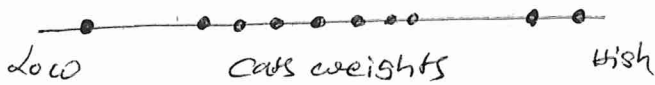


Motivation for Maximum Likelihood Estimation

Let's say we weighted a bunch of cats

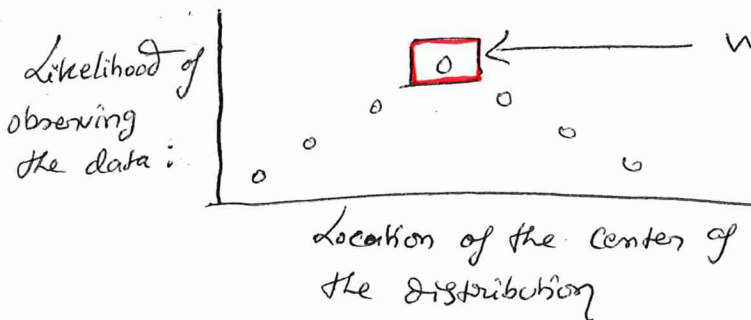
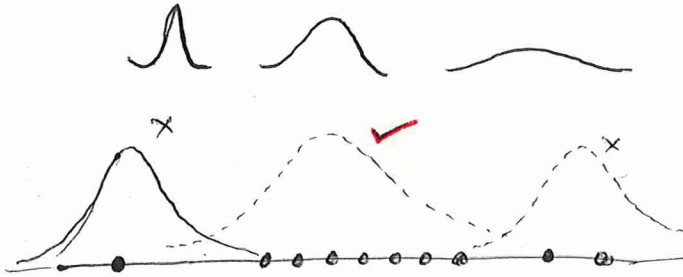


The reason you want to fit a distribution to your data is it can be easier to work with and it's also more general - it applies to every experiment of same type.

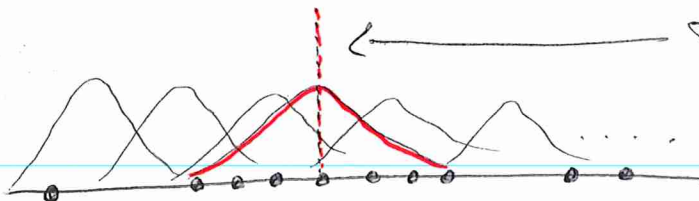
There are lots of different types of distributions for different types of data

Normal Exp Gamma

↳ In this case, you can think that the weights ~~are~~ might be Normally distributed. Now, "Normally distributed" means we might expect most of the measurements to be close to mean, and the measurements to be relatively symmetrical around the mean.



We want the location that "maximizes the likelihood" of observing the weights we measured.



This location for the mean "maximizes the likelihood" of observing the weights we measured.

In statistics, "likelihood" refers to the situation, where we are trying to find the optimal value for the parameters of a distⁿ given a bunch of measured observations

Thus, it is the "maximum likelihood estimate for the mean". ★

Similar thing happens for the "MLE for the standard deviation".

Introduction to Maximum Likelihood Estimation.

↳ Maximum Likelihood Principle:

Given a dataset (assumed to be from a distribution), choose the parameter of interest θ , in a way that the data are most likely.

Consider a pdf/pmf : $f(x; \theta) = C \cdot P_x(x; \theta)$

\uparrow Random Variable \uparrow Coefficient $\perp \theta$ \uparrow parameter (assumed known)
 \uparrow Everything else

likelihood function : $L(\theta) = P_x(x; \theta) = \frac{1}{C} f(x; \theta)$

$$0 \leq L(\theta; x) < \infty$$

For Example 1- Let, $X \sim \text{Bin}(n, p)$

Then, pmf, $f(x) = \binom{n}{x} p^x (1-p)^{n-x}$; $x \in \{0, 1, \dots, n\}$
 $0 \leq p \leq 1$.

(likelihood function)

$$L(p) = \binom{n}{x} p^x (1-p)^{n-x}; 0 \leq p \leq 1$$

$$\Rightarrow L(p) = p^x (1-p)^{n-x}; 0 \leq p \leq 1$$

Maximum Likelihood Estimate (MLE):

↳ The MLE of a parameter θ is the value $\hat{\theta}$ that maximizes the likelihood, $L(\theta)$, given that x .

$$\text{i.e., } \hat{\theta} = \arg \max_{\theta} L(\theta) \Rightarrow \left. \frac{d}{d\theta} L(\theta) \right|_{\hat{\theta}} = 0$$

$$\Rightarrow L(\hat{\theta}) \geq L(\theta)$$

Log-likelihood: $\ell(\theta) = \log [L(\theta)]$ or $\ln [L(\theta)]$

It is valid because, $\log(x)$ is monotonically increasing
★ Exercise =

↳ Back to $X \sim \text{Bin}(n, p)$

$$L(p) = p^x (1-p)^{n-x} ; \text{ find the MLE } \hat{p}.$$

Taking log on both sides,

$$\log [L(p)] = \log [p^x (1-p)^{n-x}]$$

$$\text{or, } \ell(p) = \log(p^x) + \log((1-p)^{n-x}) \quad \star$$
$$= x \log(p) + (n-x) \log(1-p) ; 0 < p < 1$$

Now, differentiating both sides,

$$\ell'(p) = \frac{x}{p} - \frac{n-x}{1-p} ; 0 < p < 1$$

$$\ell'(p) \Big|_{\hat{p}} = \left[\frac{x}{p} - \frac{n-x}{1-p} \right] \Big|_{\hat{p}} = 0$$

$$\Rightarrow \frac{x}{\hat{p}} - \frac{n-x}{1-\hat{p}} = 0 \Rightarrow \boxed{\hat{p} = x/n} \quad \text{Sample proportion}$$

$$\left\{ \begin{array}{l} \text{2nd derivative test: We want to show } \ell''(\hat{p}) < 0 \\ \ell''(p) = -\frac{x}{p^2} - \frac{n-x}{(1-p)^2} < 0 \quad \forall p \in (0, 1). \\ \text{For boundaries; } L(0) = 0, L(1) = 0 ; L(\hat{p}) \geq 0 \end{array} \right.$$

Some more defs:

The score function : $S(\theta) = \frac{d}{d\theta} l(\theta)$ or $l'(\theta)$

Information function : $I(\theta) = -\frac{d^2}{d\theta^2} l(\theta) = -\frac{d}{d\theta} S(\theta)$
or, $S'(\theta)$.

Comments: On MLE $\hat{\theta}$, $S(\hat{\theta}) = 0 \leftarrow$ defn of MLE
 $I(\hat{\theta}) > 0 \leftarrow$ 2nd derivative test.

Example:- Let X be a random variable with pmf.

$$f(x; \mu) = \frac{e^{-\mu x} (\mu x)^{x-1}}{x!} ; \begin{matrix} x \in \{1, 2, \dots\} \\ \text{or} \\ x \in \mathbb{N}^+ ; \mu \in [0, 1] \end{matrix}$$

Find MLE $\hat{\mu}$.

Step 1:- The likelihood function, $[L(\theta) = \frac{1}{c} f(\dots)]$
 $\mu \in [0, 1]$
 $L(\mu) = e^{-\mu x} \mu^{x-1}$

Now, the log-likelihood function,

$$l(\mu) = \ln[e^{-\mu x} \mu^{x-1}]$$

$$\Rightarrow l(\mu) = -\mu x + (x-1) \ln(\mu) \quad , \quad 0 < \mu < 1$$

$\leftarrow (i)$

Step-2:- Now, differentiate both sides of (i) ,

$$l'(\mu) = -x + \frac{x-1}{\mu}$$

Step-3:- $l'(\hat{\mu}) = 0$

$$\Rightarrow -x + \frac{x-1}{\mu} = 0 \Rightarrow \boxed{\hat{\mu} = \frac{x-1}{x}}$$

Step-4:- 2nd derivative test !!