

Question 1.3 - 2

```
In [1]: import numpy as np
def f_objective(theta, X, y, l2_param):
    """
    Args:
        theta: 1D numpy array of size num_features
        X: 2D numpy array of size (num_instances, num_features)
        y: 1D numpy array of size num_instances
        l2_param: regularization parameter

    Returns:
        objective: scalar value of objective function
    """
    features=X.shape[1]
    instances=X.shape[0]
    loss=0
    for i in range(instances):
        loss+=np.logaddexp(0,-y[i]*np.dot(theta,X[i]))
    loss=loss/instances
    reg=l2_param*(np.dot(theta,theta))
    return(loss+reg)
```

Question 1.3 - 3

```
In [2]: from scipy.optimize import minimize
import pandas as pd
from functools import partial
def fit_logistic_reg(X, y, objective_function, l2_param):
    """
    Args:
        X: 2D numpy array of size (num_instances, num_features)
        y: 1D numpy array of size num_instances
        objective_function: function returning the value of the objective
        l2_param: regularization parameter

    Returns:
        optimal_theta: 1D numpy array of size num_features
    """
    features= X.shape[1]
    obj_func = partial(objective_function, X = X, y=y, l2_param = l2_param)
    w_0 = np.ones(features)
    w_ = minimize(obj_func, w_0).x
    return w_
```

```
In [3]: x_train = pd.read_csv('/Users/parthvi/Downloads/X_train.txt')
x_val = pd.read_csv('/Users/parthvi/Downloads/X_val.txt')
y_train = pd.read_csv('/Users/parthvi/Downloads/y_train.txt')
y_val = pd.read_csv('/Users/parthvi/Downloads/y_val.txt')
```

```
In [4]: from sklearn.preprocessing import StandardScaler
X_train = StandardScaler().fit_transform(x_train)
X_val = StandardScaler().fit_transform(x_val)
X_train = np.hstack((X_train, np.ones((X_train.shape[0],1))))
X_val = np.hstack((X_val, np.ones((X_val.shape[0],1))))
y_train[y_train == 0] = -1
y_val[y_val == 0] = -1

#y_train = y_train.astype(int)
#y_train = np.array(y_train)
```

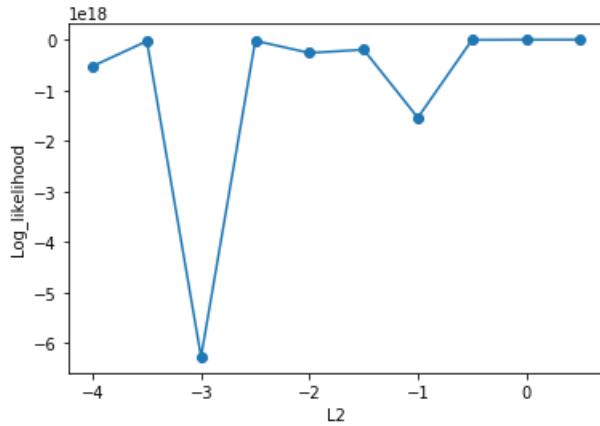
```
In [5]: y_trainn = y_train.to_numpy()
y_train = y_trainn.reshape(y_train.shape[0])
y_val = y_val.to_numpy()
y_val = y_val.reshape(y_val.shape[0])
```

```
In [6]: theta = fit_logistic_reg(X_train, y_train, f_objective, l2_param=1)
```

Question 1.3 - 4

```
In [32]: L2=list([i for i in np.arange(-4,1,0.5)])
log_likelihood_list=[]
for l2reg in L2:
    theta = fit_logistic_reg(X_train, y_train, f_objective, l2_param=l2reg)
    result = f_objective(theta,X_val,y_val, l2reg)
    log_likelihood_list.append(result)
```

```
In [33]: import matplotlib.pyplot as plt
plt.plot(L2,log_likelihood_list, 'o-')
plt.xlabel('L2')
plt.ylabel('Log_likelihood')
plt.show()
```



The lowest log_likelihood is achieved at L2 = -3

```
In [ ]:
```

DSCA 1003 - HW 5 -

Q. 2 Bayesian Logistic Regression with Gaussian priors.

1) posterior density $p(\omega | D)$

$$\rightarrow p(\omega | D) = p(\omega | x, y)$$

using Bayes rule,

$$p(y|x, \omega) \cdot p(\omega) / p(y)$$

$$\propto p(y|x, \omega) p(\omega)$$

$$\Rightarrow \prod_{i=1}^N \sigma(\omega^\top x_i)^{y_i} (1 - \sigma(\omega^\top x_i))^{1-y_i}$$

Negative log likelihood:

$$\text{Nlog } L(\omega) = - \sum_{i=1}^N y^{(i)} \log \sigma(\omega^\top x_i) + (1-y^{(i)}) \log (1 - \sigma(\omega^\top x_i))$$

$$\left\{ \boxed{P(Y=1 | x=x) = \sigma(z)} \quad \left[\text{where } \sigma = \frac{1}{1+e^{-z}} \right] \right\}$$

$$p(\omega | D) = \frac{p(D|\omega) \cdot p(\omega)}{p(D)} = \frac{\exp(-NLL_0(\omega)) p(\omega)}{p(D)}$$

\hookrightarrow prior density $p(\omega)$,

Q. 1.2

1. Decision boundary of logistic regression is a set of all points

x that satisfy:

$$P(Y=1|x) = P(Y=0|x) = \frac{1}{2}$$

$$\therefore P(y=1|x) = \frac{1}{1+e^{-\alpha^T x}}$$

$$\therefore \frac{1}{1+e^{-0.5x_+}} = \frac{1}{2}$$

$$\therefore \theta^T x_+ = 0$$

$$\therefore \partial_1 y_1 + \dots + \partial_d y_d = 0$$

for x_1, x_2

$$\theta_1 x_1 + \theta_2 x_2 = 0$$

$$\therefore x_2 = \frac{-\delta_1 x_1}{\delta_2}$$

$$\therefore \cancel{\alpha} = \boxed{x_2 = x^T \theta}, \text{ where } \theta = \frac{\theta_1}{\theta_2}$$

In general: $x = x^T Q$

In our case $\theta = w$

$$\therefore x = x^T w$$

$$\stackrel{1.2}{(2)} L(c\bar{w}) = \sum_{i=1}^n y_i \log p(x_i) + (1-y_i) \log 1-p(x_i)$$

$$= \sum_{i=1}^n \log 1-p(x_i) + \sum_{i=1}^n y_i (x_i \cdot c\bar{w})$$

$$= \sum_{i=1}^n -\log 1 + e^{c\bar{w}x_i} + \sum_{i=1}^n y_i (x_i \cdot c\bar{w})$$

$$\therefore \frac{\partial L}{\partial c} \Rightarrow - \sum_{i=1}^n \frac{1}{1+e^{c\bar{w}x_i}} \cdot e^{c\bar{w}x_i} \cdot x_{ij} + \sum_{i=1}^n y_i x_{ij} w$$

$$= \sum_{i=1}^n (y_i - p(x_i; c\bar{w})) x_{ij} w$$

Now, we can see that this is not a closed form solution. We can only approximate it numerically. As $c \uparrow$ & $p(x_i; c\bar{w}) \uparrow$.

Q. 1

1.1.

ERM with logistic loss:

$$F_{\text{Scne}} = \{x \rightarrow x^T w \mid w \in \mathbb{R}^d\}$$

$$\text{logistic} = \log(1 + \exp(-y w^T x))$$

$$\rightarrow \hat{L}_n(w) = \frac{1}{n} \sum_{i=1}^n \text{logistic}(y_i w^T x_i)$$

$$= \frac{1}{n} \sum_{i=1}^n \log(1 + \exp(-y_i w^T x_i))$$

MLE with bernoulli

$$\rightarrow p(y=1 \mid x; w) = \frac{1}{1 + \exp(-x^T w)}$$

$$\text{NLL}(w) = \sum_{i=1}^n [-y_i \log \sigma(w^T x_i)] + (y_i - 1) \log(1 - \sigma(w^T x_i))$$

st
 \hat{w} minimize $\text{NLL}(w)$; $x \rightarrow \sigma(x^T \hat{w})$
is Max. likelihood prediction fⁿ.

$$\begin{aligned}
 \text{NLL}(\omega) &= \sum_{i=1}^n [-\log \sigma(\omega^\top \alpha_i)] \\
 &= -\sum_{i=1}^n \log (1 + \exp(-y_i \omega^\top \alpha_i)) \\
 &= n \underline{\hat{R}_n}(\omega) = \text{ERM}
 \end{aligned}$$

\therefore The approaches are equivalent.

Q. 1.3

1. $J(\omega) = \hat{R}_n(\omega) + \lambda \|\omega\|^2$

→ We know, $\lambda \|\omega\|^2$ is a convex function.

We also know, exponential f's are convex.

& log of a convex function is convex.

\therefore The objective f^n is a convex f.

$$\nabla^2 J(\omega) \geq 0$$

Q. 2

No, this is ^{not} a conjugate prior to the likelihood given by logistic regression.

\because Product of gaussian prior on w & likelihood \neq posterior

3

$$-\log(p(w|D)) = NLL - \log p(w) + \log p(D)$$

We know,

$$NLL = \frac{1}{n} \hat{R}_n$$

$$\therefore -\log p(w) = -\frac{1}{2} \log((2\pi\Sigma)) + \underbrace{\frac{1}{2} w^T \Sigma^{-1} w}_{= n \lambda w^T w}$$

$$\therefore \boxed{\Sigma = \frac{1}{2n} \bar{I}}$$

Q. 4 $\frac{1}{2n\lambda} = 1 \quad (\text{if } \Sigma = \bar{I})$

$$\therefore \boxed{\lambda = \frac{1}{2n}}$$

Q. 3

$$\text{I. } \phi(Z=H|\theta_1) = \theta_1$$

$$\phi(X=H|Z=H, \theta_2) = \theta_2$$

$$P(X=T | Z=T) = 1.$$

Show that: $P(X=H | \theta_1, \theta_2) = \theta_1 \theta_2$

$$\rightarrow P(X=H) = \underbrace{P(X=H | Z=H)}_{\theta_2} \underbrace{P(Z=H)}_{\theta_1} + \underbrace{P(X=H | Z=T)}_{0} \cdot \underbrace{P(Z=T)}_{1-\theta_1}$$

$$\Rightarrow \theta_2 \cdot \theta_1$$

||

$$\Leftrightarrow N_H = n_H + n_T$$

① $P(D | \theta_1, \theta_2)$ using binomial pmf

$$\rightarrow P(D | \theta_1, \theta_2) = P(D | \theta_1) \cdot P(D | \theta_2)$$

$$= \binom{n_H}{n_H} \theta_1^{n_H} (1-\theta_1)^{n_T} \binom{N_H}{n_H} \theta_2^{n_H} (1-\theta_2)^{n_T}$$

$$\log(P(D|\theta_1, \theta_2))$$

$$= \log \left[\binom{nr}{nh} \theta_1^{nh} (1-\theta_1)^{nt} \right] \left[\binom{nr}{nh} \theta_2^{nh} (1-\theta_2)^{nt} \right]$$

$$= \log \left(\binom{nr}{nh} \theta_1^{nh} (1-\theta_1)^{nt} \right) + \log \left(\binom{nr}{nh} \theta_2^{nh} (1-\theta_2)^{nt} \right)$$

$$= \log \binom{nr}{nh} + nh \log \theta_1 + n t \log (1-\theta_1)$$

$$+ \log \binom{nr}{nh} + nh \log \theta_2 + nt \log (1-\theta_2)$$

taking ∇ w.r.t θ_1 :

$$\frac{nh}{\theta_1} + \frac{nt}{1-\theta_1} (-1) = 0$$

$$\therefore \frac{nh}{\theta_1} = \frac{nt}{1-\theta_1}$$

$$\therefore nh - \theta_1 nh = \theta_1 nt$$

$$\therefore nh = \theta_1 (nh + nt)$$

$$\therefore \theta_1 = \frac{nh}{nh + nt}$$

similarly w- θ_2

$$\theta_2 = \frac{n_h}{n_h + n_t}$$

Yes

$$\begin{aligned} 3 \quad L(\theta_1, \theta_2) &= P(D_r, D_c | \theta_1, \theta_2) \\ &= P(D_r | \theta_1, \theta_2) * P(D_c | \theta_1, \theta_2) \end{aligned}$$

$$\begin{aligned} P(D_c | \theta_1, \theta_2) &= P(O_c | \theta_1) \sim \text{Binomial} \\ &= n_{C_{\theta_1}} \theta_1^{n_h} (1 - \theta_1)^{n_t} \end{aligned}$$