

Sexually Antagonistic Selection on Genetic Variation is Rare in Humans

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ABSTRACT

Sex and sexual differentiation are ubiquitous across the tree of life. Because females and males often have substantially different functional requirements, we would expect selection to differ between the sexes. Recent studies in diverse species, including humans, suggest sexually antagonistic viability selection creates allele frequency differences between the sexes at many different loci. However, theory and population-level simulations suggest that sex-specific differences in viability would need to be very extreme in order to produce and maintain reported levels of between-sex genetic differentiation. In this study, we evaluate evidence for human sexually antagonistic selection on genomic variation in two independent, large biobanks (BioVU, $n = 93,864$, and UK Biobank, $n = 438,427$). We performed association tests between genetically ascertained sex and genotypes and, while we found dozens of genome-wide significant associations, none replicated across samples. Moreover, closer inspection revealed that all associations are likely due to cross-hybridization with regions of the sex chromosome during genotyping. Therefore, we find no compelling evidence for sexual antagonism on the autosomes, despite being well-powered to detect sex-specific allelic differences of down to 0.8% between the sexes. This study not only demonstrates a lack of strong ongoing sexually antagonistic selection on variation in humans, but also highlights significant challenges in identifying the genetic basis of sex-specific fitness differences using genome-wide scans in any species.

INTRODUCTION

Understanding the relationship between genotype and sexually dimorphic phenotypes, and how selection shapes this relationship, is fundamental to understanding sex-specific responses in aging¹, fertility², disease susceptibility³⁻⁵, and treatment⁶. Sexual dimorphism is common across a range of plant and animal taxa^{7,8}. Differences in optimal trait values between the sexes may result in sexually antagonistic selection⁹ – i.e., selection on autosomal variants that affect the trait in different directions for each sex. The repeated evolution of sexual dimorphism suggests sexually antagonistic selection may be common. Yet, we still lack an understanding of how frequently sexually antagonistic selection may occur in natural populations or how this process shapes genomic variation within and between species. One hypothesized consequence of the differences in sex-specific survival generated by conflicting effects of an allele within each sex is allele frequency differences between the sexes among adults^{10,11}.

Recent research has looked for these effects by identifying alleles with high male-female F_{ST} ^{10,12,13}, a normalized measure of allele frequency difference. Such studies across a range of taxa have suggested that potentially hundreds of autosomal loci are subject to ongoing sexually antagonistic selection with many differentiated loci having male-female divergence values of at least 10%¹⁴⁻¹⁸, and some reaching even as high as 45%¹⁹. The production and maintenance of such large male-female differences would require strong, ongoing selection, because alleles are homogenized every generation by meiotic segregation^{12,13}. Theory suggests that a male-female F_{ST} value of 1% requires at least a 33% viability cost per sex per generation¹³. Alternatively, population structure, sampling variance introduced by small sample sizes, and/or bioinformatic artifacts could contribute to the high divergence observed¹³. Of particular concern are the small sample sizes ($n = 15$ to 100 individuals) used by many of these previous studies. Detecting the level of allelic differentiation expected at sexually antagonistic loci with moderate sex-specific mortality ($\leq 10\%$ per sex) requires substantially larger sample sizes and other confounding effects, such as population structure, also need to be taken into account.

To overcome previous limitations, we use two large-scale biobanks, the UK Biobank and the Vanderbilt Biobank (BioVU). Compared to previous studies examining sexual antagonism, these datasets significantly improve our statistical power to detect allele frequency differences among the sexes by providing the largest available sample sizes to date – several orders greater than previous studies in humans^{12,16} and non-model taxa¹⁴⁻¹⁹. We analyze >500,000 human genomes for signals of male-female divergence driven by sexually antagonistic selection. Our association framework differs from traditional association studies as genetic sex is the phenotype of interest and the mechanism generating a true effect would be sex-specific viability. After controlling for multiple confounders, we find no clear evidence for sexually antagonistic variants.

RESULTS

Throughout this paper when we refer to an individual's sex, we are referencing that individual's sex chromosome composition as estimated in each biobank dataset and binarized (i.e., metadata reports each individual as XY or XX, although the datasets almost certainly include individuals not falling into these two categories²⁰). We make no statements in relation to gender, which is determined by many factors beyond genetics.

Seventy-seven variants show genome-wide significance as candidates for sexually antagonistic selection

To identify autosomal variants that could be under sexually antagonistic selection, we performed a genome-wide association study (GWAS) between females and males in two large, independent cohorts (BioVU: 34,269 females and 27,491 males; UK Biobank: 264,813 females and 223,478 males). We first applied standard quality control steps to remove samples with high relatedness, discordant sex, or high heterozygosity and excluded genotyped variants with high overall missing rate (Methods). We account for potential confounders by including age and 12 principal components for population stratification as covariates. The resulting p-values are well-calibrated, as verified by permuting the sex labels in the UK Biobank cohort (Supplementary Figure 1A), and so the standard genome-wide significance threshold of $P < 5E-8$ is appropriate for the association analysis (Methods, Supplementary Figure 1B). Applying this threshold resulted in five and 72 genome-wide significant variants in BioVU and UK Biobank, respectively.

Since different rates of variant missingness between cases and controls can lead to spurious associations²¹, we tested variants for a statistically significant difference in the missing rate between females and males (Methods). This control excluded what would have been 64 genome-wide significant variants in the UK Biobank and none in the BioVU cohort (Supplemental Figure 2, Supplementary File 1), leaving us with eight and five variants in the two datasets, respectively (Figure 1, Table 1). One intriguing genome-wide significant variant in the UK Biobank cohort, (rs11032483; OR = 1.25, $P < 1.3E-53$), which lies in a known regulatory region on chromosome 11 and has evidence from association studies for increasing risk in males and being protective in females for a number of sex-specific reproductive pathologies²².

No candidate loci replicate across BioVU and the UK Biobank

Comparing the five autosomal significant hits from BioVU to the eight from the UK Biobank, none of the associations are genome-wide significant in both cohorts (Table 1). Furthermore, none of the significant hits in one cohort even meet a nominal significance threshold ($P < 0.05$) in the other cohort. For example, the variant with the strongest association in the UK Biobank cohort (rs11032483) had no evidence for association with sex in the BioVU cohort ($P = 0.99$).

The regions surrounding each of the significant variants do not exhibit the expected association signal clusters arising from variants in strong linkage disequilibrium (LD) with the causal variant. For example, the most strongly associated variant overall (rs9870157) has 33 variants with R^2 of at least 0.8 in the 1000 Genomes Phase 3 European-ancestry (EUR) populations. However, there are no other strong associations among these variants. The lack of replication across the two cohorts and the missing association peaks among variants in strong LD suggest that these signals could be false positives driven by technical or biological artifacts.

Significant associations are likely due to mis-hybridization with sex chromosome regions

Genotyping error can occur due to probe cross-reactivity between different regions of the genome. Sex-biased error has been observed in array-based studies of DNA methylation²³ and has been reported in the canid genome²⁴ and stickleback genome¹⁷. For instance, if an autosomal variant is assayed with a probe sequence that has sufficient sequence similarity to a Y chromosome region carrying the reference allele, then males homozygous for the alternate

allele at the autosomal locus may instead be genotyped as heterozygous for the alternate allele. Females would not be subject to this bias, and thus there would appear to be an allele frequency difference between the sexes. Similarly, an autosomal variant with a probe sequence with high similarity to the X chromosome could result in a lack of homozygotes for the allele not on the X chromosome in both sexes, but the strength of this effect would differ between females and males. Furthermore, such cross-reactivity can lead the normalized intensities produced by genotyping arrays to lie outside of the regions corresponding to each genotype, and thus a missing genotype²⁵. Cross reactivity to a sex chromosome could therefore cause a differential missingness rate between the sexes. Indeed, we observe an almost complete lack of minor allele homozygotes in males across all thirteen genome-wide significant SNPs, as well as for females in all but four genome-wide significant SNPs (Supplemental Table 1). The same explanation is likely behind the 64 SNPs discarded for association between missingness and sex, as 26 of these SNPs have almost no male minor allele homozygotes and 47 have a p-value for lack of minor allele homozygotes of less than 1E-6.

To quantify the potential for mis-hybridization of sex chromosome regions to autosomal probes, we used BLAT²⁶ to find all regions across the genome with high sequence similarity to autosomal probe sequences on the MEGAEx (BioVU) and UK Axiom/BiEVE (UK Biobank) genotyping arrays (Methods). We assign each probe sequence the sex chromosome region with the highest BLAT score.

The probes for each significantly associated variant have high sequence similarity to a sex chromosome region (Figure 2, Table 2). In contrast, the majority of probes (79% in UK Biobank, 89% in BioVU) do not have any detectable similarity (BLAT score ≥ 20) to a sex chromosome sequence. Compared to the distribution of BLAT scores for probes with a match to a sex chromosome region, all genome-wide significant variants had BLAT scores greater than the 99th and 95th percentile for BioVU and UK Biobank respectively (inset Figure 2A, 2B). Using a stricter criteria to define potential sex chromosome sequence similarity (Methods), we find that all genome-wide significant variants in BioVU (Supplemental Figure 3A) and six out of eight genome-wide significant variants in UK Biobank (Supplemental Figure 3B) still have strong sequence similarity to a sex chromosome region (Table 2). Only 0.57% (4,587 probes) and 3.3% (20,528) of all probes in BioVU and UK Biobank respectively have such a sex chromosome match (Supplemental Figure 3). The difference in percentage is likely due to the UK Biobank arrays having longer probe sequences. Probes of genome-wide significant variants have similar BLAT matching properties as non-significant probes (Supplementary Figure 4) in UK Biobank and BioVU. Many of the 64 SNPs discarded for between-sex differences in missingness also demonstrated high sequence similarity to sex chromosome regions (Supplementary Table 8). Overall, the lack of homozygotes and the high sequence similarity between significant probes and sex chromosomes strongly suggests that sex-specific genotyping error is the source of the significant associations rather than sexually antagonistic selection.

The lack of sex-specific allele frequency differences is not due to being statistically underpowered

To determine if the lack of significant associations might be a result of being underpowered to detect plausible effect sizes, we conducted a power analysis (Methods). Based on the large

cohort sizes, we have 95% power to detect a variant with a true allele frequency difference greater than 2% between the sexes in the BioVU cohort and greater than 0.8% in the UK Biobank (Figure 3A). A frequency difference of $f\%$ caused by sex-specific antagonistic selection at a locus requires a mortality of roughly $f/2\%$, so we should be able to detect segregating variants with sex-specific mortality effects of at least 0.4%. For comparison, a cohort of 100 individuals, as used in a previous HapMap study¹⁶, only has 95% power to detect allele frequency differences between the sexes of 38% or greater (Figure 3B).

DISCUSSION

Understanding how sex-specific effects are transmitted by autosomal variation is critical for understanding how sexual dimorphisms arise and fix in populations. Sexually antagonistic selection maintains sexual dimorphisms and is predicted to be a pervasive driver of genome evolution⁹. Yet, empirically, the genomic footprint of this process is not well characterized. In this study, we sought to identify the extent of sexually antagonistic autosomal variation in human populations, capitalizing on two of the largest available biobanks. We performed genome-wide association tests for genetic sex that failed to identify and replicate any genome-wide significant variants. On closer inspection, a number of promising genome-wide significant SNPs were driven by technical artifacts, most likely due to high sequence similarity to a sex chromosome. Based on these results, we conclude there is no conclusive signal in these data of sexually antagonistic selection on genetic variants.

These results stand in contrast to previous studies, that have reported tens to hundreds of significantly differentiated variants with mean autosomal male-female F_{ST} values ranging from 1% to 45%¹⁴⁻¹⁹. These studies suggest strong, pervasive sexually antagonistic viability selection acting across the genomes of various species, which would be puzzling in light of theoretical observations and simulations indicating that strong allelic divergence between the sexes requires high sex-specific mortality rates to overcome the homogenizing effect of meiotic segregation occurring every generation¹³. In contrast to previous studies, the sample size of our study provided statistical power to distinguish true signal of plausible magnitude from stochastic noise. Additionally, our use of larger sample sizes provided power to detect smaller allelic divergence between the sexes – within the range predicted to be generated by weak sexually antagonistic selection.

We found strict quality control measures for population structure multiple testing essential. In particular, rigorous testing for sequence similarity to the sex chromosomes and showed that all significant SNPs had strong sequence matches. The potential for high sequence similarity between autosomes and sex chromosomes to cause sex-biased genotyping error is well established^{23,27}, though the potential for sex chromosome artifacts has not been fully appreciated until recently^{17,24}. In particular, probe sequences with high sequence similarity to one the sex chromosomes can lead to skewed allele frequency estimates in a sex-specific manner due to sequence mis-hybridization and the different sex chromosome content between females and males. This problem extends beyond SNP-based genotyping to read-based sequencing data, where inaccurate mapping of reads to an autosome instead of the sex chromosome could generate a similar skew in allele frequencies. This sex chromosome effect is potentially very common, and therefore, must be explicitly considered in any sex-specific or sex-stratified analyses to prevent technical and bioinformatic artifacts from generating false signals.

Participation bias rather than differential mortality can also generate a signal of male-female divergence²⁸, though this source of error is not relevant in this study since we did not find candidate SNPs for sexually antagonistic selection that passed our quality controls. Such artifacts will be especially problematic in species with new sex chromosomes, poorly assembled genomes, or rapidly evolving sex chromosome systems. In our case, filtering out SNPs with large differences in missingness between sexes and/or lack of homozygotes was sufficient to remove problematic SNPs.

Comparison of sequence similarity and match length for all probes indicates that thousands of other probes have similarly strong sex chromosome matches as the candidate variants analyzed here (Supplementary Figure 2). This high sequence similarity could suggest that more variants should show false positive signatures of sex-specific allele frequency differences. However, multiple factors contribute to the potential for mis-hybridization and inaccurate genotyping. For example, hybridization strength and kinetics are determined by sequence attributes beyond simple sequence identity, including local GC content and the potential for DNA secondary structures to form²⁹. Furthermore, the sequence region matched on the sex chromosome (i.e., pseudo-autosomal versus non-recombining) also matters. It is also likely that different quality control strategies used on different genotyping array platforms filter different problematic sites.

Although sexually antagonistic selection is certainly an important selective pressure, we see no evidence of it generating autosomal allelic divergence between the sexes in human populations. This strong negative result is unusual, as genome-wide association studies for most traits on a biobank-scale find significantly associated SNPs. We know that humans have the opportunity for sexually antagonistic effects, as seen through sex-specific mortality and disease susceptibility^{3,6}. However, randomization of alleles every generation by meiotic segregation means that a large selective pressure is required to create a large difference in allele frequencies, and thus, this genetic process makes it harder to detect the results of sexually antagonistic selection. Furthermore, some sexually antagonistic variants are not stably polymorphic; we would not detect these because they move rapidly to fixation.

Given the confounding factors, technical artifacts, and high sampling variance, identifying variants with small sex-specific effect sizes is a formidable challenge. We strongly recommend that future studies avoid simple metrics, like the male-female F_{ST} , and instead incorporate strict quality filters and control for known confounders into their association test. Sexually antagonistic viability selection is not the only action of sex-specific selection nor is male-female allelic divergence at a single locus the only signature of sexual antagonism. Given the extent of sexual dimorphisms in nature, there are almost surely autosomal loci subject to sexually antagonistic selection. However, our work illustrates that the field must reconsider our assumptions and develop new metrics for identifying signatures of sexual antagonism in the light of theoretical expectations to understand how this process affects the genome. Such studies will help us understand the translation of sex across the genotype-phenotype map and apply this to human health.

MATERIALS AND METHODS

Genotyping and quality control in BioVU

The DNA biobank at Vanderbilt University, consists of DNA extracted from blood collected during routine clinical testing. For 93,864 individuals, GWAS-level genotyping was performed using the Illumina MEGA-Ex chip which includes >2 million common and rare variants before imputation. We obtained genotyped data in PLINK format from the Vanderbilt sequencing core after the following quality control steps: excluding samples and variants with $\geq 5\%$ missingness, mismatched identifiers as detected by identity by descent checks, and non-concordance between reported gender and genetically determined sex. Overlapping variants with 1000 Genomes demonstrated $\geq 99.98\%$ variant call concordance using HapMap sample aliquots. Using PLINKv1.90b3s³⁰, we additionally performed the following quality control steps. We excluded duplicate samples and those with high missing rate ($\geq 5\%$), high heterozygosity on autosomes (> 3 S.D. from observed data), or high relatedness ($\%IBD \geq 0.2$). Next, we removed duplicated variants and variants with high missing rate ($\geq 5\%$) or significantly different missing rate between cases and controls ($p < 0.00001$, Fisher's Exact test). We then included only samples with a self or third party reported race as 'white' and variants with minor allele frequency > 0.01 . This additional quality control resulted in a final dataset of 61,760 samples (34,269 females and 27,491 males) and 1,763,607 variants. We calculated the top 12 principal components on this cohort. We imputed variants that reached nominal or genome-wide statistical significance ($P < 5E-8$) in the UK Biobank data but were not genotyped in the BioVU cohort. These variants were imputed using the Michigan Imputation Server (v1.2.4)³¹ using the HRC (Version r1.1 2016) reference panel and retaining variants with $R^2 > 0.3$. Imputed allele dosages were converted to hard calls using PLINK/2.00-alpha2³⁰ and filtered to exclude variants with minor allele frequency $< 1\%$ and genotyping rate $< 95\%$. All PLINK code is available on the GitHub repository XXX.

Genotyping and quality control in the UK Biobank

The UK Biobank is an international health resource with data from approximately 500,000 participants. Genotyping and quality control procedures have previously been described in detail by Bycroft et al. (2018). Briefly, two arrays – the UK Biobank Axiom Array ($n = 438,427$ participants) and the UK BiLEVE Axiom Array ($n = 49,950$ participants) – were used to genotype participants (71bp oligos). Quality control procedures carried out before the data were released, included: removal of participants with excess heterozygosity or missingness, removal of markers with batch, plate, array, or sex effects, and removal of markers with discordance across control replications³². The removal of sex effects, namely allele frequency differences between females and males at a given marker, does not preclude our analysis as the conservative threshold ($P < 10E-12$) removed only eight markers and the sex differences at these markers were due to technical artifacts, such as the probe sequence mapping to the Y chromosome (C. Bycroft pers. comm.). The released genotype data contains 805,462 markers from 488,377 participants (Field IDs 22100-22124). Additionally, the genetic sex (Field ID 22001), year of birth (Field ID 34), date of assessment (Field ID 53), and assessment center (Field ID 54) were requested for each participant. The top 40 genetic principal components (Field ID 22009) were previously calculated using fastPCA³².

Using PLINKv1.90b3s³⁰, we additionally performed the following quality control steps. We excluded samples with high missing rate ($\geq 5\%$) and high heterozygosity on autosomes (> 3 S.D. from observed data). Next, we pruned markers in linkage-disequilibrium (window size = 50kb, step rate = 5, r^2 threshold = 0.2). Finally, we removed variants with significantly different missing rate between cases and controls ($P < 0.00001$, Fisher's Exact test). We included only variants with minor allele frequency > 0.01 to exclude inaccurate calls made for low frequency alleles^{33,34}. This additional quality control resulted in a final dataset of 488,291 samples (264,813 females and 223,478 males) and 653,632 variants. A binomial test was used to test for a lack of minor allele homozygotes relative to that expected under HWE (this is conservative, because most human population dynamics is expected to lead to an excess of homozygosity). All PLINK code is available on the GitHub repository XXX.

Imputed genotype and phased haplotype values were requested to compare significant loci in the BioVU cohort, which were not directly genotyped in the UK BIOBANK arrays. Again, imputation was completed prior to the data release using the Haplotype Reference Consortium and UK10K haplotype resource. The imputation methods are described in detail in Bycroft et al. (2018). Imputed allele dosages were converted to hard calls using PLINK/2.00-alpha2³⁰.

Genome-wide association for an individual's sex

We performed a GWAS in UK Biobank and BioVU separately using a logistic regression testing the association between an individual's sex (binary variable, concordant with their genetic sex) and the effect allele, defined as the minor allele by PLINKv1.90b3s³⁰, using an additive model. For BioVU analysis, we controlled for genetic ancestry using 12 genetic principal components and included year of birth as a covariate. For the UK Biobank analysis, we again controlled for genetic ancestry using 12 genetic principal components, along with age at assessment as a continuous covariate and UK Biobank sampling center as a discrete covariate. All genome wide association tests were done using PLINKv1.90b3s³⁰. We focused our analyses on the autosomes, where genomic divergence between the sexes is not confounded by sex chromosome processes. During our quality control steps before association testing, we did not remove variants based on deviations from Hardy Weinberg Equilibrium (HWE) since theory indicates that sex-specific selection can violate the assumptions of HWE¹³.

Resampling of sex and generating a null distribution

To determine if p-values were well calibrated (i.e., uniformly distributed on $[0,1]$) at non-associated variants, we performed a permutation analysis to calculate the distribution of p-values within the UK Biobank cohort. We resampled genetic sex 100 times per chromosome to generate a set of random associations between genotype and this phenotype. We then reran the logistic regression, again including 12 genetic principal components, age, and sampling center as covariates, for only those variants that had a p-value < 0.01 in the original association analysis ($n = 8,868$ SNPs). These analyses generated a distribution of 100 p-values at each variant. Permuted p-values were uniformly distributed (Supplementary Figure 1A), even when the values were small (Supplementary Figure 1B), indicating the p-values for this association analysis were well-calibrated and therefore a genome-wide Bonferroni significance threshold of $P < 5E-8$ was appropriate. All R and PLINK code are available on the GitHub repository XXX.

Identifying SNPs with sequence similarity to sex chromosomes

Incorrectly mapped sex-chromosome variants to an autosomal region can result in statistically significant GWAS hits for an individual's sex due to the different effects on allele counts between females and males. We used BLAT²⁶ with default parameters (stepSize=5, repMatch=2253, minScore=20, minIdentity=0) to identify sequence similarity between the probe sequences used on the genotyping array and sex chromosome regions. The MEGA-Ex array probe sequences used to genotype the BioVU cohort were obtained directly from Illumina. Probes sequence for the UK Axiom Biobank array (Resource 149601) and UK BiLEVE array (Resource 149600) were download from <https://biobank.cts.ox.ac.uk/crystal/label.cgi?id=263>. MEGA-Ex probes are 50 base pair sequences adjacent to the variant being tested; MEGA-Ex uses single base extension to detect the variant allele. UK Biobank array probes are 71 base pairs long with the variant being genotyped located in the middle. BLAT hits to the X or Y chromosome were further filtered to identify regions likely to cross-hybridize by requiring at least 40 base pair overlap, sequence similarity $\geq 90\%$, and that the matching sequence overlaps (UK Biobank arrays) or flanks (MEGA-Ex array) the variant being tested. Similar criteria were used in a previous a study that reported cross-hybridization on the Illumina Infinium HumanMethylation27K microarray platform²⁷. Next, we identified the best BLAT hit to a sex chromosome for each probe sequence by selecting the hit with the highest BLAT score, which accounts for match length and sequence similarity. For this step, we considered the UK Axiom and UK BiLEVE array together thus selecting the probe sequence with the highest BLAT score from one of the two arrays per variant tested in the GWAS. In the BioVU (MEGA-Ex array) and UK Biobank arrays, 83,083 out of 798,051 and 128,090 out of 620,040 autosomal probes had at least one BLAT match (BLAT score ≥ 20) to a sex chromosome region. To further focus on sequence similarity with potential to cause genotyping error, we identified sex chromosome matches with the following criteria²⁷: 1) ≥ 40 base pairs in length, 2) $\geq 90\%$ sequence similarity, and 3) overlap between the match and the variant being genotyped.

Power Analysis

We conducted a power analysis to determine how the minimum allelic divergence between the sexes that could be detected within the BioVU and UK Biobank cohorts. Specifically, we determined the probability that we would reject the null observation that the population frequency of each allele is equal at a p-value threshold of $P = 1E-8$. Suppose we have N males and M women, and the allele frequencies in the two groups are P and Q . Since the cohort sample sizes are large, if the population frequencies are p and q , then $P \sim \text{Normal}(p, p(1-p)/2N)$ and $Q \sim \text{Normal}(q, q(1-q)/2M)$. The difference in population allele frequencies is then given by $P - Q \sim \text{Normal}(p - q, p(1-p)/2N + q(1-q)/2M)$. The variance is maximized when $p = q = 1/2$, so is at most the variance in the population is: $V = (1/N + 1/M)/8$. The two-sided p-value for $P-Q$ being nonzero will be below $1E-8$ if $|P-Q|$ is larger than $z(0.5e-8) * \sqrt{V}$, where $z(p)$ is the p-th quantile for the standard Normal distribution. Even, then, an allele with $|P-Q| = z(0.5E-8) * \sqrt{V}$ will only have a two-sided p-value half the time; alleles must be slightly farther apart (by $z(0.025) * \sqrt{V}$) to have a 95% probability that statistical noise does not put them above the $p=1e-8$ threshold. Therefore, we will have 95% power to detect any SNP with true $|p-q| > (z(0.5E-8) + z(0.025)) * \sqrt{V}$.

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Contributions

K.R.K. and P.C.P. devised the project. K.R.K., A.A., P.L.R., A.D.K., J.A.C., and P.C.P. designed the analyses. K.R.K. and A.A. performed analyses. K.R.K., A.A., P.L.R., and J.A.C wrote the manuscript with the support of the other authors.

The authors declare no competing interests.



Figure 1. Genome-wide association tests for genetic sex reveals candidate variants for sexually antagonistic selection. To identify candidate variants for sexually antagonistic selection, we performed genome-wide association tests between females (cases) and males (controls) in two large biobank cohorts: **(A)** BioVU (females = 34,269, males = 27,491) and **(B)** UK Biobank (females = 264,813, males = 223,478). After standard quality control and sex-specific missingness filters (Methods), we identified five variants with genome-wide statistically significant associations ($P < 5E-8$, solid red line) in BioVU and eight in the UK Biobank. None of the significant variants in BioVU and UK Biobank replicated at genome-wide or nominal significance ($P < 0.05$) across the two cohorts (Table 1). The probe sequence for each associated variant (except rs11032483) had >90% sequence identity to at least one sequence on a sex chromosome (Table 2). Each point represents one variant. Each variant is colored by whether the best match of its probe sequence to a sex chromosome (according to BLAT score) is on X (pink) or Y (green). If it has no strong match to either sex chromosome it is colored black. The size of each point indicates the degree of sequence similarity.

Table 1. Genome-wide significant variants in BioVU and UK Biobank cohorts. Variants passing genome-wide significance ($P < 5E-8$) in the BioVU or UK Biobank cohorts are reported. Genome-wide significant variants did not replicate across the cohorts. Location is reported in GRCh37/hg19 coordinates. Allele refers to the effect allele with which odds ratio (OR) is calculated. Individuals refers to the total number of individuals tested for the variant.

Location (chr:position)	SNP ID	Allele	Max. OR	Individuals	BioVU p-value	UK Biobank p-Value
<i>BioVU Significant SNPs</i>						
3:16652240	rs9870157	T	1.31	61,709	2.82e-83	0.42
7:100351596	rs145369881	T	0.78	60,499	3.25e-08	0.06
7:121147858	rs77638744	A	1.12	61,361	1.52e-10	0.13
13:20119336	rs9508454	C	1.19	61,694	1.26e-31	0.87
14:35761675	rs1048990	G	1.16	61,712	1.94e-20	0.75
<i>UK Biobank Significant SNPs</i>						
1:162075684	rs75745570	T	0.92	471,060	0.66	7.60e-14
4:88457099	rs114928327	T	0.89	413,257	0.30	4.09e-22
10:39006198	rs11598874	T	1.07	478,329	0.44	1.94e-12
11:4515024	rs11032483	T	1.25	482,581	0.99	1.33e-53
11:34104213	rs75212444	T	0.88	482,788	0.20	2.35e-13
12:118926685	rs7298104	T	0.91	475,771	0.93	5.97e-10
19:53535248	rs116890400	A	0.88	485,047	0.51	1.02e-11
21:18068575	rs73196350	A	0.94	479,137	0.76	5.19e-10

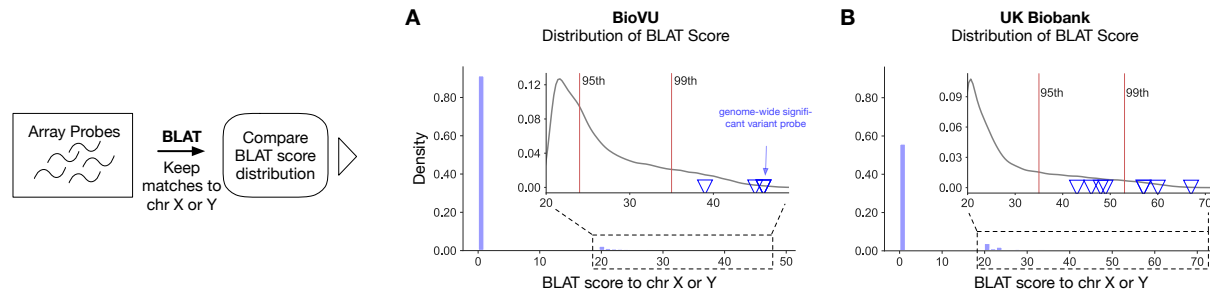


Figure 2. Probes for autosomal variants associated with genetic sex show high sequence similarity to sex chromosomes. We searched probe sequences used to genotype autosomal variants in the BioVU (798,051 autosomal probes) and UK Biobank (620,040 autosomal probes) cohorts for high sequence similarity to sex chromosome regions using BLAT (Methods). **(A)** More than 80% of BioVU autosomal probes do not have any sequence similarity (BLAT score ≤ 20) to a sex chromosome region; these are plotted at 0. Among the 83,083 BioVU probes with similarity to a sex chromosome sequence (inset), the probes for the variants with genome-wide significant associations with sex (blue triangles) are all in the tail of the distribution beyond the 99th percentile of the BLAT match score. **(B)** Patterns are similar for the UK Biobank probes; however, a higher fraction (20%, 128,090) have detectable similarity to a sex chromosome, likely due to their greater length than the BioVU probes.

Table 2. Best sex chromosome sequence match for genome-wide significant variant probes. Variants with genome-wide significant associations with genetic sex are reported with GWAS p-value (P-value) and the matched sex chromosome region (Matched Sex Chromosome) with the highest BLAT score (BLAT score). The sequence similarity and length of the matching region (Match Length) are also reported.

Dataset	SNP ID	Location (chr:position)	P-value	Matched Sex Chromosome	BLAT Score	Sequence Similarity (%)	Match Length (bp)
BioVU	rs9870157	3: 16652240	2.82E-83	Y: 26964471-26964521	46	96.0	50
BioVU	rs145369881	7: 100351596	3.25E-8	X: 26864979-2685116	39	90.0	50
BioVU	rs77638744	7: 121147858	1.52E-10	Y: 23315613-23316169	45	98.0	50
BioVU	rs9508454	13: 20119336	1.26E-31	Y: 28612640-28612689	46	91.9	50
BioVU	rs1048990	14: 35761675	1.94E-20	Y: 15398460-15398510	46	96.0	50
UKBB	rs75745570	1: 162075684	7.60E-14	X: 121952043-121952114	57	90.2	71
UKBB	rs114928327	4: 88457099	4.09E-22	X: 79084149-79084223	60	93.0	71
UKBB	rs11598874	10: 39006198	1.94E-12	Y: 13568059-13568130	67	97.2	71
UKBB	rs11032483	11: 4515024	1.33E-53	Y: 19070733-19070803	48	84.3	70
UKBB	rs75212444	11: 34104213	2.35E-13	X: 36967486-36967536	46	96.0	50
UKBB	rs7298104	12: 118926685	5.97E-10	Y: 1513524-1513899	43	93.9	59
UKBB	rs116890400	19: 53535248	1.02E-11	X: 38605892-38605963	57	90.2	71
UKBB	rs73196350	21: 18068575	5.19E-10	X: 80780038-80780101	49	88.9	63

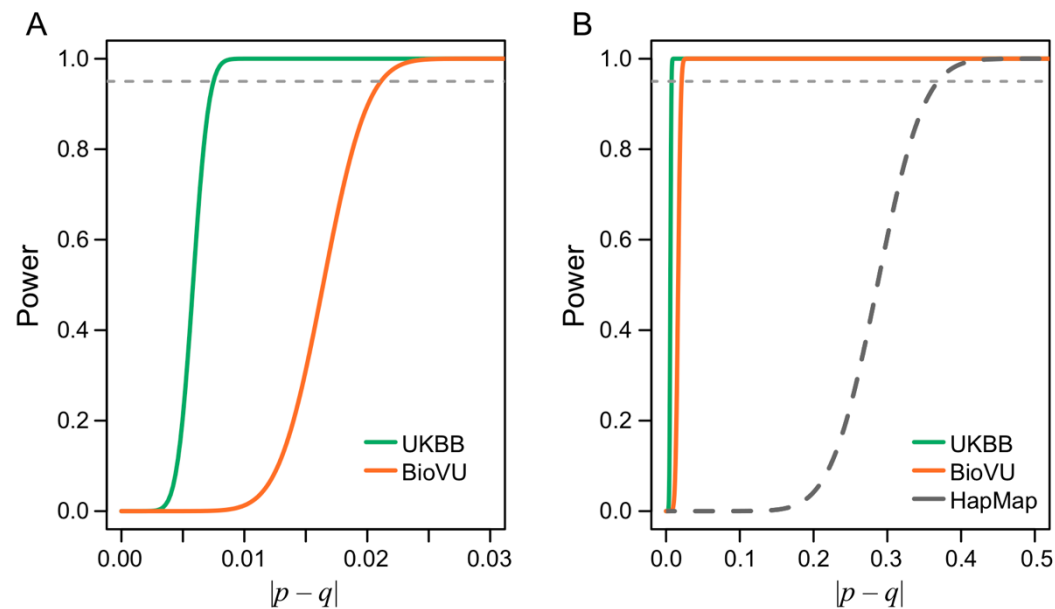


Figure 3. Statistical power was sufficient to detect small allelic divergence between the sexes. (A) The power to detect different levels of allelic divergence between the sexes was calculated for the BioVU (blue) and UKBB (green) cohorts. The dashed line shows the 95% power threshold. **(B)** Statistical power for the analyzed cohorts compared to previous analysis of human sequences¹⁶ based on approximately 100 individuals per HapMap population (gray).

REFERENCES

1. Archer, C. R., Recker, M., Duffy, E. & Hosken, D. J. Intralocus sexual conflict can resolve the male-female health-survival paradox. *Nat Commun* **9**, 1–7 (2018).
2. Farquhar, C. M. *et al.* Female subfertility. *Nat Rev Dis Primers* **5**, 1–22 (2019).
3. Morrow, E. H. The evolution of sex differences in disease. *Biol Sex Differ* **6**, 1–7 (2015).
4. Ferretti, M. T. *et al.* Sex differences in Alzheimer disease - the gateway to precision medicine. *Nat Rev Neurol* **14**, 457–469 (2018).
5. Dumitrescu, L. *et al.* Sex differences in the genetic predictors of Alzheimer's pathology. *Brain* **142**, 2581–2589 (2019).
6. Khramtsova, E. A., Davis, L. K. & Stranger, B. E. The role of sex in the genomics of human complex traits. *Nat Rev Genet* **62**, 1–190 (2018).
7. Rowe, L., Chenoweth, S. F. & Agrawal, A. F. The Genomics of Sexual Conflict. *Am Nat* **192**, 274–286 (2018).
8. Deegan, D. F. & Engel, N. Sexual Dimorphism in the Age of Genomics: How, When, Where. *Front Cell Dev Biol* **7**, 1–7 (2019).
9. Arnqvist, G. & Rowe, L. *Sexual conflict*. (Princeton University Press, 2005).
10. Kasimatis, K. R., Nelson, T. C. & Phillips, P. C. Genomic Signatures of Sexual Conflict. *J. Hered.* **108**, 780–790 (2017).
11. Mank, J. E. Population genetics of sexual conflict in the genomic era. *Nat Rev Genet* **7**, 1–10 (2017).
12. Cheng, C. & Kirkpatrick, M. Sex-Specific Selection and Sex-Biased Gene Expression in Humans and Flies. *PLoS Genetics* **12**, e1006170 (2016).
13. Kasimatis, K. R., Ralph, P. L. & Phillips, P. C. Limits to Genomic Divergence Under Sexually Antagonistic Selection. *G3* **9**, 3813–3824 (2019).
14. Wright, A. E. *et al.* Male-biased gene expression resolves sexual conflict through the evolution of sex-specific genetic architecture. *Evol Letters* **215**, 403–10 (2018).
15. Dutoit, L. *et al.* Sex-biased gene expression, sexual antagonism and levels of genetic diversity in the collared flycatcher (*Ficedula albicollis*) genome. *Mol Ecol* **27**, 3572–3581 (2018).
16. Lucotte, E. A., Laurent, R., Heyer, E., Ségurel, L. & Toupance, B. Detection of Allelic Frequency Differences between the Sexes in Humans: A Signature of Sexually Antagonistic Selection. *Genome Biol Evol* **8**, 1489–1500 (2016).
17. Bissegger, M., Laurentino, T. G., Roesti, M. & Berner, D. Widespread intersex differentiation across the stickleback genome - The signature of sexually antagonistic selection? *Mol Ecol* **77**, 1–10 (2019).
18. Flanagan, S. P. & Jones, A. G. Genome-wide selection components analysis in a fish with male pregnancy. *Evol* **71**, 1096–1105 (2017).
19. Vaux, F. *et al.* Sex matters: Otolith shape and genomic variation in deacon rockfish (*Sebastes diaconus*). *Ecol Evol* **27**, 477–21 (2019).
20. Lanfranco, F., Kamischke, A., Zitzmann, M. & Nieschlag, E. Klinefelter's syndrome. *Lancet* **364**, 273–283 (2004).
21. Moskvina, V., Craddock, N., Holmans, P., Owen, M. J. & O'Donovan, M. C. Effects of differential genotyping error rate on the type I error probability of case-control studies. *Hum. Hered.* **61**, 55–64 (2006).
22. Cortes, A., Dendrou, C. A., Fugger, L. & McVean, G. Systematic classification of shared components of genetic risk for common human diseases. *bioRxiv* 1–22 (2018). doi:10.1101/374207
23. Chen, Y.-A. *et al.* Discovery of cross-reactive probes and polymorphic CpGs in the Illumina Infinium HumanMethylation450 microarray. *Epigenetics* **8**, 203–209 (2013).

24. Tsai, K. L., Evans, J. M., Noorai, R. E., Starr-Moss, A. N. & Clark, L. A. Novel Y Chromosome Retrocopies in Canids Revealed through a Genome-Wide Association Study for Sex. *Genes* **10**, 320–11 (2019).
25. Zhao, S. *et al.* Strategies for processing and quality control of Illumina genotyping arrays. *Brief Bioinform* **19**, 765–775 (2018).
26. Kent, W. J. BLAT - The BLAST-like alignment tool. *Genome Res.* **12**, 656–664 (2002).
27. Chen, Y.-A. *et al.* Cross-Reactive DNA Microarray Probes Lead to False Discovery of Autosomal Sex-Associated DNA Methylation. *Am. J. Hum. Genet.* **91**, 762–764 (2012).
28. Pirastu, N. *et al.* Genetic analyses identify widespread sex-differential participation bias. *bioRxiv* 1–54 (2020). doi:10.1101/2020.03.22.001453
29. Zhang, J. X. *et al.* Predicting DNA hybridization kinetics from sequence. *Nat Chem* **10**, 91–98 (2018).
30. Chang, C. C. *et al.* Second-generation PLINK: rising to the challenge of larger and richer datasets. *Gigascience* **4**, 2–16 (2015).
31. Das, S. *et al.* Next-generation genotype imputation service and methods. *Nat Genet* **48**, 1284–1287 (2016).
32. Bycroft, C. *et al.* The UK Biobank resource with deep phenotyping and genomic data. *Nature* **562**, 203–209 (2018).
33. Wright, C. F. *et al.* Assessing the Pathogenicity, Penetrance, and Expressivity of Putative Disease-Causing Variants in a Population Setting. *Am. J. Hum. Genet.* **104**, 275–286 (2019).
34. Weedon, M. N. *et al.* Assessing the analytical validity of SNP-chips for detecting very rare pathogenic variants: implications for direct-to-consumer genetic testing. *bioRxiv* 696799 (2019). doi:10.1101/696799

SUPPLEMENTARY FILES

Supplementary File 1: “TableS1_gwas_significant_hits.xlsx” - Summary statistics for genome-wide significant variants (including those removed for uneven missing rate between the sexes) associated with genetic sex in BioVU and UK Biobank. MALE_HOM1, MALE_HET, and MALE_HOM2 are the counts of genotypes of the minor allele homozygote, heterozygote, and major allele homozygote genotype calls in males, respectively; MALE_MISSING is the number of missing genotypes in males reported by plink. FEM_ prefixes similar columns for females. MISSING_PVAL gives the p-value from Fisher’s exact test comparing proportions of missing genotypes between males and females as reported by plink. HOM1_PVAL gives the p-value for a binomial test for the proportion of minor allele homozygotes (of either sex) being equal to the marginal allele frequency squared. OR, STAT and ASSOC_PVAL gives the maximum odds ratio, t-statistic and p-value from logistic regression as described in the text.

Supplementary File 2: “TableS2_HWE_genotype_counts.xlsx” - Genome-wide significant variants associated with genetic sex in BioVU and UK Biobank with Hardy-Weinberg Equilibrium statistics.

Supplementary File 3: “TableS3_raw_blat_xy_hits_bv.tsv” - Raw BLAT results for hits to chromosome X or Y for probes in the MEGAex genotyping array.

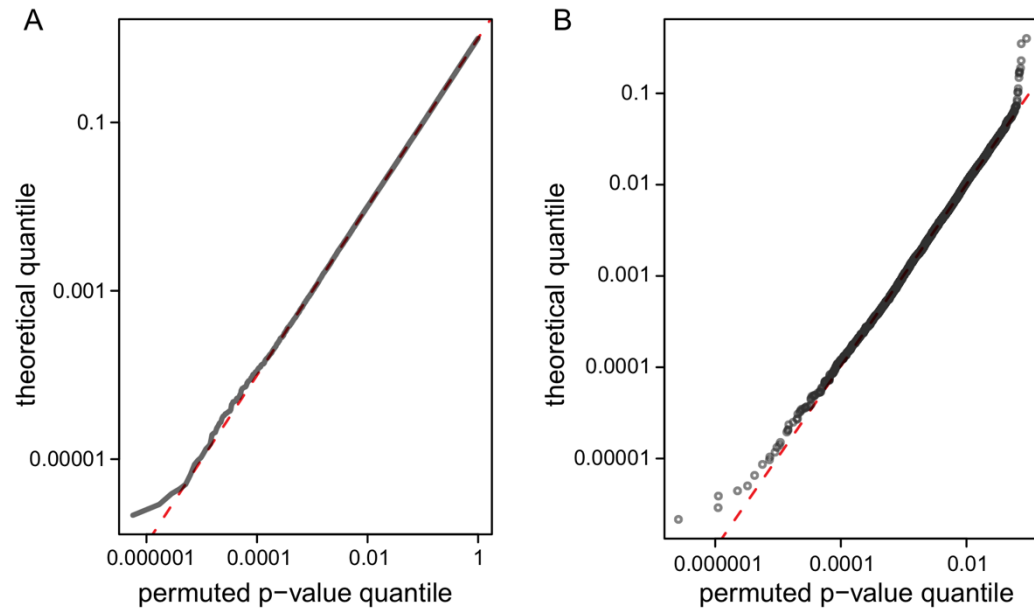
Supplementary File 4: “TableS4_raw_blat_xy_hits_ukaxiom.tsv” - Raw BLAT results for hits to chromosome X or Y for probes in the UK Biobank Axiom genotyping array.

Supplementary File 5: “TableS5_raw_blat_xy_hits_ukbil.tsv” - Raw BLAT results for hits to chromosome X or Y for probes in the UK BiLEVE genotyping array.

Supplementary File 6: “TableS6_best_blat_score_xy_hit_length_filtered_bv_gwas.tsv” - Best BLAT matches to chromosome X or Y based on highest BLAT score for each probe in the MEGAex genotyping array.

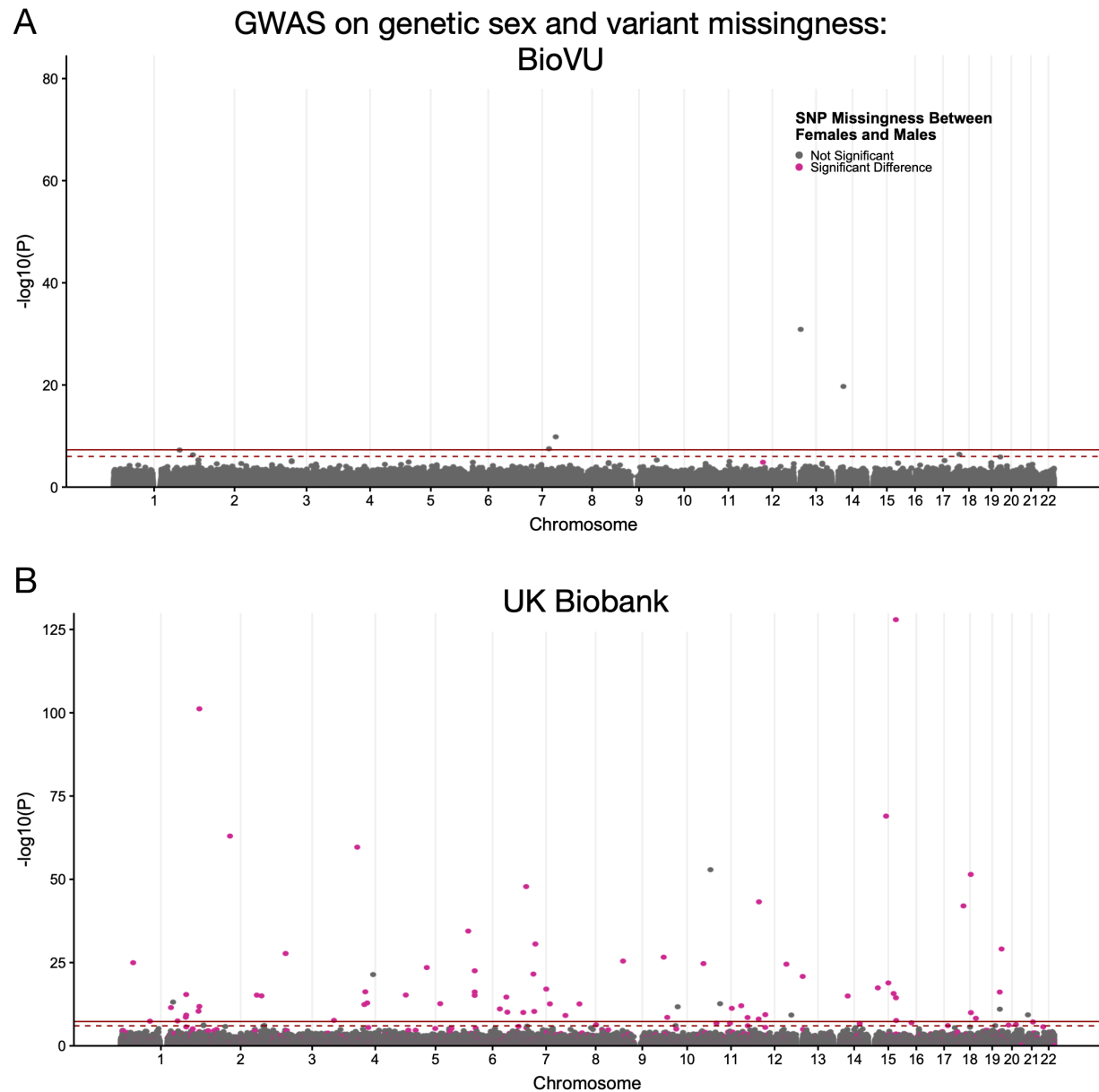
Supplementary File 7: “TableS7_best_blat_score_xy_hit_length_filtered_uk_gwas.tsv” - Best BLAT matches to chromosome X or Y based on highest BLAT score for each probe in the UK Biobank. BLAT results for hits to chromosome X or Y for probes are chosen after pooling across UK Biobank Axiom and UK BiLEVE genotyping arrays.

Supplementary File 8: “TableS8_uk_var_w_missingness_best_blat_score_xy.tsv” – Best sex chromosome match based on highest BLAT score for UK Biobank genome-wide significant variants with statistically significant difference in missing rate between females and males.



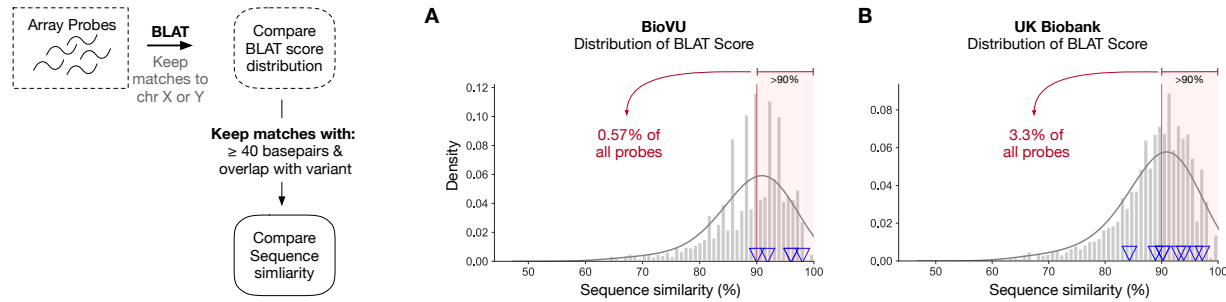
Supplementary Figure 1. Permutation of genetic sex to generate a null distribution

demonstrates that p-values are well calibrated. We randomly permuted genetic sex and ran a genome-wide association test between the permuted females and males in the UKBB cohort 100 times. Only those variants with a p-value < 0.01 under the association with the true genetic sex are considered ($n = 8,868$ SNPs). **(A)** Q-Q plot of all the permuted variants shows they are uniformly distributed. **(B)** Permuted variants are uniformly distributed even at very small p-values.

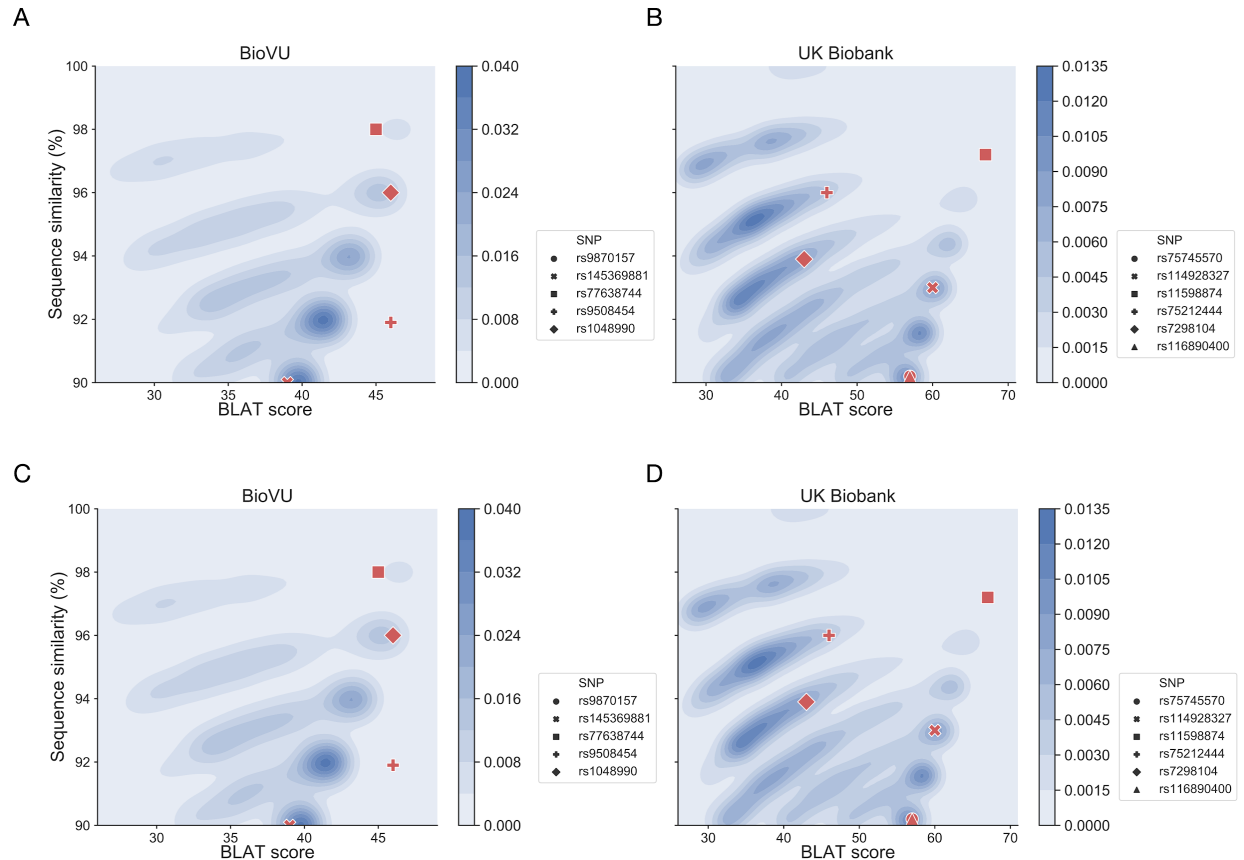


Supplementary Figure 2. Significantly different variant missingness between females and males contribute many spurious association in the UK Biobank GWAS for genetic sex.

After running a GWAS for genetic sex in **(A)** BioVU and **(B)** UK Biobank cohorts, we identify five and 72 variants with genome-wide significant associations (solid red line, $P < 5E-8$; dashed red line $P < 5E-6$) respectively. Variants with a statistically significant difference ($p < 0.00001$, Fisher's Exact test) in the missing rate between females and males are colored in red. In the UK Biobank cohort, 64 genome-wide significant variants also have a statistically significant difference in the missingness between cases and control, suggesting that these associations are spurious.



Supplementary Figure 3. Sequence similarity distribution of probes after applying strict matching criteria to a sex chromosome. To identify probes most likely to mis-hybridize between autosomal and sex chromosome sequences, we filtered those whose best BLAT match met the following criteria in (A) BioVU and (B) UK Biobank: sex chromosome match ≥ 40 base pairs in length, $\geq 90\%$ sequence similarity, and overlap of the matching region with the genotyped variant. Out of all autosomal probes, 0.57% and 3.3% met the aforementioned criteria in BioVU and UK Biobank, respectively.



Supplementary Figure 4. Probes of genome-wide significant variants with a match to sex chromosome regions have similar matching properties as non-significant variant when comparing BLAT score and match length to sequence similarity. Using BLAT, we identify array probe sequences with high sequence similarity ($\geq 90\%$) to a sex chromosome region, have a match length ≥ 40 base pairs, and overlap or is adjacent on the probe sequence to the variant being genotyped. We plot bivariate kernel density estimates comparing (A) BLAT score and (B) match length against sequence similarity (y-axis) for BioVU and UK Biobank probe sequences. Darker blue represents areas of higher density. The position of probe sequences for genome-wide significant variants are overlaid as red markers on each plot. Comparing against the densities of the non-significant variants, probes of genome-wide significant variants occur in areas of high density suggesting they have similar matching properties as non-significant probes.