

Homework Assignment 12: Applied Probabilistic Models

Generating Functions

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1 Exercises

Exercises solved in this work are provided in the book Grinstead and Snell [1].

1.1 Generating Functions for Discrete Densities

For these exercises generating functions for discrete densities are discussed. These exercises are related to branching processes

1.1.1 Exercise 1 page 392

The exercise describes a branching process. Let $h(z)$ be the ordinary generating function for the p_i :

$$h(z) = p_0 + p_1z + p_2z^2 + \cdots. \quad (1)$$

By Theorem 10.2 in Grinstead and Snell [1], if the mean number m of offspring produced by a single parent is ≤ 1 , then $d = 1$ and the process dies out with probability 1. If $m > 1$ then $d < 1$ and the process dies out with probability d .

(a) For this case $h'(z)|_{z=1} = m = \frac{1}{4} + \frac{1}{2}(1) = \frac{3}{4}$. Since $m < 1$, then $d = 1$.

(b) For this case $h'(z)|_{z=1} = m = \frac{1}{3} + \frac{2}{3}(1) = 1$. Since $m = 1$, then $d = 1$.

(c) For this case $h'(z)|_{z=1} = m = \frac{4}{3}$. Since $m > 1$, then the process dies out with probability d . To find this value, the exercise states that at most two offspring can be produced, therefore the condition $z = h(z)$ yields the equation

$$d = p_0 + p_1d + p_2d^2, \quad (2)$$

which is satisfied by $d = 1$ and $d = p_0/p_2$. Thus, in addition to the root $d = 1$, this second root $d = \frac{1}{2}$, and represents the probability that the process will die out.

(d) For this case $h'(z)|_{z=1} = m = \sum_{j=0}^{\infty} \left(\frac{n}{2^{n+1}} \right)$. It looks like $m \rightarrow 1$ as it is added up to ∞ offspring, then $d = 1$.

(e) For this case $h'(z)|_{z=1} = m = \sum_{j=0}^{\infty} \frac{j}{3} \left(\frac{2}{3}\right)^j$. This sumation is greater than 1 then, $m > 1$. To

calculate d if notice this geometric series has the form $\frac{1}{3-2z}$, then the condition $z = h(z)$ yields the equation

$$\begin{aligned}(3-2z)z &= 1, \\ 2z^2 - 3z + 1 &= 0.\end{aligned}$$

which is satisfied by $d = 1$ and $d = \frac{1}{2}$.

(f) To estimate d numerically it is used R software [2], with the following code:

```
1 pj <- function (j){
2   return (exp(-2)*2^j/factorial(j)) #Probability pj of j offspring
3 }
4
5 d <- pj(0)
6
7 for (i in 0:100){
8   d_new <- 0
9   for (j in 0:100){ #Sumation of d for j offspring
10    d_new <- d_new + pj(j)*(d^j)
11  }
12  d <- d_new
13 }
```

branching.R

This experiment estimates a value of $d \approx 0.2032$.

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1.1.2 Exercise 3 page 392

In the chain letter problem the expected number of letters we send is $m = p_1 + 2p_2$ and the expected payoff is equal to $-100 + 50(m + m^{12})$. The expected profit is asked to be calculated.

Case a:

If $p_0 = \frac{1}{2}$, $p_1 = 0$, and $p_2 = \frac{1}{2}$; then $m = 0 + 2\left(\frac{1}{2}\right) = 1$. Therefore:

$$\mathbb{E}(\text{Profit}) = -100 + 50(1 + 1^{12}) = 0.$$

Case b:

If $p_0 = \frac{1}{6}$, $p_1 = \frac{1}{2}$, and $p_2 = \frac{1}{3}$; then $m = \frac{1}{2} + 2\left(\frac{1}{3}\right) = \frac{7}{6}$. Therefore:

$$\mathbb{E}(\text{Profit}) = -100 + 50 \left[\frac{7}{6} + \left(\frac{7}{6}\right)^{12} \right] = 276.26.$$

If $p_0 > \frac{1}{2}$ we cannot expect to make a profit because let say $p_0 = 0.51$, then p_1 and p_2 are forced to take other two values, assume $p_1 = 0.16$, and $p_2 = 0.33$. Therefore:

$$\mathbb{E}(\text{Profit}) = -100 + 50 \left[0.49 + (0.49)^{12} \right] = -75.49.$$

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1.2 Generating Functions for Continuous Densities

For the continuous case, the moment generating function $g(t)$ for X is defined as:

$$g(t) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx. \quad (3)$$

1.2.1 Exercise 1 page 401

For this exercise let X be a continuous random variable with values in $[0, 2]$ and for each case there is a given density function f_X .

Case a:

$$\begin{aligned} g(t) &= \int_0^2 \frac{1}{2} e^{tx} dx \\ &= \frac{1}{2} \left[\frac{e^{tx}}{t} \right]_0^2 = \frac{1}{2} \left(\frac{e^{2t} - 1}{t} \right) = \frac{e^{2t} - 1}{2t}. \end{aligned}$$

Case b:

$$\begin{aligned} g(t) &= \int_0^2 \frac{1}{2} x e^{tx} dx \\ &= \frac{1}{2} \left[\frac{(tx - 1)e^{tx}}{t^2} \right]_0^2 = \frac{1}{2} \left[\frac{(2t - 1)e^{2t} + 1}{t^2} \right] = \frac{(2t - 1)e^{2t} + 1}{2t^2}. \end{aligned}$$

Case c:

$$\begin{aligned} g(t) &= \int_0^2 \left(1 - \frac{x}{2}\right) e^{tx} dx \\ &= \left[\frac{e}{2t^2} - \frac{(x - 2)e^{tx}}{2t} \right]_0^2 = \left[-\frac{(t(x - 2) - 1)e^{tx}}{2t^2} \right]_0^2 = \frac{e^{2t}}{2t^2} - \frac{2t + 1}{2t^2} = \frac{e^{2t} - 2t - 1}{2t^2}. \end{aligned}$$

Case d:

$$\begin{aligned} g(t) &= \int_0^2 |1 - x| e^{tx} dx \\ &= \left[\frac{(x - 1)(tx - t - 1)e^{tx}}{t^2|x - 1|} \right]_0^2 = \left[\frac{(t - 1)e^{2t}}{t^2} + \frac{2e^t}{t^2} - \frac{t + 1}{t^2} \right] = \frac{(t - 1)e^{2t} + 2e^t - t - 1}{t^2}. \end{aligned}$$

Case e:

$$\begin{aligned} g(t) &= \int_0^2 \frac{3}{8} x^2 e^{tx} dx \\ &= \left[\frac{3(t^2 x^2 - 2tx + 2)e^{tx}}{8t^3} \right]_0^2 = \frac{(6t^2 - 6t + 3)e^{2t} - 3}{4t^3}. \end{aligned}$$

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1.2.2 Exercise 6 page 402

According to Grinstead and Snell [1], $k_X(\tau)$, called the characteristic function of X is the Fourier transform of f_X , and has an inverse given by

$$f_X(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-i\tau x} k_X(\tau) d\tau. \quad (4)$$

Therefore:

$$\begin{aligned}
 f_X(x) &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-i\tau x} e^{-|\tau|} d\tau, \\
 &= \frac{\frac{ix}{x^2+1} + \frac{1}{x^2+1}}{2\pi} + \frac{\frac{1}{x^2+1} - \frac{ix}{x^2+1}}{2\pi} \\
 &= \frac{1}{\pi(x^2+1)}
 \end{aligned}$$

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1.2.3 Exercise 10 page 403

(a)

$$\begin{aligned}
 \mathbb{E}(X) &= \int_{-\infty}^{\infty} x \frac{e^{-|x|}}{2} dx \\
 &= \frac{1}{2} \int_{-\infty}^{\infty} \frac{x}{|x|} e^{-|x|} |x| dx \\
 &= -\frac{e^{-|x|} |x|}{2} - \frac{e^{-|x|}}{2} \\
 &= \left[\frac{e^{-|x|} (-|x| - 1)}{2} \right]_{-\infty}^{\infty} \\
 &= 0.
 \end{aligned}$$

$$\begin{aligned}
 \mathbb{V}(X) &= \int_{-\infty}^{\infty} x^2 \frac{e^{-|x|}}{2} dx \\
 &= \frac{1}{2} \int_{-\infty}^{\infty} x^2 e^{-|x|} dx \\
 &= \frac{1}{2} \left(\frac{x}{|x|} \right) \int_{-\infty}^{\infty} x^2 e^{-|x|} \frac{|x|}{x} dx \\
 &= \frac{1}{2} \left(\frac{x}{|x|} \right) \int_{-\infty}^{\infty} x e^{-|x|} |x| dx \\
 &= \frac{1}{2} \left[\frac{x(-2e^{-|x|} |x| - x^2 e^{-|x|} - 2e^{-|x|})}{|x|} \right] \\
 &= - \left[\frac{x e^{-|x|} |x| (2|x| - x^2 + 2)}{2|x|} \right]_{-\infty}^{\infty} \\
 &= 2.
 \end{aligned}$$

(b) Let X_1 be a trial process. The moment generating function would be

$$\begin{aligned}
g(t) &= \mathbb{E}(e^{tx}) \\
&= \int_{-\infty}^{\infty} e^{tx} \left[\frac{e^{-|x|}}{2} \right] dx \\
&= \frac{1}{2} \left[\int_{-\infty}^0 e^{xt+x} dx + \int_0^{\infty} e^{xt-x} dx \right] \\
&= \frac{1}{2} \left[\left[\frac{e^{x(t+1)}}{t+1} \right]_{-\infty}^0 - \left[\frac{e^{x(1-t)}}{1-t} \right]_0^{\infty} \right] \\
&= \frac{1}{2} \left[\frac{1}{t+1} + \frac{1}{1-t} \right] \\
&= \frac{1}{2} \left[\frac{2}{1-t^2} \right] \\
&= \frac{1}{1-t^2}.
\end{aligned}$$

If S_n is $X_1 + X_2 + \dots + X_n$, then the moment generating function is given by

$$\begin{aligned}
[g(t)]^n &= \left(\frac{1}{1-t^2} \right)^n \\
&= \frac{1}{(1-t^2)^n}.
\end{aligned}$$

Then, if $S_n^* = \frac{(S_n - n\mu)}{\sqrt{n\sigma^2}}$, the moment generating function is given by

$$\begin{aligned}
\left[g\left(\frac{t}{\sqrt{n}}\right) \right]^n &= \left(\frac{1}{1-t^2} \right)^n \\
&= \left[\frac{1}{\left[1 - \left(\frac{t}{\sqrt{n}} \right)^2 \right]} \right]^n \\
&= \frac{1}{\left[1 - \left(\frac{t}{\sqrt{n}} \right)^2 \right]^n}
\end{aligned}$$

(c) According to the obtained expression of S_n^* as $n \rightarrow \infty$ then the expression may reduce to $\frac{1}{1^n}$. When the limit is calculated it is equal to 1.

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References

- [1] Charles Miller Grinstead and James Laurie Snell. *Introduction to probability*. American Mathematical Soc., 2012.
- [2] The R Foundation. The R Project for Statistical Computing. <https://www.r-project.org/>, 2020.