# Optimal estimation of linear non-Gaussian structure equation models

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# Main Objectives and Outline

#### **Main Objectives**

- Establish the optimal sample complexity for learning Linear non-Gaussian acyclic models (LiNGAM),  $n = \Theta\left(d_{in}\log\frac{p}{d_{in}}\right)$ .
- Develop a LiNGAM learning algorithm using distance covariance that achieves the optimal sample complexity without assuming (parental) faithfulness or a known indegree.
- Introduction
- · Recent Works on LiNGAMs
- New Properties of LiNGAMs
- Proposed Algorithm
- Numerical Experiments & Real Data Analysis

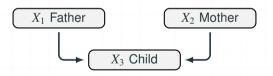


# **Directed Acyclic Graphical Models**

By recovering the directed acyclic graph (DAG)  $\mathcal{G}^*$  from the data  $\mathcal{D}$ ,

- we can learn the conditional dependency structure between variables.
- we can estimate a statistical model of the underlying distribution. In this case, the model can be used to generalize new instances.

Suppose that there are three variables of family gene information,  $X_3 = f(X_1, X_2)$  (functional):



$$X_1 \perp \!\!\!\perp X_2$$
,  $X_1 \not\perp X_2 \mid X_3$ ,  $X_1 \not\perp X_3$ ,  $X_1 \not\perp X_3 \mid X_2$ ,  $X_2 \not\perp X_3$ ,  $X_2 \not\perp X_3 \mid X_1$ .

#### **Definitions**

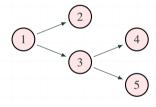


Figure 1: Directed graph G

- Parent (Pa):  $1 \to \{2, 3\}$  and Pa(3) =  $\{1\}$ .
- Child (Ch):  $3 \rightarrow \{4, 5\}$  and Ch(3) =  $\{4, 5\}$ .
- Descendant (De):  $De(1) = \{2, 3, 4, 5\}.$
- Non-descendant (Nd): Nd(3) = {1, 2}, Nd(5) = {1, 2, 3, 4}.
- Topological layers (A<sub>t</sub>): containing nodes whose longest distance to a source node is t.

$$\mathcal{A}_0 = \{1\}, \ \mathcal{A}_1 = \{2, 3\}, \ \mathcal{A}_2 = \{4, 5\}.$$

- ▶ For each  $j \in \mathcal{A}_r$ ,  $Pa(j) \subset \bigcup_{t=0}^{r-1} \mathcal{A}_t \subset Nd(j)$ .
- ▶ Ordering  $(\pi)$ :  $\pi = (1, 2, 3, 4, 5)$  or  $\pi = (1, 3, 2, 5, 4)$ .
- Maximum indegree (d<sub>in</sub>): maximum number of incoming edges,
   d<sub>in</sub> = 1.

#### **Definition of Linear SEMs**

#### **Linear Structural Equation Models**

A Linear SEM is a DAG model where each variable is expressed as a linear function of its parents variables plus an independent error.

$$X_j = \sum_{k \in \mathsf{Pa}(j)} \beta_{k,j} X_k + \epsilon_j, \qquad \forall j = 1, \dots, p,$$

$$(X_1, X_2, ..., X_p)^{\top} = B(X_1, X_2, ..., X_p)^{\top} + (\epsilon_1, \epsilon_2, ..., \epsilon_p)^{\top}.$$

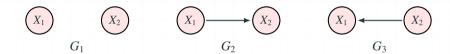
where  $[B]_{j,k} \neq 0$  for all  $k \in Pa(j)$ , otherwise  $[B]_{j,k} = 0$ .

- The error distributions in sub-Gaussian LiNGAM models are continuous and sub-Gaussian, but explicitly non-Gaussian.
  - ▶ If error distributions are uniform, then it is a sub-Gaussian LiNGAM.

# Recent Works on LiNGAMs

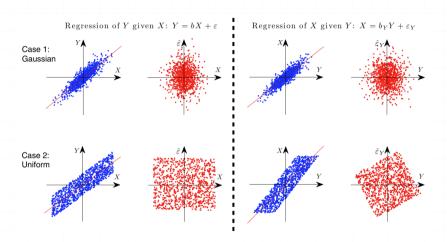
## **Model Identifiability**

A critical issue of graphical models: *Identifiability* 



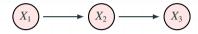
- We can distinguish  $G_2$  and  $G_3$  from  $G_1$ .
- We cannot identify the direction of an edge. Hence, we cannot distinguish  $G_2$  from  $G_3$ .
- Existing algorithms recover the skeleton or CPDAG. (e.g., PC, GES algorithms)

## **Model Identifiability of LiNGAMs**



• The direction of an edge between two variables can be determined by assessing the dependency between (residualized) variables.

# **Ordering Recovery of LiNGAMs**



$$X_1 = \epsilon_1,$$
  $X_2 = \beta_1 X_1 + \epsilon_2$   $X_3 = \beta_2 X_2 + \epsilon_3$   
=  $\beta_1 \epsilon_1 + \epsilon_2,$   $= \beta_1 \beta_2 \epsilon_1 + \beta_2 \epsilon_2 + \epsilon_3,$ 

where  $\epsilon_j$ 's are independent and non-Gaussian.

- (1) Choose an exogenous variable based on non-Gaussianity and independence.  $\rightarrow X_1 = \epsilon_1$ .
- (2) Remove the effect of the exogenous variable from the other variables.

$$\rightarrow e_{2,\{1\}} = X_2 - \Sigma_{2,1}(\Sigma_{1,1})^{-1}X_1 = \epsilon_2, e_{3,\{1\}} = X_3 - \Sigma_{3,1}(\Sigma_{1,1})^{-1}X_1 = \beta_2\epsilon_2 + \epsilon_3.$$

(3) The iteration of effect removal and identification of exogenous variables recovers the ordering of a graph.

$$\rightarrow e_{2,\{1\}} = \epsilon_2$$
 becomes an exogenous variable, whereas  $e_{3,\{1\}} = \beta_2 \epsilon_2 + \epsilon_3$  does not.

# **Details of Ordering Recovery of LiNGAMs**

$$\begin{array}{ccc}
X_1 & \xrightarrow{\beta_1} & X_2 & \xrightarrow{\beta_2} & X_3 \\
\epsilon_1 & & \beta_1 \epsilon_1 + \epsilon_2 & & \beta_1 \beta_2 \epsilon_1 + \beta_2 \epsilon_2 + \epsilon_3
\end{array}$$

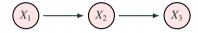
$X_1 = \epsilon_1$	$e_{2,\{1\}}=\epsilon_2$	$e_{3,\{1\}} = \beta_2 \epsilon_2 + \epsilon_3$
$e_{1,\{2\}} = (1 - \Sigma_{1,2}(\Sigma_{2,2})^{-1}\beta_1)\epsilon_1 - \Sigma_{1,2}(\Sigma_{2,2})^{-1}\epsilon_2$	$X_2 = \beta_1 \epsilon_1 + \epsilon_2$	$e_{3,\{2\}}$
e <sub>1,{3}</sub>	$e_{2,\{3\}}$	$X_3 = \beta_1 \beta_2 \epsilon_1 + \beta_2 \epsilon_2 + \epsilon_3$

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$e_{2,\{1\}}=\epsilon_2$	$e_{3,\{1\}} - \frac{Cov_{(e_{2,\{1\}},e_{3,\{1\}})}}{Var_{(e_{2,\{1\}})}} e_{2,\{1\}} = \beta_2 \epsilon_2 + \epsilon_3 - \beta_2 \epsilon_2 = \epsilon_3$
$e_{2,\{1\}} - \frac{Cov_{(e_{3,\{1\}},e_{2,\{1\}})}}{Var_{(e_{3,\{1\}})}} e_{3,\{1\}} = \epsilon_2 - \frac{Cov_{(e_{3,\{1\}},e_{2,\{1\}})}}{Var_{(e_{3,\{1\}})}} (\beta_2 \epsilon_2 + \epsilon_3)$	$e_{3,\{1\}} = \beta_2 \epsilon_2 + \epsilon_3$

 $\checkmark$  For the algorithm, a least squares regression with p-1 predictors is needed, making n < p infeasible.

# Strategies for the scenario when n < p (1)



$$X_1 = \epsilon_1,$$
  $X_2 = \beta_1 X_1 + \epsilon_2$   $X_3 = \beta_2 X_2 + \epsilon_3$   
=  $\beta_1 \epsilon_1 + \epsilon_2,$   $= \beta_1 \beta_2 \epsilon_1 + \beta_2 \epsilon_2 + \epsilon_3,$ 

where  $\epsilon_j$ 's are independent and non-Gaussian.

#### Wang and Drton (2020)

- This study interprets the effect removal step as the elimination of confounding effects.
- By identifying a set of variables that eliminates all confoundings between variables, the ordering is recovered in a top-down manner.
- The assumption of parental faithfulness is required.

# Strategies for the scenario when n < p (2)



$$X_1 = \epsilon_1,$$
  $X_2 = \beta_1 X_1 + \epsilon_2$   $X_3 = \beta_2 X_2 + \epsilon_3$   
=  $\beta_1 \epsilon_1 + \epsilon_2,$   $= \beta_1 \beta_2 \epsilon_1 + \beta_2 \epsilon_2 + \epsilon_3,$ 

where  $\epsilon_j$ 's are independent and non-Gaussian.

#### Zhao et al. (2022)

• The ordering is recovered in a bottom-up manner by utilizing the independence of the exogenous term of terminal nodes from all preceding variables.

$$ightharpoonup \epsilon_3 \perp \!\!\! \perp X_1$$
 and  $\epsilon_3 \perp \!\!\! \perp X_2$ , but  $\epsilon_2 \not\perp \!\!\! \perp X_3$ .

• The exogenous term is revealed by projecting each variable onto its Markov blanket.

# New Properties of LiNGAMs

# **Expanding Scope of Independence Analysis**

For  $j, k \in V$  and  $C \subset V \setminus \{j, k\}$ , denote the residuals as

• 
$$e_{j,C} = X_j - \Sigma_{j,C}(\Sigma_{C,C})^{-1}X_C$$
.

• 
$$r_C(j,k) = e_{j,C} - \frac{\mathsf{Cov}(e_{j,C},e_{k,C})}{\mathsf{Var}(e_{k,C})} e_{k,C}.$$

Examine the independence relationships among (residualized) variables for  $C \subset S$ .

$$(\begin{bmatrix} S \subset \mathsf{Nd}(j) \\ X_{\ell}, \ \ell \in S \end{bmatrix}) \longrightarrow (\begin{bmatrix} V \setminus (S \cup \{j\}) \\ r_{C}(j,k), \ k \in V \setminus (S \cup \{j\}) \end{bmatrix})$$

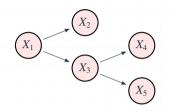
# **Optimizing Exogeneity Identification**

#### **Proposition 1**

Let P(X) be generated from a LiNGAM with DAG G and the topological layers  $\bigcup_{t=0}^{T-1} \mathcal{A}_t$ . For any  $r \in \{1, 2, ..., T-1\}, j \in \mathcal{A}_r, k \in V \setminus \bigcup_{t=0}^r \mathcal{A}_t$ , and  $S_{r-1} = \bigcup_{t=0}^{r-1} \mathcal{A}_t$ ,

- (i) there exists  $C \subset S_{r-1}$  satisfying  $\mathsf{Pa}(j) \subset C$ , and then  $e_{j,C} \perp \!\!\! \perp r_C(\ell,j)$  for all  $\ell \in V \setminus (S_{r-1} \cup \{j\})$ .
- (ii) if  $C \subset S_{r-1}$  satisfies  $e_{j,C} \perp \!\!\! \perp X_{\ell}$  for all  $\ell \in S_{r-1}$ , then  $\text{Pa}(j) \subset C$ .
- (iii) for any  $C \subset S_{r-1}$ , there exists  $\ell \in V \setminus (S_{r-1} \cup \{k\})$  such that  $e_{k,C} \not\perp r_C(\ell,k)$ .
  - Topological layers of a LiNGAM can be recovered iteratively by (i) and (iii) in a top-down manner.
    - ▶ The set of variables required to make a variable exogenous is reduced to its parent set rather than the entire set of preceding elements in the causal ordering or Markov blanket.
  - By (ii), the minimal conditioning set C ⊂ S<sub>r-1</sub>, for which e<sub>i,C</sub> is independent of X<sub>ℓ</sub> for all ℓ ∈ S<sub>r-1</sub>, is Pa(j).
    - ▶ Notably, (ii) does not rely on the assumption of parental faithfulness.

# **Example**



 $\epsilon_{\!\scriptscriptstyle j}$  are independent and non-Gaussian.

$$X_{1} = \epsilon_{1}$$

$$X_{2} = \beta_{1,2}X_{1} + \epsilon_{2}$$

$$X_{3} = \beta_{1,3}X_{1} + \epsilon_{3}$$

$$X_{4} = \beta_{3,4}X_{3} + \epsilon_{4}$$

$$X_{5} = \beta_{3,5}X_{3} + \epsilon_{5}$$

- Since  $X_1$  is the only exogenous variable,  $\mathcal{A}_0 = \{1\}$ .
- For r = 1,

(i) 
$$Pa(2) = Pa(3) = S_0 = \{1\}.$$

$$\Rightarrow$$
  $e_{2,\{1\}} = \epsilon_2$  and  $e_{3,\{1\}} = \epsilon_3$  are exogenous.

$$\Rightarrow e_{2,\{1\}} \perp r_{\{1\}}(\ell,2)$$
 for  $\ell \in \{3,4,5\}$ , and similarly for  $e_{3,\{1\}}$ .

(iii) 
$$Pa(4) = Pa(5) = \{3\} \not\subset S_0$$
.

$$\Rightarrow e_{4,\{1\}} = \beta_{3,4}\epsilon_3 + \epsilon_4$$
 is dependent of

$$r_{\{1\}}(3,4) = \epsilon_3 - \frac{\beta_{3,4} \text{Var}(\epsilon_3)}{\beta_{3,4}^2 \text{Var}(\epsilon_3) + \text{Var}(\epsilon_4)} (\beta_{3,4} \epsilon_3 + \epsilon_4),$$

and similarly for  $e_{5,\{1\}}$ .

(ii) 
$$e_{2,\{1\}} \perp X_1$$
,  $e_{3,\{1\}} \perp X_1$  and  $Pa(2) = Pa(3) = \{1\}$ .

# **Topological Layer Recovery**

#### Theorem 2

Let P(X) be generated from a LiNGAM with DAG G and the topological layers  $\bigcup_{t=0}^{T-1} \mathcal{A}_t = V$ . Consider any  $r \in \{1, 2, ..., T-1\}$  and  $S_{r-1} = \bigcup_{t=0}^{r-1} \mathcal{A}_t$ . Then

$$\mathcal{H}_0 = \left\{ j \in V : X_j \perp \!\!\!\!\perp e_{k,(j)} \text{ for all } k \in V \setminus \{j\} \right\}, \text{ and}$$

$$\mathcal{H}_r = \left\{ j \in V \setminus S_{r-1} : \exists C_j \subset S_{r-1} \text{ s.t. } X_k \perp \!\!\!\!\perp e_{j,C_i} \text{ for all } k \in S_{r-1} \text{ and } e_{j,C_i} \perp \!\!\!\!\perp r_{C_i}(\ell,j) \text{ for all } \ell \in V \setminus (S_{r-1} \cup \{j\}) \right\}.$$

Moreover, for each  $j \in \mathcal{A}_r$  and the corresponding set

$$C_j = \{C \subset S_{r-1} : X_k \perp \!\!\!\perp e_{j,C} \text{ for all } k \in S_{r-1} \text{ and } e_{j,C} \perp \!\!\!\perp r_C(\ell,j) \text{ for all } \ell \in V \setminus (S_{r-1} \cup \{j\})\},$$

 $Pa(j) \subset C \subset Nd(j)$  for any  $C \in C_i$ .

• Theorem 2 also ensures that the true ordering can be correctly identified, by substituting  $\pi_r \in \pi$  as a true ordering and setting  $S_{r-1} = \{\pi_1, ..., \pi_{r-1}\}$  with  $S_0 = \emptyset$ .

# Proposed Algorithm

# **Dependency Score**

For  $j \in V \setminus \mathcal{R}$ , define a dependency score relative to the preceding elements in the ordering  $\mathcal{R}$ .

$$\widehat{\mathcal{S}}(j,C) := \#\{k \in \mathcal{R} : \hat{e}_{j,C} \not\perp_{\text{test }} \mathbf{x}_k\} + \#\{\ell \in V \setminus (\mathcal{R} \cup \{j\}) : \hat{e}_{j,C} \not\perp_{\text{test }} \hat{r}_C(\ell,j)\},$$

where

$$\hat{e}_{j,C} = \mathbf{x}_j - \mathbf{x}_C(\widehat{\Sigma}_{C,C})^{-1}\widehat{\Sigma}_{Cj}, \quad \hat{r}_C(j,k) = \hat{e}_{j,C} - \frac{\widehat{\Sigma}_{j,k} - \widehat{\Sigma}_{j,C}(\widehat{\Sigma}_{C,C})^{-1}\widehat{\Sigma}_{C,k}}{\widehat{\Sigma}_{k,k} - \widehat{\Sigma}_{k,C}(\widehat{\Sigma}_{C,C})^{-1}\widehat{\Sigma}_{C,k}} \hat{e}_{k,C}.$$

- - ▶ (e.g.) Hilbert-Schmidt independence criterion, distance covariance measure.
- $\exists C \subset \mathcal{R}$  for which  $\widehat{\mathcal{S}}(j, C) = 0$ .
  - $\Leftrightarrow$  *j* is a source node in the subgraph obtained after removing  $\mathcal R$  by Proposition 1 (i) and (iii).
  - $\Rightarrow$  Pa(j)  $\subset$  C  $\subset$   $\mathcal{R}$  by Proposition 1 (ii).

# **OptLiNGAM Algorithm**

#### **OptLiNGAM Algorithm**

**Input:** n i.i.d. samples  $\mathbf{x}^{1:n}$  and significance level  $\alpha$ .

**Output:** Estimated graph,  $\widehat{G} = (V, \widehat{E})$ .

Step (1): Source nodes estimation.

$$\widehat{\mathcal{A}}_0 = \left\{ j \in V : \mathbf{x}_j \perp \!\!\!\! \perp_{\text{test}} \widehat{e}_{k, \{j\}} \text{ for all } k \in V \setminus \{j\} \right\}.$$

**Step (2):** Directed edges (parent) estimation Initialize  $\mathcal{R} = \widehat{\mathcal{A}}_0$ .

While  $V \setminus \mathcal{R} \neq \emptyset$ :

For 
$$q \in \{1, 2, ..., |\mathcal{R}|\}$$
:

Set 
$$\mathcal{R}_0 = \emptyset$$
.

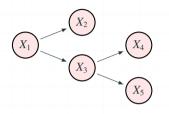
For 
$$j \in V \setminus \mathcal{R}$$
:

If 
$$\min_{|C|=q,C\subset\mathcal{R}}\widehat{\mathcal{S}}(j,C)=0$$
: Update  $\mathcal{R}_0=\mathcal{R}_0\cup\{j\}$  and  $\widehat{\mathsf{Pa}}(j)=\mathop{\mathrm{arg\,min}}_{|C|=q,C\subset\mathcal{R}}\widehat{\mathcal{S}}(j,C)$ .

If  $\mathcal{R}_0 \neq \emptyset$ : Update  $\mathcal{R} = \mathcal{R} \cup \mathcal{R}_0$  and **Break**.

**Return:** 
$$\widehat{E} = \{(k,j) : j \in V, k \in \widehat{\mathsf{Pa}}(j)\}.$$

# **Details of the OptLiNGAM Algorithm (1)**



 $\epsilon_{\!\scriptscriptstyle j}$  are independent and non-Gaussian.

$$X_1 = \epsilon_1$$
  
 $X_2 = \beta_{1,2}X_1 + \epsilon_2$   
 $X_3 = \beta_{1,3}X_1 + \epsilon_3$   
 $X_4 = \beta_{3,4}X_3 + \epsilon_4$   
 $X_5 = \beta_{3,5}X_3 + \epsilon_5$ 

- Since  $X_1$  is the only exogenous variable,  $\mathcal{A}_0 = \{1\}$ .
- $e_{2,\{1\}} = \epsilon_2$  and  $e_{3,\{1\}} = \epsilon_3$  are exogenous.

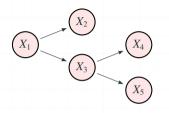
$$\Rightarrow S(2, \{1\}) = S(3, \{1\}) = 0.$$

•  $e_{4,[1]}=\beta_{3,4}\epsilon_3+\epsilon_4$  and  $e_{5,[1]}=\beta_{3,5}\epsilon_3+\epsilon_5$ , and they are dependent of

$$\begin{split} r_{\{1\}}(3,4) &= \epsilon_3 - \frac{\beta_{3,4} \text{Var}(\epsilon_3)}{\beta_{3,4}^2 \text{Var}(\epsilon_3) + \text{Var}(\epsilon_4)} (\beta_{3,4} \epsilon_3 + \epsilon_4), \text{ and} \\ r_{\{1\}}(3,5) &= \epsilon_3 - \frac{\beta_{3,5} \text{Var}(\epsilon_3)}{\beta_{3,5}^2 \text{Var}(\epsilon_3) + \text{Var}(\epsilon_5)} (\beta_{3,5} \epsilon_3 + \epsilon_5). \\ &\Rightarrow \mathcal{S}(4,\{1\}), \mathcal{S}(5,\{1\}) \geq 1. \end{split}$$

•  $\mathcal{A}_1 = \{2, 3\}$  and  $Pa(2) = Pa(3) = \{1\}$ .

# **Details of the OptLiNGAM Algorithm (2)**



 $\epsilon_{j}$  are independent and non-Gaussian.

$$X_{1} = \epsilon_{1}$$

$$X_{2} = \beta_{1,2}X_{1} + \epsilon_{2}$$

$$X_{3} = \beta_{1,3}X_{1} + \epsilon_{3}$$

$$X_{4} = \beta_{3,4}X_{3} + \epsilon_{4}$$

$$X_{5} = \beta_{3,5}X_{3} + \epsilon_{5}$$

Note that

$$e_{4,\{3\}} = e_{4,\{2,3\}} = e_{4,\{1,3\}} = e_{4,\{1,2,3\}} = \epsilon_4,$$
  
 $e_{5,\{3\}} = e_{5,\{2,3\}} = e_{5,\{1,3\}} = e_{5,\{1,2,3\}} = \epsilon_5.$ 

- We have  $S(4, \{3\}) = S(5, \{3\}) = 0$ .
- The minimal set  $\{3\}$  is identified as  $Pa(4) = Pa(5) = \{3\}$ .
- $\checkmark$  The number of predictors required for projection decreases from p-1 or d to  $d_{in}+1$ .

#### **Theoretical Results**

#### Theorem 5: Consistency of the OptLiNGAM Algorithm

Consider a sub-Gaussian LiNGAM with  $d_{in} \leq \frac{p}{2}$ . The proposed algorithm utilizes a distance covariance-based independence test a the significance level of  $\alpha_n = 2(1 - \Phi(\sqrt{n\epsilon}))$ , where  $\epsilon \in (0, \tau_1/2)$ . Then, under regularity conditions, there exist positive  $A_1 > 0$  and  $A_2 > 0$  such that

$$P(\widehat{G} = G) \ge 1 - A_1(p/d_{in})^{d_{in}} \exp(-A_2 n).$$

#### Lemma 6: Lower Bound of Sample Complexity for Arbitrary Estimator

For any  $0 < \delta < 1/2$ , there are positive  $B_1 > 0$  and  $B_2 > 0$  such that for any estimator  $\widehat{G}$ ,

$$n \leq (1 - 2\delta)B_1 d_{in} \log(p/d_{in}) \implies \sup_{F \in \mathcal{F}_{p,d_{in}}} P(\widehat{G} \neq G(F)) \geq \delta - \frac{B_2}{p d_{in} \log(p/d_{in})},$$

where  $\mathcal{F}_{p,d_{in}}$  is a class of p-node sub-Gaussian LiNGAMs with  $d_{in}$ , which satisfy regularity conditions.

#### Corollary 7: Optimality of the OptLiNGAM Algorithm

The proposed algorithm is optimal in sample complexity  $n = d_{in} \log(p/d_{in})$  under regularity conditions.

# **Comparison to Existing Algorithms**

Identifiability	Sample complexity	Assumption
Non-Gaussianity	$\Omega(d_{in}\log(p/d_{in}))$	-
Non-Gaussianity	$\Omega(T^{c_1}d^6\log(p)^{c_2})$	Incoherence
Non-Gaussianity	$\Omega((\log p)^{2K})$	Parental faithfulness
Forward condition	$\Omega(d_{in}\log(p/d_{in}))$	Knowledge of $d_{in}$
Backward condition	$\Omega(d^2 \log p)$	Incoherence
Backward condition	$\Omega(d^2 \log p)$	Incoherence
	Non-Gaussianity Non-Gaussianity Non-Gaussianity Forward condition Backward condition	$\begin{array}{lll} & \text{Non-Gaussianity} & \Omega(d_{in}\log(p/d_{in})) \\ & \text{Non-Gaussianity} & \Omega(T^{c_1}d^6\log(p)^{c_2}) \\ & \text{Non-Gaussianity} & \Omega((\log p)^{2K}) \\ & \text{Forward condition} & \Omega(d_{in}\log(p/d_{in})) \\ & \text{Backward condition} & \Omega(d^2\log p) \end{array}$

TL (Zhao et al., 2022), MDirect (Wang and Drton, 2020), OptFGSM (Gao et al., 2022), HLSM (Park et al., 2021), LISTEN (Ghoshal and Honorio, 2018).

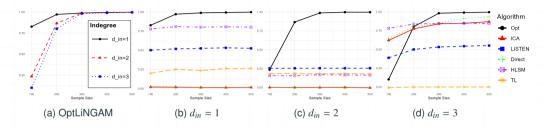
- The current sample complexity for LiNGAMs surpasses that of Gaussian SEMs, limiting its application in fields where high-dimensional causal discovery is challenged by very limited sample sizes.
- OptLiNGAM is the most sample efficient among LiNGAM learning algorithms without assuming faithfulness or a known indegree.
- This is the first result to establish optimal sample complexity for high-dimensional LiNGAMs, with an upper bound optimal up to constant factors.



## **Simulation Settings**

- Three types of hub graphs are generated with  $d_{in} \in \{1, 2, 3\}$ .
  - ▶ The number of hub nodes corresponds to  $d_{in}$ .
  - $\triangleright \lfloor \log p \rfloor$  nodes are isolated, and the remaining nodes are children of the hub nodes.
- The data generation process is repeated 30 times.
  - $\triangleright$  Error terms are drawn from Beta(0.5, 0.5).
  - ▶ Nonzero edge weights are sampled uniformly from  $[-1.5, -0.5] \cup [0.5, 1.5]$ .
- All algorithms are evaluated using the Matthews correlation coefficient (MCC), which measures the
  accuracy of the estimated directed edges.
  - $\triangleright$  +1 indicates a perfect prediction, 0 represents an average random prediction, and -1 signifies an inverse prediction.

# **Verification of Consistency**



**Figure 1:** Average MCC for 50-node sparse hub graphs by varying sample size  $n \in \{100, 200, ..., 500\}$ .

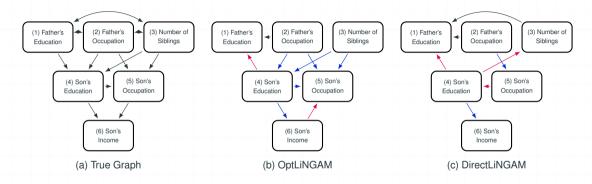
- As *n* increases, the average MCC converges to 1, signifying perfect graph recovery.
- OptLiNGAM requires fewer samples to recover sparse graphs with lower maximum indegree.
- OptLiNGAM significantly outperforms other methods as sample size increases, underscoring its advantages: being sample optimal and not requiring restrictive conditions.

# Real Data Analysis

# **General Social Survey Data**

- General Social Survey of U.S. adults(http://www.norc.org/GSS+Website) with six variables:
  - (1) Father's Education, (2) Father's Occupation, (3) Number of Siblings,
  - (4) Son's Education, (5) Son's Occupation, and (6) Son's Income.
- Duncan et al. (1972) has provided the true graph based on domain knowledge.
- This data was analyzed by DirectLiNGAM in Shimizu et al. (2011), specifically focused on samples for 45 years, from 1972 to 2006.
- This analysis employs data for five years, from 2002 to 2006, consisting of 355 observations, to demonstrate the ability of OptLiNGAM for sample efficient LiNGAM recovery.

# **Graphs Estimated by OptLiNGAM and DirectLiNGAM**



- OptLiNGAM successfully identifies most causal relationships between variables, except for two reversed edges, (4, 1) and (6, 5).
- DirectLiNGAM fails to capture important connections, such as (2, 4) and (3, 5), and incorrectly assigns directions to edges, such as (4, 3) and (5, 4), all of which are correctly identified by OptLiNGAM.

#### **Discussion**

#### **Summary**

- Exogeneity identification optimized by reducing the conditioning set to the parent set.
- Sample efficient algorithm with sample complexity based on d<sub>in</sub> without requiring incoherence or parental faithfulness.
- Sample optimality demonstrated by aligning lower and upper bounds of sample complexity, without prior knowledge of d<sub>in</sub>.

#### **Future works**

- Computationally efficient algorithms while maintaining comparable sample efficiency.
- Consistency and potential optimality of the proposed algorithm under heavy-tailed errors.

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Thank you!

## **Appendix - Darmois-Skitovitch Theorem**

#### Darmois-Skitovitch Theorem (Darmois, 1953; Skitovitch, 1953)

Define two random variables  $u_1$  and  $u_2$  as linear combinations of independent random variables  $s_i, i = 1, ..., m$ , such that

$$u_1 = \sum_{i=1}^m c_{1,i} s_i$$
 and  $u_2 = \sum_{i=1}^m c_{2,i} s_i$ .

If  $u_1$  and  $u_2$  are independent, all variables  $s_i$  with  $c_{1,i}c_{2,i} \neq 0$  are Gaussian distributed.

# **Appendix - Distance Covariance**

$$T(e_{j,C}, X_k) = \frac{\mathsf{dcov}^2(e_{j,C}, X_k)}{I_{jk,2}}, \quad \text{and} \quad \widehat{T}(\hat{e}_{j,C}, \mathbf{x}_k) = \frac{\widehat{\mathsf{dcov}}^2(\hat{e}_{j,C}, \mathbf{x}_k)}{\widehat{I}_{jk,2}}.$$

Here, for any set  $S \subset V \setminus \{j\}$ ,  $dcov^2(e_{j,C}, X_S) = I_{jS,1} + I_{jS,2} - 2I_{jS,3}$  with

$$\begin{split} I_{jS,1} &= \mathbb{E}[|e_{j,C} - e'_{j,C}|||X_S - X'_S||], \\ I_{jS,2} &= \mathbb{E}[|e_{j,C} - e'_{j,C}|]\mathbb{E}[||X_S - X'_S||], \\ I_{jS,3} &= \mathbb{E}\left[\mathbb{E}[|e_{j,C} - e'_{j,C}| \mid e_{j,C}]\mathbb{E}[||X_S - X'_S|| \mid X_S]\right], \end{split}$$

where the notation 'denotes and independent copy of the corresponding random vector.

Additionally,  $\widehat{\text{dcov}}^2(\hat{e}_{j,C}, \mathbf{x}_S) = \widehat{I}_{jS,1} + \widehat{I}_{jS,2} - 2\widehat{I}_{jS,3}$ , where

$$\begin{split} \widehat{I}_{jS,1} &= \frac{1}{n^2} \sum_{i,h=1}^{n} |\widehat{e}_{j,C}^{(i)} - \widehat{e}_{j,C}^{(h)}||\mathbf{x}_S^{(i)} - \mathbf{x}_S^{(h)}||, \\ \widehat{I}_{jS,2} &= \left(\frac{1}{n^2} \sum_{i,h=1}^{n} |\widehat{e}_{j,C}^{(i)} - \widehat{e}_{j,C}^{(h)}|\right) \left(\frac{1}{n^2} \sum_{i,h=1}^{n} ||\mathbf{x}_S^{(i)} - \mathbf{x}_S^{(h)}||\right), \\ \widehat{I}_{jS,3} &= \frac{1}{n^3} \sum_{i,h,m=1}^{n} |\widehat{e}_{j,C}^{(i)} - \widehat{e}_{j,C}^{(m)}|||\mathbf{x}_S^{(h)} - \mathbf{x}_S^{(m)}||. \end{split}$$

# **Appendix - Comparison to State-of-the-art Algorithms**

Table 1: Empirical probability of successful graph recovery of all algorithms in high-dimensional settings.

(n,p)	Method	Recovery Rate	(n, p)	Method	Recovery Rate
(300,400)	Opt	0.5333	(300,500)	Opt	0.3
	TL	0		TL	0
	LISTEN	0		LISTEN	0
	HLSM	0		HLSM	0

- The proposed method achieves optimal sample complexity at a high computational cost; if graph recovery exceeds 24 hours, it is considered a failure.
- OptLiNGAM achieves empirical probabilities of successful graph recovery of 0.5333 and 0.3 for p = 400 and p = 500, respectively.
- None of the comparison methods successfully recover the true graph.

# **Appendix - Graph Structure Learning in Different Simulation Setting**

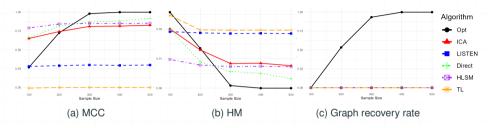


Figure 2: Average MCC, HM, and true graph recovery rate for 50-node graphs with  $d_{in} = 1$  and  $n \in \{100, 200, ..., 500\}$ .

- Start with a three-node graph that has directed edges from node 1 to nodes 2 and 3.
- A new node is added at each step with a directed edge, where the probability of an existing node connecting to the new node depends on its number of neighbors.
- Only OptLiNGAM consistently achieves true graph recovery across all sample sizes considered.