# A new efficient simulation method for multiple access channel

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August 14, 2023

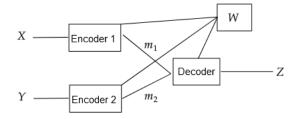
#### 1 Introduction

#### 1.1 Background and Literature Review

In recent years, the multiple access channel is getting more and more important among people. A lot of studies have been conducted on this topic. For example, [4, 6] discribed the rate region of the Quadratic Gaussian CEO problem, which is based on a setting of multiple access channel. However, while how to deal with the point-to-point channel is well known to people, such as [5] and [1], the channel simulation method for multiple access channel is not well researched so far. For this paper, we proposed a new efficient simulation method for multiple access channel based on our understanding of the point-to-point channel. This paper servers as an introduction to the multiple access channel simulation to increase the efficiency of doing this.

#### 1.2 Setting and Goal

In this section, we will focus on the setting of this problem. In this problem, X and Y are correlated sources, W is the common randomness. Encoder 1 observes W and encodes X into descriptions  $m_1$ . Encoder 2 observes W and encodes Y into descriptions  $m_2$ . Then  $m_1$  and  $m_2$  are transmitted to the decoder. This is a one-shot setting.



We set X and Y to be Jointly Gaussian. That means:

$$\begin{pmatrix} X \\ Y \end{pmatrix} \sim N \left( 0, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} \right).$$

Here  $\rho$  is the correlation coefficient taking value between -1 and 1.

Encoder 1 generates  $m_1$  by a function  $f_1$ , i.e.,  $m_1 = f_1(X, W)$ . Encoder 2 generates  $m_2$  by a function  $f_2$ , i.e.,  $m_2 = f_2(X, W)$ .

The decoder recieves  $m_1$  and  $m_2$  and generates the output  $\hat{Z}$ . The output  $\hat{Z}$  is determined by  $m_1$ ,  $m_2$  and W, i.e.,  $\hat{Z} = g(m_1, m_2, W)$ . We want the output  $\hat{Z}$  follow the same distribution as Z = aX + bY + N, where N is the Gaussian noise,  $N \sim N(0, \sigma_N)$ .

The requirement of the desired output is to satisfy a small error in the total variation distance, i.e., we want the output satisfy  $\delta_{TV}((X,Y,\hat{Z}),(X,Y,Z)) \leq \epsilon$ .

Our goal is to find the trade off between  $R_1 = H(m_1|W)$ , which is the average length of  $m_1$  encoded using Huffman conditional on W, and  $R_2 = H(m_2|W)$ , which is the average length of  $m_2$  encoded using Huffman conditional on W where  $\{(R1, R2) : \exists f_1, f_2, g \text{ s.t. the requirement is satisfied}\}$ .

This problem is already tackled and solved properly in point-to-point channel, such as [1] [5]. With multiple access channel getting more and more common in practice, however, this problem has not been well studied. This study aims to start the discussion of this problem and solve for some basic results with the experience we have in point-to-point channel.

# 2 Scheme Design

#### 2.1 Basic Idea

In section 1.1, we discussed the setting and the goal for this problem. In this section, we are going to start to work on a scheme for this problem. From 1.1, we knows that Z is a linear combination of X and Y with some Gaussian noise, we can suppose that  $\hat{X}$  and  $\hat{Y}$  are noisy version of X and Y. Hence, we have:

1. Do a channel simulation to allow Encoder 1 and Decoder to simulate an additive noise channel:  $\hat{X} - X \sim N(0, \sigma_{N_X}^2)$ ,  $\hat{X} - X \perp \!\!\! \perp X$ .

- 2. Do a channel simulation to allow Encoder 2 and Decoder to simulate an additive noise channel:  $\hat{B} - B \sim N(0, \sigma_{N_B}^2), \hat{B} - B \perp \!\!\! \perp B$ .
- 3. Solve  $\sigma_{N_X}, \sigma_{N_B}, \tilde{a}$  and  $\tilde{b}$  such that  $\hat{Z} = \tilde{a}\hat{X} + \tilde{b}\hat{Y}$  follows the same distribution as Z.

where

$$N_X \sim N(0, \sigma_{N_X}^2),$$
  
 $N_B \sim N(0, \sigma_{N_B}^2).$ 

We know that the decoder obtains  $\hat{X}$  and  $\hat{B}$ , and X and Y are correlated. To find an estimation of Y, what we need to do is:

$$\hat{Y} = \underset{Y:Y \text{ mod } \alpha = \hat{B}}{\arg \min} |Y - E[Y|\hat{X}]|. \tag{1}$$

Finally, we will show that the following theorem holds:

**Theorem:** The requirement and the goal is satisfied if  $R_1 \leq \frac{1}{2} \log(1 + \frac{1}{\sigma_N^2})$ ,

 $R_2 \leq \frac{1}{2}\log(1+\frac{\alpha^2}{4\sigma_{N_B}^2})$  and  $\epsilon \leq 2Q(\frac{\alpha}{2\sigma_{\theta}})$ , where  $\sigma_{\theta}^2 = 1-\frac{\rho^2}{1+\sigma_{N_X}^2}+\sigma_{N_B}^2$ . This holds when  $\tilde{a}^2\sigma_{N_X}^2 + \tilde{b}^2\sigma_{N_B}^2 = \sigma_N^2$ ,  $\hat{Z}$  follows the same distribution as Z. We will prove this in later parts.

#### Computation of $E[Y|\hat{X}]$ 2.2

The following sections are talking about the method to achieve this and evalu-

We need to calculate  $E[Y|\hat{X}]$  because we know that Y and X are correlated in Chapter 1. To calculate this conditional expectation easily, we can use a technique called linear MMSE estimator. Suppose:

$$Y = \gamma \hat{X} + T, T \perp \!\!\!\perp \hat{X}. \tag{2}$$

Then,

$$E[Y|\hat{X}] = \gamma \hat{X}. \tag{3}$$

By the property of covariance and variance of linear MMSE, we have:

$$Cov(\hat{X}, Y) = \gamma \sigma_{\hat{X}}^2. \tag{4}$$

$$\gamma = \frac{\operatorname{Cov}(\hat{X}, Y)}{\sigma_{\hat{X}}^2} = \frac{\rho}{1 + \sigma_{N_X}^2}.$$
 (5)

$$Var(Y - \gamma \hat{X}) = \sigma_Y^2 - \frac{Cov(\hat{X}, Y)^2}{\sigma_{\hat{X}}^2} = 1 - \frac{\rho^2}{1 + \sigma_{N_X}^2}.$$
 (6)

 $\operatorname{Var}(Y - \gamma \hat{X})$  is also called the error covariance. By the property of linear MMSE, we can put everything into the expression for  $\hat{Y}$ :

$$\hat{Y} = \frac{\text{Cov}(X, Y)}{\sigma_X^2} X. \tag{7}$$

Now, it is time for the decoder to "guess" a possible value of Y. Let the estimated value for Y of the decoder be  $\hat{Y}$ . We have calculated  $\mathrm{E}[Y|\hat{X}]$  in the previous part. By the property of Gaussian distribution, we know that the distribution of  $Y|\hat{X}$  is:

$$Y|\hat{X} \sim N(\gamma \hat{X}, 1 - \frac{\rho^2}{1 + \sigma_{N_Y}^2}).$$
 (8)

Since  $Y = K\alpha + B$  and the decoder knows  $\hat{B} = B + N_B$ , what the decoder guess is  $\tilde{K}$ . As we already know the distribution of  $Y|\hat{X}$  and it is Gaussian distribution, we can have the following conclusion:

$$\hat{Y} = \underset{Y:Y \text{ mod } \alpha = \hat{B}}{\arg \min} |Y - \gamma \hat{X}|. \tag{9}$$

$$\hat{Y} = \tilde{K}\alpha + \hat{B}.\tag{10}$$

Since  $\hat{B} = B + N_B$ ,  $\widetilde{K} = K$  if and only if

$$|Y - \gamma \hat{X} + N_B| \le \frac{\alpha}{2}.\tag{11}$$

By the property of Gaussian distribution and variance, we know that:

$$Var(Y - \gamma \hat{X} + N_B) = 1 - \frac{\rho^2}{1 + \sigma_{N_E}^2} + \sigma_{N_B}^2.$$
 (12)

$$Y - \gamma \hat{X} + N_B \sim N(0, 1 - \frac{\rho^2}{1 + \sigma_{N_X}^2} + \sigma_{N_B}^2).$$
 (13)

# 3 Analysis and Evaluation

### 3.1 Analysis on Link Capacity

After the basic setting of this problem, we can analyse the link capacity for this multiple access channel. We can get an upper bound by taking the mutual information. For example, we know that the link at the bottom transmits B, and the decoder receives  $B + N_B$ . Define  $\hat{B} = B + N_B$ . Hence, we can take the mutual information to get an upper bound of the bottom link capacity [2]:

$$I(B; \hat{B}) = h(\hat{B}) - h(\hat{B}|B) \tag{14}$$

$$=h(\hat{B}) - h(N_B) \tag{15}$$

$$=h(B+N_B)-h(N_B) \tag{16}$$

$$\leq \frac{1}{2}\log(2\pi e(\frac{\alpha^2}{4} + \sigma_{N_B}^2)) - \frac{1}{2}\log(2\pi e \sigma_{N_B}^2)$$
 (17)

$$\leq \frac{1}{2}\log(1+\frac{\alpha^2}{4\sigma_{N_B}^2}).$$
(18)

With the same method, we can get the top link capacity:

$$I(X; \widetilde{X}) \le \frac{1}{2} \log(1 + \frac{1}{\sigma_{N_Y}^2}).$$
 (19)

#### 3.2 Total Variation Distance Estimation

It is fine to estimate the distribution of  $\hat{Y} - Y$ . However, it turns out that the result will be a convolution of a Gaussian random variable with a bunch of delta functions, which is hard to solve and estimate precisely. Instead of finding the distribution of  $\hat{Y} - Y$ , we can focus on the total variation distance.

In section 2.2, the decoder estimated  $\hat{Y} = K\alpha + \hat{B}$ . According to this result, the actual output will be  $\hat{Z} = \tilde{a}\hat{X} + \tilde{b}\hat{Y}$ . Suppose in section 2.2, the decoder estimated K correctly, i.e.,  $\tilde{K} = K$ . Then  $\hat{Y}' = K\alpha + \hat{B}$ . We want to create  $\hat{Z}'$  that has the correct distribution, and make  $\hat{Z}' = \hat{Z}$  with high probability. Suppose  $\hat{Z}' = \tilde{a}\hat{X} + \tilde{b}\hat{Y}'$ .

We can estimate the total variation distance between  $(X, Y, \hat{Z})$  and (X, Y, Z):

$$\delta_{TV}((X, Y, \hat{Z}), (X, Y, Z)) \le \delta_{TV}((X, Y, \hat{Z}'), (X, Y, Z)) + \delta_{TV}((X, Y, \hat{Z}'), (X, Y, \hat{Z}))$$

(20)

$$\langle P(K \neq \widetilde{K}) \rangle$$
 (21)

$$=2Q(\frac{\alpha}{2\sigma_{\theta}}). \tag{22}$$

where  $\sigma_{\theta}^2 = 1 - \frac{\rho^2}{1 + \sigma_{N_X}^2} + \sigma_{N_B}^2$ . This holds when  $\tilde{a}^2 \sigma_{N_X}^2 + \tilde{b}^2 \sigma_{N_B}^2 = \sigma_N^2$ .

#### 3.3 One-Shot Channel Simulation

This part aims to find the minimum amount of communication needed to simulate this setup. Suppose  $m_1$  and  $m_2$  are prefix-free descriptions of X and Y. The problem is to find the minimum expected description length of  $m_1$  and  $m_2$ , that is to say,  $E[L(m_1)]$  and  $E[L(m_2)]$ . Several setting of this problem have been studied. One of them is [3]. The study shows that for one-shot channel

simulation with unlimited common randomness setup in which Alice and Bob share unlimited common randomness W, Alice observes  $X \sim P_X$  and sends a prefix-free description M to Bob via a noiseless channel such that Bob can generate Y (from M and W) according to a prescribed conditional distribution  $P_{Y|X}$ , the expected length is

$$E[L(M)] \le I(X;Y) + \log(I(X;Y) + 1) + 5. \tag{23}$$

To apply to our setting, let  $m_1$  and  $m_2$  be prefix-free descriptions (Huffman code for example), the expected length of  $m_1$  is

$$E[L(m_1)] \le I(X; \widetilde{X}) + \log(I(X; \widetilde{X}) + 1) + 5 \tag{24}$$

$$\leq \frac{1}{2}\log(1+\frac{1}{\sigma_{N_X}^2}) + \log(\frac{1}{2}\log(1+\frac{1}{\sigma_{N_X}^2}) + 1) + 5. \tag{25}$$

With the same method, we can get the expected length of m2

$$E[L(m_2)] \le I(B; \hat{B}) + \log(I(B; \hat{B}) + 1) + 5 \tag{26}$$

$$\leq \frac{1}{2}\log(1 + \frac{\alpha^2}{4\sigma_{N_B}^2}) + \log(\frac{1}{2}\log(1 + \frac{\alpha^2}{4\sigma_{N_B}^2}) + 1) + 5. \tag{27}$$

#### 4 Conclusion

With all of the results above, we can prove that the theorem stated in section 2.1 is true. The requirement and the goal is satisfied if  $R_1 \leq \frac{1}{2} \log(1 + \frac{1}{\sigma_{N_X}^2})$ ,  $R_2 \leq \frac{1}{2} \log(1 + \frac{\alpha^2}{4\sigma_{N_B}^2})$  and  $\epsilon \leq 2Q(\frac{\alpha}{2\sigma_{\theta}})$ , where  $\sigma_{\theta}^2 = 1 - \frac{\rho^2}{1 + \sigma_{N_X}^2} + \sigma_{N_B}^2$ . This holds when  $\tilde{a}^2 \sigma_{N_X}^2 + \tilde{b}^2 \sigma_{N_B}^2 = \sigma_N^2$ ,  $\hat{Z}$  follows the same distribution as Z.

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Academic Year 2022-2023

Term 2

Course Code ENGG-0001-A

Course Title FDC

Assignment Marker TJONG, Joyce

Assignment Number 2

Due Date (provided by student) 2023-08-14

Submitted File Name CUI Yicheng\_Final Report.pdf

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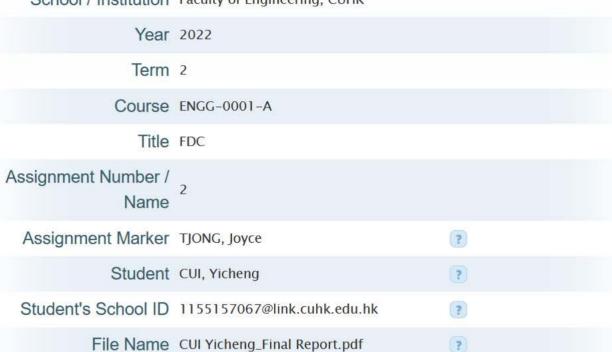
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