```
ELBO \doteq \mathbb{E}_{\theta \sim q} [\log p(y|\theta)] - KL(q||p(\cdot)) \leq \log p(y)
1.1 Usefull math
                                                                                 f \sim GP(\mu, k) then: f|y_{1:n}, x_{1:n} \sim GP(\tilde{\mu}, \tilde{k})
                                                                                                                                                                   5.3 Markov Chain Monte Carlo
\varphi is convex \Rightarrow \varphi(\mathbb{E}[X]) \leq \mathbb{E}[\varphi(X)]
                                                                                 \tilde{\mu}(x) = \mu(x) + K_{A_n}^T (K_{AA} + \epsilon I_n)^{-1} (y_A - \mu_A)
                                                                                                                                                                   Idea: All we need is sampling from postirior
Hoeffding: Z_1, ... iid, Z_i \in [0, C], \mathbb{E}[Z_i] = \mu
                                                                                                                                                                   Ergodic Markov Chain:
                                                                                 \tilde{k}(x, x') = k(x, x') - K_{A, x}^{T} (K_{AA} + \epsilon I_n)^{-1} K_{A, x'}
                                                                                                                                                                   \exists t \text{ s.t. } \mathbb{P}(i \to j \text{ in t steps}) > 0 \ \forall i, j \Rightarrow
\Rightarrow P(|\mu - \frac{1}{n}\sum_{i=1}^{n} Z_i| > \epsilon) \le 2 \exp(-2n\frac{\epsilon^2}{C}) \le \delta
                                                                                 Where: K_{A,x} = [k(x_1, x)...k(x_n, x)]^T
                                                                                                                                                                   \exists ! \pi = \lim_{N \to \infty} \mathbb{P}(X_n = x) Limit distribution
\Rightarrow n \ge \frac{C}{2c^2} \log \frac{2}{\delta}
                                                                                 [K_{AA}]_{ij} = k(x_i, x_j) \text{ and } \mu_A = [\mu(x_1 \dots x_n)]^T
                                                                                                                                                                   Ergodic Theorem: if (X_i)_{i \in \mathbb{N}} is ergodic:
Robbins Monro \alpha_t \xrightarrow{RM} 0: \sum \alpha_t = \infty, \sum \alpha_t^2 < \infty
                                                                                 4.2 Kernels
                                                                                                                                                                   \lim_{N\to\infty} \frac{1}{n} \sum_{i=1}^{N} f(X_i) = \mathbb{E}_{x\sim\pi} [f(x)]
                                                                                 k(x,y) is a kernel if it's symmetric semidefinite
1.2 Multivariate Gaussian
                                                                                 positive:
f(x) = \frac{1}{2\pi\sqrt{|\Sigma|}}e^{-\frac{1}{2}(x-\mu)^T\Sigma^{-1}(x-\mu)}
                                                                                                                                                                   Detailed Blanced Equation:
                                                                                 \forall \{x_1, \dots, x_n\} then for the Gram Matrix
                                                                                                                                                                   P(x|x') is the transition model of a MC:
                                                                                 [K]_{ij} = k(x_i, x_j) \text{ holds } c^T K c \ge 0 \forall c
Suppose we have a Gaussian random vector
                                                                                                                                                                   if R(x)P(x'|x) = R(x')P(x|x') then R is the limit
\begin{bmatrix} X_A \\ X_B \end{bmatrix} \sim \mathcal{N} \left( \begin{bmatrix} \mu_A \\ \mu_B \end{bmatrix}, \begin{bmatrix} \Sigma_{AA} & \Sigma_{AB} \\ \Sigma_{BA} & \Sigma_{BB} \end{bmatrix} \right) \Rightarrow X_A | X_B = x_B \sim
                                                                                 Some Kernels: (h is the bandwidth hyperp.)
                                                                                                                                                                   distribution of the MC
                                                                                 Gaussian (rbf): k(x,y) = \exp(-\frac{||x-y||^2}{L^2})
                                                                                                                                                                   Metropolis Hastings Algo: Sample from a MC
\mathcal{N}\left(\mu_A + \Sigma_{AB}\Sigma_{BB}^{-1}(x_B - \mu_B), \Sigma_{AA} - \Sigma_{AB}\Sigma_{BB}^{-1}\Sigma_{BA}\right)
                                                                                 Exponential: k(x,y) = \exp(-\frac{\|x-y\|}{h})
                                                                                                                                                                  which has P(x) = \frac{Q(x)}{Z} as limit dist.
1.3 Information Theory elements:
                                                                                 Linear kernel: k(x,y) = x^T y (here K_{AA} = XX^T)
Entropy: H(X) \doteq -\mathbb{E}_{x \sim p_X} [\log p_X(x)]
                                                                                 4.3 Optimization of Kernel Parameters
H(X|Y) \doteq -\mathbb{E}_{(x,y) \sim p_{(X,Y)}} \left| \log p_{Y|X}(y|x) \right|
                                                                                 Given a dataset A, a kernel function k(x, y; \theta).
                                                                                 y \sim N(0, K_v(\theta)) where K_v(\theta) = K_{AA}(\theta) + \sigma_n^2 I
if X \sim \mathcal{N}(\mu, \Sigma) \Rightarrow H(X) = \frac{1}{2} \log \left[ (2\pi e)^d \det(\Sigma) \right]
                                                                                 \hat{\theta} = \arg\max_{\theta} \log p(y|X;\theta)
Chain Rule: H(X,Y) = H(Y|X) + H(X)
Mutual Info: I(X,Y) \doteq KL(p_{(X,Y)}||p_Xp_Y)
                                                                                 In GP: \hat{\theta} = \arg\min_{\theta} y^T K_v^{-1}(\theta) y + \log |K_v(\theta)|
                                                                                 We can from here \nabla \downarrow:
I(X,Y) = H(X) - H(X|Y)
                                                                                 \nabla_{\theta} \log p(y|X;\theta) = \frac{1}{2} tr((\alpha \alpha^{T} - K^{-1}) \frac{\partial K}{\partial \theta}), \alpha = K^{-1} y
if X \sim \mathcal{N}(\mu, \Sigma), Y = X + \epsilon, \epsilon \sim \mathcal{N}(0, \sigma^2 I):
then I(X, Y) = \frac{1}{2} \log \left| \det \left( I + \frac{1}{\sigma^2} \Sigma \right) \right|
                                                                                 Or we could also be baysian about \theta
                                                                                 4.4 Aproximation Techniques
1.4 Kullback-Leiber divergence
                                                                                 4.4.1 Local method
                                                                                 k(x_1, x_2) = 0 if ||x_1 - x_2|| \gg 1
KL(p||q) = \mathbb{E}_p \left| \log \frac{p(x)}{q(x)} \right|
                                                                                 4.4.2 Random Fourier Features
if p_0 \sim \mathcal{N}(\mu_0, \Sigma_0), p_1 \sim \mathcal{N}(\mu_1, \Sigma_1) \Rightarrow KL(p_0 || p_1)
                                                                                 Need k(x, y) = \kappa(x - y); let p(w) = \mathcal{F} \{\kappa(\cdot), w\}.
=\frac{1}{2}\left(tr\left(\Sigma_{1}^{-1}\Sigma_{0}\right)+(\mu_{1}-\mu_{0})^{T}\Sigma_{1}^{-1}(\mu_{1}-\mu_{0})-k+\log\frac{|\Sigma_{1}|}{|\Sigma_{0}|}\right)
                                                                                 Then p(w) can be normalized to be a density.
\hat{q} = \arg\min_{q} KL(p||q) \Rightarrow \text{overconservative}
                                                                                 \kappa(x-y) = \mathbb{E}_{p(w)} \left[ \exp \left\{ i w^T (x-y) \right\} \right] antitransform
\hat{q} = \arg\min_{a} KL(q||p) \Rightarrow \text{overconfident}
                                                                                 \kappa(x - y) = \mathbb{E}_{b \sim \mathcal{U}([0, 2\pi]), w \sim p(w)} [z_{w,b}(x) z_{w,b}(y)]
                                                                                 where z_{w,b}(x) = \sqrt{2}cos(w^Tx + b). I can MC ex-
2 Bayesian Regression
                                                                                 tract features z. If # features is \ll n then this is
w \sim N(0, \sigma_n^2 I), \ \epsilon \sim N(0, \sigma_n^2 I), \ y = Xw + \epsilon
                                                                                 faster (X^T X \text{ vs } XX^T)
y|w \sim N(Xw, \sigma_n^2 I)
                                                                                 4.4.3 Inducing points
w|y \sim N((X^TX + \lambda I)^{-1}X^Ty,(X^TX + \lambda I)^{-1}\sigma_y^2)
                                                                                 We a vector of inducing variables u
                                                                                 f_A|_u \sim N(K_{Au}K_uu^{-1}u, K_{AA} - K_{Au}K_uu^{-1}K_{uA})
3 Kalman Filter
                                                                                 f_*|_{u} \sim N(K_{*u}K_{u}u^{-1}u, K_{**} - K_{*u}K_{u}u^{-1}K_{u*})
\begin{cases} X_{t+1} = FX_t + \epsilon_t & \epsilon_t \sim N(0, \Sigma_x) \\ Y_t = HX_t + \eta_t & \eta_t \sim N(0, \Sigma_y) \end{cases} X_1 \sim N(\mu_p, \Sigma_p)
                                                                                 Subset of Regressors (SoR): \blacksquare \rightarrow 0
                                                                                 FITC: \blacksquare \rightarrow its diagonal
Then if X_0 is Gaussian then X_t|Y_{1:t} \sim N(\mu_t, \sigma_t):
                                                                                 5 Approximate inference
\mu_{t+1} = F\mu_t + K_{t+1}(y_{t+1} - HF\mu_t)
                                                                                 5.1 Laplace Approximation
\Sigma_{t+1} = (I - K_{t+1}H)(F\Sigma_t F^T + \Sigma_x)
                                                                                 \theta = \arg\max_{\theta} p(\theta|y)
K_{t+1} = (F\Sigma_t F^T + \Sigma_x) H^T (H(F\Sigma_t F^T + \Sigma_x) H^T + \Sigma_v)^{-1}
                                                                                 \Lambda = -\nabla_{\theta} \nabla_{\theta} \log p(\theta|y)|_{\theta = \hat{\theta}}
4 Gaussian Processes
                                                                                 p(\theta|y) \simeq q(\theta) = N(\hat{\theta}, \Lambda^{-1})
f \sim GP(\mu, k) \Rightarrow \forall \{x_1, \dots, x_n\} \ \forall n < \infty
                                                                                 5.2 Variationa Inverence
[f(x_1)...f(x_n)] \sim N([\mu(x_1)...\mu(x_n)], K)
                                                                                 \hat{q} = \arg\min_{q \in O} KL(q||p(\cdot|y))
where K_{ij} = k(x_i, x_j)
                                                                                 \hat{q} = \arg \max_{q \in O} ELBO Evidence Lower Bound
```

1 Review of useful concepts and Introduction 4.1 Gaussian Process Regression

```
MCMC but cannot store all the \theta^{(i)}:
      Result: \{X_i\}_{i\in\mathbb{N}} sampled from the MC
                                                                          1) Subsampling: Only store a subset of the \theta^{(i)}
      init: R(x|x')
                                                                          2) Gaussian Aproximation: We only keep:
      /* Good R choice \rightarrow fast convergence */
                                                                          \mu_i = \frac{1}{T} \sum_{i=1}^{T} \theta_i^{(j)} and \sigma_i = \frac{1}{T} \sum_{i=1}^{T} (\theta_i^{(j)} - \mu_i)^2
      init: X_0 = x_0
      for t \leftarrow 1, 2, \dots do
                                                                          And updete them online.
            x' \sim R(\cdot, x_{t-1})
                                                                          Predictive Esnable NNs:
            \alpha = \min \left\{1; \frac{Q(x')R(x_{t-1}|x')}{Q(x_{t-1})R(x'|x_{t-1})} \right\}
                                                                          Let \mathcal{D} = \{(x_i, y_i)\}_{i=1:n} be our dataset.
                                                                          Train \theta_i^{MAP} on \mathcal{D}_i with i = 1, ..., m
            with probability α do
            X_t = x';
                                                                          \mathcal{D}_i is a Bootstrap of \mathcal{D} of same size
           otherwise X_t = x_{t-1};
                                                                          and p(y^*|x^*, D) \simeq \frac{1}{m} \sum_{i=1}^{m} p(y^*|x^*, \theta_i^{MAP})
                                                                          6.3 Model calibration
Metropolis Adj. Langevin Algo (MALA):
                                                                          Train \hat{q} on \mathcal{D}_{train}
Energy function: P(x) = \frac{Q(x)}{Z} = \frac{1}{Z} \exp(-f(x))
                                                                          Evaluate \hat{q} on \mathcal{D}_{val} = \{(y', x')\}_{i=1:m}
We chose: R(x|x') = \mathcal{N}(x' - \tau \nabla f(x), 2\tau I)
                                                                          Held-Out-Likelihood \doteq \log p(y_{1:m}^{\prime m}|x_{1:m}^{\prime}, \mathcal{D}_{train})
Stoch. Grad. Langevin Dynamics (SGLD):
                                                                          \geq \mathbb{E}_{\theta \sim \hat{q}} \left[ \sum_{i=1}^{m} \log p(y_i'|x_i', \theta) \right] (Jensen)
We use SGD to Approximate \nabla f. Converges
                                                                          \simeq \frac{1}{k} \sum_{i=1}^{k} \sum_{i=1}^{m} \log p(y_i'|x_i', \theta^{(j)}), \ \theta^{(j)} \sim \hat{q}
 also without acceptance step
Hamilton MC: SGD performance improoved
                                                                          Evaluate predicted accuracy: We divide \mathcal{D}_{val}
by adding momentum (consider last step \nabla f)
                                                                          into bins according to predicted confidence va-
Gibbs sampling: Practical when X \in \mathbb{R}^n
                                                                          lues. In each bin we compare accuracy with
 Used when P(X_{1:n}) is hard but P(X_i|X_{-i}) is easy.
                                                                          confidence
                                                                          7 Active Learning
      init: x_0 \in \mathbb{R}^n; (x_0^{(B)} = x^{(B)}) B is our data
                                                                          Let \mathcal{D} be the set of observable points.
      for t = 1, 2, ... do
\begin{vmatrix} x_t = x_{t-1} \\ \text{with } i \sim \mathcal{U}(\{1:n\} \setminus B) * \text{ do} \end{vmatrix}
                                                                          We can observe S \subseteq \mathcal{D}, |S| \leq R
                                                                          Information Gain: \hat{S} = \arg \max_{S} F(S) = I(f, y_{S})
                x_{t-1}^{(i)} \sim P(x^{(i)}|x^{(-i)})
                                                                          For GPs: F(S) = \frac{1}{2} \log \left| I + \frac{1}{\sigma^2} K_{SS} \right|
                                                                          This is NP Hard, \Rightarrow Greedy Algo:
                                                                                init: S^* = \emptyset
 * if we do it \forall i \notin B no DBE but more practical
                                                                                for t = 1 : R do
 5.4 Variable elimination for MPE (most pro-
                                                                                      x_t = \operatorname{arg\,max}_{x \in \mathcal{D}} F(\mathcal{S}^* \cup \{x\})
        bable explanation):
                                                                                      (x_t = \operatorname{arg\,max}_{x \in \mathcal{D}} \sigma_x^2 | \mathcal{S} \text{ for GPs})
With loopy graphs, BP is often overconfi-
 dent/oscillates.
                                                                                      \left(x_t = \arg\max_{x \in \mathcal{D}} \frac{\sigma_{f|S}^2(x)}{\sigma_x^2(x)} \text{ for heter. GPs}\right)
 6 Bayesian Neural Nets
Likelihood: p(y|x;\theta) = \mathcal{N}(f_1(x,\theta), \exp(f_2(x,\theta)))
Prior: p(\theta) = \mathcal{N}(0, \sigma_n^2)
```

 $\theta_{MAP} = \arg\max\log(p(y,\theta))$ **6.1 Variation inference:** 

 $\blacksquare$   $\rightarrow$  Aletoric,  $\blacksquare$   $\rightarrow$  Epistemic

 $Q = \{q(\cdot|\lambda) = \prod_i q_i(\theta_i|\lambda), \lambda \in \mathbb{R}^d\}$ 

6.2 MCMC:

where  $q_i(\theta_i|\lambda) = p\delta_0(\theta_i) + (1-p)\delta_{\lambda_i}(\theta_i)$ 

Usually we use Q = Set of Gaussians

 $\hat{q} = \arg\max ELBO$  Reparameterization trick

*q* approx. the posterior but how to predict?

Gaussian Mixture distribution:  $\mathbb{V}(y^*|x^*, \mathcal{D}) \simeq$ 

 $\simeq \frac{1}{m} \sum_{i=1}^{m} \sigma^{2}(x^{*}, \theta^{(i)}) + \frac{1}{m} \sum_{i=1}^{m} \left( \mu(x^{*}, \theta^{(i)} - \overline{\mu}(x^{*})) \right)$ 

Dropouts Regularization: Random ignore no-

des in SGD iteration: Equavalent to VI with

This allows to do Dropouts also in prediction

 $p(y^*|x^*, \mathcal{D}) \simeq \frac{1}{m} \sum_{i=1}^{m} p(y^*|x^*, \theta^{(i)}), \ \theta \sim \hat{q}(\theta)$ 

F is <b>Submodular</b> if: $\forall x \in \mathcal{D}$ , $\forall A \subseteq B \subseteq D$ holds	9.4.2 Value iteration
that: $F(A \cup \{x\}) - F(A) \ge F(B \cup \{x\}) - F(B)$	while $  V_t - V_{t-1}   \le \epsilon$ do
F is Submodular $\Rightarrow F(\mathcal{S}^*) \ge \left(1 - \frac{1}{e}\right) F(\hat{\mathcal{S}})$	foreach $x \in \mathcal{X}$ , $a \in \mathcal{A}$ do $Q_t(x, a) \leftarrow r(x, a) +$
8 Bayesian Optimization	
Like Active Learning but we only want to find	foreach $x \in \mathcal{X}$ do
the optima. We pick $x_1, x_2,$ from $\mathcal{D}$ and observe $y_i = f(x_t) + \epsilon_t$ .	$V_t(x) \leftarrow \max_a Q_t(x, a)$
T	$\hat{\pi} = \pi_{\mathcal{G}}^{V_T}$ ; where $V_T$ last found Value
Comulative regret: $R_T = \sum_{t=1}^{I} \left( \max_{x \in \mathcal{D}f(x) - f(x_t)} \right)$	
$\mathbf{Oss:} \frac{R_T}{T} \to 0 \Rightarrow \max_t f(x_t) \xrightarrow{t-1} \max_{x \in \mathcal{D}} f(x)$	9.5 Partialy Observable MDP (POMDP) POMDP can be seen as MDP where:
8.1 Upper Confidence Sampling	1) $\mathcal{X}_{POMDP}$ are prob. distribution over $\mathcal{X}_{MDP}$
With GP $x_t = \arg\max_{x \in \mathcal{D}} \mu_{t-1}(x) + \beta_t \sigma_{t-1}(x)$	2) the actions are the same 3) $r_{POMDP}(b, a) = \mathbb{E}_{x \sim b}[r_{MDP}(x, a)]$
Chosing the correct $\beta_t$ we get: $\frac{R_T}{T} = \mathcal{O}\left(\sqrt{\frac{\gamma_T}{T}}\right)$ .	4) Trans. model: $b_{t+1}(x) = \mathbb{P}(X_{t+1} = x   y_{1:t+1}, a_t)$
Where $\gamma_t = \max_{ S  < T} I(f; y_S)$ . On d dims:	$b_{t+1}(x) = \frac{1}{Z} p(y_{t+1} X_{t+1} = x) \sum_{x' \in \mathcal{X}_{MDP}} p(x x', a_t) b_t(x')$
Linear: $\gamma_T = \mathcal{O}(d \log T)$ RBF: $\gamma_T = \mathcal{O}((l \log T)^{d+1})$	<b>How to solve?</b> Discretize $\mathcal{X}_{POMDP}$ and treat it as a MDP or Policy gradient techniques
<b>Optimal</b> $\beta_t = \mathcal{O}(\ f\ _K^2 + \gamma_t \log^3 T)$	10 Non Parametric RL
Oss: $\beta \uparrow =$ more exploration	It is an MDP with unknown $p(x' x, a)$ and $r(x, a)$
8.2 Thompson Samling	10.1 Model-based RL
$x_t = \operatorname{argmax}_{x \in \mathcal{D}} \tilde{f}(x), \ \tilde{f} \sim p(f x_{1:n}, y_{1:n})$	From all steps $X_{t+1}$ , $R_t   X_t$ , $A_t$ we can learn:
9 Markov Decision Process (MDP)	$p(x' x,a) \simeq \hat{p}_{x' x,a} = \frac{Count(X_{t+1}=x', X_t=x, A_t=a)}{Count(X_t=x, A_t=a)}$
9.1 Definitions	$r(x,a) \simeq \hat{r}_{x,a} = \frac{1}{Count(X_t=x, A_t=a)} \sum_{t X_t=x, A_t=a} R_t$
$\mathcal{X} = \{1, \dots, n\}$ states; $\mathcal{A} = \{1, \dots, m\}$ actions;	How to chose $a_t$ ?
p(x' x,a) transition probability;	<b>10.1.1</b> $\epsilon$ -greedy (On-Policy) With probability $\epsilon$ , pick random action.
$r(x,a)$ reward (can be random); $\pi: \mathcal{X} \to \mathcal{A}$ policy; $T^{\pi} \in \mathbb{R}^{n \times n}, \ T^{\pi}_{ij} = p(j i,\pi(i))$ Transition Matrix:	With probability $1 - \epsilon$ , pick $a = \arg \max Q(x, a)$ .
$J(\pi) = \mathbb{E}\left[\sum_{i=0}^{\infty} \gamma^{i} r(X_{i}, \pi(X_{i}))\right]$ Expected value:	<b>Oss:</b> $Q$ is calculated from $(\hat{p}, \hat{r})$
$V^{\pi}: \mathcal{X} \to \mathbb{R}, \ x \mapsto J(\pi X_0 = x)$ Value function;	<b>Th:</b> If $\epsilon_t \xrightarrow{RM} 0$ then $(\hat{r}, \hat{p}) \xrightarrow{a.s.} (r, p)$
$Q^V(x,a) = r(x,a) + \gamma \sum_{x \in \mathcal{X}} p(x' x,a)V(x)$ Q func;	10.1.2 Softmax (On-Policy)
$\pi_G^V(x) = \arg\max_a Q^V(x, a)$ greedy policy w.r.t. V;	Draw $a \sim q(a x) = \operatorname{softmax} \frac{Q(x,a)}{\tau}$
9.2 Value function Theorem	If $\tau \uparrow$ it means I trust less $Q$
$V^{\pi}(x) = r(x, \pi(x)) + \gamma \sum_{x' \in \mathcal{X}} p(x' x, \pi(x)) V^{\pi}(x')$	10.1.3 $R_{max}$ algorithm (On-Policy)
Matrix formulation: $(I - \gamma T^{\pi})V^{\pi} = r^{\pi}$	We add a fairy state $x^*$
9.3 Bellman Theorem	init: $r(x,a) = R_{max} \ \forall x \in \mathcal{X} \cup \{x^*\}, a \in \mathcal{A}$ init: $p(x^* x,a) = 1 \ \forall x \in \mathcal{X}, a \in \mathcal{A}$
1) $\pi^*$ , $V^*$ are optimal policy and it's value func.	<b>init:</b> $\pi$ = optimal policy w.r.t. $p$ , $r$
2) $\pi^* = \pi_G^{V^*}$	repeat   Execute $\pi$ and get $x_{t+1}$ and $r_t$
3) $V^*(x) = max_a [r(x, a) + \gamma \sum_{x \in \mathcal{X}} p(x' x, a)V^*(x)]$	Update belief of $r(x_t, \pi(x_t))$ and
$1) \Leftrightarrow 2) \Leftrightarrow 3)$	$p(x_{t+1} x_t,\pi(x_t))$
9.4 Algorithms	If obeserved 'enough' in $(x, a)$
9.4.1 Policy iteration	recompute $\pi$ using the updated belief only in $(x, a)$
while no more changes do	until;
$\pi \leftarrow \pi_{\mathcal{G}}^{V} \text{ (Update the Policy)}$	'Enough'? See Hoeffding's inequality
$V \leftarrow (I - \gamma T^{\pi})^{-1} r^{\pi}) \text{ (Update the value)}$	$(\hat{p} \in [0,1], \hat{r} \in [0,R_{max}]).$
_ value)	<b>PAC bound:</b> With probability $1 - \delta$ , $R_{max}$ will reach an $\epsilon$ -optimal policy in a number of steps
	reach an e optimal policy in a number of steps

```
Learn \pi^* only via V^* or Q^{V^*}
10.2.1 TD-learning (On-Policy)
Given a policy \pi we want to learn V^{\pi}
V^{\pi}(x) = \mathbb{E}_{R \sim r(x, \pi(x)), X' \sim p(\cdot | x, \pi(x))} [R + \gamma V^{\pi}(X')]
After seeing (x_{t+1}, r_t | x_t, \pi(x_t)) we update:
V_{t+1}(x_t) \leftarrow (1 - \alpha_t)V_t(x_t) + \alpha_t(r_t + \gamma V_t^{\pi}(x_{t+1}))
Where \alpha_t is a regulizer term (only 1 samlple)
Th: If \alpha_t \xrightarrow{RM} 0 then V \xrightarrow{a.s.} V^{\pi}
10.2.2 Q-learning (Off Policy)
Given experience we want to learn Q^* = Q^{V^*}
Q^*(x, a) = \mathbb{E}_{R \sim r(x, \pi(x))} [R + \gamma \max_{a'} Q^*(X', a')]
                   X' \sim p(\cdot | x, \pi(x))
After seeing (x_{t+1}, r_t | x_t, a_t) we update:
 Q(x_t, a_t) \leftarrow (1 - \alpha_t)Q(x_t, a_t) + \alpha_t(r_t + \gamma \max_{a'} Q(x_{t+1}, a'))
Th: If \alpha_t \xrightarrow{RM} 0 then Q \xrightarrow{a.s.} Q^* Optimistic Q learning:
Initialize: Q(x,a) = \frac{R_{max}}{1-\nu} \prod_{t=1}^{T_{init}} (1-\alpha_t)^{-1}
Same convergence time as with R_{max}. Memory
O(|X||A|). Comp: O(|A|).
11 Parametric Model Free RL
11.1 Parametric TD-learning
11.1.1 TD-learinging as SGD
TD-learing = 1 sample (x', r|x, \pi(x)) SGD on:
\bar{l}_2(V;x,r) = \frac{1}{2} \left( V - r - \gamma \mathbb{E}_{x' \sim p(\cdot|x,\pi(x))} \left[ \hat{V}^{\pi}(x') \right] \right)^2
1 sample estimate of \nabla_V \bar{l}_2 = \delta = V - r - \gamma \hat{V}^{\pi}(x')
\Rightarrow V \leftarrow V - \alpha_t \delta \text{ where } V = \hat{V}^{\pi}(x)
11.1.2 TD-parametric
If \hat{V}^{\pi}(x) = V(x, \theta) then:
\delta = [V(x;\theta) - r - \gamma V(x';\theta_{old})] \nabla_{\theta} V(x,\theta)
11.2 Parametric O-learining
\delta(\theta, \theta_{old}) = (Q(x, a; \theta) - r - \gamma \max_{a'} Q(x', a'; \theta_{old})) 12.1 Planning
We don't differiantiate with regard to \theta_{old}
The SGD step is: \theta \leftarrow \theta - \alpha_t \delta(\theta, \theta) \nabla_{\theta} Q(x, a; \theta)
Deep Q Networks (DQN): Version of Q-
learning where we update Q only each batch:
L(\theta) = \sum_{(x,a,r,x')\in\mathcal{D}} (r + \gamma \max_{a'} Q(x',a';\theta_{old}) - Q(x,a;\theta))^2
Double DQN (better):
L(\theta) = \sum_{(x,a,r,x') \in \mathcal{D}} (r + \gamma Q(x', a^*(\theta); \theta_{old}) - Q(x,a;\theta))^2
where: a^*(\theta) \doteq \arg \max_{a'} Q(x', a'; \theta)
                                                                                 one
11.3 Policy-Search method
\pi(x) = \pi(x;\theta); \ r(\tau) = \sum_{t=0}^{T} \gamma^t r(x_t, a_t)
J(\theta) \doteq J(\pi(\cdot; \theta)) = \mathbb{E}_{\tau \sim \pi_{\theta}}[r(\tau)] \simeq \frac{1}{m} \sum_{i=1}^{m} r(\tau^{(i)})
with \pi_{\theta}(\tau) := p(x_0) \prod_{t=1}^{T} \pi(a_t | x_t; \theta) p(x_{t+1} | x_t, a_t)
\hat{\theta} = \arg\max_{\theta} J(\theta) \Rightarrow SGD
```

that is polynomial in |X|, |A|, T,  $1/\epsilon$  and  $log(1/\delta)$ . 11.3.1 Theory results:

Memory  $O(|X|^2|A|)$ .

10.2 Model-free RL

```
11.3.2 REINFORCE
b(\tau_{0:t-1}) = \sum_{t'=0}^{t-1} \gamma^{t'} r_{t'}; \ r(\tau) - b(\tau_{0:t-1}) = \gamma^t G_t
Let G_t = \sum_{t'=t}^{T} \gamma^{t'-t} r_{t'} = r_t + \gamma r_{t+1} + \gamma^2 r_{t+2} + \dots
\nabla_{\theta} J(\theta) = \mathbb{E}_{\tau \sim \pi_{\theta}} \left[ \sum_{t=0}^{T} \gamma^{t} G_{t} \nabla_{\theta} \log \pi(a_{t} | x_{t}; \theta) \right]
SGD: \theta \leftarrow \theta + \eta_t \gamma^t G_t \nabla_{\theta} \pi(a_t | x_t; \theta) \quad \forall t = 1 : T
Oss: G_t \leftarrow G_t - \frac{1}{T} \sum_{t'=0}^{T-1} G_t' \Rightarrow \text{Less Variance}
11.4 Actor Critic
After seeing (x, a, r, x'):
\theta_{\pi} \leftarrow \theta_{\pi} + \eta_t Q(x, a, \theta_O) \nabla_{\theta_{\pi}} \log \pi(a|x; \theta_{\pi})
\theta_O \leftarrow \theta_O - \eta_t \delta \nabla_{\theta_O} Q(x, a; \theta_O)
        \delta = Q(x, a; \theta_{O}) - r - \gamma Q(x', \pi(x', \theta_{\pi}), \theta_{O})
11.4.1 Theory justification:
\mathbb{E}_{\tau_{t+1}, \infty} \sim \pi_{\theta} \left[ G_t | a_t, x_t \right] = Q^{\pi_{\theta}} (x_t, a_t)
\nabla_{\theta_{\pi}} J(\theta_{\pi}) = \mathbb{E}_{\tau \sim \pi_{\theta_{\pi}}} \left[ \sum_{t=0}^{T} \gamma^{t} \nabla_{\theta_{\pi}} \log \pi(a_{t} | x_{t}; \theta_{\pi}) Q(x_{t}, a_{t}) \right]
\nabla_{\theta_{\pi}} J(\theta_{\pi}) = \mathbb{E}_{\tau \sim \pi_{\theta_{\pi}}, x \sim \rho(\cdot)} \left[ \nabla_{\theta_{\pi}} \log \pi(a_t | x_t; \theta_{\pi}) Q(x_t, a_t) \right]
where \rho(x) = \sum_{t=0}^{\infty} \gamma^t p(x_t = x).
11.4.2 Advantage Function Baseline
We can us V(x;\theta_V) as baseline for Q(x,a;\theta_O)
A^{\pi}(x,a) \doteq Q^{\pi}(x,a) - V^{\pi}(x)
Oss: \forall \pi \forall x \max_a A^{\pi}(x, a) \geq 0
Oss: \pi optimal \forall x A^{\pi}(x, \pi(x)) \leq 0.
11.4.3 Off policy variation (DDPG)
\hat{\theta}_{\pi} = \arg \max J(\theta_{\pi}); J(\theta_{\pi}) = \mathbb{E}_{x \sim \mu} \left[ Q(x, \pi(x; \theta_{\pi}); \theta_{Q}) \right]
where \mu visits all states. SGD can now be done
in Batches and Off-policy also for \theta_{\pi}:
\theta_{\pi} \leftarrow \theta_{\pi} + \eta \nabla_{\theta_{\pi}} \frac{1}{|B|} \sum_{(x,a,r,x') \in B} Q(x,\pi(x,\theta_{\pi});\theta_{Q})
Oss: \pi randomized \Rightarrow reparametrization trick.
J_{\lambda}(\theta) = J(\theta) + \lambda H(\pi_{\theta}) to encourage exploration
12 Parametric Model Based RL
J(a_{0:\infty}) := \sum_{t=0}^{\infty} \gamma^t r(x_t, a_t)
12.1.1 Reciding Horizon (MPC)
 \max \sum_{\tau=t}^{t+H-1} \gamma^{\tau-t} r_{\tau}(x_{\tau}, a_{\tau}) \text{ s.t. } x_{\tau+1} = f(x_{\tau}, a_{\tau})
And then we carry out action a_t.
\nabla \uparrow : \nabla of f \circ \cdots \circ f easy vanishes or explodes.
Random shooting: Pick some set of actions at
random and chose the first action of the best
12.1.2 Value Function estimates
We look at value function after the horizon H:
We maximize J_H(a_{t:t+H-1}) :=
\sum_{\tau=t}^{t+H-1} \gamma^{\tau-t} r_{\tau}(x_{\tau}(a_{t:\tau-1}), a_{\tau}) + \gamma^{H} V(x_{t+H})
Microth: If V is correct then the chosen action
is optimal \forall H
```

 $\nabla_{\theta} J(\theta) = \nabla_{\theta} \mathbb{E}_{\tau \sim \pi_{\theta}} [r(\tau)] = \mathbb{E}_{\tau \sim \pi_{\theta}} [r(\tau) \nabla_{\theta} \log \pi_{\theta}(\tau)] =: \clubsuit$ 

Oss: Value can estimated with off-policy Techniques (TD-learing)

## 12.1.3 Planning (stochastic model)

$$\mathbb{E}_{x_{t+1:t+H}} \left[ \sum_{\tau=t}^{t+H-1} \gamma^{\tau-t} r_{\tau}(x_{\tau}, a_{\tau}) + \gamma^{H} V(x_{t+H}) | a_{t:t+H-1} \right]$$

maximise this doing SGD by sampling or random shooting.

## 12.1.4 Parametric policy (stoch. model)

$$\mathbb{E}_{x_{t+1:t+H}} \left[ \sum_{\tau=t}^{t+H-1} \gamma^{\tau-t} r_{\tau}(x_{\tau}, \pi_{\theta}(x_{\tau})) + \gamma^{H} Q(x_{t+H}, \pi_{\theta}(x_{t+H})) | \theta \right]$$

Oss:  $H=0 \Rightarrow DDPG$  objective

### 12.2 Model Learning

**Insight:** Markov  $\Rightarrow \bar{D} = \{(x_i, a_i, r_i, x_{i+1})_i\}_{i=1:t}$  is a conditional  $\perp$  dataset

Learn 
$$f$$
,  $\mathcal{R}$  s.t.  $x_{i+1} \leftarrow f(x_i, a_i)$ ;  $r_i \leftarrow \mathcal{R}(x_i, a_i)$ 

#### 12.2.1 Neural Networks

$$x_{t+1} \sim \mathcal{N}(\mu(x_t, \underline{a_t}, \theta), \Sigma(x_t, a_t, \theta))$$

Where  $\Sigma = CC^T$ , *C* lower triangular.

**Pittfall:** using MAP estimate we don't capture epistemic uncertainty ⇒ Bayesian Learning

#### Greedy exploitation:

## repeat

plan  $\pi$  to maximize  $\mathbb{E}_{f \sim p(\cdot | \mathcal{D})} J_H(\pi, f)$ 

rollout  $\pi$  to collect more data update  $p(\cdot|\mathcal{D})$ 

### until:

To find  $\pi$  we use SGD or Trajectory samling (sampling also  $f \sim p(\cdot|\mathcal{D})$ )

### 12.2.2 Exploration?

**Thompson:**  $\mathbb{E}_{f \sim p(\cdot | \mathcal{D})} J_H(\pi, f) \simeq J_H(\pi, \hat{f}), \ \hat{f} \sim p(\cdot | \mathcal{D})$ 

**Optimism:** Let  $M(\mathcal{D})$  be set of most plausible models given  $\mathcal{D}$ . then:

$$\mathbb{E}_{f \sim p(\cdot | \mathcal{D})} J_H(\pi, f) \simeq \max_{f \in M(\mathcal{D})} J_H(\pi, f)$$

We do Greedy exploitation but joint maximization  $\max_{\pi} \max_{f \in M(\mathcal{D})} J_H(\pi, f)$  is hard:

H-URCL: 
$$\tilde{f}(x,a) = \mu(x,a) + \beta_{t-1}\Sigma(x,a)\eta(x,a)$$
.  
then  $\pi_t^{H-URCL} = \arg\max_{\pi} J_H(\tilde{f},\pi)$ 

# 12.3 Safe Exploration planning

With confidence bounds I make  $1-\delta$  sure in H step I can be safe. Issue, long term dependency **Liapunov:** I make sure I can reach a Liapunov state