

Section 5.1: Inverting tests to Derive Confidence Intervals.

Defn: Let, θ_0 be the true parameter value of θ .

$[A, B]$ is $100(1-\alpha)\%$ Confidence Interval (CI)

$$\text{if, } P[A \leq \theta_0 \leq B | \theta = \theta_0] = 1 - \alpha$$

"Converge probability"

i.e. $1-\alpha$ is the proportion of C.I's, estimated from repeated random samples, that include the true parameter value θ_0 .

You can invert a Hypothesis test (like, a LRT, for example) to obtain/derive a Confidence interval.

Now, let's look at the Example 5.1.1 from the complete lecture notes. \rightarrow

$$\hookrightarrow n = 50,$$

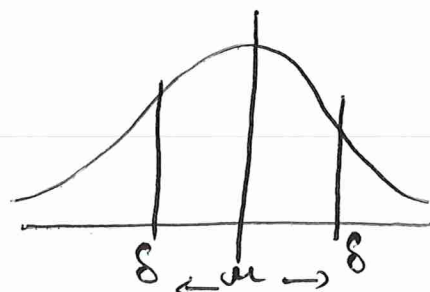
H_0 : No effect on profits.

X_i : Business's change in profits, $i = 1, 2, \dots, 50$

$$\bar{x} = -1000$$

$$X_i \stackrel{iid}{\sim} N(\mu, \sigma^2 = 600^2)$$

$$H_0: \mu = 0$$



Basic Model :-

$$L(\mu) = \prod_{i=1}^n \text{Exp}\left[-\frac{1}{2\sigma^2}(x_i - \mu)^2\right]$$

★ See Ex 4.2.1 for derivation

$$l(\mu) = -\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

Since, σ is known, μ is only unknown
 $\therefore \boxed{K=1}$; $\hat{\mu} = \bar{x}$

Hypothesized Model :-

$$\sigma = 600; \mu_0 = 0$$

So, $\boxed{q=0}$

Performing LRT :-

$$D = 2[\ell(\hat{\mu}) - \ell(\mu_0)]$$

$$\Rightarrow D = 2 \left[-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \bar{x})^2 - \left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu_0)^2 \right) \right]$$

$$\Rightarrow D = \frac{n(\bar{x} - \mu_0)^2}{\sigma^2}$$

$$d_{obs} = \frac{50(-1000 - 0)^2}{600^2} = \cancel{85} 138.889$$

$$p\text{-value} = P[D \geq d_{obs}] = P[\chi^2_{(1)} \geq \cancel{85}^{138.889}] < \cancel{0.0004}^{0.01}$$

We have very strong evidence against H_0 , with $p\text{-value} < 0.01$
 The data are not consistent with the hypothesis that no effects on profits.

Question : What range of μ_0 , would give us a $p\text{-value} \geq 0.05$?
 \hookrightarrow these would be "reasonably consistent" with the data.

$$p\text{-value} = P(\chi^2_{(1)} \geq d_{obs}) = 0.05$$

$$d_{obs} = 3.843, \text{ from tables.}$$

$$p\text{-value} \geq 0.05 \iff d_{obs} \leq 3.843$$

$$\frac{n(\bar{x} - \mu_0)^2}{s^2} \leq 3.843$$

$$\Rightarrow (\bar{x} - \mu_0)^2 \leq 3.843 \frac{s^2}{n}$$

$$\Rightarrow 1.96 \frac{s}{\sqrt{n}} \leq \bar{x} - \mu_0 \leq -1.96 \frac{s}{\sqrt{n}}$$

$$\Rightarrow \bar{x} - 1.96 \frac{s}{\sqrt{n}} \leq \mu_0 \leq \bar{x} + 1.96 \frac{s}{\sqrt{n}}$$

or, $\boxed{\bar{x} \pm 1.96 \frac{s}{\sqrt{n}}}$ is our 95% C.I for μ_0 .

$$95\% = 100(1-\alpha)\%$$

0.05

Putting the values, we get,

$\{ [-1166, -834] \}$ Any μ value in this interval is what we'd consider a "reasonable" estimate of μ_0 , given the data.

Since, 0 is not in the 95% C.I, it is not a plausible value for μ_0 , given the data. Hence, evidence against the null hypothesis.