

ECE 50024

Homework 3

Parth Sagar Hasabnis

phasabni@purdue.edu

Exercise 1:

$$a) \quad \begin{matrix} x \in \mathbb{R}^{d \times 1} \\ A \in \mathbb{R}^{d \times d} \end{matrix} \quad \therefore x^T \in \mathbb{R}^{1 \times d}$$

$$\begin{matrix} x^T & A & x \\ 1 \times d & d \times d & d \times 1 \end{matrix} \in \mathbb{R}^{1 \times 1}$$

$\therefore x^T A x$ is a scalar

$$\text{tr}[x^T A x] = x^T A x, \quad \text{Trace of a scalar is itself}$$

$$\text{tr}[x^T A x] = \text{tr}[A x x^T], \quad \text{tr}[AB] = \text{tr}[BA]$$

$$\therefore x^T A x = \text{tr}[A x x^T]$$

b) The likelihood funcⁿ is given by

$$p(\mathcal{D} | \Sigma) = \prod_{i=1}^N \left(\frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right\} \right)$$

$$= \frac{1}{(2\pi)^{Nd/2} |\Sigma|^{-N/2}} \exp \left\{ \sum_{i=1}^N -\frac{1}{2} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right\}$$

$$= \frac{1}{(2\pi)^{Nd/2}} |\Sigma^{-1}|^{N/2} \exp \left\{ -\frac{1}{2} \sum_{i=1}^N \text{tr} \left[\Sigma^{-1} (x_i - \mu) (x_i - \mu)^T \right] \right\} \quad x^T A x = \text{tr}[A x x^T]$$

$$= \frac{1}{(2\pi)^{Nd/2}} |\Sigma^{-1}|^{N/2} \exp \left\{ -\frac{1}{2} \text{tr} \left[\Sigma^{-1} \sum_{i=1}^N (x_i - \mu) (x_i - \mu)^T \right] \right\} \quad \text{tr}(A+B) = \text{tr}(A) + \text{tr}(B)$$

$$\begin{aligned}
 c) \quad p(D|\Sigma) &= \frac{1}{(2\pi)^{Nd/2}} \times \frac{1}{|\Sigma|^{N/2}} \exp \left\{ -\frac{1}{2} \operatorname{tr} \left[\Sigma^{-1} \sum_{i=1}^N (x_i - \mu)(x_i - \mu)^T \right] \right\} \\
 &= \frac{1}{(2\pi)^{Nd/2}} \times \frac{1}{|\Sigma|^{N/2}} \exp \left\{ -\frac{1}{2} \operatorname{tr} \left[\Sigma^{-1} N \tilde{\Sigma} \right] \right\} \quad \dots \quad \tilde{\Sigma} = \frac{1}{N} \sum_{n=1}^N (x_n - \mu)(x_n - \mu)^T \\
 &= \frac{1}{(2\pi)^{Nd/2}} \times \frac{1}{|\Sigma|^{N/2}} \exp \left\{ -\frac{N}{2} \operatorname{tr} [A] \right\} \quad \operatorname{tr}(NA) = N \operatorname{tr}(A), \quad A = \Sigma^{-1} \tilde{\Sigma}
 \end{aligned}$$

$$\begin{aligned}
 A &= \Sigma^{-1} \tilde{\Sigma} \\
 \Sigma A &= \tilde{\Sigma} \\
 \Sigma &= \tilde{\Sigma} A^{-1} \\
 |\Sigma| &= |\tilde{\Sigma}| |A^{-1}|
 \end{aligned}$$

$$\begin{aligned}
 |A^{-1}| &= |A|^{-1} \\
 |A| &= \prod_{i=1}^d \lambda_i \quad \lambda_1, \dots, \lambda_d \text{ are the eigenvalues of } A
 \end{aligned}$$

$$\therefore p(D|\Sigma) = \frac{1}{(2\pi)^{Nd/2}} \frac{1}{|\Sigma|^{N/2}} \left(\prod_{i=1}^d \lambda_i \right)^{N/2} \exp \left\{ -\frac{N}{2} \sum_{i=1}^d \lambda_i \right\} \quad \operatorname{tr}[A] = \sum_{i=1}^d \lambda_i$$

$$d) \quad \lambda_i = \operatorname{argmax}_{\lambda_i} p(D|\Sigma)$$

$$= \operatorname{argmin}_{\lambda_i} -\log p(D|\Sigma)$$

$$-\log p(D|\Sigma) = \frac{Nd}{2} \log(2\pi) + \frac{N}{2} \log |\tilde{\Sigma}| - \frac{N}{2} \sum_{i=1}^d \log \lambda_i + \frac{N}{2} \sum_{i=1}^d \lambda_i$$

$$\begin{aligned}
 \frac{d(-\log p(D|\Sigma))}{d\lambda_i} &= -\frac{N}{2} \times \frac{1}{\lambda_i} + \frac{N}{2} = 0 \\
 \frac{N}{2} &= \frac{N}{2\lambda_i}
 \end{aligned}$$

$$\therefore \lambda_i = 1$$

$$\therefore \lambda = [\lambda_1, \lambda_2, \dots, \lambda_d]^T = [1 \quad 1 \quad 1 \quad \dots \quad 1]_{1 \times d}^T$$

e) Consider the eigen decomposition of A

$$A = Q \Lambda Q^{-1}$$

$$\Lambda = \begin{bmatrix} \lambda_1 & & 0 \\ & \lambda_2 & \\ 0 & & \ddots \\ & & & \lambda_i \end{bmatrix} = \begin{bmatrix} 1 & & 0 \\ & \ddots & \\ 0 & & 1 \end{bmatrix} = I$$

$$\therefore A = Q Q^{-1}$$

$$\text{But } A = \Sigma^{-1} \tilde{\Sigma}$$

$$\Sigma^{-1} \tilde{\Sigma} = Q Q^{-1}$$

This implies that $\Sigma^{-1} = \tilde{\Sigma}^{-1}$

$$\therefore \Sigma = \tilde{\Sigma}$$

This is the value of Σ that maximizes $p(O|\Sigma)$

$$\therefore \tilde{\Sigma} = \hat{\Sigma}_{ML} = \frac{1}{N} \sum_{n=1}^N (x_n - \mu)(x_n - \mu)^T$$

f) An alternative way of finding $\hat{\Sigma}_{ML}$ would be :

$$\hat{\Sigma}_{ML} = \underset{\Sigma}{\operatorname{argmax}} p(O|\Sigma)$$

$$= \underset{\Sigma}{\operatorname{argmin}} -\log(p(O|\Sigma))$$

We can take the gradient of $-\log(p(O|\Sigma))$ and equate it to 0. Under first order optimality conditions and if $-\log(p(O|\Sigma))$ is convex, we can get $\hat{\Sigma}_{ML}$

g) The unbiased estimate of Σ is given by:

$$\hat{\Sigma}_{\text{unbias}} = \frac{1}{N-1} \sum_{n=1}^N (x_n - \hat{\mu})(x_n - \hat{\mu})^T$$

Exercise 2:

a)

$$\rho_{y|x}(c_i|x) = \frac{\rho_{x|y}(x|c_i) \rho(c_i)}{\rho_x(x)} \quad \rho_{y|x}(c_0|x) = \frac{\rho_{x|y}(x|c_0) \rho(c_0)}{\rho_x(x)}$$

$$\rho_x(x) = \rho_{x|y}(x|c_1) + \rho_{x|y}(x|c_0)$$

$$\rho_{y|x}(c_i|x) \stackrel{c_i}{\geq} \rho_{y|x}(c_0|x)$$

$$\frac{\rho_{x|y}(x|c_i) \rho(c_i)}{\rho_x(x)} \stackrel{c_i}{\geq} \frac{\rho_{x|y}(x|c_0) \rho(c_0)}{\rho_x(x)}$$

$$\pi_1 \frac{1}{(2\pi)^{d/2} |\Sigma_1|^{1/2}} \exp \left\{ -\frac{1}{2} (x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1) \right\} \stackrel{c_i}{\geq} \pi_0 \frac{1}{(2\pi)^{d/2} |\Sigma_0|^{1/2}} \exp \left\{ -\frac{1}{2} (x - \mu_0)^T \Sigma_0^{-1} (x - \mu_0) \right\}$$

Taking log on both sides

$$-\frac{1}{2} (x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1) + \log \pi_1 - \frac{1}{2} \log |\Sigma_1| \stackrel{c_i}{\geq} -\frac{1}{2} (x - \mu_0)^T \Sigma_0^{-1} (x - \mu_0) + \log \pi_0 - \frac{1}{2} \log |\Sigma_0|$$

b)

- The first few values in μ_0 are: [[0.48249575] [0.4864399]]

- The first few values in μ_1 are: [[0.44080734] [0.43871359]]

- The first few values in Σ_0 are:

[[0.064484 0.0369168]

[0.0369168 0.06623457]]

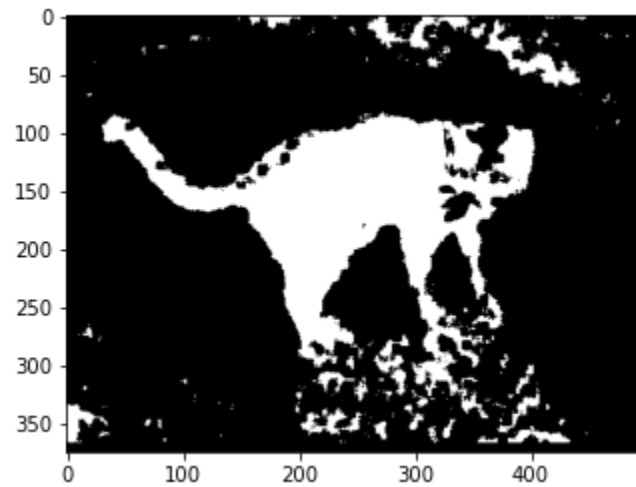
- The first few values in Σ_1 are:

[[0.04307832 0.03535405]

[0.03535405 0.0424875]]

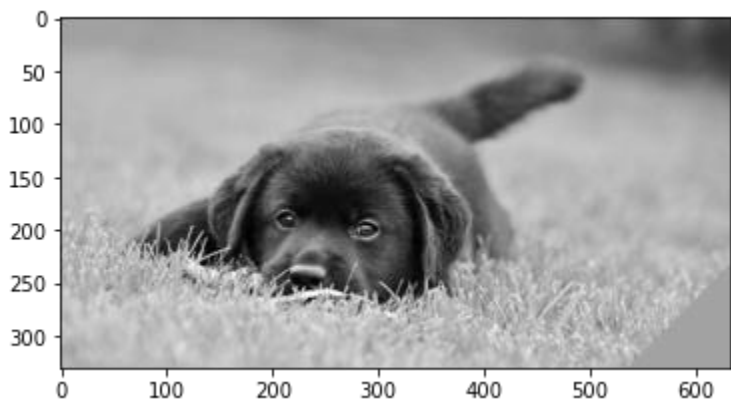
- $\pi_0 = 0.83, \pi_1 = 0.17$

c) Predicted Binary Mask

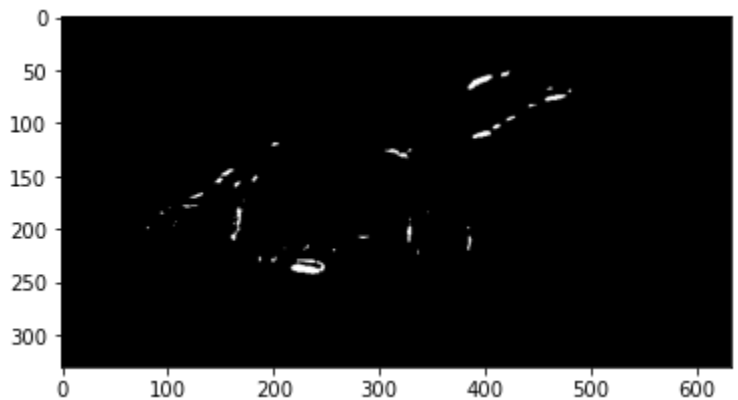


d) Mean Absolute Error = 0.087642

e) Test image and mask



Test image



Test image mask

The testing image does not perform well with the given Bayesian decision rule. This could be because the Class 1 pixels in the test image do not follow a similar distribution to the training pixels, and hence as misclassified.

Exercise 3:

a)

$$\frac{P_{X|Y}(x|C_1)}{P_{X|Y}(x|C_0)} \underset{C_0}{\overset{C_1}{\geq}} \tau$$

$$\frac{P_{X|Y}(x|C_1)}{P_{X|Y}(x|C_0)} = \frac{(2\pi)^{d/2} |\Sigma_0|^{1/2} \exp\left\{-\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1}(x-\mu_1)\right\}}{(2\pi)^{d/2} |\Sigma_1|^{1/2} \exp\left\{-\frac{1}{2}(x-\mu_0)^T \Sigma_0^{-1}(x-\mu_0)\right\}}$$

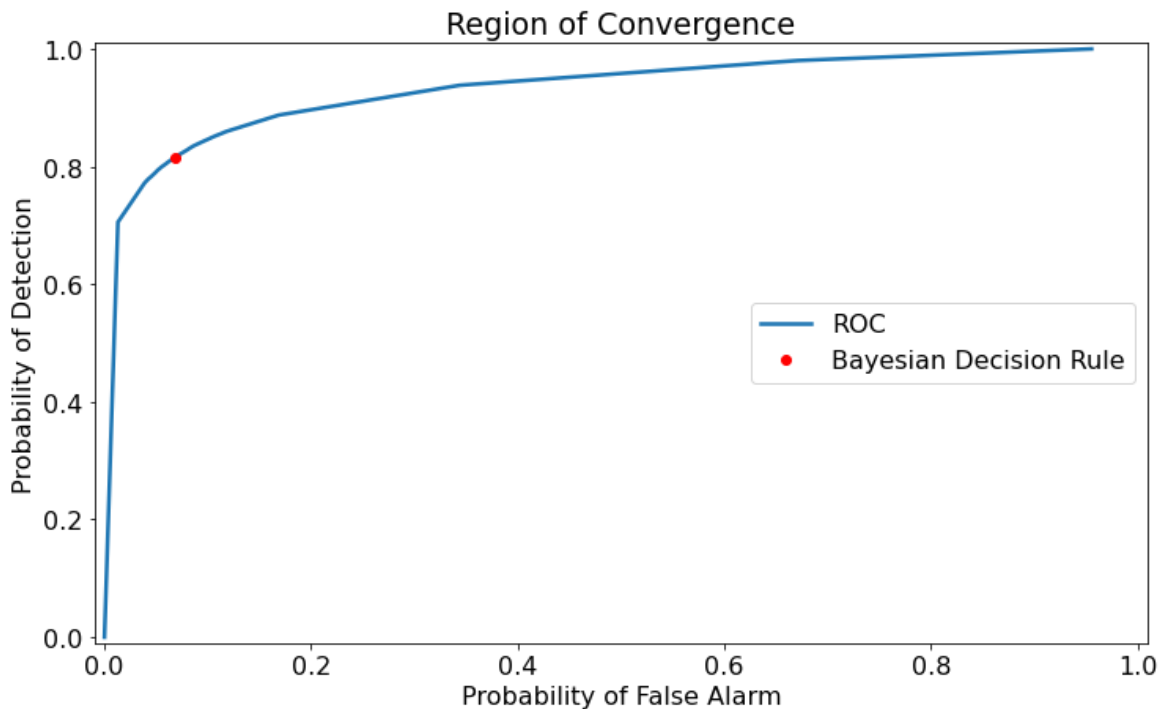
$$= \left(\frac{|\Sigma_0|}{|\Sigma_1|} \right)^{1/2} \exp\left\{ \frac{1}{2} \left[\begin{array}{c} (x-\mu_0)^T \Sigma_0^{-1}(x-\mu_0) \\ - (x-\mu_1)^T \Sigma_1^{-1}(x-\mu_1) \end{array} \right] \right\} = \tau$$

$$\text{But, } \frac{(2\pi)^{d/2} |\Sigma_0|^{1/2} \exp\left\{-\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1}(x-\mu_1)\right\}}{(2\pi)^{d/2} |\Sigma_1|^{1/2} \exp\left\{-\frac{1}{2}(x-\mu_0)^T \Sigma_0^{-1}(x-\mu_0)\right\}} = \frac{\pi_0}{\pi_1}$$

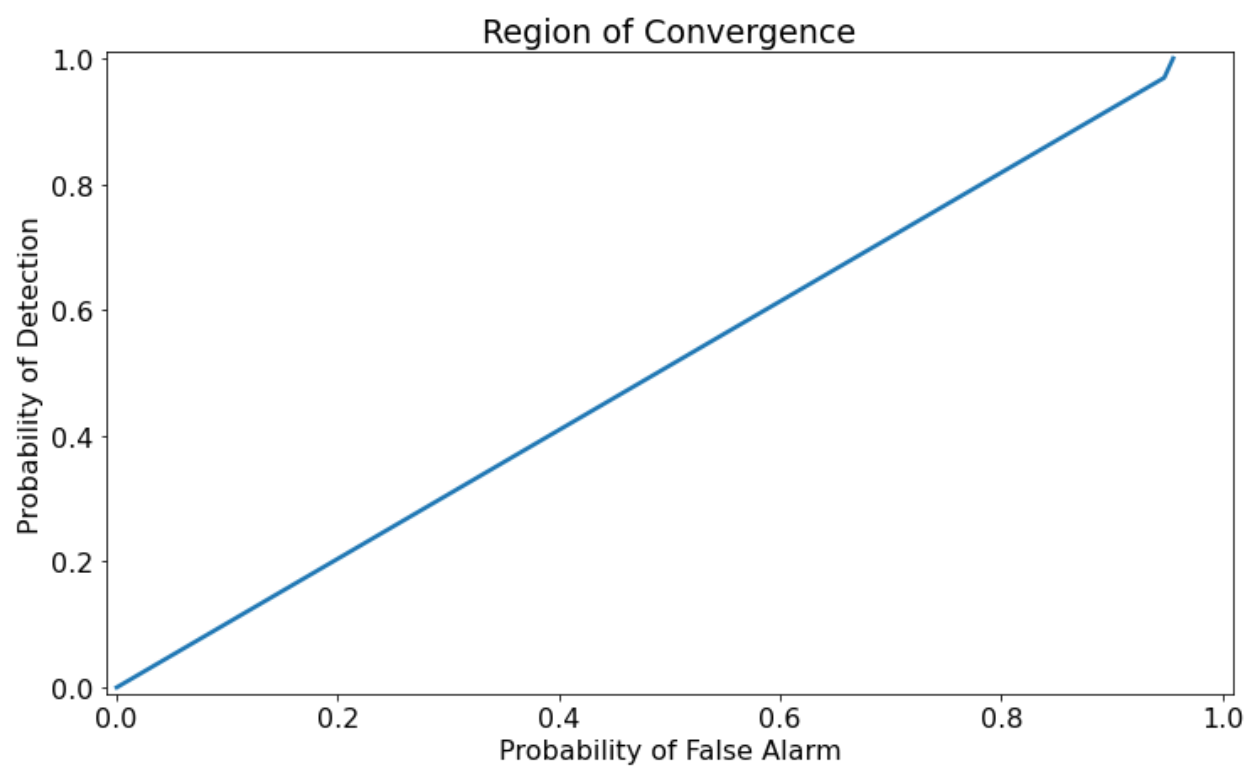
(From Exercise 2)

$$\therefore \boxed{\tau = \frac{\pi_0}{\pi_1}}$$

c)



d)



APPENDIX