

# Using Google Community Mobility Reports to investigate the growth of COVID-19 in the United States

Antonio Paez<sup>\*,a</sup>

<sup>a</sup>*School of Geography and Earth Sciences, McMaster University, Hamilton, ON, L8S 4K1, Canada*

## Abstract

In 2020 Google released a set of Community Mobility Reports (GCMR). These reports are based on the company's location-tracking capabilities and measure changes in mobility with respect to a baseline. This novel source of data offers an opportunity to investigate potential correlations between mobility and transmission of COVID-19. Using data from the New York Times on COVID-19 cases and GCMR, this paper presents an analysis of mobility levels and new daily reported cases of COVID-19 by state in the US. The results provide insights about the utility and interpretability of GCMR for COVID-19 research and decision-making.

## Research Questions and Hypotheses

The main policy tool to control the spread of the COVID-19 pandemic has been stay-at-home orders, which in the United States have been implemented on a state-by-state basis, with considerable variations in compliance. Concurrently, numerous initiatives have been developed to track the progress and the impact of the pandemic. As a result, there are new sources of data such as the recently-released Google Community Mobility Reports (GCMR)<sup>1</sup>, as well as The New York Times repository of COVID-19 data<sup>2</sup>. These two open data sets offer novel opportunities to investigate in quasi-real time the relationship between mobility patterns and transmission of COVID-19.

This paper investigates the potential of Google Community Mobility Reports to assess the impact of mobility on COVID-19. The following questions are posed:

- Do changes in mobility according to GCMR correlate with the transmission of COVID-19?
- And if so, what do we learn about mobility and the spread of the disease?

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\*Corresponding Author

Email address: [paezha@mcmaster.ca](mailto:paezha@mcmaster.ca) (Antonio Paez)

<sup>1</sup><https://www.google.com/covid19/mobility/>

<sup>2</sup><https://github.com/nytimes/covid-19-data>

This paper is a reproducible research document. The source is an R markdown file available in a public repository<sup>3</sup>.

## Methods and Data

GCMR use aggregated and anonymized data to chart changes in mobility with respect to different classes of places (see Table 1). Mobility indicators are calculated based on the frequency and length of visits. The reports give percentage change from a baseline level, which corresponds to the median value of mobility of identical days of the week during the period between January 3 and Feb 6, 2020. Covid-19 data is compiled by The New York Times based on reports from state and local health agencies.

Table 1: Descriptive statistics of the data set

Variable	Definition	min	median	max	sd
New Cases	New Daily Cases of COVID-19	0	90	12312	1051.67
date	Date	2020-02-27	2020-04-04	2020-05-02	
retail	Mobility trends for places like restaurants, cafes, shopping centers, theme parks, museums, libraries, and movie theaters	0.34	0.66	1.16	0.21
groceries	Mobility trends for places like grocery markets, food warehouses, farmers markets, specialty food shops, drug stores, and pharmacies	0.66	0.94	1.26	0.13
parks	Mobility trends for places like local parks, national parks, public beaches, marinas, dog parks, plazas, and public gardens	0.36	1.15	2.1	0.26
transit	Mobility trends for places like public transport hubs such as subway, bus, and train stations	0.24	0.72	1.14	0.23
work	Mobility trends for places of work	0.34	0.63	1.07	0.2
residential	Mobility trends for places of residence	0.97	1.14	1.27	0.08

*Note:*

All mobility indicators are lagged 11-day moving averages

For analysis, all mobility indicators are centered so that the value of 1 is the base mobility, and a 0.01 deviation corresponds to a 1% change. The incubation

<sup>3</sup>See folder Covid-19-Google-CMR-US in <https://github.com/paezha/Google-Mobility-Reports-and-COVID-19-US/>

time of the disease is between 2 and 12 days (95% interval; see Lauer et al., 2020). Given this, it is to be expected that any changes in mobility will have a lagged effect (if any) on the discovery of new cases. For this reason, lagged moving averages of the mobility indicators are calculated. Furthermore, it is possible that mobility and reports of new cases of COVID-19 are endogenous, if the public adjust their mobility according to reports of the incidence. Therefore, in addition to being consistent with an incubation period, use of lagged indicators also helps to break this potential endogeneity.

The lagged indicators are calculated as the mean of the mobility indicator using the values from date-minus-12-days to date-minus-2-days. Furthermore, using the cumulative number of reported COVID-19 cases, the total number of new daily cases is calculated. This variable (log-transformed after adding a small constant) is paired with the corresponding lagged moving average of the mobility indicators. The log-transformation is useful to avoid negative values of daily new cases when making predictions. Table 1 shows the descriptive statistics of the data set. Analysis is based on correlation analysis, multivariate regression, and data visualization.

## Findings

Table 2 shows that the mobility indicators are highly correlated with each other. Two variables are selected for multivariate analysis: parks- and work-related mobility. Work has a high correlation with the outcome variable, and its correlation with parks is relatively weak, which increases the information content of the two variables. Furthermore, parks- and work-related mobility represent two dimensions of out-of-home activities: mandatory and discretionary travel.

Table 2: Simple correlation between log(New Cases) and the mobility indicators

	log_new_cases	retail	groceries	parks	transit	work	residential
log_new_cases	1.00	-0.68	-0.48	-0.27	-0.68	-0.69	0.71
retail	-0.68	1.00	0.87	0.41	0.93	0.98	-0.98
groceries	-0.48	0.87	1.00	0.44	0.87	0.87	-0.87
parks	-0.27	0.41	0.44	1.00	0.49	0.39	-0.41
transit	-0.68	0.93	0.87	0.49	1.00	0.93	-0.95
work	-0.69	0.98	0.87	0.39	0.93	1.00	-0.99
residential	0.71	-0.98	-0.87	-0.41	-0.95	-0.99	1.00

*Note:*

All mobility indicators are lagged 11-day moving averages

A regression model is estimated with the log of new daily cases as the dependent variable. The covariates enter the regression in the form of a second order polynomial expansion. In addition, the date (centered on April 5) is introduced to account for the temporal trend of the pandemic. Finally, an indicator variable for the state of New York and an interaction with the date are used to distinguish the unusually high incidence of the disease there. The

results appear in Table 3. The model provides a good fit to the data and all variables reported are significant at  $p < 0.05$  or better.

There is an overall temporal trend that indicates a growing number of cases over time, at an accelerating rate (see positive sign of  $\text{date}^2$ ). Mobility related to parks and to work both tend to increase the number of new cases; the negative sign for their interaction is indicative of the trade-offs between these two forms of mobility and their impact on new cases. The influence of parks-related mobility was relatively weaker early in the pandemic (negative sign of the  $\text{parks} \times \text{date}$  term) but has become more important over time (positive sign of the  $\text{parks} \times \text{date}^2$  term). The opposite happens with work-related mobility, which started with a greater effect (positive sign of  $\text{work} \times \text{date}$  term), but whose impact has declined over time (negative sign of  $\text{work} \times \text{date}^2$  term). New York has on average more new daily cases than the rest of the states, but this has declined over time (see negative sign of  $\text{NY} \times \text{date}$ ).

Table 3: Results of estimating regression model. Dependent variable is  $\log(\text{New Cases} + 0.0001)$ .

Variable	Coefficient Estimate	p-value
date	0.1085	<0.001
$\text{date}^2$	0.0070	<0.001
$\text{parks}^2$	2.1756	<0.001
parks	8.9203	<0.001
$\text{parks} \times \text{work}$	-23.6010	<0.001
work	6.3156	0.0035
$\text{work}^2$	12.0539	<0.001
$\text{parks} \times \text{date}$	-0.2133	<0.001
$\text{parks} \times \text{date}^2$	0.0079	<0.001
$\text{work} \times \text{date}$	0.1376	0.0272
$\text{work} \times \text{date}^2$	-0.0238	<0.001
NY	3.5466	<0.001
$\text{NY} \times \text{date}$	-0.0527	<0.001

*Note:*

Coefficient of Determination  $R^2 = 0.831$

Adjusted Coefficient of Determination  $R^2 = 0.83$

Standard Error  $\sigma = 2.081$

Visualization is the most effective way to understand the trend according to the mobility indicators and date. Figure 1 shows the prediction surfaces on three different dates: March 21, when the first states began implementing stay-at-home orders; April 5, fifteen days later; and April 20, fifteen more days later, and at a time when some states started to consider relaxing stay-at-home orders.

On March 21 there were still only minor departures from baseline mobility (recall that these are temporally lagged); the prediction surface at this point is relatively flat. This changes by April 5, when work-based mobility has declined substantially. Although every state registers lower work-based mobility, there are large variations in parks-based mobility, with some states seeing increases of

up to 60% for this class of mobility.

The prediction surface indicates an expectation of a greater number of new daily cases as either class of mobility increases, but the effect is not linear. On the last date examined, April 20, the trend becomes more steep for park-based mobility, even as this indicator continues to display large variations from the baseline in both directions. The white dashed lines in the plots are the folds of the saddles, and represent, for each date, the combination of parks- and work-related mobility levels that tended to minimize the emergence of new cases.

The results suggest that over time the benefits of reduced work-related mobility can be offset by parks-related mobility. For example, on March 21, New Jersey and Idaho had similar levels of park-related mobility. By April 20, New Jersey had substantially reduced park-related mobility, whereas Idaho's was even higher than in March. The estimated and actual number of new cases grew in the intervening period; however, New Jersey's growth in new cases (from March 21 to April 20) was only 894%, whereas Idaho's was 1030%.

These results suggest the potential of GCMR to investigate the spread of COVID-19, but also point at some limitations. The baseline level is not defined in a metric that is amenable to policy development (e.g., person-km travelled). Without a clearer understanding of the absolute levels of these variables, their use can suggest trends, but their potential for applied policy analysis appears to be more limited.

## References

Lauer, S.A., Grantz, K.H., Bi, Q., Jones, F.K., Zheng, Q., Meredith, H.R., Azman, A.S., Reich, N.G., Lessler, J., 2020. The incubation period of coronavirus disease 2019 (covid-19) from publicly reported confirmed cases: Estimation and application. *Annals of Internal Medicine*. doi:10.7326/m20-0504

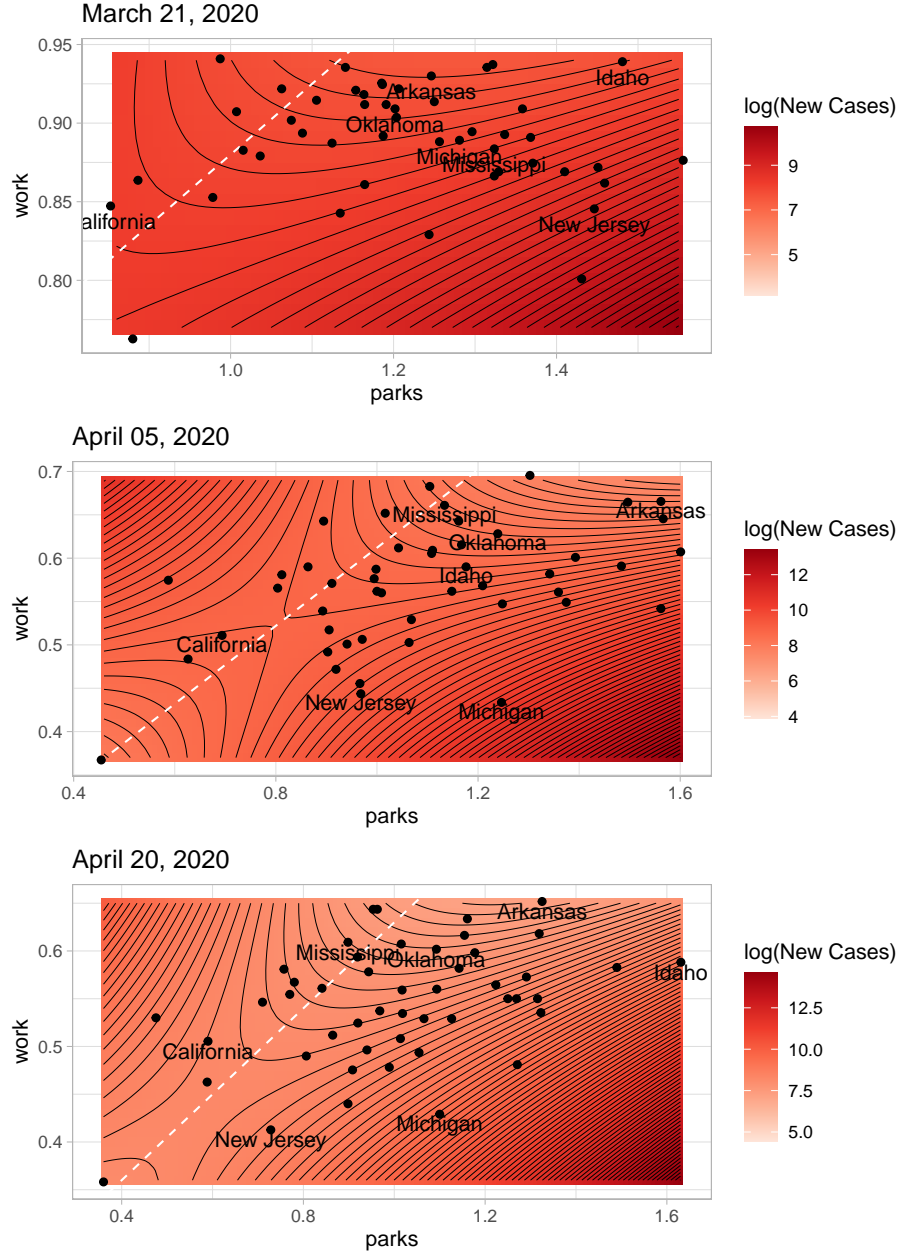


Figure 1: Prediction surfaces at three points during the pandemic according to the model; the dots are a scatterplot of the parks- and work-related mobility indicators of the states on that date; the white dashed line is the fold of the saddle.