Coursework_MAP501_2021

02/12/2021

- (1a)
- (1b)
- (1c)
- (1d)
- (1e)
- (2a)
- (2b)
- (2c)
- (0-1)
- (2d)(2e)
- (3a)
- (3b)
- (3c)
- (3d)
- (3e)
- (3f)
- (4a)
- (4b)
- (4c)
- (4d)
- (4e)
- (4f)
- (4g)

library("car") library("rio") library("dplyr") library("tidyr") library("magrittr") library("ggplot2") library("pROC") library("nnet") library("caret") library("lme4") library("AmesHousing") library("Lahman") library("tidyverse") library("here") library("janitor") library("readxl") library("lindia") library("glmnet") library("caret") library("lme4")



```
People %>%
  select(c(playerID, birthYear, nameFirst, nameLast, weight, height, bats, throws, debut, birth
Country)) %>%
  rename("bornUSA" = "birthCountry") %>%
  mutate(bornUSA = as.logical(as.factor(bornUSA) =="USA")) -> Peopledata
```

(1b)

```
Batting %>%
  filter(yearID == 1985 | yearID == 2015)  %>%
  select(!c(G, teamID, lgID)) %>%
  mutate(batav = case_when(H == 0 ~ 0, H > 0 ~ H/AB)) -> Battingdata

Battingdata %>% sapply(function(x) sum(is.na(x)))

Fielding %>%
  filter(yearID == 1985 | yearID == 2015) %>%
  select(!c(G, teamID, lgID)) -> Fieldingdata1

Fieldingdata1 %>% sapply(function(x) sum(is.na(x)))

Fieldingdata1 %>%
  select(!c(PB, WP, SB, CS, ZR)) -> Fieldingdata

Fieldingdata
Battingdata
```

| p] | layerID | year: | | stint | | АВ | R | | H | | X2 | | X | 3B |
|------------------|-------------|----------|--------|-------|------------|----------|-------------|-----|-----------|---|---------|---|---|----|
| | 0 | | 0 | 0 | | 0 | 0 | | 9 | | | 0 | | 0 |
| | HR | RI | BI | SB | | CS | ВВ | | SC | | IB | | H | BP |
| | 0 | | 0 | 0 | | 0 | 0 | | 6 |) | | 0 | | 0 |
| | SH | 9 | SF | GIDP | ba | atav | | | | | | | | |
| | 0 | | 0 | 0 | | 0 | | | | | | | | |
| p] | layerID | year: | | stint | | POS | GS | Inr | 10uts | | P | | | Α |
| | 0 | | 0 | 0 | | 0 | 0 | | 6 | | | 0 | | 0 |
| | E | [| OP | PB | | WP | SB | | CS | 5 | Z | R | | |
| | 0 | | 0 | 2995 | 3 | 3205 | 2995 | | 2995 | 5 | 320 | 5 | | |
| | play | erID y | yearID | stint | POS | GS | InnOuts | PO | Α | Ε | DP | | | |
| 1 | aase | do01 | 1985 | 1 | Р | 0 | 264 | 8 | 10 | 0 | 0 | | | |
| 2 | abreg | jo01 | 1985 | 1 | Р | 5 | 72 | 1 | 6 | 1 | 0 | | | |
| 3 | acker | • | 1985 | 1 | Р | 0 | 259 | 10 | 16 | 0 | 1 | | | |
| 4 | adams | ri02 | 1985 | 1 | 2B | 3 | 84 | 9 | 13 | 1 | 1 | | | |
| 5 | adams | ri02 | 1985 | 1 | 3B | 10 | 337 | 2 | 31 | 1 | 3 | | | |
| 6 | adams | ri02 | 1985 | 1 | SS | 19 | 476 | 24 | 57 | 3 | 9 | | | |
| 7 | agost | ju01 | 1985 | 1 | Р | 0 | 181 | 10 | 15 | 1 | 0 | | | |
| 8 | aguay | lu01 | 1985 | 1 | 2B | 3 | 192 | 27 | 25 | 1 | 5 | | | |
| 9 | aguay | lu01 | 1985 | 1 | 3B | 3 | 126 | 4 | 16 | 0 | 1 | | | |
| 16 | aguay | lu01 | 1985 | 1 | SS | 36 | 1052 | 61 | 117 | 8 | 21 | | | |
| 11 | L aguil | ri01 | 1985 | 1 | Р | 19 | 367 | 8 | 16 | 0 | 1 | | | |
| 12 | _ | | 1985 | 1 | Р | 36 | 782 | 28 | 32 | 1 | 4 | | | |
| 13 | | | 1985 | 1 | С | 11 | 271 | 39 | 2 | 0 | 0 | | | |
| 14 | | _ | 1985 | 1 | Р | 1 | 87 | 2 | 5 | 0 | 0 | | | |
| 15 | | | 1985 | 2 | Р | 0 | 88 | 3 | 3 | 0 | 0 | | | |
| 16 | | | 1985 | 1 | 1B | 3 | 74 | 25 | 1 | 2 | 3 | | | |
| 17 | | | 1985 | 1 | 3B | 4 | 111 | 2 | 4 | 1 | 0 | | | |
| 18 | | | 1985 | 1 | 0F | 23 | 507 | 27 | 2 | 0 | 0 | | | |
| 19 | | | 1985 | 1 | SS | 31 | 920 | | 101 | 2 | 19 | | | |
| 26 | | | 1985 | 1 | 2B | 1 | 36 | 4 | 2 | 0 | 1 | | | |
| 21 | | | 1985 | 1 | | | | 28 | | 6 | | | | |
| 22 | | | 1985 | 1 | SS SB | 39 20 | 1100 566 | 29 | 107 78 | 3 | 10 9 | | | |
| 23 | | | 1985 | 1 | 33 P | 0 | 219 | 5 | 21 | 2 | 2 | | | |
| 24 | | | | 1 | P | 38 | 809 | 8 | 45 | 6 | 8 | | | |
| | • | • | 1985 | | | | | | | | | | | |
| 25 | | | 1985 | 1 | OF D | 79 0 | 1962 | 173 | 3 | 3 | 1 | | | |
| 26 | | | 1985 | 1 | P | 0 | 44 | 212 | 1 27 | 0 | 0 | | | |
| 27 | - | | 1985 | 1 | C | 55 | 1414 | 312 | 37 | 8 | 1 | | | |
| 28 | | | 1985 | 1 | P | 10 | 314 | 4 | 6 | 1 | 0 | | | |
| 29 | - | | 1985 | 1 | OF | 19 | 294 | 21 | 1 | 2 | 0 | | | |
| 36 | | | 1985 | 1 | | 122 | 3384 | | 370 | 7 | 76 | | | |
| 31 | | | 1985 | 1 | SS | 0 | 6 | 1 | 0 | 0 | 0 | | | |
| 32 | | | 1985 | 1 | 1B | 0 | 7 | 1 | 1 | 0 | 0 | | | |
| 33 | | | 1985 | 1 | С | 96 | 2651 | 565 | 51 | | 6 | | | |
| 34 | | | 1985 | 1 | 2B | 6 | 209 | 18 | 30 | 0 | 5 | | | |
| 35 | | | 1985 | 1 | 3B | 18 | 642 | 14 | 63 | 3 | 6 | | | |
| 36 | | | 1985 | 1 | OF | 0 | 1 | 0 | 0 | 0 | 0 | | | |
| 37 | | | 1985 | 1 | SS | 2 | 75 | 3 | 10 | 0 | 1 | | | |
| 38 | | | 1985 | 1 | OF | 158 | 4193 | 318 | 8 | 2 | 2 | | | |
| 39 | 9 bair | do01 | 1985 | 1 | Р | 3 | 147 | 4 | 9 | 0 | 0 | | | |
| 46 | ð bair | do01 | 1985 | 2 | Р | 0 | 6 | 0 | 0 | 0 | 0 | | | |
| 41 | L baker | do01 | 1985 | 1 | 2B | 0 | 6 | 0 | 0 | 0 | 0 | | | |
| 42 | 2 baker | do01 | 1985 | 1 | SS | 6 | 177 | 12 | 12 | 1 | 2 | | | |
| 43 | | | 1985 | 1 | 1 B | 53 | 1258 | 400 | 26 | 3 | 33 | | | |
| P _f Q | cessing met | hiu1092% | 1985 | 1 | OF | 25 | 699 | 65 | 3 | 2 | 0 | | | |
| | | | | | | | | | | | | | | |

```
      2474
      0
      1
      8
      0.27011494

      2475
      2
      0
      1
      0.16883117

      2476
      0
      0
      0.00000000

      2477
      0
      0
      0.00000000

      2478
      6
      0
      0.15873016

      2479
      0
      10
      13
      0.24855491

      2480
      0
      0
      0.00000000

      2481
      0
      3
      5
      0.26808511

      2482
      0
      2
      3
      0.28448276

      2483
      8
      2
      6
      0.17428571

      2484
      0
      0
      0.00000000
```

(1c)

```
Salaries %>%
  filter(yearID == 1985 | yearID == 2015) %>%
  inner_join(Fieldingdata, by = c("yearID", "playerID"), keep = FALSE) %>%
  mutate(allstar = playerID %in% AllstarFull$playerID) %>%
  inner_join(Battingdata, by = c("yearID", "playerID"), keep = FALSE) %>%
  inner_join(Peopledata, by = c("playerID" = "playerID"), keep = FALSE) %>%
  mutate(age = yearID - birthYear) %>%
  rename(stint = stint.x) %>%
  drop_na() %>%
  droplevels() -> Playerdata
Playerdata
```

| | yearID | teamID | lgID | playerID | salary | stint | POS | | InnOuts | PO | Α | Ε |
|----|---------------|--------|------|---------------|---------|-------|------------|-----|---------|------|-----|----|
| 1 | 1985 | ATL | NL | barkele01 | 870000 | 1 | Р | 18 | 221 | 2 | 9 | 1 |
| 2 | 1985 | ATL | NL | | 550000 | 1 | Р | 37 | 620 | 13 | 23 | 4 |
| 3 | 1985 | ATL | NL | benedbr01 | 545000 | 1 | C | 67 | 1698 | 314 | 35 | 4 |
| 4 | 1985 | ATL | NL | campri01 | 633333 | 1 | Р | 2 | 383 | 7 | 13 | 4 |
| 5 | 1985 | ATL | NL | ceronri01 | 625000 | 1 | C | 76 | 2097 | 384 | 48 | 6 |
| 6 | 1985 | ATL | NL | chambch01 | 800000 | 1 | 1B | 27 | 814 | 299 | 25 | 1 |
| 7 | 1985 | ATL | NL | dedmoje01 | 150000 | 1 | Р | 0 | 258 | 9 | 27 | 2 |
| 8 | 1985 | ATL | NL | forstte01 | 483333 | 1 | Р | 0 | 178 | 2 | 7 | 1 |
| 9 | 1985 | ATL | NL | garbege01 | 772000 | 1 | Р | 0 | 292 | 11 | 17 | 0 |
| 10 | 1985 | ATL | NL | harpete01 | 250000 | 1 | OF | 124 | 3299 | 215 | 10 | 5 |
| 11 | 1985 | ATL | NL | hornebo01 | 1500000 | 1 | 1B | 85 | 2239 | 892 | 58 | 0 |
| 12 | 1985 | ATL | NL | hornebo01 | 1500000 | 1 | 3B | 40 | 957 | 25 | 61 | 11 |
| 13 | 1985 | ATL | NL | hubbag101 | 455000 | 1 | 2B | 130 | 3425 | 339 | 539 | 10 |
| 14 | 1985 | ATL | NL | mahleri01 | 407500 | 1 | Р | 39 | 800 | 21 | 45 | 4 |
| 15 | 1985 | ATL | NL | mcmurcr01 | 275000 | 1 | Р | 6 | 135 | 2 | 12 | 2 |
| 16 | 1985 | ATL | NL | mumphje01 | 775000 | 1 | OF | 113 | 3009 | 248 | 6 | 8 |
| 17 | 1985 | ATL | NL | murphda05 | 1625000 | 1 | OF | 161 | 4264 | 334 | 8 | 7 |
| 18 | 1985 | ATL | NL | oberkke01 | 616667 | 1 | 2B | 12 | 275 | 18 | 37 | 1 |
| 19 | 1985 | ATL | NL | oberkke01 | 616667 | 1 | 3B | 101 | 2766 | 70 | 220 | 11 |
| 20 | 1985 | ATL | NL | perezpa01 | 450000 | 1 | Р | 22 | 286 | 7 | 9 | 1 |
| 21 | 1985 | ATL | NL | perryge01 | 120000 | 1 | 1 B | 50 | 1319 | 541 | 37 | 9 |
| 22 | 1985 | ATL | NL | perryge01 | 120000 | 1 | OF | 0 | 6 | 0 | 0 | 0 |
| 23 | 1985 | ATL | NL | ramirra01 | 750000 | 1 | SS | 130 | 3466 | 214 | 451 | 32 |
| 24 | 1985 | ATL | NL | suttebr01 | 1354167 | 1 | Р | 0 | 265 | 5 | 13 | 0 |
| 25 | 1985 | ATL | NL | washicl01 | 800000 | 1 | OF | 90 | 2459 | 122 | 3 | 5 |
| 26 | 1985 | BAL | AL | boddimi01 | 625000 | 1 | Р | 32 | 610 | 26 | 46 | 2 |
| 27 | 1985 | BAL | AL | dauerri01 | 480000 | 1 | 1B | 0 | 3 | 2 | 0 | 0 |
| 28 | 1985 | BAL | AL | dauerri01 | 480000 | 1 | 2B | 63 | 1504 | 117 | 181 | 3 |
| 29 | 1985 | BAL | AL | dauerri01 | 480000 | 1 | 3B | 8 | 250 | 7 | 21 | 1 |
| 30 | 1985 | BAL | AL | davisst02 | 437500 | 1 | Р | 28 | 525 | 15 | 20 | 0 |
| 31 | 1985 | BAL | AL | dempsri01 | 512500 | 1 | С | 113 | 3024 | 575 | 49 | 8 |
| 32 | | BAL | AL | dwyerji01 | 375000 | 1 | OF | 61 | 1494 | 131 | 4 | 1 |
| 33 | 1985 | BAL | AL | flanami01 | 641667 | 1 | Р | 15 | 258 | 4 | 11 | 0 |
| 34 | 1985 | BAL | AL | grosswa01 | 483333 | 1 | 1B | 5 | 132 | 40 | 4 | 0 |
| 35 | | BAL | | grosswa01 | 483333 | 1 | 3B | 57 | 1337 | 41 | 98 | 10 |
| 36 | | BAL | AL | lacyle01 | 725000 | 1 | OF | 112 | 2946 | 231 | 9 | 4 |
| 37 | | BAL | AL | lynnfr01 | 1090000 | 1 | | 121 | 3139 | 314 | 6 | 2 |
| 38 | | BAL | | martide01 | 560000 | 1 | Р | 31 | 540 | 17 | 26 | 1 |
| 39 | | BAL | | martiti01 | 440000 | 1 | Р | 0 | 210 | 9 | 10 | 1 |
| 40 | | BAL | | mcgresc01 | 547143 | 1 | Р | 34 | 612 | 13 | 26 | 1 |
| 41 | | BAL | | murraed02 | 1472819 | 1 | 1B | | | 1338 | | |
| 42 | | BAL | | nolanjo01 | 341667 | 1 | C | 4 | 97 | 22 | 2 | 0 |
| 43 | | BAL | | rayfofl01 | 128500 | 1 | 3B | 66 | 1829 | | 145 | 6 |
| 44 | | BAL | | rayfofl01 | 128500 | 1 | C | 22 | 575 | 114 | 7 | 1 |
| 45 | | BAL | | ripkeca01 | 800000 | 1 | | 161 | 4282 | | 474 | |
| 46 | | BAL | | roeniga01 | 558333 | 1 | OF | 53 | 1541 | 134 | 6 | 1 |
| 47 | | BAL | | sheetla01 | 60000 | 1 | 1B | 1 | 24 | 5 | 1 | 0 |
| 48 | | BAL | | sheetla01 | 60000 | 1 | 0F | 6 | 141 | 7 | | 1 |
| 49 | | BAL | | shelbjo01 | 130000 | 1 | 2B | 0 | 3 | 0 | 1 | 0 |
| 50 | | BAL | | shelbjo01 | 130000 | 1 | 0F | 43 | 1262 | 148 | 3 | 3 |
| 51 | | BAL | | stewasa01 | 581250 | 1 | P | 1 | 389 | 12 | 13 | 0 |
| 52 | | BAL | | youngmi01 | 121000 | 1 | OF | 83 | 2239 | 190 | 6 | 5 |
| 53 | | BOS | | armasto01 | 915000 | 1 | 0F | 79 | 1962 | 173 | 3 | 3 |
| | cessing 1986: | | | barrema02 | 272500 | 1 | | 150 | 4007 | | 479 | |
| 74 | - 1203 | ديو | AL | Jul 1 Cilladz | 212300 | | 20 | 100 | -007 | درر | 713 | |

| 1 | 2/2021, 17:52 | | | | |
|---|---------------|---|------------|-------|----|
| | 2330 | L | 2008-08-06 | TRUE | 30 |
| | 2331 | R | 2012-04-28 | TRUE | 23 |
| | 2332 | R | 2006-04-27 | TRUE | 34 |
| | 2333 | R | 2003-04-17 | TRUE | 39 |
| | 2334 | R | 2009-07-05 | FALSE | 31 |
| | 2335 | R | 2012-04-29 | TRUE | 28 |
| | 2336 | R | 2012-04-29 | TRUE | 28 |
| | 2337 | R | 2012-04-29 | TRUE | 28 |
| | 2338 | R | 2010-05-02 | FALSE | 28 |
| | 2339 | R | 2013-04-21 | TRUE | 25 |
| | 2340 | R | 2013-04-21 | TRUE | 25 |
| | 2341 | R | 2013-08-07 | TRUE | 29 |
| | 2342 | L | 2012-06-08 | TRUE | 30 |
| | 2343 | L | 2012-06-08 | TRUE | 30 |
| | 2344 | L | 2012-06-08 | TRUE | 30 |
| | 2345 | R | 2008-04-29 | TRUE | 31 |
| | 2346 | L | 2008-04-06 | TRUE | 31 |
| | 2347 | R | 2009-05-21 | TRUE | 31 |
| | 2348 | R | 2010-05-17 | TRUE | 28 |
| | 2349 | R | 2010-06-08 | TRUE | 27 |
| | 2350 | R | 2014-08-12 | TRUE | 24 |
| | 2351 | L | 2004-06-27 | TRUE | 39 |
| | 2352 | R | 2014-04-12 | TRUE | 27 |
| | 2353 | R | 2006-04-03 | TRUE | 35 |
| | 2354 | R | 2006-04-03 | TRUE | 35 |
| | 2355 | R | 2002-09-01 | TRUE | 36 |
| | 2356 | R | 2009-04-20 | TRUE | 29 |
| | 2357 | R | 2005-09-01 | TRUE | 31 |
| | 2358 | R | 2005-09-01 | TRUE | 31 |
| | | | | | |

(1d)

```
Salaries %>%
  group_by(teamID, yearID) %>%
  summarise(Rostercost = sum(salary), meansalary = mean(salary), rostersize = n_distinct(player
ID)) -> TeamSalaries
TeamSalaries
```

```
# A tibble: 918 x 5
# Groups:
           teamID [35]
  teamID yearID Rostercost meansalary rostersize
   <fct>
          <int>
                      <int>
                                 <dbl>
                                            <int>
1 ANA
            1997
                  31135472
                             1004370.
                                               31
2 ANA
            1998
                  41281000
                             1214147.
                                               34
 3 ANA
           1999
                  55388166
                             1384704.
                                               40
4 ANA
           2000
                  51464167
                             1715472.
                                               30
           2001
                 47535167
                                               30
5 ANA
                             1584506.
 6 ANA
            2002
                 61721667
                             2204345.
                                               28
7 ANA
           2003
                 79031667
                             2927099.
                                               27
8 ANA
            2004 100534667
                             3723506.
                                               27
9 ARI
           1998
                  32347000
                              898528.
                                               36
10 ARI
            1999
                   68703999
                              2020706.
                                               34
# ... with 908 more rows
```

(1e)

```
Teams %>%
  filter(yearID >= 1984, yearID <=2016) %>%
  inner_join(TeamSalaries, by = c("yearID", "teamID"), keep = FALSE) %>%
  drop_na() -> Teamdata
Teamdata
```

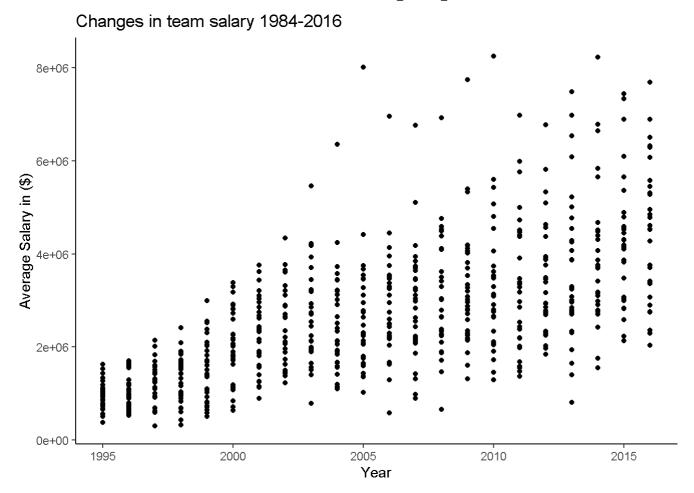
| | - | _ | | franchID | | | | Ghome | W | | DivWin | | LgWin |
|----------|--------------|----------|------------|------------|--------|---|------------|----------|----------|-----------|--------|--------|--------|
| 1 | 1995 | NL | ATL | ATL | Е | | 144 | 72 | 90 | 54 | Υ | N | Υ |
| 2 | 1995 | AL | BAL | BAL | E | | 144 | 72 | 71 | 73 | N | N | N |
| 3 | 1995 | AL | BOS | BOS | Е | | 144 | 72 | 86 | 58 | Υ | N | N |
| 4 | 1995 | AL | CAL | ANA | W | | 145 | 72 | 78 | 67 | N | N | N |
| 5 | 1995 | AL | CHA | CHW | С | | 145 | 72 | 68 | 76 | N | N | N |
| 6 | 1995 | NL | CHN | CHC | C | | 144 | 72 | 73 | 71 | N | N | N |
| 7 | 1995 | NL | CIN | CIN | С | | 144 | 72 | 85 | 59 | Υ | N | N |
| 8 | 1995 | AL | CLE | CLE | С | | 144 | | 100 | 44 | Y | N | Υ |
| 9 | 1995 | NL | COL | COL | W | | 144 | 72 | 77 | 67 | N | Υ | N |
| 10 | 1995 | AL | DET | DET | E | | 144 | 72 | 60 | 84 | N | N | N |
| 11 | 1995 | NL | FLO | FLA | E | | 143 | 71 | 67 | 76 | N | N | N |
| 12 | 1995 | NL | HOU | HOU | C | | 144 | 72 | 76 | 68 | N | N | N |
| 13 | 1995 | AL | KCA | KCR | C | | 144 | 72 | 70 | 74 | N | N | N |
| 14 | 1995 | NL | LAN | LAD | W | | 144 | 72 | 78 | 66 | Y | N | N |
| 15 | 1995 | AL | MIN | MIN | C | | 144 | 72 | 56 | 88 | N | N | N |
| 16 | 1995 | AL | ML4 | MIL | C | | 144 | 72 | 65 | 79 | N | N | N |
| 17 | 1995 | NL | MON | WSN | E | | 144 | 72 | 66 | 78 | N | N | N |
| 18 | 1995 | AL | NYA | NYY | E | | 145 | 73 | 79 | 65 | N | Y | N |
| 19 | 1995 | NL | NYN | NYM | E | | 144 | 72 | 69 | 75 | N | N | N |
| 20 | 1995 | AL | OAK | OAK | W | | 144 | 72 | 67 | 77 | N | N | N |
| 21 | 1995 | NL | PHI | PHI | Е | | 144 | 72 | 69 | 75 | N | N | N |
| 22 | 1995 | NL | PIT | PIT | C | | 144 | 72 | 58 | 86 | N | N | N |
| 23 | 1995 | NL | SDN | SDP | W | | 144 | 72 | 70 | 74 | N | N | N |
| 24 | 1995 | AL | SEA | SEA | W | | 145 | 73 | 79 | 66 | Υ | N | N |
| 25 | 1995 | NL | SFN | SFG | W | | 144 | 72 | 67 | 77 | N | N | N |
| 26 | 1995 | NL | SLN | STL | C | | 143 | 72 | 62 | 81 | N | N | N |
| 27 | 1995 | AL | TEX | TEX | W | | 144 | 72 | 74 | 70 | N | N | N |
| 28 | 1995 | AL | TOR | TOR | E _ | | 144 | 72 | 56 | 88 | N | N | N |
| 29 | 1996 | NL | ATL | ATL | E | | 162 | 81 | 96 | 66 | Y | N | Y |
| 30 | 1996 | AL | BAL | BAL | E | | 163 | 82 | 88 | 74 | N | Y | N |
| 31 | 1996 | AL | BOS | BOS | E | | 162 | 81 | 85 | 77 | N | N | N |
| 32 | 1996 | AL | CAL | ANA | W | | 161 | 81 | 70 | 91 | N | N | N |
| 33 | 1996 | AL | CHA | CHW | C | | 162 | | 85 | 77 | N | N | N |
| 34 | 1996 | NL | CHN | CHC | C | | 162 | 81 | 76 | 86 | N | N | N |
| 35 | 1996 | NL | CIN | CIN | C | | 162 | 81 | 81 | 81 | N | N | N |
| 36 | 1996 1996 | AL | CLE | CLE | C | | 161 | 80 | 99 | 62 | Y | N | N |
| 37 | | NL | COL | COL | W | | 162 | 81 | 83 | 79 | N | N | N |
| 38 39 | 1996 1996 | AL NL | DET FLO | DET FLA | E E | | 162 162 | 81 81 | 80 | 109 82 | N N | N N | N N |
| 40 | 1996 | NL | HOU | HOU | C | | 162 | 81 | 82 | 80 | | | |
| 41 | 1996 | | KCA | KCR | C | | 161 | 80 | 75 | | N | N | N |
| 42 | 1996 | AL NL | LAN | LAD | W | | 162 | 81 | 90 | 86 72 | N N | N Y | N N |
| 43 | 1996 | AL | MIN | MIN | C | | 162 | 82 | 78 | 84 | | | |
| 44 | 1996 | AL | ML4 | MIL | C | | 162 | 81 | 80 | 82 | N N | N N | N N |
| 45 | 1996 | NL | MON | WSN | E | | 162 | 81 | 88 | 74 | N | N | N |
| 46 | 1996 | AL | NYA | NYY | E | | 162 | 80 | 92 | 70 | Y | N | Y |
| 47 | 1996 | NL | | NYM | E | | | | 71 | | | | |
| 47 | 1996 | AL | NYN | OAK | W | | 162 162 | 81 81 | 71 78 | 91 84 | N | N N | N |
| 48 | 1996 | NL | OAK PHI | PHI | w E | | 162 | 81 | 78 67 | 95 | N N | N N | N N |
| 50 | 1996 | NL | PIT | PIT | C | | 162 | 80 | 73 | 89 | N | N | N |
| 51 | 1996 | NL NL | SDN | SDP | W | | 162 | 81 | 91 | 71 | N Y | N N | N N |
| 52 | 1996 | AL | SEA | SEA | W | | 161 | 81 | 85 | 71 76 | r N | N | N |
| 53 | 1996 | NL | _ SFN | SFG | W | | 162 | 82 | 68 | 94 | N | N | N |
| | essing gaath | | | STL | C | | 162 | 81 | 88 | 74 | Y | N | N |
| J4 | ±950 | INL | JLIN | JIL | C | 1 | 102 | OI | 50 | 74 | 1 | IN | IN |

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|----------------|-----|-----|-----------|-------------|-------------|--|
| 627 | BAL | BAL | 161863456 | 5581498.5 | 29 | |
| 628 | BOS | BOS | 188545761 | 6501578.0 | 29 | |
| 629 | CHA | CHA | 112998667 | 4519946.7 | 25 | |
| 630 | CHN | CHN | 154067668 | 5312678.2 | 29 | |
| 631 | CIN | CIN | 88940059 | 3066898.6 | 29 | |
| 632 | CLE | CLE | 74311900 | 2752292.6 | 27 | |
| 633 | COL | COL | 112645071 | 3413487.0 | 33 | |
| 634 | DET | DET | 194876481 | 6286338.1 | 31 | |
| 635 | HOU | HOU | 94893700 | 3389060.7 | 28 | |
| 636 | KCA | KCA | 131487125 | 4534038.8 | 29 | |
| 637 | ANA | ANA | 137251333 | 5278897.4 | 26 | |
| 638 | LAN | LAN | 221288380 | 6322525.1 | 35 | |
| 639 | FLO | MIA | 77314202 | 2761221.5 | 28 | |
| 640 | ML4 | MIL | 68775237 | 2292507.9 | 30 | |
| 641 | MIN | MIN | 102583200 | 4274300.0 | 24 | |
| 642 | NYA | NYA | 222997792 | 7689579.0 | 29 | |
| 643 | NYN | NYN | 133889129 | 4958856.6 | 27 | |
| 644 | OAK | OAK | 86806234 | 2893541.1 | 30 | |
| 645 | PHI | PHI | 58980000 | 2033793.1 | 29 | |
| 646 | PIT | PIT | 103778833 | 3706386.9 | 28 | |
| 647 | SDN | SDN | 101424814 | 3756474.6 | 27 | |
| 648 | SEA | SEA | 135683339 | 4845833.5 | 28 | |
| 649 | SFN | SFN | 172253778 | 6890151.1 | 25 | |
| 650 | SLN | SLN | 143053500 | 4614629.0 | 31 | |
| 651 | TBA | TBA | 57097310 | 2039189.6 | 28 | |
| 652 | TEX | TEX | 176038723 | 6070300.8 | 29 | |
| 653 | TOR | TOR | 138701700 | 4782817.2 | 29 | |
| 654 | MON | WAS | 141652646 | 5448178.7 | 26 | |
| | | | | | | |

(2a)

```
# Regular Plot

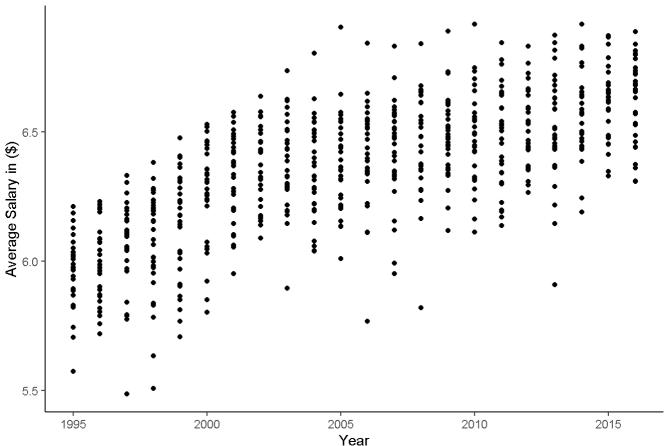
Teamdata %>%
    ggplot(mapping = aes(x = yearID, y = meansalary)) +
    geom_point() +
    labs(x = "Year", y = "Average Salary in ($)") +
    ggtitle("Changes in team salary 1984-2016") +
    theme_classic()
```



```
# Log Plot

Teamdata %>%
    ggplot(mapping = aes(x = yearID, y = log10(meansalary))) +
    geom_point() +
    labs(x = "Year", y = "Average Salary in ($)") +
    ggtitle("Changes in team salary 1984-2016") +
    theme_classic()
```





Two reasons a linear model using log base 10 as opposed to the raw salary figures here could be:

- Using log 10 axis in a linear model allows for a cleaner visualization of the data which is easier to interpret and present to a stakeholder, the exponential values on an axis require are too difficult to quickly read and understand.
- Our log scale also allows us to plot values which are significantly higher/lower on the same chart
 without them warping the visualization. For example in the regular plot there were very low salaries in
 95 and so it is difficult to see the spread of data down there when it's on the same chart as the high
 salaries later on; the log scale also allows for outliar values such as those seen in 2005 not to warp the
 scale to such a significant degree.

(2b)

```
lm(log10(meansalary) ~ yearID, data = Teamdata) -> linmod1
linmod1
summary(linmod1)
linmod1$coefficients
```

```
Call:
lm(formula = log10(meansalary) ~ yearID, data = Teamdata)
Coefficients:
               yearID
(Intercept)
 -51.22242
               0.02871
Call:
lm(formula = log10(meansalary) ~ yearID, data = Teamdata)
Residuals:
             1Q
                Median
                             3Q
                                    Max
-0.66345 -0.11692 0.00644 0.13394 0.55976
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
24.92 <2e-16 ***
            0.028711 0.001152
yearID
---
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.1858 on 652 degrees of freedom
Multiple R-squared: 0.4878,
                            Adjusted R-squared: 0.487
F-statistic: 620.9 on 1 and 652 DF, p-value: < 2.2e-16
(Intercept)
                yearID
-51.22241645
            0.02871141
```

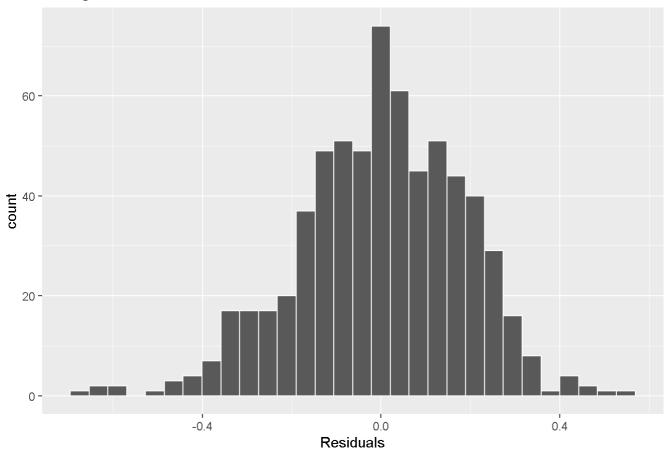
```
log10(meansalary) = -51.22242 + 0.02871 \times Year
```

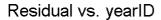
The multiple R squared here tells us that 48.78% of the variance seen in the log10(meansalary) value can be explained by the current year.

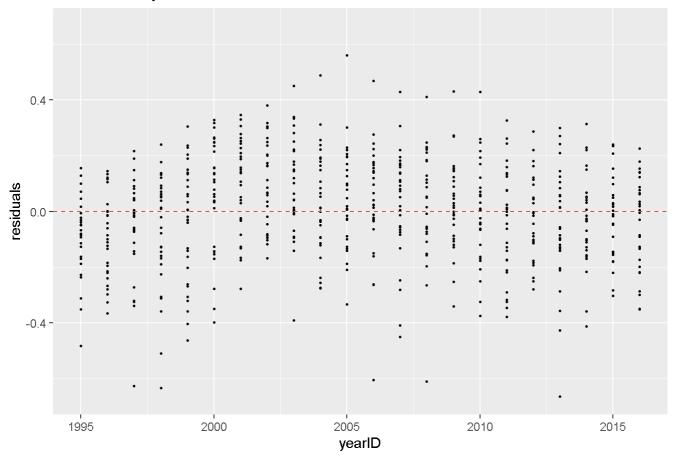
(2c)

```
linmod1 %>%
  gg_diagnose(max.per.page = 1)
```

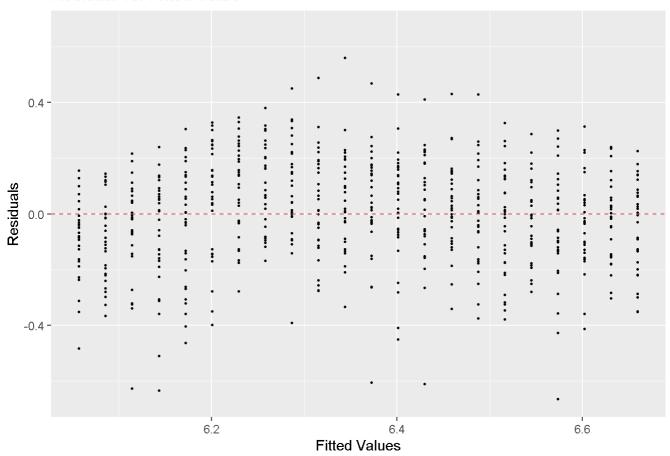
Histogram of Residuals

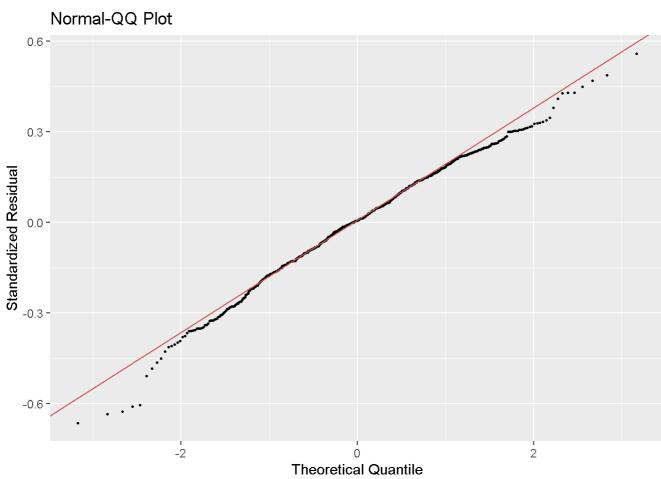




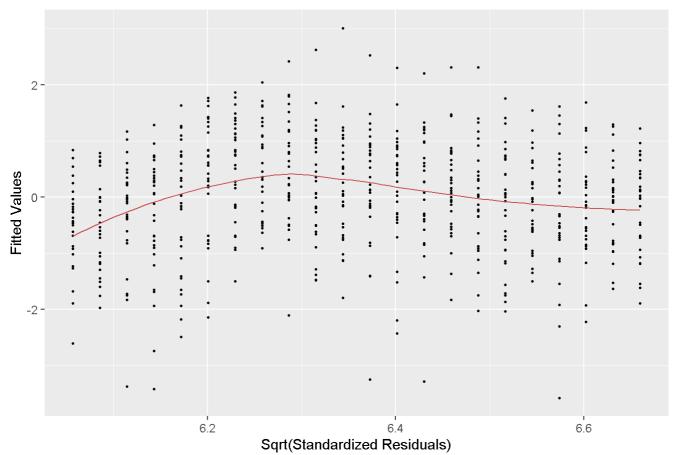


Residual vs. Fitted Value

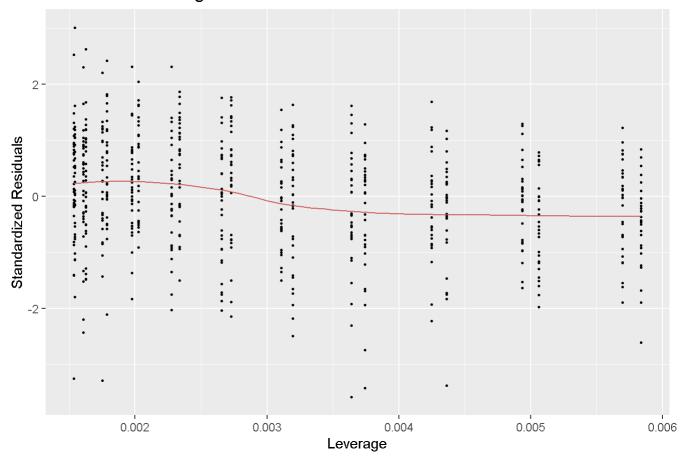




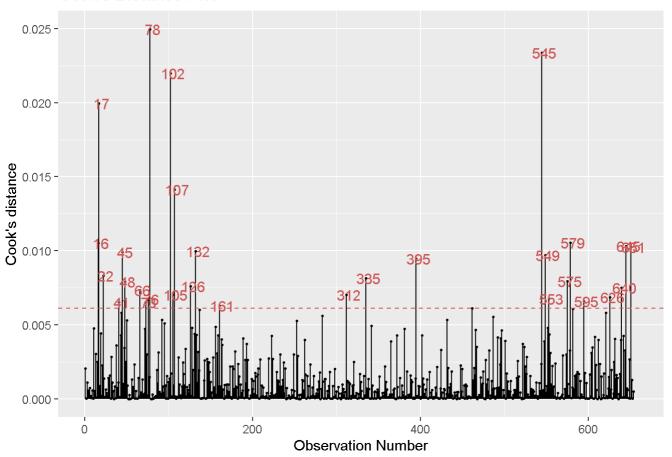
Scale-Location Plot



Residual vs. Leverage



Cook's Distance Plot



- 1. Linearity: to begin with we can look at both scatter plots in question 2a which both would indicate some form of linearity at first a glance. A look at the residual plots in the diagnose function would further support this assumption, I would say that we have a strong case for linearity here.
- 2. Homoscedasticity: for this assumption we have to analyse the variance in residual values for our data. In the above diagnose function, looking at the "Residual vs yearID" I believe this is a reasonable residual plot to assume homoscedasticity.
- 3. Normality: we can look at the QQ plot in the diagnose function, and seeing as the points are almost all hugging the straight line on the plot we can say that the residuals are approximately normally distributed. This is further supported by the histogram plot which we can see clearly indicated normality.
- 4. Independence: as we have time series data we have to watch out for autocorrelation, i.e. are data points easily predicted or known based on the data that's come in the previous X value/s. Looking at the residual plot for yearID, it doesn't seem that there is any autocorrelation occurring, I would say that generally the results seem independent of each other; however it would be interesting to see if we had a larger data set in the future whether or not time based factors such as economic cycle crashes would influence this data.

(2d)

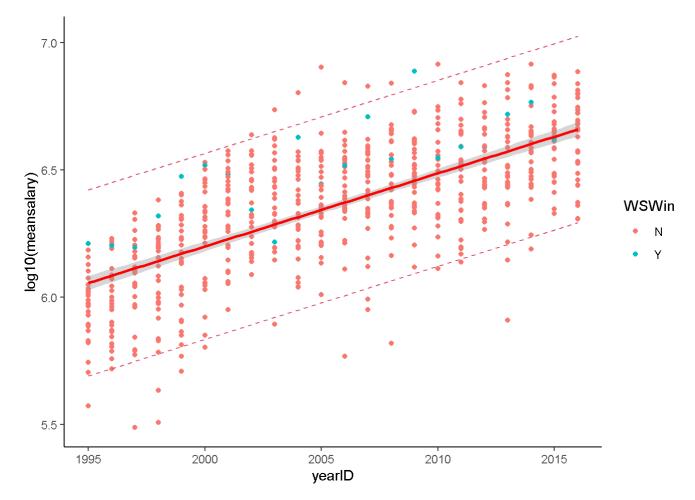
```
confint(linmod1) -> confidenceinterval1

predict(linmod1, interval = "prediction") -> predictionband1

as.tibble(predictionband1) -> predictiontibble

bind_cols(Teamdata, predictiontibble) -> TeamSalary2d

TeamSalary2d %>%
    ggplot(aes(x = yearID, y = log10(meansalary), colour = WSWin)) +
    geom_point() +
    geom_smooth(method = lm, colour = "red") +
    geom_line(aes(y = lwr), color = 2, lty = 2) +
    geom_line(aes(y = upr), color = 2, lty = 2) +
    theme_classic()
```



From this plot we can see that your chance of a world series win is definitely influenced by the amount you pay your teams, however I would also say that it seems post 2000 this is slowly becoming less of the case; perhaps due to the fact that more teams have the money to hand out (just speculating lol I have no clue about baseball).

(2e)

```
TeamSalary2d %>%
filter(log10(meansalary) > upr) %>%
select(yearID, name) %>%
Processing (math: 190%
```

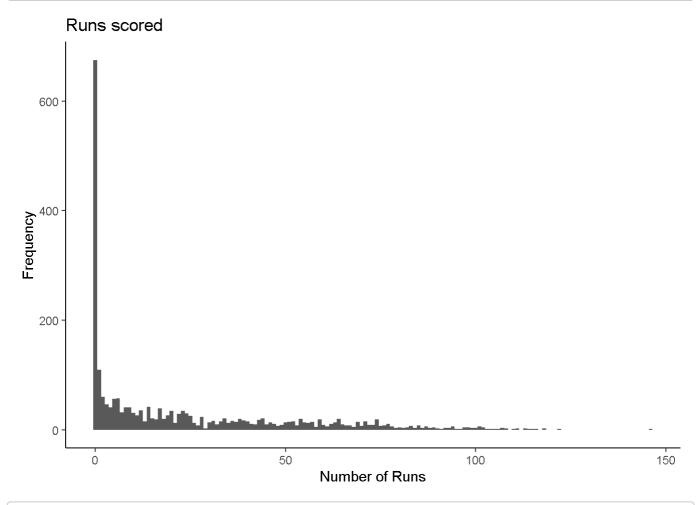
```
name n
1 New York Yankees 9
```

The teams appearing above the top prediction band are always the New York Yankees - a total of 9 times.

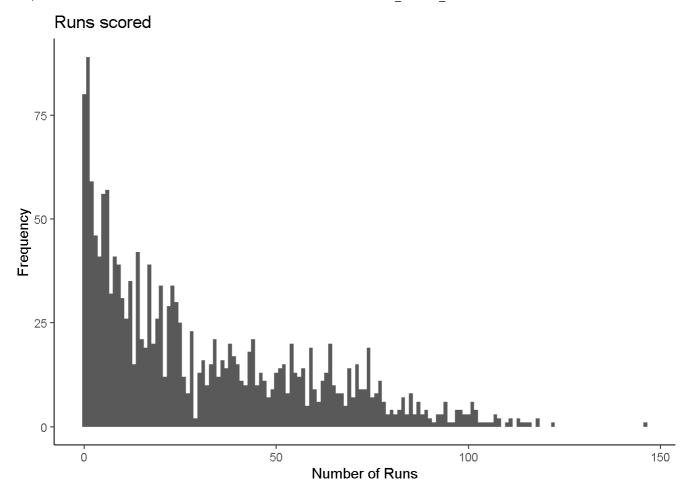
(3a)

```
Playerdata %>%
  ggplot(aes(R)) +
  geom_histogram(binwidth = 1) +
  labs(x = "Number of Runs", y = "Frequency", title = "Runs scored") +
  theme_classic() -> RunsScored

RunsScored
```



```
Playerdata %>%
  filter(H > 0) %>%
  ggplot(aes(R)) +
  geom_histogram(binwidth = 1) +
  labs(x = "Number of Runs", y = "Frequency", title = "Runs scored") +
  theme_classic() -> HitRuns
HitRuns
```



The second data set is more reasonable to use when creating our model as it is pointless to include players who haven't had the chance to score a run in a research question looking at runs.

Including the full set of hits = 0 will skew the data visualization making it harder to read.

(3b)

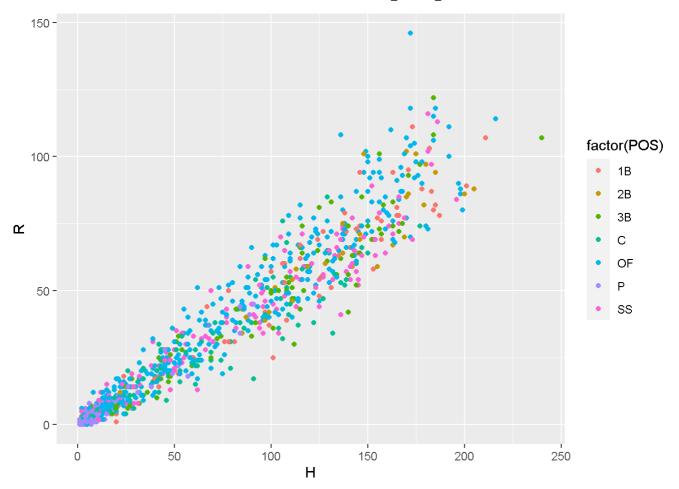
```
Playerdata %>%
  filter(H > 0) %>%
  mutate("yearID" = as.factor(yearID)) -> OnBaseP

glm(formula = R ~ H + as.factor(yearID) + POS + height + age, family = "poisson", data = OnBase
P) -> glm1

glm1

# I'm going to evaluate the data we have checking whether any positions shouldn't be included i
n this model (I'm pretty sure baseball has their own version of a useless role like goalkeeper)

glm1 %>%
  ggplot(aes(x= H, y = R, colour = factor(POS))) +
  geom_point()
```



Have found that POS = P (pitchers) don't get many hits or runs (I think they only get to hit if you rotate through every other player on the team? And I guess because rarely do it as well they usually suck). So I will remove Pitchers from the Hit model.

```
Playerdata %>%
  filter(H > 0) %>%
  mutate("yearID" = as.factor(yearID)) %>%
  filter(POS != "P") -> OnBase

glm(formula = R ~ H + as.factor(yearID) + POS + height + age, family = "poisson", data = OnBas
```

 $e) \rightarrow glm2$

```
Call: glm(formula = R ~ H + as.factor(yearID) + POS + height + age,
    family = "poisson", data = OnBaseP)
Coefficients:
          (Intercept)
                                           H as.factor(yearID)2015
             2.347998
                                    0.013332
                                                            0.024172
                                                                POSC
                POS2B
                                       POS3B
            -0.005030
                                                           -0.061520
                                    0.008146
                POSOF
                                        POSP
                                                               POSSS
             0.069327
                                                           -0.014180
                                   -1.161731
               height
                                          age
            -0.002828
                                    0.005467
Degrees of Freedom: 1739 Total (i.e. Null); 1729 Residual
Null Deviance:
                    44650
Residual Deviance: 6734
                            AIC: 14680
```

(3c)

```
Anova(glm2)
```

```
Analysis of Deviance Table (Type II tests)
Response: R
                 LR Chisq Df Pr(>Chisq)
Н
                  24778.8 1 < 2.2e-16 ***
as.factor(yearID)
                     1.7 1
                              0.197407
POS
                    79.6 5 1.034e-15 ***
                             0.988469
height
                     0.0 1
                              0.008928 **
                     6.8 1
age
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

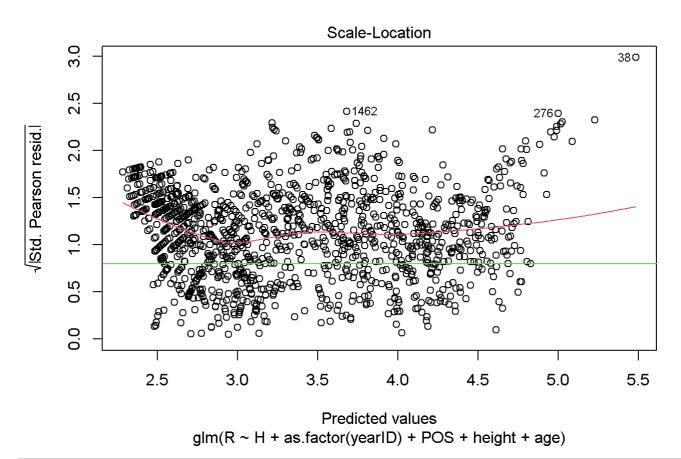
The hypothesis being tested by each p value is whether not the full model is a better method for prediction compared to a reduced model (which would be the NULL hypothesis). Each p value indicates whether or not that variable is a relevant predictor of the number of runs (R) scored.

We can see that the p value for POS (1.034e-15 or 0.00000000000001034) is extremely small, almost 0. So we have strong evidence that including position within our model generates more accuracy than excluding it.

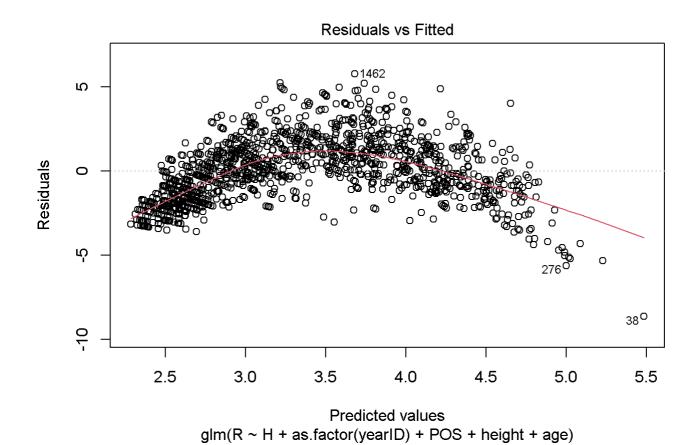
We can also see that the p value for height (0.988469) is the greatest value and is above the typical threshold of 0.05. This indicates that height is not an important predictor of runs scored.

(3d)

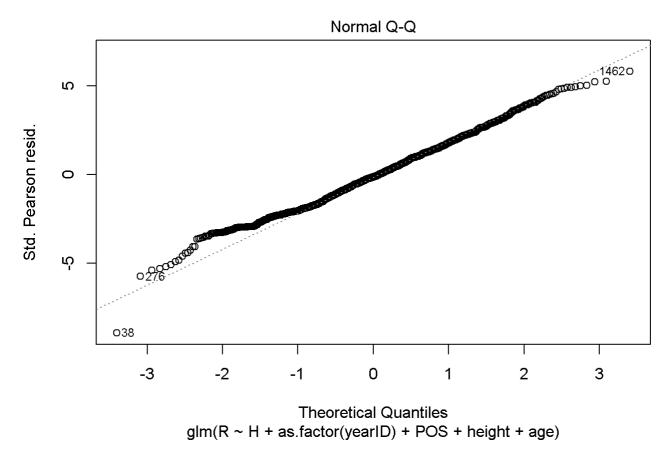
```
plot(glm2,which=3)
abline(h=0.8,col=3)
```



Linearity
plot(glm2,which=1)



Poisson distribution
plot(glm2,which=2)



The first assumption we can test is the dispersion assumption (testing if variance = mean), so the variance should rise when our measured values rise; winning teams with skyrocketing salaries may create an environment where the smaller teams can't keep up with the paychecks - widening the spread of data. We can see the red line, in the dispersion chart above is slightly higher than 0.8 meaning the data is overdispersed.

For our linearity assumption it seems that the data follows a parabolic curve rather than a straight line. Meaning that we will have to look for a better model and test that one.

For the distribution assumption we can look at the QQ plot and see that the data follows the line nicely so we won't have to use robust confidence intervals.

Finally there is the independence assumption, in our data set there is not natural order to how the data is collected so we cannot check this assumption. (Will data points be influenced by the data point before it)

(3e)

```
glmer(R ~ H + yearID + POS + age + (1|teamID), data = OnBase, family = "poisson") -> mmod1
mmod1
glmer(R ~ H + yearID + POS + age + (1|teamID), data = OnBase, family = "poisson", nAGQ = 0) -> mmod2
mmod2

# statistical significance, there are some warnings
# glmer(R ~ H + yearID + POS + age + (1|teamID), data = OnBase, family = "poisson") -> mmod3
# confint(mmod3)
# this code takes really long to run so I commented it when knitting
```

```
Generalized linear mixed model fit by maximum likelihood (Laplace
 Approximation) [glmerMod]
 Family: poisson (log)
Formula: R ~ H + yearID + POS + age + (1 | teamID)
   Data: OnBase
     AIC
              BIC logLik deviance df.resid
12602.48 12655.61 -6291.24 12582.48
Random effects:
Groups Name
                    Std.Dev.
teamID (Intercept) 0.1055
Number of obs: 1500, groups: teamID, 33
Fixed Effects:
                                                                        POSC
(Intercept)
                           yearID2015
                                             POS2B
                                                          POS3B
                                                       0.010591
   2.248067
                0.013191
                             0.015829
                                         -0.001358
                                                                   -0.064989
      POSOF
                   POSSS
                                  age
   0.067566
               -0.013711
                             0.002408
optimizer (Nelder Mead) convergence code: 0 (OK); 0 optimizer warnings; 2 lme4 warnings
Generalized linear mixed model fit by maximum likelihood (Adaptive
  Gauss-Hermite Quadrature, nAGQ = 0) [glmerMod]
 Family: poisson (log)
Formula: R ~ H + yearID + POS + age + (1 | teamID)
   Data: OnBase
                BIC
      ATC
                       logLik deviance df.resid
12602.482 12655.614 -6291.241 12582.482
                                             1490
Random effects:
Groups Name
                    Std.Dev.
teamID (Intercept) 0.1055
Number of obs: 1500, groups: teamID, 33
Fixed Effects:
                                                                        POSC
(Intercept)
                           yearID2015
                                             POS2B
                                                          POS3B
                       Н
   2.248398
                0.013191
                             0.015832
                                         -0.001356
                                                       0.010591
                                                                   -0.064989
      POSOF
                   POSSS
                                  age
   0.067567
                             0.002408
               -0.013710
```

Runs ~ Pois(2.248398 + 0.013191 * H + POS2B * -0.001356 + POS3B * 0.010591 + POSC * -0.064989 + POSOF * 0.067567 + POSSS * -0.013710 + age * 0.002408 + yearID2015 * 0.015832 + u)

```
4-rock ging math 5500%
```

We can see that the standard deviation of the effect of team on number of runs was 0.1055 and average runs are 2.248398.

So this means being in the top club as opposed to an average team a player will score about 0.57 more runs, just based on team as a stand alone. (e(2*0.1055)). Which is a reasonably large difference.

You can check the statistical significance of this effect by using confidence intervals.

(3f)

```
predict(mmod2, newdata=data.frame(age = 30, height = 72, teamID = "BAL", POS = "OF", yearID =
"2015", H = 20), type = "response")
```

```
1
16.76264
```

You would expect a mean of 16.763 runs for such a player.

(4a)

```
Teamdata %>%
    select(!c(teamID, name, park, teamIDBR, teamIDlahman45, teamIDretro, lgID, Rank, franchID, di
vID, WCWin, LgWin, WSWin)) -> DivWinners

DivWinners
set.seed(123)

DivWinners$DivWin %>%
    createDataPartition(p = 0.8, list = FALSE) -> training.samples

DivWinners[training.samples, ] -> train.data
DivWinners[-training.samples, ] -> test.data
```

| | yearID | G | Ghome | W | L | DivWin | R | AB | Н | X2B | ХЗВ | HR | ВВ | SO | SB | cs |
|----------|--------------|--------|-------------|----------|----------|--------|-----|--------------|------|-----|-----|-----|-----|-------------|------------|----------|
| 1 | 1995 | | 72 | 90 | 54 | | | 4814 | | | _ | 168 | | 933 | | 43 |
| 2 | 1995 | 144 | 72 | 71 | 73 | N | 704 | 4837 | 1267 | 229 | 27 | 173 | 574 | 803 | 92 | 45 |
| 3 | 1995 | 144 | 72 | 86 | 58 | Υ | 791 | 4997 | 1399 | 286 | 31 | 175 | 560 | 923 | 99 | 44 |
| 4 | 1995 | 145 | 72 | 78 | 67 | N | 801 | 5019 | 1390 | 252 | 25 | 186 | 564 | 889 | 58 | 39 |
| 5 | 1995 | 145 | 72 | 68 | 76 | N | 755 | 5060 | 1417 | 252 | 37 | 146 | 576 | 767 | 110 | 39 |
| 6 | 1995 | 144 | 72 | 73 | 71 | N | 693 | 4963 | 1315 | 267 | 39 | 158 | 440 | 953 | 105 | 37 |
| 7 | 1995 | 144 | 72 | 85 | 59 | Υ | 747 | 4903 | 1326 | 277 | 35 | 161 | 519 | 946 | 190 | 68 |
| 8 | 1995 | | 72 | 100 | 44 | Υ | 840 | 5028 | 1461 | 279 | 23 | 207 | 542 | 766 | 132 | 53 |
| 9 | 1995 | | 72 | 77 | 67 | N | | 4994 | | _ | | 200 | | | 125 | 59 |
| 10 | 1995 | | 72 | 60 | 84 | N | 654 | 4865 | 1204 | 228 | | 159 | | 987 | 73 | 36 |
| 11 | 1995 | | 71 | 67 | 76 | N | | 4886 | | | | 144 | | | 131 | |
| 12 | 1995 | | 72 | 76 | 68 | N | | 5097 | | | | 109 | | | 176 | |
| 13 | 1995 | | 72 | _ | 74 | | - | 4903 | _ | _ | | 119 | | | 120 | |
| 14 | 1995 | | 72 | 78 56 | 66 | Y | | 4942 | | | | | | 1023 | | |
| 15 | 1995 | | 72 | 56 | 88 | N | | 5005 | | | | 120 | | | 105 | |
| 16 17 | 1995 1995 | | 72 72 | 65 66 | 79 78 | N | | 5000 4905 | | | | 128 | | | 105 120 | _ |
| | | | | | | N | | 4905 | | | | 122 | 400 | | | |
| 18 19 | 1995 1995 | | 73 72 | 79 69 | 65 75 | N N | | 4947 | | | _ | 125 | | 851 994 | | 30 39 |
| 20 | 1995 | | 72 | 67 | 73 77 | N N | | 4916 | | | | 169 | | | 112 | |
| 21 | 1995 | | 72 | 69 | 75 | N | | 4950 | | | 30 | | 497 | 884 | | 25 |
| 22 | 1995 | | 72 | | 86 | N | | 4937 | | | | 125 | | 972 | | 55 |
| 23 | 1995 | | 72 | 70 | 74 | N | | 4950 | | | | | 447 | | 124 | |
| 24 | 1995 | | 73 | 79 | 66 | Υ | | 4996 | | | | 182 | | | 110 | |
| 25 | 1995 | | 72 | 67 | 77 | N | | 4971 | | | | | | 1060 | | |
| 26 | 1995 | | 72 | 62 | 81 | N | | 4779 | | | | 107 | | 920 | | 46 |
| 27 | 1995 | 144 | 72 | 74 | 70 | N | 691 | 4913 | 1304 | 247 | 24 | 138 | 526 | 877 | 90 | 47 |
| 28 | 1995 | 144 | 72 | 56 | 88 | N | 642 | 5036 | 1309 | 275 | 27 | 140 | 492 | 906 | 75 | 16 |
| 29 | 1996 | 162 | 81 | 96 | 66 | Υ | 773 | 5614 | 1514 | 264 | 28 | 197 | 530 | 1032 | 83 | 43 |
| 30 | 1996 | 163 | 82 | 88 | 74 | N | 949 | 5689 | 1557 | 299 | 29 | 257 | 645 | 915 | 76 | 40 |
| 31 | 1996 | 162 | 81 | 85 | 77 | N | 928 | 5756 | 1631 | 308 | 31 | 209 | 642 | 1020 | 91 | 44 |
| 32 | 1996 | 161 | 81 | 70 | 91 | N | 762 | 5686 | 1571 | 256 | 24 | 192 | 527 | 974 | 53 | 39 |
| 33 | 1996 | 162 | 81 | 85 | 77 | N | 898 | 5644 | 1586 | 284 | 33 | 195 | 701 | 927 | 105 | 41 |
| 34 | 1996 | | 81 | 76 | 86 | N | 772 | 5531 | 1388 | 267 | | | | 1090 | | |
| 35 | 1996 | | | | 81 | | | 5455 | | | | | | 1134 | | |
| 36 | 1996 | | 80 | | 62 | | | 5681 | | | | | | 844 | | |
| 37 | 1996 | | 81 | | 79 | | | 5590 | | | | | | 1108 | | |
| 38 | 1996 | | | | 109 | | | 5530 | | | | | | 1268 | _ | 50 |
| 39 | 1996 | | | | 82 | | | 5498 | | | | | | 1122 | | |
| 40 | 1996 | | 81 | | 80 | N | | 5508 5542 | | | | | | 1057 | | |
| 41 42 | 1996 1996 | | 80 91 | | 86 72 | | _ | 5538 | | | | | | 943 1190 | | |
| 42 | 1996 | | 81 82 | | 84 | | | 5673 | | | | | 576 | 958 | | |
| 44 | 1996 | | | | 82 | | | 5662 | | | | | | 986 | | |
| 45 | 1996 | | 81 | | 74 | | | 5505 | | | | | | 1077 | | |
| 46 | 1996 | | 80 | | 70 | | | 5628 | | | | | | 909 | | 46 |
| 47 | 1996 | | 81 | | 91 | | | 5618 | | | | | | 1069 | | 48 |
| 48 | 1996 | | 81 | | 84 | | | 5630 | | | | | | 1114 | | 35 |
| 49 | 1996 | | 81 | 67 | 95 | N | | 5499 | | | | | | 1092 | | |
| 50 | 1996 | | 80 | 73 | 89 | N | | 5665 | | | | | | 989 | | |
| 51 | 1996 | | 81 | | 71 | | | 5655 | | | | | | 1014 | | |
| 52 | 1996 | | 81 | 85 | 76 | | | 5668 | | | | | | 1052 | | |
| 53 | 1996 | | 82 | 68 | 94 | N | 752 | 5533 | 1400 | 245 | 21 | 153 | 615 | 1189 | 113 | 53 |
| P596 | essing goodh | 1:1100 | % 81 | 88 | 74 | Υ | 759 | 5502 | 1468 | 281 | 31 | 142 | 495 | 1089 | 149 | 58 |
| | | | | | | | | | | | | | | | | |

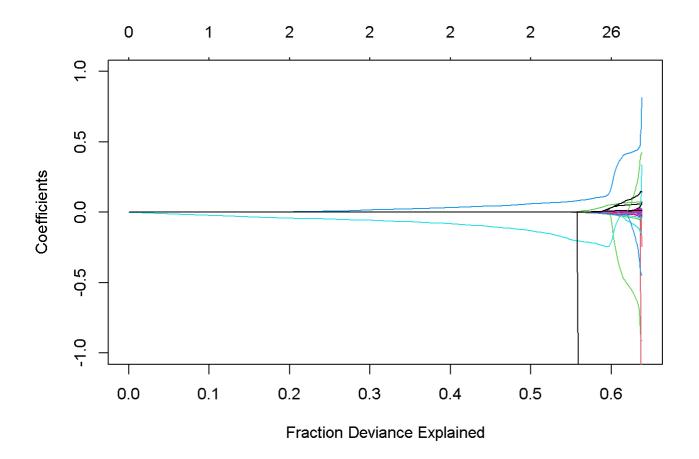
| 649 | 3365256 103 102 | 172253778 | 6890151.1 | 25 |
|-----|-----------------|-----------|-----------|----|
| 650 | 3444490 100 99 | 143053500 | 4614629.0 | 31 |
| 651 | 1286163 93 94 | 57097310 | 2039189.6 | 28 |
| 652 | 2710402 106 105 | 176038723 | 6070300.8 | 29 |
| 653 | 3392099 111 110 | 138701700 | 4782817.2 | 29 |
| 654 | 2481938 100 98 | 141652646 | 5448178.7 | 26 |
| | | | | |

(4b)

```
as.vector(train.data$DivWin) -> divwin1
model.matrix(~ . -1, train.data[,-c(6)]) -> divpredict1

glmnet(divpredict1, divwin1, family = "binomial") -> divwinfit

plot(divwinfit, xvar = "dev",ylim = c(0,0))
```



(4c)

```
divwinfit
```

```
# First over 0.5: 21  2 50.02 0.038030
# First over 0.6: 53 26 60.13 0.001937

coef(divwinfit, s = 0.038030) -> divwin4c50
divwin4c50@Dimnames[[1]][1+divwin4c50@i]

coef(divwinfit, s = 0.001937) -> divwin4c60
divwin4c60@Dimnames[[1]][1+divwin4c60@i]
```

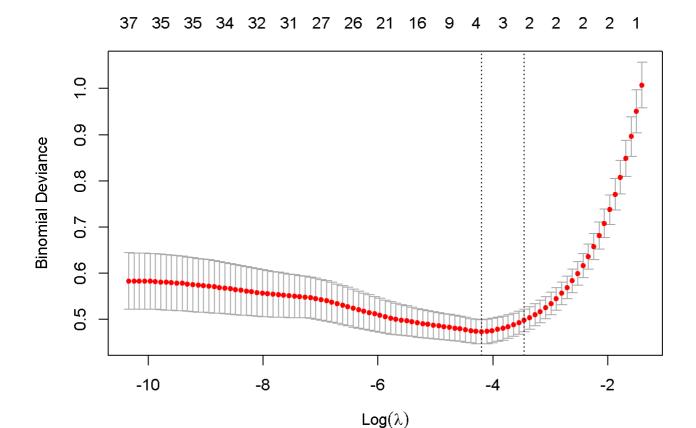
```
glmnet(x = divpredict1, y = divwin1, family = "binomial")
   Df %Dev
              Lambda
    0 0.00 0.244500
1
2
    1 6.30 0.222800
3
    1 11.67 0.203000
    1 16.31 0.184900
5
    2 20.45 0.168500
    2 24.13 0.153500
6
7
    2 27.40 0.139900
    2 30.31 0.127500
8
    2 32.92 0.116100
10
   2 35.27 0.105800
11 2 37.38 0.096430
    2 39.29 0.087860
   2 41.02 0.080060
14 2 42.58 0.072940
   2 44.00 0.066460
16 2 45.27 0.060560
   2 46.43 0.055180
18 2 47.47 0.050280
19 2 48.41 0.045810
20 2 49.26 0.041740
21 2 50.02 0.038030
22 2 50.70 0.034650
23 3 51.31 0.031580
24 3 51.91 0.028770
25
   3 52.45 0.026220
26 3 52.92 0.023890
    3 53.34 0.021760
27
28 3 53.71 0.019830
29 3 54.04 0.018070
30 3 54.33 0.016460
31 3 54.58 0.015000
32 4 54.80 0.013670
33
    6 55.04 0.012450
   7 55.29 0.011350
   8 55.74 0.010340
   9 56.15 0.009421
37 9 56.51 0.008584
38 12 56.84 0.007822
39 14 57.18 0.007127
40 14 57.52 0.006494
41 15 57.81 0.005917
42 15 58.10 0.005391
43 16 58.36 0.004912
44 18 58.61 0.004476
45 19 58.84 0.004078
46 19 59.04 0.003716
47 19 59.21 0.003386
48 20 59.36 0.003085
49 21 59.49 0.002811
50 22 59.63 0.002561
Processing math: 100%02334
```

```
52 25 59.93 0.002126
53 26 60.13 0.001937
54 27 60.29 0.001765
55 26 60.44 0.001609
56 27 60.61 0.001466
57 27 60.80 0.001335
58 28 61.04 0.001217
59 28 61.28 0.001109
60 28 61.48 0.001010
61 27 61.62 0.000920
62 29 61.79 0.000839
63 29 61.98 0.000764
64 30 62.21 0.000696
65 31 62.42 0.000634
66 31 62.60 0.000578
67 31 62.76 0.000527
68 31 62.89 0.000480
69 31 63.00 0.000437
70 32 63.10 0.000398
71 32 63.18 0.000363
72 32 63.25 0.000331
73 32 63.30 0.000301
74 32 63.35 0.000275
75 33 63.40 0.000250
76 33 63.43 0.000228
77 34 63.46 0.000208
78 34 63.49 0.000189
79 34 63.51 0.000172
80 34 63.53 0.000157
81 34 63.55 0.000143
82 34 63.56 0.000130
83 34 63.57 0.000119
84 35 63.58 0.000108
85 35 63.59 0.000099
86 35 63.60 0.000090
87 35 63.61 0.000082
88 36 63.62 0.000075
89 36 63.65 0.000068
90 35 63.66 0.000062
91 35 63.67 0.000056
92 35 63.68 0.000051
93 36 63.69 0.000047
94 37 63.73 0.000043
95 37 63.78 0.000039
96 37 63.81 0.000035
97 37 63.81 0.000032
[1] "(Intercept)" "W"
                                                                "L"
                                                 "W"
 [1] "(Intercept)" "yearID"
                                  "Ghome"
                                  "X2B"
                                                                "HR"
 [6] "AB"
                    "H"
                                                 "X3B"
                                  "SB"
                                                 "CS"
[11]
    "BB"
                    "S0"
                                                                "HBP"
                                                 "SV"
                                  "CG"
                                                               "HA"
[16] "SF"
                    "RA"
                    "BBA"
                                  "SOA"
                                                 "DP"
                                                                "FP"
[21] "HRA"
                    "PPF"
[26] "attendance"
                                  "rostersize"
```

We can see from the plot that 50% of the deviance can be explained by 2 parameters(Wins(W) and Losses(L)), and 60% can be explained by 26 parameters ("yearID", "Ghome", "W", "L", "AB", "H", "X2B", "X3B", "HR", "BB", "SO", "SB", "CS", "HBP", "SF", "RA", "CG", "SV", "HA", "HRA", "BBA", "SOA", "DP", "FP", "attendance", "PPF", "roster size".)

(4d)

set.seed(312)
cv.glmnet(divpredict1, divwin1, family = "binomial") -> crossvalidation
plot(crossvalidation)



coef(divwinfit, s = crossvalidation\$lambda.1se) -> divwincoef

divwincoef

setdiff(divwincoef@Dimnames[[1]][1+divwincoef@i], divwin4c50@Dimnames[[1]][1+divwin4c50@i])

```
38 x 1 sparse Matrix of class "dgCMatrix"
(Intercept) 3.674316e+00
yearID
G
Ghome
W
             6.536722e-02
L
             -1.417659e-01
R
AB
Н
X2B
X3B
HR
BB
S0
SB
CS
HBP
SF
RA
ER
ERA
CG
SHO
SV
IPouts
HA
HRA
BBA
SOA
Ε
DΡ
FΡ
attendance
             1.444007e-10
BPF
PPF
Rostercost
meansalary
rostersize
[1] "attendance"
```

Looking at the plot of cross validation we can see that around 3~5 variables seems to be where the binomial deviance is at its lowest. So going off this I have cross validated against our previous previous analysis where s = 0.038030 (lamda value showing 2 variables) as this is closer to our amount of values suggested here. Looking at our coefficient results we can see that "attendance" is suggested to be included within the model. So for my conservative model I will include: W, L, attendance.

(4e)

```
DivWinners
train.data
test.data

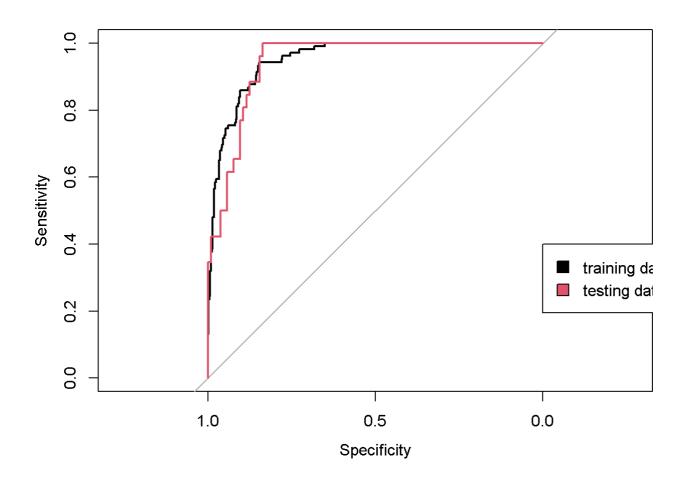
glm(as.factor(DivWin) ~ W + L + attendance, data = train.data, family = "binomial") -> train.mo
del4

train.model4 %>%
  predict(type = "response") -> predtrain

train.model4 %>%
  predict(newdata = test.data, type = "response") -> predtest

roc(response = train.data$DivWin, predictor = predtrain, plot = TRUE, auc = TRUE) -> roctrain

roc(response = test.data$DivWin, predictor = predtest, plot = TRUE, auc = TRUE, add = TRUE, col = 2)
legend(0,0.4,legend=c("training data","testing data"),fill=1:2)
```



```
2148808 94 95 85556990 3168777.4
                                                  27
583
      2423852 100 101 180944967 5654530.2
                                                  32
585
595
      2080145 107 106
                                                  29
                      61834000 2132206.9
604
      2726048 97 98 172284750 6891390.0
                                                  25
605
      2153585 97 99
                      72256200 2580578.6
                                                  28
      2708549 104 103 112107025 4152112.0
                                                  27
606
      1831080 98 98 111693000 4295884.6
615
                                                  26
      2498596 99 97
                      88892499 3065258.6
                                                  29
616
      2193581 92 94 122208700 4888348.0
                                                  25
618
628
      2955434 108 106 188545761 6501578.0
                                                  29
      3016142 95 95 137251333 5278897.4
637
                                                  26
      1521506 90 91
                      86806234 2893541.1
                                                  30
644
      2351422 99 99 101424814 3756474.6
                                                  27
647
652
      2710402 106 105 176038723 6070300.8
                                                  29
roc.default(response = test.data$DivWin, predictor = predtest, auc = TRUE, plot = TRUE, add
= TRUE, col = 2)
Data: predtest in 104 controls (test.data$DivWin N) < 26 cases (test.data$DivWin Y).
Area under the curve: 0.9442
```

We can see that these ROC curves are very close, almost on top of each other, so this indicates that the model did not overfit the data.

(4f)

```
coords(roctrain, "b", best.method = "youden", transpose = TRUE) -> youdenroctrain

youdenroctrain

ifelse(predict(train.model4, newdata = train.data, type = "response") >= 0.1836071, "Y", "N") ->
    confusiontrain

table(confusiontrain, train.data$DivWin)

ifelse(predict(train.model4, newdata = test.data, type = "response") >= 0.1836071, "Y", "N") ->
    test.data$confusiontest

table(test.data$confusiontest, test.data$DivWin)
```

Looking at the confusion matrix for testing data we can see that 25/26 = 96.15% DivWinners were identified correctly (Y). We can also see that 87/104 = 83.65% of Non-DivWinners were identified correctly. So this is a great model in terms of both false positives and false negatives.

(4g)

```
merge(x = test.data, y = Teamdata[ , c("yearID","W","L","DivWin","attendance","divID")], by = c
("yearID", "W", "L", "DivWin", "attendance"), all.x=TRUE) -> testdivID
testdivID %>%
  filter(divID == "C") -> testdivIDC
testdivID %>%
  filter(divID == "W") -> testdivIDW
testdivID %>%
  filter(divID == "E") -> testdivIDE
table(testdivIDC$confusiontest, testdivIDC$DivWin) -> tableC
sensitivity(tableC) + specificity(tableC) -> divC
table(testdivIDW$confusiontest, testdivIDW$DivWin) -> tableW
sensitivity(tableW) + specificity(tableW) -> divW
table(testdivIDE$confusiontest, testdivIDE$DivWin) -> tableE
sensitivity(tableE) + specificity(tableE) -> divE
data.frame(div = c("C", "W", "E"),
           SS = c(1.794715 , 1.866667, 1.757576)) \rightarrow divdf
divdf %>%
  ggplot(aes(x = div, y = SS)) +
  geom_col()
```

