

# HW4, MATH 423

Owen Queen

September 28, 2020

## 1 Example 3.8.2

Suppose that  $X$  and  $Y$  are two independent binomial random variables, each with the same success probability but defined on  $m$  and  $n$  trials, respectively. i.e.:

$$p_X(k) = \binom{m}{k} p^k (1-p)^{m-k}, k = 0, 1, \dots, m$$

and

$$p_Y(k) = \binom{n}{k} p^k (1-p)^{n-k}, k = 0, 1, \dots, n$$

We are to find  $p_W(w)$  where  $W = X + Y$ .

By Theorem 3.8.3, we know that  $p_W(w) = \sum_x p_X(x) p_Y(w-x)$ . The binomial probability distribution is defined to be 0 when  $x < 0$ . Therefore, we only want to sum values where  $p_X(x)$  and  $p_Y(w-x)$  are nonzero, i.e. where  $(w-x) \geq 0$ . This occurs when we sum over the values  $x \in \{0, 1, \dots, w\}$ . Thus:

$$\begin{aligned} p_W(w) &= \sum_{x=0}^w p_X(x) p_Y(w-x) = \sum_{x=0}^w \binom{m}{x} p^x (1-p)^{m-x} \binom{n}{w-x} p^{w-x} (1-p)^{n-(w-x)} \\ &= \sum_{x=0}^w \binom{m}{x} \binom{n}{w-x} p^w (1-p)^{n+m-w} \end{aligned}$$

Consider an urn having  $m$  red chips and  $n$  white chips. If  $w$  chips are drawn without replacement, then the probability that exactly  $x$  red chips are drawn is given by the hypergeometric distribution:

$$P(x) = \frac{\binom{m}{x} \binom{n}{w-x}}{\binom{m+n}{w}}$$

Since this is a probability density function (pdf), we know:

$$\sum_{x=0}^w \frac{\binom{m}{x} \binom{n}{w-x}}{\binom{m+n}{w}} = 1$$

Thus,

$$\sum_{x=0}^w \binom{m}{x} \binom{n}{w-x} = \binom{m+n}{w}$$

Plugging this into our  $p_W(w)$  equation, we get:

$$p_W(w) = \binom{m+n}{w} p^w (1-p)^{n+m-w}$$

This is a binomial distribution with parameters  $m+n$  and  $p$ .

## 2 Theorem 3.8.4

**Thm. 3.8.4:** Let  $X$  and  $Y$  be independent continuous random variables, with pdf's  $f_X(x)$  and  $f_Y(y)$ , respectively. Assume that  $X$  is zero for at most a set of isolated points. Let  $W = X/Y$ . Then:

$$f_W(w) = \int_{-\infty}^{\infty} |x| f_X(x) f_Y(wx) dx$$

**Proof:**

$$F_W(w) = P(Y/X \leq w)$$

$$= P(Y/X \leq w \cap X \geq 0) + P(Y/X \geq w \cap X < 0)$$

$$= P(Y \leq wX \cap X \geq 0) + P(Y \geq wX \cap X < 0)$$

because

$$(Y \geq wX \cap X < 0)^c = (Y \leq wX \cap X < 0)$$

$$P((Y \geq wX \cap X < 0)^c) = 1 - P(Y \leq wX \cap X < 0)$$

Thus,

$$F_W(w) = P(Y \leq wX \cap X \geq 0) + 1 - P(Y \leq wX \cap X < 0)$$

By definition of cumulative distribution function:

$$F_W(w) = \int_0^{\infty} \int_{-\infty}^{wx} f_X(x) f_Y(y) dy dx + 1 - \int_{-\infty}^0 \int_{-\infty}^{wx} f_X(x) f_Y(y) dy dx$$

Differentiate:  $\frac{d}{dw} F_W(w) = f_W(w)$

$$f_W(w) = \frac{d}{dw} \int_0^{\infty} \int_{-\infty}^{wx} f_X(x) f_Y(y) dy dx - \frac{d}{dw} \int_{-\infty}^0 \int_{-\infty}^{wx} f_X(x) f_Y(y) dy dx$$

$$= \int_0^\infty f_X(x) \left[ \frac{d}{dw} \int_{-\infty}^{wx} f_Y(y) dy \right] dx - \int_{-\infty}^0 f_X(x) \left[ \frac{d}{dw} \int_{-\infty}^{wx} f_Y(y) dy \right] dx$$

By the Fundamental Theorem of Calculus, we have that:

$$\frac{d}{dw} F_W(w) = \frac{d}{dw} \int_{-\infty}^{wx} f_Y(y) dy = f_Y(wx) \frac{d}{dw} wx = x f_Y(wx)$$

$$\begin{aligned} f_W(w) &= \int_0^\infty x f_X(x) f_Y(wx) dx + \int_{-\infty}^0 (-x) f_X(x) f_Y(wx) dx \\ &= \int_0^\infty |x| f_X(x) f_Y(wx) dx + \int_{-\infty}^0 |x| f_X(x) f_Y(wx) dx \end{aligned}$$

Thus,

$$f_W(w) = \int_{-\infty}^\infty |x| f_X(x) f_Y(wx) dx \blacksquare$$