

Section 6.2 - One Sample Models (6.2.1, 6.2.2)

Hypothesis tests and C.I's for μ

Set-up: Suppose $Y_1, Y_2, \dots, Y_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ random sample.

Recall the joint log-likelihood:

$$l(\mu, \sigma^2) = -\frac{n}{2} \ln(\sigma^2) - \frac{\sum_{i=1}^n (y_i - \mu)^2}{2\sigma^2}$$

We want to test, $H_0: \mu = \mu_0 \leftarrow \text{some value.}$

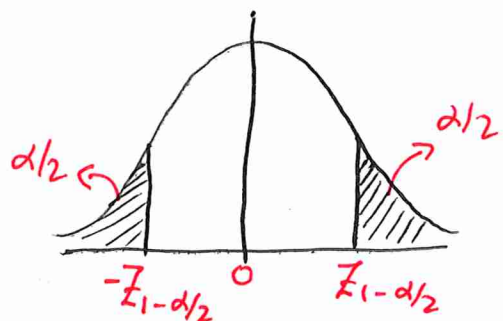
When σ^2 is known

$$\text{LRS ; } D = 2 \left[\underbrace{l(\hat{\mu}, \sigma^2)}_{k=1} - \underbrace{l(\mu_0, \sigma^2)}_{q=0} \right]$$

..... Algebra

$$= (\hat{\mu} - \mu_0)^2 / \sigma^2 / n = Z^2 \sim \chi^2_{(1)} \left[Z \sim N(0,1) \right]$$

$$\begin{aligned} p\text{-value} &= P(D \geq d_{\text{obs}}) \\ &= P(Z^2 \geq Z_{\text{obs}}^2) \\ &= P(|Z| \geq |Z_{\text{obs}}|) \leftarrow \text{2-tailed} \\ &= 2 \cdot P(Z \geq |Z_{\text{obs}}|), \text{ because } N(0,1) \text{ is symmetric.} \end{aligned}$$



100(1- α)% C.I for μ :

$$\hat{\mu} \pm Z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}$$

Labels: $\hat{\mu} = \bar{y}$ (Standard Error), $Z_{1-\alpha/2}$ (Critical value), $\frac{\sigma}{\sqrt{n}}$ (Margin of Error).

When S^2 unknown $H_0: \mu = \mu_0$

$$D = 2 \left[\ell(\hat{\mu}, \hat{S}^2) - \ell(\mu_0, \tilde{S}^2) \right] \approx \chi^2_{(1)}$$

$k=2$ Estimated $q=1$

D is a 1-1 increasing function of T^2 ,

where, $T = \frac{\hat{\mu} - \mu_0}{S/\sqrt{n}} \sim t_{(n-1)}$ Student's t Distribution

$$S^2 = \sum_{i=1}^n (y_i - \hat{\mu})^2 / n-1 \rightarrow \text{Sample Variance.}$$

$$D = n \ln \left[1 + \frac{1}{n-1} T^2 \right] \quad \star \text{Exercise: Derive this!}$$

$$\begin{aligned} P\text{-value} &= P(D \geq D_{\text{obs}}) \\ &= P(T^2 \geq t_{\text{obs}}^2) \\ &= \dots = 2 P(T \geq |t_{\text{obs}}|), \text{ by symmetry of } t\text{-distribution.} \end{aligned}$$

100(1- α)% C.I for μ :

$$P(|T| \geq t_{1-\alpha/2, n-1}) \geq \alpha \iff t_{\text{obs}}(\mu_0) \leq t_{1-\alpha/2, (n-1)}$$

$$\Rightarrow \boxed{\hat{\mu} \pm t_{1-\alpha/2, (n-1)} \cdot S/\sqrt{n}} \quad (\hat{\mu} = \bar{y})$$

Solve for set of μ_0 that satisfy this.

N.B., Let look at the Example 6.2.1 from the Complete Lecture Notes \rightarrow

Plotting your data before analysis important

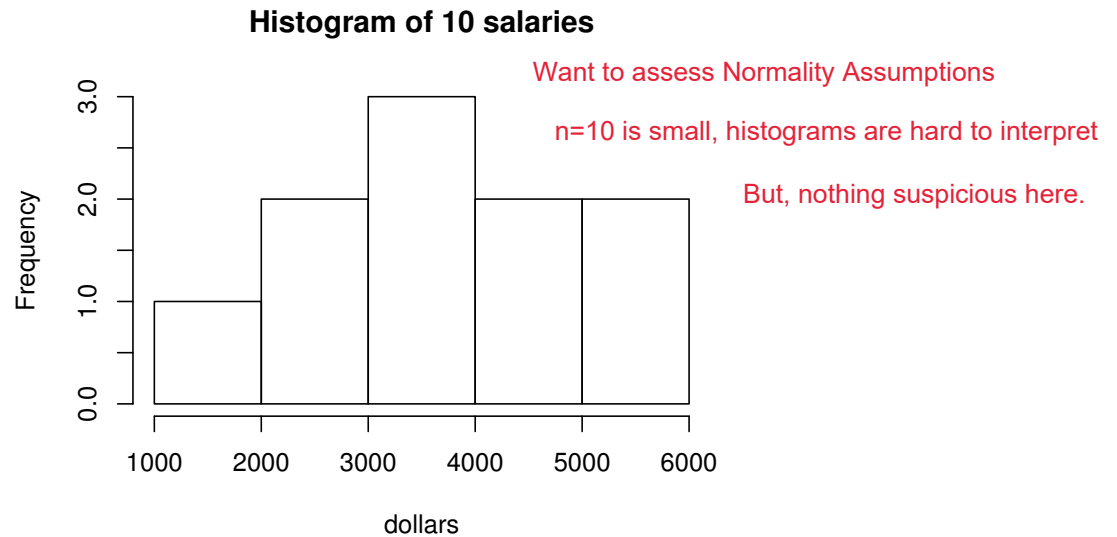


Figure 6.1: Histogram of 10 Salaries

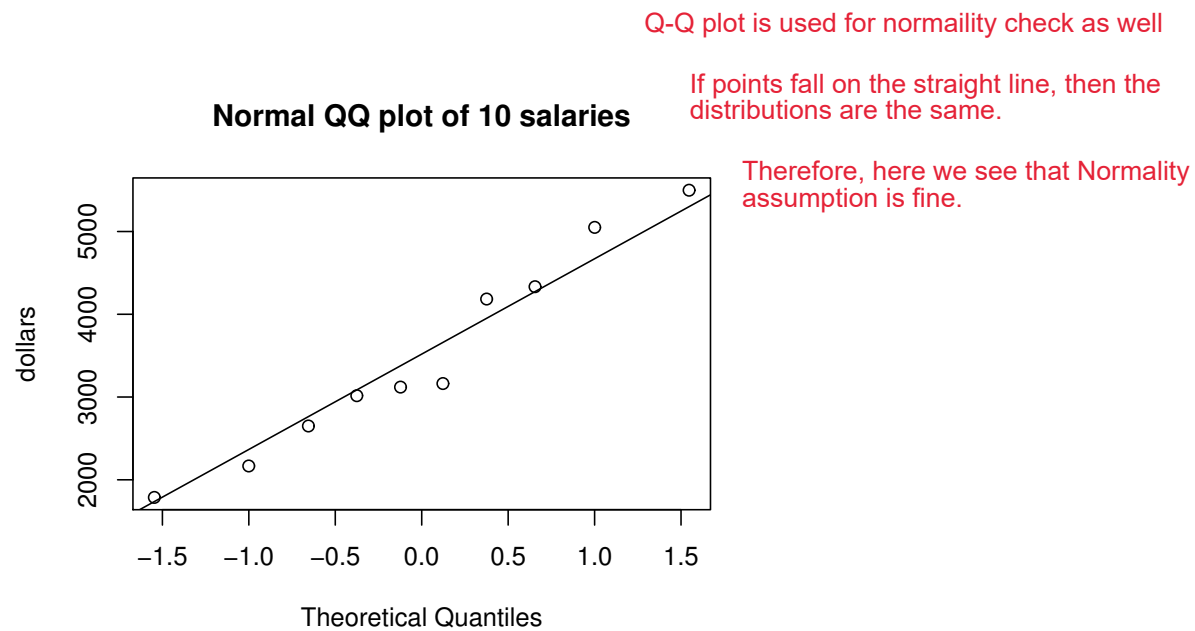


Figure 6.2: QQ plot of 10 Salaries

↳ Since, 95% C.I $\Rightarrow \alpha = 0.05$

Because, σ^2 is not assumed known, we will use the $t_{(n-1)}$ version of the confidence interval formula.

$$\hat{\mu} = \bar{y} = \$3496.90$$

$$s = \sqrt{\frac{\sum_{i=1}^n (y_i - \hat{\mu})^2}{n-1}} = \$1224.116$$

$$t_{1-\alpha/2, (n-1)} = t_{0.975, (9)} = 2.262 \quad \text{in R, } qt(0.975, 9)$$

percentile of interest
↓
Degrees of Freedom.

$$\hookrightarrow \hat{\mu} \pm t_{0.975, (9)} \frac{s}{\sqrt{n}} = [\$2621.22, \$4372.58]$$

Note:-

Consider $X \sim N(0, 1)$, and the following equation:

$$P = P(X \leq X_{\text{obs}})$$

in R, this is saying:

$$[P = pnorm(X_{\text{obs}}, mu=0, sd=1)]$$

and also,

$$[X_{\text{obs}} = qnorm(p, mu=0, sd=1)]$$

So, $[q()]$ functions return quantiles, given probabilities.]

and, $[p()]$ functions return probabilities, given quantiles.]

i.e., $pnorm(\overset{X_{\text{obs}}}{0}, 0, 1) = 0.5$

$$qnorm(\underset{p}{0.5}, 0, 1) = 0$$

Let's look at the Example 6.2.2 from the Complete Lecture Notes ...

↪ The likelihood ratio test for testing

$$H_0: \mu = \$3000 \text{ is}$$

$$D = n \ln \left[1 + \frac{1}{n-1} T^2 \right], \quad T = \frac{\bar{Y} - 3000}{S/\sqrt{n}}$$

$$p\text{-value} = P[D \geq D_{\text{obs}}]$$

$$= 2P(T \geq |t_{\text{obs}}|)$$

$$= 2P(T \geq 1.2836)$$

$$> 2(0.1), \text{ where } 0.1 \text{ was found using } t\text{-tables}$$
$$= 0.2.$$

But, before jumping into calculation,

Ask yourself, if your confidence interval from Ex: 6.2.1 tells us anything about the p-value of this test,

$$95\% = [\$2621.22, \$4372.58]$$

So, ... ?