

Theory of Probability HW #2

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Part A

Problems taken from Chapters 2 and 3 of the textbook.

Chapter 2

Problem 31a

Problem: A 3-person basketball team consists of a guard, a forward, and a center. If a person is chosen at random from each of three different such teams, what is the probability of selecting a complete team? (*use the inclusion-exclusion principle instead of calculating directly*)

Solution: Let G denote the chosen team has 1 guard, F denote that it was onw forward, and C denote that it has one center. Then the probability we seek is $P(GFC)$. Note that we can express this in terms of the probability of the union of events:

$$\begin{aligned} P(GFC) &= P(((GFC)^c)^c) && \text{(involutory property)} \\ &= P((G^c \cup F^c \cup C^c)^c) && \text{(DeMorgan's law)} \\ &= 1 - P(G^c \cup F^c \cup C^c) && \text{(complement of event)} \end{aligned}$$

Via the inclusion-exclusion principle for $n = 3$ we have:

$$\begin{aligned} P(G^c \cup F^c \cup C^c) &= P(G^c) + P(F^c) + P(C^c) \\ &\quad - P(G^c F^c) - P(G^c C^c) - P(F^c C^c) \\ &\quad + P(G^c F^c C^c) \end{aligned}$$

Now note that the probability that any particular position is not chosen from any particular team is $\frac{2}{3}$ since there are 2 valid options out of the 3 members. Since there are 3 teams the principle of counting gives us:

$$P(G^c) = P(F^c) = P(C^c) = \frac{2}{3} \cdot \frac{2}{3} \cdot \frac{2}{3} = \frac{2^3}{3^3}$$

In a similar vain, the probability that any 2 different positions are not chosen from any particular team is $\frac{1}{3}$ since there is only 1 valid option out of the 3 members. Since there are 3 teams the principle of counting gives us:

$$P(G^c F^c) = P(G^c C^c) = P(F^c C^c) = \frac{1}{3} \cdot \frac{1}{3} \cdot \frac{1}{3} = \frac{1}{3^3}$$

And taken to the extreme, the probability that none of the three positions are chosen from any particular team is 0 since each member has a position and at least 1 member from each team must be chosen:

$$P(G^{\mathfrak{C}} F^{\mathfrak{C}} C^{\mathfrak{C}}) = 0 \cdot 0 \cdot 0 = 0$$

Plugging this into the inclusion-exclusion principle we have:

$$\begin{aligned} P(G^{\mathfrak{C}} \cup F^{\mathfrak{C}} \cup C^{\mathfrak{C}}) &= \frac{2^3}{3^3} + \frac{2^3}{3^3} + \frac{2^3}{3^3} - \frac{1}{3^3} - \frac{1}{3^3} - \frac{1}{3^3} + 0 \\ &= \frac{3 \cdot 2^3}{3^3} - \frac{3}{3^3} = \frac{7}{9} \end{aligned}$$

And so our desired probability is given by:

$$P(GFC) = 1 - P(G^{\mathfrak{C}} \cup F^{\mathfrak{C}} \cup C^{\mathfrak{C}}) = 1 - \frac{7}{9} = \frac{2}{9}$$

Problem 45

Problem: A woman has n keys, of which only 1 will open her door. **a)** If she tries keys at random, discarding those that do not work, what's the probability she will open the door on her k th try? **b)** What if she doesn't discard the keys?

Solution: Let E_k be the event she opens the door on the k th key. **a)** Note that the first try has a $\frac{1}{n}$ chance of being correct, the second has a $P(E_1^{\mathfrak{C}}) \frac{1}{n-1}$ chance since the first try had to fail and since 1 key has been discarded, and so on. This pattern gives us:

$$\begin{aligned} P(E_1) &= \frac{1}{n} \\ P(E_2) &= \underbrace{\frac{n-1}{n}}_{P(E_1^{\mathfrak{C}})} \cdot \frac{1}{n-1} \\ P(E_3) &= \underbrace{\frac{n-1}{n}}_{P(E_1^{\mathfrak{C}})} \cdot \underbrace{\frac{n-2}{n-1}}_{P(E_2^{\mathfrak{C}})} \cdot \frac{1}{n-2} \\ &\vdots \\ P(E_k) &= \underbrace{\frac{n-1}{n}}_{P(E_1^{\mathfrak{C}})} \cdot \underbrace{\frac{n-2}{n-1}}_{P(E_2^{\mathfrak{C}})} \cdots \underbrace{\frac{n-k+1}{n-k+2}}_{P(E_{k-1}^{\mathfrak{C}})} \cdot \frac{1}{n-k+1} \end{aligned}$$

It is quite plain to see that this is a finite telescoping product. Canceling out the like factors, we are left with the very simple:

$$P(E_k) = \frac{1}{n}$$

b) If she doesn't discard the keys, then each try is independent of any other. And so the probability of her opening the door on the k th try is the probability that she failed the first $k-1$ tries then succeeded k th one:

$$P(E_k) = \left(\frac{n-1}{n} \right)^{k-1} \frac{1}{n}$$

Problem 55a

Problem: What's the probability that a hand of 13 cards contains both ace and king of at least one suit?

Solution: Let E_{\heartsuit} denote the event that our hand has both an ace and king of hearts, E_{\spadesuit} denote the event that our hand has both an ace and king of diamonds, and so on. Thus, we desire the probability of the following event $E_{\heartsuit} \cup E_{\spadesuit} \cup E_{\diamondsuit} \cup E_{\clubsuit}$. The cardinality of this event is given by the inclusion-exclusion principle:

$$|E_{\heartsuit} \cup E_{\spadesuit} \cup E_{\diamondsuit} \cup E_{\clubsuit}| = \left| \bigcup_{i \in \{\heartsuit, \spadesuit, \diamondsuit, \clubsuit\}} E_i \right| = \sum_{k=1}^4 \left((-1)^{k-1} \sum_{\substack{I \subseteq \{\heartsuit, \spadesuit, \diamondsuit, \clubsuit\} \\ |I|=k}} \left| \bigcap_{i \in I} E_i \right| \right)$$

Now note that each suit we have determines 2 of the cards in our hand. For example, $|E_{\heartsuit}| = \binom{52-2}{13-2}$ since 2 of the cards have been chosen for us. Similarly, we have $|E_{\heartsuit} E_{\spadesuit}| = \binom{52-4}{13-4}$ since 4 of the cards (2 from hearts and 2 from spades) have been chosen for us. Generalizing this, we have for $|I| = k$:

$$\left| \bigcap_{i \in I} E_i \right| = \binom{52-2k}{13-2k}$$

Plugging this into our expression for the inclusion-exclusion principle gives us:

$$\left| \bigcup_{i \in \{\heartsuit, \spadesuit, \diamondsuit, \clubsuit\}} E_i \right| = \sum_{k=1}^4 \left((-1)^{k-1} \sum_{\substack{I \subseteq \{\heartsuit, \spadesuit, \diamondsuit, \clubsuit\} \\ |I|=k}} \binom{52-2k}{13-2k} \right)$$

Now, noting that there are $\binom{4}{k}$ k -combinations of the set $\{\heartsuit, \spadesuit, \diamondsuit, \clubsuit\}$, we have:

$$\left| \bigcup_{i \in \{\heartsuit, \spadesuit, \diamondsuit, \clubsuit\}} E_i \right| = \sum_{k=1}^4 \left((-1)^{k-1} \binom{4}{k} \binom{52-2k}{13-2k} \right) = 139565328072$$

We have calculated the number of hands with at least 1 pair of an ace and heart of a single suit. And because this is a discrete uniform distribution, i.e. each hand is equally likely, the probability of drawing such a hand is simply:

$$P(E_{\heartsuit} \cup E_{\spadesuit} \cup E_{\diamondsuit} \cup E_{\clubsuit}) = \frac{|E_{\heartsuit} \cup E_{\spadesuit} \cup E_{\diamondsuit} \cup E_{\clubsuit}|}{|\Omega|} = \frac{139565328072}{\binom{52}{13}} = \frac{9895443}{45023650} \approx 0.219783$$

Chapter 3

Problem 12

Problem: Suppose distinct values are written on each of 3 cards, which are then randomly named a, b , and c . Given that $a < b$ find the probability that $a < c$.

Solution: Note that we can consider the sample space Ω of this experiment the $3!$ passive permutations of $\{a, b, c\}$:

$$\Omega = \{(a, b, c), (a, c, b), (b, a, c), \dots\}$$

The event $E_{a < b}$ is given by the set of all triplets in Ω in which a precedes b :

$$E_{a < b} = \{(a, b, c), (a, c, b), (c, a, b)\}$$

The event $E_{a < c}$ is given by the set of all triplets in Ω in which a precedes c :

$$E_{a < c} = \{(a, b, c), (a, c, b), (b, a, c)\}$$

The intersection of these two events $E_{a < b}E_{a < c}$ is thus given by:

$$E_{a < b}E_{a < c} = \{(a, b, c), (a, c, b)\}$$

And now, noting that this is a discrete uniform distribution, we can finally calculate the desired conditional probability of $a < c$ assuming $a < b$:

$$P(E_{a < c} | E_{a < b}) = \frac{|E_{a < b}E_{a < c}|}{|E_{a < b}|} = \frac{2}{3}$$

Problem 30

Problem: Suppose that a deck of 52 cards is shuffled and the cards are then turned over one at a time until the first ace appears. Given that the first ace is the 20th card, what is the conditional probability that the card following it is **a)** the ace of spades? **b)** the two of clubs?

Solution: Let A denote the event that the first ace drawn is the 20th card, B denote that the 21st card is the ace of spades, and C denote that the 21st card is the two of clubs. The cardinality of A is given by:

$$|A| = \underbrace{48 \cdot 47 \cdot 46 \cdots 30}_{\substack{\text{non-ace} \\ \text{cards} \\ c1 \quad c2 \quad c3 \quad \dots \quad c19}} \cdot \underbrace{4}_{\substack{\text{ace} \\ \text{cards} \\ c20}} \cdot \underbrace{32!}_{\substack{\text{all other} \\ \text{cards} \\ c21-c52}} = 48^{19} \cdot 4 \cdot 32!$$

Via similar reasoning, the cardinality of BA is given by:

$$|BA| = \underbrace{48 \cdot 47 \cdot 46 \cdots 30}_{\substack{\text{non-ace} \\ \text{cards} \\ c1 \quad c2 \quad c3 \quad \dots \quad c19}} \cdot \underbrace{3}_{\substack{\text{aces of} \\ \diamond, \clubsuit, \heartsuit \\ c21}} \cdot \underbrace{1}_{\substack{\text{ace of } \spadesuit \\ c21}} \cdot \underbrace{31!}_{\substack{\text{all other} \\ \text{cards} \\ c22-c52}} = 48^{19} \cdot 3 \cdot 1 \cdot 31!$$

Yet again via similar reasoning, the cardinality of CA is given by:

$$|CA| = \underbrace{47 \cdot 46 \cdot 45 \cdots 29}_{\substack{\text{non-ace/two of } \clubsuit \\ \text{cards} \\ c1 \quad c2 \quad c3 \quad \dots \quad c19}} \cdot \underbrace{4}_{\substack{\text{ace} \\ \text{cards} \\ c20}} \cdot \underbrace{1}_{\substack{\text{two of } \clubsuit \\ c21}} \cdot \underbrace{31!}_{\substack{\text{all other} \\ \text{cards} \\ c22-c52}} = 47^{19} \cdot 4 \cdot 31!$$

The probabilities we seek are **a)** $P(B|A)$ and **b)** $P(C|A)$ and since this is a discrete uniform distribution, we have:

$$P(B|A) = \frac{|BA|}{|A|} = \frac{48^{19} \cdot 3 \cdot 1 \cdot 31!}{48^{19} \cdot 4 \cdot 32!} = \frac{3}{4 \cdot 32} = \frac{3}{128}$$

$$P(C|A) = \frac{|CA|}{|A|} = \frac{47^{19} \cdot 4 \cdot 31!}{48^{19} \cdot 4 \cdot 32!} = \frac{29}{48 \cdot 32} = \frac{29}{1536}$$

Problem 35

Problem: On rainy days Joe is late to work with probability 0.3, and on nonrainy days he is late with probability 0.1. It will rain tomorrow with probability 0.7. **a)** Find the probability that Joe is early tomorrow. **b)** Given that Joe is early, what is the conditional probability that it rained?

Solution: Let L denote the event that Joe is late, and R denote the event that it is raining. The problem statement and the complement rule gives us the following probabilities:

$$\begin{aligned}P(R) &= 0.7 & P(R^c) &= 0.3 \\P(L|R) &= 0.3 & P(L|R^c) &= 0.1\end{aligned}$$

Now note that via the law of total probability, $P(L)$ is given by:

$$\begin{aligned}P(L) &= P(L|R)P(R) + P(L|R^c)P(R^c) \\&= 0.3 \cdot 0.7 + 0.1 \cdot 0.3 = 0.24\end{aligned}$$

And so for **a)** the probability that Joe is early (i.e. not late) is given by:

$$P(L^c) = 1 - P(L) = 1 - 0.24 = 0.76$$

For **b)** the desired probability is given by $P(R|L^c)$. We can solve for this probability via Bayes' theorem:

$$\begin{aligned}P(R|L^c) &= \frac{P(L^c|R)P(R)}{P(L^c)} && \text{(Bayes' theorem)} \\&= \frac{(1 - P(L|R))P(R)}{P(L^c)} && \text{(complement of conditional)} \\&= \frac{(1 - 0.3)0.7}{0.76} = \frac{49}{76} \approx 0.644734\end{aligned}$$

Problem 47a

Problem: There is a 30% chance that A can fix her busted computer. If A cannot, then there is a 40% chance that her friend B can fix it. Find the probability it will be fixed by either A or B .

Solution: Let A denote the probability that she fixes the computer, and B denote that the friend fixes the computer. Thus the problem statement gives us:

$$P(A) = 0.3 \quad P(B|A^c) = 0.4$$

Now, note that by Bayes' theorem we have the following chain of equalities:

$$P(B|A^c) = \frac{P(A^c|B)P(B)}{P(A^c)} \quad (\text{Bayes' theorem})$$

$$0.4 = \frac{P(A^c|B)P(B)}{1 - 0.3}$$

$$0.4 \cdot 0.7 = P(A^c|B)P(B)$$

$$0.28 = (1 - P(A|B))P(B) \quad (\text{complement of conditional})$$

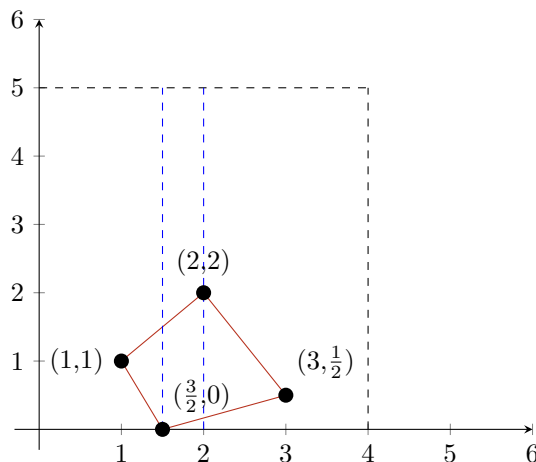
$$0.28 = P(B) - P(A|B)P(B)$$

$$0.28 = P(B) - P(AB) \quad (\text{def. of conditional probability})$$

Now, via the inclusion-exclusion principle for $n = 2$, we can solve for our desired probability, that either A or B fix the computer, $P(A \cup B)$:

$$P(A \cup B) = P(A) + P(B) - P(AB) = 0.3 + 0.28 = 0.58$$

Part B



Problem a

Problem: Express the area in the red bounded region E as a sum of integrals.

Solution: As we can see, the red bounded region can be expressed as the 3 regions split up by the blue dashed lines. The functions of each side, starting with the top left line and going clockwise, are given by:

$$\begin{aligned} y &= x \\ y &= \frac{-3x}{2} + 5 \\ y &= \frac{x}{3} - \frac{1}{2} \\ y &= -2x + 3 \end{aligned}$$

And so we can express our integral sum like so:

$$\int_1^{\frac{3}{2}} \int_{-2x+3}^x dy dx + \int_{\frac{3}{2}}^2 \int_{\frac{x}{3}-\frac{1}{2}}^x dy dx + \int_2^3 \int_{\frac{x}{3}-\frac{1}{2}}^{\frac{-3x}{2}+5} dy dx$$

Problem b

Problem: Compute the area of E .

Solution: Computing the integral sum, we have:

$$\begin{aligned} & \int_1^{\frac{3}{2}} \int_{-2x+3}^x dy dx + \int_{\frac{3}{2}}^2 \int_{\frac{x}{3}-\frac{1}{2}}^x dy dx + \int_2^3 \int_{\frac{x}{3}-\frac{1}{2}}^{\frac{-3x}{2}+5} dy dx \\ &= \int_1^{\frac{3}{2}} [y]_{-2x+3}^x dx + \int_{\frac{3}{2}}^2 [y]_{\frac{x}{3}-\frac{1}{2}}^x dx + \int_2^3 [y]_{\frac{x}{3}-\frac{1}{2}}^{\frac{-3x}{2}+5} dx \\ &= \int_1^{\frac{3}{2}} 3x - 3 dx + \int_{\frac{3}{2}}^2 \frac{2x}{3} + \frac{1}{2} dx + \int_2^3 \frac{-11x}{6} + \frac{11}{2} dx \\ &= \left[\frac{3x^2}{2} - 3x \right]_1^{\frac{3}{2}} + \left[\frac{2x^2}{6} + \frac{x}{2} \right]_{\frac{3}{2}}^2 + \left[\frac{-11x^2}{12} + \frac{11x}{2} \right]_2^3 \\ &= \frac{3}{8} + \frac{5}{6} + \frac{11}{12} = \frac{17}{8} \end{aligned}$$

Problem c

Problem: Suppose the rectangle R (demarcated by the black dashed lines) represents the backyard of a house where a drone is trying to drop a (very small) parcel, and E represents a swimming pool in the backyard. If the drone drops the parcel in R at random, what is the probability that the parcel will not fall in the swimming pool?

Solution: Assuming the parcel is point-like and that the area R represents a uniform distribution, we can imagine that the probability the parcel lands in the pool $P(L)$ is given by:

$$P(L) = \frac{\text{Area of } E}{\text{Area of } R} = \frac{17}{8 \cdot 20} = \frac{17}{160}$$

And so the probability it doesn't fall in is simply the complement:

$$P(L^c) = 1 - P(L) = 1 - \frac{17}{160} = \frac{143}{160}$$