

Introduction to Gaussian Mixture Models and Expectation Maximization Algorithm (DRAFT)

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Abstract—This manuscript introduces

I. GAUSSIAN MIXTURE MODELS (1-DIMENSION)

A. Data Points from K Independent Sources

Let $X^{(k)}$, $k = 1, 2, \dots, K$, be the (latent) random variable (RV) associated with the k -th source (or cluster) \mathcal{C}_k . The RV $X^{(k)}$ is Gaussian distributed with the probability density function (PDF)

$$p_X(x|\mathcal{C}_k) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{(x-\mu_k)^2}{2\sigma_k^2}} \triangleq \mathcal{N}(x; \mu_k, \sigma_k^2).$$

The probability of the source being active is $\pi_k \triangleq p(\mathcal{C}_k)$, where $\sum_{k=1}^K p(\mathcal{C}_k) = 1$. Let $x_1^{(k)}, x_2^{(k)}, \dots, x_{n_k}^{(k)}$ be n_k samples from the k -th source \mathcal{C}_k . For K sources, we observe $N = \sum_{k=1}^K n_k$ data points in set

$$\mathcal{X} = \{x_1^{(k)}, x_2^{(k)}, \dots, x_{n_k}^{(k)} \mid k = 1, 2, \dots, K\}.$$

B. Mixture Models

Often times we don't observe the sources \mathcal{C}_k ($k = 1, 2, \dots, K$). This is why the $X^{(k)}$ is sometimes called latent RV. We observe only the data points in set \mathcal{X} which is rewritten as

$$\mathcal{X} = \{x, x_2, \dots, x_n, \dots, x_N\}.$$

From the law of total probability, we know that the marginal probability of x_n is

$$p(x_n) = \sum_{k=1}^K p(x_n, \mathcal{C}_k) = \sum_{k=1}^K p(\mathcal{C}_k) p(x_n|\mathcal{C}_k) = \sum_{k=1}^K \pi_k p(x_n|\mathcal{C}_k).$$

which is the mixture model of the observation. Here $\pi_k = p(\mathcal{C}_k)$ are called mixture proportions or mixture weights. We call $p(x_n|\mathcal{C}_k)$ the mixture component, and it represents the distribution of x_n assuming it came from component \mathcal{C}_k . The mixture components in this note are Gaussian distributions.

If we observe N independent samples x_1, x_2, \dots, x_N from the mixture, the likelihood function is

$$p(x_1, x_2, \dots, x_N) = \prod_{n=1}^N p(x_n) = \prod_{n=1}^N \sum_{k=1}^K \pi_k p(x_n|\mathcal{C}_k).$$

Taking the logarithm yields the following log-likelihood function

$$\log p(x_1, x_2, \dots, x_N) = \sum_{n=1}^N \log \sum_{k=1}^K \pi_k p(x_n|\mathcal{C}_k). \quad (1)$$

Now assume we are in the Gaussian mixture model setting where the k -th component $p(x_n|\mathcal{C}_k)$ is Gaussian $\mathcal{N}(x_n; \mu_k, \sigma_k^2)$ and the mixture proportions are π_k . A natural next question to ask is how to estimate the parameters $\{\mu_k, \sigma_k^2, \pi_k\}$ from our observations x_1, x_2, \dots, x_N . In other words, how to label (or classify) each observation to its cluster \mathcal{C}_k , $k = 1, 2, \dots, K$.

II. EM ALGORITHM (1-DIMENSION)

Given N observations x_1, x_2, \dots, x_N , we will estimate the parameters $\{\mu_k, \sigma_k^2, \pi_k\}$, $k = 1, 2, \dots, K$. Furthermore, we label (or classify) each observation to its cluster \mathcal{C}_k , $k = 1, 2, \dots, K$.

1. Initialization: Given K , randomly set $\{\mu_k^0, (\sigma_k^0)^2, \pi_k^0\}$, $k = 1, 2, \dots, K$

2. Expectation-Step (E-Step): at the ℓ -th iteration,

$$\begin{aligned} \underbrace{p(\mathcal{C}_k^\ell | x_n)}_{c_{k,n}^\ell} &= \frac{p(\mathcal{C}_k^{\ell-1}, x_n)}{\sum_{k=1}^K p(\mathcal{C}_k^{\ell-1}, x_n)} \\ &= \frac{p(\mathcal{C}_k^{\ell-1}) p(x_n | \mathcal{C}_k^{\ell-1})}{\sum_{k=1}^K p(\mathcal{C}_k^{\ell-1}) p(x_n | \mathcal{C}_k^{\ell-1})} \\ &= \frac{\pi_k^{\ell-1} p(x_n | \mathcal{C}_k^{\ell-1})}{\sum_{k=1}^K \pi_k^{\ell-1} p(x_n | \mathcal{C}_k^{\ell-1})} \\ &= \frac{\pi_k^{\ell-1} \mathcal{N}(x_n; \mu_k^{\ell-1}, \sigma_k^{\ell-1})}{\sum_{k=1}^K \pi_k^{\ell-1} \mathcal{N}(x_n; \mu_k^{\ell-1}, \sigma_k^{\ell-1})}. \end{aligned} \quad (2)$$

3. Maximization-Step (M-Step):

$$\begin{aligned} \pi_k^\ell &\triangleq p(\mathcal{C}_k^\ell) = \sum_{n=1}^N p(\mathcal{C}_k^\ell, x_n) \\ &= \sum_{n=1}^N p(x_n) p(\mathcal{C}_k^\ell | x_n) = \frac{1}{N} \sum_{n=1}^N \underbrace{p(\mathcal{C}_k^\ell | x_n)}_{c_{k,n}^\ell} \end{aligned} \quad (3)$$

Since

$$\begin{aligned}
p(x_n | \mathcal{C}_k^\ell) &= \frac{p(\mathcal{C}_k^\ell, x_n)}{\sum_{n=1}^N p(\mathcal{C}_k^\ell, x_n)} = \frac{p(x_n)p(\mathcal{C}_k^\ell | x_n)}{\sum_{n=1}^N p(x_n)p(\mathcal{C}_k^\ell | x_n)} \\
&= \frac{(1/N)p(\mathcal{C}_k^\ell | x_n)}{\sum_{n=1}^N (1/N)p(\mathcal{C}_k^\ell | x_n)} = \frac{p(\mathcal{C}_k^\ell | x_n)}{\sum_{n=1}^N p(\mathcal{C}_k^\ell | x_n)} \\
&= \frac{p(\mathcal{C}_k^\ell | x_n)}{N\pi_k^\ell}
\end{aligned}$$

we have

$$\mu_k^\ell = \sum_{n=1}^N p(x_n | \mathcal{C}_k^\ell) x_n = \sum_{n=1}^N \frac{\overbrace{p(\mathcal{C}_k^\ell | x_n)}^{c_{k,n}^\ell}}{N\pi_k^\ell} x_n. \quad (4)$$

$$\begin{aligned}
(\sigma_k^\ell)^2 &= \sum_{n=1}^N p(x_n | \mathcal{C}_k^\ell) (x_n - \mu_k^\ell)^2 \\
&= \sum_{n=1}^N \frac{\overbrace{p(\mathcal{C}_k^\ell | x_n)}^{c_{k,n}^\ell}}{N\pi_k^\ell} (x_n - \mu_k^\ell)^2. \quad (5)
\end{aligned}$$

4. Evaluation-Step (Eva-Step)

We valuate (1) at each iteration of E-Step and M-Step, and check the convergence of the algorithm.

5. Result of the estimation

The parameters

$$\{\mu_k^{\ell \rightarrow \infty}, (\sigma_k^2)^{\ell \rightarrow \infty}, \pi_k^{\ell \rightarrow \infty}\}, k = 1, 2, \dots, K$$

are estimated such that the LLF of (1) is maximal. The probability that x_n belongs to \mathcal{C}_k is

$$\Pr(x_n \in \mathcal{C}_k) \triangleq p(\mathcal{C}_k^{\ell \rightarrow \infty} | x_n) = c_{k,n}^{\ell \rightarrow \infty}, (\text{soft clustering}) \\
n = 1, 2, \dots, N, k = 1, 2, \dots, K.$$

The observation x_n is labeled (or clustered) to \mathcal{C}_k ,

$$x_n \in \mathcal{C}_{k^*}, (\text{hard clustering}) \\
n = 1, 2, \dots, N$$

where

$$k^* = \operatorname{argmax}_{k=1,2,\dots,K} p(\mathcal{C}_k^{\ell \rightarrow \infty} | x_n).$$

Finally, we summery the EM algorithm in Algorithm 1.

Algorithm 1: EM algorithm

Result: Labeling x_1, \dots, x_N to \mathcal{C}_k , $k = 1, \dots, K$;

input: x_1, \dots, x_N ;

initialization: $K, \{\mu_k^0, (\sigma_k^0)^2, \pi_k^0\}, L$;

while $\ell \leq L$ **do**

 compute $c_{k,n}^\ell$ in (2) ;

 compute $\pi_k^\ell, \mu_k^\ell, (\sigma_k^\ell)^2$, in (3), (4), (5), respectively;

If $(c_{k,n}^\ell - c_{k,n}^{\ell-1})^2 < \epsilon$ **break**

end

outputs: $\Pr(x_n \in \mathcal{C}_k) = c_{k,n}^\ell$;

$x_n \in \mathcal{C}_{k^*}$ where $k^* = \operatorname{argmax}_{k=1,2,\dots,K} p(\mathcal{C}_k^\ell | x_n)$;

III. EM ALGORITHM (d -DIMENSION)

We observe N independent samples

$$\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n, \dots, \mathbf{x}_N$$

from a Gaussian mixture model

$$p(\mathbf{x}_n) = \sum_{k=1}^K \pi_k p(\mathbf{x}_n | \mathcal{C}_k)$$

where \mathbf{x}_n are length- d real vector, π_k are mixture weights, and mixture components are d -variable Gaussian distribution with PDF

$$\begin{aligned}
p(\mathbf{x} | \mathcal{C}_k) &= \frac{1}{(2\pi)^{d/2} |\Sigma_k|^{1/2}} \exp \left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) \right) \\
&\triangleq \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_k, \Sigma_k).
\end{aligned}$$

Here the exponent part is represented as

$$(\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) = \sum_{j=1}^d \sum_{i=1}^d (x_{n,j} - \mu_{k,j}) (\Sigma_k^{-1})_{j,i} (x_{n,i} - \mu_{k,i}),$$

and

$$\mathbf{x}_n = (x_{n,1}, \dots, x_{n,d})^T, \quad \boldsymbol{\mu}_k = (\mu_{k,1}, \dots, \mu_{k,d})^T.$$

Σ_k is $d \times d$ covariance matrix, Σ_k^{-1} is the inverse of Σ_k , $(\Sigma_k^{-1})_{j,i}$ is the (j,i) -th element of Σ_k^{-1} , and $|\Sigma_k|$ is the determinant of Σ_k .

The log-likelihood function is

$$\log p(\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N) = \sum_{n=1}^N \log \sum_{k=1}^K \pi_k p(\mathbf{x}_n | \mathcal{C}_k). \quad (6)$$

Similar to 1-dimension, the EM algorithm is as follows:

1. Initialization: Given K , randomly set $\{\boldsymbol{\mu}_k^0, \Sigma_k^0, \pi_k^0\}$, $k = 1, 2, \dots, K$

2. E-Step:

$$\overbrace{p(\mathcal{C}_k^\ell | \mathbf{x}_n)}^{c_{k,n}^\ell} = \frac{\pi_k^{\ell-1} \mathcal{N}(\mathbf{x}_n; \boldsymbol{\mu}_k^{\ell-1}, \Sigma_k^{\ell-1})}{\sum_{k=1}^K \pi_k^{\ell-1} \mathcal{N}(\mathbf{x}_n; \boldsymbol{\mu}_k^{\ell-1}, \Sigma_k^{\ell-1})}. \quad (7)$$

3. M-Step:

$$\pi_k^\ell = \frac{1}{N} \sum_{n=1}^N \underbrace{p(\mathcal{C}_k^\ell | \mathbf{x}_n)}_{c_{k,n}^\ell} \quad (8)$$

$$\boldsymbol{\mu}_k^\ell = \sum_{n=1}^N \frac{\overbrace{p(\mathcal{C}_k^\ell | \mathbf{x}_n)}^{c_{k,n}^\ell}}{N \pi_k^\ell} \mathbf{x}_n. \quad (9)$$

$$((\Sigma_k^\ell)_{j,i}) = \sum_{n=1}^N \frac{\overbrace{p(\mathcal{C}_k^\ell | \mathbf{x}_n)}^{c_{k,n}^\ell}}{N \pi_k^\ell} (x_{n,j} - \mu_{k,j}^\ell)(x_{n,i} - \mu_{k,i}^\ell) \quad (10)$$

4. Eva-Step

We evaluate (6) at each iteration of E-Step and M-Step, and check the convergence of the algorithm.

5. Result of the estimation

The probability that \mathbf{x}_n belongs to \mathcal{C}_k is

$$\Pr(\mathbf{x}_n \in \mathcal{C}_k) \triangleq p(\mathcal{C}_k^{\ell \rightarrow \infty} | \mathbf{x}_n) = c_{k,n}^{\ell \rightarrow \infty}, \text{ (soft clustering)} \\ n = 1, 2, \dots, N, k = 1, 2, \dots, K.$$

The observation \mathbf{x}_n is labeled (or clustered) to \mathcal{C}_k ,

$$\mathbf{x}_n \in \mathcal{C}_{k*}, \text{ (hard clustering)} \\ n = 1, 2, \dots, N$$

where

$$k^* = \operatorname{argmax}_{k=1,2,\dots,K} p(\mathcal{C}_k^{\ell \rightarrow \infty} | \mathbf{x}_n).$$