

# Math Statistics

## Weekly HW 5

Ozaner Hansha

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### Question 1

Suppose we have a population described by a Poisson distribution with parameter  $\lambda$ . That is, with pdf  $f(x; \lambda) = \frac{\lambda^x e^{-\lambda}}{x!}$  for a nonnegative integer  $x$ .

**Part a:** Compute the Fisher information about  $\lambda$  for a single observation  $X$ .

**Solution:** The Fisher information  $\mathcal{I}(\lambda)$  is given by:

$$\begin{aligned}\mathcal{I}(\lambda) &= E \left[ \left( \frac{\partial}{\partial \lambda} \ln f(X; \lambda) \right)^2 \right] && \text{(def. of Fisher information)} \\ &= E \left[ \left( \frac{\partial}{\partial \lambda} \ln \frac{\lambda^X e^{-\lambda}}{X!} \right)^2 \right] \\ &= E \left[ \left( \frac{\partial}{\partial \lambda} (X \ln \lambda - \lambda - \ln X!) \right)^2 \right] \\ &= E \left[ \left( \frac{X}{\lambda} - 1 \right)^2 \right] \\ &= \text{Var} \left( \frac{X}{\lambda} \right) && \text{(def. of variance } (E[X/\mu_X] = 1)) \\ &= \frac{1}{\lambda^2} \text{Var}(X) \\ &= \frac{\lambda}{\lambda^2} = \boxed{\frac{1}{\lambda}}\end{aligned}$$

**Part b:** Show that the sample mean  $\bar{X}$  (for sample size  $n \in \mathbb{Z}^+$ ) is a minimum variance unbiased estimator for  $\lambda$

**Solution:** First we will prove that our estimator is unbiased by showing that its mean equals  $\lambda$ :

$$\begin{aligned}E[\bar{X}] &= E \left[ \frac{\sum_{i=1}^n X_i}{n} \right] && \text{(def. of } \bar{X}) \\ &= \frac{\sum_{i=1}^n E[X_i]}{n} && \text{(linearity of expectation)} \\ &= \frac{\sum_{i=1}^n \lambda}{n} && (X_i \sim \text{Pois}(\lambda)) \\ &= \frac{n\lambda}{n} = \boxed{\lambda}\end{aligned}$$

Now recall that if the following is satisfied then  $\bar{X}$  is an MVU estimator:

$$\text{Var}(\bar{X}) = \frac{1}{n\mathcal{I}(\lambda)}$$

Where each  $X_i \sim \text{Pois}(\lambda)$  is i.i.d. And indeed the above does hold:

$$\begin{aligned} \text{Var}(\bar{X}) &= \text{Var}\left(\frac{\sum_{i=1}^n X_i}{n}\right) && (\text{def. of } \bar{X}) \\ &= \text{Var}\left(\frac{\sum_{i=1}^n X_i}{n}\right) \\ &= \frac{1}{n^2} \text{Var}\left(\sum_{i=1}^n X_i\right) \\ &= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) && (X_i \text{ are independent}) \\ &= \frac{1}{n^2} \sum_{i=1}^n \lambda && (X_i \sim \text{Pois}(\lambda)) \\ &= \frac{n\lambda}{n^2} = \frac{1}{n(1/\lambda)} = \frac{1}{n\mathcal{I}(\lambda)} && (\text{part a}) \end{aligned}$$

And so,  $\bar{X}$  is a MVU estimator for  $\lambda$ .

## Question 2

**Problem:** If  $\hat{\theta}$  is an unbiased estimator for  $\theta$ , is  $\hat{\theta}^2$  is an unbiased estimator for  $\theta^2$ ?

**Solution:** No. Consider a population  $X$  with mean  $\mu$  and variance  $\sigma^2$ . Recall that the sample mean (for sample size  $n$ ) is an unbiased estimator of a population's mean  $\mu$ :

$$E[\bar{X}] = \mu$$

However, we find that  $E[\bar{X}^2] \neq \mu^2$ :

$$\begin{aligned} \text{Var}(\bar{X}) &= E[\bar{X}^2] - E[\bar{X}]^2 && (\text{true of any RV}) \\ E[\bar{X}^2] &= \text{Var}(\bar{X}) + E[\bar{X}]^2 \\ &= \frac{\sigma^2}{n} + \mu^2 && (\text{mean/variance of sample mean}) \\ &\neq \mu^2 \end{aligned}$$

And so, as we have shown, it is not necessarily the case that the square of an unbiased estimator is itself an unbiased estimator of the parameter squared.

### Question 3

**Problem:** Suppose that we have a population with known mean  $\mu$  and unknown variance  $\sigma^2$ . Show that  $\sum_{i=1}^n \frac{(X_i - \mu)^2}{n}$  is an unbiased estimator for  $\sigma^2$ .

**Solution:** Note the following chain of equalities:

$$\begin{aligned} E \left[ \sum_{i=1}^n \frac{(X_i - \mu)^2}{n} \right] &= \frac{1}{n} \sum_{i=1}^n E[(X_i - \mu)^2] && \text{(linearity of expectation)} \\ &= \frac{1}{n} \sum_{i=1}^n \text{Var}(X_i) && \text{(def. of variance)} \\ &= \frac{1}{n} \sum_{i=1}^n \sigma^2 && \text{(def. } \sigma^2) \\ &= \frac{n\sigma^2}{n} = \sigma^2 \end{aligned}$$

And so, since the mean of this estimator is indeed the parameter  $\sigma^2$ , it is unbiased.

### Question 4

**Problem:** Show that  $\bar{X}^2 - \frac{S^2}{n}$  is an unbiased estimator for  $\mu^2$ .

**Solution:** Now consider the following chain of equalities:

$$\begin{aligned} E \left[ \bar{X}^2 - \frac{S^2}{n} \right] &= E[\bar{X}^2] - \frac{E[S^2]}{n} && \text{(linearity of expectation)} \\ &= \text{Var}(\bar{X}) + E[\bar{X}]^2 - \frac{E[S^2]}{n} && (E[\bar{X}^2] = \text{Var}(\bar{X}) + E[\bar{X}]^2) \\ &= \frac{\sigma^2}{n} + \mu^2 - \frac{\sigma^2}{n} && (E[S^2] = \sigma^2) \\ &= \mu^2 \end{aligned}$$

And so, since the mean of this estimator is indeed the parameter  $\mu^2$ , it is unbiased.

Where  $S^2$  is the corrected sample variance  $\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$ .