

1 All tests are imperfect: proper interpretation makes them
2 useful

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10 **Abstract**

11 Commentary is a very flexible format; Commentaries may be on policy, science and society
12 or purely scientific issues. The main criteria are that they should be of immediate interest
13 to a broad readership and should be written in an accessible, non-technical style. Their
14 length is typically 1-4 pages, although some may be longer. Because the content is variable,
15 the format is also flexible. Commentaries do not normally contain primary research data,
16 although they may present 'sociological' data (funding trends, demographics, bibliographic
17 data, etc.). As a guideline, Commentaries allow up to 25 references, and article titles are
18 omitted from the reference list.

19 Although all tests are most likely imperfect, proper interpretation of the test
20 result is the key to make imperfect tests useful. Proper interpretation starts with an
21 understanding of the purpose of the test.

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22 1. Introduction

23 In scientific research, as well as in routine daily decision-making, we depend on
24 results of various tests. Tests come in all forms and shapes. For example, doctors may
25 test a patient's blood for the presence of a disease marker, environmental engineers
26 may test water at a swimming beach for the presence of fecal coliform bacteria
27 as an indicator of sewage pollution, ecologists may survey for invasive species, a
28 geologist may drill a test well exploring for oil, a pollster may take an opinion poll
29 to evaluate the viability of a political candidate, and so on. A unifying feature of
30 these tests is that they are imperfect: the test result is likely correct but not always.
31 In most cases, we can simplify a test result to be either positive or negative. For
32 example, a positive blood test indicates the existence of the disease marker and a
33 positive opinion poll result indicate that the candidate is likely to win (with more
34 than 50% popular support). A test result is imperfect because a test is always based
35 on a sample (e.g., a blood sample, a water sample, or a sample of 1000 potential
36 voters). A sample can misrepresent the population. A water sample from a polluted
37 water source may contain no fecal coliform bacteria by chance, thereby leading to
38 a false negative result. Likewise, a water sample from a clean water source may
39 be contaminated unintentionally by researchers during the sampling process or in
40 the lab. The subsequent positive result is then classified as a false positive (i.e., the
41 water source is incorrectly classified as "polluted"). Almost all tests used in scientific
42 research are imperfect and therefore, such sampling and experimental errors are
43 unavoidable. Given this imperfection of tests, we face a challenge when interpreting
44 the test result: how can we use a potentially incorrect result to draw inference or
45 make a decision? Is a positive fecal coliform test result truly indicating that the
46 water is polluted by domestic sewage, or is it a false positive? How we interpret the
47 result will affect how we decide, for example, when to issue public health warning
48 for an affected recreational beach.

49 The imperfection of a test is routinely quantified with two error rates: the rate of

50 a false positive and the rate of a false negative. Scientists in all fields recognize the
 51 imperfection, and used various terms to describe the two types errors. For example,
 52 in the early days of World War II, the US Army developed the receiver operat-
 53 ing characteristic (ROC) curve to identify the optimal threshold when determining
 54 whether a radar signal was from a Japanese aircraft. The two axes of the ROC curve
 55 are errors of omission (false negative) and commission (false positive). For simplify-
 56 ing the discussion, we will define the following terms. First, we use present or absent
 57 to represent the state of the world we are trying to infer: a present indicates the
 58 presence of an agent of interest (e.g., fecal coliform, more than 50% popular support,
 59 an enemy aircraft, and so on) and an absent means the absence of the agent. When
 60 conducting a test, the test result is either positive (indicating the state of the world
 61 is present) or negative (absent). A false positive rate tells us how likely a test would
 62 lead to a positive result when the state of the world is absent and a false negative rate
 63 is the likelihood of a negative result when the state of the world is present. When
 64 a test is carried out, we want to use the test result (either positive or negative) to
 65 infer the state of the world. The imperfection of the test leads to uncertainty in the
 66 subsequent interpretation and inference.

67 **2. Conditional Probability and the Bayes' Theorem**

68 Properly handling the uncertainty of the test result is the realm of probability
 69 and statistics. We use probability to quantify the uncertainty and rules of probability
 70 allow us to make inferences. Using the probability language to describe an imperfect
 71 test, the rate of a false positive is the probability of a positive test result when the
 72 underlying state of the world is absent. Likewise, the rate of a false negative is the
 73 probability of a negative test result when the state of the world is present. These two
 74 probabilities are examples of a *conditional* probability. To summarize the rules of
 75 conditional probabilities, we use “ p ” to represent the state of the world being present,
 76 “ a ” to represent absent (Figure 1(i)), “ $+$ ” to represent a positive test result, and “ $-$ ”

77 a negative result (Figure 1(ii)). A false positive probability is symbolized as $\Pr(+|a)$,
 78 and a false negative is $\Pr(-|p)$. These two conditional probabilities characterize the
 79 quality of the test.

80 When a test is carried out, we observed either a “+” or a “−”. What we want
 81 to know is how likely the state of the world is p when observing a + and how
 82 likely the state of the world is a when observing a “−”. These are also conditional
 83 probabilities: $\Pr(p|+)$ and $\Pr(a|-)$. These two conditional probabilities are the basis
 84 for interpretation and inference of imperfect tests.

85 The Bayes’ rule [1] is the probability rule connecting these two groups of condi-
 86 tional probabilities (Figure 1(c)):

$$\Pr(p|+) = \frac{\Pr(p) \Pr(+|p)}{\Pr(p) \Pr(+|p) + \Pr(a) \Pr(+|a)} \quad (1)$$

87 and

$$\Pr(a|-) = \frac{\Pr(a) \Pr(-|a)}{\Pr(a) \Pr(-|a) + \Pr(p) \Pr(-|p)} \quad (2)$$

88 We will focus on equation (1). The test result is either positive or negative,
 89 consequently, $\Pr(+|p) = 1 - \Pr(-|p)$ and $\Pr(a) = 1 - \Pr(p)$. The Bayes’ rule (eq.
 90 (1)) suggests that in addition to the false positive and false negative probabilities we
 91 also must know $\Pr(p)$ in order to calculate $\Pr(p|+)$. In statistics, $\Pr(p)$ is a marginal
 92 probability – the probability of the state of the world being present regardless of the
 93 test result (or before we carried out the test). This probability can be interpreted
 94 as, for example, the prevalence of a disease in a population or our uncertainty with
 95 regard to the state of the world before a test is carried out. For example, when seeing
 96 a new patient, a doctor would treat her as a random sample from the population
 97 and interpret $\Pr(p)$ as the prevalence of the disease in the population (Figure 1(iii)).
 98 The patient, however, would use $\Pr(p)$ to describe her uncertainty with regard to the
 99 chance of her being infected given her knowledge of her own risk factors. In Bayesian

100 statistics, we call $\Pr(p)$ the prior probability of being present.

101 Add a separate box for the rattle snake example

Box A: Example: Fungal disease in Michigan rattle snake population

102

Box B: A simple explanation of the Bayes' rule

- The Bayes rule is a rule about conditional probability. We use a simple numeric example to explain equation (1). Suppose we are using a qPCR to test for the presence of a fungal disease in a rattle snake in Michigan. The test is imperfect with a false positive rate of 5% and a false negative rate of 3%. Suppose we also know, based on similar studies on the same snake species in the Midwest, the prevalence of the disease is about 4%. In order to interpret a positive or negative test result, we can use the Bayes' rule. We can explain the Bayes' rule as a straightforward account of the expected number of true positives and false positives. Assuming there are 10,000 snakes in the population, our prior knowledge suggests that 400 snakes have the disease and 9,600 snakes do not. With a false negative probability of 3%, we expect to have 12 false negatives and 388 true positives. Likewise, among the 9,600 healthy snakes, a 5% false positive rate will result in 480 false positives. If we test all 10,000 snakes, we would expect 868 positives, and only 388 of them are truly infected. For a randomly caught snake, a positive result would suggest that the chance that the snake is infected is $388/(388+480)$ or 0.447. In other words, a snake with a positive test result is (slightly) less likely to be infected with the disease. The Bayes' rule will lead to the same result. The numerator of the Bayes rule is $\Pr(p) \Pr(+|p)$, where $\Pr(p)$ is the prevalence (4%), and $\Pr(+|p) = 1 - \Pr(-|p) = 1 - 0.03 = 0.97$. That is, the numerator is 0.0388. We can easily verify that the denominator is $0.0388 + 0.048$. Multiplying the population number with both the

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numerator and denominator, we see that the Bayes rule simply tallies the number of true positives and total positives.

- When the test is imperfect, we are uncertain about the result. The *posterior probability* $\Pr(p|+)$, the probability of present after observing a positive result, is used to summarize the uncertainty we have on the result. The uncertainty, however, is a function of the prior and the two probabilities characterizing the performance of the test. In the numerical example here, if the prior is 1%, the posterior would be $\Pr(p|+) = 0.164$. This outcome is expected as the number of true positives is now a smaller fraction of the total number of positives.

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105 3. The Prior Probability and Statistical Inference

106 A point of contention in using the Bayes rule is the meaning of the prior $\Pr(p)$. In
107 Box B, we interpreted the prior as the fraction of individual snakes in the population
108 that are infected with the fungal disease. When this fraction changes, the proportion
109 of true positives (positives from infected individuals) also changes. When the prior
110 has a clear physical meaning and can be measured, the use of the Bayes' rule is
111 widely accepted [2]. When the prior is difficult to estimate or the physical meaning
112 is ambiguous, the use of prior used to be controversial. Increasingly, we recognize
113 that estimating the prior is a means for properly use of relevant information in an
114 analysis (Box C). Regardless of the meaning of the prior, the Bayes rule highlights
115 the need of quantifying the prior $\Pr(p)$ in order to properly interpret a test result;
116 whether we call the quantity a prior, a marginal probability, or the prevalence is
117 irrelevant. The proper interpretation of the test result and the use of the test result
118 for inference requires a proper statistical treatment. As in all statistical applications,
119 the first step is to represent the scientific hypothesis using a statistical model (with
120 parameters). The proposed statistical model, in turn, will decide how we use data

Box C: The meaning of a prior

- Different people may have different prior probabilities for the same event. This is likely because they use different references. For example, when a doctor tests for a communicable disease in a public health exhibition at a county fair, her prior should be the prevalence of the disease in the population because she would consider the test subject a random sample from the population. To the patient, the general population may not be a good reference because he knows more about himself. In this case, he knows which risk factors apply to him. As a result, the relevant population would be people with the same risk factors (e.g., smokers).
- The meaning of the prior to the patient and to the doctor may be the same (e.g., prevalence of the disease in a population). But deciding which population to use to form the prior requires more information. The subjective nature of a Bayesian prior is often criticized. We contend that a prior is simply a means for a scientist to properly sort out the relevant facts/information using their knowledge of their study system. In this regard, when we define the prior probability as a degree of belief, we are really trying to make use of all our knowledge to ensure the outcome is most relevant.
- When we cannot confidently identify a sub-population for inference, we must step back to a larger population. The resulting prior is likely less relevant. As a result, the estimated posterior probability is less accurate. For example, in the numerical example in Box B, the posterior probability of present given a positive results is less than 0.5. But the test result puts the snake into a smaller population (the 388+480 would-be positive snakes). If we choose to conduct a follow-up test, our prior would be

the posterior from the first test. This iterative process is an appealing characteristic of the Bayesian method for many applied scientists.

4. The Purpose of A Test

Understanding and defining the objective of a test is the key to deciding what statistical model to use. When analyzing results from an imperfect test, there are at least two different objectives: (1) testing for an individual subject, that is, to estimate $\Pr(p|+)$ (e.g., whether a patient has a particular disease agent), and (2) testing to learn about a population, that is, to estimate $\Pr(p)$ (e.g., the prevalence of fungal disease in the rattlesnake population in Michigan).

4.1. Tests for Individual Subjects

When a doctor tests a patient for a disease, the objective is to determine the likelihood that the patient has the disease. With known characteristics of the test (i.e., $\Pr(+|a)$ and $\Pr(-|p)$) and the prevalence of the disease in the population, we can use the Bayes' rule to calculate the conditional probability $\Pr(p|+)$ for a positive result. In this situation, we want a test that will result in a $\Pr(p|+)$ larger than 0.5 ($\Pr(p|+) > 0.5$). That is, a positive result should suggest that the patient is more likely to have the disease than not. Another way to express this condition is that the odds ratio should be larger than 1: $\frac{\Pr(p|+)}{1-\Pr(p|+)} > 1$. More generally, we can require that $\frac{\Pr(p|+)}{1-\Pr(p|+)} > \gamma$ before we consider the test to be useful (e.g., prescribe treatment upon a positive result). Using the Bayes' rule, we can express this requirement in terms of the rates of false positive and false negative, as well as the prevalence.

$$\Pr(+|a) < \frac{1-\gamma}{\gamma} \frac{\Pr(p)}{1-\Pr(p)} (1 - \Pr(-|p)) \quad (3)$$

In other words, the inequality set by equation 3 must be met before a test can be used.

145 4.2. Tests for Estimating Population Parameter(s)

146 In many cases, we test multiple individuals of a population in order to understand
 147 the characteristics of the population. In the rattlesnake example (Box A), we are
 148 interested in estimating the prevalence of the fungal disease in the population. That
 149 is, we want to estimate the prior $\theta = \Pr(p)$. There are two approaches to estimate θ .
 150 One is using inverse Bayes formulae (IBF) [5], which is a simple and straightforward
 151 method for situations where the true state of the world of the sampled individuals
 152 can be ascertained (perhaps with additional experimentation). Specifically, once we
 153 know the true state of the world of the samples, we can use the test results to estimate
 154 the posterior directly from the data (Box D). That is, we have $\Pr(p|+)$, $\Pr(+|a)$, and
 155 $\Pr(-|p)$, we can calculate $\Pr(p)$ using IBF.

Box D: Inverse Bayes Formulae (IBF)

IBF can be derived directly using the Bayes' rule or, more parsimoniously, using the joint distribution formula:

$$\Pr(X, Y) = \Pr(Y|X) \Pr(X) = \Pr(X|Y) \Pr(Y)$$

which yields

$$\Pr(Y) = \frac{\Pr(Y|X)}{\Pr(X|Y)} \Pr(X)$$

We set X to be the variable of the state of the world (taking values p or a) and Y the variable of test results ($+$ or $-$). Because $\Pr(Y = +) + \Pr(Y = -) = 1$, $\Pr(Y = -) = \frac{\Pr(Y=-|X)}{\Pr(X|Y=-)} \Pr(X)$, and $\Pr(Y = +) = \frac{\Pr(Y=+|X)}{\Pr(X|Y=+)} \Pr(X)$, we have

$$\frac{\Pr(Y = +|X)}{\Pr(X|Y = +)} \Pr(X) + \frac{\Pr(Y = -|X)}{\Pr(X|Y = -)} \Pr(X) = 1 \quad (4)$$

From equation (4), we can express the quantity of interest $\Pr(X = p)$ as:

$$\Pr(X = p) = \frac{1}{\frac{\Pr(Y=+|X=p)}{\Pr(X=p|Y=+)} + \frac{\Pr(Y=-|X=p)}{\Pr(X=p|Y=-)}}$$

or, using the notation in Box B:

$$\Pr(p) = \frac{1}{\frac{\Pr(+|p)}{\Pr(p|+)} + \frac{\Pr(-|p)}{\Pr(p|-)}} \quad (5)$$

Assuming that the false negative probability ($\Pr(-|p)$) is a known quantity characterizing the test, we then know two of the four probabilities ($\Pr(-|p)$ and $\Pr(+|p) = 1 - \Pr(-|p)$). The other two probabilities ($\Pr(p|-)$, $\Pr(p|+)$) can be estimated if the true state of the world of an individual test subject can be ascertained. Suppose that we test N individuals and record the test result. In the follow up test we determine the true status of these individuals and record the outcome in a two by two table:

	$X = p$	$X = a$	total
$y = +$	n_{11}	n_{12}	N_1
$y = -$	n_{21}	n_{22}	N_2
total	M_1	M_2	N

From the above table, we have $\Pr(p|-) = n_{21}/N_2$ and $\Pr(p|+) = n_{11}/N_1$. In the rattlesnake example, we can ascertain the status of the tested snakes by keeping them in the lab for a period of time to observe which ones eventually succumb to the fatal disease.

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158 When the true state of the world cannot be ascertained, we need to approach
 159 the problem from a different angle. The objective of the study is now the estimation
 160 of $\theta = \Pr(p)$, a continuous variable, based on observed positives and negatives. In
 161 statistical terms, we tested n snakes and observed y positives, from which, we wish

162 to estimate the prevalence θ . We start the process by proposing a statistical model
163 describing the data generating process. In the case of analyzing testing results, the
164 data are the number of positive results from a total number of subjects. The statis-
165 tical model describes the distribution of the data is the binomial distribution. The
166 model is parameterized by a single parameter – the probability of observing a positive
167 result. The quantity of interest is the probability of infection. How the parameter
168 of interest and the binomial model parameter are linked depends on what we know
169 (Box E). If the test is perfect, that is, we know the rates of false positive and false
170 negative are both 0, we have a simple binomial-beta model and the parameter can
171 be easily estimated. This model is often the first model in an introductory Bayesian
172 statistics book (e.g., McElreath [4]). When the complexity of the data generation
173 process increases, the simple model needs to be modified. If the test is imperfect
174 and rates of false positive and false negative are known, the posterior distribution
175 of θ cannot be represented by a commonly seen probability distribution. But the
176 posterior distribution can be numerically evaluated and graphed for inference. When
177 one of the two error rates is unknown and need to be estimated, we must estimate
178 a two-dimensional joint probability distribution. Computation is more intense, al-
179 though we can still graphically display the joint distribution. When both rates are
180 unknown, the posterior distribution becomes a three-dimensional joint distribution,
181 and we must resort to modern Bayesian computation method to estimate the param-
182 eters and numerical summary of these parameters to understand the distribution.

183 The steps of moving from the simplest binomial-beta model to the model simul-
184 taneously estimating three parameters are typical in statistical learning. These steps
185 should also be taken in applying statistics. When using imperfect tests, we need to
186 understand why and how the test is imperfect and derive the appropriate model ac-
187 cordingly. Such a process is often tedious and iterative, a message we almost always
188 miss when teaching or taking a statistics course.

Box E: Bayesian inference

- Using Bayesian statistics to estimate an unknown variable consists of three steps.
 1. Propose a statistical model that describes the data. This model includes the unknown variable as a parameter. For example, the statistical model we use to estimate the prevalence of fungal disease in the Michigan rattlesnake population is a binomial model. The test result is either positive or negative and the probability of a positive snake is the prevalence.
 2. Using the statistical model, we derive the likelihood function of the data – the likelihood of observing the data if the proposed model is correct. If the test is perfect, that is, both false positive and false negative rates are 0, the prevalence is the probability of observing a positive result. The statistical model of observing y positive in n snakes is

$$\Pr(y|\theta) = \binom{n}{y} \theta^y (1 - \theta)^{n-y} \quad (6)$$

This is the likelihood of observing the data if the prevalence is θ . In classical statistics, the estimated parameter $\hat{\theta}$ is the value that maximizes the likelihood.

3. Estimating the posterior distribution of θ using the Bayes rule of a continuous variable:

$$\pi(\theta|y) = \frac{\pi(\theta) \Pr(y|\theta)}{\int_{\theta} \pi(\theta) \Pr(y|\theta) d\theta} \quad (7)$$

where $\pi()$ represent a probability density function. $\pi(\theta)$ is the prior distribution, representing the uncertainty we have about θ

and $\pi(\theta|y)$ is the posterior distribution of θ after observing data.

- The derivation of the posterior distribution is often the difficult part of Bayesian inference because of the integral in the denominator. In some cases, the derivation can be simplified if we choose a suitable prior distribution. For example, when we choose the beta distribution as the prior for θ , the prior density (with parameters α and β) is $\pi(\theta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}\theta^{\alpha-1}(1-\theta)^{\beta-1}$, the posterior is

$$\begin{aligned}\pi(\theta|y) &= \frac{\frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}\theta^{\alpha-1}(1-\theta)^{\beta-1}\binom{n}{y}\theta^y(1-\theta)^{n-y}}{\int_{\theta} \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}\theta^{\alpha-1}(1-\theta)^{\beta-1}\binom{n}{y}\theta^y(1-\theta)^{n-y}d\theta} \\ &\propto \theta^{y+\alpha-1}(1-\theta)^{(n-y)+\beta-1}\end{aligned}\quad (8)$$

The posterior is the beta distribution with parameters $y + \alpha$ and $(n - y) + \beta$. The mean of the posterior is $\frac{y+\alpha}{n+\alpha+\beta}$. From this distribution, we can also calculate the credible interval to summarize the uncertainty.

- When the test is imperfect and we know the rate of false positive (f_p) and rate of false negative (f_n), the model becomes more complicated because the probability of observing a positive result is now $\theta^* = \theta(1 - f_n) + (1 - \theta)f_p$. The statistical model of the data is still the binomial distribution, but the probability of observing a positive result is now θ^* . As a result, the posterior distribution is:

$$\begin{aligned}\pi(\theta|y) &\propto \theta^{\alpha-1}(1-\theta)^{n-y} \times \\ &\quad (\theta(1 - f_n) + (1 - \theta)f_p)^y \times \\ &\quad (1 - \theta(1 - f_n) - (1 - \theta)f_p)^{n-y}\end{aligned}\quad (9)$$

Box F: Bayesian computation

- In Box E we derived two posterior distributions of the parameter θ . One is summarized by a standard probability distribution, of which we know how to derive needed statistics to summarize the uncertainty about the parameter. The other is an algebraic expression that cannot be represented by a known form of probability distribution. Because θ is the only unknown parameter, we can graphically draw the posterior distribution on a two dimensional space. By numerically rescaling the curve such that the area under the curve is 1, we have a graphical representation of the probability distribution (Figure 2), from which we can draw inference about the parameter. When one of the two error rates (e.g., the false positive rate f_p) is unknown, the problem is to estimate the joint distribution of the two unknowns. We can still use the same numerical approximation method and graphically present the posterior using a contour plot (Figure 3). The same computational framework can be used if both both error rates are unknown. But graphical display of the distribution is no longer feasible. Frequently, we use the modern Bayesian computational method for this type of problem. Computational details of this paper can be found at <https://github.com/songsqian/imperfect>.

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192 5. Practical Implications

193 Although our discussions were made with specific examples, the problem of an
194 imperfect test is universal. How we interpret results from an imperfect test depends
195 on how the data were collected and for what purpose. When individual test subjects
196 are of concern, the uncertainty should be presented in terms of a conditional prob-
197 ability (i.e., $\Pr(p|+)$). The quality of the test is the key for proper interpretation of
198 the test result. This is not just a problem of medical diagnostic tests. The statis-

199 tical principle is the same in any situation where results from an imperfect test are
 200 applied to a specific subject. When an imperfect test is used to infer a population
 201 parameter such as the prevalence of fungal disease in the rattlesnake population in
 202 Michigan, test results are raw data to be further processed to estimate the parameter
 203 of interest. The uncertainty associated with the imperfect test is represented by the
 204 posterior probability distribution. Depending on the nature of the test and the ease
 205 of determining the true status of a test subject, we have different computational
 206 needs. In other words, the proper use of an imperfect test requires us to fully con-
 207 sider all available information and properly structure the statistical analysis based
 208 on the objective of the study, just as the interpretation of the p -value from a null
 209 hypothesis test with low power [3] should be case-specific.

210 Summary Table

Test Target	An individual subject	A population
Test Objective	$\Pr(p +)$ Does the subject have the disease?	$\Pr(p)$ The disease prevalence in the population
Test Result	+/-	+/- (multiple subjects)
211 Knowledge needed to interpret test results	$\Pr(p)$, quality of the test $\Pr(+ a)$, $\Pr(- p)$	$\Pr(+ a)$, $\Pr(- p)$, and an educated guess of $\Pr(p)$
Additional steps (beyond test results)	Bayes rule	IBF if true state of each subject can be ascertained, or Bayesian computation when true state is infeasible
How can the test be improved	improve test quality	... and a better guess of $\Pr(p)$

References

- [1] M.H. DeGroot. *Probability and Statistics*. Addison-Wesley Publishing Company, Inc., second edition, 1986.
- [2] B. Efron and C. Morris. Stein's paradox in statistics. *Scientific American*, 236: 119–127, 1977.
- [3] L.G. Halsey, G. Curran-Everett, S.L. Vowler, and G.B. Drummond. The fickle P value generates irreproducible results. *Nature Methods*, 12:179–185, 2015.
- [4] R. McElreath. *Statistical Rethinking: A Bayesian Course with Examples in R and Stan*. Chapman & Hall/CRC Press, 2016.
- [5] M.T. Tan, G.L. Tian, and K.W. Ng. *Bayesian Missing Data Problems: EM, Data Augmentation and Noniterative Computation*. Chapman and Hall/CRC Press, Boca Raton, FL, USA, 2010.

Figures

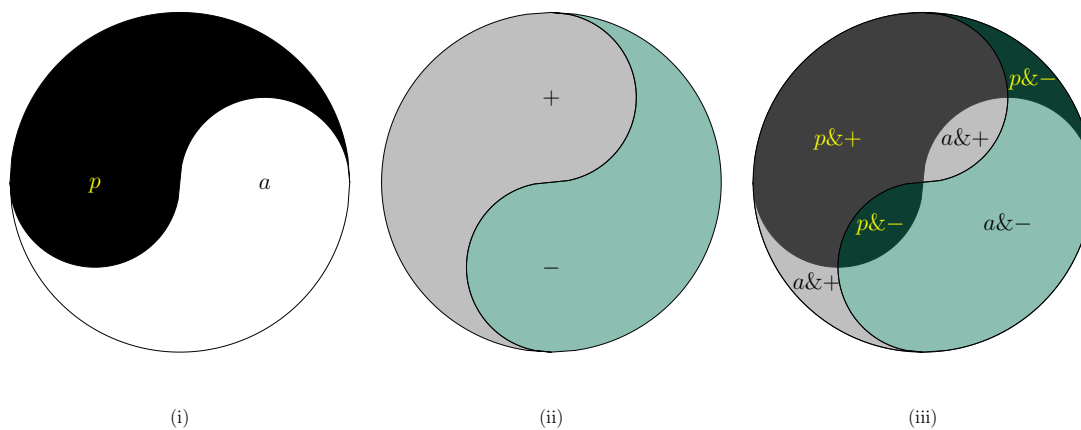


Figure 1: A graphical depiction of an imperfect test. The real state of the world is either p (present) or a (absent) (i), and the test result is either $+$ or $-$ (ii). The imperfection of the test makes the interpretation of a test result contingent on information regarding the accuracy of the test, as well as the relative size of p and a .

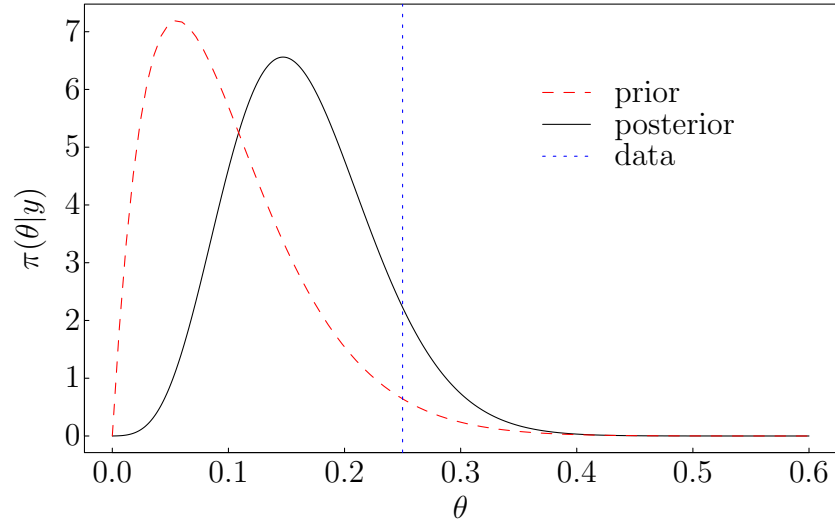


Figure 2: Numerically estimated probability density function of the posterior distribution of θ

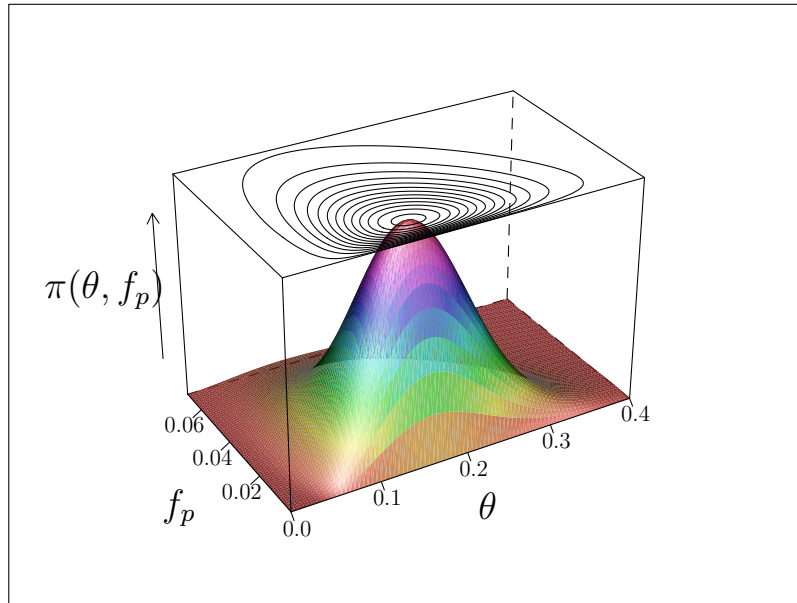


Figure 3: A perspective plot of the numerically estimated joint distribution of f_p and θ

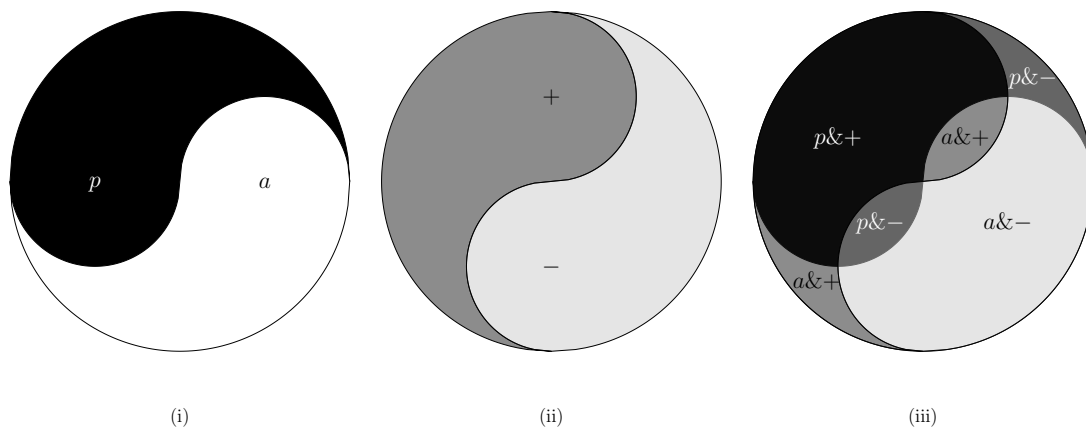


Figure 4: A graphical depiction of an imperfect test. The real state of the world is either p (present) or a (absent) (i), and the test result is either $+$ or $-$ (ii). The imperfection of the test makes the interpretation of a test result contingent on information regarding the accuracy of the test, as well as the relative size of p and a .

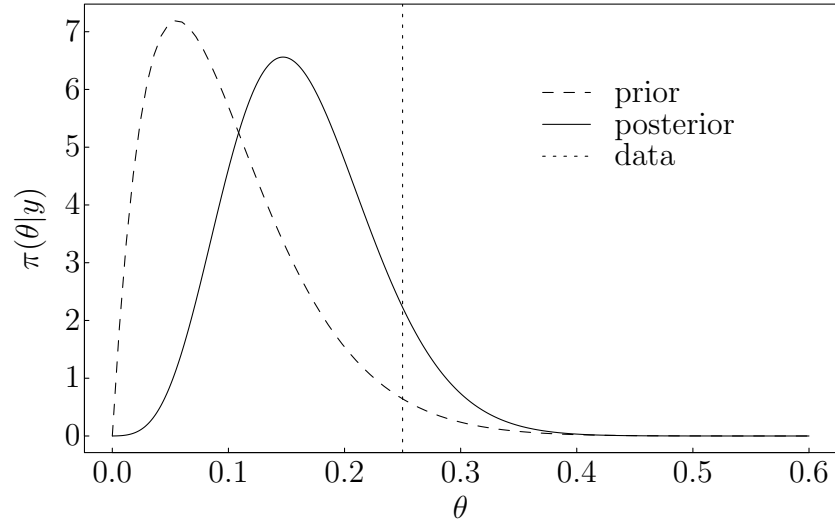


Figure 5: Numerically estimated probability density function of the posterior distribution of θ

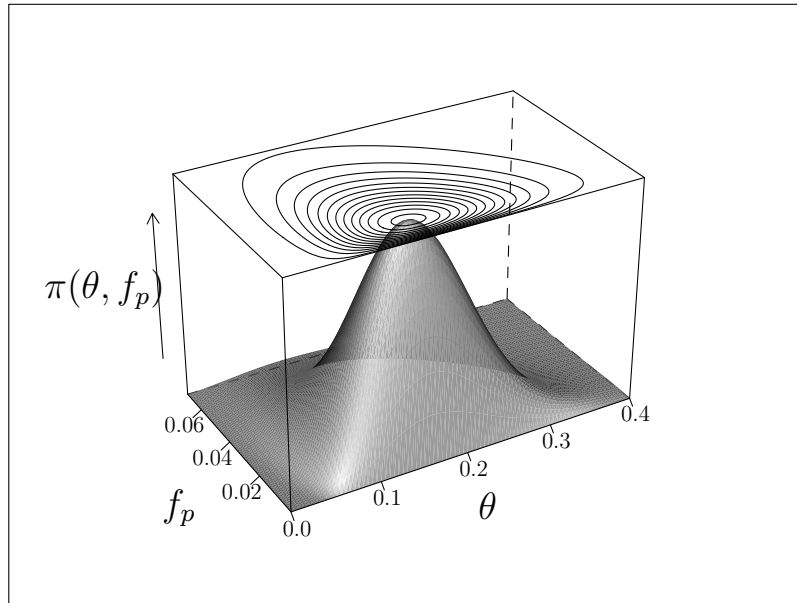


Figure 6: A perspective plot of the numerically estimated joint distribution of f_p and θ