TIME SERIES FORECASTING BUSINESS REPORT – ROSE WINE SALES

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PROBLEM 2

For this particular assignment, the data of different types of wine sales in the 20th century is to be analysed. Both of these data are from the same company but of different wines (Sparkling & Rose). As an analyst in the ABC Estate Wines, you are tasked to analyse and forecast Rose Wine Sales in the 20th century.

2.0. Read the data as an appropriate Time Series data and plot the data

Dataset Head:

| | YearMonth | Rose |
|---|-----------|-------|
| 0 | 1980-01 | 112.0 |
| 1 | 1980-02 | 118.0 |
| 2 | 1980-03 | 129.0 |
| 3 | 1980-04 | 99.0 |
| 4 | 1980-05 | 116.0 |

Dataset Tail:

| | YearMonth | Rose |
|-----|-----------|------|
| 182 | 1995-03 | 45.0 |
| 183 | 1995-04 | 52.0 |
| 184 | 1995-05 | 28.0 |
| 185 | 1995-06 | 40.0 |
| 186 | 1995-07 | 62.0 |

The dataset contents 187 observations across 02 columns in total.

Date Time Index:

```
DatetimeIndex(['1980-01-31', '1980-02-29', '1980-03-31', '1980-04-30', '1980-05-31', '1980-06-30', '1980-07-31', '1980-08-31', '1980-09-30', '1980-10-31', ...
'1994-10-31', '1994-11-30', '1994-12-31', '1995-01-31', '1995-02-28', '1995-03-31', '1995-04-30', '1995-05-31', '1995-06-30', '1995-07-31'], dtype='datetime64[ns]', length=187, freq='M')
```

Time Stamp:

Rose

| Time_Stamp | | | | |
|------------|-------|--|--|--|
| 1980-01-31 | 112.0 | | | |
| 1980-02-29 | 118.0 | | | |
| 1980-03-31 | 129.0 | | | |
| 1980-04-30 | 99.0 | | | |
| 1980-05-31 | 116.0 | | | |

We do not require the column of "YearMonth" as we have created a Time Stamp for the same and made it as our index column as well. Hence, we have dropped "YearMonth" from our dataset.

Missing Values:

YearMonth 0 Rose 2 dtype: int64

| | YearMonth | Rose |
|-----|-----------|------|
| 174 | 1994-07 | NaN |
| 175 | 1994-08 | NaN |

We observe that there a 2 months of July and August 1994 where the Rose Wine sales figures are missing.

Dataset Description:

| | Rose |
|-------|------------|
| count | 185.000000 |
| mean | 90.394595 |
| std | 39.175344 |
| min | 28.000000 |
| 25% | 63.000000 |
| 50% | 86.000000 |
| 75% | 112.000000 |
| max | 267.000000 |

We observe from a historical record of 187 months since Jan 1980 until July 1995 that average sales over the period of Rose wine was 90 bottles. The least being 28 bottles and highest being 267 bottles. The data however is missing the sales for the 2 months of July and August 1994 which we will fill accurately through interpolate function later. Now, we have our data ready for the Time Series Analysis.

Time Series Plot

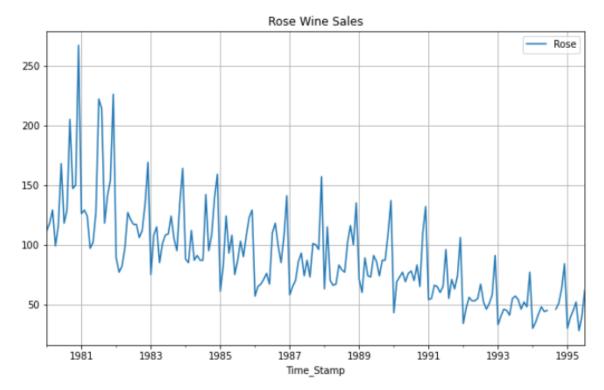


Figure 1: Time Series Plot

For above figure, we observe presence of trend and seasonality throughout the time series. The trend is downward indicating that the sales figures of the Rose Wine have been dropping over the years. We can also observe a slight cut between the plot line between the years 1994 and 1995 due to the missing values.

2.1 Perform appropriate Exploratory Data Analysis to understand the data and also perform decomposition.

Yearly Boxplot:

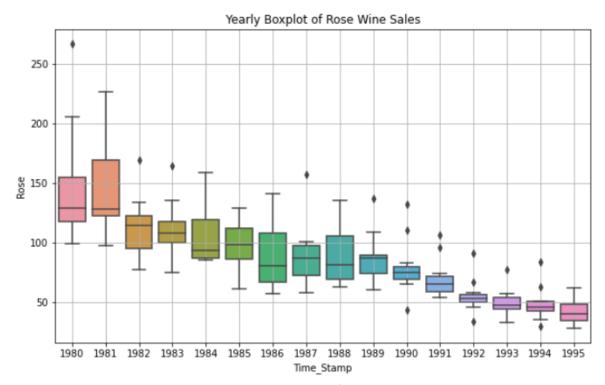


Figure 2 – Yearly Boxplot of Rose Wine Sales

The sales figures can be seen to pretty much ranges between 25 bottles to 175 bottles from 1982 till 1995 with outlier present in some years. The sales figures were higher in 1980 and 1981 after which there has been a significant drop in the sales figures. The sales have been on a continuous drop in most years dropping to the lowest in 1995 which can be clearly seen from the downward trend.

Monthly Boxplot:

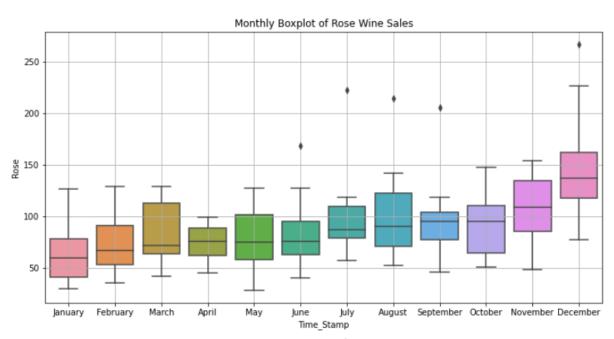


Figure 3 – Monthly Boxplot of Rose Wine Sales

The sales have been better in the 2nd half of the year throughout the time series especially from July onwards. The least sales have come in the month of April and highest from December throughout various years!

Outliers are present for June, July, August, September and December months.

Monthly Boxplot with Median Values as Red Line:

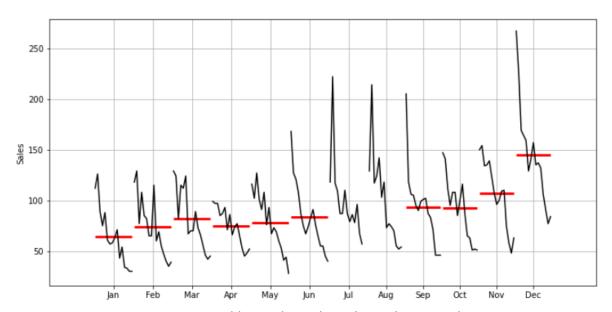


Figure 4 - Monthly Boxplot with Median Values as Red Line

The median values are within 50 and 100 bottles until October and 100 to 150 bottles for November and December. There is actually a huge spike in the median value for December compared to November . Also, we can see that the median values for July and August are not seen as there are missing values for these months in the time series dataset.

Pivot Table Monthly Sales:

| Time_Stamp | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
|------------|-------|-------|--------|--------|----------|-----------|---------|---------|-------|-------|-------|-------|
| Time_Stamp | | | | | | | | | | | | |
| 1980 | 112.0 | 118.0 | 129.0 | 99.0 | 116.0 | 168.0 | 118.0 | 129.0 | 205.0 | 147.0 | 150.0 | 267.0 |
| 1981 | 126.0 | 129.0 | 124.0 | 97.0 | 102.0 | 127.0 | 222.0 | 214.0 | 118.0 | 141.0 | 154.0 | 226.0 |
| 1982 | 89.0 | 77.0 | 82.0 | 97.0 | 127.0 | 121.0 | 117.0 | 117.0 | 106.0 | 112.0 | 134.0 | 169.0 |
| 1983 | 75.0 | 108.0 | 115.0 | 85.0 | 101.0 | 108.0 | 109.0 | 124.0 | 105.0 | 95.0 | 135.0 | 164.0 |
| 1984 | 88.0 | 85.0 | 112.0 | 87.0 | 91.0 | 87.0 | 87.0 | 142.0 | 95.0 | 108.0 | 139.0 | 159.0 |
| 1985 | 61.0 | 82.0 | 124.0 | 93.0 | 108.0 | 75.0 | 87.0 | 103.0 | 90.0 | 108.0 | 123.0 | 129.0 |
| 1986 | 57.0 | 65.0 | 67.0 | 71.0 | 76.0 | 67.0 | 110.0 | 118.0 | 99.0 | 85.0 | 107.0 | 141.0 |
| 1987 | 58.0 | 65.0 | 70.0 | 86.0 | 93.0 | 74.0 | 87.0 | 73.0 | 101.0 | 100.0 | 96.0 | 157.0 |
| 1988 | 63.0 | 115.0 | 70.0 | 66.0 | 67.0 | 83.0 | 79.0 | 77.0 | 102.0 | 116.0 | 100.0 | 135.0 |
| 1989 | 71.0 | 60.0 | 89.0 | 74.0 | 73.0 | 91.0 | 86.0 | 74.0 | 87.0 | 87.0 | 109.0 | 137.0 |
| 1990 | 43.0 | 69.0 | 73.0 | 77.0 | 69.0 | 76.0 | 78.0 | 70.0 | 83.0 | 65.0 | 110.0 | 132.0 |
| 1991 | 54.0 | 55.0 | 66.0 | 65.0 | 60.0 | 65.0 | 96.0 | 55.0 | 71.0 | 63.0 | 74.0 | 106.0 |
| 1992 | 34.0 | 47.0 | 56.0 | 53.0 | 53.0 | 55.0 | 67.0 | 52.0 | 46.0 | 51.0 | 58.0 | 91.0 |
| 1993 | 33.0 | 40.0 | 46.0 | 45.0 | 41.0 | 55.0 | 57.0 | 54.0 | 46.0 | 52.0 | 48.0 | 77.0 |
| 1994 | 30.0 | 35.0 | 42.0 | 48.0 | 44.0 | 45.0 | NaN | NaN | 46.0 | 51.0 | 63.0 | 84.0 |
| 1995 | 30.0 | 39.0 | 45.0 | 52.0 | 28.0 | 40.0 | 62.0 | NaN | NaN | NaN | NaN | NaN |
| | | | Figure | 5 – Pi | ivot Tab | ole of Ro | ose Wir | e Sales | | | | |

Monthly Sales across Years:

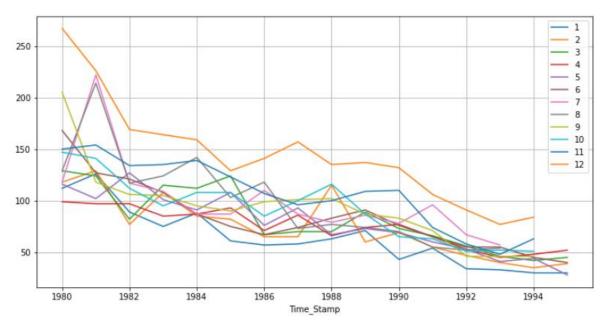


Figure 6 –Rose Wine Monthly Sales across Years

December has the highest sales across all years

Empirical Cumulative Distribution:

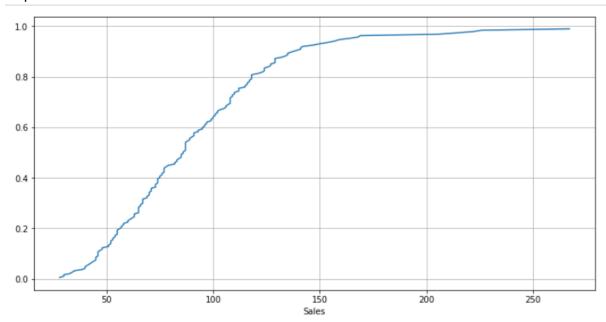


Figure 7 –Rose Wine Sales Empirical Cumulative Distribution Plot

Average Sales and Percentage Change of Sales:

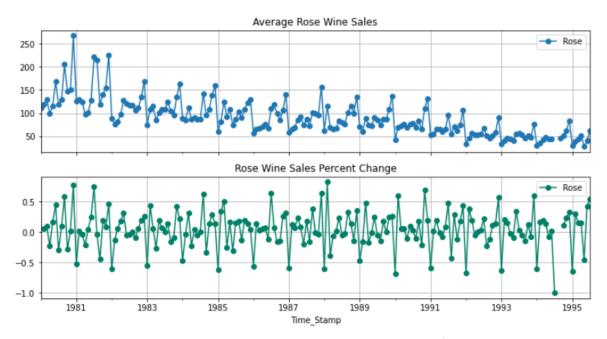


Figure 8 – Average Sales and Percentage Change of Sales

Sum of Sales of each year in Plot:

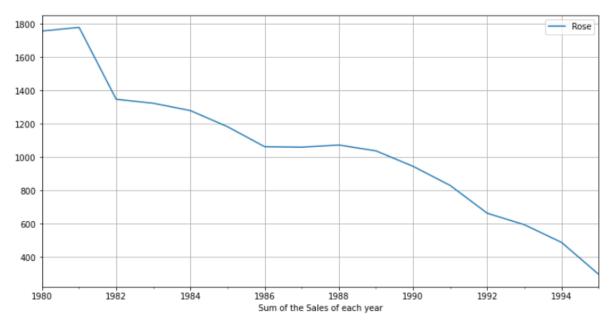


Figure 9 –Sum of Rose Wine Sales for each Year

The sales have dropped over the years, only exception being 1981 where they have slightly increased.

Sum of Sales of each year in Figures:

| | Rose |
|------------|--------|
| Time_Stamp | |
| 1980-12-31 | 1758.0 |
| 1981-12-31 | 1780.0 |
| 1982-12-31 | 1348.0 |
| 1983-12-31 | 1324.0 |
| 1984-12-31 | 1280.0 |
| 1985-12-31 | 1183.0 |
| 1986-12-31 | 1063.0 |
| 1987-12-31 | 1060.0 |
| 1988-12-31 | 1073.0 |
| 1989-12-31 | 1038.0 |
| 1990-12-31 | 945.0 |
| 1991-12-31 | 830.0 |
| 1992-12-31 | 663.0 |
| 1993-12-31 | 594.0 |
| 1994-12-31 | 488.0 |
| 1995-12-31 | 296.0 |

The highest sales were recorded in the year 1981 with total 1780 bottles sold and the least sales were for 1995 with total 296 bottles sold within the first seven months. On a 12 monthly basis, the least sales can be seen for the year 1994 with 488 bottles sold in total.

Mean of Sales of each year in Plot:

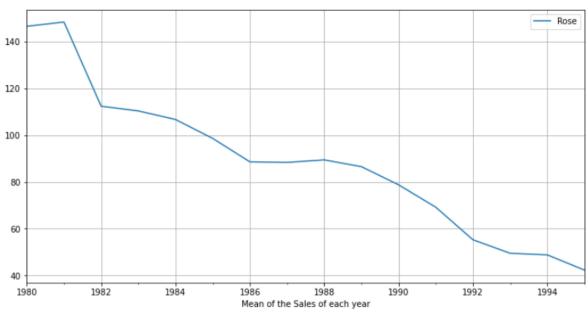


Figure 10 – Mean of Sparkling Wine Sales for each Year

Mean of Sales of each year in Figures:

| | Rose |
|------------|------------|
| Time_Stamp | |
| 1980-12-31 | 146.500000 |
| 1981-12-31 | 148.333333 |
| 1982-12-31 | 112.333333 |
| 1983-12-31 | 110.333333 |
| 1984-12-31 | 106.666667 |
| 1985-12-31 | 98.583333 |
| 1986-12-31 | 88.583333 |
| 1987-12-31 | 88.333333 |
| 1988-12-31 | 89.416667 |
| 1989-12-31 | 86.500000 |
| 1990-12-31 | 78.750000 |
| 1991-12-31 | 69.166667 |
| 1992-12-31 | 55.250000 |
| 1993-12-31 | 49.500000 |
| 1994-12-31 | 48.800000 |
| 1995-12-31 | 42.285714 |

The mean sales ranged between 42 and 146.5 from 1980 till 1995. With lowest in 1995 and highest in 1981 similar to earlier observation for sum of monthly sales.

Interpolate Missing Values:

| Time_Stamp | |
|------------|------|
| 1994-01-31 | 30.0 |
| 1994-02-28 | 35.0 |
| 1994-03-31 | 42.0 |
| 1994-04-30 | 48.0 |
| 1994-05-31 | 44.0 |
| 1994-06-30 | 45.0 |
| 1994-07-31 | 45.0 |
| 1994-08-31 | 45.0 |
| 1994-09-30 | 46.0 |
| 1994-10-31 | 51.0 |
| 1994-11-30 | 63.0 |
| 1994-12-31 | 84.0 |

Rose

We have interpolated the missing values for July and August 1994 as the same sales figures for June of 45 bottles. Now, that we have interpolated the missing values, we can now decompose the time series.

Decomposing the Time Series: Additive Method

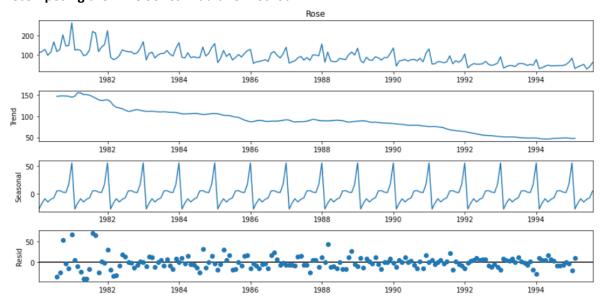


Figure 11 – Additive Decomposition

As per the 'additive' decomposition, we see that there is a pronounced trend until 1991. There is a seasonality as well. A lot of residuals are located around 0 from the plot of the residuals in the decomposition.

Trend, Seasonality and Residual:

| Trend | | | | Residual | |
|--------------|----------------|--------------------------|----------------------|--------------|----------------|
| Time_Stamp | | Seasonality | | Time_Stamp | |
| 1980-01-31 | NaN | Time Stamp | | 1980-01-31 | NaN |
| 1980-02-29 | NaN | 1980-01-31 | -27.908647 | 1980-02-29 | NaN |
| 1980-03-31 | NaN | 1980-02-29 | -17.435632 | 1980-03-31 | NaN |
| 1980-04-30 | NaN | 1980-03-31 | -9.285830 | 1980-04-30 | NaN |
| 1980-05-31 | NaN | 1980-04-30 | -15.098330 | 1980-05-31 | NaN |
| 1980-06-30 | NaN | 1980-05-31 | -10.196544 | 1980-06-30 | NaN |
| 1980-07-31 | 147.083333 | 1980-06-30 | -7.678687 | 1980-07-31 | -33.980241 |
| 1980-08-31 | 148.125000 | 1980-07-31 1980-08-31 | 4.896908 5.499686 | 1980-08-31 | -24.624686 |
| 1980-09-30 | 148.375000 | 1980-09-30 | 2.774686 | 1980-09-30 | 53.850314 |
| 1980-10-31 | 148.083333 | 1980-10-31 | 1.871908 | 1980-10-31 | -2.955241 |
| 1980-11-30 | 147.416667 | 1980-11-30 | 16.846908 | 1980-11-30 | -14.263575 |
| 1980-12-31 | 145.125000 | 1980-12-31 | 55.713575 | 1980-12-31 | 66.161425 |
| Name: trend, | dtype: float64 | Name: season | al, dtype: float64 | Name: resid, | dtype: float64 |

2.2 Split the data into training and test. The test data should start in 1991

Training Data Shape (132, 1) Testing Data Shape (55, 1)

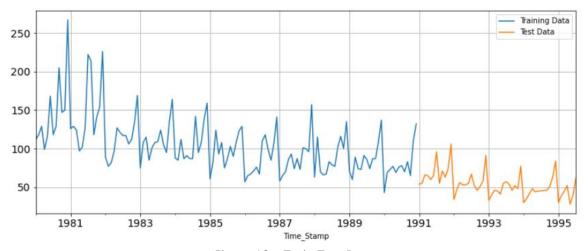


Figure 12 – Train Test Dataset

Training and Test Time Instances:

```
Training Time instance
[1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14, 15, 16, 17, 18, 19, 20, 21, 22, 23, 24, 25, 26, 27, 28, 29, 30, 31, 32, 33, 3
4, 35, 36, 37, 38, 39, 40, 41, 42, 43, 44, 45, 46, 47, 48, 49, 50, 51, 52, 53, 54, 55, 56, 57, 58, 59, 60, 61, 62, 63, 64, 65,
66, 67, 68, 69, 70, 71, 72, 73, 74, 75, 76, 77, 78, 79, 80, 81, 82, 83, 84, 85, 86, 87, 88, 89, 90, 91, 92, 93, 94, 95, 96, 97,
98, 99, 100, 101, 102, 103, 104, 105, 106, 107, 108, 109, 110, 111, 112, 113, 114, 115, 116, 117, 118, 119, 120, 121, 122, 123,
124, 125, 126, 127, 128, 129, 130, 131, 132]
Test Time instance
```

[133, 134, 135, 136, 137, 138, 139, 140, 141, 142, 143, 144, 145, 146, 147, 148, 149, 150, 151, 152, 153, 154, 155, 156, 157, 158, 159, 160, 161, 162, 163, 164, 165, 166, 167, 168, 169, 170, 171, 172, 173, 174, 175, 176, 177, 178, 179, 180, 181, 182, 18 3, 184, 185, 186, 187]

We see that we have successfully the generated the numerical time instance order for both the training and test set.

2.3 Build all the exponential smoothing models on the training data and evaluate the model using RMSE on the test data. Other additional models such as regression, naïve forecast models, simple average models, moving average models should also be built on the training data and check the performance on the test data using RMSE.

MODEL 1 – Linear Regression Model:

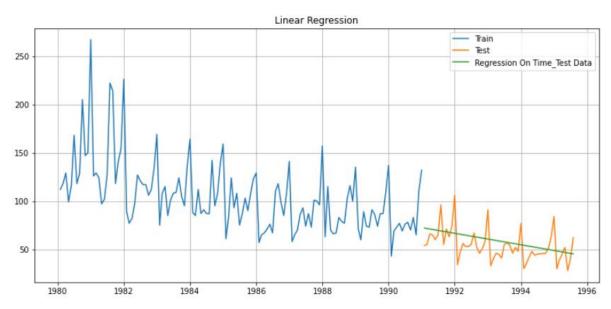


Figure 13 – Linear Regression Plot

| | Test RMSE |
|------------------|-----------|
| RegressionOnTime | 15.268955 |

MODEL 2 – Naïve Model:

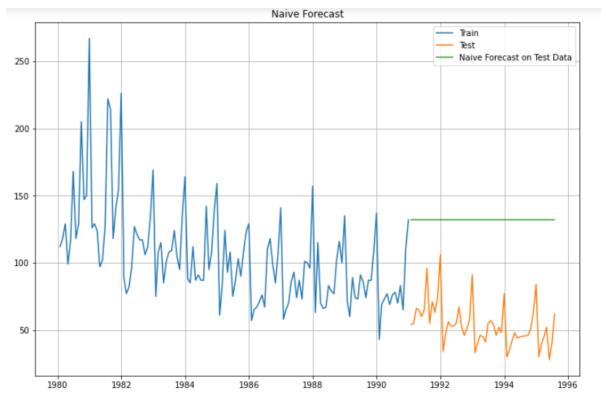


Figure 14 – Naïve Forecast Plot

NaiveModel 79.718773

MODEL 3 – Simple Average Model:

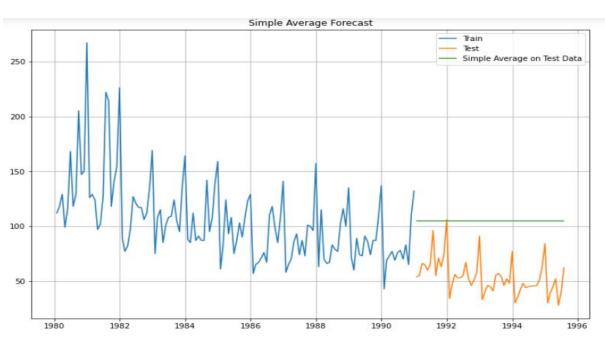


Figure 15 – Simple Average Forecast

53.46057

MODEL 4 – Moving Average Model:

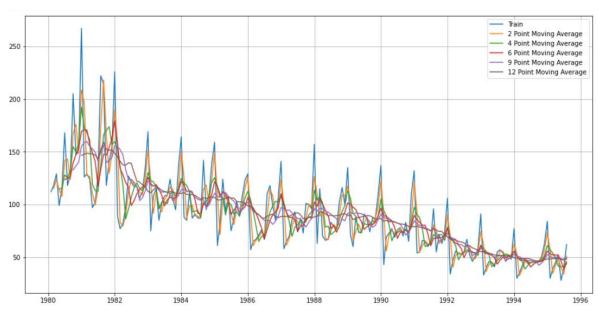


Figure 16 – Point-Wise Moving Average Forecast

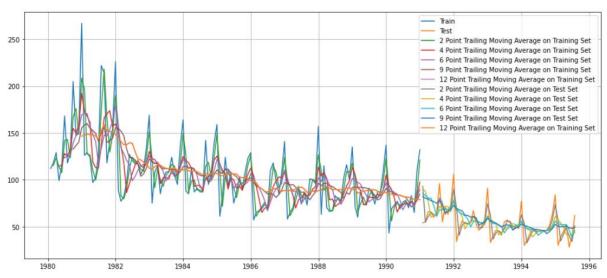


Figure 17 – Point-Wise Trailing Moving Average Forecast

| 2pointTrailingMovingAverage | 11.529278 |
|------------------------------|-----------|
| 4pointTrailingMovingAverage | 14.451403 |
| 6pointTrailingMovingAverage | 14.566327 |
| 9pointTrailingMovingAverage | 14.727630 |
| 12pointTrailingMovingAverage | 15.236052 |

Plotting of all the models:

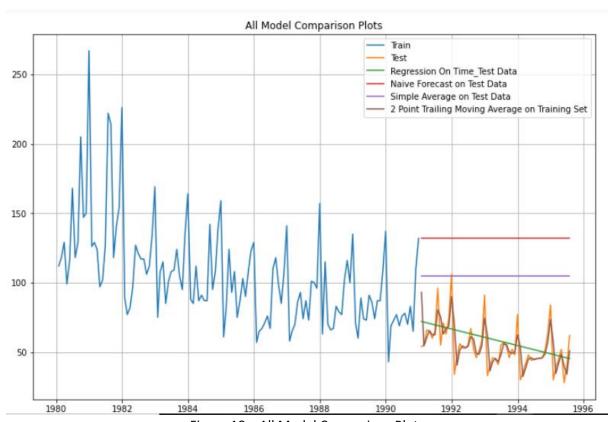


Figure 18 – All Model Comparison Plots

MODEL 5 – Simple Exponential Smoothing Model:

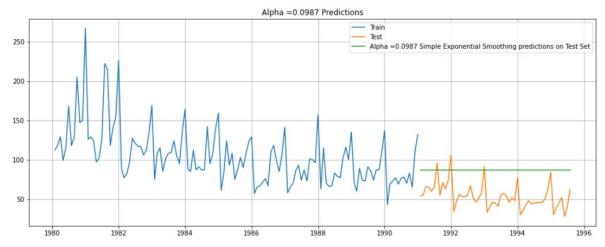


Figure 19 – Simple Exponential Smoothing Model

Alpha=0.0987,SimpleExponentialSmoothing 36.796243

MODEL 6 – Double Exponential Smoothing:

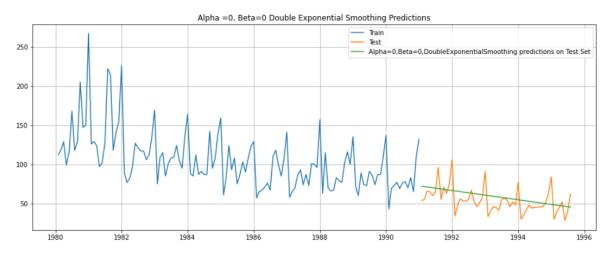


Figure 20 – Double Exponential Smoothing Model

Test RMSE

For Alpha =0, Beta = 0 DoubleExponentialSmoothing 15.268961

MODEL 7 – Triple Exponential Smoothing Additive:

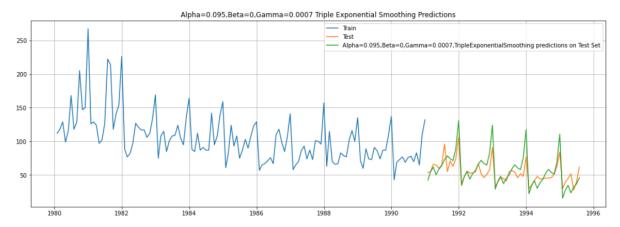


Figure 21 – Triple Exponential Smoothing Additive Model

Alpha=0.095,Beta=0,Gamma=0.0007,TripleExponentialSmoothing 14.176738

MODEL 8 – Triple Exponential Smoothing Multiplicative:

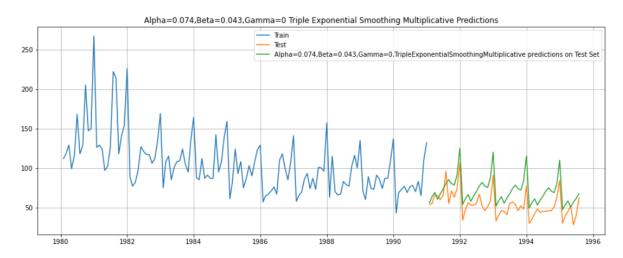


Figure 22 – Triple Exponential Smoothing Multiplicative Model

Test RMSE

Alpha=0.074,Beta=0.043,Gamma=0 TripleExponentialSmoothingMultiplicative

19.741738

Performance using RSME on Test Data:

From the observations so far, we can clearly see that Trailing Moving Average models have performed better than the other models as they have a lower RMSE on the test data. The 2 point Trailing Moving Average Model has performed the best due to its lower RMSE of 11.52.

2.4 Build all the Check for the stationarity of the data on which the model is being built on using appropriate statistical tests and also mention the hypothesis for the statistical test. If the data is found to be non-stationary, take appropriate steps to make it stationary. Check the new data for stationarity and comment.

Note: Stationarity should be checked at alpha = 0.05.

Hypothesis for Statistical Test:

Null Hypothesis – H0 = Time Series is not Stationary Alternative Hypothesis – HA = Time Series is Stationary

Stationarity Check Using Dickey-Fuller Test:

| Results of Dickey-Fuller Test: | |
|--------------------------------|------------|
| Test Statistic | -1.876699 |
| p-value | 0.343101 |
| #Lags Used | 13.000000 |
| Number of Observations Used | 173.000000 |
| Critical Value (1%) | -3.468726 |
| Critical Value (5%) | -2.878396 |
| Critical Value (10%) | -2.575756 |
| dtype: float64 | |

We observe the Time Series is non-stationary for alpha = 0.05 as the p-value is > alpha at 0.34. Hence, we fail to reject the null hypothesis.

Let us take a difference of order 1 and check whether the Time Series is stationary or not.

Stationarity Check Using Dickey-Fuller Test by taking difference of Order 1:

```
Results of Dickey-Fuller Test:
Test Statistic
                             -8.044392e+00
                             1.810895e-12
p-value
                             1.200000e+01
#Lags Used
Number of Observations Used
                            1.730000e+02
                           -3.468726e+00
Critical Value (1%)
Critical Value (5%)
                           -2.878396e+00
Critical Value (10%)
                           -2.575756e+00
dtype: float64
```

We observe the Time Series is now stationary for alpha = 0.05 as the p-value at less than alpha.

2.5 Build an automated version of the ARIMA/SARIMA model in which the parameters are selected using the lowest Akaike Information Criteria (AIC) on the training data and evaluate this model on the test data using RMSE.

MODEL 9 – ARIMA Model:

```
Some parameter combinations for the Model...

Model: (0, 1, 1)

Model: (0, 1, 2)

Model: (1, 1, 0)

Model: (1, 1, 1)

Model: (1, 1, 2)

Model: (2, 1, 0)

Model: (2, 1, 1)

Model: (2, 1, 2)
```

Above is some combination of different parameters of p and q in the range of 0 and 2

ARIMA AIC SCORES for Parameters in range of 0 & 2:

| param | AIC |
|-----------|--|
| (0, 1, 2) | 1279.671529 |
| (1, 1, 2) | 1279.870723 |
| (1, 1, 1) | 1280.574230 |
| (2, 1, 1) | 1281.507862 |
| (2, 1, 2) | 1281.870722 |
| (0, 1, 1) | 1282.309832 |
| (2, 1, 0) | 1298.611034 |
| (1, 1, 0) | 1317.350311 |
| (0, 1, 0) | 1333.154673 |
| | (0, 1, 2) (1, 1, 1) (1, 1, 1) (2, 1, 1) (2, 1, 2) (0, 1, 1) (2, 1, 0) (1, 1, 0) |

We can see the best AIC is for ARIMA (0,1,2) of 1279.67. Below, we will built our ARIMA Model for this parameter and check the performance.

ARIMA (0,1,2) Model Results:

|--|

| Dep. Varia | ble: | Ro | se No. | Observations: | | 132 | |
|------------|---------------|--------------|-------------------------|---------------|---------|----------|------|
| Model: | | ARIMA(0, 1, | Log | Likelihood | | -636.836 | |
| Date: | Wee | d, 02 Mar 20 | 22 AIC | | | 1279.672 | |
| Time: | | 00:08: | 56 BIC | | | 1288.297 | |
| Sample: | | 01-31-19 | 80 HQIC | | | 1283.176 | |
| | | - 12-31-19 | 90 | | | | |
| Covariance | Type: | 0 | pg | | | | |
| | ========= | | ====== | | | | |
| | coef | std err | Z | P> z | [0.025 | 0.975] | |
| | | | | | | | |
| ma.L1 | -0.6970 | 0.072 | -9.689 | 0.000 | -0.838 | -0.556 | |
| ma.L2 | -0.2042 | 0.073 | -2.794 | 0.005 | -0.347 | -0.061 | |
| sigma2 | 965.8407 | 88.305 | 10.938 | 0.000 | 792.766 | 1138.915 | |
| | | | | | | | |
| Ljung-Box | (L1) (Q): | | 0.14 | Jarque-Bera | (JB): | 39 | 9.24 |
| Prob(Q): | | | 0.71 | Prob(JB): | | (| 0.00 |
| Heterosked | asticity (H): | | 0.36 | Skew: | | (| 3.82 |
| Prob(H) (t | wo-sided): | | 0.00 | Kurtosis: | | | 5.13 |
| | | | | | | | |

Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).

Figure 23 – Arima Model

RMSE

ARIMA(0,1,2) 37.30648

MODEL 10 – SARIMA Model with seasonality 6:

```
Examples of some parameter combinations for Model...

Model: (0, 1, 1)(0, 0, 1, 6)

Model: (0, 1, 2)(0, 0, 2, 6)

Model: (1, 1, 0)(1, 0, 0, 6)

Model: (1, 1, 1)(1, 0, 1, 6)

Model: (1, 1, 2)(1, 0, 2, 6)

Model: (2, 1, 0)(2, 0, 0, 6)

Model: (2, 1, 1)(2, 0, 1, 6)

Model: (2, 1, 2)(2, 0, 2, 6)
```

Above is some combination of different parameters of p and q in the range of 0 and 2

SARIMA AIC SCORES for Parameters in range of 0 & 2:

| | param | seasonal | AIC |
|----|-----------|--------------|-------------|
| 53 | (1, 1, 2) | (2, 0, 2, 6) | 1041.655818 |
| 26 | (0, 1, 2) | (2, 0, 2, 6) | 1043.600261 |
| 80 | (2, 1, 2) | (2, 0, 2, 6) | 1045.220389 |
| 71 | (2, 1, 1) | (2, 0, 2, 6) | 1051.673461 |
| 44 | (1, 1, 1) | (2, 0, 2, 6) | 1052.778470 |

We can see the best AIC is for SARIMA (1,1,2) (2,0,2,6) of 1041.65. Below we will built our SARIMA Model for this parameter and check the performance.

SARIMA (1,1,2) (2,0,2,6) Model Results:

| | | | SARIMAX Re | sults | | | |
|-------------|---------|-------------|---------------|---------|---------------|---------|----------|
| Dep. Variab | le: | | | v No. (| Dbservations: | | 13: |
| Model: | | MAX(1, 1, 2 |)x(2, 0, 2, 6 | , | | | -512.828 |
| Date: | | | d, 02 Mar 202 | | | | 1041.656 |
| Time: | | | 00:15:3 | 4 BIC | | | 1063.685 |
| Sample: | | | | 0 HQIC | | | 1050.598 |
| | | | - 13 | 2 | | | |
| Covariance | Type: | | ор | g | | | |
| | | | | | | | |
| | coef | std err | Z | P> z | [0.025 | 0.975] | |
| ar.L1 | -0.5939 | 0.152 | -3.914 | 0.000 | -0.891 | -0.296 | |
| ma.L1 | -0.1954 | 188.566 | -0.001 | 0.999 | -369.777 | 369.387 | |
| ma.L2 | -0.8046 | 151.765 | -0.005 | 0.996 | -298.258 | 296.649 | |
| ar.S.L6 | -0.0625 | 0.035 | -1.794 | 0.073 | -0.131 | 0.006 | |
| ar.S.L12 | 0.8451 | 0.039 | 21.889 | 0.000 | 0.769 | 0.921 | |
| ma.S.L6 | 0.2226 | 188.635 | 0.001 | 0.999 | -369.495 | 369.940 | |
| ma.S.L12 | -0.7774 | 146.598 | -0.005 | 0.996 | -288.104 | 286.549 | |

| sigma2 | 335.1965 | 0.906 | 369.902 | 0.000 | 333.420 | 336.973 |
|-----------|-----------------|-------|---------|-------------|---------|---------|
| | | | | | | |
| Ljung-Box | x (L1) (Q): | | 0.07 | Jarque-Bera | (JB): | 56.68 |
| Prob(Q): | | | 0.78 | Prob(JB): | | 0.00 |
| Heterosk | edasticity (H): | | 0.47 | Skew: | | 0.52 |
| Prob(H) | (two-sided): | | 0.02 | Kurtosis: | | 6.26 |
| | | | | | | |

- [1] Covariance matrix calculated using the outer product of gradients (complex-step).
 [2] Covariance matrix is singular or near-singular, with condition number 1.37e+21. Standard errors may be unstable.

Figure 24 – Sarima Model with seasonality 6

Summary Frame for Alpha = 0.05:

| y | mean | mean_se | mean_ci_lower | mean_ci_upper |
|---|-----------|-----------|---------------|---------------|
| 0 | 62.839941 | 18.848279 | 25.897993 | 99.781888 |
| 1 | 67.629885 | 19.300121 | 29.802343 | 105.457427 |
| 2 | 74.746081 | 19.412682 | 36.697924 | 112.794238 |
| 3 | 71.324859 | 19.475628 | 33.153329 | 109.496389 |
| 4 | 76.016791 | 19.483907 | 37.829034 | 114.204548 |

RMSE

SARIMA(1, 1, 2)(2,0,2,6) 26.134254

MODEL 11 - SARIMA Model with seasonality 12:

Examples of some parameter combinations for Model... Model: (0, 1, 1)(0, 0, 1, 12)

Model: (0, 1, 2)(0, 0, 2, 12) Model: (1, 1, 0)(1, 0, 0, 12)

Model: (1, 1, 1)(1, 0, 1, 12)

Model: (1, 1, 2)(1, 0, 2, 12)

Model: (2, 1, 0)(2, 0, 0, 12) Model: (2, 1, 1)(2, 0, 1, 12)

Model: (2, 1, 2)(2, 0, 2, 12)

Above is some combination of different parameters of p and q in the range of 0 and 2

SARIMA AIC SCORES for Parameters in range of 0 & 2:

| | param | seasonal | AIC |
|----|-----------|---------------|------------|
| 26 | (0, 1, 2) | (2, 0, 2, 12) | 887.937509 |
| 80 | (2, 1, 2) | (2, 0, 2, 12) | 890.668798 |
| 69 | (2, 1, 1) | (2, 0, 0, 12) | 896.518161 |
| 53 | (1, 1, 2) | (2, 0, 2, 12) | 896.686897 |
| 78 | (2, 1, 2) | (2, 0, 0, 12) | 897.346444 |

We can see the best AIC is for SARIMA (0,1,2) (2,0,2,12) of 887.94. Below we will built our SARIMA Model for this parameter and check the performance.

SARIMA (0,1,2) (2,0,2,12) Model Results:

| Model: SARIMAX(0, 1, 2)x(2, 0, 2, 12) Log Likelihood -436.96 Date: Wed, 02 Mar 2022 AIC 887.95 Time: 00:24:11 BIC 906.44 Sample: 0 HQIC 895.43 Covariance Type: opg coef std err z P> z [0.025 0.975] ma.L1 -0.8427 189.512 -0.004 0.996 -372.279 370.593 ma.L2 -0.1573 29.773 -0.005 0.996 -58.512 58.197 ar.S.L12 0.3467 0.079 4.375 0.000 0.191 0.502 ar.S.L24 0.3023 0.076 3.996 0.000 0.154 0.451 ma.S.L24 0.0767 0.133 0.577 0.564 -0.184 0.337 ma.S.L24 -0.0726 0.146 -0.498 0.618 -0.358 0.213 sigma2 251.3136 4.76e+04 0.005 0.996 -9.31e+04 9.36e+04 Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Prob(Q): | | | | SARIMAX | Results | | | |
|---|-------------|-------------|------------|-------------|------------|---------------|----------|----------|
| Date: | Dep. Variab | le: | | | y No. | Observations: | | 13 |
| Date: | Model: | SAR | IMAX(0, 1, | 2)x(2, 0, 2 | , 12) Log | Likelihood | | -436.969 |
| Sample: 0 HQIC 895.43 Covariance Type: opg Coef std err z P> z [0.025 0.975] | Date: | | , , , | | | | | 887.938 |
| - 132 | Time: | | | 00: | 24:11 BIC | | | 906.448 |
| Covariance Type: opg coef std err z P> z [0.025 0.975] ma.L1 -0.8427 189.512 -0.004 0.996 -372.279 370.593 ma.L2 -0.1573 29.773 -0.005 0.996 -58.512 58.197 ar.S.L12 0.3467 0.079 4.375 0.000 0.191 0.502 ar.S.L24 0.3023 0.076 3.996 0.000 0.154 0.451 ma.S.L12 0.0767 0.133 0.577 0.564 -0.184 0.337 ma.S.L24 -0.0726 0.146 -0.498 0.618 -0.358 0.213 sigma2 251.3136 4.76e+04 0.005 0.996 -9.31e+04 9.36e+04 Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Prob(Q): 0.75 Prob(JB): 0.31 | Sample: | | | | | 0 | | 895.437 |
| coef std err z P> z [0.025 0.975] ma.L1 -0.8427 189.512 -0.004 0.996 -372.279 370.593 ma.L2 -0.1573 29.773 -0.005 0.996 -58.512 58.197 ar.S.L12 0.3467 0.079 4.375 0.000 0.191 0.502 ar.S.L24 0.3023 0.076 3.996 0.000 0.154 0.451 ma.S.L12 0.0767 0.133 0.577 0.564 -0.184 0.337 ma.S.L24 -0.0726 0.146 -0.498 0.618 -0.358 0.213 sigma2 251.3136 4.76e+04 0.005 0.996 -9.31e+04 9.36e+04 Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Prob(Q): 0.75 Prob(JB): 0.31 | | | | | - 132 | | | |
| ma.L1 -0.8427 189.512 -0.004 0.996 -372.279 370.593 ma.L2 -0.1573 29.773 -0.005 0.996 -58.512 58.197 ar.S.L12 0.3467 0.079 4.375 0.000 0.191 0.502 ar.S.L24 0.3023 0.076 3.996 0.000 0.154 0.451 ma.S.L12 0.0767 0.133 0.577 0.564 -0.184 0.337 ma.S.L24 -0.0726 0.146 -0.498 0.618 -0.358 0.213 sigma2 251.3136 4.76e+04 0.005 0.996 -9.31e+04 9.36e+04 Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Prob(Q): 0.75 Prob(JB): 0.31 | Covariance | Type: | | | opg | | | |
| ma.L2 | | coef | std err | | P> z | [0.025 | 0.975] | |
| ma.L2 | | | | | | | | |
| ar.S.L12 | | | | | | | | |
| ar.S.L24 0.3023 0.076 3.996 0.000 0.154 0.451 ma.S.L12 0.0767 0.133 0.577 0.564 -0.184 0.337 ma.S.L24 -0.0726 0.146 -0.498 0.618 -0.358 0.213 sigma2 251.3136 4.76e+04 0.005 0.996 -9.31e+04 9.36e+04 Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Prob(Q): 0.75 Prob(JB): 0.31 | | | | | | | | |
| ma.S.L12 | | | | | | | | |
| ma.S.L24 | | | | | | | | |
| sigma2 251.3136 4.76e+04 0.005 0.996 -9.31e+04 9.36e+04 Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Prob(Q): 0.75 Prob(JB): 0.31 | | | | | | | | |
| Ljung-Box (L1) (Q): 0.10 Jarque-Bera (JB): 2.33 Prob(Q): 0.75 Prob(JB): 0.31 | ma.S.L24 | -0.0726 | 0.146 | | | | | |
| Prob(Q): 0.75 Prob(JB): 0.31 | sigma2 | 251.3136 | 4.76e+04 | 0.005 | 0.996 | -9.31e+04 | 9.36e+04 | |
| Prob(Q): 0.75 Prob(JB): 0.31 | | | | | | | | |
| () | Ljung-Box (| (L1) (Q): | | 0.10 | Jarque-Ber | a (JB): | | 2.33 |
| Untranslandantisity (U). 0.99 Charry | Prob(Q): | | | 0.75 | Prob(JB): | | | 0.31 |
| neteroskedasticity (n): 0.00 Skew: 0.57 | Heteroskeda | sticity (H) | : | 0.88 | Skew: | | | 0.37 |
| Prob(H) (two-sided): 0.70 Kurtosis: 3.03 | Prob(H) (tv | ιο-sided): | | 0.70 | Kurtosis: | | | 3.03 |

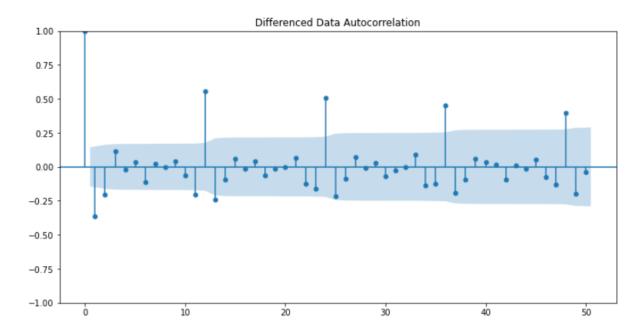
Figure 25 – Sarima Model with Seasonality 12

Summary Frame for Alpha = 0.05:

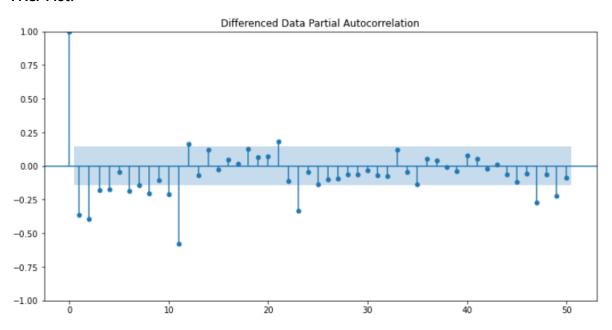
| y | mean | mean_se | mean_ci_lower | mean_ci_upper |
|---|-----------|-----------|---------------|---------------|
| 0 | 62.867261 | 15.928500 | 31.647975 | 94.086547 |
| 1 | 70.541189 | 16.147658 | 38.892362 | 102.190017 |
| 2 | 77.356410 | 16.147655 | 45.707587 | 109.005233 |
| 3 | 76.208813 | 16.147655 | 44.559990 | 107.857637 |
| 4 | 72.747397 | 16.147655 | 41.098574 | 104.396220 |

2.6 Build ARIMA/SARIMA models based on the cut-off points of ACF and PACF on the training data and evaluate this model on the test data using RMSE.

ACF Plot:



PACF Plot:



Here, we have taken alpha=0.05.

- The Auto-Regressive parameter in an ARIMA model is 'p' which comes from the significant lag before which the PACF plot cuts-off to 4.
- The Moving-Average parameter in an ARIMA model is 'q' which comes from the significant lag before the ACF plot cuts-off to 2.

MODEL 12 – Manual ARIMA Model based on cut-Off point (4,1,2):

| SARIMAX Results | | | | | | |
|--|----------|-------------|-------------------------|---------------|---------|----------|
| | | | | | | |
| Dep. Variable | : | Ro | se No. | Observations: | | 132 |
| Model: | Δ | RIMA(4, 1, | Log | Likelihood | | -635.859 |
| Date: | Sun | , 06 Mar 20 | 22 AIC | | | 1285.718 |
| Time: | | 12:51: | 01 BIC | | | 1305.845 |
| Sample: | | 01-31-19 | 80 HQIC | | | 1293.896 |
| • | | - 12-31-19 | 90 | | | |
| Covariance Ty | pe: | o | pg | | | |
| | | | | | | |
| | coef | std err | Z | P> z | [0.025 | 0.975] |
| | | | | | | |
| ar.L1 | -0.3838 | 0.923 | -0.416 | 0.677 | -2.192 | 1.425 |
| ar.L2 | 0.0046 | 0.258 | 0.018 | 0.986 | -0.502 | 0.511 |
| ar.L3 | 0.0414 | 0.113 | 0.366 | 0.714 | -0.180 | 0.263 |
| ar.L4 | -0.0054 | 0.177 | -0.031 | 0.976 | -0.353 | 0.342 |
| ma.L1 | -0.3239 | 0.933 | -0.347 | 0.729 | -2.153 | 1.505 |
| ma.L2 | -0.5407 | 0.874 | -0.619 | 0.536 | -2.254 | 1.172 |
| sigma2 | 951.1524 | 93.870 | 10.133 | 0.000 | 767.170 | 1135.135 |
| - | | | | | | |
| Ljung-Box (L1) (Q): 0.02 Jarque-Bera (JB): | | | | | 32.85 | |
| Prob(Q): | | | 0.88 | Prob(JB): | | 0.00 |
| Heteroskedasticity (H): | | | 0.37 | Skew: | | 0.77 |
| Prob(H) (two-sided): 0.00 Kurtosis: 4.9 | | | | | 4.91 | |
| | | | | | | |

Warnings:

Figure 26 – Arima Model based on ACF and PACF cut-off points

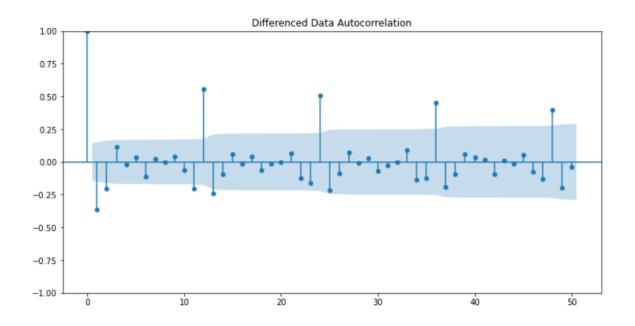
RMSE

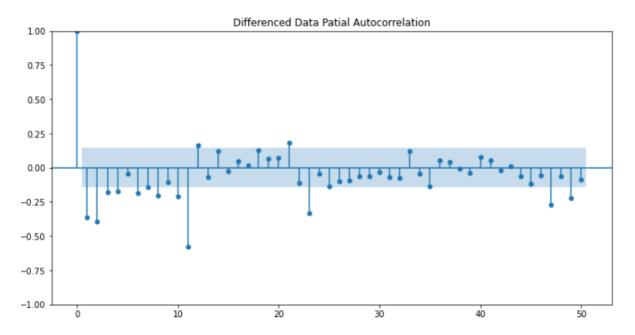
ARIMA(4,1,2) 37.037639

MODEL – 13 Manual SARIMA Model based on cut-Off point (4,1,2):

ACF Plot:

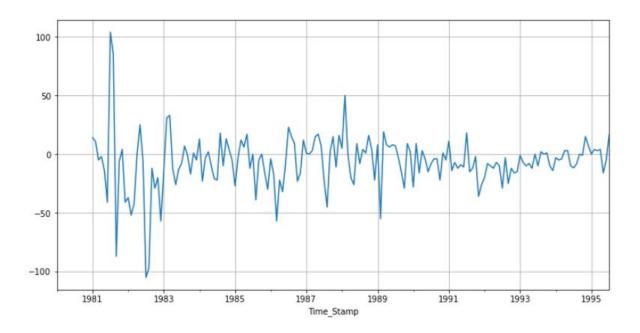
^[1] Covariance matrix calculated using the outer product of gradients (complex-step).



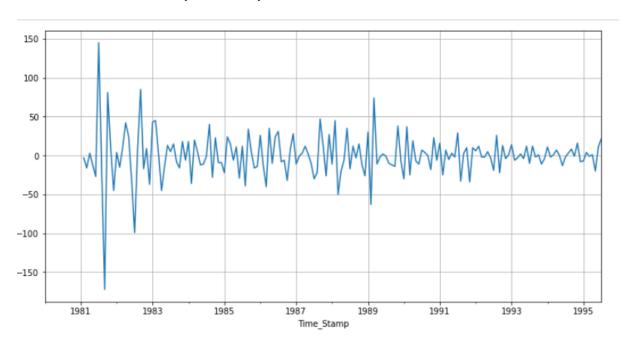


We see that there can be a seasonality of 12. We will run our auto SARIMA models by setting seasonality for 12.

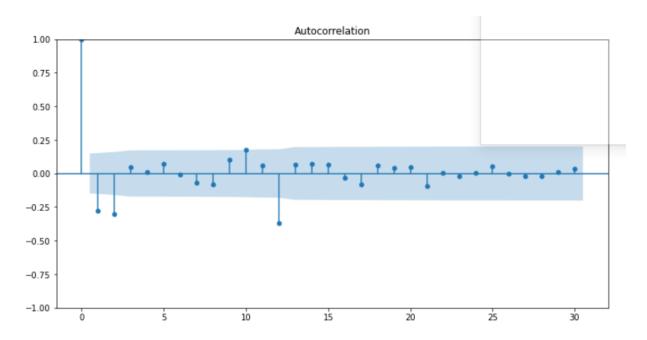
Plot with Seasonal Difference of 12:



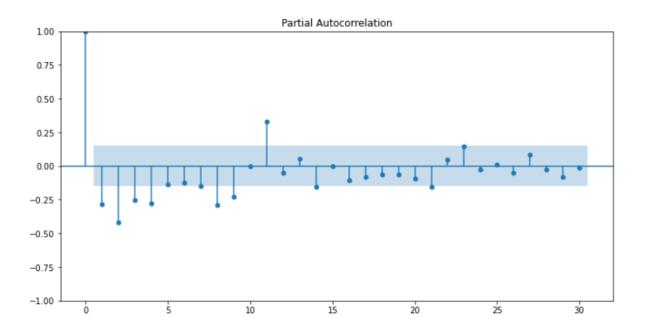
Plot without Trend and Only Seasonality:



ACF plot for the new modified Time Series with Stationarity:



PACF plot for the new modified Time Series with Stationarity:



Manual SARIMA (4,1,2) (4,1,2,12) Model Results:

| SARIMAX Results | | | | | | | |
|----------------------|-------------|------------|-------------|-------------|--------------------|----------|----------|
| | | | | | | | |
| Dep. Variab | le: | | | y No. | Observations 0 4 1 | : | 132 |
| Model: | SAR | IMAX(4, 1, | 2)x(4, 1, 2 | , 12) Log | Likelihood | | -277.661 |
| Date: | | | Sun, 06 Mar | 2022 AIC | | | 581.322 |
| Time: | | | 13: | 16:32 BIC | | | 609.983 |
| Sample: | | | | 0 HQI | C | | 592.663 |
| | | | | - 132 | | | |
| Covariance | Type: | | | opg | | | |
| | | | | | | | |
| | coef | std err | Z | P> z | [0.025 | 0.975] | |
| | | | | | | | |
| ar.L1 | | | | | -1.364 | | |
| ar.L2 | | | | | -0.670 | | |
| ar.L3 | | 0.277 | | | -0.647 | | |
| ar.L4 | -0.1285 | 0.162 | -0.794 | | -0.446 | 0.189 | |
| ma.L1 | | 299.757 | 0.001 | | -587.353 | 587.674 | |
| | | 251.668 | -0.003 | | -494.099 | 492.420 | |
| ar.S.L12 | -0.1441 | 0.364 | -0.396 | | -0.858 | 0.569 | |
| ar.5.L24 | -0.3596 | 0.227 | -1.587 | 0.113 | -0.804 | 0.085 | |
| ar.S.L36 | -0.2153 | 0.106 | -2.039 | 0.041 | -0.422 | -0.008 | |
| ar.S.L48 | -0.1195 | 0.093 | -1.281 | 0.200 | -0.302 | 0.063 | |
| ma.S.L12 | -0.5158 | 0.343 | -1.502 | 0.133 | -1.189 | 0.157 | |
| ma.S.L24 | 0.2085 | 0.373 | 0.559 | 0.576 | -0.523 | 0.940 | |
| sigma2 | 215.3526 | 6.46e+04 | 0.003 | 0.997 | -1.26e+05 | 1.27e+05 | |
| | | | | | | | |
| Ljung-Box (L1) (Q): | | | | Jarque-Bera | a (JB): | | 2.41 |
| Prob(Q): | | | 0.86 | Prob(JB): | | | 0.30 |
| Heteroskeda | sticity (H) | : | 0.49 | Skew: | | | 0.32 |
| Prob(H) (two-sided): | | | 0.10 | Kurtosis: | | | 3.68 |
| | | | | | | | |
| | | | | | | | |

Figure 27 – Sarima Model with Seasonality 12 based on ACF and PACF cut-off points

Summary Frame for Alpha = 0.05:

| у | mean | mean_se | mean_ci_lower | mean_ci_upper |
|---|-----------|-----------|---------------|---------------|
| 0 | 46.385325 | 14.770634 | 17.435414 | 75.335237 |
| 1 | 62.932846 | 14.989764 | 33.553449 | 92.312244 |
| 2 | 63.527810 | 14.999436 | 34.129455 | 92.926165 |
| 3 | 66.473786 | 15.179628 | 36.722261 | 96.225311 |
| 4 | 63.540811 | 15.180505 | 33.787567 | 93.294055 |

RMSE

SARIMA(4,1,2)(4,1,2,12) 17.529039

2.7 Build a table with all the models built along with their corresponding parameters and the respective RMSE values on the test data.

| MODELS | TEST RMSE |
|--|-----------|
| 2pointTrailingMovingAverage | 11.529278 |
| Alpha=0.095,Beta=0,Gamma=0.0007,TripleExponentialSmoothing | 14.176738 |
| 4pointTrailingMovingAverage | 14.451403 |

Warnings:
[1] Covariance matrix calculated using the outer product of gradients (complex-step).

| 6pointTrailingMovingAverage | 14.566327 |
|---|-----------|
| 9pointTrailingMovingAverage | 14.72763 |
| 12pointTrailingMovingAverage | 15.236052 |
| RegressionOnTime | 15.268955 |
| For Alpha =0, Beta = 0 DoubleExponentialSmoothing | 15.268961 |
| SARIMA(4,1,2)(4,1,2,12) | 17.529039 |
| SARIMA(1, 1, 2)(2,0,2,6) | 26.134254 |
| Alpha=0.074,Beta=0.043,Gamma=0 TripleExponentialSmoothingMultiplicative | 19.741738 |
| SARIMA(0, 1, 2)(2, 0, 2, 12) | 26.928361 |
| Alpha=0.0987,SimpleExponentialSmoothing | 36.796243 |
| ARIMA(4,1,2) | 37.037639 |
| ARIMA(0,1,2) | 37.30648 |
| Simple Average | 53.46057 |
| Naive Model | 79.718773 |

Table 1 – All Models with RMSE

2.8 Based on the model-building exercise, build the most optimum model(s) on the complete data and predict 12 months into the future with appropriate confidence intervals/bands.

The best model built is **2pointTrailingMovingAverage** with Test RMSE of **11.529278**. However, moving average models are somewhat outdated and there are new models such as Arima/Sarima which are more robust and more advanced. Also, it is somewhat incorrect to use a 2point trailing moving average to future forecast 12 months. Hence, we will be building our model on the 2nd best optimum model, "Alpha=0.095,Beta=0,Gamma=0.0007,TripleExponentialSmoothing" which has the Test RMSE as **14.176** .Now we will built the best optimum full model on the same parameters

Future 12 Months Sales Forecast:

| 1995-08-31 | 49.878500 |
|---------------|-------------|
| 1995-09-30 | 46.705164 |
| 1995-10-31 | 45.439534 |
| 1995-11-30 | 60.040383 |
| 1995-12-31 | 98.313260 |
| 1996-01-31 | 13.835254 |
| 1996-02-29 | 24.144658 |
| 1996-03-31 | 31.704142 |
| 1996-04-30 | 24.511248 |
| 1996-05-31 | 27.880899 |
| 1996-06-30 | 33.379915 |
| 1996-07-31 | 44.016469 |
| Freq: M, dtyp | oe: float64 |

Future 12 Months Forecast Plot:

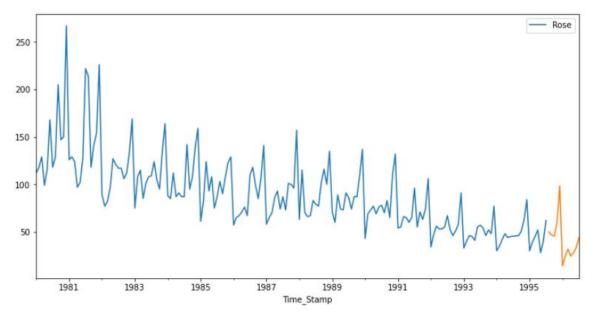


Figure 28 – Future 12 months Forecast Plot

Summary Frame of next 12 months for Alpha = 0.05:

| | lower_CI | prediction | upper_ci |
|------------|------------|------------|------------|
| 1995-08-31 | 15.161545 | 49.878500 | 84.595455 |
| 1995-09-30 | 11.988209 | 46.705164 | 81.422119 |
| 1995-10-31 | 10.722579 | 45.439534 | 80.156489 |
| 1995-11-30 | 25.323428 | 60.040383 | 94.757338 |
| 1995-12-31 | 63.596305 | 98.313260 | 133.030215 |
| 1996-01-31 | -20.881701 | 13.835254 | 48.552209 |
| 1996-02-29 | -10.572297 | 24.144658 | 58.861613 |
| 1996-03-31 | -3.012813 | 31.704142 | 66.421097 |
| 1996-04-30 | -10.205707 | 24.511248 | 59.228203 |
| 1996-05-31 | -6.836056 | 27.880899 | 62.597854 |
| 1996-06-30 | -1.337040 | 33.379915 | 68.096870 |
| 1996-07-31 | 9.299514 | 44.016469 | 78.733424 |

Forecast for next 12 months along with confidence band:

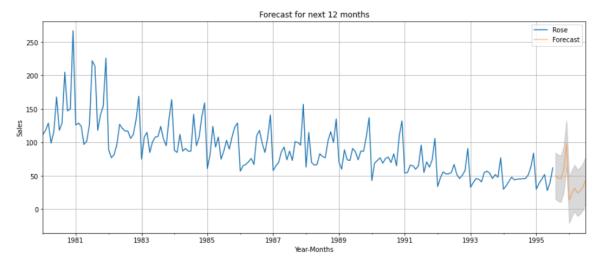


Figure 29 - Future 12 months Forecast with Confidence Bands

2.9 Based Comment on the model thus built and report your findings and suggest the measures that the company should be taking for future sales.

- The sales for Rose Wine has been dropping over the years. Hence, there is a need to reintroduce this variant and maybe with a little bit of refreshing touch, perhaps a different breed of grapes could be tried or cultivation at more cooler places, indoor cultivation at specific temperatures could work.
- It also could be evaluated if transportation of the grapes/bottles is having any effect on the quality of the wine or if there could be modifications made to the barrels that are used.
- The future forecast shows a slight uptick which is a good sign but it still remains a question if the popularity of Rose wine is sliding down or there is something which can turn around and increase the popularity.
- The company could also look at marketing strategies during peak season to boost sales even further which can maybe cover up for the lower sales during off-season or/and they can come up with impressive offers during off-season to increase the sales