

Prove Gamma-Poisson conjugation

likelihood $\hat{x} \sim \text{Poisson}(\lambda)$ pmf: $p(x|\lambda) = \frac{\lambda^x e^{-\lambda}}{x!}$

prior $\lambda \sim \text{Gamma}(\alpha, \beta)$ $p(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$

$$p(\lambda|x) \propto p(x|\lambda) \cdot p(\lambda)$$

$$\propto \frac{\lambda^x e^{-\lambda}}{x!} \cdot \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$$

$$\propto \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{(\alpha+x)-1} e^{-(\beta+1)\lambda}$$

$$\text{Gamma}(\lambda|\alpha', \beta') = \frac{\beta'^{\alpha'}}{\Gamma(\alpha')} \lambda^{\alpha'-1} e^{-\beta'\lambda} \quad \alpha' = \alpha + x \quad \beta' = \beta + 1$$

\therefore Gamma prior & Poisson pmf의 합부분적분

Posterior prior은 Gamma prior

Prove mean and variance with Gaussian prior $w \sim N(\mu_0, \Lambda_0^{-1})$

$$\text{prior } p(w) = N(\mu_0, \Lambda_0^{-1})$$

$$p(w) \propto \exp\left(-\frac{1}{2}(w - \mu_0)^T \Lambda_0 (w - \mu_0)\right)$$

$$= \exp\left(-\frac{1}{2}[w^T \Lambda_0 w - 2\mu_0^T \Lambda_0 w + \mu_0^T \Lambda_0 \mu_0]\right)$$

$$(w^T \Lambda_0 w)^{-1} = w^T \phi^T \phi w$$

$$\text{likelihood } p(y|w) = N(\phi w, \sigma^2 I)$$

$$p(y|w) \propto \exp\left(-\frac{1}{2\sigma^2} (y - \phi w)^T (y - \phi w)\right)$$

$$= \exp\left(-\frac{1}{2\sigma^2} [y^T y - 2y^T \phi w + w^T \phi^T \phi w]\right)$$

Bayes theorem $p(w|y) \propto p(y|w) \cdot p(w)$

$$p(w|y) = \exp\left(-\frac{1}{2}[\mathbf{w}^T \Lambda_0 \mathbf{w} - 2\mathbf{y}^T \Lambda_0 \mathbf{w} + \mathbf{y}^T \Lambda_0 \mathbf{y}]\right) \times \exp\left(-\frac{1}{2a}[\mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \Phi \mathbf{w} + \mathbf{w}^T \Phi^T \Phi \mathbf{w}]\right)$$
$$\mu_0^T \Lambda_0 \mathbf{w} + \frac{1}{a} \mathbf{y}^T \Phi^T \Phi \mathbf{w}$$
$$\log p(w|y) = -\frac{1}{2}[\mathbf{w}^T \Lambda_0 \mathbf{w} - 2\mathbf{y}^T \Lambda_0 \mathbf{w} + \mathbf{y}^T \Lambda_0 \mathbf{y}] + \mathbf{w}^T \mu_0 + \frac{1}{a} \mathbf{y}^T \Phi^T \Phi \mathbf{w}$$
$$= -\frac{1}{2} \mathbf{w}^T (\Lambda_0 + \frac{1}{a} \Phi^T \Phi) \mathbf{w} + \mathbf{w}^T \left(\frac{1}{a} \Phi^T \mathbf{y} + \Lambda_0 \mu_0\right) + \text{constant}$$

$$\log p(w) = -\frac{1}{2} \mathbf{w}^T \Lambda w + \mathbf{w}^T b$$
$$\Lambda = \Sigma_N^{-1}$$
$$\mu_N = \Lambda^{-1} \cdot b$$

$$\Sigma_N = (\Lambda_0 + \frac{1}{a} \Phi^T \Phi)^{-1}$$

$$\mu_N = \Sigma_N \left(\frac{1}{a} \Phi^T \mathbf{y} + \Lambda_0 \mu_0 \right)$$

Bayes Linear Regression

precision 精度

initial prior $\mathbf{w} \sim N(0, b^{-1} \mathbf{I}) \Rightarrow \mu = 0, \Lambda_0 = b \cdot \mathbf{I}$

$$\Sigma_N = (b \mathbf{I} + \frac{1}{a} \Phi^T \Phi)^{-1}$$

每筆資料帶來的資訊

$$\mu = \Sigma_N \left(\frac{1}{a} \Phi^T \mathbf{y} \right)$$

Φ column vector ($n \times 1$)

每筆資料對 mean by 贏得 $m = \frac{1}{a} \sum y_i \Phi(x_i)$