

Paco (n°2) TP n°4 - Improve the Metropolis-Hastings algorithm

Bayesian analysis of one-way random effects model:

$$\begin{aligned} \textcircled{1} \quad p(X, \mu, \sigma^2, Z^2 | Y) &= p(Y | X, \mu, \sigma^2, Z^2) p(X, \mu, \sigma^2, Z^2) / p(Y) \\ &= C p(Y | X, \mu, \sigma^2, Z^2) p(X | \mu, \sigma^2, Z^2) \prod_{\text{prior}} (\mu, \sigma^2, Z^2) \end{aligned}$$

We have:

$$p(Y | X, \mu, \sigma^2, Z^2) = \prod_{i=1}^N \prod_{j=1}^{k_i} f_{X_i, Z^2}(y_{ij})$$

$$p(X | \mu, \sigma^2, Z^2) = \prod_{i=1}^N f_{\mu, \sigma^2}(x_i)$$

$$\prod_{\text{prior}} (\mu, \sigma^2, Z^2) = K \frac{1}{\sigma^{2(1+\alpha)} e^{-\beta/\sigma^2}} \frac{1}{Z^{2(1+\gamma)} e^{-\beta/Z^2}}, \quad K \text{ a constant.}$$

With f_{μ, σ^2} the density of a $\text{Wi}(\mu, \sigma^2)$ law

Therefore:

$$p(X, \mu, \sigma^2, Z^2 | Y) \propto \prod_{i=1}^N \left[f_{\mu, \sigma^2}(x_i) \prod_{j=1}^{k_i} f_{X_i, Z^2}(y_{ij}) \right] \prod_{\text{prior}} (\mu, \sigma^2, Z^2)$$

$\textcircled{2}$ To perform Gibbs sampling, we need to compute the conditional distributions of each parameter.

$$\begin{aligned} p(X | \sigma^2, Z^2, \mu, Y) &= \frac{p(X, Y, \mu, \sigma^2, Z^2)}{p(\sigma^2, Z^2, \mu, Y)} = \frac{p(X, \mu, \sigma^2, Z^2 | Y) p(Y)}{p(\sigma^2, Z^2, \mu, Y)} \\ &\propto p(X, \mu, \sigma^2, Z^2 | Y) \end{aligned}$$

So:

$$p(x_i | \mu, \sigma^2, z^2) \propto \prod_{j=1}^{k_i} \left[\prod_{j=1}^{k_i} p_{x_i, z^2}(y_{ij}) p_{\mu, \sigma^2}(x_i) \right]$$

$$p(x_i | \mu, \sigma^2, z^2, \{y_{ij}\}_{j \in \{1, k_i\}}) \propto \prod_{j=1}^{k_i} p_{x_i, z^2}(y_{ij}) p_{\mu, \sigma^2}(x_i)$$

$$(A) = \prod_{j=1}^{k_i} \frac{1}{\sqrt{2\pi z^2}} e^{-\frac{1}{2} \frac{(y_{ij}-x_i)^2}{z^2}} \frac{1}{\sqrt{2\pi \sigma^2}} e^{-\frac{1}{2} \frac{(x_i-\mu)^2}{\sigma^2}}$$

$$= \left(\frac{1}{2\pi \sigma^2 (2\pi z^2)^{k_i}} \right)^{1/2} e^{-\frac{1}{2} \left[\frac{(x_i-\mu)^2}{\sigma^2} + \sum_{j=1}^{k_i} \frac{(y_{ij}-x_i)^2}{z^2} \right]}$$

$$\left(\frac{x_i-\mu}{\sigma} \right)^2 + \sum_{j=1}^{k_i} \frac{(y_{ij}-x_i)^2}{z^2} = \frac{1}{\sigma^2} (x_i^2 - 2\mu x_i + \mu^2) + \frac{1}{z^2} \left(\sum_{j=1}^{k_i} y_{ij}^2 - 2x_i y_{ij} + x_i^2 \right)$$

$$= x_i^2 \left(\frac{1}{\sigma^2} + \frac{k_i}{z^2} \right) - 2x_i \left(\mu + \frac{1}{z^2} \sum_{j=1}^{k_i} y_{ij} \right) + \text{cste}$$

$$= \left(\frac{1}{\sigma^2} + \frac{k_i}{z^2} \right) \left[x_i^2 - 2x_i \left(\mu + \frac{1}{z^2} \sum_{j=1}^{k_i} y_{ij} \right) \left(\frac{1}{\sigma^2} + \frac{k_i}{z^2} \right)^{-1} \right] + \text{cste}$$

$$= \left(\frac{1}{\sigma^2} + \frac{k_i}{z^2} \right) \left[x_i - \underbrace{\left(\mu + \frac{1}{z^2} \sum_{j=1}^{k_i} y_{ij} \right) \left(\frac{1}{\sigma^2} + \frac{k_i}{z^2} \right)^{-1}}_{\mu_i} \right]^2 + \text{cste}$$

Therefore:

$$(x_i | \mu, \sigma^2, z^2, \{y_{ij}\}_{j \in \{1, k_i\}}) \sim \mathcal{N}(\mu_i, \sigma_i^2)$$

$$\text{with } \mu_i = \left(\frac{1}{\sigma^2} + \frac{1}{z^2} \sum_{j=1}^{k_i} y_{ij} \right) \sigma_i^{-2}$$

$$\sigma_i^2 = \frac{\sigma^2 z^2}{z^2 + k_i \sigma^2}$$

The x_i are iid and follow a $W(\mu, \sigma^2)$
 Therefore: $\mu | X, \sigma^2, Z^2, Y \sim W(x_n, \frac{\sigma^2}{N})$

With simple computations, we obtain $\sigma^2 | X, \mu, Z^2, Y$ and
 $Z^2 | \mu, X, \sigma^2, Y$, from $p(X, \mu, \sigma^2, Z^2 | Y)$

$$\sigma^2 | X, Y, \mu, Z^2 \sim \Gamma^{-1} \left(\alpha + \frac{N}{2}, \beta + \frac{1}{2} \sum_{i=1}^N (x_i - \mu)^2 \right)$$

$$Z^2 | X, Y, \mu, \sigma^2 \sim \Gamma^{-1} \left(\frac{\alpha + k}{2}, \beta + \frac{1}{2} \sum_{i=1}^N \sum_{j=1}^{k_i} (y_{ij} - x_i)^2 \right)$$

These distributions being known, we can implement a
 Gibbs sampler which updates in turn σ^2, Z^2, μ, X .

③ We have to determine the distribution of $(X, \mu) | \sigma^2, Z^2, Y$.

$$\begin{aligned} p(X, \mu | \sigma^2, Z^2, Y) &= p(X, \mu, \sigma^2, Z^2 | Y) / p(\sigma^2, Z^2, Y) \\ &= p(Y | X, \mu, \sigma^2, Z^2) p(X, \mu, \sigma^2, Z^2) / p(\sigma^2, Z^2, Y) \\ &= p(Y | X, \mu, \sigma^2, Z^2) p(X | \mu, \sigma^2, Z^2) \underbrace{p(\mu, \sigma^2, Z^2)}_{p(\sigma^2, Z^2, Y)} \\ &\quad \times p(Y | X, \mu, \sigma^2, Z^2) p(X | \mu, \sigma^2, Z^2) \end{aligned}$$

case of (X, μ) !

And we have:

$$p(X | \mu, \sigma^2, Z^2) = \prod_{i=1}^N f_{\mu, \sigma^2}(x_i)$$

$$p(Y | X, \mu, \sigma^2, Z^2) = \prod_{i=1}^N \prod_{j=1}^{k_i} f_{x_i, Z^2}(y_{ij})$$

Therefore:

$$p(X, \mu | \sigma^2, \Sigma, \gamma) \propto \prod_{i=1}^N e^{-\frac{1}{2} \frac{(x_i - \mu)^2}{\sigma^2}} \prod_{j=1}^{k_i} e^{-\frac{1}{2} \frac{(y_{ij} - \mu_{ij})^2}{\Sigma_{jj}}}$$

$$\log(p(X, \mu | \sigma^2, \Sigma, \gamma)) = \sum_{i=1}^N \left[-\frac{1}{2} \frac{(x_i - \mu)^2}{\sigma^2} + \sum_{j=1}^{k_i} -\frac{1}{2} \frac{(y_{ij} - \mu_{ij})^2}{\Sigma_{jj}} \right] + \text{const}$$

$$= -\frac{1}{2} \sum_{i=1}^N \left[\frac{1}{\sigma^2} (x_i^2 - 2x_i\mu + \mu^2) + \frac{1}{\Sigma^2} \sum_{j=1}^{k_i} (y_{ij}^2 - 2y_{ij}\mu_{ij} + \mu_{ij}^2) \right] + \text{const}$$

$$= -\frac{1}{2} \sum_{i=1}^N \left[x_i^2 \left(\frac{1}{\sigma^2} + \frac{\mu_{ii}}{\Sigma^2} \right) + x_i \left(-2\mu + 2 \sum_{j=1}^{k_i} \mu_{ij} \right) + \frac{1}{\Sigma^2} \sum_{j=1}^{k_i} y_{ij}^2 \right]$$

$$-\frac{N\mu^2}{2\sigma^2} + \text{const.}$$

$$= -\frac{1}{2} \frac{N\mu^2}{\sigma^2} - \frac{1}{2} \sum_{i=1}^N \left[x_i^2 \left(\frac{1}{\sigma^2} + \frac{\mu_{ii}}{\Sigma^2} \right) - 2x_i \left(\mu + \frac{1}{\Sigma^2} \sum_{j=1}^{k_i} y_{ij} \right) \right] + \text{const}$$

By identification:

$X, \mu | \sigma^2, \Sigma, \gamma \sim \mathcal{W}(\boldsymbol{\mu}_0, \boldsymbol{\Sigma})$, with

$$\boldsymbol{\Sigma} = \begin{pmatrix} \frac{1}{\sigma^2} & & & \\ & 0 & & \\ & & 0 & -\frac{1}{\sigma^2} \\ & & -\frac{1}{\sigma^2} & 0 \end{pmatrix}$$

$$\boldsymbol{\mu}_0 = \frac{1}{\Sigma^2} \begin{pmatrix} y_{1,1} \\ y_{2,1} \\ \vdots \\ y_{N,1} \end{pmatrix}$$

$\mathbb{E}_{\boldsymbol{\mu}}(\boldsymbol{\mu})$