

Bayesian Inference

Chapter 13

Review

Making rational decisions when faced with uncertainty:

- *Probability*
the precise representation of knowledge and uncertainty
- *Probability theory*
how to optimally update your knowledge based on new information
- *Decision theory: probability theory + utility theory*
how to use this information to achieve maximum expected utility

Marginalization

Weather =	sunny	rain	cloudy	snow
Cavity = T	0.144	0.02	0.016	0.02
Cavity = F	0.576	0.08	0.064	0.08

Weather =	sunny	rain	cloudy	snow
	0.72	0.1	0.08	0.1



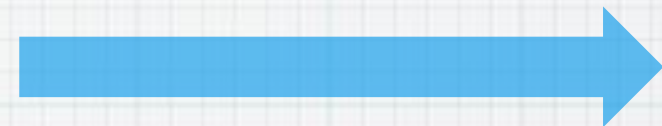
$$P(W) = \sum_{i=1}^2 P(C = c_i, W)$$

$$\begin{aligned} P(W = \text{sunny}) \\ = P(\text{cavity}=\text{t}, W=\text{sunny}) + P(\text{cavity}=\text{f}, W=\text{sunny}) \end{aligned}$$

Marginalization

Weather =	sunny	rain	cloudy	snow
Cavity = T	0.144	0.02	0.016	0.02
Cavity = F	0.576	0.08	0.064	0.08

Cavity = T	0.2
Cavity = F	0.8



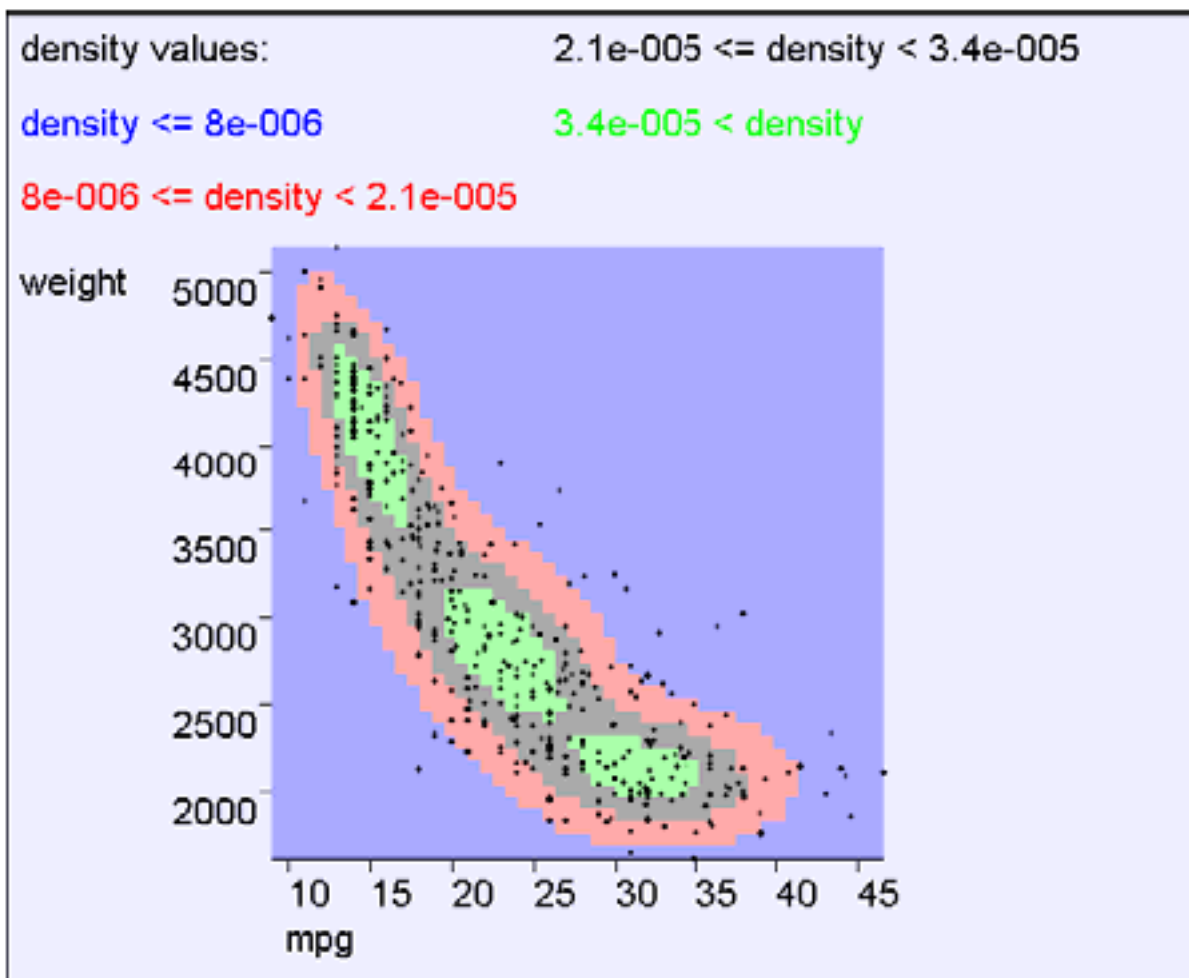
$$P(C) = \sum_{i=1}^4 P(C, W = w_i)$$

Continuous Densities

In 2
dimensions

Let X, Y be a pair of continuous random variables, and let R be some region of (X, Y) space...

$$P((X, Y) \in R) = \iint_{(x, y) \in R} p(x, y) dy dx$$



Probability of Joint Event

In 2
dimensions

Let X, Y be a pair of continuous random variables, and let R be some region of (X, Y) space...

$$P((X, Y) \in R) = \iint_{(x, y) \in R} p(x, y) dy dx$$

density values:

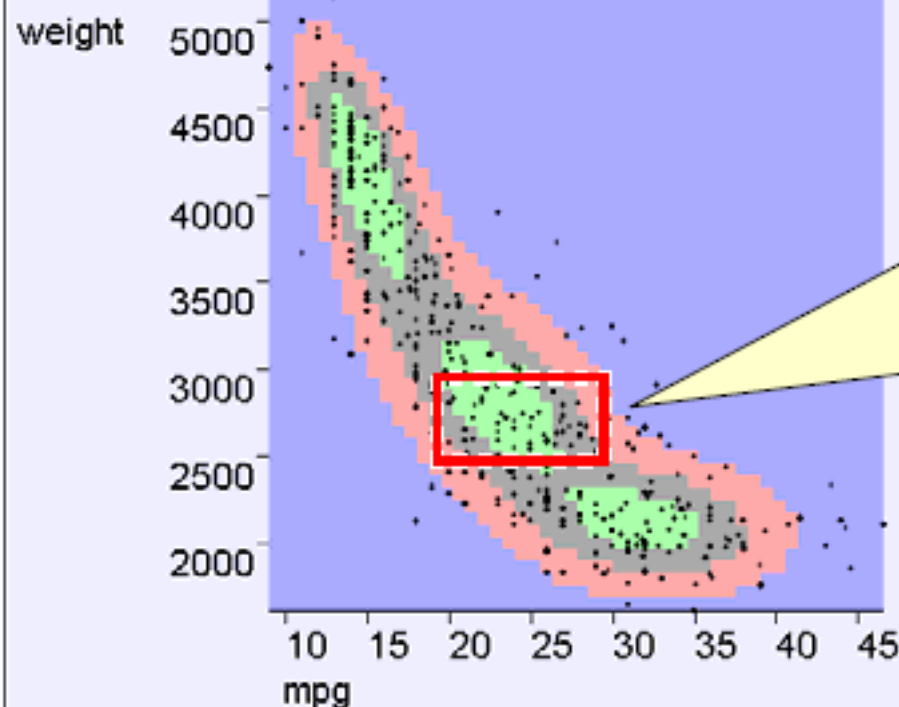
$2.1\text{e-}005 \leq \text{density} < 3.4\text{e-}005$

$(x, y) \in R$

$\text{density} \leq 8\text{e-}006$

$3.4\text{e-}005 < \text{density}$

$8\text{e-}006 \leq \text{density} < 2.1\text{e-}005$

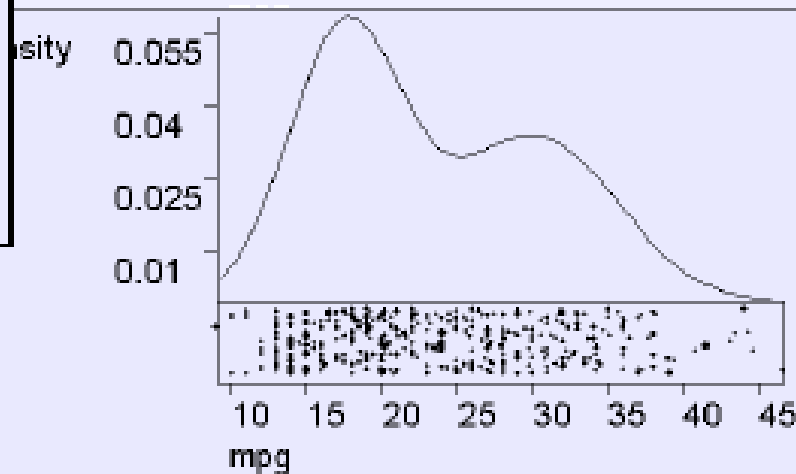
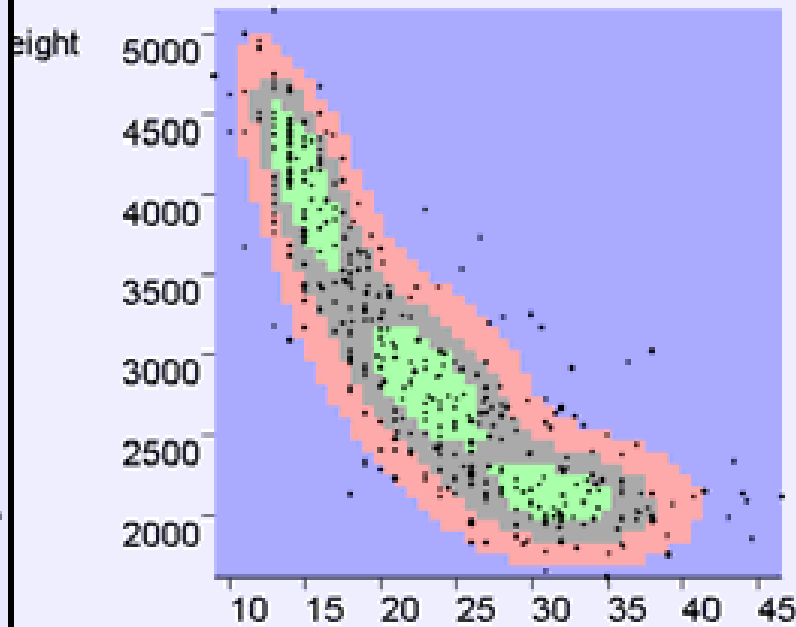
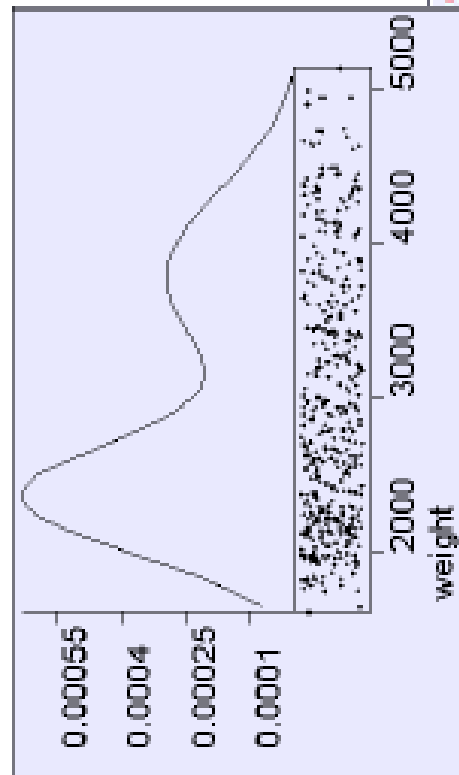


$P(20 < \text{mpg} < 30 \text{ and } 2500 < \text{weight} < 3000) =$

area under the 2-d surface within the red rectangle

Marginal Distributions

density values: $2.1\text{e-}005 \leq \text{density} < 3.4\text{e-}005$
 $\text{density} \leq 8\text{e-}006$ $3.4\text{e-}005 < \text{density}$
 $8\text{e-}006 \leq \text{density} < 2.1\text{e-}005$



$$p(x) = \int_{y=-\infty}^{\infty} p(x, y) dy$$

Inference by Enumeration

Start with the joint distribution:

	<i>toothache</i>		\neg <i>toothache</i>	
	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
<i>cavity</i>	.108	.012	.072	.008
\neg <i>cavity</i>	.016	.064	.144	.576

For any proposition ϕ , sum the atomic events where it is true:

$$P(\phi) = \sum_{\omega: \omega \models \phi} P(\omega)$$

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$$P(\text{toothache}) = 0.108 + 0.012 + 0.016 + 0.064 = 0.2$$

Inference by Enumeration

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	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
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For any proposition ϕ , sum the atomic events where it is true:

$$P(\phi) = \sum_{\omega: \omega \models \phi} P(\omega)$$

$$P(\text{cavity} \vee \text{toothache}) = 0.108 + 0.012 + 0.072 + 0.008 + 0.016 + 0.064 = 0.28$$

Inference by Enumeration

Start with the joint distribution:

	<i>toothache</i>		\neg <i>toothache</i>	
	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
<i>cavity</i>	.108	.012	.072	.008
\neg <i>cavity</i>	.016	.064	.144	.576

Can also compute conditional probabilities:

$$\begin{aligned} P(\neg \text{cavity} | \text{toothache}) &= \frac{P(\neg \text{cavity} \wedge \text{toothache})}{P(\text{toothache})} \\ &= \frac{0.016 + 0.064}{0.108 + 0.012 + 0.016 + 0.064} = 0.4 \end{aligned}$$

Normalization

	<i>toothache</i>		\neg <i>toothache</i>	
	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
<i>cavity</i>	.108	.012	.072	.008
\neg <i>cavity</i>	.016	.064	.144	.576

Denominator can be viewed as a normalization constant α

$$\begin{aligned}\mathbf{P}(Cavity|toothache) &= \alpha \mathbf{P}(Cavity, toothache) \\ &= \alpha [\mathbf{P}(Cavity, toothache, catch) + \mathbf{P}(Cavity, toothache, \neg catch)] \\ &= \alpha [\langle 0.108, 0.016 \rangle + \langle 0.012, 0.064 \rangle] \\ &= \alpha \langle 0.12, 0.08 \rangle = \langle 0.6, 0.4 \rangle\end{aligned}$$

General idea: compute distribution on query variable
by fixing **evidence variables** and summing over **hidden variables**

Independence

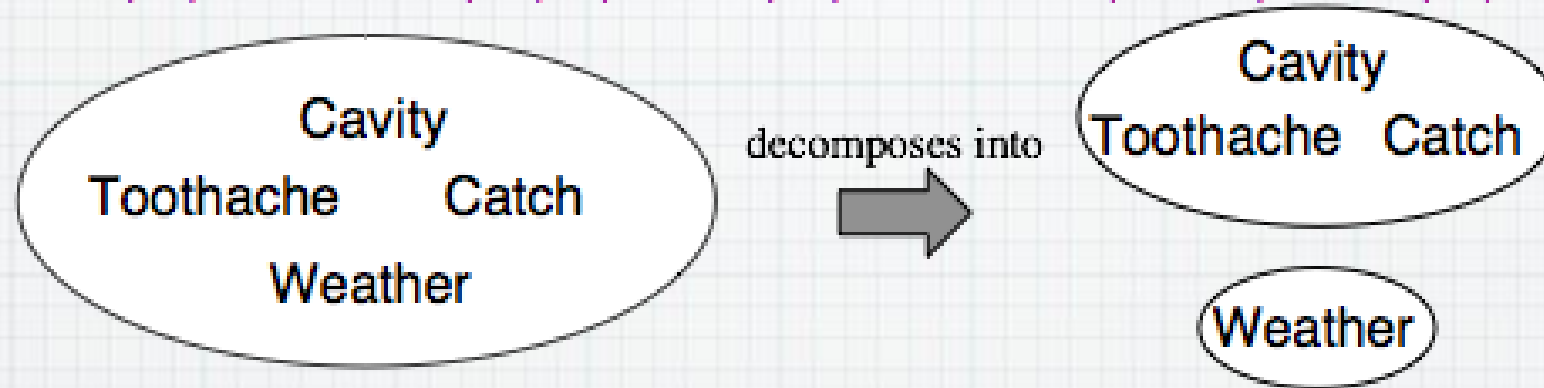
A and B are independent iff

$$\mathbf{P}(A|B) = \mathbf{P}(A) \quad \text{or} \quad \mathbf{P}(B|A) = \mathbf{P}(B) \quad \text{or} \quad \mathbf{P}(A, B) = \mathbf{P}(A)\mathbf{P}(B)$$

Independence

A and B are independent iff

$$P(A|B) = P(A) \quad \text{or} \quad P(B|A) = P(B) \quad \text{or} \quad P(A, B) = P(A)P(B)$$

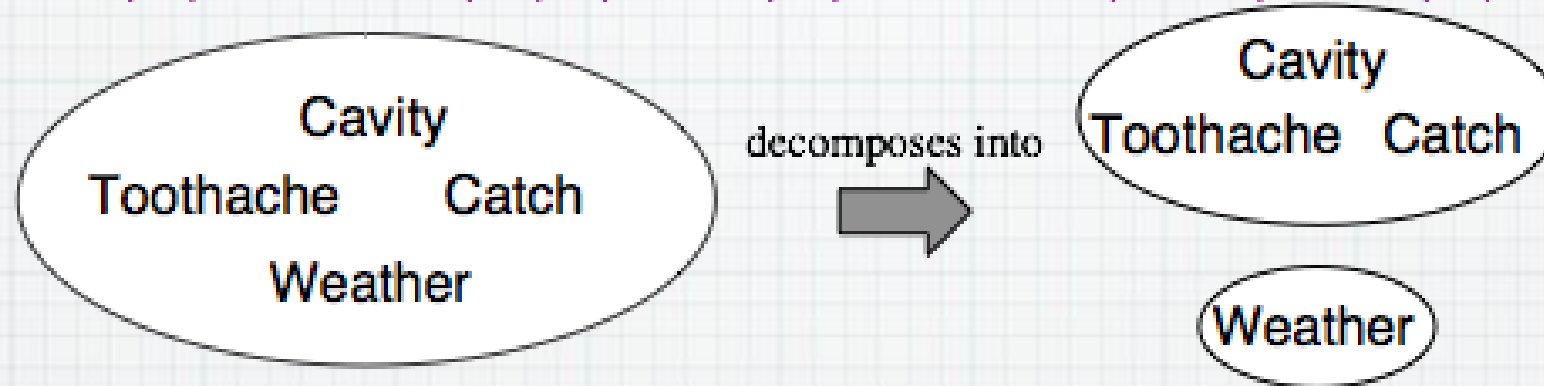


$$\begin{aligned} P(\textit{Toothache}, \textit{Catch}, \textit{Cavity}, \textit{Weather}) \\ = P(\textit{Toothache}, \textit{Catch}, \textit{Cavity})P(\textit{Weather}) \end{aligned}$$

Independence

A and B are independent iff

$$P(A|B) = P(A) \quad \text{or} \quad P(B|A) = P(B) \quad \text{or} \quad P(A, B) = P(A)P(B)$$



$$\begin{aligned} P(\textit{Toothache}, \textit{Catch}, \textit{Cavity}, \textit{Weather}) \\ = P(\textit{Toothache}, \textit{Catch}, \textit{Cavity})P(\textit{Weather}) \end{aligned}$$

32 entries reduced to 12; for n independent biased coins, $2^n \rightarrow n$

Absolute independence powerful but rare

Dentistry is a large field with hundreds of variables, none of which are independent. What to do?

Inference Using Bayes Rule

Simple example: medical test results

- Test report for rare disease is positive, 90% accurate
- What's the probability that you have the disease?
- What if the test is repeated?
- This is the simplest example of reasoning by combining sources of information.

How do we model the problem?

- Which is the correct description of “Test is 90% accurate” ?

$$P(T = \text{true}) = 0.9$$

$$P(T = \text{true} | D = \text{true}) = 0.9$$

$$P(D = \text{true} | T = \text{true}) = 0.9$$

- What do we want to know?

$$P(T = \text{true})$$

$$P(T = \text{true} | D = \text{true})$$

$$P(D = \text{true} | T = \text{true})$$

- More compact notation:

$$P(T = \text{true} | D = \text{true}) \rightarrow P(T | D)$$

$$P(T = \text{false} | D = \text{false}) \rightarrow P(\bar{T} | \bar{D})$$

Evaluating the posterior probability through Bayesian inference

- We want $P(D|T)$ = “The probability of the having the disease given a positive test”
- Use Bayes rule to relate it to what we know: $P(T|D)$

$$\text{posterior } P(D|T) = \frac{\text{likelihood } P(T|D) \text{ prior } P(D)}{\text{normalizing constant } P(T)}$$

- What's the prior $P(D)$?
- Disease is rare, so let's assume

$$P(D) = 0.001$$

- What about $P(T)$? $P(T) = P(T, D) + P(T, \bar{D})$
- What's the interpretation of that?

Evaluating the normalizing constant

$$\text{posterior } P(D|T) = \frac{\text{likelihood } P(T|D) \text{ prior } P(D)}{\text{normalizing constant } P(T)}$$

- $P(T)$ is the marginal probability of $P(T,D) = P(T|D) P(D)$
- So, compute with summation

$$P(T) = \sum_{\text{all values of } D} P(T|D)P(D)$$

- For true or false propositions:

$$P(T) = P(T|D)P(D) + P(T|\bar{D})P(\bar{D})$$

- Our complete expression is

$$P(D|T) = \frac{P(T|D)P(D)}{P(T|D)P(D) + P(T|\bar{D})P(\bar{D})}$$

- Plugging in the numbers we get:

$$P(D|T) = \frac{0.9 \times 0.001}{0.9 \times 0.001 + 0.1 \times 0.999} = 0.0089$$

$$P(D) = 0.001$$

- Does this make intuitive sense?

After positive test, 9 times more likely to have disease

Playing around with the numbers

$$P(D|T) = \frac{P(T|D)P(D)}{P(T|D)P(D) + P(T|\bar{D})P(\bar{D})}$$

- What if the test were 100% reliable?

$$P(D|T) = \frac{1.0 \times 0.001}{1.0 \times 0.001 + 0.0 \times 0.999} = 1.0$$

- What if the test was the same, but disease wasn't so rare?

$$P(D|T) = \frac{0.9 \times 0.1}{0.9 \times 0.1 + 0.1 \times 0.999} = 0.5$$

Repeating the test

- We can relax, $P(D|T) = 0.0089$, right?
- Just to be sure the doctor recommends repeating the test.
- How do we represent this?

$$P(D|T_1, T_2)$$

- Again, we apply Bayes' rule

$$P(D|T_1, T_2) = \frac{P(T_1, T_2|D)P(D)}{P(T_1, T_2)}$$

- How do we model $P(T_1, T_2|D)$?

Modeling repeated tests

$$P(D|T_1, T_2) = \frac{P(T_1, T_2|D)P(D)}{P(T_1, T_2)}$$

- Easiest is to assume the tests are *independent*.

$$P(T_1, T_2|D) = P(T_1|D)P(T_2|D)$$

- This also implies:

$$P(T_1, T_2) = P(T_1)P(T_2)$$

- Plugging these in, we have

$$P(D|T_1, T_2) = \frac{P(T_1|D)P(T_2|D)P(D)}{P(T_1)P(T_2)}$$

Evaluating the normalizing constant again

- Expanding as before we have

$$P(D|T_1, T_2) = \frac{P(T_1|D)P(T_2|D)P(D)}{\sum_{D=\{t,f\}} P(T_1|D)P(T_2|D)P(D)}$$

- Plugging in the numbers gives us

$$P(D|T) = \frac{0.9 \times 0.9 \times 0.001}{0.9 \times 0.9 \times 0.001 + 0.1 \times 0.1 \times 0.999} = 0.075$$

- Another way to think about this:
 - What's the chance of 1 false positive from the test?
 - What's the chance of 2 false positives?
- The chance of 2 false positives is still 10x more likely than the a prior probability of having the disease.

Simpler: Combining information the Bayesian way

- Let's look at the equation again:

$$P(D|T_1, T_2) = \frac{P(T_1|D)P(T_2|D)P(D)}{P(T_1)P(T_2)}$$

- If we rearrange slightly:

$$P(D|T_1, T_2) = \frac{P(T_2|D) \boxed{P(T_1|D)P(D)}}{P(T_2)P(T_1)}$$

We've seen
this before!

- It's the posterior for the first test, which we just computed

$$P(D|T_1) = \frac{P(T_1|D)P(D)}{P(T_1)}$$

The old posterior is the new prior

- We can just plugin the value of the old posterior
- It plays exactly the same role as our old prior

$$P(D|T_1, T_2) = \frac{P(T_2|D)P(T_1|D)P(D)}{P(T_2)P(T_1)}$$

$$P(D|T_1, T_2) = \frac{P(T_2|D) \times 0.0089}{P(T_2)}$$

- Plugging in the numbers gives the same answer:

$$P(D|T) = \frac{P(T|D)P'(D)}{P(T|D)P'(D) + P(T|\bar{D})P'(\bar{D})}$$

$$P(D|T) = \frac{0.9 \times 0.0089}{0.9 \times 0.0089 + 0.1 \times 0.9911} = 0.075$$

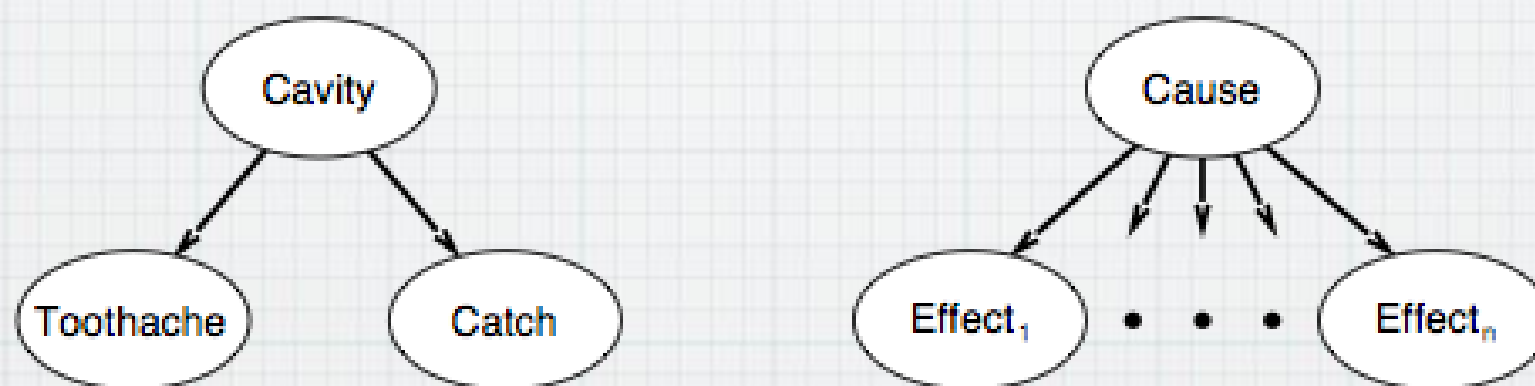
This is how Bayesian reasoning combines old information with new information to update our belief states.

Bayes' Rule and Conditional Independence

$$\begin{aligned} & \mathbf{P}(Cavity|toothache \wedge catch) \\ &= \alpha \mathbf{P}(toothache \wedge catch|Cavity)\mathbf{P}(Cavity) \\ &= \alpha \mathbf{P}(toothache|Cavity)\mathbf{P}(catch|Cavity)\mathbf{P}(Cavity) \end{aligned}$$

This is an example of a **naive Bayes** model:

$$\mathbf{P}(Cause, Effect_1, \dots, Effect_n) = \mathbf{P}(Cause) \prod_i \mathbf{P}(Effect_i|Cause)$$



Total number of parameters is **linear** in n

Wumpus Example

1,4	2,4	3,4	4,4
1,3	2,3	3,3	4,3
1,2 B OK	2,2	3,2	4,2
1,1 OK	2,1 B OK	3,1	4,1

$P_{ij} = \text{true}$ iff $[i, j]$ contains a pit

$B_{ij} = \text{true}$ iff $[i, j]$ is breezy

Include only $B_{1,1}$, $B_{1,2}$, $B_{2,1}$ in the probability model

Wumpus Probability Model

The full joint distribution is $\mathbf{P}(P_{1,1}, \dots, P_{4,4}, B_{1,1}, B_{1,2}, B_{2,1})$

Apply product rule: $\mathbf{P}(B_{1,1}, B_{1,2}, B_{2,1} \mid P_{1,1}, \dots, P_{4,4}) \mathbf{P}(P_{1,1}, \dots, P_{4,4})$

(Do it this way to get $P(\textit{Effect} \mid \textit{Cause})$.)

First term: 1 if pits are adjacent to breezes, 0 otherwise

Second term: pits are placed randomly, probability 0.2 per square:

$$\mathbf{P}(P_{1,1}, \dots, P_{4,4}) = \prod_{i,j=1,1}^{4,4} \mathbf{P}(P_{i,j}) = 0.2^n \times 0.8^{16-n}$$

for n pits.

Observations and Query

We know the following facts:

$$b = \neg b_{1,1} \wedge b_{1,2} \wedge b_{2,1}$$

$$known = \neg p_{1,1} \wedge \neg p_{1,2} \wedge \neg p_{2,1}$$

Query is $\mathbf{P}(P_{1,3}|known, b)$

Define $Unknown = P_{ij}$ s other than $P_{1,3}$ and $Known$

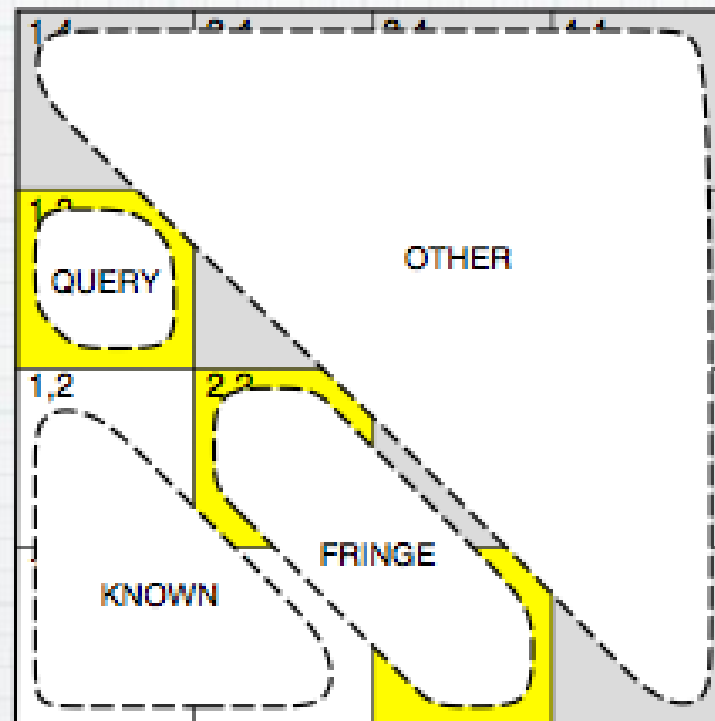
For inference by enumeration, we have

$$\mathbf{P}(P_{1,3}|known, b) = \alpha \sum_{unknown} \mathbf{P}(P_{1,3}, unknown, known, b)$$

Grows exponentially with number of squares!

Using Conditional Independence

Basic insight: observations are conditionally independent of other hidden squares given neighbouring hidden squares



Define $Unknown = Fringe \cup Other$

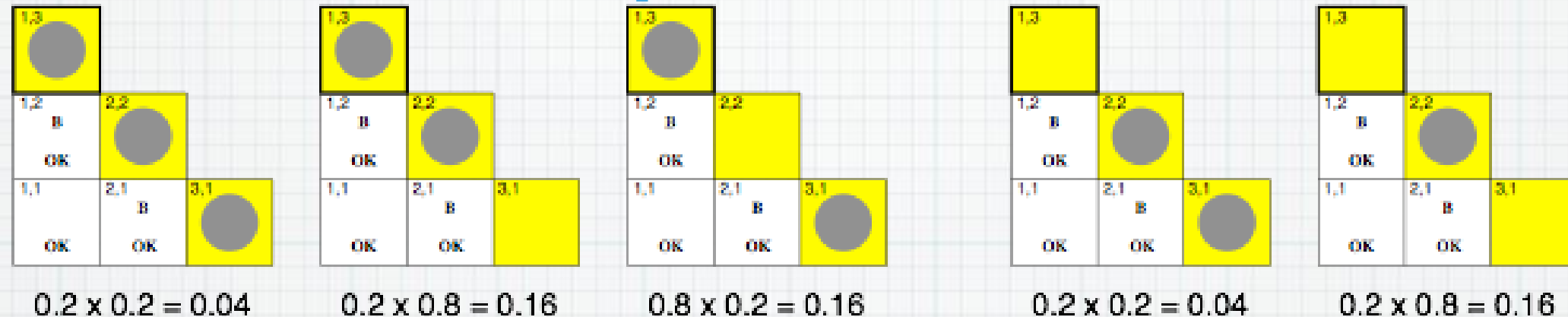
$$\mathbf{P}(b|P_{1,3}, \textit{Known}, \textit{Unknown}) = \mathbf{P}(b|P_{1,3}, \textit{Known}, \textit{Fringe})$$

Manipulate query into a form where we can use this!

Using Conditional Independence

$$\begin{aligned}\mathbf{P}(P_{1,3}|\textit{known}, b) &= \alpha \sum_{\textit{unknown}} \mathbf{P}(P_{1,3}, \textit{unknown}, \textit{known}, b) \\&= \alpha \sum_{\textit{unknown}} \mathbf{P}(b|P_{1,3}, \textit{known}, \textit{unknown}) \mathbf{P}(P_{1,3}, \textit{known}, \textit{unknown}) \\&= \alpha \sum_{\textit{fringe}} \sum_{\textit{other}} \mathbf{P}(b|\textit{known}, P_{1,3}, \textit{fringe}, \textit{other}) \mathbf{P}(P_{1,3}, \textit{known}, \textit{fringe}, \textit{other}) \\&= \alpha \sum_{\textit{fringe}} \sum_{\textit{other}} \mathbf{P}(b|\textit{known}, P_{1,3}, \textit{fringe}) \mathbf{P}(P_{1,3}, \textit{known}, \textit{fringe}, \textit{other}) \\&= \alpha \sum_{\textit{fringe}} \mathbf{P}(b|\textit{known}, P_{1,3}, \textit{fringe}) \sum_{\textit{other}} \mathbf{P}(P_{1,3}, \textit{known}, \textit{fringe}, \textit{other}) \\&= \alpha \sum_{\textit{fringe}} \mathbf{P}(b|\textit{known}, P_{1,3}, \textit{fringe}) \sum_{\textit{other}} \mathbf{P}(P_{1,3}) P(\textit{known}) P(\textit{fringe}) P(\textit{other}) \\&= \alpha P(\textit{known}) \mathbf{P}(P_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(b|\textit{known}, P_{1,3}, \textit{fringe}) P(\textit{fringe}) \sum_{\textit{other}} P(\textit{other}) \\&= \alpha' \mathbf{P}(P_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(b|\textit{known}, P_{1,3}, \textit{fringe}) P(\textit{fringe})\end{aligned}$$

Using Conditional Independence



$$\mathbf{P}(P_{1,3} | \text{known}, b) = \alpha' \langle 0.2(0.04 + 0.16 + 0.16), 0.8(0.04 + 0.16) \rangle \\ \approx \langle 0.31, 0.69 \rangle$$

$$\mathbf{P}(P_{2,2} | \text{known}, b) \approx \langle 0.86, 0.14 \rangle$$

Summary

Probability is a rigorous formalism for uncertain knowledge

Joint probability distribution specifies probability of every atomic event

Queries can be answered by summing over atomic events

For nontrivial domains, we must find a way to reduce the joint size

Independence and conditional independence provide the tools