

UiO: Matematisk institutt

Det matematisk-naturvitenskapelige fakultet

STK-4051/9051 Computational Statistics Spring 2022 Chaper 2

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Course originally made by Geir Storvik



Optimization

Focus maximum likelihood

But methods are general

 $\max_{\theta} L(\theta | \mathbf{y})$

- Different settings
 - Continuous vs discrete
 - One vs multi-dimensional
 - Unconstrained vs constrained
 - Common: $y \sim N(\mu, \sigma^2)$, μ unconstrained, $\sigma^2 > 0$
- Can we compute the derivative analytically?

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One dimensional ML, Newton's method

Common to consider log likelihood:

| Log | Common to consider log likelihood: | Log | Common that pec this is what they use they use they use the consider log likelihood: | Color)

$$- \underset{\theta}{\operatorname{argmax}} L(\theta | \mathbf{y}) = \underset{\theta}{\operatorname{argmax}} \underbrace{\log(L(\theta | \mathbf{y}))}_{\theta(\theta | \mathbf{y})}$$

 $^{ert}\!\ell(hetaertoldsymbol{
u})$ or just $\;\ell(heta)$

$$\ell(\theta) \approx \ell(\theta^*) + (\theta - \theta^*)\ell'(\theta^*) + \frac{1}{2}(\theta - \theta^*)^2\ell''(\theta^*)$$
$$\ell(\theta^*) + (\theta - \theta^*)s(\theta^*) - \frac{1}{2}(\theta - \theta^*)^2J(\theta^*)$$

Taylor expansion around θ^*

Score function:

 $s(\theta) = \ell'(\theta)$

Observed information: $I(\theta) = -\ell''(\theta)$

— Solving the maximum of the approximation:

$$\theta = \theta^* + \frac{s(\theta^*)}{J(\theta^*)} = \theta^* - \frac{\ell'(\theta^*)}{\ell''(\theta^*)}$$

Example \mathbb{R}

$$L(\mu) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2} \left(\frac{x_i - \mu}{\sigma}\right)^2\} \qquad \sigma \text{ known}$$

$$l(\mu) = \sum_{i=1}^{n} -\frac{1}{2} \log 2\pi - \frac{1}{2} \log \sigma^{2} - \frac{1}{2} \left(\frac{x_{i} - \mu}{\sigma} \right)^{2}$$

$$s(\mu) = l'(\mu) = \sum_{i=1}^{n} -0 - 0 - \frac{x_i - \mu}{\sigma^2} = \frac{1}{\sigma^2} \sum_{i=1}^{n} (x_i - \mu)$$

$$J(\mu) = -l''(\mu) = -s'(\mu) = \frac{-1}{\sigma^2} \sum_{i=1}^{n} -1 = \frac{n}{\sigma^2}$$

Iterations in Newton's method

•
$$\theta^{(t+1)} = \theta^{(t)} + \frac{s(\theta^{(t)})}{J(\theta^{(t)})}$$
 $\ell(\theta^1) + (\theta - \theta^1)s(\theta^1) - \cdots$

$$\ell(\theta) = \frac{r_0}{65}$$

$$\ell(\theta^0) + (\theta - \theta^0)s(\theta^0) - \frac{1}{2}(\theta - \theta^0)^2 J(\theta^0)$$

•
$$\boldsymbol{\theta}^{(t+1)} = \boldsymbol{\theta}^{(t)} + \boldsymbol{J}(\boldsymbol{\theta}^{(t)})^{-1} \boldsymbol{s}(\boldsymbol{\theta}^{(t)})$$

Multidimensional extension

Common to consider log likelihood:

$$\underset{\theta}{\operatorname{argmax}} \ L(\boldsymbol{\theta}|\boldsymbol{y}) = \underset{\theta}{\operatorname{argmax}} \ L(\boldsymbol{\theta}|\boldsymbol{y}) = \underset{\theta}{\operatorname{argmax}} \log(L(\boldsymbol{\theta}|\boldsymbol{y}))$$

$$\ell(\boldsymbol{\theta}) \approx \ell(\boldsymbol{\theta}^*) + (\boldsymbol{\theta} - \boldsymbol{\theta}^*) \ell'^{(\boldsymbol{\theta}^*)} + \frac{1}{2} (\boldsymbol{\theta} - \boldsymbol{\theta}^*)^T \boldsymbol{H}(\boldsymbol{\theta}^*) (\boldsymbol{\theta} - \boldsymbol{\theta}^*)$$
$$\ell(\boldsymbol{\theta}^*) + (\boldsymbol{\theta} - \boldsymbol{\theta}^*)^T \boldsymbol{s}(\boldsymbol{\theta}^*) - \frac{1}{2} (\boldsymbol{\theta} - \boldsymbol{\theta}^*)^T \boldsymbol{J}(\boldsymbol{\theta}^*) (\boldsymbol{\theta} - \boldsymbol{\theta}^*)$$

Score function:
$$s(\theta) = \nabla \ell(\theta) = \frac{\partial}{\partial \theta} \ell(\theta)$$
 p - vector Observed information: $J(\theta) = -\nabla^2 \ell(\theta) = \frac{\partial^2}{\partial \theta^2} \ell(\theta)$ $p \times p$ - matrix

– Solving the maximum of the approximation:

$$\boldsymbol{\theta} = \boldsymbol{\theta}^* + \boldsymbol{I}(\boldsymbol{\theta}^*)^{-1}\boldsymbol{s}(\boldsymbol{\theta}^*) = \boldsymbol{\theta}^* - \boldsymbol{H}(\boldsymbol{\theta}^*)^{-1}\nabla\ell(\boldsymbol{\theta}^*)$$

Example \mathbb{R}^p

$$L(\mu) = \prod_{i=1}^{n} \frac{1}{\sqrt{(2\pi)^{p} |\Sigma|}} \exp\{-\frac{1}{2} (x_{i} - \mu)^{T} \Sigma^{-1} (x_{i} - \mu)\} \quad \Sigma \text{ known}$$

$$l(\mu) = \sum_{i=1}^{n} -\frac{p}{2} \log 2\pi - \frac{1}{2} \log |\Sigma| - \frac{1}{2} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

$$s(\mu) = \nabla l(\mu) = \sum_{i=1}^{n} \Sigma^{-1}(x_i - \mu) = \Sigma^{-1} \sum_{i=1}^{n} (x_i - \mu)$$

$$J(\mu) = -\nabla^2 l(\mu) = -\sum_{i=1}^n -\Sigma^{-1} = n\Sigma^{-1} = \left(\frac{1}{n}\Sigma\right)^{-1}$$

Stopping criteria

- Absolute convergence
 - $|x^{(t+1)} x^{(t)}| < \epsilon \text{ or } ||x^{(t+1)} x^{(t)}|| < \epsilon$
 - If x is large this might iterate too long
- Relative convergence

$$- \frac{|x^{(t+1)} - x^{(t)}|}{|x^{(t)}|} < \epsilon \quad \text{or } \frac{\|x^{(t+1)} - x^{(t)}\|}{\|x^{(t)}\|} < \epsilon$$

- Unstable if $|x^{(t)}|$ is small
 - · usually not a problem in a multivariate setting
- After N iterations (use as additional criteria)
- If not converged do not trust result
- There is in general no theorem that tells you in advance how many iterations you need
- Try different methods and starting points

Fisher scoring and ascent algorithms

- Newton's method require $\ell''(\theta) < 0$ or $J(\theta) > 0$ Multivariate: $J(\theta)$ need to be positive definite
- Note: $\mathbf{J}(\theta)$ is stochastic (depend on data)
- $I(\theta) = E[J(\theta)]$ is the expected information matrix
- Can show: $I(\theta) = \text{Var}[\mathbf{s}(\theta)]$, always positive (semi-)definite
- Fisher scoring algorithm:

$$\boldsymbol{\theta}^{(t+1)} = \boldsymbol{\theta}^{(t)} + [\mathbf{I}(\boldsymbol{\theta}^{(t)})]^{-1}\mathbf{s}(\boldsymbol{\theta}^{(t)})$$

- Will typically be more stable than Newton's method
- Can be both computationally and analytically easier
- Generalized linear models (STK3100/4100): $\mathbf{I}(\theta) = \mathbf{J}(\theta)$.
- Alternative: Ascent algorithms

$$\boldsymbol{\theta}^{(t+1)} = \boldsymbol{\theta}^{(t)} + \alpha^{(t)} \mathbf{s}(\boldsymbol{\theta}^{(t)})$$

By choosing $\alpha^{(t)}$ small enough, decrease in likelihood value can be avoided.

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Example:
$$I(\mu) = E(J(\mu)) = Var(s(\mu))$$

$$s(\mu) = \Sigma^{-1} \sum_{i=1}^{n} (x_i - \mu)$$

$$J(\mu) = \left(\frac{1}{n} \Sigma\right)^{-1}$$

1
$$E(J(\mu)) = E\left(\left(\frac{1}{n}\Sigma\right)^{-1}\right) = \left(\frac{1}{n}\Sigma\right)^{-1}$$

Independent observations
$$Var(s(\mu)) = Var\left(\Sigma^{-1} \sum_{i=1}^{n} (x_i - \mu)\right) = \sum_{i=1}^{n} \Sigma^{-1} Var(x_i - \mu) \Sigma^{-1}$$

$$= \sum_{i=1}^{n} \Sigma^{-1} \Sigma \Sigma^{-1} = n \Sigma^{-1}$$

$$= \left(\frac{1}{n} \Sigma\right)^{-1}$$

Gauss-Newton method

Assume we have a model

$$Y_i = f(\mathbf{z}_i; \boldsymbol{\theta}) + \varepsilon_i$$

and want to maximize $g(\theta) = -\sum_{i=1}^{n} (y_i - f(\mathbf{z}_i; \theta))^2$

- Newton's method: Approximate $g(\theta)$
- Gauss-Newton: Approximate $f(z_i; \theta)$:

$$\tilde{f}(\mathbf{z}_i; \mathbf{\theta}; \mathbf{\theta}^{(t)}) pprox f(\mathbf{z}_i; \mathbf{\theta}^{(t)}) + (\mathbf{\theta} - \mathbf{\theta}^{(t)}) \nabla_{\mathbf{\theta}} f(\mathbf{z}_i, \mathbf{\theta}^{(t)})$$

Gauss-Newton step: Maximize

Solution

$$\boldsymbol{\theta}^{(t+1)} = \boldsymbol{\theta}^{(t)} + [(\boldsymbol{A}^{(t)})^T \boldsymbol{A}^{(t)}]^{-1} (\boldsymbol{A}^{(t)})^T [\boldsymbol{y} - \boldsymbol{f}(\boldsymbol{z}; \boldsymbol{\theta}^{(t)})]$$

Advantage: Only need first derivatives!

ution
$$\theta^{(t+1)} = \theta^{(t)} + [(\boldsymbol{A}^{(t)})^T \boldsymbol{A}^{(t)}]^{-1} (\boldsymbol{A}^{(t)})^T [\boldsymbol{y} - \boldsymbol{f}(\boldsymbol{z}; \theta^{(t)})] \qquad A^{(t)} = \begin{bmatrix} \nabla_{\theta} f(\boldsymbol{z}_1, \theta) \\ \vdots \\ \nabla_{\theta} f(\boldsymbol{z}_n, \theta) \end{bmatrix}$$
$$n \times p$$

Other optimization methods

- Newton-type methods require derivatives
- Secant methods: Replace $J(\theta) = -\ell''(\theta)$ by finite difference approximation
- Fixed-point methods $(\max_x g(x))$
 - Find function G(x) such that $G(x) = x \Leftrightarrow g'(x) = 0$
 - Use updating scheme $x^{(t+1)} = G(x^{(t)})$
 - Obvious choice: $G(x) = \alpha g'(x) + x \Rightarrow x^{(t+1)} = x^{(t)} + \alpha g'(x^{(t)})$
 - Requirements for convergence:
 - 1 $x \in [a, b] \Rightarrow G(x) \in [a, b]$ 2 $|G(x_1) - G(x_2)| \le \lambda |x_1 - x_2|$ for all $x_1, x_2 \in [a, b]$ for some $\lambda \in (0, 1)$.
 - Newton-type methods can be seen as special cases of fixed point methods

Example fixed point

- Maximize $g(x) = x \log(x) x + 0.5x^2$, $g'(x) = \log(x) + x$
- Possible choices of G:

$$G_1(x) = g'(x) + x = \log(x) + 2x$$

 $G_2(x) = -\log(x)$
 $G_3(x) = \exp(-x)$
 $G_4(x) = (x + \exp(-x))/2$

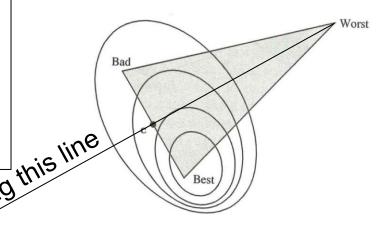
fixed_point_example.R

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Nelder - Mead

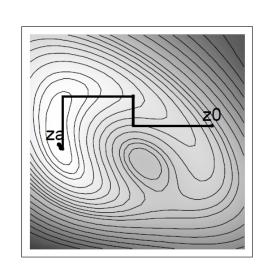
- Starts with p + 1 distinct points $\mathbf{x}_1, ..., \mathbf{x}_{p+1}$
- Points ranked through $g(\mathbf{x}_1), ..., g(\mathbf{x}_{p+1})$
- \mathbf{x}_{best} and \mathbf{x}_{worst} best and worst points
- Calculate $\mathbf{c} = \frac{1}{p} \left[\sum_{i=1}^{p+1} \mathbf{x}_i \mathbf{x}_{worst} \right]$
- Find new value $\mathbf{x}_r = \mathbf{c} + \alpha(\mathbf{c} \mathbf{x}_{worst})$, replace with \mathbf{x}_{worst}
- Require no derivatives



If you do not find improvement shrink all points towards best

Gauss-Seidel

- Aim: maximize $g(\theta)$, $\theta = (\theta_1, ..., \theta_p)$
- Procedure: For j = 1, ..., p,
 - Maximize $g(\theta)$ with respect to θ_j keeping the other θ_k 's fixed
- Reduce the multivariate problem to many univariate problems



Pick your favorite 1D optimization

BFGS-algorithm Broyden-Fletcher-Goldfarb-Shanno

• Quasi-Newton (variable metric) method (argmax g(x))

$$x_{k+1} = x_k - \alpha_k M_k^{-1} \nabla g(x_k)$$

- M_k is an approximation to the Hessian
- α_k obtained by line-search
- Do a rank 1 update of M_k to M_{k+1} using quantities computed during iterations (see book)
- Note: even though x_k converges, M_k may not converge to Hessian in optimum

optim in R

- Nelder-Mead: Default. Robust, but can be slow.
- BFGS:
 - $\mathbf{x}^{t+1} = \mathbf{x}^{(t)} (\mathbf{M}^{(t)})^{-1} \mathbf{g}'(\mathbf{x}^{(t)}), \mathbf{M}^{(t)}$ approximation of $\mathbf{g}''(\mathbf{x}^{(t)})$
 - $\mathbf{M}^{(t)}$ updated by a low-rank operation
- CG (Conjugate gradient): Optimize along gradient direction (iteratively).
- L-BFGS-B: Modification of BFGS to allow for constraints
- SANN: Simulated annealing (to be covered later)
- Brent: One-dimensional method

Recursive approaches

- Optimisation of $g(\mathbf{x})$
- Iterative approach: $\mathbf{x}^{(t+1)} = T(\mathbf{x}^{(t)})$
- Stochastic iterative approach: $\mathbf{x}^{(t+1)} = T(\mathbf{x}^{(t)}, \boldsymbol{\varepsilon}^{(t+1)})$
 - $\mathbf{x}^{(t+1)}$ only depend on $\mathbf{x}^{(t)}$ and not the previous values
 - This is called a Markov process
 - If $\mathbf{x}^{(t)}$ is discrete: Markov chain (STK2030)

Brief review of Markov chains

- Consider a stochastic sequence $X^{(t)}$, t = 0, 1, ...
- $X^{(t)} \in S$, a finite (or countable) set
- In general:

$$P(X^{(0)}, X^{(1)}, X^{(2)}, ..., X^{(n)})$$

$$= P(X^{(0)}) P(X^{(1)}, |X^{(0)}) P(X^{(2)}|X^{(0)}, X^{(1)}) \cdots P(X^{(n)}|X^{(0)}, X^{(1)}, ...X^{(n-1)})$$

Markov assumption:

$$P(X^{(t)}|X^{(0)}, X^{(1)}, ...X^{(t-1)}) = P(X^{(t)}|X^{(t-1)})$$

- Denote $P_{ij}^t = P(X^{(t)} = j | X^{(t-1)} = i)$, defines a transition matrix
- Time-homogeneous Markov chain: $P_{ij}^t = P_{ij}^1$ for all t
- A Markov chain is irreducible if any state $j \in S$ can be reached from any state $i \in S$ in a finite number of transitions.

Next time:

- Iterative re-weighted least square
- ADMM
 - Lasso example
- Combinatorial optimization (chapter 3)

Exercise