STAT 3006: Statistical Computing Lecture 9*

13 March

7.3 Hybrid Gibbs Sampler

In the last subsection, we have known that Gibbs sampler is sampling from full conditional functions iteratively to approximate samples from the target distribution $f(\mathbf{x})$ of our interest. In some case, for a (or some) specific conditional function $f_i(x_i|x_1,\ldots,x_{i-1},x_{i+1},\ldots,x_p)$, it does not correspond to a common distribution. Directly drawing from $f_i(x_i|x_1,\ldots,x_{i-1},x_{i+1},\ldots,x_p)$ is difficult. We can generate a MH sample in this step, which requires a proposal distribution $q_i(y|x_i)$. The algorithm combines the Gibbs sampler and the MH algorithm into a hybird one, thus called hybrid Gibbs sampler.

Algorithm: Hybrid Gibbs sampler to sample from $f(\mathbf{x})$.

Input: the pdf (or pmf) $f(\mathbf{x})$ and a starting point $\mathbf{x}^{(0)} = (x_1^{(0)}, \dots, x_p^{(0)})$. f_1, f_2, \dots, f_p are full conditional functions of $f(\mathbf{x})$. Without loss of generality, f_1, \dots, f_{p-1} are common distributions that are easy to sample, but f_p is not a standard distribution to be sampled using MH algorithm.

Initialize: $t \leftarrow 0$.

Repeat

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\begin{array}{l} \text{generate } x_1^{(t+1)} \sim f_1(x_1|x_2^{(t)}, x_3^{(t)}, \dots, x_p^{(t)}); \\ \vdots \\ \text{generate } x_{p-1}^{(t+1)} \sim f_{p-1}(x_{p-1}|x_1^{(t+1)}, x_2^{(t+1)}, \dots, x_p^{(t)}); \\ (\text{MH step}) \quad \text{generate a proposal } y \text{ from } q_p(y|x_p^{(t)}); \\ \text{calculate } r, \text{ the value of } \min\left(\frac{f_p(y|x_1^{(t+1)}, \dots, x_{p-1}^{(t+1)})}{f_p(x_p^{(t)}|x_1^{(t+1)}, \dots, x_{p-1}^{(t+1)})} \frac{q(x_p^{(t)}|y)}{q(y|x_p^{(t)})}, 1\right). \\ \text{accept } x_p^{(t+1)} \text{ as } y \text{ with probability } r; \text{otherwise, reject } y \text{ and let } x_p^{(t+1)} \text{ be } x_p^{(t)}. \\ t \leftarrow t+1; \end{array}
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Until some criteria are met.

Output:Given a large number B, $\{\mathbf{x}^{(B)}, \mathbf{x}^{(B+1)}, \ldots\}$ are samples from the distribution f.

Remark 1. MH step only implements MH algorithm once in one Gibbs sampler cycle.

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7.4 One Example of Implementing Gibbs Sampler

There are n subjects. For each subject j(j = 1, ..., n), it has a G dimensional vector $(x_{1j}, x_{2j}, ..., x_{Gj})$. We know n subjects can be clustered into K groups. We use C_j to represent the cluster to which subject j belongs, and π_k to represent the proportion of cluster k. Given $C_j = k$, $(x_{gj} \sim Poi(\lambda_{gk}))$ for g = 1, ..., G. $\{x_{gj} : g = 1, ..., G, j = 1, ..., n\}$ are observed, and the missing data is $\mathbf{C} = \{C_j : j = 1, ..., n\}$. The parameters are $\pi = \{\pi_1, ..., \pi_K\}$ and $\Lambda = \{\lambda_{gk}\}(g = 1, ..., G, k = 1, ..., K)$.

Based on the information, we have the complete-data likelihood function

$$f(\mathbf{X}, \mathbf{C} | \pi, \Lambda) = \prod_{j=1}^{n} \prod_{k=1}^{K} \left[\pi_k \prod_{q=1}^{G} \frac{\lambda_{gk}^{x_{gj}}}{x_{gj}!} e^{-\lambda_{gk}} \right]^{I(C_j = k)}.$$

We assign the following priors on (π, Λ) :

$$(\pi_1, \dots, \pi_K) \sim Dirichlet(\alpha_1, \alpha_2, \dots, \alpha_K);$$

$$\lambda_{qk} \sim Gamma(\alpha, \beta).$$

Subsequently, the posterior distribution for $(\pi, \Lambda, \mathbf{C})$ is

$$p(\pi, \lambda, \mathbf{C}|\mathbf{X}) \propto f(\mathbf{X}, \mathbf{C}|\pi, \Lambda)p(\pi)p(\Lambda).$$

The Gibbs sampler to simulate $p(\pi, \lambda, \mathbf{C}|\mathbf{X})$ proceeds as follows.

- $p(\lambda_{gk}|-) \propto \lambda_{gk}^{\sum_{j=1}^{n} I(C_j=k)x_{gj}+\alpha-1} e^{-\lambda_{gk}(\beta+\sum_{j=1}^{n} I(C_j=k))}$, so $\lambda_{gk} \sim Gamma(\sum_{j=1}^{n} I(C_j=k))$.
- $p(\pi|-) \propto \prod_{k=1}^K \pi_k^{\sum_{j=1}^n I(C_j=k) + \alpha_k 1}$, so $\pi \sim Dirichlet(\sum_{j=1}^n I(C_j=1) + \alpha_1, \dots, \sum_{j=1}^n I(C_j=1) + \alpha_k)$.
- $p(C_j = k|-) = \frac{\pi_k \prod_{g=1}^G \lambda_{gk}^{xgj} e^{-\lambda_{gk}}}{\sum_{l=1}^K \pi_l \prod_{g=1}^G \lambda_{gl}^{xgj} e^{-\lambda_{gl}}}.$

"-" means given other variables.