## Math Statistics Semiweekly HW 15

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November 24, 2020

## Question 1

**Problem:** Suppose we have a Poisson population  $X_i \sim \text{Pois}(\lambda)$ . Our null hypothesis  $H_0$  is that  $\lambda = 2$  and our alternative hypothesis  $H_1$  is that  $\lambda = 1$ . Given an i.i.d. sample X of n = 7, use the Neyman-Pearson lemma to produce a most powerful critical region C for  $\alpha = .01$  to test these hypotheses.

**Solution:** Note that the ratio of the likelihood of the null and alternative hypotheses must be less than some finite constant k (since we can only observe items in the support):

$$k \geq \frac{L(\lambda_0)}{L(\lambda_1)}$$

$$= \frac{\prod_{i=1}^n P(X_i; \lambda = \lambda_0)}{\prod_{i=1}^n P(X_i; \lambda = \lambda_1)} \qquad \text{(def. of likelihood)}$$

$$= \prod_{i=1}^n \frac{P(X_i; \lambda = \lambda_0)}{P(X_i; \lambda = \lambda_1)} \qquad \text{(associativity)}$$

$$= \prod_{i=1}^n \frac{\lambda_0^{X_i} e^{-\lambda_0} / X_i!}{\lambda_1^{X_i} e^{-\lambda_1} / X_i!} \qquad \text{(pmf of Poisson RV)}$$

$$= \prod_{i=1}^n \left(\frac{\lambda_0}{\lambda_1}\right)^{X_i} e^{\lambda_1 - \lambda_0}$$

$$= \lim_{i=1}^n \left(\frac{\lambda_0}{\lambda_1}\right)^{X_i} e^{\lambda_1 - \lambda_0} \qquad \text{(In is monotone increasing)}$$

$$= \sum_{i=1}^n X_i (\ln \lambda_0 - \ln \lambda_1) + \lambda_1 - \lambda_0$$

$$= n(\lambda_1 - \lambda_0) + (\ln \lambda_0 - \ln \lambda_1) \sum_{i=1}^n X_i$$

$$\frac{\ln k - n(\lambda_1 - \lambda_0)}{\ln \lambda_0 - \ln \lambda_1} \geq \sum_{i=1}^n X_i$$

$$\frac{\ln k + 7}{\ln 2} \geq \sum_{i=1}^n X_i$$

$$k' \geq \sum_{i=1}^n X_i \qquad \text{(let LHS} = k')$$

Note that since both our hypotheses are simple, the Neyman-Pearson lemma tells us that the most powerful test for testing them is given by the above inequality for some constant k':

$$\sum_{i=1}^{n} X_i \le k'$$

What that k' is depends on our  $\alpha$  which, in this case, is equal to .01. Below we will solve for k':

$$\alpha = P(\text{Type I error}) \qquad (\text{def. of } \alpha)$$

$$= P(\neg \hat{H}_0 \mid H_0) \qquad (\text{def. of Type I error})$$

$$= P\left(\sum_{i=1}^n X_i > k' \mid H_0\right) \qquad (\text{Neyman-Pearson lemma})$$

$$= P\left(\sum_{i=1}^n X_i > k' \mid \lambda = \lambda_0\right) \qquad (\text{given null hypothesis})$$

$$= 1 - p_Y(k') \qquad (Y = \sum_{i=1}^n X_i \sim \text{Pois}(n\lambda_0), \text{ i.i.d } X_i)$$

$$.01 = 1 - p_Y(k') \qquad (\text{desired } \alpha)$$

$$.99 = p_Y(k')$$

And so to solve for k' we must find the 99th quantile of the Poisson cdf with  $\lambda = n\lambda_0 = 14$ . Using a calculator, we find that:

$$k' = q(.99; \lambda = 14) \approx 22.88387$$

And so our most powerful test of the hypotheses that splits  $\mathbb{R}$  into regions  $R_0$  and  $R_1$  is given by:

$$\sum_{i=1}^{7} X_i \le 22.88387$$
 
$$\bar{X} \le 3.26912$$
 (divide both sides by  $n=7$ )

If the sample passes this test we accept the alternative hypothesis, else we reject it.