mgcv: GAMs in R

Simon Wood

Mathematical Sciences, University of Bath, U.K.

mgcv, gamm4

- mgcv is a package supplied with R for generalized additive modelling, including generalized additive mixed models.
- ▶ The main GAM fitting routine is gam.
- bam provides an alternative for very large datasets.
- The main GAMM fitting is gamm which uses PQL based on package nlme.
- gamm4 is an R package available from cran.r-project.org supplying gamm4, a version of gamm which uses 1me4 for GAMM fitting, and avoids PQL. It is really an extension package for mgcv.
- ► The packages are loaded into R using e.g. library(mgcv).

mgcv help

- ➤ To get 'overview' help type help.start() within R and follow the Packages link to mgcv.
- The help page mgcv-package is a good page to start to get an overview.
- Note other overview pages such as gam.models and gam.select.
- All user visible functions are documented.
- Once a library is loaded its help pages are accessible via, e.g. help("gam") or ?gam.
- ► Technical references on the underlying methods are given in ?gam and via citation("mgcv").
- ▶ Wood (2006) Generalized additive models:an introduction with R CRC/Taylor&Francis provides further information.

gam

- Use of the function gam is similar to the use of function glm, except for the following
 - 1. The model formula can contain smooth terms s and tensor product smooth terms te in the linear predictor.
 - There are extra arguments controlling smoothing parameter estimation, notably method for choosing between "REML", "ML", "GCV.Cp" and "GACV.Cp" smoothness selection.
 - 3. The family argument can also be Tweedie or negbin.
- gam returns and object of class "gam", which can be further interrogated using method functions such as print, summary, anova, plot, predict, residuals etc.
- ► The front end design of gam and its associated functions is based heavily on Trevor Hastie's original gam function for S. The underlying model representation and numerical methods are very different, however, being based on the penalized regression spline methods covered in this course.

family arguments usable with gam

- gaussian (default) is useful for real valued response data.
- Gamma is useful for strictly positive real valued data. The default link is only useful in some waiting time applications, and the log link is more often used.
- poisson is useful when the response is count data of some sort.
- binomial is used most often for binary (logistic) regression, but is applicable to any response that is the number of successes from a known number of trials.
- inverse.gaussian is for strictly positive real response variables: useful for various 'time to event' data.
- quasi does not define a full distribution, but allows inference when only the mean variance relationship can be well approximated. quasipoisson and quasibinomial are special cases. Not useable with likelihood based smoothness selection.
- Tweedie is an alternative to quasi when $var(y) = \phi \mu^p$, 1 , and a full distribution is required (for a non-negative real response).
- negbin is useful for overdispersed count data, but computation is slow.

Specifying models: some examples

- ▶ $y_i = f(x_i) + \epsilon_i$ where $\epsilon_i \sim N(0, \sigma^2)$ gam(y ~ s(x))
- ▶ $\log{\mathbb{E}(y_i)} = f_1(x_i) + f_2(z_i) + f_3(v_i) + w_i$ where $y_i \sim \text{Poi.}$ gam(y ~ s(x) + s(z) + s(v) + w,family=poisson)
- $\sqrt{\mathbb{E}(y_i)} = f_1(\mathsf{time}_i, \mathsf{distance}_i) + f_2(w_i), \ y_i \sim \mathsf{gamma}.$ $\mathsf{gam}(\mathsf{y} \sim \mathsf{te}(\mathsf{time}, \mathsf{distance}) + \mathsf{s}(\mathsf{w}),$ $\mathsf{family=Gamma}(\mathsf{link=sqrt}))$

uses a scale invariant tensor product smooth for f_1 .

▶ logit{ $\mathbb{E}(y_i)$ } = $w_i f_1(x_i, z_i) + f_2(v_i)$, $y_i \sim$ binary. gam(y ~ s(x,z,by=w) + s(v), family=binomial) Here f_1 is isotropic (a thin plate spline).

s term details

- s(x,k=20,id=2,bs="tp") is an example smooth specifier, used in a formula. (Some) arguments are . . .
 - x is the covariate of the smooth (can have any name!): some types of smooth can have several covariates (e.g. "tp").
 - bs is the type of basis-penalty smoother.
 - k is the basis dimension for the smooth (before imposing any identifiability constraints).
 - id used to allow different smooths to be forced to use the same basis and smoothing parameter.
 - sp allows the smoothing parameter to be supplied.
 - fx if TRUE then the term is unpenalized.
 - by allows specification of interactions of the smooth with a factor or metric variable.
 - m specifies the penalty order for some bases.

Smooth classes

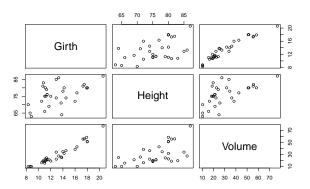
- Built in smooth classes (i.e. options for the bs argument of s) are:
 - "cr" a penalized cubic regression spline ("cc" for cyclic version).
 - "ps" Eilers and Marx style P-splines ("cp" for cyclic).
 - "ad" adaptive smoothers based on "ps".
 - "tp" Optimal low rank approximation to thin plate spline, any dimension and permissable penalty order is possible.
- In addition the "re" class implements simple random effects. For example s(x,z,bs="re") specifies a random effect **Zb** where **b** ~ N(**0**, Iσ_b²). **Z** is given by model.matrix(~x:z-1). This approach is slow for large numbers od random effects, however.
- ▶ New classes can be added. See ?smooth.construct

Tensor product smoothing in mgcv

- ▶ Tensor product smooths are constructed automatically from marginal smooths of lower dimension. The resulting smooth has a penalty for each marginal basis.
- mgcv can construct tensor product smooths from any single penalty smooths useable with s terms.
- ▶ te terms within the model formula invoke this construction. For example:
 - ► te(x,z,v,bs="ps",k=5) creates a tensor product smooth of x, z and v using rank 5 P-spline marginals: the resulting smooth has 3 penalties and basis dimension 125.
 - ▶ te(x,z,t,bs=c("tp","cr"),d=c(2,1),k=(20,5)) creates a tensor product of an isotropic 2-D TPS with a 1-D smooth in time. The result is isotropic in x,z, has 2 penalties and a basis dimension of 100. This sort of smooth would be appropriate for a location-time interaction.

Simple data example

➤ To illustrate the basics of gam, consider a very simple dataset relating the timber volume of cherry trees to their height and trunk girth.



trees initial gam fit

A possible model is

$$\log(\mu_i) = f_1(\mathtt{Height}_i) + f_2(\mathtt{Girth}_i), \ \ \mathtt{Volume}_i \sim \mathsf{Gamma}(\mu_i, \phi)$$

- ▶ The results are stored in class "gam" object ct1.
- ► For the full contents of a "gam" object see ?gamObject

print.gam

- ► Typing ct1 causes R to pass ct1 to the print method function.
- ► For class "gam" object ct1 this means printing by print.gam.
 > ct1

```
Family: Gamma
Link function: log

Formula:
Volume ~ s(Height) + s(Girth)

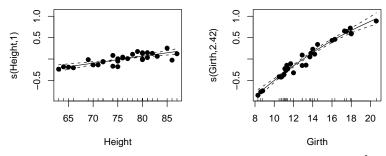
Estimated degrees of freedom:
1.0000 2.4222 total = 4.422254
```

GCV score: 0.008082356

Notice how the EDFs for each term and the GCV score are reported.

plot.gam

par(mfrow=c(1,2))
plot(ct1,residuals=TRUE,pch=19) ## calls plot.gam

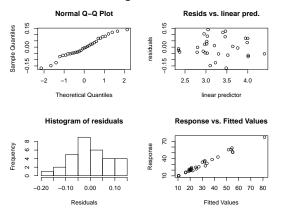


- **partial** residuals for f_j are weighted working residuals from PIRLS added to \hat{f}_j . Systematic departure from f_i indicates a problem.
- Rug plot shows values of predictors.
- EDF for term reported in y axis lable.
- ▶ 95% Bayesian CIs shown (constraint causes vanishing CI on left).



Basic model checking: gam.check

> gam.check(ct1) ## note QQ beefed up for next mgcv version ## smoothness selection convergence info omitted

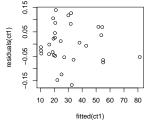


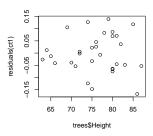
- Deviance residuals are used: often approximately normal.
- ▶ Plots are utterly useless for binary data!

residuals

Other residual plots should be examined. e.g.

```
plot(fitted(ct1),residuals(ct1))
plot(trees$Height,residuals(ct1))
```





- OK, but possibly the short trees are less variable than Gamma is suggesting.
- ► Alternatives for residual argument type are "deviance" (default), "pearson", "scaled.pearson", "working" and (the unstandardized) "response".

Checking basis dimensions k

Can simply increase suspect k and see if fit measures change...

```
> gam(Volume~s(Height)+s(Girth,k=20),
+ family=Gamma(link=log),data=trees)

Family: Gamma
Link function: log

Formula:
Volume~s(Height) + s(Girth, k = 20) ## increase k from 10 to 20

Estimated degrees of freedom:
1.0000 2.4248 total = 4.424817

GCV score: 0.008083182
```

- ▶ ... no evidence for change in fit.
- ▶ Note that a 20 dimensional basis includes a larger space of EDF 2.5 functions, than a 10 dimensional basis...

Cheaper k check

- Refitting complex models with increased k can be costly.
- ▶ A cheaper alternative is to check the adequacy of k by checking for unmodelled pattern in the residuals...

```
> rsd <- residuals(ct1,type="deviance")
> gam(rsd~s(Girth,k=20)-1,data=trees,select=TRUE)
Family: gaussian
Link function: identity

Formula:
rsd ~ s(Girth, k = 20) - 1

Estimated degrees of freedom:
4.839e-09 total = 4.839038e-09

GCV score: 0.005940881
```

▶ ... no signal detected ⇒ k large enough for Girth! (select allows 0 model to be selected, -1 suppresses intercept).

Is smoothness selection robust?

➤ To check robustness of smoothness selection, fit with an alternative smoothness selection criterion. e.g.

```
> ct1 <- gam(Volume~s(Height)+s(Girth),
+ family=Gamma(link=log),data=trees,method="ML")
> ct1
...
Formula:
Volume~s(Height) + s(Girth)

Estimated degrees of freedom:
1.0000 2.5097 total = 4.509708
ML score: 69.64346
```

- ... EDFs much the same.
- ► A very different criterion is best (i.e. "ML" or "REML" versus "GCV.Cp" or "GACV.Cp").
- Remember:likelihood based methods tend to be more robust.

summary.gam

Once checking suggests that the model is acceptable, then we can proceed to more formal inference. e.g.

```
> summary(ct1)
  Parametric coefficients:
             Estimate Std. Error t value Pr(>|t|)
  (Intercept) 3.27568 0.01493 219.4 <2e-16 ***
  Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
  Approximate significance of smooth terms:
            edf Ref.df F p-value
  s(Height) 1.00 1.000 31.04 7.04e-06 ***
  s(Girth) 2.51 3.150 211.93 < 2e-16 ***
  Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
  R-sq.(adj) = 0.974 Deviance explained = 97.8\%
  ML score = 69.643 Scale est. = 0.0069092 n = 31
```

anova.gam

If you must have p-values, then anova is better for any model containing factor variables

$\hat{oldsymbol{eta}}$: coef and vcov

▶ The estimated coefficients, $\hat{\beta}$, and the Bayesian covariance matrix for $\beta|\mathbf{y}$...

```
> b <- coef(ct1) ## estimated coefficients

> V <- vcov(ct1) ## Bayesian cov matrix

> b[1:4] ## first 4 coefs

(Intercept) s(Height).1 s(Height).2 s(Height).3

3.275684e+00 -1.449735e-07 1.936084e-07 -8.372687e-08

> V[1:4,1:4] (Intercept) s(Height).1 s(Height).2 s(Height).3

(Intercept) 2.228777e-04 1.938385e-23 -5.369032e-23 7.371339e-24

s(Height).1 1.938385e-23 1.172984e-08 1.944762e-09 -1.629145e-09

s(Height).2 -5.369032e-23 1.944762e-09 2.797615e-08 -6.769112e-09

s(Height).3 7.371339e-24 -1.629145e-09 -6.769112e-09 3.914123e-09
```

ightharpoonup vcov(ct1,freq=TRUE) extracts the frequentist covariance matrix for \hat{eta} .

The smoothing parameters λ

▶ We can extract the λ_i estimates, and, for RE/ML smoothness selection, the covariance matrix of the log λ_i estimates.

 Alternatively, extract information as (root) variance components . . .

```
> gam.vcomp(ct1)
```

Standard deviations and 0.95 confidence intervals:

```
        std.dev
        lower
        upper

        s(Height)
        1.175867e-05
        2.267390e-86
        6.098041e+75

        s(Girth)
        1.748033e-02
        5.683986e-03
        5.375839e-02

        scale
        7.969330e-02
        6.129519e-02
        1.036137e-01
```

Rank: 3/3

Predicting at new covariate values: predict.gam

- predict(ct1) returns the linear predictor corresponding to the original data.
- predict.gam's main use is to predict from the model, given new values for the predictor variables...

predict has several useful arguments. e.g.

More predict.gam

Predictions can also be returned by model term...

...or on the response scale: predict(ct1,pd,type="response").

Linear prediction matrices: lpmatrix

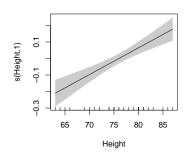
predict.gam can return the matrix mapping the estimated coefficients to the linear predictor. e.g.

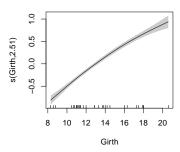
- ► Obviously predict(ct1,type="lpmatrix") returns the original model matrix used for fitting (take a look at the code for model.matrix.gam sometime).
- Linear predictor matrices are the key to inference about *any* quantity that can be predicted by the model. More later.

More visualization: customizing plot.gam

- ▶ plot.gam has quite a few options. Here are three examples
 - 1. The component-wise CIs have good coverage, except when smooths are close to straight lines. A solution is to include the uncertainty in the intercept when plotting.
 - 2. Each smooth can have its own y axis scale.
 - 3. Some people like their confidence regions to be shaded...

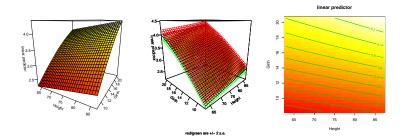
```
par(mfrow=c(1,2))
plot(ct1,shade=TRUE,seWithMean=TRUE,scale=0)
```





More visualization: vis.gam

- ➤ Sometimes it is helpful to see how the linear predictor or expected response would vary with 2 predictors, if all the others were held fixed at some value. vis.gam allows this.
- vis.gam(ct1,theta=30,ticktype="detailed")
 vis.gam(ct1,theta=-45,ticktype="detailed",se=2)
 vis.gam(ct1,plot.type="contour")



Alternative tree models

Would a 2D smooth of Height and Girth be better?

- ▶ ...apparently not. Model has more degrees of freedom for a lower marginal likelihood, and the AIC is worse than before.
- ▶ But an isotropic smooth is not really appropriate here.

A tensor product smooth for trees

```
> ct2 <- gam(Volume~te(Height,Girth),</p>
                family=Gamma(link=log),data=trees,method="ML")
  +
  > ct1;ct2
  Estimated degrees of freedom:
  1.0000 \ 2.5097 \ \text{total} = 4.509708
  ML score: 69.64346
   . . .
  Estimated degrees of freedom:
  3 \text{ total} = 4.000024
  ML score: 66.71353
  > AIC(ct1,ct2)
             df
                   ATC
  ct.1 5.509708 142.8381
  ct2 5.000024 143.4271
  > 2*66.71353+10 ## "ATC" ct2
   [1] 143.4271
  > 2*69.64346+8 ## "AIC" ct1
   [1] 147.2869
this is weird ....
```

Which AIC is correct?

- ▶ AIC(ct1,ct2) selects the additive model, ct1. But ct1 has lower marginal likelihood and more degrees of freedom, implying a higher AIC. Consider the two AIC computations...
 - 1. AIC(ct1,ct2) is based on the log likelihood of $\hat{\beta}$ treated as fixed effects, but with the model EDF in place of dim(β).
 - Alternatively, base AIC on the marginal likelihood of the fixed parameters of the model with the random effect components of the smooths integrated out. Then the appropriate degrees of freedom is the number of fixed effects.
- ▶ Why do the 2 give different answers and which is 'right'?
- ▶ The difference lies in the model assumed. For option 1, we are being Bayesian, and would expect broadly the same smooths on resampling of the data. For option 2 we are implicitly being fully frequentist and assuming that the smooths would look completely different on resampling.
- ▶ So, use option 2 if *really* believe its model. I don't.

Other alternative models

Having rejected a 2D smooth model, let's consider a couple of simple alternatives, based on discretizing Height.

```
trees$Hclass <- factor(floor(trees$Height/10)-5,</pre>
                             labels=c("Small","Medium","Large"))
ct3 <- gam(Volume~s(Girth)+Hclass,
             family=Gamma(link=log),data=trees,method="ML")
par(mfrow=c(1,2))
plot(ct3,all.terms=TRUE)
                                           Partial for Hclass
         s(Girth, 2.49)
                                              0.2
             0.0
                                              0.1
             -0.5
                          Girth
                                                           Helass
```

Again AIC etc, suggest that this is no improvement.

A random effect for Height

▶ Finally try $\log \mu_i = f(Girth_i) + b_k$ if tree i is in height class k. The b_k are i.i.d. $N(0, \sigma_b^2)$ and $Volume_i \sim Gamma(\mu_i, \phi)$. > ct4 <- gam(Volume~s(Girth)+s(Hclass,bs="re"),</pre> family=Gamma(link=log),data=trees,method="ML") > ct4 Estimated degrees of freedom: $2.4876 \ 1.6294 \ \text{total} = 5.117026$ ML score: 79.12945 > gam.vcomp(ct4) Standard deviations and 0.95 confidence intervals: std.dev lower upper s(Girth) 0.02108980 0.00615994 0.0722052 s(Hclass) 0.07889864 0.02726095 0.2283484 scale 0.09911287 0.07492786 0.1311042

Again AIC etc suggest that this is a worse model.

Summary

- gam is like glm, but with extra facilities to allow smooth functions of covariates in the linear predictor.
- Residual checking is as for a GLM, but plot.gam plots the smooths, not residual plots.
- Additional model checking should check the smoothing basis dimension choice, and the reasonableness of the smoothness selection.
- Sometimes model selection is similar to model selection for ordinary GLMs.