

29th April

Finding marginals, posterior probabilities

- ① Given P_X , $P_{Y|X}$. (Prior and likelihood).
 X and Y are discrete.

X is Poisson
 $Y = X + Z$
 Z is Poisson
 $\&$ independent
of X

$$P_Y(y) = \sum_{x \in \mathcal{X}_0} P_X(x) P_{Y|X}(y|x).$$

$$P_{X|Y}(x|y) = \frac{P_{Y|X}(y|x) P_X(x)}{P_Y(y)}$$

$$E[Y] = E[E[Y|X]].$$

$$= \sum_x P_X(x) E[Y|X=x]$$

$$E[Y|X=x] = \sum_{y \in \mathcal{Y}} y P_{Y|X}(y|x).$$

- ② X is discrete, Y is continuous.

Example:

$$P_X(x), \quad f_{Y|X}(y|x)$$

$Y = XZ$
 X is discrete
 $\in \{+1, -1\}$
 Z is normal.

$$f_Y(y) = \sum_{x \in \mathcal{X}_0} f_{Y|X}(y|x) P_X(x)$$

$$P_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) P_X(x)}{f_Y(y)}.$$

$$E[Y] = E[E[Y|X]]$$

$$= \sum_x P_X(x) E[Y|X=x]$$

$$E[Y|X=x] = \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy.$$

③ X is continuous, Y is discrete

$$X \sim U[0,1]$$

$$Y|X=x \sim \text{Binomial}(n, x)$$

↓ # heads on n tosses for a coin with bias x .

Given

$$f_X(x) = 1 \quad x \in [0,1]$$

$$P_{Y|X}(y|x) = \binom{n}{y} x^y (1-x)^{n-y}$$

$$P_Y(y) = \int_{-\infty}^{\infty} P_{Y|X}(y|x) f_X(x) dx$$

$$f_{X|Y}(x|y) = \frac{P_{Y|X}(y|x) f_X(x)}{P_Y(y)}$$

$$E[Y] = E[E[Y|X=x]]$$

$$= \int_{-\infty}^{\infty} f_X(x) E[Y|X=x] dx$$

$$E[Y|X=x] = \sum_{y \in \mathcal{Y}} P_{Y|X}(y|x) y$$

④ X and Y are continuous

X is Gaussian
 $Y = X + Z$
 Z is Gaussian independent of X .

$$f_X(x)$$

$$f_{Y|X}(y|x) = f_Z(y-x)$$

$$f_Y(y) = \int_{-\infty}^{\infty} f_X(x) f_{Y|X}(y|x) dx$$

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)}$$

$$E[Y] = \int_{-\infty}^{\infty} f_X(x) E[Y|X=x] dx$$

$$E[Y|X=x] = \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy$$

ML estimator

Y is continuous.

$$\arg \max_x f_{Y|X}(y|x), \quad := \hat{x}_{ML}(y).$$

Y is discrete

$$\hat{x}_{ML}(y) = \arg \max_x P_{Y|X}(y|x)$$

MAP estimator

X is continuous.

Best estimator
to minimize
decision error probability:
 $P(X \neq \hat{X})$.

$$\begin{aligned} \arg \max_x f_{X|Y}(x|y) &= \arg \max_x \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)} \\ &= \arg \max_x f_{Y|X}(y|x) f_X(x). \\ &= \hat{x}_{MAP}(y). \end{aligned}$$

MMSE estimator

$$\hat{x}_{MMSE}(y) = E[X|Y=y]$$

$$\begin{aligned} &\text{MSE} \\ &E[(X - \hat{X})^2] \end{aligned}$$

$$= E[E[(X - \hat{X})^2|Y]]$$

Linear MMSE

$$\hat{x}_{LMMSE}(y) = ay + b.$$

$$b = E[X - aY].$$

$$a = \frac{\text{Cov}(X, Y)}{\text{Var}(Y)}.$$

Multiple measurements

$$\hat{x}_{ML}(y_1, y_2, \dots, y_n) = \arg \max_x f_{Y_1 \dots Y_n | X}(y_1 \dots y_n | x)$$

$$\hat{x}_{MAP}(y_1, \dots, y_n) = \arg \max_x f_{X | Y_1 \dots Y_n}(x | y_1 \dots y_n)$$

$$\hat{x}_{MMSE}(y_1, \dots, y_n) = E[X | Y_1 = y_1, \dots, Y_n = y_n]$$

$$\hat{x}_{LMMSE}(y_1, \dots, y_n) = \underbrace{R_{XY}}_{\substack{\text{Cov}(X, Y_1) \text{ Cov}(X, Y_2) \dots \text{Cov}(X, Y_n)}} K_Y^{-1} y.$$

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}.$$

MSE error
of constant
estimate

\geq

MSE error
of LMMSE

\geq

MSE error of
MMSE estimate

\wedge

MSE error of
ML, MAP
estimates.

Functions of one R.V

Continuous RVs.

$$Y = g(X)$$

invertible fn.

$$f_Y(y) = \frac{f_X(x)}{|g'(x)|}$$

\nwarrow
 $x: y = g(x)$

$$Y = X^2.$$

x_1, x_2, \dots, x_n result in $g(x_i) = y$.

$$f_Y(y) = \sum_{i=1}^n \frac{f_X(x_i)}{|g'(x_i)|}$$

One Function of two R.V.s

$$F_Z(z) = F_X(z) F_Y(z)$$

$$Z = \min(x, y)$$

$$Z = x + y$$

$$Z = \max(x, y)$$

$$1 - F_Z(z) = (1 - F_X(x)) (1 - F_Y(y))$$

$$F_Z(z) = P(X + Y \leq z)$$

Two Functions of two RVs

$$Z = g_1(x, y)$$

$$W = g_2(x, y)$$

$$f_{Z,W}(z, w) = \frac{f_{X,Y}(x, y)}{J(x, y)}$$

(x, y) s.t.
 $z = g_1(x, y)$
 $w = g_2(x, y)$

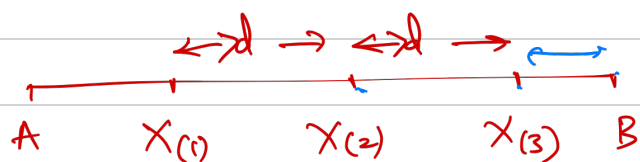
$$= f_{X,Y}(x, y) J(z, w)$$

max and min joint density.
Order statistics

$X_{(1)}, X_{(2)}, \dots, X_{(n)}$ is the ordered statistics of X_1, \dots, X_n

\downarrow \downarrow \downarrow
 smallest second smallest largest

$$f_{X_{(1)}, \dots, X_{(n)}}(x_1, \dots, x_n) = n! \prod_{i=1}^n f_X(x_i) \cdot x_1 < x_2 < \dots < x_n$$



$$\begin{array}{ccc}
 X_1, & X_2, & X_3 \\
 \downarrow & & \\
 X_{(1)} & X_{(2)} & X_{(3)}
 \end{array}$$

$$\begin{aligned}
 P \left(X_{(2)} \geq X_{(1)} + d, \quad X_{(3)} \geq X_{(2)} + d \right) \\
 = \int_0^{1-2d} \int_{x_1+d}^{1-d} \int_{x_2+d}^1 f_{X_{(1)} X_{(2)} X_{(3)}}(x_1, x_2, x_3) dx_3 dx_2 dx_1
 \end{aligned}$$

MGF ,

Concentration Inequality ,

Convergence of RVs ; Gaussian Vectors