# Time Series Analysis 2

#### Blake Pappas

#### December 19, 2023

#### Seasonal Component Estimation

Now let's consider the situation that a time series consists of seasonal components only (assuming the trend has been estimated/removed). That is

$$Y_t = s_t + \eta_t$$
.

with  $\{s_t\}$  having period d (i.e.,  $s_t = s_{t+jd}$  for all integers j and t).  $\sum_{t=1}^{d} s_t = 0$  and  $\mathbb{E}[\eta_t] = 0$ . We can use a harmonic regression or a seasonal factor model to estimate the seasonal components or to use seasonal-differencing to remove the seasonality.

#### Harmonic Regression

A harmonic regression model has the form

$$s_t = \sum_{i=1}^k A_k \cos(2\pi f_j + \phi_j).$$

For each  $j = 1, \dots, k$ :

- $A_i > 0$  is the amplitude of the jth cosine wave.
- $f_j$  controls the the frequency of the j-th cosine wave (how often waves repeat).
- $\phi_j \in [-\pi, \pi]$  is the *phase* of the j-th wave (where it starts)

The above can be expressed as

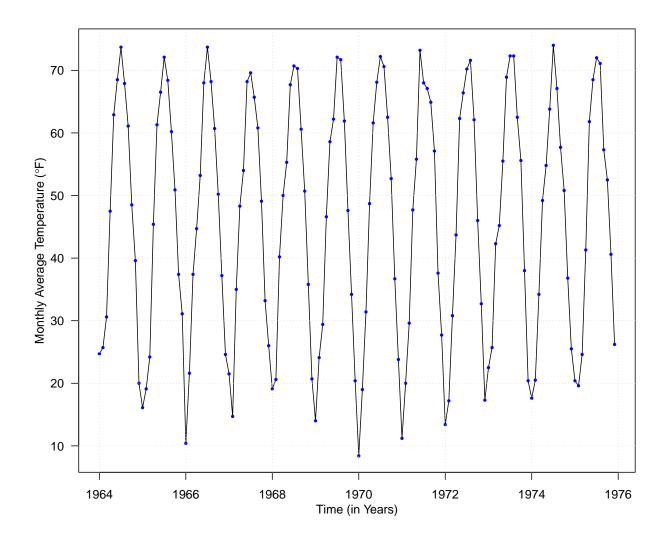
$$\sum_{j=1}^{k} (\beta_{1j} \cos(2\pi f_j) + \beta_{2j} \sin(2\pi f_j)),$$

where  $\beta_{1j} = A_j \cos(\phi_j)$  and  $\beta_{2j} = A_j \sin(\phi_j)$ . Therefore, if the frequencies  $\{f_j\}_{j=1}^k$  are known, we can use regression techniques to estimate the parameters  $\{\beta_{1j}, \beta_{2j}\}_{j=1}^k$  by treating  $\{\cos(2\pi f_j)\}_{j=1}^k$  and  $\{\sin(2\pi f_j)\}_{j=1}^k$  as predictor variables.

Let's use the monthly average temperature (in degrees Fahrenheit) recorded in Dubuque, IA from Jan. 1964 - Dec. 1975.

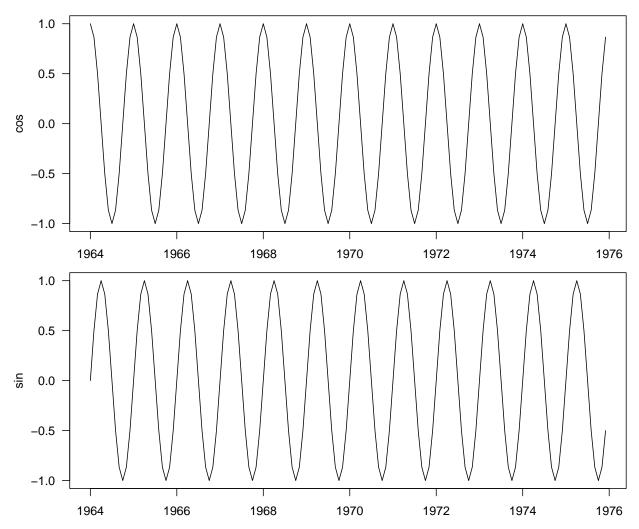
```
# install.packages("TSA")
library(TSA)
data(tempdub)
time <- as.numeric(time(tempdub))
par(mar = c(4, 4, 0.8, 0.6))</pre>
```

```
plot(time, tempdub, type = "l", las = 1, xlab = "", ylab = "")
points(time, tempdub, pch = 16, col = "blue", cex = 0.6)
grid()
mtext("Time (in Years)", side = 1, line = 2)
mtext(expression(paste("Monthly Average Temperature (", degree, "F)")), side = 2, line = 2)
```



First, we need to set up the harmonics (assuming yearly cycle):

```
harmonics <- harmonic(tempdub, 1)
time <- as.numeric(time(tempdub))
par(mfrow = c(2, 1), las = 1, mar = c(2, 4, 0.8, 0.6))
plot(time, harmonics[, 1], type = "l", ylab = "cos")
plot(time, harmonics[, 2], type = "l", ylab = "sin")</pre>
```



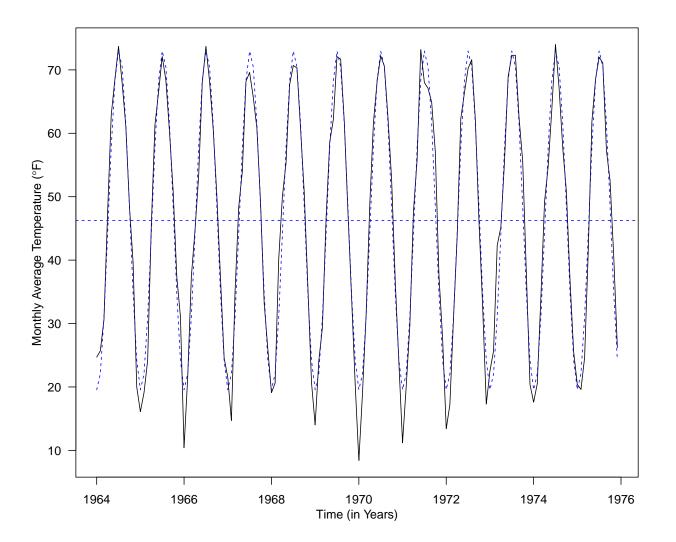
Next, perform a linear regression using the harmonics as the predictors:

```
harReg <- lm(tempdub ~ harmonics)
summary(harReg)</pre>
```

```
##
## Call:
## lm(formula = tempdub ~ harmonics)
##
## Residuals:
##
        Min
                                            Max
                  1Q
                       Median
                                    ЗQ
   -11.1580 -2.2756 -0.1457
                                2.3754
                                       11.2671
##
##
##
  Coefficients:
##
                        Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                         46.2660
                                     0.3088 149.816 < 2e-16 ***
                                     0.4367 -61.154 < 2e-16 ***
## harmonicscos(2*pi*t) -26.7079
                                            -4.968 1.93e-06 ***
## harmonicssin(2*pi*t)
                         -2.1697
                                     0.4367
##
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
##
```

```
## Residual standard error: 3.706 on 141 degrees of freedom
## Multiple R-squared: 0.9639, Adjusted R-squared: 0.9634
## F-statistic: 1882 on 2 and 141 DF, p-value: < 2.2e-16</pre>
```

```
par(mar = c(3.6, 3.6, 0.8, 0.6))
plot(time, tempdub, type = "l", las = 1, xlab = "", ylab = "")
mtext("Time (in Years)", side = 1, line = 2)
mtext(expression(paste("Monthly Average Temperature (", degree, "F)")), side = 2, line = 2)
time <- as.numeric(time(tempdub))
lines(time, harReg$fitted.values, col = "blue", lty = 2)
abline(h = harReg$coefficients[1], lty = 2, col = "blue")</pre>
```



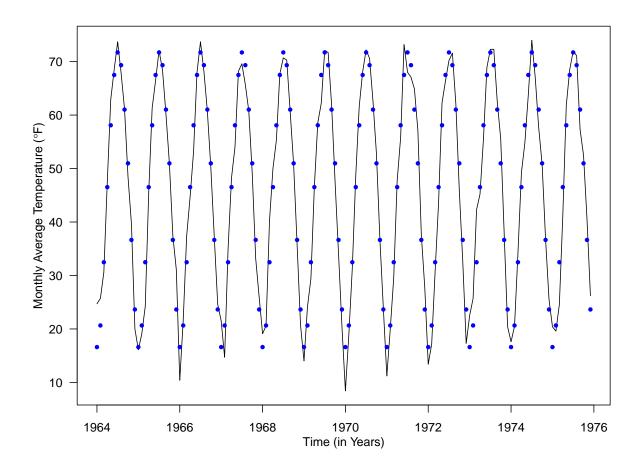
#### **Seasonal Factors**

Harmonic regression assumes the seasonal pattern has a regular shape (i.e. the height of the peaks is the same as the depth of the troughs). Assuming the seasonal pattern repeats itself every d time points, a less restrictive approach is to model it as

```
s_t = \left\{ \begin{array}{ll} \beta_1 & \text{for } t=1,1+d,1+2d,\cdots; \\ \beta_2 & \text{for } t=2,2+d,2+2d,\cdots; \\ \vdots & \vdots; \\ \beta_d & \text{for } t=d,2d,3d,\cdots. \end{array} \right.
```

```
month = season(tempdub)
season_means <- lm(tempdub ~ month - 1)
summary(season_means)</pre>
```

```
##
## Call:
## lm(formula = tempdub ~ month - 1)
## Residuals:
##
       Min
                1Q Median
                                3Q
                                       Max
## -8.2750 -2.2479 0.1125 1.8896
                                   9.8250
##
## Coefficients:
##
                  Estimate Std. Error t value Pr(>|t|)
## monthJanuary
                    16.608
                                0.987
                                        16.83
                                                <2e-16 ***
## monthFebruary
                    20.650
                                0.987
                                        20.92
                                                <2e-16 ***
## monthMarch
                    32.475
                                0.987
                                        32.90
                                                <2e-16 ***
## monthApril
                    46.525
                                0.987
                                        47.14
                                                <2e-16 ***
## monthMay
                    58.092
                                0.987
                                        58.86
                                                <2e-16 ***
## monthJune
                    67.500
                                0.987
                                        68.39
                                                <2e-16 ***
                                0.987
                                        72.66
                                                <2e-16 ***
## monthJuly
                    71.717
                                        70.25
## monthAugust
                    69.333
                                0.987
                                                <2e-16 ***
## monthSeptember
                                0.987
                                        61.83
                                                <2e-16 ***
                    61.025
                                        51.65
                                                <2e-16 ***
## monthOctober
                    50.975
                                0.987
## monthNovember
                    36.650
                                0.987
                                        37.13
                                                <2e-16 ***
## monthDecember
                    23.642
                                        23.95
                                                <2e-16 ***
                                0.987
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
## Residual standard error: 3.419 on 132 degrees of freedom
## Multiple R-squared: 0.9957, Adjusted R-squared: 0.9953
## F-statistic: 2569 on 12 and 132 DF, p-value: < 2.2e-16
plot(time, tempdub, type = "l", las = 1, xlab = "", ylab = "")
mtext("Time (in Years)", side = 1, line = 2)
mtext(expression(paste("Monthly Average Temperature (", degree, "F)")), side = 2, line = 2)
points(time, season_means\fitted.values, col = "blue", pch = 16, cex = 0.8)
```



# Let's Put Trend and Seasonal Variation Together

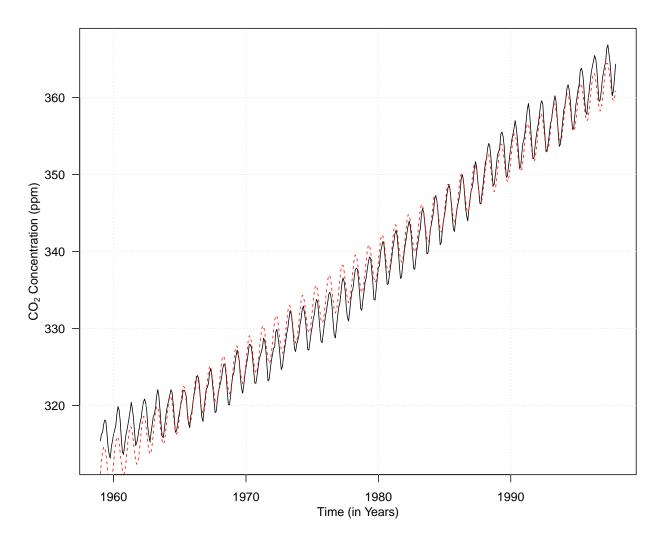
Here we use the  $CO_2$  concentration time series as an example. First, we perform a linear regression with both time and the harmonics as the covariates.

```
time <- as.numeric(time(co2))
harmonics <- harmonic(co2, 1)

lm_trendSeason <- lm(co2 ~ time + harmonics)
summary(lm_trendSeason)</pre>
```

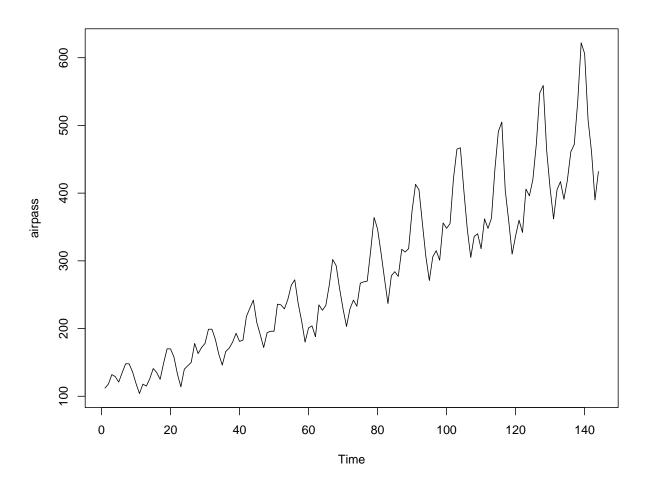
```
##
## Call:
## lm(formula = co2 ~ time + harmonics)
##
## Residuals:
## Min 1Q Median 3Q Max
## -3.433 -1.323 -0.282 1.221 4.615
##
```

```
## Coefficients:
##
                         Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                       -2.256e+03 1.391e+01 -162.155 < 2e-16 ***
                        1.311e+00 7.033e-03 186.382 < 2e-16 ***
## time
## harmonicscos(2*pi*t) -3.889e-01 1.120e-01
                                               -3.474 0.00056 ***
## harmonicssin(2*pi*t)
                       2.772e+00 1.120e-01
                                               24.760 < 2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.712 on 464 degrees of freedom
## Multiple R-squared: 0.987, Adjusted R-squared: 0.9869
## F-statistic: 1.173e+04 on 3 and 464 DF, p-value: < 2.2e-16
par(mar = c(3.8, 4, 0.8, 0.6))
plot(time, co2, type = "1", las = 1, xlab = "", ylab = "")
# points(co2, col = "blue", pch = 16, cex = 0.25)
mtext("Time (in Years)", side = 1, line = 2)
mtext(expression(paste("CO"[2], " Concentration (ppm)")), side = 2, line = 2.5)
grid()
lines(time, lm_trendSeason$fitted.values, col = "red", lty = 2)
```



### Read the Data

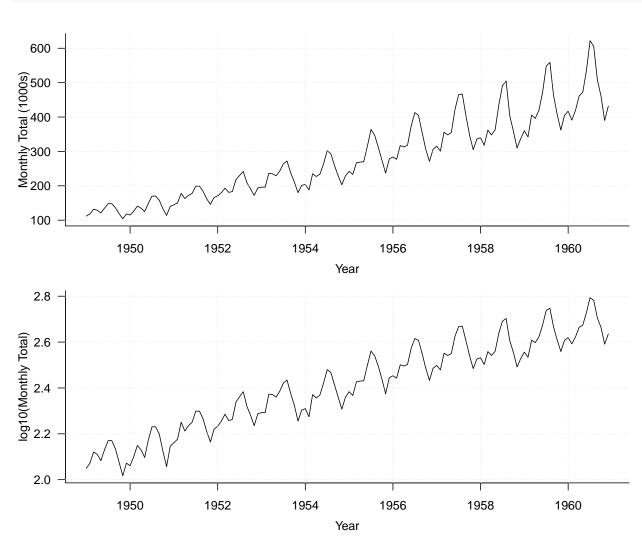
```
airpass <- read.csv("airpass.csv", header = FALSE)
ts.plot(airpass)</pre>
```



```
## Create the Time Variable
yr <- seq(1949, 1960 + 11 / 12, by = 1 / 12)
```

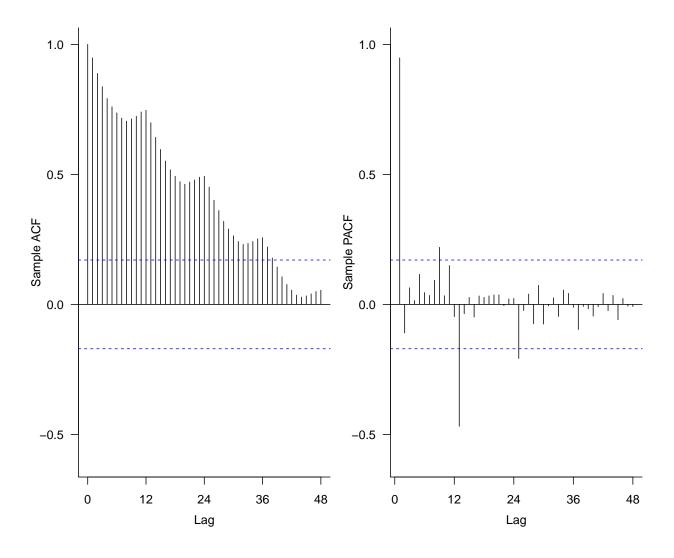
### Plot the Time Series

```
## Take a Log (To the Base 10) of the Air Passenger Data
log.airpass <- log10(airpass)
plot(yr, log.airpass$V1, type = "l", xlab = "Year", ylab = "log10(Monthly Total)")
grid()</pre>
```



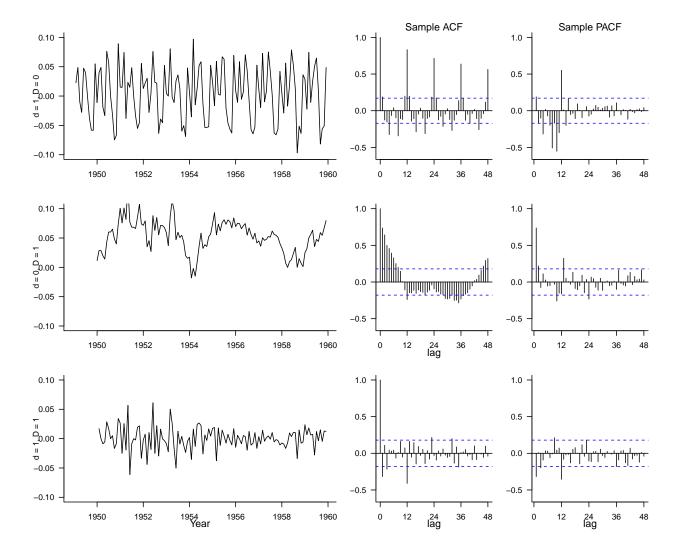
# Plot Sample ACF and PACF

```
log.shortair <- log.airpass$V1[1:132]
shortyears <- yr[1:132]
par(bty = "L", mar = c(3.6, 3.5, 0.8, 0.6), mgp = c(2.4, 1, 0), las = 1, mfrow = c(1, 2))
stats::acf(log.shortair, ylab = "Sample ACF", main = "", lag.max = 48, ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
stats::pacf(log.shortair, ylab = "Sample PACF", main = "", lag.max = 48, ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))</pre>
```



### Trying Different Orders of Differencing

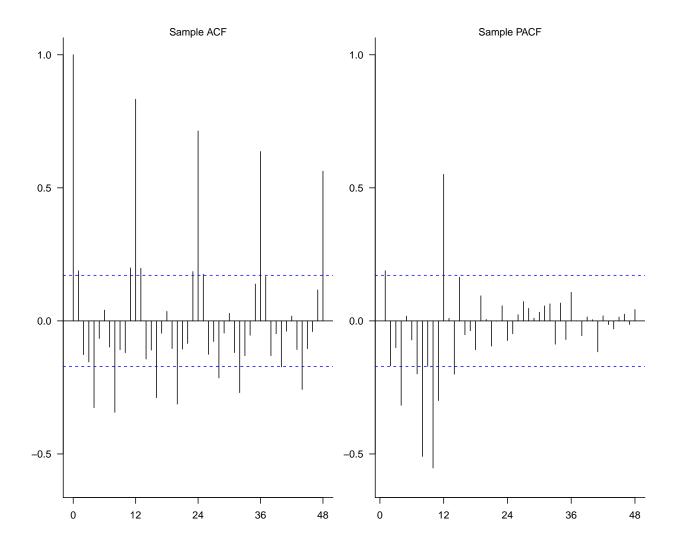
```
mtext("Sample ACF", side = 3, line = 0, cex = 0.8)
stats::pacf(diff.1.0, lag.max = 48, ylab = "", xlab = "", main = "", ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
mtext("Sample PACF", side = 3, line = 0, cex = 0.8)
plot(shortyears[-c(1:12)], diff.0.1, xlab = "", ylab = "d = 0, D = 1",
     type = "l", vlim = c(-0.1, 0.1), vlim = range(shortyears)
stats::acf(diff.0.1, lag.max = 48, ylab = "", xlab = "", main = "", ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
mtext("lag", side = 1, line = 1.8, cex = 0.8)
stats::pacf(diff.0.1, lag.max = 48, ylab = "", xlab = "", main = "", ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
plot(shortyears[-c(1:13)], diff.1.1, xlab = "", ylab = "d = 1, D = 1",
     type = "l", ylim = c(-0.1, 0.1), xlim = range(shortyears))
mtext("Year", side = 1, line = 1.8, cex = 0.8)
stats::acf(diff.1.1, lag.max = 48, ylab = "", xlab = "", main = "", ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
mtext("lag", side = 1, line = 1.8, cex = 0.8)
stats::pacf(diff.1.1, lag.max = 48, ylab = "", xlab = "", main = "", ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
mtext("lag", side = 1, line = 1.8, cex = 0.8)
```



### Show the ACF and PACF for the d = 1, D = 0 Case

```
par(mfrow = c(1, 2), cex = 0.8, bty = "L", mar = c(3.6, 3, 1, 0.6), mgp = c(2.4, 1, 0), las = 1)
stats::acf(diff.1.0, lag.max = 48, ylab = "", xlab = "", main = "", ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
mtext("Sample ACF", side = 3, cex = 0.8)

stats::pacf(diff.1.0, lag.max = 48, ylab = "", xlab = "", main = "", ylim = c(-0.6, 1), xaxt = "n")
axis(side = 1, at = seq(0, 48, 12))
mtext("Sample PACF", side = 3, cex = 0.8)
```

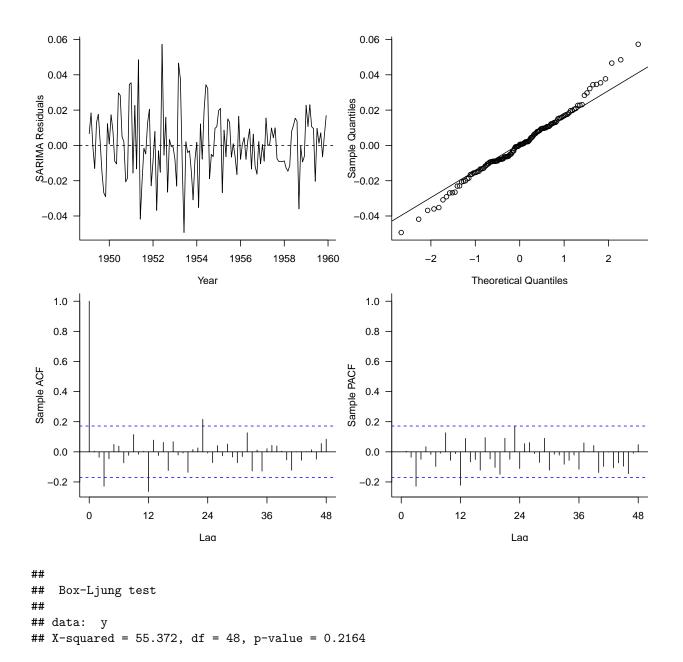


A Useful Function for the Model Diagnostics (Courtesy of Peter Craigmile at  $\operatorname{OSU})$ 

```
plot(x, y, type = "1", ...)
if (mean.line) abline(h = 0, lty = 2)
qqnorm(y, main = "", las = 1); qqline(y)
if (is.null(lags)) {
  acf(y, main = "", lag.max = lag.max, xlim = c(0, lag.max), ylim = acf.ylim,
      ylab = "Sample ACF", las = 1)
 pacf(y, main = "", lag.max = lag.max, xlim = c(0, lag.max), ylim = acf.ylim,
      ylab = "Sample PACF", las = 1)
else {
 stats::acf(y, main = "", lag.max = lag.max, xlim = c(0, lag.max), ylim = acf.ylim,
      ylab = "Sample ACF", xaxt = "n", las = 1)
 axis(side = 1, at = lags)
 stats::pacf(y, main = "", lag.max = lag.max, xlim = c(0, lag.max), ylim = acf.ylim,
       ylab = "Sample PACF", xaxt = "n", las = 1)
 axis(side = 1, at = lags)
Box.test(y, lag.max, type = "Ljung-Box")
```

#### Fitting the SARIMA $(1,1,0) \times (1,0,0)$ Model

```
fit1 \leftarrow arima(diff.1.0, order = c(1, 0, 0), seasonal = list(order = c(1, 0, 0), period = 12))
fit1
##
## Call:
## arima(x = diff.1.0, order = c(1, 0, 0), seasonal = list(order = c(1, 0, 0),
##
       period = 12))
## Coefficients:
##
             ar1
                    sar1 intercept
##
         -0.2667 0.9291
                             0.0039
## s.e.
        0.0865 0.0235
                             0.0096
##
## sigma^2 estimated as 0.0003298: log likelihood = 327.27, aic = -648.54
Box.test(fit1$residuals, lag = 48, type = "Ljung-Box")
## Box-Ljung test
## data: fit1$residuals
## X-squared = 55.372, df = 48, p-value = 0.2164
par(mfrow = c(2, 2), cex = 0.8, bty = "L", mar = c(3.6, 4, 0.8, 0.6),
    mgp = c(2.8, 1, 0), las = 1)
plot.residuals(shortyears[-1], resid(fit1), lag.max = 48,
               ylab = "SARIMA Residuals", xlab = "Year", lags = seq(0, 48, 12))
```



Fitting the SARIMA $(0,1,0) \times (1,0,0)$  Model

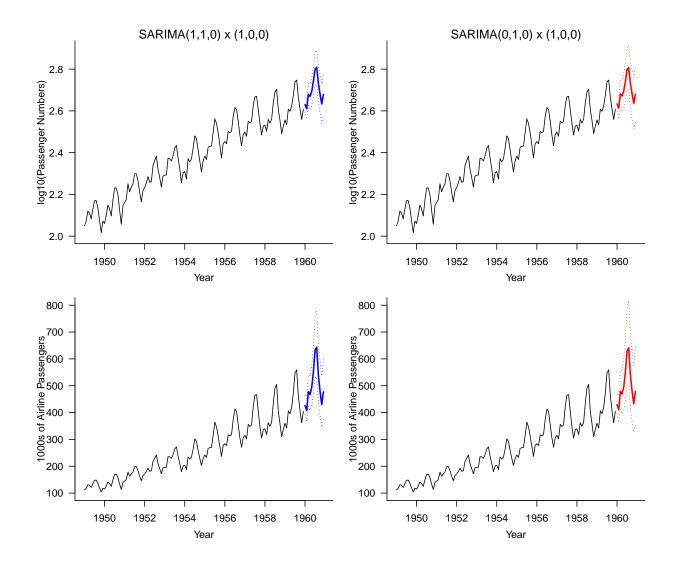
```
(fit2 \leftarrow arima(diff.1.0, seasonal = list(order = c(1, 0, 0), period = 12)))
##
## Call:
  arima(x = diff.1.0, seasonal = list(order = c(1, 0, 0), period = 12))
##
##
   Coefficients:
##
                  intercept
           sar1
##
         0.9081
                     0.0040
## s.e.
         0.0278
                     0.0108
```

```
##
## sigma^2 estimated as 0.0003616: log likelihood = 322.75,
Box.test(fit2$residuals, lag = 48, type = "Ljung-Box")
##
##
    Box-Ljung test
##
## data: fit2$residuals
## X-squared = 80.641, df = 48, p-value = 0.002209
par(mfrow = c(2, 2), cex = 0.8, bty = "L", mar = c(3.6, 4, 0.8, 0.6),
    mgp = c(2.8, 1, 0), las = 1)
plot.residuals(shortyears[-1], resid(fit2), lag.max = 48,
                 ylab = "SARIMA Residuals", xlab = "Year", lags = seq(0, 48, 12))
  0.06
                                                       0.06
                                                                                                00
  0.04
                                                       0.04
SARIMA Residuals
0.00
0.00
20.0
                                                     -0.04
                                                      -0.04
                                                                0
                                                             0
           1950
                                                                  -2
                                                                                 0
                                                                                                2
                  1952
                         1954
                                 1956
                                        1958
                                                1960
                                                                         -1
                            Year
                                                                         Theoretical Quantiles
   1.0
                                                        1.0
                                                        8.0
   8.0
   0.6
                                                        0.6
                                                    Sample PACF
Sample ACF
   0.4
                                                        0.4
   0.2
                                                       0.2
   0.0
                                                        0.0
  -0.2
                                                       -0.2
         0
                                                             0
                  12
                            24
                                      36
                                                48
                                                                       12
                                                                                 24
                                                                                          36
                                                                                                    48
                            Lag
                                                                                Lag
##
##
    Box-Ljung test
##
## data: y
## X-squared = 80.641, df = 48, p-value = 0.002209
```

#### Forecasting the 1960 Data

```
## Fit the First Full Model
fit1 <- arima(log.shortair, order = c(1, 1, 0),</pre>
                     seasonal = list(order = c(1, 0, 0), period = 12))
fit1
##
## Call:
## arima(x = log.shortair, order = c(1, 1, 0), seasonal = list(order = c(1, 0,
       0), period = 12))
##
## Coefficients:
             ar1
                    sar1
         -0.2665 0.9298
##
## s.e. 0.0866 0.0233
##
## sigma^2 estimated as 0.0003299: log likelihood = 327.19, aic = -650.38
## Fit the Second Full Model
fit2 \leftarrow arima(log.shortair, order = c(0, 1, 0),
                     seasonal = list(order = c(1, 0, 0), period = 12))
fit2
##
## arima(x = log.shortair, order = c(0, 1, 0), seasonal = list(order = c(1, 0, 1, 0))
       0), period = 12))
##
## Coefficients:
##
           sar1
##
         0.9088
## s.e. 0.0276
## sigma^2 estimated as 0.0003617: log likelihood = 322.69, aic = -643.38
## Define the Forecasting Time Points
fyears <- yr[133:144]
preds1 <- predict(fit1, 12)</pre>
forecast1 <- preds1$pred</pre>
flimits1 <- qnorm(0.975) * preds1$se
preds2 <- predict(fit2, 12)</pre>
forecast2 <- preds2$pred</pre>
flimits2 <- qnorm(0.975) * preds2$se
par(mfrow = c(2, 2), cex = 0.8, bty = "L", mar = c(3.6, 4, 1, 0.6),
    mgp = c(2.4, 1, 0), las = 1)
plot(shortyears, log.shortair, type = "l", xlab = "Year",
     ylab = "log10(Passenger Numbers)", xlim = range(yr), ylim = c(2, 2.9))
```

```
mtext("SARIMA(1,1,0) x (1,0,0)")
## Plot the Forecasts
lines(fyears, forecast1, lwd = 2, col = "blue")
## Plot the 95% Prediction Intervals
lines(fyears, forecast1 + flimits1, lty = 3, col = "blue")
lines(fyears, forecast1 - flimits1, lty = 3, col = "blue")
plot(shortyears, log.shortair, type = "l", xlab = "Year",
     ylab = "log10(Passenger Numbers)", xlim = range(yr), ylim = c(2, 2.9))
mtext("SARIMA(0,1,0) x (1,0,0)")
## Plot the Forecasts
lines(fyears, forecast2, lwd = 2, col = "red")
## Plot the 95% Prediction Intervals
lines(fyears, forecast2 + flimits2, lty = 3, col = "red")
lines(fyears, forecast2 - flimits2, lty = 3, col = "red")
plot(shortyears, 10^log.shortair, type = "l", xlab = "Year",
     ylab="1000s of Airline Passengers", xlim = range(yr), ylim = c(100, 800))
lines(fyears, 10^forecast1, lwd = 2, col = "blue")
lines(fyears, 10^(forecast1 + flimits1), lty = 3, col = "blue")
lines(fyears, 10^(forecast1 - flimits1), lty = 3, col = "blue")
plot(shortyears, 10^log.shortair, type = "1", xlab = "Year",
     ylab="1000s of Airline Passengers", xlim = range(yr), ylim = c(100, 800))
lines(fyears, 10^forecast2, lwd = 2, col = "red")
lines(fyears, 10^(forecast2 + flimits2), lty = 3, col = "red")
lines(fyears, 10^(forecast2 - flimits2), lty = 3, col = "red")
```



### **Evaluating Forecast Performance**

```
## Calculate the Root Mean Square Error (RMSE)
sqrt(mean((10^forecast1 - 10^log.airpass$V1[133:144])^2))

## [1] 30.36384

sqrt(mean((10^forecast2 - 10^log.airpass$V1[133:144])^2))

## [1] 31.32376

## Calculate the Mean Relative Prediction Error
mean((10^forecast1 - 10^log.airpass$V1[133:144]) / 10^log.airpass$V1[133:144])
```