

Exercise 9

13.1

记

$$X = \begin{pmatrix} 1 & x_1 \\ 2 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{pmatrix}$$

$$Y = (y_1, y_2, \dots, y_n)^T$$

$$\hat{\beta} = (\hat{\beta}_0, \hat{\beta}_1)^T$$

最小二乘法的结果为

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$

其中回归结果的残差为

$$\begin{aligned} \hat{\varepsilon} &= Y - X\hat{\beta} \\ &= Y - X(X^T X)^{-1} X^T Y \\ &= (I - X(X^T X)^{-1} X^T) Y \\ &= (I - X(X^T X)^{-1} X^T) (X\beta + \varepsilon) \\ &= (I - X(X^T X)^{-1} X^T) \varepsilon \end{aligned}$$

残差平方和为

$$\begin{aligned} \sum_{i=1}^n \hat{\varepsilon}_i^2 &= \hat{\varepsilon}^T \hat{\varepsilon} \\ &= \varepsilon^T (I - X(X^T X)^{-1} X^T)^T (I - X(X^T X)^{-1} X^T) \varepsilon \\ &= \varepsilon^T (I - X(X^T X)^{-1} X^T) \varepsilon \end{aligned}$$

求期望，注意到 ε_i 之间的独立性

$$\begin{aligned} E \sum_{i=1}^n \hat{\varepsilon}_i^2 &= E \varepsilon^T (I - X(X^T X)^{-1} X^T) \varepsilon \\ n E \hat{\varepsilon}_i^2 &= E \text{tr}(\varepsilon^T (I - X(X^T X)^{-1} X^T) \varepsilon) \\ &= \text{tr}(E(\varepsilon \varepsilon^T) (I - X(X^T X)^{-1} X^T)) \\ &= n \sigma^2 \text{tr}(I - X(X^T X)^{-1} X^T) \\ &= n \sigma^2 (n - \text{tr}(X^T X (X^T X)^{-1})) \\ &= n \sigma^2 (n - 2) \end{aligned}$$

也就是

$$\sigma^2 = \frac{1}{n-2} \sum_{i=1}^n \hat{\varepsilon}_i^2 = \frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

13.2

最小二乘法的结果为

$$\hat{\beta} = (X^T X)^{-1} X^T Y$$

其中

$$Y = X\beta + \varepsilon$$

于是

$$\mathbb{E}(\hat{\beta}|X^n) = \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}$$

下面计算方差

$$\begin{aligned} \hat{\beta}\hat{\beta}^T &= (X^T X)^{-1} X^T Y [(X^T X)^{-1} X^T Y]^T \\ &= (X^T X)^{-1} X^T Y Y^T X (X^T X)^{-1} \end{aligned}$$

其中

$$\mathbb{E}(Y Y^T) = \mathbb{E}(X\beta + \varepsilon)(X\beta + \varepsilon)^T = \mathbb{E}(X\beta)(X\beta)^T + \mathbb{E}\varepsilon\varepsilon^T = X\beta\beta^T X^T + \sigma^2 I$$

于是

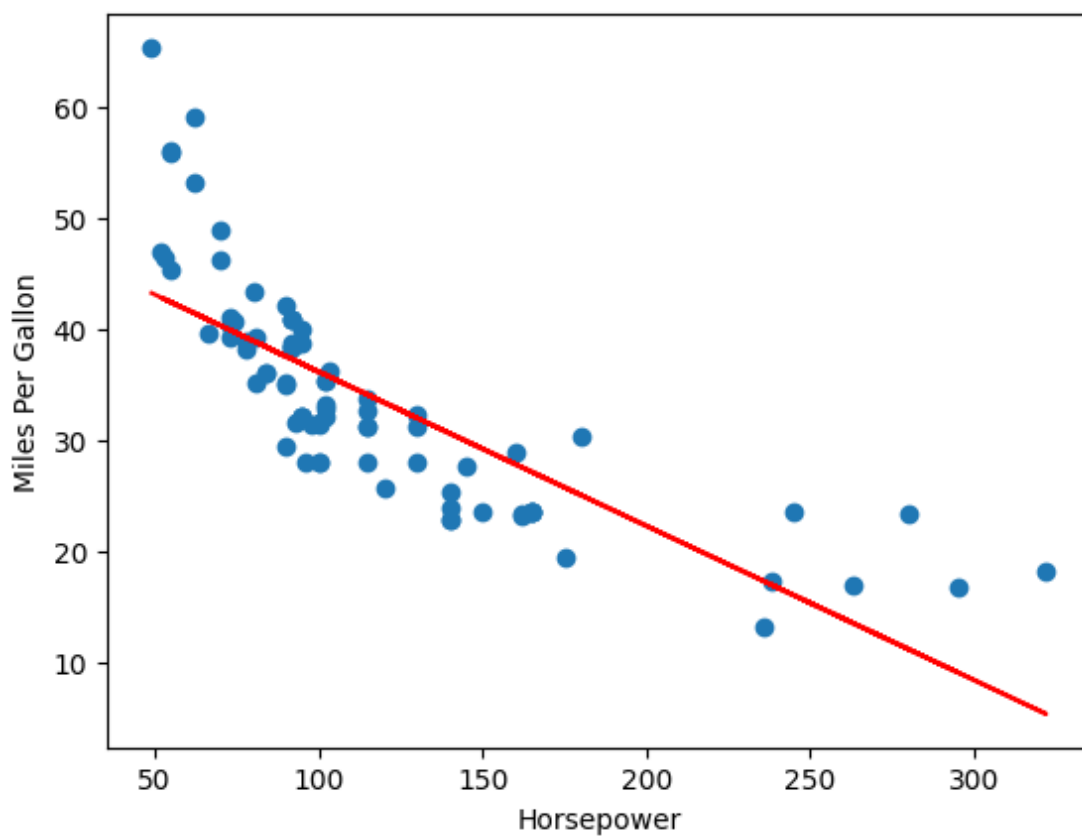
$$\begin{aligned} \mathbb{E}(\hat{\beta}\hat{\beta}^T) &= (X^T X)^{-1} X^T (X\beta\beta^T X^T + \sigma^2 I) X (X^T X)^{-1} \\ &= \beta\beta^T + \sigma^2 (X^T X)^{-1} \end{aligned}$$

$$\begin{aligned} \mathbb{V}(\hat{\beta}|X^n) &= \mathbb{E}(\hat{\beta}\hat{\beta}^T) - \mathbb{E}\hat{\beta}\mathbb{E}\hat{\beta}^T \\ &= \sigma^2 (X^T X)^{-1} \\ &= \frac{\sigma^2}{ns_X^2} \begin{pmatrix} \frac{1}{n} \sum_{i=1}^n x_i^2 & -\bar{X}_n \\ -\bar{X}_n & 1 \end{pmatrix} \end{aligned}$$

13.6

(a)

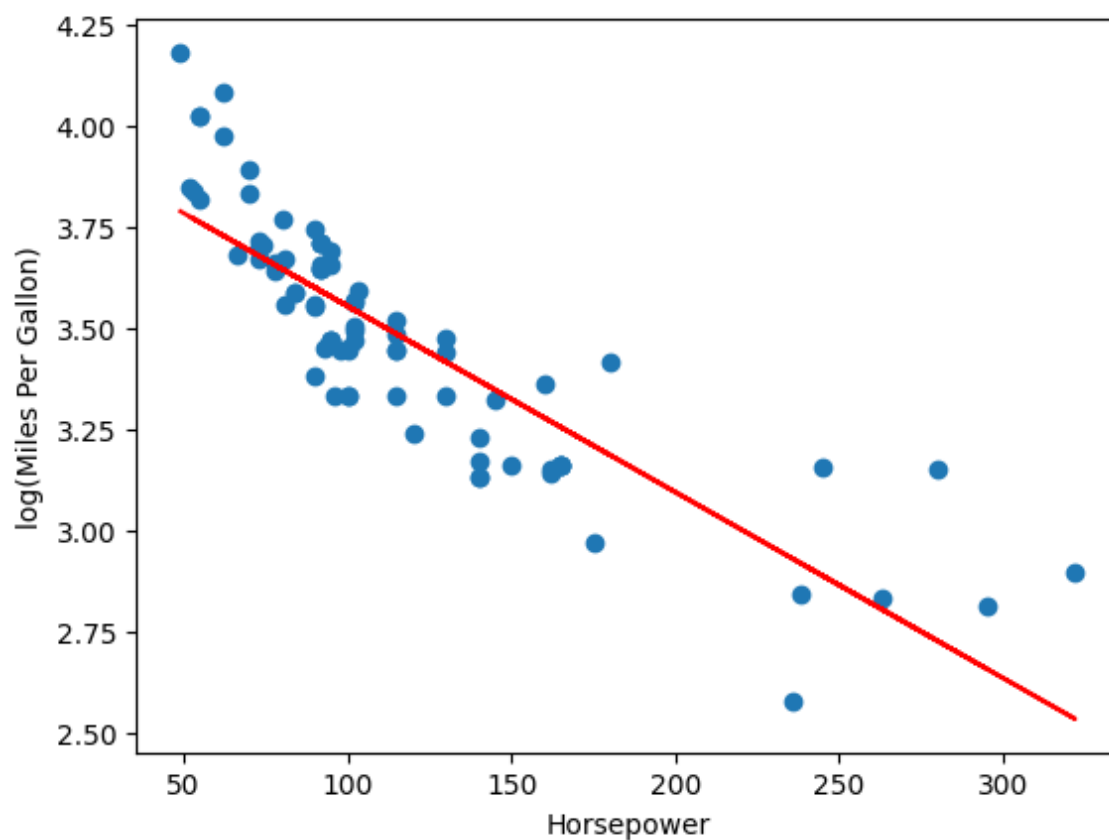
回归结果如为 $\text{MPG} = -0.13902326 \text{ HP} + 50.06607807$.



此时相关系数为-0.79，线性回归结果在两端有较大偏差。

(b)

回归结果如为 $\log \text{MPG} = -0.0045889 \text{ HP} + 4.0132294$.



此时相关系数为-0.85，可见取对数后相关系数更接近-1，拟合效果更好。

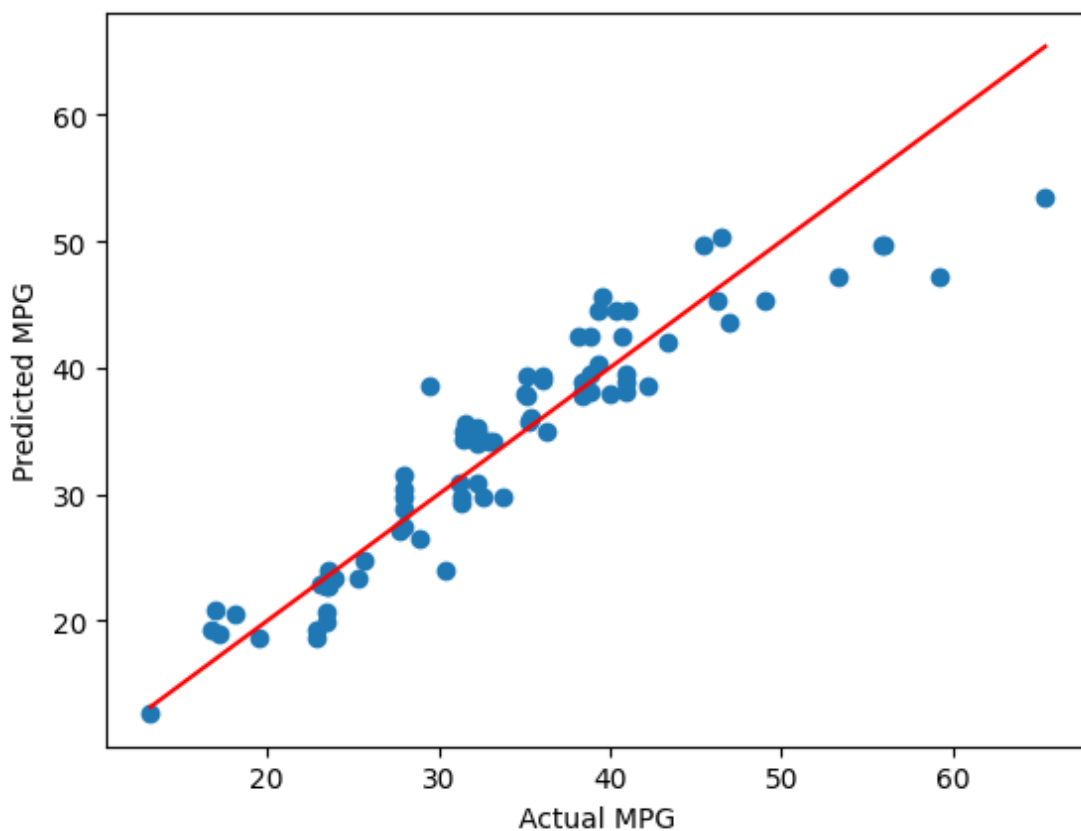
13.7

回归结果如下：

$$\text{MPG} = 192.437753 - 0.0156450113\text{HP} + 0.392212315\text{SP} - 1.29481848\text{WT} - 1.85980373\text{VOL}$$

对比上一问中仅使用HP的回归结果，其prediction risk为3049；使用多维回归结果的prediction risk为1027，可见我们获得了更好的拟合效果。

下图展示了所有数据点真实值与预测值的对比：



4.2

(a)

$$f(x, y) = \frac{1}{2\pi\sigma_{11}\sigma_{22}\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}\left[\left(\frac{x-\mu_1}{\sigma_{11}}\right)^2 - 2\rho\left(\frac{x-\mu_1}{\sigma_{11}}\right)\left(\frac{y-\mu_2}{\sigma_{22}}\right) + \left(\frac{y-\mu_2}{\sigma_{22}}\right)^2\right]}$$

$$= \frac{1}{2\sqrt{3}\pi} e^{-\frac{2}{3}\left[\left(\frac{x}{2}\right)^2 - \left(\frac{x}{2}\right)(y-2) + (y-2)^2\right]}$$

(b)

$$\Sigma_{12} = \Sigma_{21} = \rho\sigma_{11}\sigma_{22} = 1$$

$$\Sigma = \begin{pmatrix} 4 & 1 \\ 1 & 1 \end{pmatrix}$$

所以 squared generalized distance 为

$$\frac{x_1^2}{3} - \frac{2x_1x_2}{3} + \frac{4x_1}{3} + \frac{4x_2^2}{3} - \frac{16x_2}{3} + \frac{16}{3}$$

4.5

(a)

$$\mu_{1|x_2} = \mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(x_2 - \mu_2) = x_2 - 2$$

$$\Sigma_{11|x_2} = \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21} = 3$$

于是 $X_1|x_2 \sim N(x_2 - 2, 3)$

$$f(X_1|x_2) = \frac{1}{\sqrt{6\pi}} e^{-\frac{1}{6}(X_1 - x_2 + 2)^2}$$

(b)

$$\begin{aligned}\mu_{2|x_1, x_3} &= 1 - 2(x_1 + 3) = -2x_1 - 5 \\ \Sigma_{22|x_1, x_3} &= 5 - 4 = 1\end{aligned}$$

于是 $X_2|x_1, x_3 \sim N(-2x_1 - 5, 1)$

$$f(X_2|x_1, x_3) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(X_2+2x_1+5)^2}$$

(c)

$$\begin{aligned}\mu_{1|x_2, x_3} &= 0.5x_1 + 0.5x_2 + 1.5 \\ \Sigma_{33|x_1, x_2} &= 2 - 1.5 = 0.5\end{aligned}$$

于是 $X_3|x_1, x_2 \sim N(0.5x_1 + 0.5x_2 + 1.5, 0.5)$

$$f(X_3|x_1, x_2) = \frac{1}{\sqrt{\pi}} e^{-(X_3-0.5x_1-0.5x_2-1.5)^2}$$

4.16

(a) 由于 X_1, X_2, X_3, X_4 相互独立, 所以

$$\begin{aligned}V_1 &= \frac{1}{4}X_1 - \frac{1}{4}X_2 + \frac{1}{4}X_3 - \frac{1}{4}X_4 \sim N_p(0, \frac{1}{4}\Sigma) \\ V_2 &= \frac{1}{4}X_1 + \frac{1}{4}X_2 - \frac{1}{4}X_3 - \frac{1}{4}X_4 \sim N_p(0, \frac{1}{4}\Sigma)\end{aligned}$$

(b)

V_1, V_2 的协方差为

$$\begin{aligned}\text{cov}(V_1, V_2) &= \mathbb{E}[(V_1 - \mathbb{E}V_1)(V_2 - \mathbb{E}V_2)^T] \\ &= \mathbb{E}[V_1 V_2^T] \\ &= \mathbb{E}\left[\left(\frac{1}{4}X_1 - \frac{1}{4}X_2 + \frac{1}{4}X_3 - \frac{1}{4}X_4\right)\left(\frac{1}{4}X_1 + \frac{1}{4}X_2 - \frac{1}{4}X_3 - \frac{1}{4}X_4\right)^T\right]\end{aligned}$$

其中

$$\begin{aligned}\mathbb{E}X_i X_i^T &= \Sigma + \mu\mu^T \\ \mathbb{E}X_i X_j^T &= \mu\mu^T (i \neq j)\end{aligned}$$

于是

$$\text{cov}(V_1, V_2) = 0_{p \times p}$$

有

$$\begin{aligned}\mu_{V_1, V_2} &= 0_{2p \times 1} \\ \Sigma_{V_1, V_2} &= \begin{pmatrix} \frac{1}{4}\Sigma & 0 \\ 0 & \frac{1}{4}\Sigma \end{pmatrix}\end{aligned}$$

所以

$$\begin{aligned}\begin{pmatrix} V_1 \\ V_2 \end{pmatrix} &\sim N_{2p}\left(0, \begin{pmatrix} \frac{1}{4}\Sigma & 0 \\ 0 & \frac{1}{4}\Sigma \end{pmatrix}\right) \\ f(V_1, V_2) &= \frac{1}{(2\pi)^p \sqrt{|\Sigma|}} e^{-2V_1^T \Sigma^{-1} V_1 - 2V_2^T \Sigma^{-1} V_2}\end{aligned}$$