

# Moment Generating Func<sup>n</sup> (MGF) of $X$

8/3  
①

$$\begin{aligned}\hat{\Phi}_X(s) &= E[e^{sX}] \\ &= \int_{-\infty}^{\infty} e^{sx} f_X(x) dx\end{aligned}$$

eg:-  $X$  is exponential ( $\lambda$ )

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{o/w} \end{cases}$$

$$\begin{aligned}\hat{\Phi}_X(s) &= \int_0^{\infty} \lambda e^{-\lambda x} e^{sx} dx \\ &= \lambda \int_0^{\infty} e^{-(\lambda-s)x} dx\end{aligned}$$

$$= \lambda \left. \frac{e^{-(\lambda-s)x}}{-(\lambda-s)} \right|_0^\infty$$

②

$$\Phi_X(s) = \frac{\lambda}{\lambda - s}$$

$X_1$  &  $X_2$  are independent

$$X = X_1 + X_2$$

pdf of  $X$ ?

MGF of  $X$

$$\Phi_X(s) = E[e^{sX}]$$

$$= E[e^{s(X_1 + X_2)}]$$

$$= E[e^{sX_1} e^{sX_2}]$$

$$= E[e^{sX_1}] E[e^{sX_2}]$$

$X_1$  &  $X_2$   
are independent

(3)

$$\Phi_X(s) = \Phi_{X_1}(s) \Phi_{X_2}(s)$$

MGF of the sum of independent RVs is the product of the individual MGFs

eg:-  $X_1$  is exponential ( $\lambda_1$ )  
 $X_2$  " " ( $\lambda_2$ )

$X_1, X_2$  are independent

Find the pdf of  $X = X_1 + X_2$

$$\Phi_X(s) = \Phi_{X_1}(s) \Phi_{X_2}(s)$$

$$= \frac{\lambda_1}{\lambda_1 - s} \cdot \frac{\lambda_2}{\lambda_2 - s}$$

$$= \frac{A \lambda_1}{\lambda_1 - s} + \frac{B}{\lambda_2 - s}$$

partial fractions

A & B are constants  
 $\rightarrow (\lambda_1 \neq \lambda_2)$

$$f_X(x) = \begin{cases} \frac{A}{\lambda_1} \cdot \lambda_1 e^{-\lambda_1 x} + \frac{B}{\lambda_2} \cdot \lambda_2 e^{-\lambda_2 x} & x \geq 0 \\ 0, \text{ o/w.} \end{cases} \quad (4)$$

EXAM

Q① Two RVs

↓

Marginals, Conditionals, Expected values

Correlation, Covariance, correlation coeff.

Jointly Gaussian.

~~Q②~~

② Central Limit Thm.

③ Estimation → Linear Regression

④ Confidence Interval calculations

eg:- Joint pdf of  $x$  &  $y$  is

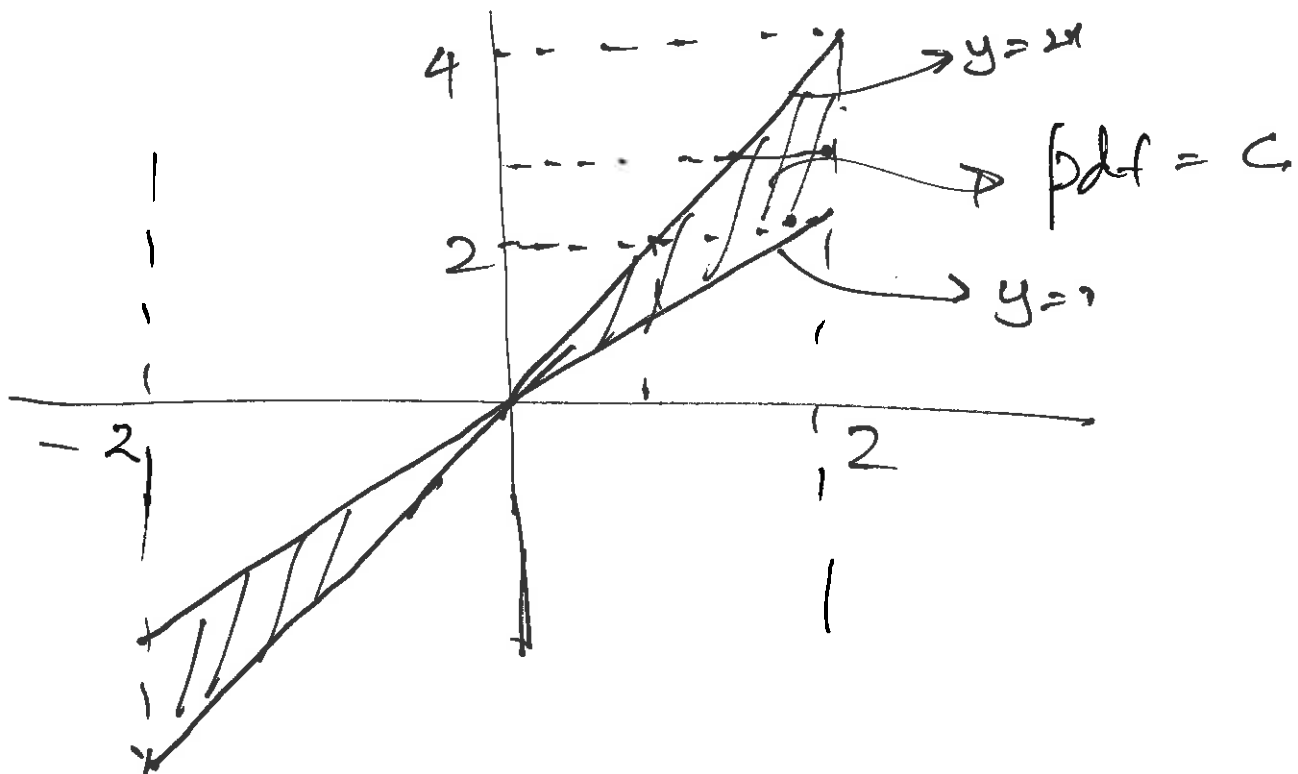
$$f_{x,y}(x,y) = \begin{cases} C, & -2 \leq x \leq 2, \quad x \leq y \leq 2x \\ & y \text{ lies between } x \text{ & } 2x \\ 0, & \text{o.w.} \end{cases}$$

Find  $C$

Are  $x$  &  $y$  uncorrelated

"  $x$  &  $y$  independent  
linearly

Estimate  $x$  in terms of  $y \rightarrow$  Regression line?



$$C[\text{Area}] = 1$$

⑥

$$\rightarrow 2 \left[ \frac{1}{2} (2) (4) - \frac{1}{2} (2) (2) \right]$$

$$C[8-4] = 1 \rightarrow C = \frac{1}{4}$$

check the independence

$$1. f_{X,Y}(x,y) = f_X(x) f_Y(y)$$

$$2. f_{X|Y}(x|y) = f_X(x)$$

$$3. f_{Y|X}(y|x) = f_Y(y)$$

varies  
from  
-2 to 2

$$\rightarrow f_{X|Y}(x|2) \rightarrow \text{varies from } 1 \text{ to } 2$$

$f_{X|Y}(x|y)$  depends on  $y$

$\therefore X$  &  $Y$  are not independent

Uncorrelated?

⑦

Need  $\text{Cov}[X, Y]$

$$\text{Cov}[X, Y] = E[XY] - \mu_x \mu_y$$

$$E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \cdot f_{X,Y}(x, y) dy dx$$

$$= 2 \int_0^2 \int_0^{2x} xy \cdot c dy dx$$

$$= 2c \int_0^2 x \left. \frac{y^2}{2} \right|_0^{2x} dx$$

$$= \frac{2c}{2} \int_0^2 x (4x^2 - x^2) dx$$

$$= c \left. \frac{3x^4}{4} \right|_0^2 =$$

$$\frac{3c}{4} (4) = \frac{12c}{4} = 3c$$

(8)

$$\mu_x = \int_0^2 \int_x^{2x} x \cdot c \, dy \, dx + \int_{-2}^0 \int_{2x}^x x \cdot c \, dy \, dx$$

$$= 0.$$

$f_x(a)$

$$\text{Cov}[X, Y] = E[XY] - \mu_x \mu_y$$

$$= 2c \neq 0$$

$X$  &  $Y$  are correlated.

$$[x_2(y) - \mu_x] = \left( \frac{\rho_{x,y} \sigma_x}{\sigma_y} \right) [y - \mu_y]$$

Find  $\mu_y, \sigma_x, \sigma_y$  →  $\rho_{x,y}$

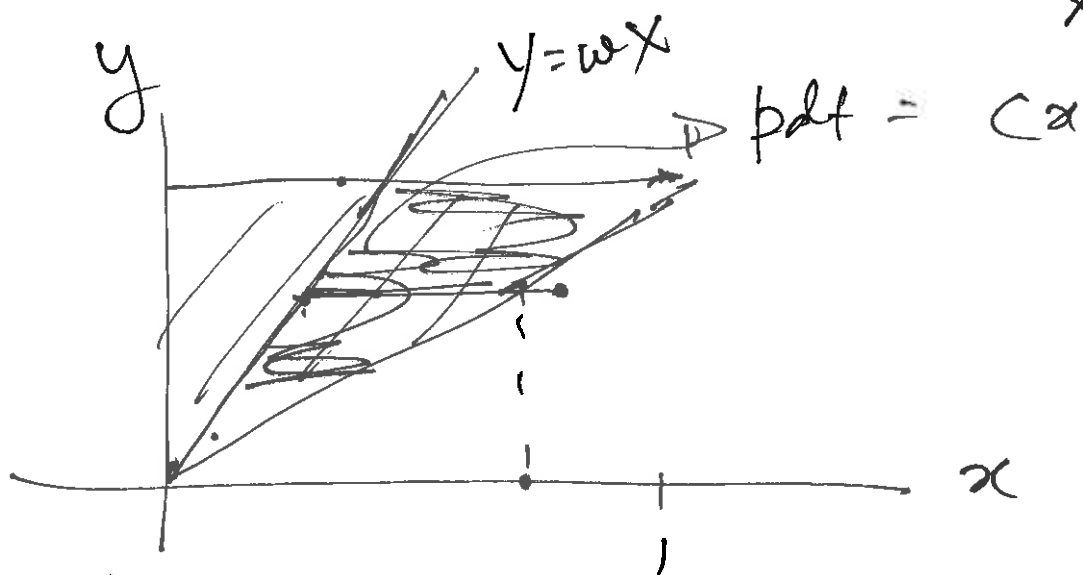
$$\frac{\text{Cov}[X, Y]}{\text{Var}[Y]}$$



eg:-

$$f_{X,Y}(x,y) = \begin{cases} cx, & 0 \leq x \leq y \leq 1 \\ 0, & \text{o/w.} \end{cases} \quad (9)$$

Find  $c$  & the pdf of  $\frac{Y}{X}$



$$\int_0^1 \int_x^1 cx \, dy \, dx = 1$$

$$c \int_0^1 x(1-x) \, dx = 1$$

$$c \left[ \frac{x^2}{2} - \frac{x^3}{3} \right]_0^1 = 1 \rightarrow$$

$$c \left[ \frac{1}{2} - \frac{1}{3} \right] = 1$$

$c = \checkmark$

$$W = \frac{Y}{X} \rightarrow W \text{ varies from } 1 \text{ to } \infty \quad (10)$$

$$\text{CDF of } W, \quad F_W(w) = P[W \leq w]$$

$$= P\left[\frac{Y}{X} \leq w\right]$$

$$\underline{w \geq 1}$$

$$= P[Y \leq wX]$$

$$= \iint_R f_{X,Y}(x,y) dy dx$$

$$= \int_0^1 \int_{\frac{y}{w}}^y cx dx dy$$

$$= c \int_0^1 \left. \frac{x^2}{2} \right|_{\frac{y}{w}}^y dy$$

$$f_W(w) =$$

(11)

$$F_W(\omega) = \frac{C}{2} \int_0^1 \left( y^2 - \frac{y^2}{\omega^2} \right) dy$$

$$= \frac{C}{2} \left[ \frac{y^3}{3} - \frac{y^3}{3\omega^2} \right]_0^1$$

$$F_W(\omega) = \frac{C}{6} \left[ 1 - \frac{1}{\omega^2} \right]$$

$$F_W(\omega) = \begin{cases} 0, & \omega < 1 \\ \frac{C}{6} \left[ 1 - \frac{1}{\omega^2} \right] & \omega \geq 1 \end{cases}$$

$$f_W(\omega) = \begin{cases} \frac{C}{6} \left[ \frac{2}{\omega^3} \right], & \omega \geq 1 \\ 0, & \text{ok.} \end{cases}$$

Ex:  $X$  &  $Y$  are jointly Gaussian (12)  
with joint pdf

$$N(2, 3; 1, 2; 0)$$

Find  $P[|X| > 3 \text{ \& } Y < 2]$

$X$  &  $Y$  are uncorrelated &  
Jointly Gaussian

$\therefore X$  &  $Y$  are independent

$$= P[|X| > 3] \cdot P[Y < 2]$$