Stochastic Processes

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Moment Generating Function

- Moment Generating Function: $\mathbb{E}[e^{tX}]$
 - Property:

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$$\mathbb{E}[e^{tX}] = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx$$

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$$\mathbb{E}[e^{tX}] = \sum_{k=0}^{\infty} E[X^k] \frac{t^k}{k!}$$

$$e^{tx} = \sum_{k=0}^{\infty} \frac{(tx)^k}{k!}$$

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$$E[e^{tX}] = E[\sum_{k=0}^{\infty} \frac{(tX)^k}{k!}] = \sum_{k=0}^{\infty} E[X^k] \frac{t^k}{k!}$$

$$\mathbb{E}[e^{tX}] = \mathbb{E}[X^k] \frac{t^k}{k!}$$

$$* \frac{d\mathbb{E}[e^{tX}]}{dt} = \mathbb{E}[X]$$

$$* \mathbb{E}[e^{t(aX+b)}] = e^t b \mathbb{E}[e^{taX}]$$

- * Not all random variables have Moment generating function
- Characteristic Function: $\mathbb{E}[e^{itX}]$
 - Property:
 - * All random variables have Moment generating function
- Joint Moment Generating Function: $G(x,y) = \mathbb{E}[e^{xX}e^{yY}]$
- Property:
 - (Joint) moment generating function uniquely determines the (joint) CDF
- Example
 - Trapped miner's random walk
 - * Miner has probability of $\frac{1}{3}$ to waste 3 hours in vain, $\frac{1}{3}$ to waste 5 hours in vain, and $\frac{1}{3}$ to spend 2 hours to go out of the mine.
 - * X is the random variables of the hours to go out of the mine
 - * Y_i is the random variables of the hours for the *i*-th action.

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$$\mathbb{E}[e^{tX}] = \mathbb{E}[e^{tX}|Y_1 = 2] + \mathbb{E}[e^{tX}|Y_1 = 3] + \mathbb{E}[e^{tX}|Y_1 = 5]$$

= $\mathbb{E}[e^{2t}] + \mathbb{E}[e^{t(X+3)}] + \mathbb{E}[e^{t(X+5)}]$

* Find expectation and variance by joint moment generating function

Expectation 2

- \bullet N i.i.d. events, when N is a random variable
 - Suppose N is a integer random variable
 - Suppose $X_1, \ldots, X_i, \ldots, X_N$ are i.i.d random variables with mean μ and variance σ^2

$$-Y = \sum_{i=1}^{N} X_i$$

$$- \mathbb{E}[Y] = \mathbb{E}[N]\mu$$

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$$\mathbb{E}[Y] = \sum_{n=1}^{\infty} \mathbb{E}[\sum_{i=1}^{N} X_i | N = n] P[N = n]$$

$$= \mu \times \sum_{n=1}^{\infty} n P[N = n] = \mathbb{E}[N] \mu$$

$$\begin{split} - & \ \mathbb{E}[Y^2] = \mathbb{E}[N]\mathbb{E}[X^2] + \mathbb{E}[N^2]\mu^2 - \mathbb{E}[N]\mu^2 \\ & * \ \mathbb{E}[Y^2] = \sum_{n=1}^{\infty} \mathbb{E}[(\sum_{i=1}^{N} X_i)^2 | N = n] P[N = n] = \sum_{n=1}^{\infty} (n\mathbb{E}[X_i^2] + n(n-1)\mu^2) P[N = n] \\ & = \mathbb{E}[N]\mathbb{E}[X^2] + \mathbb{E}[N^2]\mu^2 - \mathbb{E}[N]\mu^2 \\ & - \ Var(Y) = \mathbb{E}[N]\sigma^2 + Var(N)\mu^2 \end{split}$$

- Expectation by P[X > x]
 - $-\mathbb{E}[X] = \sum_{x} P[X > x]$, when X is a non-negative discrete random variable

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$$\mathbb{E}[X] = \sum_{x} P[X > x]$$
, when X is a non-negative discrete random variable
* $\mathbb{E}[X] = \sum_{x=0}^{\infty} x P[X = x] = \sum_{x=0}^{\infty} \sum_{y=0}^{x-1} P[X = x] = \sum_{y=0}^{\infty} \sum_{x=y+1}^{\infty} P[X = x] = \sum_{y=0}^{\infty} P[X > y]$

 $-\mathbb{E}[X] = \int_0^\infty P[X > x] dx$, when X is a non-negative continuous random variable

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$$\mathbb{E}[X] = \int_0^\infty x f_X(x) dx = \int_0^\infty \int_0^x f_X(x) dy dx = \int_0^\infty \int_y^\infty f_X(x) dx dy = \int_0^\infty P[X > y] dy$$

3 Inequality

• Markov Inequality

Definition:

- Suppose
$$X \geq 0$$
, then $P[X \geq \epsilon] \leq \frac{\mathbb{E}[X]}{\epsilon}$

Proof:

1.
$$\mathbb{E}[X] = \int_0^\infty x f_X(x) \ge \int_{\epsilon}^\infty x f_X(x) \ge \epsilon \int_{\epsilon}^\infty f_X(x) = \epsilon P[X \ge \epsilon]$$

2.
$$X(\omega) \ge \epsilon \mathbb{1}_{X(\omega) > \epsilon}, \forall \omega \in S$$

- Calculate expectation on both side.

$$- \mathbb{E}[X] \ge \epsilon P[X \ge \epsilon]$$

Property:

- The equality happens when $P[X = k] = 0, \forall k \notin \{0, \epsilon\}.$
- Chebyshev Inequality

Definition:

- Suppose
$$m = \mathbb{E}[X], \sigma^2 = Var(X)$$
, then $P[|X - m| \ge \epsilon] \le \frac{\sigma^2}{\epsilon^2}$

Proof:

$$-P[|X-m| > \epsilon] = P[(X-m)^2 > \epsilon^2]$$

–
$$P[(X-m)^2 \ge \epsilon^2] \le \frac{\mathbb{E}[(X-m)^2]}{\epsilon^2}$$
 (by Markov Inequality)

Property:

- The equality happens when $P[X = k] = 0, \forall k \notin \{m \epsilon, m, m + \epsilon\}.$
- Might be tighter than Markov Inequality since it requires m, σ
- Chernoff Inequality

Definition:

– Suppose X_1, \ldots, X_n are independent identically distributed Bernoulli random variable with probability p and $X = \sum_{i=1}^n X_i$

$$-P[X \ge \epsilon] \le \frac{(pe^t + 1 - p)^n}{e^{t\epsilon}} \le \frac{e^{np(e^t - 1)}}{e^{t\epsilon}}$$

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$$P[X \ge \epsilon] = P[e^{tX} \ge e^{t\epsilon}] \le \frac{E[e^{tX}]}{e^{t\epsilon}} = \frac{(E[e^{tX_i}])^n}{e^{t\epsilon}} = \frac{(pe^t + 1 - p)^n}{e^{t\epsilon}} \le \frac{e^{np(e^t - 1)}}{e^{t\epsilon}}$$

$$-\ P[X \geq np(1+\epsilon)] \leq (\frac{e^{\epsilon}}{(1+\epsilon)^{1+\epsilon}})^{np} \leq \left\{ \begin{array}{ll} e^{\frac{-\epsilon^2 np}{3}} & \text{if } 0 \leq \epsilon \leq 1 \\ e^{\frac{-\epsilon^2 np}{(2+\epsilon)}} & \text{if } \epsilon \geq 1 \end{array} \right.$$

- * Substitude ϵ with $np(1+\epsilon)$
- * Substitude t with $\log(1+\epsilon)$
- * the last inequality is without proof

$$- \ P[X \le \epsilon] \le \frac{(pe^t + 1 - p)^n}{e^{t\epsilon}} \le \frac{e^{np(e^t - 1)}}{e^{t\epsilon}}$$

$$*\ P[X \le \epsilon] = P[e^{-tX} \ge e^{-t\epsilon}] \le \frac{E[e^{-tX}]}{e^{-t\epsilon}} = \frac{(E[e^{-tX_i}])^n}{e^{-t\epsilon}} = \frac{(pe^{-t} + 1 - p)^n}{e^{-t\epsilon}} \le \frac{e^{np(e^{-t} - 1)}}{e^{-t\epsilon}}$$

$$- P[X \le np(1-\epsilon)] \le (\frac{e^{-\epsilon}}{(1-\epsilon)^{1-\epsilon}})^{np} \le e^{\frac{-\epsilon^2 np}{2}}$$

- * Substitude ϵ with $np(1-\epsilon)$
- * Substitude t with $-\log(1-\epsilon)$
- * the last inequality is without proof
- Chernoff/ Hoeffding Lemma

Definition:

- Suppose X_1, \ldots, X_n are independent distributed random variable and $a_i \leq X_i \leq b_i$
- Suppose $X = \sum_{i=1}^{n} X_i$ and $\mu = \mathbb{E}[X]$

$$-P[|X-\mu| \geq \epsilon] \leq 2e^{\frac{-2\epsilon^2}{\sum_{i=1}^n (b_i-a_i)^2}}$$
 without proof

- Application:
 - Balls in Bins

Definition: Throw n balls into n bins, find bounds for the maximum number of balls in all bins

- * P[maximum number of balls in all bins $\geq \epsilon]$
 - $=P[\cup_{i=1}^n \text{ number of balls in } i\text{-th bin } \geq \epsilon]$
 - $\leq n \times P[$ number of balls in one bin $\geq \epsilon]$
- * By Markov inequality:
 - · P[number of balls in one bin $\geq \epsilon$] $\leq \frac{1}{\epsilon} \rightarrow$ useless
- * By Chebyshev inequality:
 - · P[number of balls in one bin $\geq \epsilon] \leq \frac{(1-\frac{1}{n})}{\epsilon^2}$
 - · $P[\text{ maximum number of balls in all bins } \ge n^{\frac{1}{2}+\epsilon}] \le \frac{(1-\frac{1}{n})}{n^{2\epsilon}}$
 - · when $n \to \infty$, the maximum number of balls should less than $n^{\frac{1}{2}+\epsilon}$
- * By Chernoff inequality:
 - + P[number of balls in one bin $\geq 2\log n] \leq \frac{e^{np(e^t-1)}}{n^{2t}}$
 - · $P[\text{ maximum number of balls in all bins } \ge 2\log n] \le \frac{e^{np(e^t-1)}}{n^{2t-1}}$
 - · when t is a constant ≥ 0.5 and $n \to \infty$, the maximum number of balls should less than $2 \log n$

4 Law of Large Numbers

- $\{X_i\}_{i=1}^{\infty}$ is a sequence of pairwise uncorrelated random variable with $\mathbb{E}[X_i] = m$, $Var(X_i) = \sigma_i^2$.
- $\bullet \ M_n = \frac{1}{n} \sum_{i=1}^n X_i$
- $M_n \to m$ almost surely, in mean square and in probability.

5 Memoryless

- Definition: $P[X > x_1 + x_2 | X > x_1] = P[X > x_2]$
- Property:
 - Exponential random variable is the only continuous memoryless random variable
 - Bernoulli random variable is the only discrete memoryless random variable