

# CFDS® – Chartered Financial Data Scientist Introduction to Python

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# 4 Financial Time Series

- Time series are ubiquitous in finance.
- pandas is the main library in Python to deal with time series.

# 4.1 Financial Data

#### **Financial data**

- For the time being we work with locally stored data files.
- These are in .csv -files (comma-separated values), where the data entries in each row are separated by commas.
- · Some initialisation:

#### In [ ]:

```
import matplotlib.pyplot as plt
import seaborn as sns
```

## **Data import**

- pandas provides a numer of different functions and DataFrame methods for importing and exporting data.
- Here we use pd.read csv().
- The file that we load contains end-of-day data for different financial instruments retrieved from Thomson Reuters.

```
# If using colab, then uncomment the line below and comment the line after that
#filename = 'https://raw.githubusercontent.com/packham/Python_CFDS/main/data/tr_eikof
filename = './data/tr_eikon_eod_data.csv' # path and filename
f = open(filename, 'r') # this will give an error when using colab; just ignore it
f.readlines()[:5] # show first five lines
```

## **Data import**

```
In [ ]:
```

```
In [ ]:
```

```
data.info() # information about the DataFrame object
```

# Data import

```
In [ ]:
```

```
data.head()
```

## **Data import**

```
In [ ]:
```

```
data.tail()
```

# **Data import**

```
In [ ]:
```

```
data.plot(figsize=(10, 10), subplots=True);
```

### **Data import**

- The identifiers used by Thomson Reuters are so-called RIC's.
- The financial instruments in the data set are:

## **Data import**

#### In [ ]:

```
for ric, name in zip(data.columns, instruments):
    print('{:8s} | {}'.format(ric, name))
```

## **Summary statistics**

#### In [ ]:

```
data.describe().round(2)
```

# **Summary statistics**

• The aggregate() -function allows to customise the statistics viewed:

#### In [ ]:

#### Returns

- When working with financial data we typically (=always you must have good reasons to deviate from this) work with performance data, i.e., **returns**.
- · Reasoning:
  - Historical data are mainly used to make forecasts one or several time periods forward.
  - The daily average stock price over the last eight years is meaningless to make a forecast for tomorrow's stock price.
  - However, the daily returns are possible scenarios for the next time period(s).
- The function pct\_change() calculates discrete returns:

$$r_t^{\rm d} = \frac{S_t - S_{t-1}}{S_{t-1}},$$

where  $S_t$  denotes the stock price at time t.

#### Returns

data.pct\_change().round(3).head()

#### Returns

#### In [ ]:

```
data.pct change().mean().plot(kind='bar', figsize=(10, 6));
```

#### **Returns**

- In finance, **log-returns**, also called **continuous returns**, are often preferred over discrete returns:  $r_t^c = \ln\left(\frac{S_t}{S_{t-1}}\right)$ .
- The main reason is that log-return are additive over time.
- For example, the log-return from t-1 to t+1 is the sum of the single-period log-returns:

$$r_{t-1,t+1}^{c} = \ln\left(\frac{S_{t+1}}{S_{t}}\right) + \ln\left(\frac{S_{t}}{S_{t-1}}\right) = \ln\left(\frac{S_{t+1}}{S_{t}} \cdot \frac{S_{t}}{S_{t-1}}\right) = \ln\left(\frac{S_{t+1}}{S_{t-1}}\right).$$

• Note: If the sampling (time) interval is small (e.g. one day or one week), then the difference between discrete returns and log-returns is negligible.

#### Returns

```
In [ ]:
```

```
rets = np.log(data / data.shift(1)) # calculates log-returns in a vectorised way
```

In [ ]:

```
rets.head().round(3)
```

#### Returns

```
In [ ]:
```

```
rets.cumsum().apply(np.exp).plot(figsize=(10, 6)); # recover price paths from log-1
```

# 4.2 Correlation analysis and linear regression

- To further illustrate how to work with financial time series we consider the S&P 500 stock index and the VIX volatility index.
- Empirical stylised fact: As the S&P 500 rises, the VIX falls, and vice versa.
- Note: This is about correlation not causation.

## **Correlation analysis**

```
In [ ]:
```

```
# EOD data from Thomson Reuters Eikon Data API

# If using colab, then uncomment the line below and comment the line after that
#raw = pd.read_csv('https://raw.githubusercontent.com/packham/Python_CFDS/main/data/
raw = pd.read_csv('./data/tr_eikon_eod_data.csv', index_col=0, parse_dates=True)
data = raw[['.SPX', '.VIX']].dropna()
data.tail()
```

# **Correlation analysis**

```
In [ ]:
```

```
data.plot(subplots=True, figsize=(10, 6));
```

# **Correlation analysis**

• Transform both data series into log-returns:

```
In [ ]:
```

```
rets = np.log(data / data.shift(1))
rets.head()
```

```
In [ ]:
```

```
rets.dropna(inplace=True) # drop NaN (not-a-number) entries
```

# **Correlation analysis**

```
In [ ]:
```

```
rets.plot(subplots=True, figsize=(10, 6));
```

# **Correlation analysis**

```
In [ ]:
```

# **Correlation analysis**

```
In [ ]:
```

```
rets.corr()
```

## **OLS** regression

- Linear regression captures the linear relationship between two variables.
- For two variables x, y, we postulate a linear relationship:

$$y = \alpha + \beta x + \varepsilon, \quad \alpha, \beta \in \mathbb{R}.$$

- Here,  $\alpha$  is the intercept,  $\beta$  is the slope (coefficient) and  $\varepsilon$  is the error term.
- Given data sample of joint observations  $(x_1, y_1), \dots, (x_n, y_n)$ , we set

$$y_i = \hat{\alpha} + \hat{\beta}x_i + \hat{\varepsilon}_i$$

where  $\hat{\alpha}$  and  $\hat{\beta}$  are estimates of  $\alpha, \beta$  and  $\hat{\varepsilon}_1, \dots, \hat{\varepsilon}_n$  are the so-called **residuals**.

• The **ordinary least squares (OLS)** estimator  $\hat{\alpha}$ ,  $\hat{\beta}$  corresponds to those values of  $\alpha$ ,  $\beta$  that minimise the sum of squared residuals:

$$\min_{\alpha,\beta} \sum_{i=1}^{n} \varepsilon_i^2 = \sum_{i=1}^{n} (y_i - \alpha - \beta x_i)^2.$$

## **OLS** regressions

· Simplest form of OLS regression:

```
In [ ]:
```

```
reg = np.polyfit(rets['.SPX'], rets['.VIX'], deg=1) # fit a linear equation (a poly
reg.view() # the fitted paramters
```

```
2.62e-03 is scientific notation: 2.62e - 03 = 2.62 \cdot 10^{-3}.
```

```
In [ ]:
```

```
ax = rets.plot(kind='scatter', x='.SPX', y='.VIX', figsize=(8, 5))
ax.plot(rets['.SPX'], np.polyval(reg, rets['.SPX']), 'r', lw=2);
```

# **OLS** regression

• To do a more refined OLS regression with a proper analysis, use the package statsmodels.

```
In [ ]:
```

```
import statsmodels.api as sm

Y=rets['.VIX']
X=rets['.SPX']
X = sm.add_constant(X)
```

```
In [ ]:
```

```
model = sm.OLS(Y, X)
results = model.fit()
```

```
In [ ]:
```

```
results.params
```

results.predict()[0:10]

## **OLS** regression

#### In [ ]:

print(results.summary())

## **OLS regression: Interpretation of output and forecasting**

- The column <code>coef</code> lists the coefficients of the regression: the coefficient in the row labelled <code>const</code> corresponds to  $\hat{\alpha}$  (= 0.0026) and the coefficient in the row <code>.spx</code> denotes  $\hat{\beta}$  (= -6.6515).
- The estimated model in the example is thus:

$$.VIX = 0.0026 - 6.6516.SPX.$$

• The best forecast of the VIX return when observing an S&P return of 2% is therefore  $0.0026-6.6516\cdot0.02=-0.130432=-13.0432\%$ .

# OLS regression: Validation ( $R^2$ )

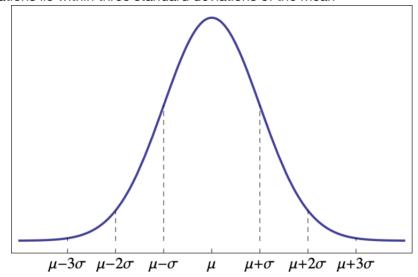
- To **validate** the model, i.e., to determine, if the model in itself and the explanatory variable(s) make sense, we look  $R^2$  and various p-values (or confidence intervals or t-statistics).
- $R^2$  measures the fraction of variance in the dependent variable Y that is captured by the regression line;  $1 R^2$  is the fraction of Y-variance that remaines in the residuals  $\varepsilon_i^2$ ,  $i = 1, \ldots, n$
- In the output above  $\mathbb{R}^2$  is given as 0.647. In other words, 64.7% of the variance in VIX returns are "explained" by SPX returns.
- A high  $\mathbb{R}^2$  (and this one is high) is necessary for making forecasts.

# OLS regression: Validation (confidence interval)

- An important hypothesis to test in any regression model is whether the explanatory variable(s) have an effect on the independent variable.
- This can be translated into testing whether  $\beta \neq 0$ . ( $\beta = 0$  is the same as saying that the X variable can be removed from the model.)
- Formally, we test the null hypothesis  $H_0: \beta = 0$  against the alternative hypothesis  $H_1: \beta \neq 0$ .
- There are several statistics to come to the same conclusion: confidence intervals, *t*-statistics and *p*-values.
- The **confidence interval** is an interval around the estimate  $\hat{\beta}$  that we are confident contains the true parameter  $\beta$ . A typial **confidence level** is 95%.
- If the 95% confidence interval does **not** contain 0, then we say  $\beta$  is **statistically significant** at the 5% (=1-95%) level, and we conclude that  $\beta \neq 0$ .

# **OLS regression: Validation (***t***-statistic)**

- The t-statistic corresponds to the **number of standard deviations** that the estimated coefficient  $\hat{\beta}$  is away from 0 (the mean under  $H_0$ ).
- For a normal distribution, we have the following rules of thumb:
  - 66% of observations lie within one standard deviation of the mean
  - 95% of observations lie within two standard deviations of the mean
  - 99.7% of observations lie within three standard deviations of the mean



- If the sample size is large enough, then the t-statistic is approximately normally distributed, and if it is large (in absolute terms), then this is an indication against  $\beta = 0$ .
- In the example above, the t-statistics is -62.559, i.e.,  $\hat{\beta}$  is approx. 63 standard deviations away from zero, which is practically impossible.

# **OLS** regression: Validation (*p*-value)

- The p-value expresses the probability of observing a coefficient estimate as extreme (away from zero) as  $\hat{\beta}$  under  $H_0$ , i.e., when  $\beta = 0$ .
- In other words, it measures the probability of observing a t-statistic as extreme as the one observed if  $\beta = 0$ .
- If the p-value (column |P| | t|) is smaller than the desired level of significance (typically 5%), then the  $H_0$  can be rejected and we conclude that  $\beta \neq 0$ .
- In the example above, the p-value is given as 0.000, i.e., it is so small, that we can conclude the estimated coefficient  $\hat{\beta}$  is so extreme (= away from zero) that is virtually impossible to obtain such an estimated if  $\beta = 0$ .
- Finally, the F-test tests the hypotheses  $H_0: R^2 = 0$  versus  $H_1: R^2 \neq 0$ . In a multiple regression with k independent variables, this is equivalent to  $H_0: \beta_1 = \cdots = \beta_k = 0$ .
- In the example above, the p-value of the F-test is 0, so we conclude that the model overall has explanatory power.