Introduction to Gaussian Mixture Models and Expectation Maximization Algorithm (DRAFT)

Jun Cheng

Dept. of Intelligent Information Eng. and Sci., Doshisha University Kyoto 610-0321 Japan jcheng@ieee.org

Abstract—This manuscript introduces the basic concept of Gaussian Mixture models and the expectation maximization algorithm for clustering.

I. GAUSSIAN MIXTURE MODELS (1-DIMENSION)

A. Data Points from K Independent Sources

Let $X^{(k)}$, k = 1, 2, ..., K, be the (latent) random variable (RV) associated with the k-th source (or cluster) \mathcal{C}_k . The RV $X^{(k)}$ is Gaussian distributed with the probability density function (PDF)

$$p_X(x|\mathcal{C}_k) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{(x-\mu_k)^2}{2\sigma_k^2}} \stackrel{\triangle}{=} \mathcal{N}(x;\mu_k,\sigma_k^2).$$

The probability of the source being active is $\pi_k \stackrel{\triangle}{=} p(\mathcal{C}_k)$, where $\sum_{k=1}^K p(\mathcal{C}_k) = 1$. Let $x_1^{(k)}, x_2^{(k)}, \dots, x_{n_k}^{(k)}$ be n_k samples from the k-th source \mathcal{C}_k . For K sources, we observe $N = \sum_{k=1}^K n_k$ data points in set

$$\mathcal{X} = \{x_1^{(k)}, x_2^{(k)}, \dots, x_{n_k}^{(k)} \mid k = 1, 2, \dots, K\}.$$

B. Mixture Models

Often times we don't observe the sources \mathcal{C}_k $(k=1,2,\ldots,K)$. This is why the $X^{(k)}$ is sometimes called <u>latent</u> RV. We observe only the data points in set \mathcal{X} which is rewritten as

$$\mathcal{X} = \{x_1, x_2, \dots, x_n, \dots, x_N\}.$$

From the law of total probability, we know that the marginal probability of x_n is

$$p(x_n) = \sum_{k=1}^{K} p(x_n, C_k) = \sum_{k=1}^{K} p(C_k) p(x_n | C_k) = \sum_{k=1}^{K} \pi_k p(x_n | C_k)$$

which is the mixture model of the observation. Here $\pi_k = p(\mathcal{C}_k)$ are called <u>mixture proportions</u> or <u>mixture weights</u>. We call $p(x_n|\mathcal{C}_k)$ the mixture component, and it represents the distribution of x_n assuming it came from component \mathcal{C}_k . The mixture components in this note are Gaussian distributions.

If we observe N independent samples x_1, x_2, \ldots, x_N from the mixture, the likelihood function is

$$p(x_1, x_2, ..., x_N) = \prod_{n=1}^{N} p(x_n) = \prod_{n=1}^{N} \sum_{k=1}^{K} \pi_k p(x_n | \mathcal{C}_k).$$

Taking the logarithm yields the following log-likelihood function (LLF)

$$\log p(x_1, x_2, \dots, x_N) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k p(x_n | \mathcal{C}_k).$$
 (1)

Now assume we are in the Gaussian mixture model setting where the k-th component $p(x_n|\mathcal{C}_k)$ is Gaussian $\mathcal{N}(x_n; \mu_k, \sigma_k^2)$ and the mixture proportions are π_k . A natural next question to ask is how to estimate the parameters $\{\mu_k, \sigma_k^2, \pi_k\}$ from our observations x_1, x_2, \ldots, x_N . In other words, how to label (or classify) each observation to its cluster \mathcal{C}_k , $k = 1, 2, \ldots, K$.

II. EM ALGORITHM (1-DIMENSION)

Given N observations x_1, x_2, \ldots, x_N , we will estimate the parameters $\{\mu_k, \sigma_k^2, \pi_k\}$, $k = 1, 2, \ldots, K$. Furthermore, we label (or classify) each observation to its cluster \mathcal{C}_k , $k = 1, 2, \ldots, K$.

- 1. Initialization: Given K, randomly set $\{\mu_k^0, (\sigma_k^0)^2, \pi_k^0\}$, $k = 1, 2, \dots, K$
 - 2. Expectation-Step (E-Step): at the ℓ -th iteration,

$$\underbrace{p(\mathcal{C}_{k}^{\ell}|x_{n})}_{c_{k,n}^{\ell}} = \frac{p(\mathcal{C}_{k}^{\ell-1}, x_{n})}{\sum_{k'=1}^{K} p(\mathcal{C}_{k'}^{\ell-1}, x_{n})} \\
= \frac{p(\mathcal{C}_{k}^{\ell-1})p(x_{n}|\mathcal{C}_{k}^{\ell-1})}{\sum_{k'=1}^{K} p(\mathcal{C}_{k'}^{\ell-1})p(x_{n}|\mathcal{C}_{k'}^{\ell-1})} \\
= \frac{\pi_{k}^{\ell-1}p(x_{n}|\mathcal{C}_{k}^{\ell-1})}{\sum_{k'=1}^{K} \pi_{k'}^{\ell-1}p(x_{n}|\mathcal{C}_{k'}^{\ell-1})} \\
= \frac{\pi_{k}^{\ell-1}\mathcal{N}(x_{n}; \mu_{k}^{\ell-1}, \sigma_{k'}^{\ell-1})^{2})}{\sum_{k'=1}^{K} \pi_{k'}^{\ell-1}\mathcal{N}(x_{n}; \mu_{k'}^{\ell-1}, \sigma_{k'}^{\ell-1})^{2})}. (2)$$

3. Maximization-Step (M-Step):

$$\pi_k^{\ell} \stackrel{\triangle}{=} p(\mathcal{C}_k^{\ell}) = \sum_{n=1}^N p(\mathcal{C}_k^{\ell}, x_n)$$

$$= \sum_{n=1}^N p(x_n) p(\mathcal{C}_k^{\ell} | x_n) = \frac{1}{N} \sum_{n=1}^N \underbrace{p(\mathcal{C}_k^{\ell} | x_n)}_{\mathcal{C}_{k,n}^{\ell}}. \tag{3}$$

Since

$$p(x_{n}|\mathcal{C}_{k}^{\ell}) = \frac{p(\mathcal{C}_{k}^{\ell}, x_{n})}{\sum_{n'=1}^{N} p(\mathcal{C}_{k}^{\ell}, x_{n'})} = \frac{p(x_{n})p(\mathcal{C}_{k}^{\ell}|x_{n})}{\sum_{n'=1}^{N} p(x_{n'})p(\mathcal{C}_{k}^{\ell}|x_{n'})}$$

$$= \frac{(1/N)p(\mathcal{C}_{k}^{\ell}|x_{n})}{\sum_{n'=1}^{N} (1/N)p(\mathcal{C}_{k}^{\ell}|x_{n'})} = \frac{p(\mathcal{C}_{k}^{\ell}|x_{n})}{\sum_{n'=1}^{N} p(\mathcal{C}_{k}^{\ell}|x_{n'})}$$

$$= \frac{p(\mathcal{C}_{k}^{\ell}|x_{n})}{N\pi^{\ell}}$$

we have

$$\mu_k^{\ell} = \sum_{n=1}^{N} p(x_n | \mathcal{C}_k^{\ell}) x_n = \sum_{n=1}^{N} \underbrace{\frac{p(\mathcal{C}_k^{\ell} | x_n)}{p(\mathcal{C}_k^{\ell} | x_n)}}_{N\pi_k^{\ell}} x_n.$$
(4)

$$(\sigma_k^2)^{\ell} = \sum_{n=1}^{N} p(x_n | \mathcal{C}_k^{\ell}) (x_n - \mu_k^{\ell})^2$$

$$= \sum_{n=1}^{N} \frac{\overbrace{p(\mathcal{C}_k^{\ell} | x_n)}^{c_{k,n}^{\ell}}}{N\pi_k^{\ell}} (x_n - \mu_k^{\ell})^2.$$
 (5)

4. Evaluation-Step (Eva-Step)

We evaluate (see (1))

$$\Gamma^{\ell} \stackrel{\triangle}{=} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k}^{\ell} p(x_{n} | \mathcal{C}_{k}^{\ell})$$

$$= \sum_{n=1}^{N} \log \left(\frac{1}{N} \sum_{k=1}^{K} \underbrace{p(\mathcal{C}_{k}^{\ell} | x_{n})}_{\mathcal{C}_{k,n}^{\ell}} \right)$$
(6)

at each iteration of E-Step and M-Step, and check the convergence of the algorithm.

5. Result of the estimation

The parameters

$$\{\mu_k^{\ell \to \infty}, (\sigma_k^2)^{\ell \to \infty}, \pi_k^{\ell \to \infty}\}, k = 1, 2, \dots, K$$

are estimated such that the LLF (see (6)) $\Gamma^{\ell\to\infty}$ converse. The probability that x_n belongs to \mathcal{C}_k is

$$\Pr(x_n \in \mathcal{C}_k) \stackrel{\triangle}{=} p(\mathcal{C}_k^{\ell \to \infty} | x_n) = c_{k,n}^{\ell \to \infty}, \text{(soft clustering)}$$

 $n = 1, 2, \dots, N, k = 1, 2, \dots, K.$

The observation x_n is labeled (or clustered) to C_k ,

$$x_n \in \mathcal{C}_{k*}$$
, (hard clustering)
 $n = 1, 2, \dots, N$

where

$$k^* = \underset{k=1,2,\dots,K}{\operatorname{argmax}} p(\mathcal{C}_k^{\ell \to \infty} | x_n).$$

Finally, we summery the EM algorithm in Algorithm 1.

Algorithm 1: EM algorithm

Result: Labeling
$$x_1, \ldots, x_N$$
 to \mathcal{C}_k , $k = 1, \ldots, K$; input: x_1, \ldots, x_N ; initialization: K , $\{\mu_k^0, (\sigma_k^0)^2, \pi_k^0\}$, L ; while $\underline{\ell} \leq \underline{L}$ do $|$ compute $c_{k,n}^\ell$ in (2); compute π_k^ℓ , μ_k^ℓ , $(\sigma_k^2)^\ell$, in (3), (4), (5), respectively; If $(\Gamma^\ell - \Gamma^{\ell-1})^2 < \epsilon$ break end outputs: $\Pr(x_n \in \mathcal{C}_k) = c_{k,n}^\ell$; $x_n \in \mathcal{C}_{k^*}$ where $k^* = \underset{k=1,2,\ldots,K}{\operatorname{argmax}} p(\mathcal{C}_k^\ell | x_n)$;

III. EM ALGORITHM (d-DIMENSION)

We observe N independent samples

$$\boldsymbol{x}_1, \boldsymbol{x}_2, \dots, \boldsymbol{x}_n, \dots, \boldsymbol{x}_N$$

from a Gaussian mixture model

$$p(\boldsymbol{x}_n) = \sum_{k=1}^K \pi_k p(\boldsymbol{x}_n | \mathcal{C}_k)$$

where x_n are length-d real vector, π_k are mixture weights, and mixture components are d-variable Gaussian distribution with PDF

$$p(\boldsymbol{x}|\mathcal{C}_k)$$

$$= \frac{1}{(2\pi)^{d/2}|\Sigma_k|^{1/2}} \exp\left(-\frac{1}{2}(\boldsymbol{x} - \boldsymbol{\mu}_k)^{\mathsf{T}} \Sigma_k^{-1} (\boldsymbol{x} - \boldsymbol{\mu}_k)\right)$$

$$\triangleq \mathcal{N}(x; \boldsymbol{\mu}_k, \Sigma_k).$$

Here the exponent part is represented as

$$(\boldsymbol{x} - \boldsymbol{\mu}_k)^{\mathrm{T}} \Sigma_k^{-1} (\boldsymbol{x} - \boldsymbol{\mu}_k) = \sum_{i=1}^d \sum_{j=1}^d (x_{n,j} - \mu_{k,j}) (\Sigma_k^{-1})_{j,i} (x_{n,i} - \mu_{k,i}),$$

and

$$\boldsymbol{x}_n = (x_{n,1}, \dots, x_{n,d})^{\mathsf{T}}, \quad \boldsymbol{\mu}_k = (\mu_{k,1}, \dots, \mu_{k,d})^{\mathsf{T}}.$$

 Σ_k is $d \times d$ covariance matrix, Σ_k^{-1} is the inverse of Σ_k , $(\Sigma_k^{-1})_{j,i}$ is the (j,i)-th element of Σ_k^{-1} , and $|\Sigma_k|$ is the determinant of Σ_k .

The log-likelihood function is

$$\log p(\boldsymbol{x}_1, \boldsymbol{x}_2, \dots, \boldsymbol{x}_N) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k p(\boldsymbol{x}_n | \mathcal{C}_k).$$
 (7)

Similar to 1-dimension, the EM algorithm is as follows:

- 1. Initialization: Given K, randomly set $\{\mu_k^0, \Sigma_k^0, \pi_k^0\}$, $k = 1, 2, \dots, K$.
 - 2. E-Step:

$$\underbrace{p(\mathcal{C}_k^{\ell}|\boldsymbol{x}_n)}_{c_{k,n}^{\ell}} = \frac{\pi_k^{\ell-1}\mathcal{N}(\boldsymbol{x}_n;\boldsymbol{\mu}_k^{\ell-1},\boldsymbol{\Sigma}_k^{\ell-1})}{\sum_{k=1}^K \pi_k^{\ell-1}\mathcal{N}(\boldsymbol{x}_n;\boldsymbol{\mu}_k^{\ell-1},\boldsymbol{\Sigma}_k^{\ell-1})}.$$
(8)

3. M-Step:

$$\pi_k^{\ell} = \frac{1}{N} \sum_{n=1}^{N} \underbrace{p(\mathcal{C}_k^{\ell} | \boldsymbol{x}_n)}_{c_{k,n}^{\ell}}$$
 (9)

$$\boldsymbol{\mu}_{k}^{\ell} = \sum_{n=1}^{N} \frac{\overbrace{p(\mathcal{C}_{k}^{\ell} | \boldsymbol{x}_{n})}^{c_{k,n}^{\ell}}}{N\pi_{k}^{\ell}} \boldsymbol{x}_{n}. \tag{10}$$

$$(\Sigma_k^{\ell})_{j,i} = \sum_{n=1}^{N} \frac{\overbrace{p(\mathcal{C}_k^{\ell} | \mathbf{x}_n)}^{c_{k,n}^{\ell}}}{N\pi_k^{\ell}} (x_{n,j} - \mu_{k,j}^{\ell}) (x_{n,i} - \mu_{k,i}^{\ell}) (11)$$

4. Eva-Step

We evaluate (see (7))

$$\Gamma^{\ell} \stackrel{\triangle}{=} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k}^{\ell} p(\boldsymbol{x}_{n} | \mathcal{C}_{k}^{\ell})$$

$$= \sum_{n=1}^{N} \log \left(\frac{1}{N} \sum_{k=1}^{K} \underbrace{p(\mathcal{C}_{k}^{\ell} | \boldsymbol{x}_{n})}_{c_{k,n}^{\ell}} \right)$$
(12)

at each iteration of E-Step and M-Step, and check the convergence of the algorithm.

5. Result of the estimation

The probability that x_n belongs to C_k is

$$\Pr(\boldsymbol{x}_n \in \mathcal{C}_k) \stackrel{\triangle}{=} p(\mathcal{C}_k^{\ell \to \infty} | \boldsymbol{x}_n) = c_{k,n}^{\ell \to \infty}, \text{(soft clustering)}$$

$$n = 1, 2, \dots, N, k = 1, 2, \dots, K.$$

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$$k^* = \underset{k=1,2,...,K}{\operatorname{argmax}} p(\mathcal{C}_k^{\ell \to \infty} | \boldsymbol{x}_n).$$

REFERENCES