

Under standing the Metropolis - Hastings Algorithm Journal of American Statistical Association - Corx lation can be introduced in sampling making the observations dependent rather than independent as a tool to reduce variance. Eg-Markov samples Acception or Rejection Sampling , x & 1Rd target density -  $\pi(\pi) = f(\pi)$ Density from which we want to sample the observations. h(x) - some other density [Researcher has to chose it] which can be simulated using some method &  $\exists c \ s.t \ f(x) \le ch(x) \ \forall \ x$ Then to obtain one random variate from sc() Stop 1: Grenvrate a candidate Z Jrom h(n)
Grencrate a value u Jrom U(0,1) If  $\frac{f(x)}{ch(x)} \leq u$  - reject Z and go to stop | Stip 2:  $\frac{f(n)}{ch(n)} \ge u - accept Z - sampled from <math>\pi(.)$ Objective - Chose run c that satisfy condition for school to decrease numbers of sujections

Drawback - Joo many rejections, hence mefficient Markov Chain Monte Carlo Simulation P(x,A) is toansition kernal for  $x \in \mathbb{R}^d$   $A \in \mathcal{B}$  S = box1 = algebra S = algebra = algebra = algebra S = algebra = algebState space is continuous P(n, A) - conditional probability of moving in set A Invariant distribution  $\pi^*(dy) = \int_{\mathbb{R}^d} P(x, dy) \pi(x) dx$ si\*(dy)= ri(y) dy [ri\*(dy) is debesque ]  $\pi (y) dy = \int_{\mathbb{R}^d} P(x, dy) \pi(x) dx$ nth iterat converges to the invarient distribution under cortain conditions Idea - Jo generaly samples from  $\pi(\cdot)$ , the algorithm while a transition kernel P(n, dy) whose  $n^{th}$  iterate converges to  $\pi(\cdot)$  for large n. Now we have to find appropriate P(x, dy) which converges to r.l.) Suppose Jor some playy)

P(x)dy) = p(x,y) dy + r(n) 8x (dy)  $\xi_{x}(dy) = 1$  y  $x \in dy$  probability stays in  $r(x) = 1 - \int_{\mathbb{R}^{d}} \rho(x,y) dy - 1$  probability x = 1 $\rho(\pi_0\pi)=0$ 

If  $\rho(x,y)$  satisfies reversibility condition  $\pi(x) \rho(x,y) = \pi(y) \rho(y,x)$ Then  $\pi(\cdot)$  is invarient distribution of  $\rho(x)$ Prove:  $\int_{A} P(n,A) \pi(n) dn$   $n, y \in A$  $= \iint_{A} \left[ \int_{A} p(x, y) dy + \gamma(x) \delta_{x}(A) \right] \pi(x) dx$ =  $\iint \int \rho(n,y) dy \pi(n) dn + \int r(n) S_n(A) \pi(n) dn$ =  $\int_{A} \left[ \int_{A} \rho(y,n) \pi(y) dy dx + \int_{A} r(n) \pi(n) dx \right]$ =  $\int_{A} \int_{A} \rho(y, n) dn \pi(y) dy + \int_{A} \sigma(n) \pi(n) dn$  $= \int_{A} \left[ 1 - \sigma(y) \right] \pi(y) dy + \int_{A} \sigma(n) \pi(n) dx$  $= \int_{A} J(y) dy$ Metropolis - Hastings Algorithm Candidali generating density = \int g(x,y) dy = 1 of (2,y). is at x, this density generate y from Objective > q(x,y) satisfy yever sibility condition J not then

π(n) q(n,y) > π(y) q(y,n) prob of going from n > y is higher than y -> n

Jo corxed it, introduct a night 2 (n,y)<1 which is multiplied on LHS to reduce probability of going from n to y. In A·R > ij xjechon, a new pair was drawn.
independently of the princes value as draws
are independent In MM-) the current value is not independendent of previous value. Thus we have q(x,y) interpreted as conditional probability rather than h(y) undependent of parcious rather a)  $\pi(n) q(n,y) \alpha(n,y) = \pi(y) q(y,n) \alpha(y,n)$ Jhus  $p_{MH}(n,y) = q(x,y) \omega(n,y)$  [P(n, A) in derivation]
is a new density that we created
using q(x,y) and  $\omega(x,y)$  which satisfy reversibility Jhin pw d(y,x) = 1 [ manimum probability] & solve for **くしれ,y)**.  $d(n,y) = \frac{T(y) q(y,x)}{I(n) q(n,y)}$ generally for both > & < care. Jhus  $d(n,y) = \min \left[ \frac{\pi(y) \varphi(y,n)}{\pi(n)}, 1 \right]$ Conditions required for convergence to invariant distri-

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