

Cryptographically Privileged State Estimation With Gaussian Keystreams

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Abstract—State estimation via public channels requires additional planning with regards to state privacy and the information leakage of involved parties. In some scenarios it is desirable to allow partial leakage of state information, thus distinguishing between privileged and unprivileged estimators and their capabilities. Existing methods identifying privileged and unprivileged or trusted and untrusted estimators typically result in reduced estimation quality for both parties or require additional secure channels of communication. We introduce a method to diminish the estimation quality at unprivileged estimators using a stream of pseudorandom Gaussian samples while leaving privileged estimation unaffected and requiring no additional transmission beyond an initial key exchange. A cryptographic definition of privileged estimation is first presented, capturing the definition of the difference between estimation qualities, before we develop a Gaussian keystream privileged estimation scheme meeting the security criteria defined. Achieving privileged estimation without additional channel requirements allows quantifiable estimation to be made available to the public while keeping the best estimation private to trusted privileged parties and can find uses in a variety of service providing and privacy-preserving scenarios.

I. INTRODUCTION

The role of state estimation and sensor data processing has become increasingly prevalent in modern systems [1]. Particularly, since the development of Kalman estimation theory, Bayesian state estimation has found common application, varying from autonomous systems to remote estimation [2], [3]. As advancements in distributed algorithms and cloud computing develop, the use of wireless and public communication channels for data transmission has become widespread, bringing to light the requirements of data privacy and state secrecy [4], [5].

Typically, use-cases for preserving data privacy over public channels involve hiding all transferred information such that eavesdroppers or present untrusted parties learn no additional information from the observed data. This is achievable using common encryption schemes such as AES [6] or RSA [7] which formally capture this security requirement by satisfying cryptographic ciphertext indistinguishability [8]. Some more advanced requirements using public channels also exist. When partial information needs to be made public, or computations need to be performed on data, homomorphic or aggregation encryption schemes can be used. [9] proposes localisation

where individual sensors and measurements remain private using homomorphic encryption. [10] has distributed control inputs aggregated without parties learning individual contributions, while [11], [12] use homomorphic encryption to allow control inputs to be computed without decryption. In the context of estimation, a quantifiable difference in estimation performance between untrusted and trusted parties can provide different levels of estimation privilege. This was achieved in the initial release of the Global Positioning System (GPS) [13], which relied on a second, encrypted, stream of information required for accurate estimation. Another example is seen in [14], where intermittent sending is proposed to increase eavesdropper estimation error while using a secure feedback channel to ensure better estimation for trusted parties. Our contribution in this work considers this context of privileged and unprivileged estimation and is comprised of a formal definition of privileged state estimation, crucial for providing cryptographic security albeit often neglected, before proposing an estimation scheme which provides privileged levels of estimation without reliance on additional secure channels or privileged estimator feedback. We accompany the method with a cryptographic proof sketch and simulation results.

In section II we introduce the relevant privileged estimation problem followed by a cryptographic formalisation of desired properties. Section III introduces our proposed privileged estimation scheme and in section IV a cryptographic proof sketch is given. A simulation of the method is then demonstrated and explained in section V, while concluding remarks and future work are discussed in section VI.

A. Notation

Throughout this work the following notation is used. Lowercase underlined characters \underline{a} denote vectors, while uppercase bold characters \mathbf{M} denote matrices. $\mathbf{M} \succ 0$ and $\mathbf{M} \succeq 0$ denote positive definitiveness and positive semi-definitiveness, respectively, and $\mathbf{M} \succ \mathbf{N}$ is used as shorthand for $\mathbf{M} - \mathbf{N} \succ 0$. Function $\text{eig}(\mathbf{M})$ gives the set of eigenvalues for matrix \mathbf{M} , $\text{Cov}[\cdot]$ computes the covariance of a random vector and \mathbf{I} is the identity matrix with size inferrable from context.

II. PROBLEM STATEMENT

We consider the estimation scenario with known process and measurement models, where state estimators are either privileged estimators possessing a secret key, or unprivileged estimators without. We aim to develop a scheme for which the difference in their estimation errors is quantifiable and cryptographically guaranteed when process and measurement models are Gaussian, linear and time-invariant.

The process model we consider gives the state $\underline{x}_k \in \mathbb{R}^n$ at a time k and is given by

$$\underline{x}_k = \mathbf{F}\underline{x}_{k-1} + \underline{w}, \quad (1)$$

with noise term $\underline{w} \sim \mathcal{N}(\underline{0}, \mathbf{Q})$ and a known covariance $\mathbf{Q} \in \mathbb{R}^{n \times n}$. Similarly, the measurement model gives the measurement \underline{y}_k at time k and is given by

$$\underline{y}_k = \mathbf{H}\underline{x}_k + \underline{v}, \quad (2)$$

with noise term $\underline{v} \sim \mathcal{N}(\underline{0}, \mathbf{R})$ and a known covariance $\mathbf{R} \in \mathbb{R}^{m \times m}$.

To capture our aim of creating a privileged and unprivileged estimator, we must first define how to assess the estimation advantage between estimators, and which algorithms are required to characterise a privileged estimation scheme. In the following section, we give relevant formal definitions, which will be used when assessing the security of our proposed scheme.

A. Formal cryptographic Problem

While we are interested in Gaussian and linear models, it is more practical to define a broader security notion that can be satisfied under arbitrary specified conditions on the models. This allows the use of the security notion in future literature and is more in-line with typical cryptographic practice. We will later show that our proposed scheme meets this security notion under the specific Gaussian and linear model assumptions.

Typical formal cryptographic security notions are given in terms of probabilistic polynomial-time (PPT) attackers and capture desired privacy properties as well as attacker capabilities [8]. The most commonly desired privacy property, cryptographic indistinguishability, is not suitable for our estimation scenario due to our desire for unprivileged estimators to gain some information from measurements. Instead, we will define security with respect to a time series of covariances, given arbitrary known Bayesian models, such that the difference in estimation error between estimators with and without the secret key is bounded by the series at all times.

To formalise this, we introduce the following notations and definitions. We assume the existence of an arbitrary process (not necessarily Gaussian or linear) following a known model exactly, with the state at time k denoted $\underline{x}_k \in \mathbb{R}^n$ and model parameters \mathcal{M}_P . Similarly, we assume the existence of a means of process measurement following a known measurement model exactly, with the measurement at time k denoted $\underline{y}_k \in \mathbb{R}^m$ and model parameters \mathcal{M}_M . We can now define a *privileged estimation scheme* as a pair of algorithms (Setup, Noise) given by

Setup($\mathcal{M}_P, \mathcal{M}_M, \kappa$) On the input of models \mathcal{M}_P and \mathcal{M}_M , and the security parameter κ , public parameters pub and a secret key sk are created.

Noise($\text{sk}, k, \mathcal{M}_P, \mathcal{M}_M, \underline{y}_1, \dots, \underline{y}_k$) On input of secret key sk , time k , models \mathcal{M}_P and \mathcal{M}_M , and measurements $\underline{y}_1, \dots, \underline{y}_k$, a noisy measurement \underline{y}'_k (with no required model constraints) is created.

In addition to the scheme description above, we also give the following definitions to help formalise our desired security notion.

Definition II.1. An *estimator* is any algorithm which produces a guess of the state \underline{x}_k for a given time k .

Definition II.2. A *negligible covariance function* is a function

$$\text{neglCov}_m(\kappa) : \mathbb{N} \rightarrow \mathbb{R}^{m \times m} \quad (3)$$

that returns a matrix \mathbf{A} such that \mathbf{A} is a valid covariance ($\mathbf{A} \succ 0$ and $\mathbf{A} = \mathbf{A}^\top$) and that for each of its eigenvalues $e \in \text{eig}(\mathbf{A})$, there exists a negligible function η such that $e \leq \eta(\kappa)$.

With the terminology above, we can now introduce the security notion which captures the formal requirements of the estimation problem that we want to solve.

Definition II.3. A privileged estimation scheme meets notion $\{\mathbf{D}_1, \mathbf{D}_2, \dots\}$ -*Covariance Privilege for Models \mathcal{M}_P and \mathcal{M}_M* if for any PPT estimator \mathcal{A} , there exists a PPT estimator \mathcal{A}' , such that

$$\begin{aligned} & \text{Cov} \left[\mathcal{A} \left(k, \kappa, \text{pub}, \mathcal{M}_P, \mathcal{M}_M, \underline{y}'_1, \dots, \underline{y}'_k \right) - \underline{x}_k \right] \\ & - \text{Cov} \left[\mathcal{A}' \left(k, \kappa, \text{pub}, \mathcal{M}_P, \mathcal{M}_M, \underline{y}_1, \dots, \underline{y}_k \right) - \underline{x}_k \right] \\ & \succeq \mathbf{D}_k + \text{neglCov}_m(\kappa) \end{aligned} \quad (4)$$

for valid covariances $\mathbf{D}_1, \mathbf{D}_2, \dots$ and some negligible covariance for all $k > 0$. Here, estimators \mathcal{A} and \mathcal{A}' are running in polynomial-time with respect to the security parameter κ , and all probabilities are taken over models \mathcal{M}_P and \mathcal{M}_M , estimators \mathcal{A} and \mathcal{A}' , and algorithms Setup and Noise.

Informally, the above definition states that no estimator that can only access noisy measurements $\underline{y}'_1, \dots, \underline{y}'_k$ can estimate a state \underline{x}_k for a time k with a mean square error (MSE) covariance less than an equivalent estimator with access to true measurements $\underline{y}_1, \dots, \underline{y}_k$, by a margin of at least \mathbf{D}_k . Next, we will propose a scheme meeting the aforementioned notion for a derivable series of covariances when models \mathcal{M}_P and \mathcal{M}_M are Gaussian and linear.

III. PRIVILEGED ESTIMATION

General idea
(picture ?)

Use a cryptographically secure key stream to generate pseudorandom Gaussian samples

Samples are used to increase the uncertainty of estimation and are known and removable only by those with the key used to generate them

A. Gaussian Keystream

To generate pseudorandom Gaussian samples, we rely on first generating a traditional pseudorandom bitstream given a secret key.

Using well-studied methods for the generation of pseudorandomness guarantees robustness and an easy means of updating only the relevant component when the methods used are no longer considered safe.

Any implementation of a cryptographic stream cipher can be used for our purpose and will produce a stream of bits typically combined with plaintexts to provide secure encryption.

Rather than encrypting plaintext, we interpret the bitstream as sequential pseudorandom integers g and use these to generate pseudorandom uniform real numbers in the range $(0,1)$. u

While the uniform real samples are only approximated by floating-point numbers in the conversion from integers, we argue this is sufficiently random and discuss this further in the Security section.

Finally, independent standard Gaussian samples can then be generated from the uniform real numbers using the Box-Muller transform, and are ready to be used by our sensor and privileged filter. z

B. Additional Gaussian Noise

To use the pseudorandom Gaussian samples at the sensor and privileged estimator, they need to be converted to multivariate Gaussian samples suitable for use in the measurement model and need a means of controlling how much uncertainty is added to the unprivileged estimators.

We define the additional noise term $Z > 0$ and can transform the Gaussian samples z into pseudorandom samples o of a multivariate zero-mean Gaussian distribution with covariance Z .

Before estimation, we assume that a secret key is shared between the sensor and the privileged estimator.

During estimation, the sensor modifies its measurements at each timestep.

There are now two estimation problems present for the privileged and unprivileged estimators respectively.

For the privileged estimator who holds the shared secret key, values z , and therefore o , can be computed at any time k and received measurements modified to their original form. This in turn results in exactly the measurement model from the problem formulation.

In the case where pseudorandomness is indistinguishable from randomness, as is the case at an unprivileged estimator when using cryptographically secure Gaussian keystreams and the secret key is not known, the measurement model noise covariance can now be written as $R + Z$.

Intuitively, we can already see that the two estimators will have an estimation error covariance differing by some value dependent on Z at each time k . In the security section, we will show that the best possible error covariances achievable by the privileged and unprivileged estimators can be computed exactly by computing the Cramér–Rao lower-bound for both

measurement models and that the difference between them will give an exact lower-bound on the difference between the two estimator error covariances.

C. Multiple Privileges

In the above scenario, we have considered a single level of estimation privileged with one private key, dividing estimation error covariance into two groups; privileged and unprivileged estimators.

As a direct extension, it may be desirable to define multiple levels of privilege, such that the best estimation performance would depend on the privilege level of the estimator.

Here we will discuss the case of multiple privilege levels where a single secret key corresponds to each level, and where noise is added in the same manner as above, for each key individually.

N noise terms are added to the original measurement equation, with variances Z_i , $0 \leq i < N$.

From the equation, we can see that obtaining any single key sk_i would lead to a measurement model where only a single pseudorandom Gaussian noise sample, of variance Z_i , is removed.

This restricts possible estimation error bounds of each privilege level due to the dependence of measurement noise at an estimator with key sk_i , on the remaining noise terms Z_j , $j \neq i$.

If we write the covariances of added measurement model noise for holder of each key sk_i as E_i , we can capture this dependance as $E_i = \sum_{j=0, j \neq i}^{N-1} Z_j$ where both $E_i > 0$ and $Z_i > 0$.

Since choosing values E_i directly controls the estimation error differences between privileged levels, we are interested in the numerical restrictions on $E_i > 0$ which will produce valid covariances $Z_j > 0$, that can be used when adding noise at the sensor.

The dependencies between the covariances can be captured by the block matrix equation.

...equation and also block matrix inequality (might need some defining as it uses \prec)

Since we require $Z_i > 0$ for all $0 \leq i < N$, the restriction on the choices of privilege level additional noises E_i can be rewritten as

$$E_i \prec \frac{1}{N-2} \sum_{i=0}^{N-1} E_i \quad (5)$$

for all $0 \leq i < N$.

Alternative methods involving multiple or overlapping keys among privilege levels may allow choices of E_i to be less restricted than in the equation above. We have chosen the case with a single shared key per privilege level due to its simplicity and ability to change privilege estimation performance without needing additional key redistribution, and leave variants with fewer restrictions than () as future work.

IV. SCHEME SECURITY

The security of the proposed scheme will be primarily considered in the single privileged estimation level as introduced in section ().

A proof sketch will be provided to show how the proposed scheme meets the cryptographic notion in section ().

The extension to multiple privilege levels as described in the section () will be informally discussed afterwards.

A. Single Privileged Case

Recalling the introduced security notion in section (2), we aim to show how our introduce privileged estimation scheme, given conditions on \mathcal{M}_P and \mathcal{M}_M , meets the desired security for a computable series $\mathbf{D}_1, \dots, \mathbf{D}_k$ dependent on the additional noise variable Z .

We consider the process model (1) and measurement model (2) exactly, that is, any linear models with known zero-mean Gaussian additive noises. This information is captured in \mathcal{M}_P and \mathcal{M}_M and defines our conditions on the models.

The two required algorithms for the privileged estimation scheme, Setup and Noise, are defined such that Setup initialises the stream cipher with security parameter κ , sets the secret key sk to that of the cipher and includes initial filter estimate \hat{x}_0 , error covariance \mathbf{P}_0 and additional noise variance Z in the public parameters pub .

Here we note that including the initial state, error covariance and added noise variance is not a requirement for the security of our scheme, but rather just a means of making relevant estimation parameters public for completeness.

The Noise function is then given by () where y'_k is the measurement after adding a pseudorandom Gaussian sample using the stream cipher at time k .

The idea behind the proof relies on the Cramér–Rao lower bound (CRLB). The CRLB gives the smallest error covariance, with respect to root mean square error (RMSE), achievable for *any* estimator when all measurements y_1, \dots, y_k are observed []. Notably, the CRLB can be computed exactly when process and measurement models are linear and Gaussian.

The CRLB can also be computed recursively for time k , in which case it reduces to the posterior estimate error covariance at time k as given by the linear Kalman Filter []. This is given by

...recursive equations for the update covariance of the KF (predict and update combined, but predict first). (Note is this true?? Might need biased CRLB - when is it the best?? What assumptions are made??)

which gives us a value \mathbf{P}_k at time k , such that

$$\mathbf{P}_k \preceq \text{Cov} \left[\mathcal{A} \left(k, \mathcal{M}_P, \mathcal{M}_M, y_1, \dots, y_k \right) - \underline{x}_k \right] \quad (6)$$

for any estimator \mathcal{A} following definition II.1 and any Gaussian, linear, time-invariant models \mathcal{M}_P and \mathcal{M}_M .

This leads us into our sketch proof.

Proof sketch: (LaTeXdescription? sub(sub)section?)

As we use a cryptographically pseudorandom stream cipher, the stream integers g and generated pseudorandom uniformly

distributed floating point numbers u are indistinguishable from random integers and floating-point number by any polynomially bound estimator.

Additionally, we argue that the uniformly distributed floating point numbers u are sufficiently close to real standard uniform numbers $\mathcal{U}(0,1)$ for all the necessary estimation properties and lower bounds to hold. This is the standard in estimation tasks due to the inexistence of true real numbers on modern computer hardware.

In turn, this leads to pseudorandom noisy measurements y'_k that are indistinguishable from real measurements following the noisy measurement model () exactly.

We can now compute the CRLB recursively for both the true measurement model (), obtaining series $\{\mathbf{P}_1, \mathbf{P}_2, \dots\}$, and the modified model (), obtaining series $\{\mathbf{P}'_1, \mathbf{P}'_2, \dots\}$. Due to models' properties $R < R + Z$, and the properties of the CRLB, taking the difference of the two series for *any* initial covariance \mathbf{P}_0 produces the infinite series of valid covariances $\mathbf{D}_1, \mathbf{D}_2, \dots$, where

$$\mathbf{D}_k = \mathbf{P}'_k - \mathbf{P}_k. \quad (7)$$

Since both series \mathbf{P}_k and \mathbf{P}'_k give the lowest possible error covariance of respective estimators, an estimator following model () can always be created for an estimator following the modified model () such that their error covariances at any time k differs by at least \mathbf{D}_k .

A reduction proof can be easily constructed where the existence of an unprivileged estimator in our scheme that can produce estimates such that (4) does not hold, can be used to construct an estimator with an error covariance lower than \mathbf{P}'_k given the modified model. As we know that no such estimator exists, we conclude that our scheme meets $\{\mathbf{D}_1, \mathbf{D}_2, \dots\}$ -Covariance Privilege for Models \mathcal{M}_P and \mathcal{M}_M , when \mathcal{M}_P and \mathcal{M}_M are Gaussian, linear and time-invariant.

End proof sketch.

In addition to the proof sketch above, we stress caution when assuming accepting a cryptographic guarantee in terms of models \mathcal{M}_P and \mathcal{M}_M when used to predict a measured physical process. The following implicit assumptions would be made in such a scenario.

Exact models When assigning a model to a physical process, our scheme involving the model only guarantees the same security of estimating the physical process when it follows the model *exactly*. It is often the case that models are chosen to be Gaussian and linear in part to simplify estimation, resulting in the possibility of better estimation with alternative, more complicated, models. While it may be likely that the estimation of an unprivileged estimator is worse even when this is the case, it cannot be proved by our methodology here.

Bayesian interpretation We also note that the model requirements and security definitions we have put forward assume a Bayesian interpretation of probability theory. Although this is the standard for state estimation problems due to its applicability and performance, without stating

this assumption it may be arguable that the security definition 4 may not hold.

B. Multiple Additional Noises

While we have not defined a security notion for multiple levels of privileged estimation, an intuitive and informal extension can be described.

The security notion desired would require that for any subset of corrupted estimators, and therefore the knowledge of any subset of different privilege level secret keys $s \subset \{\text{sk}_i, 0 \leq i < N\}$ and noisy measurements, an estimator which is given true measurements can be constructed, such that the difference between the corrupted subset's error covariance and its own is at least $\mathbf{D}_k^{(s)}$ at time k .

Although this definition requires an infinite series for every possible subset of privilege level keys s , complicating its form specification, it captures the exact advantage of every such subset making it a more general definition.

Given the structure of our scheme in section () it can be readily seen how the above notion would be met. Similarly to the single level case, the CRLB can be used to compute the minimum error covariances for all compromised key subsets as well as for an estimator with the true measurements, and the relevant infinite covariance series' can be defined.

V. SIMULATION AND RESULTS

As well as showing the theoretical security of our scheme, we have simulated the estimation problem using linear Kalman filter estimators which achieve the CRLB error covariances for the different measurement models. Simulations have been implemented in the Python programming language and use the AES block cipher in CTR mode as a cryptographically secure stream cipher (AES-CTR) [].

We have considered two simulations, both following the same two-dimensional constant-velocity process model, given by

$$\mathbf{F} = \begin{bmatrix} 1 & 0.5 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0.5 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

and

$$\mathbf{Q} = 0.01 \cdot \begin{bmatrix} 0.0417 & 0.1250 & 0 & 0 \\ 0.1250 & 0.5000 & 0 & 0 \\ 0 & 0 & 0.0417 & 0.1250 \\ 0 & 0 & 0.1250 & 0.5000 \end{bmatrix},$$

with differing measurement models. In all cases, estimators were initialised with the same initial conditions, equal to the true starting condition of the processes they were estimating.

The first measurement model measures location and leads to an observable system with bounded CRLB error covariances as $k \rightarrow \infty$. It is given by

$$\mathbf{H} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 \end{bmatrix} \text{ and } \mathbf{R} = \begin{bmatrix} 5 & 2 \\ 2 & 5 \end{bmatrix},$$

and the sensor adds pseudorandom Gaussian samples with covariance

$$\mathbf{Z} = \begin{bmatrix} 35 & 0 \\ 0 & 35 \end{bmatrix}$$

to create an estimator privilege level. Figures () and () show the average error covariance traces and RMSE of estimation from 1000 runs of our privileged estimation scheme, respectively, where the above models are followed. It can be seen that the trace of the privileged estimator error covariance is lower than that of the unprivileged one. Both traces are equivalent to the those of respective estimator CRLB covariances and their difference equal to the trace of the estimation privilege covariance \mathbf{D}_k at any time k .

The second simulation considers an unobservable system where only the velocity is measured and has an unbounded CRLB error covariance as $k \rightarrow \infty$. It is given by

$$\mathbf{H} = \begin{bmatrix} 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

and uses the same values for \mathbf{R} and \mathbf{Z} as given for the previous model. Figures () and () show the average error covariance traces and RMSE of estimation from 1000 runs using this model and capture how error covariance boundedness does not affect the privileged estimation scheme's properties.

Lastly, a simulation of multiple privilege levels was also performed using the bounded error covariance measurement model () and using pseudorandom Gaussian sample with variances $\mathbf{Z}_0 = 20 \cdot \mathbf{I}$, $\mathbf{Z}_1 = 14 \cdot \mathbf{I}$ and $\mathbf{Z}_2 = 17 \cdot \mathbf{I}$. Note that the three matrices \mathbf{Z}_i , $0 \leq i < 3$ are such that () is satisfied. Figures () and () again show the average traces and RMSE of estimation from 1000 runs and display the distinct difference in estimation error of the different privilege levels. Additionally, two special case bounding estimators are included, one holding all privilege level keys and one holding none.

All of the included figures capture the difference in estimation error of the best possible estimators given the simulated processes (with respect to RMSE) and support the proposed security proof sketch given in section ().

VI. CONCLUSION

Presented the idea of privileged estimation and gave a formal cryptographic definition.

Demonstrated a concrete privileged estimation scheme and gave a proof sketch for it meeting the desired security notion.

Additionally supported the proof sketch by simulation using the best possible estimators for the simulated processes.

Future work includes reducing the requirement () for multiple levels of private estimation, discussing the formal security in this case further and implementing privileged estimation on hardware to demonstrate the real-time capability of the method.

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