# Background

Since its first appearance in Wuhan, China, in December 2019 [1], coronavirus disease (COVID-19) has spread internationally, including to South Korea. The first COVID-19 case in South Korea was confirmed on January 20, 2020, from a traveling resident of Wuhan, China [2]. In February, the disease spread rapidly within a church community in the city of Daegu [2]. The chains of transmission that began from this cluster distinguish the epidemic in South Korea from that in any other countries: As of March 24, 2020, 9,037 cases were confirmed, of which 56% were related to the church and 27% were in their 20s [2]. South Korea’s intensive testing using novel contact tracing techniques allowed rapid identification and isolation cases and reduction of onward transmission [3-5]. Here, we describe potential roles of social distancing in mitigating the spread of COVID-19 in South Korea by using metro traffic data to compare epidemics in two major cities.

# Data description

We analyzed epidemiological data describing the COVID-19 outbreak in South Korea between January 20–March 16, 2020. Daily number of reported cases in each geographic region was transcribed from press releases by the Korea Centers for Disease Control and Prevention (KCDC) [2]. Partial line lists were translated and transcribed from press releases by the KCDC and various local and provincial governments [6-12]. All data and original reports are stored in a publicly available GitHub repository.

We compared epidemiological dynamics of COVID-19 from two cities in which the largest number of COVID-19 cases have been reported: Daegu and Seoul. Between January 20– March 16, 2020, 6,083 cases from Daegu and 248 from Seoul were reported by the KCDC. The epidemic in Daegu is characterized by a single, large peak followed by a gradual decrease, whereas the epidemic in Seoul consists of several small outbreaks (Fig. 1).

Figure 1: **Comparison of epidemiological and traffic data from Daegu and Seoul.** Solid lines represent the daily metro traffic volume in 2020 (red) and mean daily traffic volume between 2017–2019 (black). Daily traffic from previous years have been shifted by 1–3 days to align day of the weeks. Vertical lines indicate Feb 18, 2020, when the first case was confirmed in Daegu. See Supplementary Materials for details.

Daily metro traffic in Daegu and Seoul between 2017–2020 was obtained from data.go.kr and data.seoul.go.kr, respectively. We tabulated the total number of individuals who accessed the subway or monorail (using Seoul lines 1–9, and Daegu lines 1–3; Fig. 1). Soon after the first church-related case was confirmed in Daegu on Feb 18, 2020, the daily traffic volume decreased by about 80% and 50% in Daegu and Seoul, respectively.

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|  | Parameterization | Priors | Source |
| Incubation period distribution | Gamma(*µI, µ*2*I/σ*2) | *µI* ∼ Gamma(6*.*5days*,*145) *σ σ* ∼ Gamma(2*.*6*,*25) | [14] |
| Generation-interval distribution | NegativeBinomial(*µG, θ*) | *µG* ∼ Gamma(5days*,*62) *θ θ* *θ* ∼ Gamma(5*,*20) | [15,16] |

Table 1: **Assumed incubation and generation-interval distributions.** Gamma distributions are parameterized using its mean and shape. Negative binomial distributions are parameterized using its mean and dispersion. Priors are chosen such that the 95% quantiles of prior means and standard deviations are consistent with previous estimates.

# Trends in time-dependent reproduction number and traffic volume

To estimate the time-dependent reproduction number (the average number of secondary cases caused by an average individual, given conditions at time *t* [13]), we first estimated daily incidence of infection from the daily number of reported cases by the KCDC [2]. We adjusted the number of reported cases to account for changes in testing criteria, which occurred 4 times between January 20–March 16, 2020. Then, we inferred onset-to-confirmation delay distributions from the partial line list and combined them with previously estimated incubation period distribution (Table 1) to obtain probability distributions for date of infection for each reported case. We accounted for right-censoring by dividing the daily incidence by the probability that a case infected on a given day would have been reported before March 16, 2020. Implementation details are provided in the Supplementary Materials.

We estimated the time-dependent reproduction number using the renewal equation with a 14-day sliding window [13]:

where is the reconstructed incidence time series (i.e., the number of infected cases on day *t*) and is the generation-interval distribution randomly drawn from a prior distribution (Table 1). We weighted each sample of by a gamma probability distribution with a mean of 2.6 and a standard deviation of 2 to reflect prior knowledge [17] and took weighted quantiles to calculate the medians and associated 95% credible intervals. We estimated between February 2, 2020 (14 days after the first confirmed case was imported) and March 10, 2020 (after this point the effects of censoring are too strong for reliable estimates).

Fig. 2 compares the reconstructed incidence (A,B) and estimates of (C,D) in Daegu and Seoul. In Daegu, incidence peaked shortly after the first case was confirmed and then decreased (Fig. 2A). Likewise, the estimates of gradually decrease and eventually drop below 1 about a week after the reporting of its first case, coinciding with the decrease in the metro traffic volume (Fig. 2C). The initial decrease in R*t* may reflect behavior change within the church; the first confirmed case in Daegu became symptomatic on February 7, 2020, and visited the church on February 9 and 16, 2020 [2]. Our estimates of for Daegu are consistent with the estimates of for South Korea by Abbott *et al.* [17] — their estimates drop below 1 slightly later because they rely on number of symptomatic cases instead.

Figure 2: **Comparison of reconstructed incidence and time-dependent reproduction number in Daegu and Seoul.** Black lines and gray ribbons represent the median estimates of reconstructed incidence (A,B) and (C,D) and their corresponding 95% credible intervals. Bar plots show the number of reported cases. Red lines represent the normalized traffic volume. Vertical lines indicate Feb 18, 2020, when the first case was confirmed in Daegu.

In Seoul, estimates of R*t* decreases slightly but remain around 1 (Fig. 2D). Our analysis suggests that social distancing in Seoul was less intense, and this could be why reduction in spread was less sharp. Stronger distancing or further intervention will be necessary to reduce below 1.

While we find clear, positive correlations between the normalized traffic and the median estimates of in both Daegu (*r* = 0*.*90; 95% CI: 0.79–0.95) and Seoul (*r* = 0*.*76; 95% CI: 0.59–0.87), these correlations are conflated by time trends, and are also likely conflated by other measures that could have affected . We do not find clear signatures of lagged correlation between and traffic volume (Supplementary Materials). Similar patterns in the estimates of are found in directly surrounding provinces (Gyeongsangbuk-do and Gyeonggi-do), providing support for the robustness of our analysis (Supplementary Materials).

# Discussion

The South Korean experience with COVID-19 provides evidence that epidemics can be suppressed with less extreme measures than those taken by China [18]. It demonstrates the necessity of prompt identification and isolation of cases in preventing further spread [3, 4, 5]. Our analysis reveals potential roles of social distancing in mitigating the COVID-19 epidemic in South Korea. Even though social distancing alone may not be able to fully prevent the spread of the disease, its ability to flatten the epidemic curve (cf. Fig. 2B,D) reduces burden for healthcare system and provides time to plan for the future [19].

Our study is not without limitations. We did not account for differences in the delay distributions or changes in the number of tests among cities. The intensity of intervention is likely to vary across regions given that majority of COVID-19 cases in South Korea were reported from Daegu. We did not have sufficient data to account for these factors. Nonetheless, the robustness of our findings is supported by the sensitivity analyses (Supplementary Materials). We were also unable to distinguish local and imported cases, which may overestimate R*t* [20]. We were able to perform a separate analysis for Seoul that accounts for imported cases using line list provided by the Seoul Metropolitan Government; our qualitative conclusions remained robust (Supplementary Materials).

Our analysis focused on comparing metro traffic, which serves as a proxy for the degree of social distancing, with epidemiological dynamics in two cities. The 80% decrease in traffic volume suggests that the strength of social distancing in Daegu may be comparable to that in Wuhan, China [21]. However, we are not able to directly estimate the effect of social distancing on epidemiological dynamics. Other measures, such as intensive testing of core transmission groups and school closure, are also likely to have affected the changes in [2]. Future studies should consider quantifying contributions of different measures in preventing the spread.

Finally, our study highlights the importance of considering geographical heterogeneity in estimating epidemic potential. The recent decrease in the number of reported cases in South Korea is driven by the sharp decrease in Daegu. Our analysis reveals that the epidemic may still persist in other regions, including Seoul and Gyeonggi-do; reports from Seoul and Gyeonggi-do (around 10 cases almost every day between March 11–24, 2020) provide further support for our conclusion [2]. Unless the reproduction number can be reduced below 1 in all regions, small outbreaks may continue to occur in South Korea.

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