S	Z	49	39	12	7

## 5	2002	5	S	Z	44	31	8	5
## 6	2003	6	S	Z	39	19	3	1
## 7	2004	7	S	Z	32	17	4	2
## 8	2005	8	S	Z	42	8	2	0
## 9	2006	9	S	Z	44	16	6	3
## 10	2007	10	S	Z	42	16	5	3
## 11	2008	11	S	Z	53	26	7	6
## 12	2009	12	S	Z	53	55	22	7
## 13	2010	13	S	Z	35	42	15	2
## 14	2011	14	S	Z	39	59	24	6
## 15	2012	15	S	Z	31	18	5	3
## 16	2013	16	S	Z	14	9	3	1
## 17	2014	17	S	Z	20	1	1	0
## 18	2015	18	S	Z	33	8	2	2
## 19	2016	19	S	Z	38	38	5	15
## 20	2017	20	S	Z	36	36	9	4
## 21	1998	1	S	Y	41	13	4	1
## 22	1999	2	S	Y	37	16	6	2
## 23	2000	3	S	Y	30	13	5	3
## 24	2001	4	S	Y	28	8	2	2
## 25	2002	5	S	Y	20	5	3	0
## 26	2003	6	S	Y	19	6	3	0
## 27	2004	7	S	Y	7	1	0	0
## 28	2005	8	S	Y	14	0	0	0
## 29	2006	9	S	Y	11	2	0	0
## 30	2007	10	S	Y	12	0	0	0
## 31	2008	11	S	Y	16	0	0	0
## 32	2009	12	S	Y	17	4	2	0
## 33	2010	13	S	Y	18	8	3	1
## 34	2011	14	S	Y	24	12	2	1
## 35	2012	15	S	Y	26	7	2	0
## 36	2013	16	S	Y	14	8	2	0
## 37	2014	17	S	Y	16	6	3	1
## 38	2015	18	S	Y	16	2	0	0
## 39	2016	19	S	Y	18	12	3	2
## 40	2017	20	S	Y	19	9	2	0
## 41	1998	1	S	AC	36	5	1	0
## 42	1999	2	S	AC	36	7	2	0
## 43	2000	3	S	AC	34	18	6	1
## 44	2001	4	S	AC	38	15	8	2
## 45	2002	5	S	AC	29	16	8	2
## 46	2003	6	S	AC	24	7	3	0
## 47	2004	7	S	AC	21	8	3	0
## 48	2005	8	S	AC	31	4	0	0
## 49	2006	9	S	AC	31	2	0	0
## 50	2007	10	S	AC	23	6	2	0
## 51	2008	11	S	AC	37	7	2	0
## 52	2009	12	S	AC	41	23	8	1
## 53	2010	13	S	AC	28	13	6	1
## 54	2011	14	S	AC	32	11	4	1
## 55	2012	15	S	AC	34	26	6	4
## 56	2013	16	S	AC	22	34	10	2
## 57	2014	17	S	AC	13	7	2	0
## 58	2015	18	S	AC	13	3	1	1

##	59	2016	19	S	AC	27	12	4	1
##	60	2017	20	S	AC	28	15	7	0
##	61	1998	1	S	BD	35	3	1	0
##	62	1999	2	S	BD	37	4	1	0
##	63	2000	3	S	BD	23	1	1	0
##	64	2001	4	S	BD	34	1	0	0
##	65	2002	5	S	BD	39	0	0	0
##	66	2003	6	S	BD	32	0	0	0
##	67	2004	7	S	BD	25	0	0	0
##	68	2005	8	S	BD	32	2	1	0
##	69	2006	9	S	BD	23	1	0	1
##	70	2007	10	S	BD	24	2	0	0
##	71	2008	11	S	BD	33	0	0	0
##	72	2009	12	S	BD	37	2	0	0
##	73	2010	13	S	BD	33	19	8	2
##	74	2011	14	S	BD	44	13	5	1
##	75	2012	15	S	BD	43	28	13	4
##	76	2013	16	S	BD	30	25	7	4
##	77	2014	17	S	BD	17	8	3	1
##	78	2015	18	S	BD	32	8	2	2
##	79	2016	19	S	BD	28	14	3	2
##	80	2017	20	S	BD	26	23	6	3
##	81	1996	1	N	2	14	15	1	2
##	82	1997	2	N	2	20	18	4	2
##	83	1998	3	N	2	14	29	8	4
##	84	1999	4	N	2	20	21	7	2
##	85	2000	5	N	2	12	18	5	3
##	86	2001	6	N	2	13	16	7	0
##	87	2002	7	N	2	9	14	6	2
##	88	2003	8	N	2	4	15	4	2
##	89	2004	9	N	2	7	7	2	1
##	90	2005	10	N	2	13	10	3	2
##	91	2006	11	N	2	13	6	2	2
##	92	2007	12	N	2	13	10	4	1
##	93	2008	13	N	2	14	22	4	2
##	94	2009	14	N	2	19	27	8	2
##	95	2010	15	N	2	17	28	10	3
##	96	2011	16	N	2	14	26	9	2
##	97	2012	17	N	2	16	16	4	2
##	98	2013	18	N	2	16	23	9	2
##	99	2014	19	N	2	16	29	11	5
##	100	2015	20	N	2	16	37	9	3
##	101	2016	21	N	2	9	21	8	0
##	102	2017	22	N	2	9	5	3	0
##	103	1996	1	N	3	1	4	0	1
##	104	1997	2	N	3	3	5	0	0
##	105	1998	3	N	3	2	11	4	0
##	106	1999	4	N	3	3	14	3	1
##	107	2000	5	N	3	5	9	2	2
##	108	2001	6	N	3	5	10	6	1
##	109	2002	7	N	3	7	12	5	2
##	110	2003	8	N	3	3	5	3	0
##	111	2004	9	N	3	3	1	1	0
##	112	2005	10	N	3	6	1	1	0

##	113	2006	11	N	3	5	3	2	0
##	114	2007	12	N	3	6	6	4	1
##	115	2008	13	N	3	5	11	4	2
##	116	2009	14	N	3	6	10	2	3
##	117	2010	15	N	3	4	9	5	0
##	118	2011	16	N	3	4	11	3	0
##	119	2012	17	N	3	5	5	0	1
##	120	2013	18	N	3	7	8	2	0
##	121	2014	19	N	3	5	11	2	1
##	122	2015	20	N	3	7	9	1	0
##	123	2016	21	N	3	3	6	3	0
##	124	2017	22	N	3	5	4	1	0
##	125	1996	1	N	4	2	0	0	0
##	126	1997	2	N	4	3	0	0	0
##	127	1998	3	N	4	6	0	0	0
##	128	1999	4	N	4	6	10	2	2
##	129	2000	5	N	4	1	11	2	1
##	130	2001	6	N	4	5	7	2	2
##	131	2002	7	N	4	5	11	6	2
##	132	2003	8	N	4	5	5	1	0
##	133	2004	9	N	4	6	7	3	0
##	134	2005	10	N	4	7	4	3	1
##	135	2006	11	N	4	6	6	3	1
##	136	2007	12	N	4	6	5	2	0
##	137	2008	13	N	4	5	7	0	2
##	138	2009	14	N	4	7	6	2	0
##	139	2010	15	N	4	9	6	2	0
##	140	2011	16	N	4	6	11	4	0
##	141	2012	17	N	4	5	6	1	0
##	142	2013	18	N	4	1	3	1	0
##	143	2014	19	N	4	5	0	0	0
##	144	2015	20	N	4	4	2	0	0
##	145	2016	21	N	4	1	0	0	0
##	146	2017	22	N	4	3	0	0	0
##	147	1996	1	N	5	9	10	2	1
##	148	1997	2	N	5	7	9	3	0
##	149	1998	3	N	5	11	14	6	1
##	150	1999	4	N	5	11	15	3	2
##	151	2000	5	N	5	6	12	5	1
##	152	2001	6	N	5	9	12	5	2
##	153	2002	7	N	5	8	14	6	3
##	154	2003	8	N	5	7	17	8	4
##	155	2004	9	N	5	8	9	0	1
##	156	2005	10	N	5	7	12	4	0
##	157	2006	11	N	5	10	7	3	0
##	158	2007	12	N	5	11	9	3	1
##	159	2008	13	N	5	10	11	5	0
##	160	2009	14	N	5	9	14	3	1
##	161	2010	15	N	5	9	10	3	0
##	162	2011	16	N	5	11	9	3	1
##	163	2012	17	N	5	6	8	5	0
##	164	2013	18	N	5	5	5	1	0
##	165	2014	19	N	5	4	0	0	0
##	166	2015	20	N	5	2	0	0	0

##	167	2016	21	N	5	7	0	0	0
##	168	2017	22	N	5	9	0	0	0
##	169	1996	1	N	6	20	34	10	4
##	170	1997	2	N	6	26	40	10	2
##	171	1998	3	N	6	14	32	12	1
##	172	1999	4	N	6	14	14	4	3
##	173	2000	5	N	6	14	15	2	2
##	174	2001	6	N	6	9	7	3	2
##	175	2002	7	N	6	11	17	3	4
##	176	2003	8	N	6	11	9	2	1
##	177	2004	9	N	6	14	3	2	0
##	178	2005	10	N	6	14	14	4	4
##	179	2006	11	N	6	17	18	6	2
##	180	2007	12	N	6	15	29	6	4
##	181	2008	13	N	6	23	30	9	4
##	182	2009	14	N	6	26	36	16	8
##	183	2010	15	N	6	20	59	22	3
##	184	2011	16	N	6	18	52	16	4
##	185	2012	17	N	6	14	17	7	3
##	186	2013	18	N	6	8	15	6	0
##	187	2014	19	N	6	12	7	2	1
##	188	2015	20	N	6	17	18	6	0
##	189	2016	21	N	6	14	31	11	2
##	190	2017	22	N	6	19	33	11	5
##	191	1996	1	N	7	12	14	6	2
##	192	1997	2	N	7	14	16	4	3
##	193	1998	3	N	7	10	16	6	1
##	194	1999	4	N	7	16	11	5	0
##	195	2000	5	N	7	15	20	6	5
##	196	2001	6	N	7	5	16	6	2
##	197	2002	7	N	7	6	13	6	0
##	198	2003	8	N	7	5	7	4	0
##	199	2004	9	N	7	2	5	2	0
##	200	2005	10	N	7	4	2	1	0
##	201	2006	11	N	7	6	0	0	0
##	202	2007	12	N	7	8	0	0	0
##	203	2008	13	N	7	9	4	0	1
##	204	2009	14	N	7	14	8	4	1
##	205	2010	15	N	7	6	16	8	1
##	206	2011	16	N	7	8	12	4	1
##	207	2012	17	N	7	8	9	2	1
##	208	2013	18	N	7	10	6	4	0
##	209	2014	19	N	7	4	13	4	0
##	210	2015	20	N	7	5	4	2	0
##	211	2016	21	N	7	6	5	1	0
##	212	2017	22	N	7	6	6	1	0
##	213	1996	1	N	8	5	4	2	0
##	214	1997	2	N	8	7	5	1	1
##	215	1998	3	N	8	7	11	5	0
##	216	1999	4	N	8	5	1	1	0
##	217	2000	5	N	8	6	10	6	2
##	218	2001	6	N	8	6	13	7	1
##	219	2002	7	N	8	8	11	2	0
##	220	2003	8	N	8	10	4	1	0

##	221	2004	9	N	8	3	3	2	0
##	222	2005	10	N	8	3	1	0	0
##	223	2006	11	N	8	5	0	0	0
##	224	2007	12	N	8	12	1	1	0
##	225	2008	13	N	8	9	4	1	0
##	226	2009	14	N	8	9	8	4	0
##	227	2010	15	N	8	15	7	1	0
##	228	2011	16	N	8	11	16	7	1
##	229	2012	17	N	8	13	18	5	4
##	230	2013	18	N	8	10	13	3	2
##	231	2014	19	N	8	5	10	4	0
##	232	2015	20	N	8	8	5	1	2
##	233	2016	21	N	8	9	3	2	0
##	234	2017	22	N	8	6	4	2	1

```
dat %>%
  count(region)
```

```
##      region  n
## 1         2 22
## 2         3 22
## 3         4 22
## 4         5 22
## 5         6 22
## 6         7 22
## 7         8 22
## 8         AC 20
## 9         BD 20
## 10        Y 20
## 11        Z 20
```

```
dat %>%
  count(country)
```

```
##      country  n
## 1         N 154
## 2         S  80
```

Le modèle

Dans leur papier, Henrik et les collègues construisent un modèle démographique structuré en classes d'âge. J'ai pas envie de me lancer dans un truc compliqué, l'idée est simplement de comprendre comment dérouler leur approche.

On part sur un modèle exponentiel. On stipule que les effectifs N_t à l'année t sont obtenus à partir des effectifs à l'année $t - 1$ auxquels on a retranché les prélèvements H_{t-1} , le tout multiplié par le taux de croissance annuel λ :

$$N_t = \lambda(N_{t-1} - H_{t-1}).$$

Cette relation est déterministe. Pour ajouter de la variabilité démographique, on suppose que les effectifs sont distribués selon une distribution log-normale, autrement dit que les effectifs sont normalement distribués sur l'échelle log :

$$\log(N_t) \sim \text{Normale}(\mu_t, \sigma_{\text{proc}})$$

avec $\mu_t = \log(N_t) = \log(\lambda(N_{t-1} - H_{t-1}))$ et σ_{proc} l'erreur standard des effectifs sur l'échelle log. On aurait pu prendre une loi de Poisson à la place. La stochasticité environnementale est en général captée par le taux de croissance, mais pas ici puisqu'il est constant. C'est une hypothèse forte du modèle. Dans l'idéal, on pourrait coupler le modèle de capture-recapture, et le modèle qui décrit l'évolution des effectifs au cours du temps.

On ajoute une couche d'observation qui capture les erreurs sur les effectifs. Si l'on note y_t les effectifs observés, on suppose que ces comptages annuels sont distribués comme une loi de Poisson de moyenne les vrais effectifs N_t :

$$y_t \sim \text{Poisson}(N_t).$$

```
lynx_model <- function(){
  # Priors
  sigmaProc ~ dunif(0, 10)
  tauProc <- 1/sigmaProc^2
  lambda ~ dunif(0, 5)

  N[1] ~ dgamma(1.0E-6, 1.0E-6)

  # Process model
  for (t in 2:(nyears)) {
    mu[t] <- lambda * (N[t-1] - harvest[t-1])
    Nproc[t] <- log(max(1, mu[t]))
    N[t] ~ dlnorm(Nproc[t], tauProc)
  }

  # Observation model
  for (t in 1:nyears) {
    y[t] ~ dpois(N[t])
  }
}
```

Dans le papier, Henrik fait des regroupements d'aires de gestion, et applique le modèle à chacun de ces regroupements.

Northern Sweden: management areas Z, Y, BD and AC

On regroupe.

```
dat %>%
  filter(region == "Z" | region == "Y" | region == "BD" | region == "AC") %>%
  select(year, census, harvest) %>%
  group_by(year) %>%
  summarize(census = sum(census),
            harvest = sum(harvest)) -> dat1
```

```
## 'summarise()' ungrouping output (override with '.groups' argument)
```

On prépare les données.

```
bugs.data <- list(
  nyears = 20,
  y = dat1$census,
  harvest = dat1$harvest)
```

On précise les paramètres à estimer et le nombre de chaînes de MCMC (j'en prends trois ici).

```
bugs.monitor <- c("lambda", "sigmaProc", "N", "tauProc")
bugs.chains <- 3
bugs.inits <- function(){
  list(
  )
}
```

Allez zooh, on lance la machine!

```
lynx_mod <- jags(data = bugs.data,
  inits = bugs.inits,
  parameters.to.save = bugs.monitor,
  model.file = lynx_model,
  n.chains = bugs.chains,
  n.thin = 10,
  n.iter = 100000,
  n.burnin = 50000)
```

```
## module glm loaded
```

```
## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 20
##   Unobserved stochastic nodes: 22
##   Total graph size: 147
##
## Initializing model
```

Jetons un coup d'oeil aux estimations.

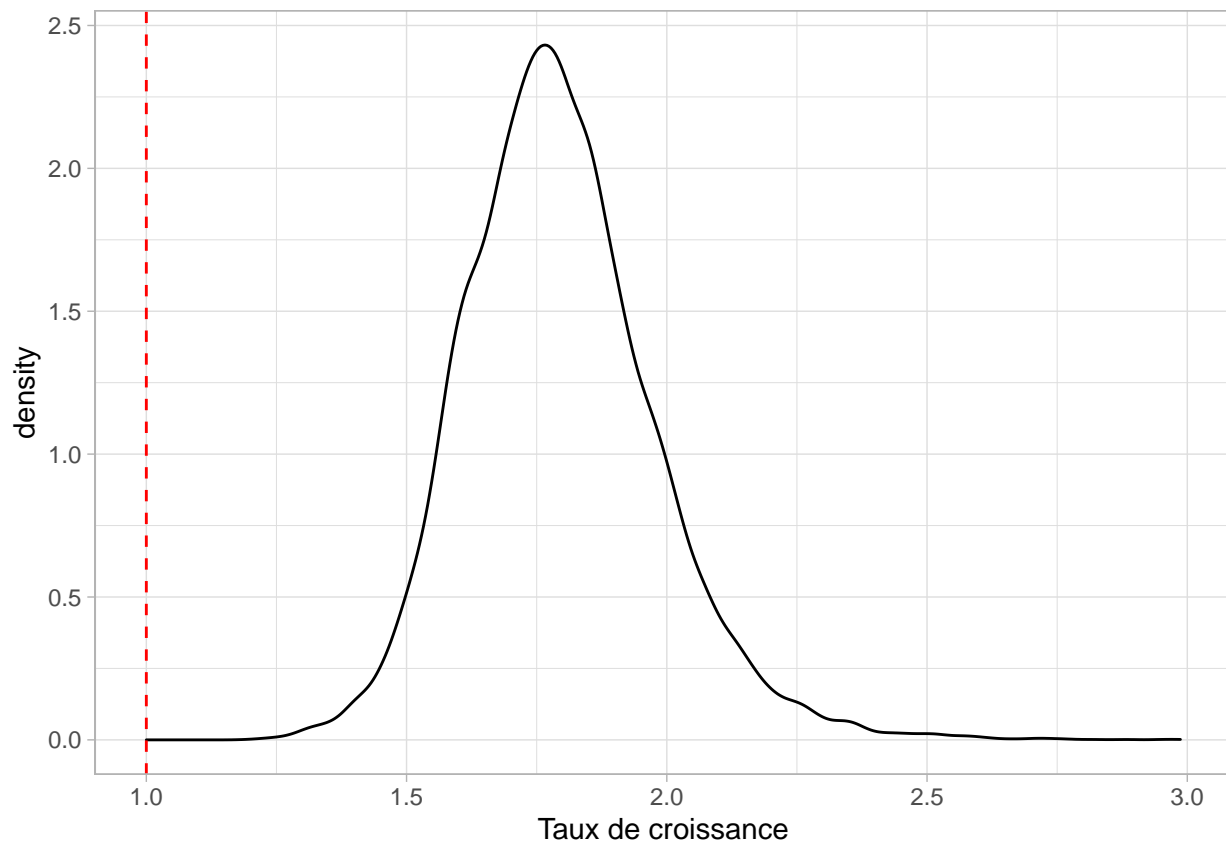
```
res <- print(lynx_mod, intervals = c(2.5/100, 50/100, 97.5/100))
```

```
## Inference for Bugs model at "/var/folders/r7/j0wqj1k95vz8w44sdxzm986c0000gn/T//RtmpZaM5Ca/model3dba1
## 3 chains, each with 1e+05 iterations (first 50000 discarded), n.thin = 10
## n.sims = 15000 iterations saved
##           mu.vect sd.vect   2.5%    50%   97.5%  Rhat n.eff
## N[1]      192.271  13.414 167.040 191.920 219.567 1.001 15000
## N[2]      190.865  13.282 165.508 190.511 217.776 1.001 15000
## N[3]      155.282  11.253 134.178 154.974 178.143 1.001 15000
## N[4]      146.586  11.284 125.655 146.203 169.886 1.001  8500
## N[5]      130.321  10.555 110.445 129.919 152.114 1.001 15000
## N[6]      111.142   9.915  92.568 110.801 131.532 1.001  6600
```



```
## N[7]      88.449   8.702  72.331  88.134 106.269 1.001 11000
## N[8]     115.140  10.323  95.866 114.772 136.367 1.001 15000
## N[9]     108.878   9.778  90.544 108.578 128.764 1.001 15000
## N[10]    103.584   9.424  86.196 103.325 122.888 1.001 15000
## N[11]    137.649  11.049 116.893 137.258 160.350 1.001 15000
## N[12]    150.902  10.993 130.452 150.571 173.215 1.001  4700
## N[13]    126.100   9.314 109.063 125.655 145.578 1.001 13000
## N[14]    144.183   9.788 126.119 143.828 164.323 1.001 14000
## N[15]    132.494   9.837 114.412 132.052 152.525 1.001 15000
## N[16]     97.744   6.134  86.743  97.446 110.526 1.001  5600
## N[17]     64.755   7.101  51.499  64.534  79.369 1.001  4400
## N[18]     92.172   8.890  75.482  91.880 110.408 1.001 15000
## N[19]    118.652   8.918 102.096 118.356 137.173 1.001 13000
## N[20]    106.611  10.240  87.608 106.297 127.726 1.001 15000
## lambda     1.793   0.182   1.477   1.781   2.192 1.001 15000
## sigmaProc   0.407   0.096   0.261   0.392   0.635 1.001  6200
## tauProc     7.036   3.169   2.481   6.503  14.721 1.001  6200
## deviance  154.850   6.456 144.156 154.241 169.152 1.001 10000
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 20.8 and DIC = 175.7
## DIC is an estimate of expected predictive error (lower deviance is better).
```

```
lynx_mod$BUGSoutput$sims.matrix %>%
  as_tibble() %>%
  # pivot_longer(cols = everything(), values_to = "value", names_to = "parameter") %>%
  # filter(str_detect(parameter, "lambda")) %>%
  ggplot() +
  aes(x = lambda) +
  geom_density() +
  geom_vline(xintercept = 1, lty = "dashed", color = "red") +
  labs(x = "Taux de croissance")
```

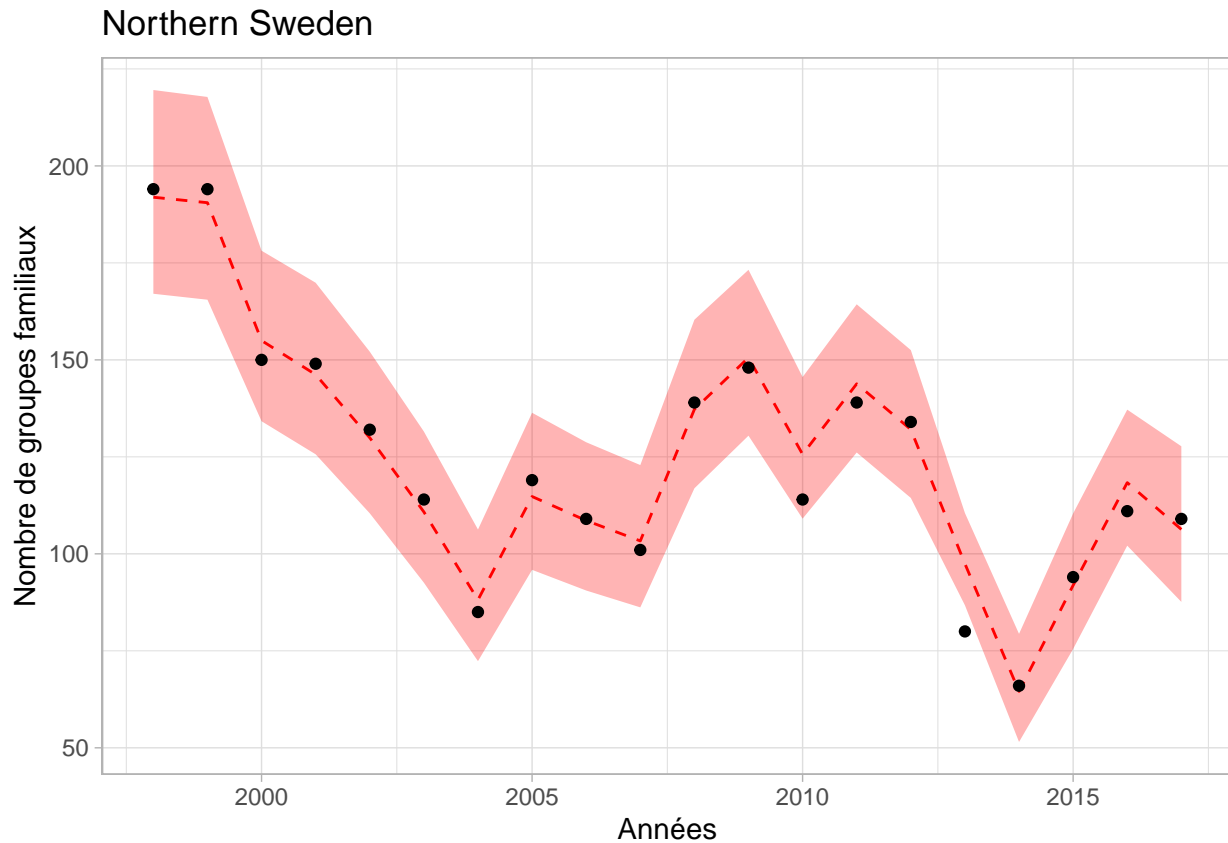


Ensuite les projections.

```
northern_sweden <- lynx_mod$BUGSoutput$sims.matrix %>%
  as_tibble() %>%
  pivot_longer(cols = everything(), values_to = "value", names_to = "parameter") %>%
  filter(str_detect(parameter, "N")) %>%
  group_by(parameter) %>%
  summarize(medianN = median(value),
            lci = quantile(value, probs = 2.5/100),
            uci = quantile(value, probs = 97.5/100)) %>%
  mutate(an = parse_number(parameter) + 1997) %>%
  arrange(an) %>%
  ggplot() +
  geom_ribbon(aes(x = an, y = medianN, ymin = lci, ymax = uci), fill = "red", alpha = 0.3) +
  geom_line(aes(x = an, y = medianN), lty = "dashed", color = "red") +
  # geom_point(aes(x = an, y = medianN), color = "red") +
  geom_point(data = bugs.data %>% as_tibble, aes(x = 1997 + 1:unique(nyears), y = y)) +
  labs(y = "Nombre de groupes familiaux",
       x = "Années",
       title = "Northern Sweden")
```

'summarise()' ungrouping output (override with '.groups' argument)

```
northern_sweden
```



Northern Norway: management areas 6, 7, 8

Idem qu'au-dessus.

```
dat %>%
  filter(region == "6" | region == "7" | region == "8") %>%
  select(year, census, harvest) %>%
  group_by(year) %>%
  summarize(census = sum(census),
            harvest = sum(harvest)) -> dat1
```

'summarise()' ungrouping output (override with '.groups' argument)

```
bugs.data <- list(
  nyears = 22,
  y = dat1$census,
  harvest = dat1$harvest)
```

On précise les paramètres à estimer et le nombre de chaînes de MCMC (j'en prends trois ici).

```
bugs.monitor <- c("lambda", "sigmaProc", "N", "tauProc")
bugs.chains <- 2
bugs.inits <- function(){
  list(
  )
}
```

Allez zooh, on lance la machine!

```
lynx_mod <- jags(data = bugs.data,
  inits = bugs.inits,
  parameters.to.save = bugs.monitor,
  model.file = lynx_model,
  n.chains = bugs.chains,
  n.thin = 10,
  n.iter = 100000,
  n.burnin = 50000)
```

```
## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 22
##   Unobserved stochastic nodes: 24
##   Total graph size: 161
##
## Initializing model
```

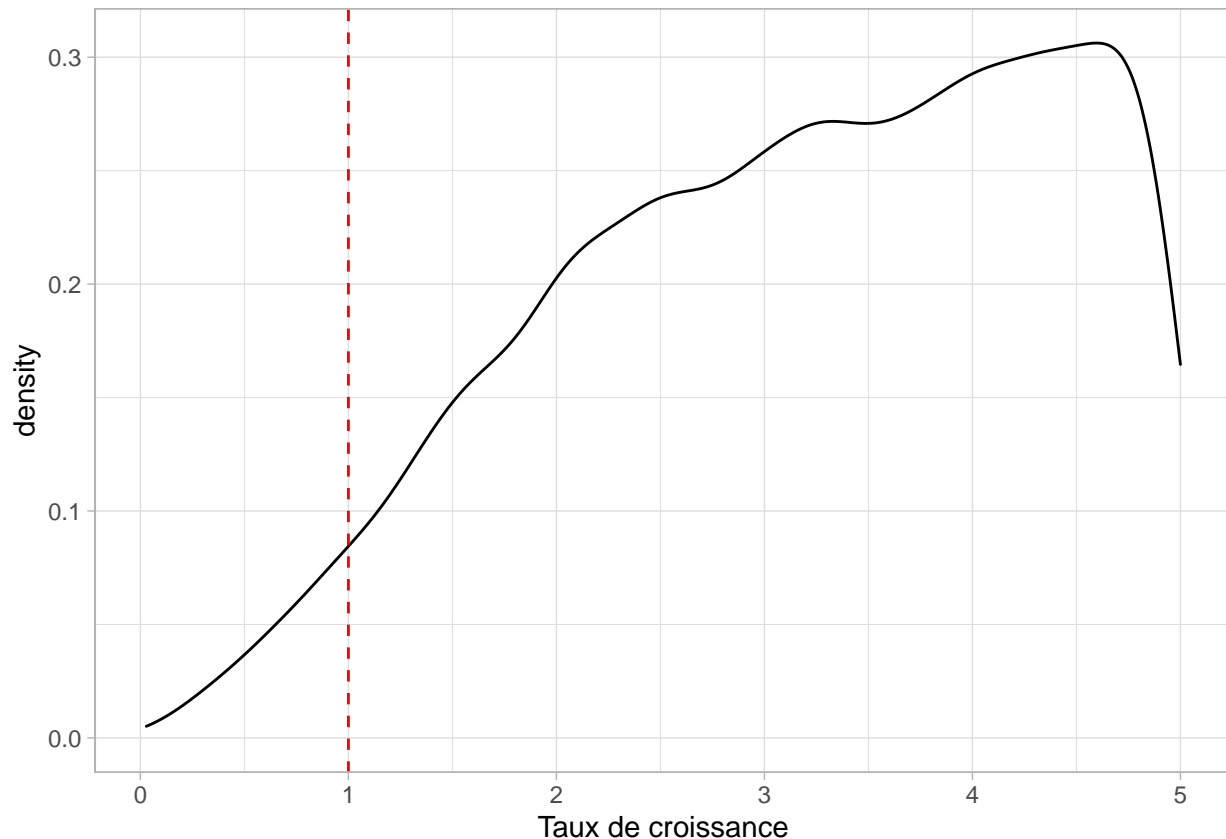
Jetons un coup d'oeil aux estimations.

```
res <- print(lynx_mod, intervals = c(2.5/100, 50/100, 97.5/100))
```

```
## Inference for Bugs model at "/var/folders/r7/j0wqj1k95vz8w44sdxzm986c0000gn/T//RtmpZaM5Ca/model13dba63"
## 2 chains, each with 1e+05 iterations (first 50000 discarded), n.thin = 10
## n.sims = 10000 iterations saved
##           mu.vect sd.vect   2.5%   50%  97.5%  Rhat n.eff
## N[1]      37.264   6.385  26.099  36.834  51.898 1.001  6500
## N[2]      46.855   7.105  34.343  46.369  62.795 1.001 10000
## N[3]      30.613   5.587  20.650  30.302  42.321 1.001 10000
## N[4]      35.179   5.565  25.424  34.757  47.311 1.001 10000
## N[5]      35.201   6.203  24.384  34.735  48.435 1.001 10000
## N[6]      19.745   4.494  11.799  19.451  29.423 1.001 10000
## N[7]      24.680   5.016  16.003  24.305  35.517 1.001 10000
## N[8]      26.087   4.819  17.325  25.676  36.728 1.001 10000
## N[9]      18.953   4.228  11.968  18.533  28.261 1.001 10000
## N[10]     21.842   4.328  13.876  21.569  31.263 1.001 10000
## N[11]     28.182   5.046  19.559  27.786  39.187 1.001  8500
## N[12]     36.322   5.488  25.565  36.003  47.720 1.001  3800
## N[13]     42.798   6.096  30.439  42.750  55.287 1.001 10000
## N[14]     50.767   7.484  36.764  50.834  65.441 1.001 10000
## N[15]     40.696   6.351  29.433  40.336  54.051 1.001 10000
## N[16]     36.548   6.124  25.575  36.152  49.751 1.001 10000
## N[17]     35.150   6.386  24.326  34.585  49.302 1.001 10000
## N[18]     28.266   5.736  18.261  27.706  40.430 1.001  8500
## N[19]     20.855   4.811  12.796  20.377  31.974 1.001 10000
## N[20]     30.910   5.288  20.892  30.744  41.994 1.001  8000
## N[21]     29.181   5.787  19.258  28.612  42.314 1.001 10000
## N[22]     30.512   5.544  20.777  30.149  42.309 1.001  9100
## lambda      3.199   1.171   0.834   3.308   4.927 1.001 10000
## sigmaProc   2.942   0.520   2.134   2.867   4.167 1.001  5100
```

```
## tauProc      0.126   0.042   0.058   0.122   0.220 1.001  5100
## deviance    138.283   6.758 127.124 137.653 153.386 1.001  3700
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 22.8 and DIC = 161.1
## DIC is an estimate of expected predictive error (lower deviance is better).
```

```
lynx_mod$BUGSoutput$sims.matrix %>%
  as_tibble() %>%
  # pivot_longer(cols = everything(), values_to = "value", names_to = "parameter") %>%
  # filter(str_detect(parameter, "lambda")) %>%
  ggplot() +
  aes(x = lambda) +
  geom_density() +
  geom_vline(xintercept = 1, lty = "dashed", color = "red") +
  labs(x = "Taux de croissance")
```



Ensuite les projections.

```
northern_norway <- lynx_mod$BUGSoutput$sims.matrix %>%
  as_tibble() %>%
  pivot_longer(cols = everything(), values_to = "value", names_to = "parameter") %>%
  filter(str_detect(parameter, "N")) %>%
  group_by(parameter) %>%
```

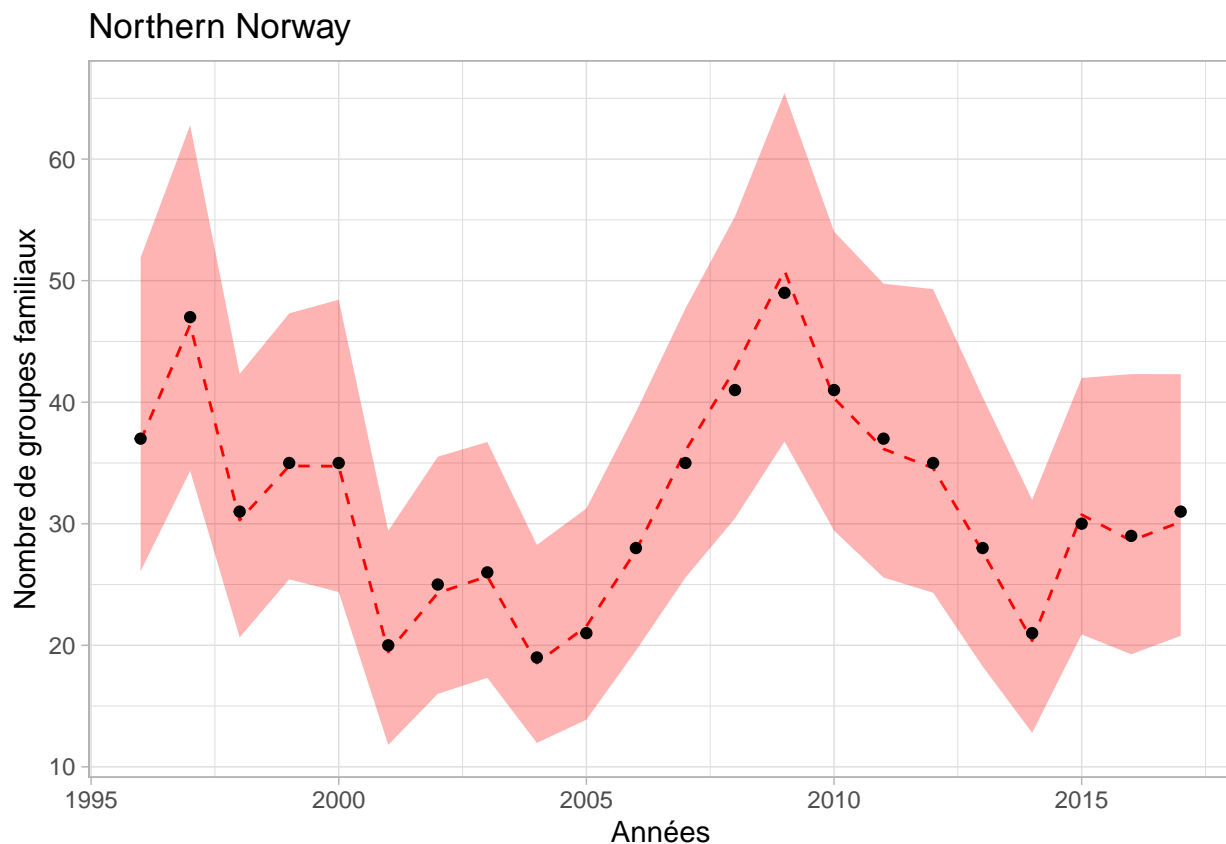
```

summarize(medianN = median(value),
          lci = quantile(value, probs = 2.5/100),
          uci = quantile(value, probs = 97.5/100)) %>%
mutate(an = parse_number(parameter) + 1995) %>%
arrange(an) %>%
ggplot() +
  geom_ribbon(aes(x = an, y = medianN, ymin = lci, ymax = uci), fill = "red", alpha = 0.3) +
  geom_line(aes(x = an, y = medianN), lty = "dashed", color = "red") +
  # geom_point(aes(x = an, y = medianN), color = "red") +
  geom_point(data = bugs.data %>% as_tibble, aes(x = 1995 + 1:unique(nyears), y = y)) +
  labs(y = "Nombre de groupes familiaux",
       x = "Années",
       title = "Northern Norway")

```

'summarise()' ungrouping output (override with '.groups' argument)

northern_norway



Southern Norway: management areas 2, 3, 4 and 5

On applique le modèle exponentiel au dernier regroupement.

```

dat %>%
  filter(region == "2" | region == "3" | region == "4" | region == "5") %>%

```

```
select(year, census, harvest) %>%
group_by(year) %>%
summarize(census = sum(census),
          harvest = sum(harvest)) -> dat1
```

'summarise()' ungrouping output (override with '.groups' argument)

```
bugs.data <- list(
  nyears = 22,
  y = dat1$census,
  harvest = dat1$harvest)
```

On précise les paramètres à estimer et le nombre de chaînes de MCMC (j'en prends trois ici).

```
bugs.monitor <- c("lambda", "sigmaProc", "N", "tauProc")
bugs.chains <- 3
bugs.inits <- function(){
  list(
  )
}
```

Allez zooh, on lance la machine!

```
lynx_mod <- jags(data = bugs.data,
                inits = bugs.inits,
                parameters.to.save = bugs.monitor,
                model.file = lynx_model,
                n.chains = bugs.chains,
                n.thin = 10,
                n.iter = 100000,
                n.burnin = 50000)
```

```
## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 22
##   Unobserved stochastic nodes: 24
##   Total graph size: 161
##
## Initializing model
```

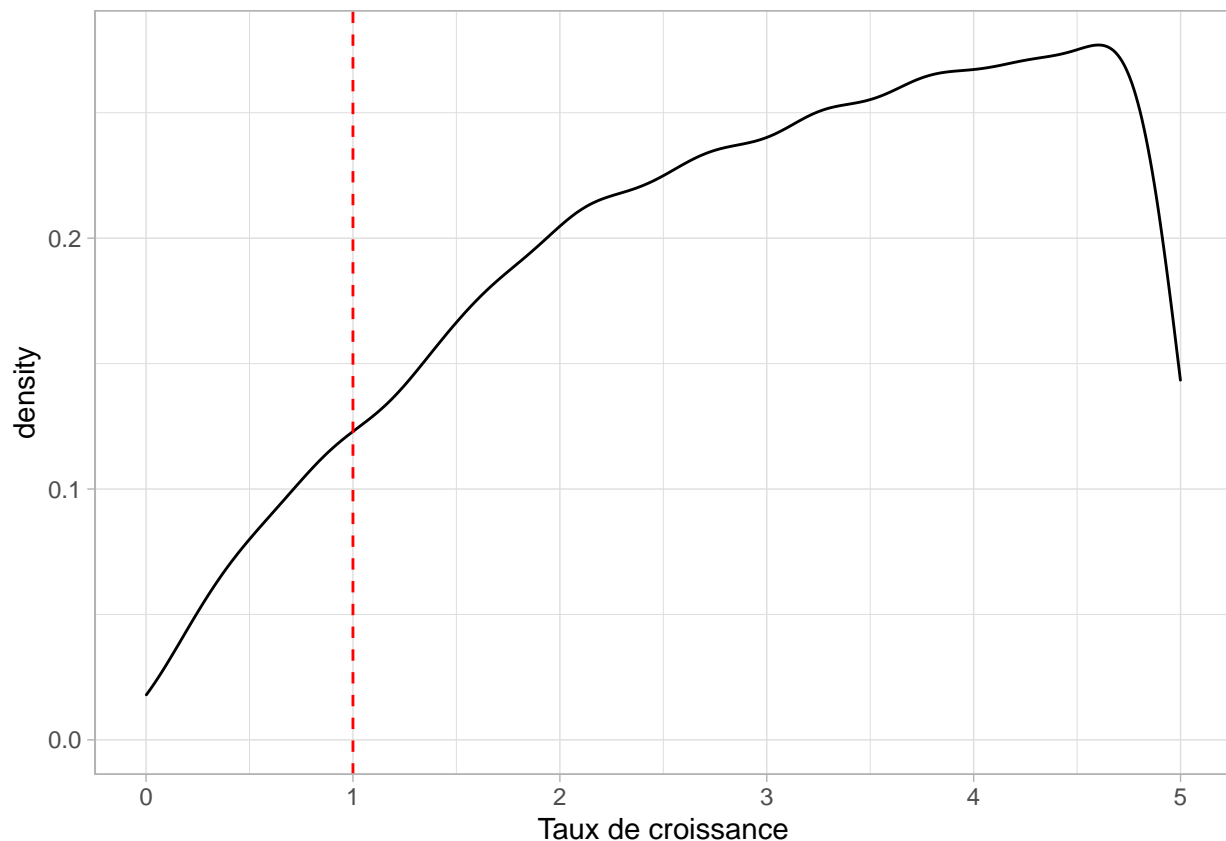
Jetons un coup d'œil aux estimations.

```
res <- print(lynx_mod, intervals = c(2.5/100, 50/100, 97.5/100))
```

```
## Inference for Bugs model at "/var/folders/r7/j0wqj1k95vz8w44sdxzm986c0000gn/T//RtmpZaM5Ca/model13dba4
## 3 chains, each with 1e+05 iterations (first 50000 discarded), n.thin = 10
## n.sims = 15000 iterations saved
##      mu.vect sd.vect   2.5%   50%  97.5%  Rhat n.eff
## N[1]    27.027   5.567  17.217  26.645  38.188 1.001 15000
```

```
## N[2]      34.026   5.807  23.022  34.106  45.649  1.001 15000
## N[3]      32.802   5.772  22.583  32.463  45.166  1.001  5900
## N[4]      39.715   6.334  28.303  39.305  53.038  1.001 11000
## N[5]      23.610   4.834  15.105  23.284  33.995  1.002  2500
## N[6]      31.804   5.789  21.530  31.376  44.694  1.001  8500
## N[7]      28.669   5.348  19.139  28.316  39.981  1.001  4900
## N[8]      18.651   4.292  11.224  18.336  27.895  1.001 15000
## N[9]      24.848   5.038  15.657  24.904  35.120  1.001 10000
## N[10]     33.657   5.432  23.398  33.341  45.158  1.001 15000
## N[11]     34.091   5.680  24.155  33.765  46.254  1.001 15000
## N[12]     36.983   5.673  26.359  36.654  49.070  1.001 14000
## N[13]     33.932   5.918  23.523  33.483  46.444  1.001  7700
## N[14]     40.706   6.490  28.990  40.279  54.691  1.001 15000
## N[15]     38.730   6.398  27.465  38.331  52.901  1.001 15000
## N[16]     34.616   5.893  24.075  34.276  47.042  1.001 15000
## N[17]     32.732   6.039  21.867  32.429  44.906  1.001 15000
## N[18]     28.988   5.656  19.222  28.497  41.683  1.001 15000
## N[19]     29.944   5.681  20.112  29.508  42.570  1.001 15000
## N[20]     28.678   5.344  19.065  28.349  40.104  1.001  7700
## N[21]     19.988   4.680  12.062  19.505  30.540  1.001 15000
## N[22]     25.667   5.051  16.708  25.304  36.413  1.001 15000
## lambda      3.013   1.273   0.489   3.140   4.910  1.001 12000
## sigmaProc    3.222   0.544   2.357   3.155   4.488  1.001 10000
## tauProc      0.104   0.033   0.050   0.100   0.180  1.001 10000
## deviance  138.223   6.754 126.969 137.569 152.993  1.001 15000
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 22.8 and DIC = 161.0
## DIC is an estimate of expected predictive error (lower deviance is better).
```

```
lynx_mod$BUGSoutput$sims.matrix %>%
  as_tibble() %>%
  # pivot_longer(cols = everything(), values_to = "value", names_to = "parameter") %>%
  # filter(str_detect(parameter, "lambda")) %>%
  ggplot() +
  aes(x = lambda) +
  geom_density() +
  geom_vline(xintercept = 1, lty = "dashed", color = "red") +
  labs(x = "Taux de croissance")
```

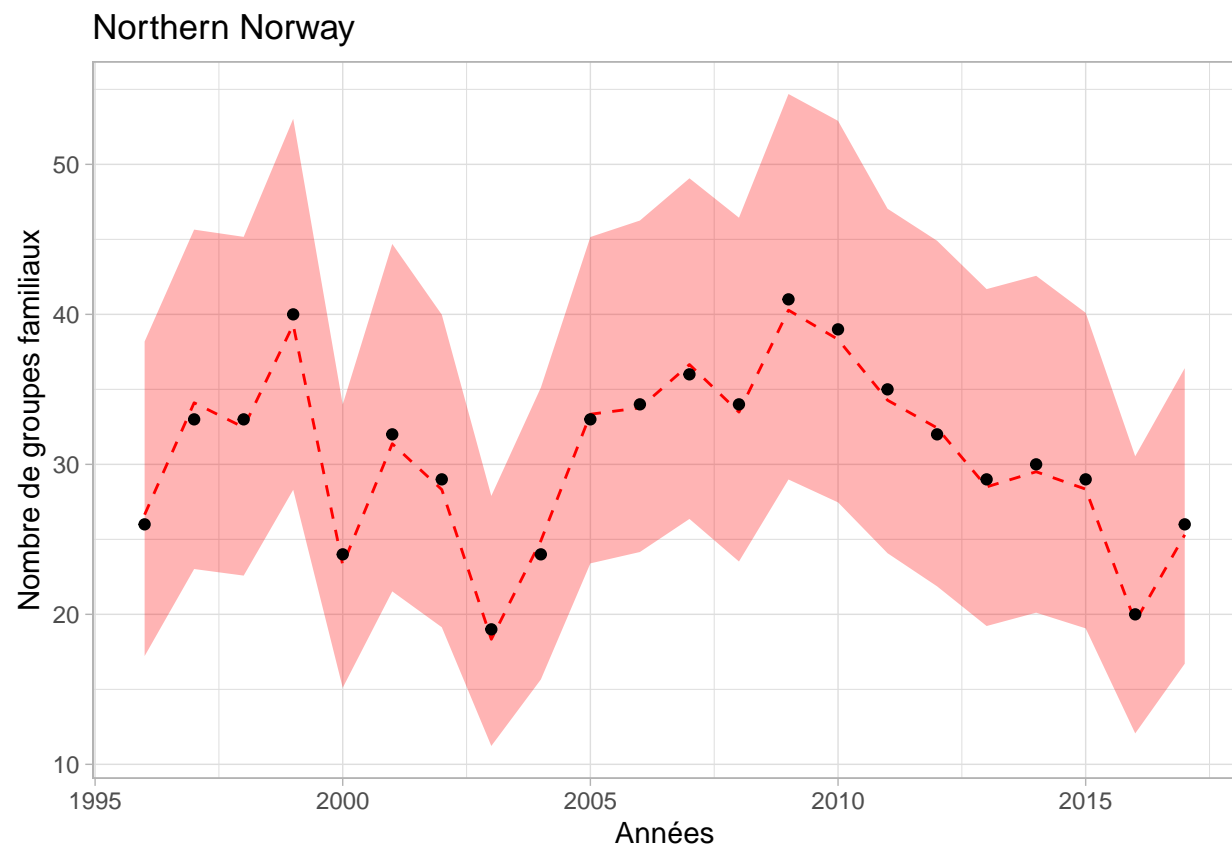



Ensuite les projections.

```
southern_norway <- lynx_mod$BUGSoutput$sims.matrix %>%
  as_tibble() %>%
  pivot_longer(cols = everything(), values_to = "value", names_to = "parameter") %>%
  filter(str_detect(parameter, "N")) %>%
  group_by(parameter) %>%
  summarize(medianN = median(value),
            lci = quantile(value, probs = 2.5/100),
            uci = quantile(value, probs = 97.5/100)) %>%
  mutate(an = parse_number(parameter) + 1995) %>%
  arrange(an) %>%
  ggplot() +
  geom_ribbon(aes(x = an, y = medianN, ymin = lci, ymax = uci), fill = "red", alpha = 0.3) +
  geom_line(aes(x = an, y = medianN), lty = "dashed", color = "red") +
  # geom_point(aes(x = an, y = medianN), color = "red") +
  geom_point(data = bugs.data %>% as_tibble, aes(x = 1995 + 1:unique(nyears), y = y)) +
  labs(y = "Nombre de groupes familiaux",
       x = "Années",
       title = "Northern Norway")
```

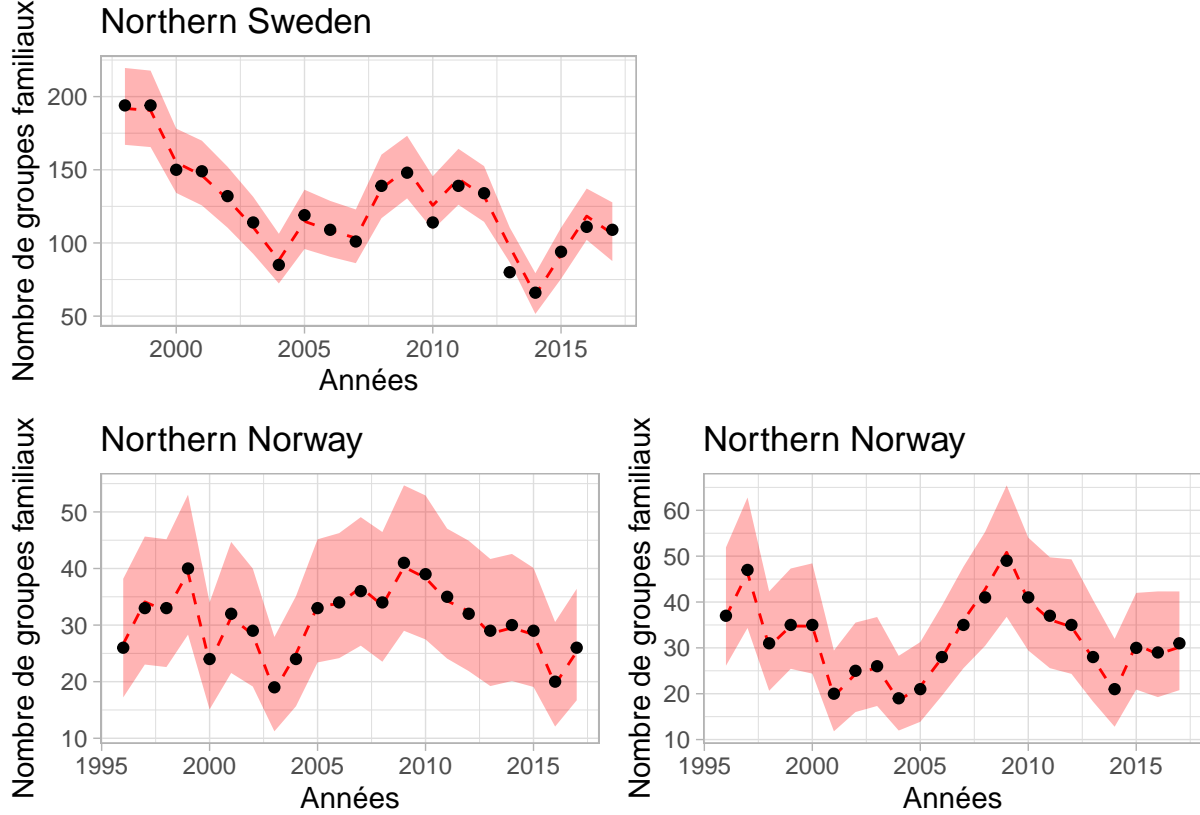
```
## 'summarise()' ungrouping output (override with '.groups' argument)
```

```
southern_norway
```



Tout ensemble - Figure 3 ou presque

```
library(patchwork)
(northern_sweden + grid::textGrob("")) / (southern_norway | northern_norway)
```



Hmm. Si l'on compare à la Figure 3 du papier, on s'aperçoit que l'ajustement du modèle exponentiel aux données est bien meilleur que celui du modèle structuré en âge développé par les auteurs. Ha!

Forecasting

Le modèle décrit l'évolution des effectifs à t en fonction des effectifs à $t-1$ et permet donc de projeter les effectifs en 2018 en connaissant les effectifs de 2017 la dernière année du suivi, puis ceux de 2019 en utilisant les effectifs prédits pour 2018, et ainsi de suite. A chaque étape, il y a des erreurs qui s'accumulent. L'approche bayésienne a l'avantage de permettre de faire ces prédictions en reportant les incertitudes d'une année à l'autre. C'est ce qui fait des modèles à espace d'états en bayésien un outil très utile pour faire des projections.

Bien. Maintenant dans le modèle utilisé, la variable effectifs prélevés est supposée connue. Il s'agit d'une donnée, et par définition on ne la connaît pas dans le futur. Il nous faut donc un modèle sur les effectifs prélevés, comme on en a un sur les effectifs comptés.

Les auteurs proposent le modèle à espace d'états suivant :

$$H_t \sim \text{log-Normale}(\max(0, \log(b_0 + b_1 y_{t-1})), \sigma_q^2)$$

et

$$q_t \sim \text{Poisson}(H_t)$$

où q_t est le quota observé au temps t et H_t l'effectif réel d'animaux prélevés. La prédiction du modèle est H_t avec une erreur de processus σ_q^2 .

On retrouve l'astuce utilisée par Guillaume pour forcer la moyenne de la normale à être supérieure ou égale à 0 avec le $\max(0, \log)$.

On a deux scénarios, ou bien un quota proportionnel aux effectifs comptés avec $b_0 = 0$ (modèle 1 : proportional quota setting strategy), ou bien des prélèvements qui augmentent proportionnellement, avec un

quota nul en-dessous d'un seuil (modèle 2 : threshold quota setting strategy). Ce seuil X se calcule en fixant $0 = b_0 + b_1X$ soit $X = -b_0/b_1$. J'ai pas tout bien compris encore à ce scénario. Ca deviendra plus clair en essayant d'ajuster les modèles je suppose.

On lit les données spécifique au modèle de décision. On a : * year – the year of census (February) * run – the run in the data * country – code for country; 1 = Sweden and 2 = Norway * census – number of lynx family groups censused in that year in that region * quota – the harvest quota for lynx based on the census result of the same year in that region * quota_1 – the harvest quota for lynx based on the census result of the year before in the region.

```
dat <- read.csv("eap2063-sup-0004-datas2.csv")
```

```
dat %>%
  filter(country == "1") %>%
  select(year, census, quota_1) -> dat_sweden
```

```
dat_sweden
```

##	year	census	quota_1
## 1	1998	194	140
## 2	1999	194	108
## 3	2000	150	73
## 4	2001	149	72
## 5	2002	132	67
## 6	2003	114	32
## 7	2004	86	15
## 8	2005	119	28
## 9	2006	109	30
## 10	2007	102	32
## 11	2008	140	99
## 12	2009	148	127
## 13	2010	114	95
## 14	2011	139	86
## 15	2012	135	62
## 16	2013	80	24
## 17	2014	66	0

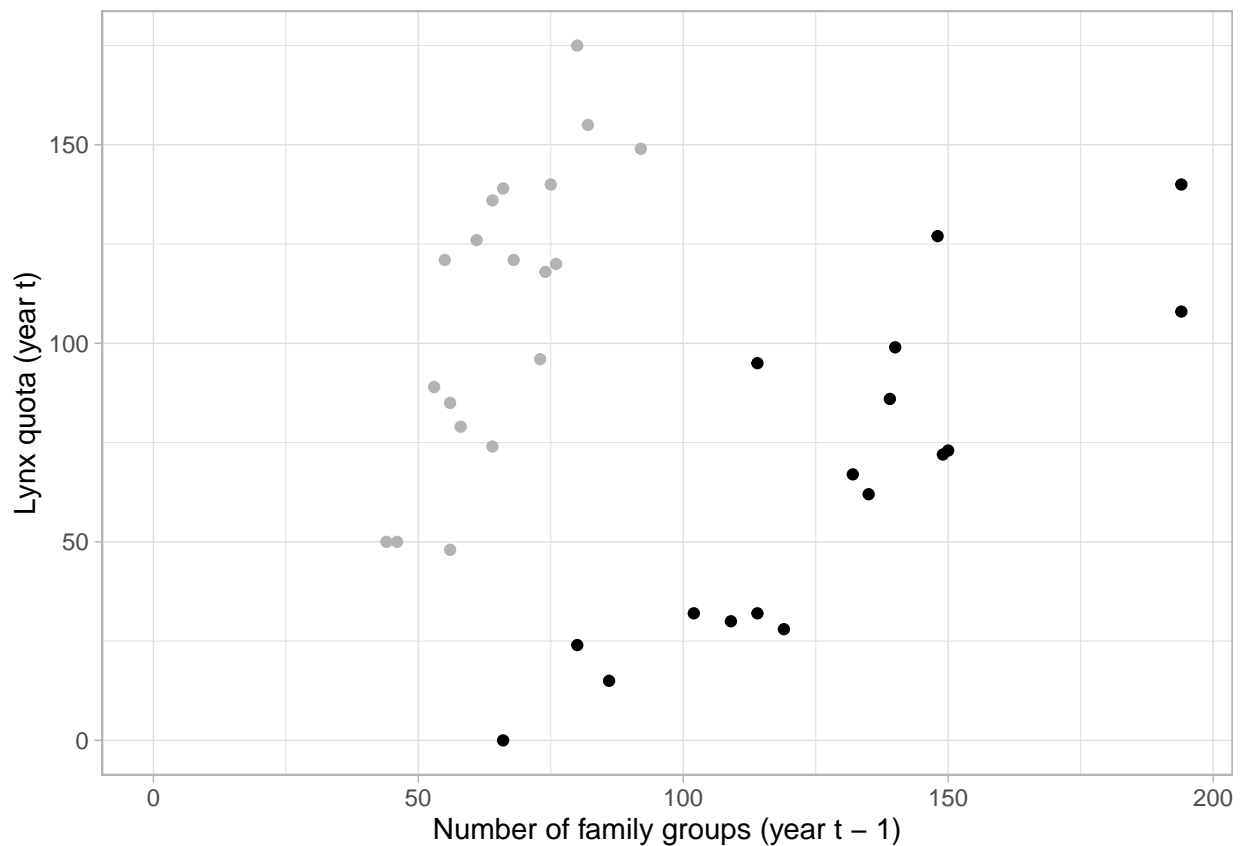
```
dat %>%
  filter(country == "2") %>%
  select(year, census, quota_1) -> dat_norway
```

```
dat_norway
```

##	year	census	quota_1
## 1	1996	64	136
## 2	1997	82	155
## 3	1998	66	139
## 4	1999	75	140
## 5	2000	61	126
## 6	2001	55	121
## 7	2002	56	85
## 8	2003	46	50
## 9	2004	44	50

```
## 10 2005    56    48
## 11 2006    64    74
## 12 2007    73    96
## 13 2008    76   120
## 14 2009    92   149
## 15 2010    80   175
## 16 2011    74   118
## 17 2012    68   121
## 18 2013    58    79
## 19 2014    53    89
```

```
ggplot() +
  geom_point(data = dat_sweden, aes(x = census, y = quota_1), color = "black") +
  geom_point(data = dat_norway, aes(x = census, y = quota_1), color = "gray70") +
  expand_limits(x = 0, y = 0) +
  labs(x = "Number of family groups (year t - 1)",
       y = "Lynx quota (year t)")
```



Modèle 1

Commençons par le modèle 1.

```
model1 <- function(){
  # Priors
```

```

sigmaProc ~ dunif(0, 4)
tauProc <- 1/sigmaProc^2
b[1] ~ dnorm(0, 3)

# Process model
for (t in 1:(nyears)) {
  mu[t] <- log(b[1] * y[t])
  Hproc[t] <- max(0, mu[t])
  H[t] ~ dlnorm(Hproc[t], tauProc)
}

# Observation model
for (t in 1:nyears) {
  q[t] ~ dpois(H[t])
}
}

```

On prépare les données pour la Suède.

```

bugs.data <- list(
  nyears = 17,
  y = dat_sweden$census,
  q = dat_sweden$quota_1)

```

On précise les paramètres à estimer et le nombre de chaînes de MCMC (j'en prends trois ici).

```

bugs.monitor <- c("b", "sigmaProc", "H")
bugs.chains <- 3
bugs.inits <- function(){
  list(
  )
}

```

Allez zooh, on lance la machine!

```

mod1_sweden <- jags(data = bugs.data,
  inits = bugs.inits,
  parameters.to.save = bugs.monitor,
  model.file = model1,
  n.chains = bugs.chains,
  n.thin = 10,
  n.iter = 100000,
  n.burnin = 50000)

```

```

## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 17
##   Unobserved stochastic nodes: 19
##   Total graph size: 106
##
## Initializing model

```

Jetons un coup d'oeil aux estimations.

```
print(mod1_sweden, intervals = c(2.5/100, 50/100, 97.5/100))
```

```
## Inference for Bugs model at "/var/folders/r7/j0wqj1k95vz8w44sdxzm986c0000gn/T//RtmpZaM5Ca/model3dba2
## 3 chains, each with 1e+05 iterations (first 50000 discarded), n.thin = 10
## n.sims = 15000 iterations saved
##      mu.vect sd.vect   2.5%   50%   97.5%  Rhat n.eff
## H[1]    138.820  11.644 116.906 138.461 162.411 1.001 15000
## H[2]    107.162  10.212  87.998 106.909 128.201 1.001  6000
## H[3]     72.540   8.450  56.840  72.234  89.975 1.001 15000
## H[4]     71.636   8.387  56.162  71.313  88.911 1.001  6800
## H[5]     66.645   8.147  51.502  66.326  83.318 1.001 15000
## H[6]     32.686   5.598  22.586  32.392  44.500 1.001  4500
## H[7]     16.515   3.928   9.734  16.237  25.075 1.001 15000
## H[8]     28.987   5.205  19.569  28.689  39.958 1.001 15000
## H[9]     30.665   5.377  20.992  30.360  41.953 1.001 12000
## H[10]    32.430   5.552  22.484  32.116  44.248 1.001 12000
## H[11]    97.887   9.821  79.558  97.614 118.025 1.001 15000
## H[12]   125.485  11.281 104.591 125.145 148.454 1.001 15000
## H[13]    93.478   9.556  75.486  93.209 113.067 1.001  7000
## H[14]    85.160   9.131  68.351  84.668 104.115 1.001 15000
## H[15]    61.757   7.742  47.618  61.433  77.647 1.001 15000
## H[16]    24.568   4.820  16.143  24.252  34.824 1.001  4100
## H[17]     3.816   2.041   0.856   3.497   8.672 1.001  9600
## b         0.402   0.079   0.264   0.395   0.578 1.001 15000
## sigmaProc 0.775   0.200   0.479   0.744   1.261 1.001 11000
## deviance 117.394   6.889 105.810 116.757 132.765 1.001 15000
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 23.7 and DIC = 141.1
## DIC is an estimate of expected predictive error (lower deviance is better).
```

Le paramètre b_1 est estimé très proche de la valeur qu'on trouve dans le Tableau 4.

```
mod1_sweden$BUGSoutput$mean$b
```

```
## [1] 0.4016313
```

Idem pour la Norvège. On prépare les données.

```
bugs.data <- list(
  nyears = 19,
  y = dat_norway$census,
  q = dat_norway$quota_1)
```

On précise les paramètres à estimer et le nombre de chaînes de MCMC (j'en prends trois ici).

```
bugs.monitor <- c("b", "sigmaProc", "H")
bugs.chains <- 3
bugs.inits <- function(){
  list(
  )
}
```

Allez zooh, on lance la machine!

```
mod1_norway <- jags(data = bugs.data,
  inits = bugs.inits,
  parameters.to.save = bugs.monitor,
  model.file = model1,
  n.chains = bugs.chains,
  n.thin = 10,
  n.iter = 100000,
  n.burnin = 50000)
```

```
## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 19
##   Unobserved stochastic nodes: 21
##   Total graph size: 118
##
## Initializing model
```

Jetons un coup d'oeil aux estimations.

```
print(mod1_norway, intervals = c(2.5/100, 50/100, 97.5/100))
```

```
## Inference for Bugs model at "/var/folders/r7/j0wqj1k95vz8w44sdxzm986c0000gn/T//RtmpZaM5Ca/model13dba2"
## 3 chains, each with 1e+05 iterations (first 50000 discarded), n.thin = 10
## n.sims = 15000 iterations saved
##           mu.vect sd.vect   2.5%    50%   97.5%  Rhat n.eff
## H[1]      131.554  10.999 110.967 131.229 154.279 1.001  5700
## H[2]      152.424  11.889 130.139 152.145 176.567 1.001 15000
## H[3]      134.849  11.144 114.036 134.445 157.925 1.001 14000
## H[4]      137.562  11.180 116.786 137.159 160.529 1.001 15000
## H[5]      121.949  10.518 102.434 121.539 143.839 1.001 15000
## H[6]      116.075  10.206  97.155 115.851 137.242 1.001 15000
## H[7]       85.564   8.646  69.626  85.283 103.406 1.001 15000
## H[8]       54.541   6.714  41.857  54.317  68.301 1.001 15000
## H[9]       54.072   6.667  41.707  53.782  67.759 1.001 15000
## H[10]      55.497   7.089  42.489  55.279  69.916 1.001 15000
## H[11]      78.110   8.295  62.628  77.861  95.160 1.001 15000
## H[12]      98.468   9.268  81.272  98.154 117.447 1.001 13000
## H[13]     119.852  10.236 100.565 119.638 140.342 1.001 15000
## H[14]     148.542  11.684 126.391 148.401 171.981 1.001 11000
## H[15]     169.836  12.555 146.086 169.553 195.171 1.001 15000
## H[16]     117.773  10.308  98.610 117.329 138.835 1.001 15000
```



```
## H[17]      119.205  10.283 100.165 118.890 140.343 1.001  6200
## H[18]      81.017   8.326  65.595  80.735  98.042 1.001  6900
## H[19]      88.024   8.613  71.970  87.767 105.559 1.001 11000
## b          1.560   0.101   1.360   1.560   1.758 1.001 15000
## sigmaProc  0.262   0.058   0.170   0.255   0.397 1.001 15000
## deviance  142.562   6.433 132.069 141.859 156.892 1.001 15000
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 20.7 and DIC = 163.3
## DIC is an estimate of expected predictive error (lower deviance is better).
```

Le paramètre b_1 est estimé proche de la valeur qu'on trouve dans le Tableau 4.

```
mod1_norway$BUGSoutput$mean$b
```

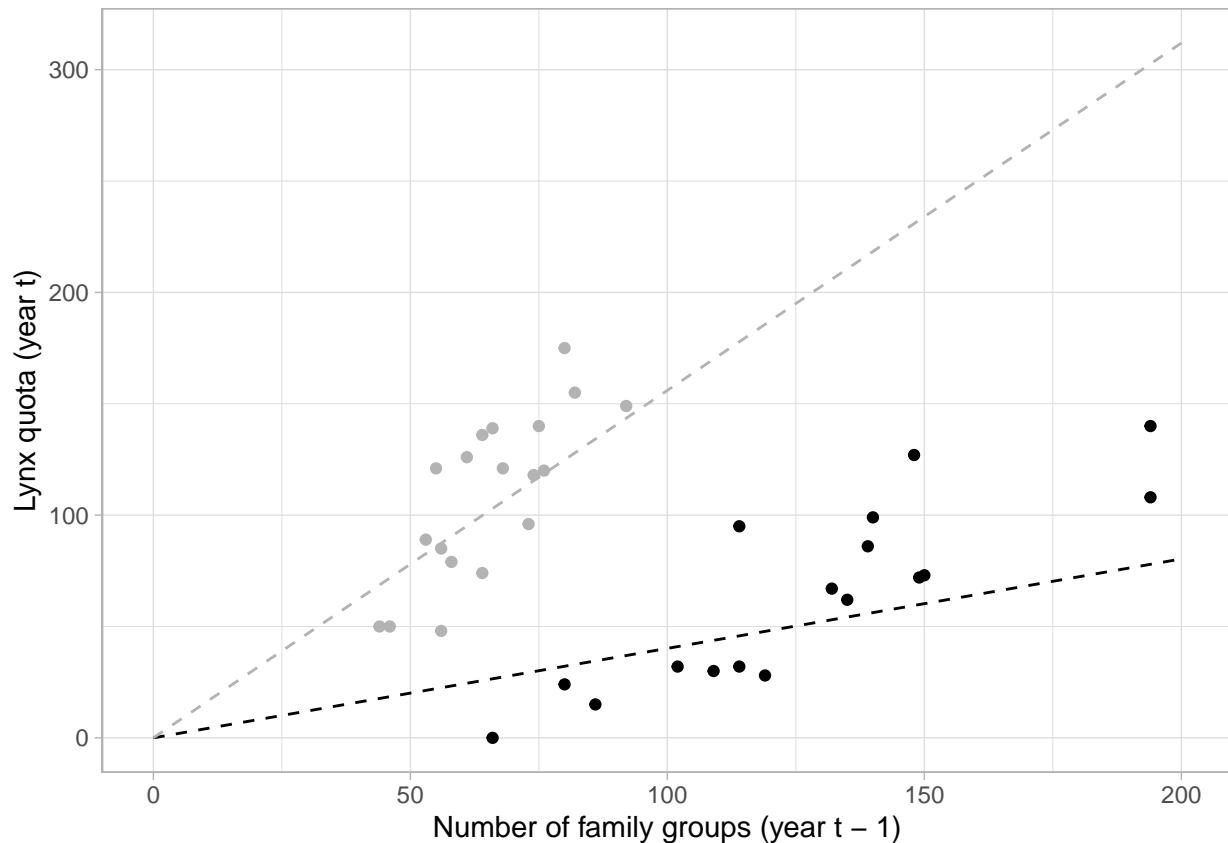
```
## [1] 1.559921
```

Graphiquement, on obtient.

```
swgrid <- seq(0, 200, length.out = length(dat_sweden$census))
nwgrid <- seq(0, 200, length.out = length(dat_norway$census))
ggplot() +
  geom_point(data = dat_sweden, aes(x = census, y = quota_1), color = "black") +
  geom_point(data = dat_norway, aes(x = census, y = quota_1), color = "gray70") +
  geom_line(data = dat_sweden, aes(x = swgrid, y = mod1_sweden$BUGSoutput$mean$b * swgrid), color = "black") +
  geom_line(data = dat_norway, aes(x = nwgrid, y = mod1_norway$BUGSoutput$mean$b * nwgrid), color = "gray70") +
  expand_limits(x = 0, y = 0) +
  labs(x = "Number of family groups (year t - 1)",
       y = "Lynx quota (year t)")
```

```
## Warning in mod1_sweden$BUGSoutput$mean$b * swgrid: Recycling array of length 1 in array-vector arithmetic
## Use c() or as.vector() instead.
```

```
## Warning in mod1_norway$BUGSoutput$mean$b * nwgrid: Recycling array of length 1 in array-vector arithmetic
## Use c() or as.vector() instead.
```



Modèle 2

On écrit le modèle. La différence avec le modèle 1 est qu'on estime une ordonnée à l'origine.

```
model2 <- function(){

  # Priors
  sigmaProc ~ dunif(0, 4)
  tauProc <- 1/sigmaProc^2
  b[1] ~ dnorm(0, 1000)
  b[2] ~ dnorm(0, 1000)
  # Process model
  for (t in 1:(nyears)) {
    mu[t] <- log(b[1] + b[2] * y[t])
    # mu[t] <- log(b[1] + b[2] * y[t]) * index[t]
    # index[t] <- - 1000 * step(y[t] + b[1] / b[2]) # step(x) = 1 if x >= 0
    # index[t] <- step(q[t]) # step(x) = 1 if x >= 0
    # mu[t] <- log(b[1] + b[2] * y[t])
    Hproc[t] <- max(0, mu[t])
    H[t] ~ dlnorm(Hproc[t], tauProc)

    # les lignes de code suivantes donnent un ajustement pas mal, mais
    # sauf qu'à l'approche de census == 0 on a harvest == 0
    # Hproc[t] <- log(b[1] + b[2] * y[t])
    # H[t] ~ dlnorm(Hproc[t], tauProc)
  }
}
```

```

# Observation model
for (t in 1:nyears) {
  q[t] ~ dpois(H[t])
}
}

```

On prépare les données pour la Suède.

```

bugs.data <- list(
  nyears = 17,
  y = dat_sweden$census,
  q = dat_sweden$quota_1)

```

On précise les paramètres à estimer et le nombre de chaînes de MCMC (j'en prends trois ici).

```

bugs.monitor <- c("b", "sigmaProc")
bugs.chains <- 3
bugs.inits <- function(){
  list(
  )
}

```

Allez zooh, on lance la machine!

```

mod2_sweden <- jags(data = bugs.data,
  inits = bugs.inits,
  parameters.to.save = bugs.monitor,
  model.file = model2,
  n.chains = bugs.chains,
  n.thin = 10,
  n.iter = 100000,
  n.burnin = 50000)

```

```

## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 17
##   Unobserved stochastic nodes: 20
##   Total graph size: 122
##
## Initializing model

```

Jetons un coup d'oeil aux estimations.

```

print(mod2_sweden, intervals = c(2.5/100, 50/100, 97.5/100))

```

```

## Inference for Bugs model at "/var/folders/r7/j0wqj1k95vz8w44sdxzm986c0000gn/T//RtmpZaM5Ca/model3dba3
## 3 chains, each with 1e+05 iterations (first 50000 discarded), n.thin = 10
## n.sims = 15000 iterations saved

```

```
##          mu.vect sd.vect   2.5%    50%   97.5%  Rhat n.eff
## b[1]         0.001  0.032 -0.062   0.001   0.063 1.001  9200
## b[2]         0.087  0.024  0.042   0.086   0.136 1.001  9900
## sigmaProc    1.956  0.485  1.210   1.885   3.114 1.001 15000
## deviance    111.594  5.872 101.977 110.971 124.688 1.001  9800
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 17.2 and DIC = 128.8
## DIC is an estimate of expected predictive error (lower deviance is better).
```

Les paramètres b sont estimés comme suit.

```
mod2_sweden$BUGSoutput$mean$b
```

```
## [1] 0.0005412535 0.0865190161
```

Le ratio se calcule comme suit.

```
- mod2_sweden$BUGSoutput$mean$b[1] / mod2_sweden$BUGSoutput$mean$b[2]
```

```
## [1] -0.006255891
```

```
lm(q ~ y, data = bugs.data)
```

```
##
## Call:
## lm(formula = q ~ y, data = bugs.data)
##
## Coefficients:
## (Intercept)          y
##      -64.757         1.009
```

```
glm(q ~ y, data = bugs.data, family = "poisson")
```

```
##
## Call: glm(formula = q ~ y, family = "poisson", data = bugs.data)
##
## Coefficients:
## (Intercept)          y
##      2.10350         0.01504
##
## Degrees of Freedom: 16 Total (i.e. Null); 15 Residual
## Null Deviance:      481.8
## Residual Deviance: 178.1    AIC: 276
```

Graphiquement, on obtient.

```

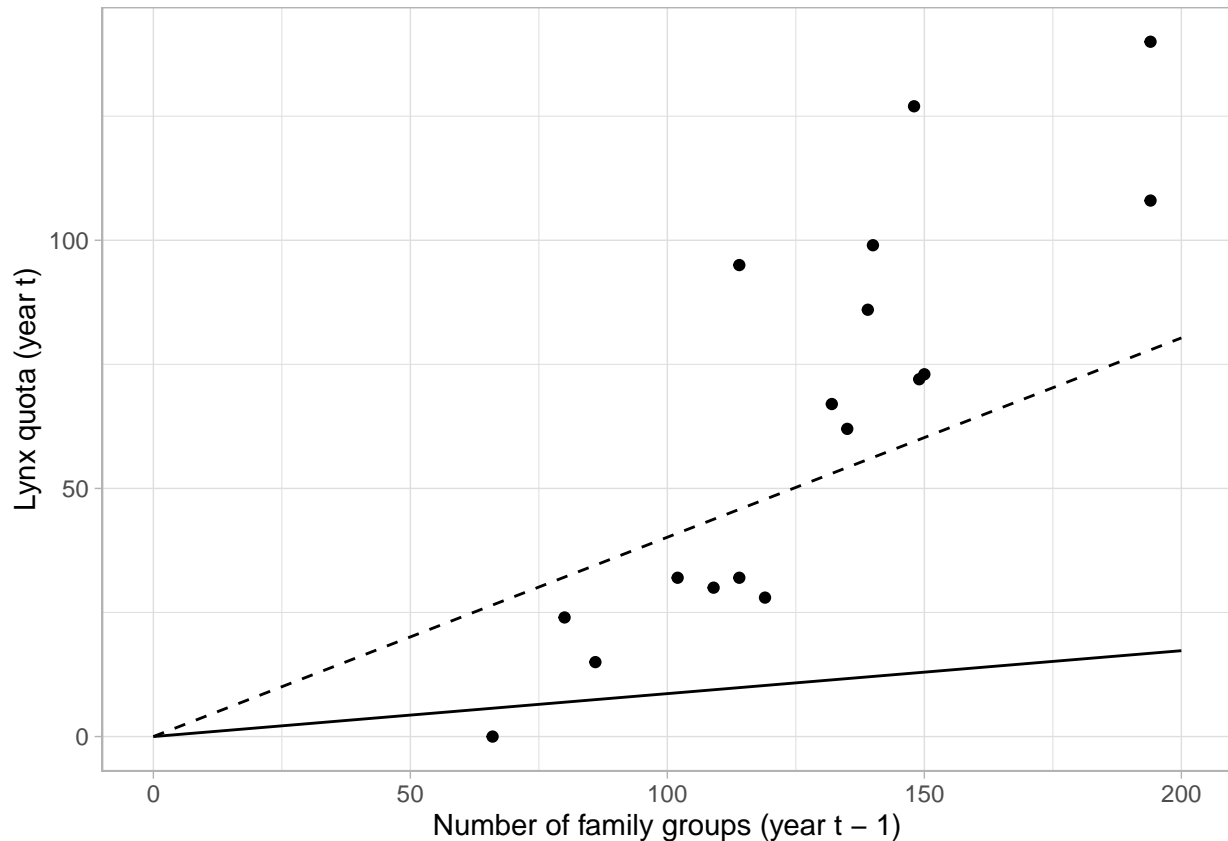
swgrid <- seq(0, 200, length.out = length(dat_sweden$census))
ggplot() +
  geom_point(data = dat_sweden, aes(x = census, y = quota_1), color = "black") +
  geom_line(data = dat_sweden, aes(x = swgrid, y = mod1_sweden$BUGSoutput$mean$b * swgrid), color = "black") +
  geom_line(data = dat_sweden, aes(x = swgrid, y = (mod2_sweden$BUGSoutput$mean$b[1] + mod2_sweden$BUGSoutput$mean$b[2] * swgrid)), color = "black") +
  expand_limits(x = 0, y = 0) +
  labs(x = "Number of family groups (year t - 1)",
       y = "Lynx quota (year t)")

```

```

## Warning in mod1_sweden$BUGSoutput$mean$b * swgrid: Recycling array of length 1 in array-vector arithmetic
## Use c() or as.vector() instead.

```



Idem pour la Norvège. On prépare les données.

```

bugs.data <- list(
  nyears = 19,
  y = dat_norway$census,
  q = dat_norway$quota_1)

```

On précise les paramètres à estimer et le nombre de chaînes de MCMC (j'en prends trois ici).

```

bugs.monitor <- c("b", "sigmaProc", "H")
bugs.chains <- 3
bugs.inits <- function(){
  list(

```

```
)
}
```

Allez zooh, on lance la machine!

```
mod2_norway <- jags(data = bugs.data,
  inits = bugs.inits,
  parameters.to.save = bugs.monitor,
  model.file = model2,
  n.chains = bugs.chains,
  n.thin = 10,
  n.iter = 100000,
  n.burnin = 50000)
```

```
## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 19
##   Unobserved stochastic nodes: 22
##   Total graph size: 136
##
## Initializing model
```

Jetons un coup d'oeil aux estimations.

```
print(mod2_norway, intervals = c(2.5/100, 50/100, 97.5/100))
```

```
## Inference for Bugs model at "/var/folders/r7/j0wqj1k95vz8w44sdxzm986c0000gn/T//RtmpZaM5Ca/model13dba5"
## 3 chains, each with 1e+05 iterations (first 50000 discarded), n.thin = 10
## n.sims = 15000 iterations saved
##           mu.vect sd.vect   2.5%    50%   97.5%  Rhat n.eff
## H[1]      135.692  11.600  113.919  135.380  159.185  1.001 15000
## H[2]      154.654  12.376  131.473  154.214  180.212  1.001 15000
## H[3]      138.665  11.766  116.793  138.306  162.531  1.001 15000
## H[4]      139.543  11.704  117.552  139.355  163.182  1.001 15000
## H[5]      125.724  11.270  104.445  125.397  148.736  1.001 15000
## H[6]      120.823  10.938  100.586  120.514  143.319  1.001 10000
## H[7]       84.684   9.203   67.455   84.328  103.645  1.001 15000
## H[8]       49.716   7.025   36.852   49.398   64.364  1.001 15000
## H[9]       49.721   7.075   36.817   49.358   64.835  1.001 12000
## H[10]      47.800   6.980   35.223   47.436   62.252  1.001 15000
## H[11]      73.576   8.579   57.782   73.234   91.432  1.001 13000
## H[12]      95.763   9.684   77.783   95.426  115.330  1.001 14000
## H[13]     119.511  10.990   99.042  119.257  141.878  1.001 15000
## H[14]     148.676  12.280  125.624  148.415  173.373  1.001 15000
## H[15]     174.670  13.356  149.538  174.325  201.547  1.001 15000
## H[16]     117.763  10.838   97.659  117.430  140.327  1.001 11000
## H[17]     120.656  11.072   99.777  120.354  143.377  1.001 15000
## H[18]      78.717   8.875   62.031   78.453   97.045  1.001  7800
## H[19]      88.776   9.418   71.326   88.446  108.044  1.001 12000
## b[1]        0.002   0.032  -0.061   0.002   0.064  1.001 15000
```

```
## b[2]          0.082   0.023   0.040   0.081   0.131 1.001 9600
## sigmaProc    3.070   0.467   2.202   3.065   3.910 1.001 15000
## deviance    141.936   6.244 131.657 141.298 156.111 1.001 15000
##
## For each parameter, n.eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor (at convergence, Rhat=1).
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 19.5 and DIC = 161.4
## DIC is an estimate of expected predictive error (lower deviance is better).
```

Le paramètre b_1 est estimé proche de la valeur qu'on trouve dans le Tableau 4.

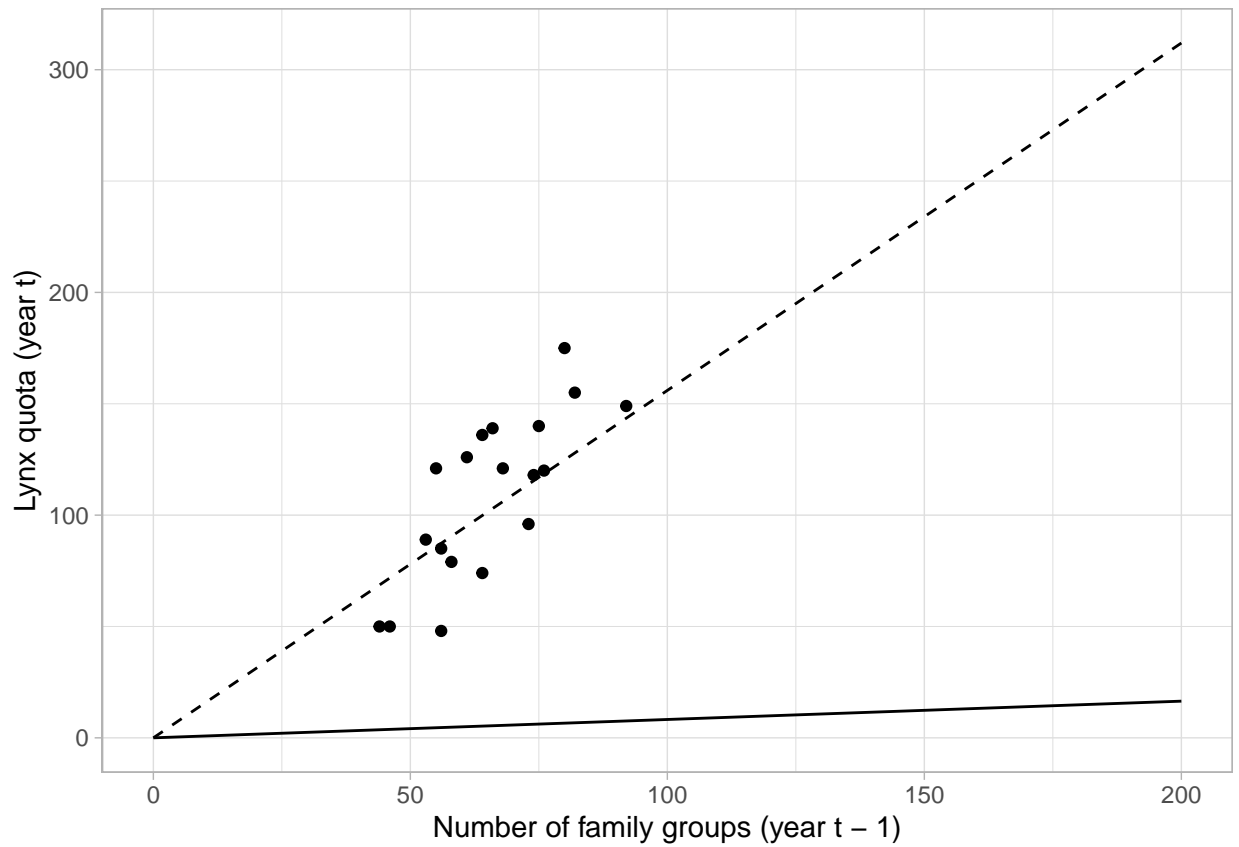
```
mod2_norway$BUGSoutput$mean$b
```

```
## [1] 0.001526982 0.082346563
```

Graphiquement, on obtient.

```
nwgrid <- seq(0, 200, length.out = length(dat_norway$census))
ggplot() +
  geom_point(data = dat_norway, aes(x = census, y = quota_1), color = "black") +
  geom_line(data = dat_norway, aes(x = nwgrid, y = mod1_norway$BUGSoutput$mean$b * nwgrid), color = "black") +
  geom_line(data = dat_norway, aes(x = nwgrid, y = (mod2_norway$BUGSoutput$mean$b[1] + mod2_norway$BUGSoutput$mean$b[2] * nwgrid)), color = "black") +
  expand_limits(x = 0, y = 0) +
  labs(x = "Number of family groups (year t - 1)",
       y = "Lynx quota (year t)")
```

```
## Warning in mod1_norway$BUGSoutput$mean$b * nwgrid: Recycling array of length 1 in array-vector arithmetic
## Use c() or as.vector() instead.
```



Je n'arrive pas reproduire ce satané modèle à seuil avec l'intercept non-nul, les lignes en trait plein de la figure 4 du papier. Grrrr. On arrive aux estimations du Tableau 3 via une simple régression hein, donc c'est pas bien grave.