

# **ESMA 6787: Exam 2**

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*Israel Almodovar*

**Alejandro Ouslan**

## Problem 1: Acceptance of Syllabus

I have read the syllabus, understand its contents, and have no questions.

## Problem 2: Definitions

- (a) **Sample Space:** The sample space is the set of all possible outcomes of an experiment. It is to say everything that could happen in the experiment
- (b) **Kolmogorov Axioms of Probability:** It is a set of rules that define the probability of an event. Those rules can be summarized by the sum of the probability of all the events is 1, the probability of a given event should be non negative and if the events are disjoint then the probability of the union of the events is the sum of the probability of the events
- (c) **Exponential family:** A family of distributions is said to be an exponential family if the pdf can be written in the form
- (d) **Convergence in distribution:**
- (e) **Convergence in Probability:**
- (f) **Almost sure convergence (or convergence with probability 1):**
- (g) **Weak law of large numbers:**
- (h) **Strong law of large numbers:**
- (i) **Characteristics functions:**

## Problem 3:

Show that each of the following families of distributions is an exponential family,

- (a) The family of Bernoulli distribution with a unknown value of the parameter  $p$ .

$$f(x; p) = p^x (1 - p)^{1-x}, \quad x \in \{0, 1\}$$

$$f(x; p) = \exp(x \log(p) + (1 - x) \log(1 - p))$$

Hence, it is in the form of an exponential family with natural parameter  $\theta = \log\left(\frac{p}{1-p}\right)$ , sufficient statistic  $T(x) = x$ , and log-partition function  $A(\theta) = \log(1 + \exp(\theta))$ .

- (b) The family of Poisson distributions with an unknown mean  $\lambda$

$$f(x; \lambda) = \frac{\lambda^x e^{-\lambda}}{x!}, \quad x \in \{0, 1, 2, \dots\}$$

$$f(x; \lambda) = \exp(x \log(\lambda) - \lambda - \log(x!))$$

Hence, it is in the form of an exponential family with natural parameter  $\theta = \log(\lambda)$ , sufficient statistic  $T(x) = x$ , and log-partition function  $A(\theta) = e^\theta$ .

- (c) The family of negative binomial distributions for which the value of  $r$  is known and the value of  $p$  is unknown.

$$f(x; p) = \binom{x+r-1}{x} p^x (1-p)^r, \quad x = 0, 1, 2, \dots$$

$$f(x; p) = \exp\left(x \log(p) + (r-1) \log(1-p) + \log\left(\binom{x+r-1}{x}\right)\right)$$

Thus, it is in the form of an exponential family with natural parameter  $\theta = \log\left(\frac{p}{1-p}\right)$ , sufficient statistic  $T(x) = x$ , and log-partition function  $A(\theta) = -r \log(1 - \exp(\theta))$ .

- (d) The family of normal distributions with an unknown mean and a known variance  $\sigma^2$ .

$$f(x; \mu) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

$$f(x; \mu) = \exp\left(-\frac{1}{2\sigma^2}(x^2 - 2\mu x + \mu^2)\right)$$

Hence, it is in the form of an exponential family with natural parameter  $\theta = \frac{\mu}{\sigma^2}$ , sufficient statistic  $T(x) = x$ , and log-partition function  $A(\theta) = \frac{\mu^2}{2\sigma^2}$ .

- (e) The family of normal distributions with an unknown variance and a known mean.

$$f(x; \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

$$f(x; \sigma^2) = \exp\left(-\frac{1}{2\sigma^2}(x - \mu)^2 - \frac{1}{2}\log(2\pi\sigma^2)\right)$$

Hence, it is in the form of an exponential family with natural parameter  $\theta = -\frac{1}{2\sigma^2}$ , sufficient statistic  $T(x) = (x - \mu)^2$ , and log-partition function  $A(\theta) = -\frac{1}{2}\log(-2\pi\theta)$ .

- (f) The family of gamma distributions for which the value of  $\alpha$  is unknown and the value of  $\beta$  is known.

$$f(x; \alpha) = \frac{x^{\alpha-1}e^{-x/\beta}}{\Gamma(\alpha)\beta^\alpha}, \quad x > 0$$

$$f(x; \alpha) = \exp\left((\alpha - 1)\log(x) - \frac{x}{\beta} - \log(\Gamma(\alpha)) - \alpha\log(\beta)\right)$$

Hence, it is in the form of an exponential family with natural parameter  $\theta = -\frac{1}{\beta}$ , sufficient statistic  $T(x) = x$ , and log-partition function  $A(\theta) = -\alpha\log(-\theta\beta)$ .

- (g) The family of gamma distributions for which the value of  $\alpha$  is known and the value of  $\beta$  is unknown.

$$f(x; \beta) = \frac{x^{\alpha-1}e^{-x/\beta}}{\Gamma(\alpha)\beta^\alpha}, \quad x > 0$$

$$f(x; \beta) = \exp\left((\alpha - 1)\log(x) - \frac{x}{\beta} - \log(\Gamma(\alpha)) - \alpha\log(\beta)\right)$$

Hence, it is in the form of an exponential family with natural parameter  $\theta = -\frac{1}{\beta}$ , sufficient statistic  $T(x) = x$ , and log-partition function  $A(\theta) = \alpha\log(-\theta)$ .

- (h) The family of beta distributions for which the value of  $\alpha$  is unknown and the value of  $\beta$  is known.

$$f(x; \alpha) = \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha, \beta)}, \quad x \in [0, 1]$$

$$f(x; \alpha) = \exp((\alpha - 1)\log(x) + (\beta - 1)\log(1 - x) - \log(B(\alpha, \beta)))$$

Hence, it is in the form of an exponential family with natural parameter  $\theta_1 = \alpha - 1$  and  $\theta_2 = \beta - 1$ , sufficient statistic  $T(x) = (\log(x), \log(1 - x))$ , and log-partition function  $A(\theta_1, \theta_2) = \log(B(\alpha, \beta))$ .

- (i) The family of beta distributions for which the value of  $\alpha$  is known and the value of  $\beta$  is unknown.

$$f(x; \beta) = \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha, \beta)}, \quad x \in [0, 1]$$

$$f(x; \beta) = \exp((\alpha - 1)\log(x) + (\beta - 1)\log(1 - x) - \log(B(\alpha, \beta)))$$

Hence, it is in the form of an exponential family with natural parameter  $\theta_1 = \alpha - 1$  and  $\theta_2 = \beta - 1$ , sufficient statistic  $T(x) = (\log(x), \log(1 - x))$ , and log-partition function  $A(\theta_1, \theta_2) = \log(B(\alpha, \beta))$ .

## Problem 4:

Let  $X$  be a random variable with a Student's  $t$  distribution with  $p$  degrees of freedom.

- (a) Derive the mean and variance of  $X$ .
- (b) Show that  $X^2$  has an  $F_{1,p}$  distribution.
- (c) Let  $f(x|p)$  denote the pdf of  $X$ . Show that

$$\lim_{p \rightarrow \infty} f(x|p) \rightarrow \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{x^2}{2}\right\}$$

at each value of  $x$ ,  $-\infty < x < \infty$ . This correctly suggests that as  $p \rightarrow \infty$ ,  $X$  converges in distributions to a  $N(0, 1)$  random variable (Hint: Use Stirling's Formula).

- (d) Use the results of parts (a) and (b) to argue that, as  $p \rightarrow \infty$ ,  $X$  converges in distribution to a  $X_1^2$  random variable.
- (e) What might you conjecture about the distributional limit, as  $p \rightarrow \infty$ , of  $F_{p,q}$

## Optional Problem 1:

Suppose that  $X$  has the log-normal distribution with parameters  $\mu$  and  $\sigma^2$ . Find the distribution of  $\frac{1}{X}$ . Let  $X$  follow a lognormal distribution, so:

$$X \sim \text{Lognormal}(\mu, \sigma^2) \quad \text{implies} \quad \log(X) \sim N(\mu, \sigma^2)$$

This means that:

$$X = \exp(Z) \quad \text{where} \quad Z \sim N(\mu, \sigma^2)$$

Now, we are asked to find the distribution of  $Y = \frac{1}{X}$ .

We can write  $Y$  as:

$$Y = \frac{1}{X} = \frac{1}{\exp(Z)} = \exp(-Z)$$

Since  $Z \sim N(\mu, \sigma^2)$ , the random variable  $-Z$  follows a normal distribution with parameters  $-\mu$  and  $\sigma^2$ , i.e.,

$$-Z \sim N(-\mu, \sigma^2)$$

Thus,  $Y = \exp(-Z)$  follows a lognormal distribution, and we conclude:

$$Y \sim \text{Lognormal}(-\mu, \sigma^2)$$

Therefore, the distribution of  $Y = \frac{1}{X}$  is a lognormal distribution with parameters  $-\mu$  and  $\sigma^2$ :

$$Y \sim \text{Lognormal}(-\mu, \sigma^2)$$

## Optional Problem 2:

Suppose  $\bar{X}$  is the mean of 100 observations from a population with mean  $\mu$  and variance  $\sigma^2 = 9$ . Find limits between which  $\bar{X} - \mu$  will lie with probability at least 90%. Use both Chebyshev's inequality and the Central Limit Theorem, and comment on each.