ECCCos from the Black Box: Faithful Explanations through Energy-Constrained Conformal Counterfactuals

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Abstract

Counterfactual Explanations offer an intuitive and straightforward way to explain black-box models and offer Algorithmic Recourse to individuals. To address the need for plausible explanations, existing work has primarily relied on surrogate models to learn how the input data is distributed. This effectively reallocates the task of learning realistic explanations for the data from the model itself to the surrogate. Consequently, the generated explanations may seem plausible to humans but need not necessarily describe the behaviour of the black-box model faithfully. We formalise this notion of faithfulness through the introduction of a tailored evaluation metric and propose a novel algorithmic framework for generating Energy-Constrained Conformal Counterfactuals (ECCCos) that are only as plausible as the model permits. Through extensive empirical studies involving multiple synthetic and real-world datasets, we demonstrate that ECCCos reconcile the need for plausibility and faithfulness. In particular, we show that it is possible to achieve state-of-the-art plausibility for models with gradient access without the need for surrogate models. To do so, our framework relies solely on properties defining the black-box model itself by leveraging recent advances in energy-based modelling and conformal prediction. To our knowledge, this is the first venture in this direction for generating faithful Counterfactual Explanations. Thus, we anticipate that ECCCos can serve as a baseline for future research. We believe that our work opens avenues for researchers and practitioners seeking tools to better distinguish trustworthy from unreliable models.

1 Introduction

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Counterfactual Explanations (CE) provide a powerful, flexible and intuitive way to not only explain 23 black-box models but also help affected individuals through the means of Algorithmic Recourse. 24 Instead of opening the Black Box, CE works under the premise of strategically perturbing model 25 inputs to understand model behaviour [33]. Intuitively speaking, we generate explanations in this 26 context by asking what-if questions of the following nature: 'Our credit risk model currently predicts 27 that this individual is not credit-worthy. What if they reduced their monthly expenditures by 10%? 28 This is typically implemented by defining a target outcome $\mathbf{y}^+ \in \mathcal{Y}$ for some individual $\mathbf{x} \in \mathcal{X} = \mathbb{R}^D$ 29 described by D attributes, for which the model $M_{\theta}: \mathcal{X} \mapsto \mathcal{Y}$ initially predicts a different outcome: 30 $M_{\theta}(\mathbf{x}) \neq \mathbf{y}^+$. Counterfactuals are then searched by minimizing a loss function that compares the 31 predicted model output to the target outcome: yloss $(M_{\theta}(\mathbf{x}), \mathbf{y}^+)$. Since Counterfactual Explanations 32 work directly with the black-box model, valid counterfactuals always have full local fidelity by 33 construction where fidelity is defined as the degree to which explanations approximate the predictions of a black-box model [21, 20].

Artificial Intelligence (XAI) than other popular approaches like LIME [26] and SHAP [17], which 37 involve local surrogate models. But even full fidelity is not a sufficient condition for ensuring 38 that an explanation faithfully describes the behaviour of a model. That is because multiple very 39 distinct explanations can all lead to the same model prediction, especially when dealing with heavily 40 parameterized models like deep neural networks, which are typically underspecified by the data [35]. 41 In the context of CE, the idea that no two explanations are the same arises almost naturally. A key focus in the literature has therefore been to identify those explanations and algorithmic recourses that 43 are most appropriate based on a myriad of desiderata such as sparsity, actionability and plausibility. 44

In situations where full fidelity is a requirement, CE offers a more appropriate solution to Explainable

- In this work, we draw closer attention to model faithfulness rather than fidelity as a desideratum for 45 counterfactuals. Our key contributions are as follows: 46
 - We show that fidelity is an insufficient evaluation metric for counterfactuals (Section 3) and propose a definition of faithfulness that gives rise to more suitable metrics (Section 4).
 - We introduce a novel algorithmic approach for generating Energy-Constrained Conformal Counterfactuals (ECCCos) in Section 5.
 - We provide extensive empirical evidence demonstrating that ECCCos faithfully explain model behaviour without sacrificing plausibility (Section 6).

Thus, we believe that our work opens avenues for researchers and practitioners seeking tools to better distinguish trustworthy from unreliable models.

Background 2

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While Counterfactual Explanations can be generated for arbitrary regression models [28], existing 56 work has primarily focused on classification problems. Let $\mathcal{Y} = (0,1)^K$ denote the one-hot-encoded 57 output domain with K classes. Then most counterfactual generators rely on gradient descent to optimize different flavours of the following counterfactual search objective:

$$\mathbf{Z}' = \arg\min_{\mathbf{Z}' \in \mathcal{Z}^L} \left\{ yloss(M_{\theta}(f(\mathbf{Z}')), \mathbf{y}^+) + \lambda cost(f(\mathbf{Z}')) \right\}$$
(1)

Here yloss denotes the primary loss function, $f(\cdot)$ is a function that maps from the counterfactual state space to the feature space and cost is either a single penalty or a collection of penalties that are used to impose constraints through regularization. Equation 1 restates the baseline approach to gradient-based counterfactual search proposed by Wachter et al. [33] in general form as introduced by Altmeyer et al. [2]. To explicitly account for the multiplicity of explanations $\mathbf{Z}' = \{\mathbf{z}_l\}_L$ denotes an L-dimensional array of counterfactual states.

The baseline approach, which we will simply refer to as **Wachter** [33], searches a single counterfactual directly in the feature space and penalises its distance to the original factual. In this case, $f(\cdot)$ is simply the identity function and Z corresponds to the feature space itself. Many derivative works of Wachter et al. [33] have proposed new flavours of Equation 1, each of them designed to address specific desiderata that counterfactuals ought to meet in order to properly serve both AI practitioners and 70 individuals affected by algorithmic decision-making systems. The list of desiderata includes but is not limited to the following: sparsity, proximity [33], actionability [31], diversity [21], plausibility [11, 25, 27], robustness [30, 24, 2] and causality [14]. Different counterfactual generators addressing 73 these needs have been extensively surveyed and evaluated in various studies [32, 13, 23, 4, 10].

Perhaps unsurprisingly, the different desiderata are often positively correlated. For example, Artelt et al. [4] find that plausibility typically also leads to improved robustness. Similarly, plausibility has also been connected to causality in the sense that plausible counterfactuals respect causal relationships [18]. Consequently, the plausibility of counterfactuals has been among the primary concerns for researchers. Achieving plausibility is equivalent to ensuring that the generated counterfactuals comply with the true and unobserved data-generating process (DGP). We define plausibility formally in this work as follows:

Definition 2.1 (Plausible Counterfactuals). Let $\mathcal{X}|\mathbf{y}^+$ denote the true conditional distribution of samples in the target class y^+ . Then for x' to be considered a plausible counterfactual, we need: $\mathbf{x}' \sim \mathcal{X} | \mathbf{y}^+$.

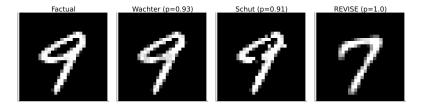


Figure 1: Counterfactuals for turning a 9 (nine) into a 7 (seven): original image (left); then from left to right the counterfactuals generated using Wachter, Schut and REVISE

To generate plausible counterfactuals, we need to be able to quantify the DGP: $\mathcal{X}|\mathbf{y}^+$. One straightforward way to do this is to use surrogate models for the task. Joshi et al. [11], for example, suggest that instead of searching counterfactuals in the feature space \mathcal{X} , we can instead traverse a latent embedding \mathcal{Z} (Equation 1) that implicitly codifies the DGP. To learn the latent embedding, they propose using a generative model such as a Variational Autoencoder (VAE). Provided the surrogate model is well-trained, their proposed approach called **REVISE** can yield plausible explanations. Others have proposed similar approaches: Dombrowski et al. [6] traverse the base space of a normalizing flow to solve Equation 1; Poyiadzi et al. [25] use density estimators ($\hat{p}: \mathcal{X} \mapsto [0, 1]$) to constrain the counterfactuals to dense regions in the feature space; and, finally, Karimi et al. [14] assume knowledge about the structural causal model that generates the data.

A competing approach towards plausibility that is also closely related to this work instead relies on 95 the black-box model itself. Schut et al. [27] show that to meet the plausibility objective we need not 96 explicitly model the input distribution. Pointing to the undesirable engineering overhead induced by 97 surrogate models, they propose that we rely on the implicit minimisation of predictive uncertainty 98 instead. Their proposed methodology, which we will refer to as **Schut**, solves Equation 1 by greedily 99 applying JSMA in the feature space with standard cross-entropy loss and no penalty at all. The 100 authors demonstrate theoretically and empirically that their approach yields counterfactuals for which 101 the model M_{θ} predicts the target label \mathbf{y}^+ with high confidence. Provided the model is well-specified, 102 these counterfactuals are plausible. This idea hinges on the assumption that the black-box model 103 provides well-calibrated predictive uncertainty estimates. 104

3 Why Fidelity is not Enough

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As discussed in the introduction, any valid Counterfactual Explanation also has full fidelity by construction: solutions to Equation 1 are considered valid as soon as the label predicted by the model matches the target class. So while fidelity always applies, counterfactuals that address the various desiderata introduced above can look vastly different from each other. The following motivating example illustrates this point and demonstrates why fidelity is an insufficient evaluation metric to assess the faithfulness of Counterfactual Explanations.

We have trained a simple image classifier M_{θ} on the well-known MNIST dataset [15]: a Multi-Layer Perceptron (MLP) with above 90 percent test accuracy. No measures have been taken to improve the model's adversarial robustness or its capacity for predictive uncertainty quantification. The far left panel of Figure 1 shows a random sample drawn from the dataset. The underlying classifier correctly 115 predicts the label 'nine' for this image. For the given factual image and model, we have used Wachter, 116 Schut and REVISE to generate one counterfactual each in the target class 'seven'. The perturbed 117 images are shown next to the factual image from left to right in Figure 1. Captions on top of the 118 119 individual images indicate the generator along with the predicted probability that the image belongs 120 to the target class. In all three cases that probability is above 90 percent and yet the counterfactuals look very different from each other. 121

Since Wachter is only concerned with proximity, the generated counterfactual is almost indistinguishable from the factual. The approach by Schut expects a well-calibrated model that can generate predictive uncertainty estimates. Since this is not the case, the generated counterfactual looks like an adversarial example. Finally, the counterfactual generated by REVISE looks much more plausible than the other two. But is it also more faithful to the behaviour of our MNIST classifier? That is much less clear because the surrogate used by REVISE introduces friction: the generated explanations no longer depend exclusively on the black-box model itself.

So which of the counterfactuals most faithfully explains the behaviour of our image classifier? Fidelity cannot help us to make that judgement, because all of these counterfactuals have full fidelity. To bridge this gap, we introduce a new notion of faithfulness in the following section.

4 A new Notion of Faithfulness

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Analogous to Definition 2.1, we propose to define faithfulness in the context of Counterfactual Explanations as follows:

Definition 4.1 (Faithful Counterfactuals). Let $\mathcal{X}_{\theta}|\mathbf{y}^{+} = p_{\theta}(\mathbf{X}_{\mathbf{y}^{+}})$ denote the conditional distribution of \mathbf{x} in the target class \mathbf{y}^{+} , where θ denotes the parameters of model M_{θ} . Then for \mathbf{x}' to be considered a conformal counterfactual, we need: $\mathbf{x}' \sim \mathcal{X}_{\theta}|\mathbf{y}^{+}$.

In doing this, we merge in and nuance the concept of plausibility (Definition 2.1) where the notion of consistent with the data' becomes 'consistent with what the model has learned about the data'.

4.1 Quantifying the Model's Generative Property

To assess counterfactuals with respect to Definition 4.1, we need a way to quantify the posterior conditional distribution $p_{\theta}(\mathbf{x}|\mathbf{y}^{+})$. To this end, we draw on recent advances in Energy-Based Modelling (EBM), a subdomain of machine learning that is concerned with generative or hybrid modelling [8?]. In particular, note that if we fix \mathbf{y} to our target value \mathbf{y}^{+} , we can conditionally draw from $p_{\theta}(\mathbf{x}|\mathbf{y}^{+})$ using Stochastic Gradient Langevin Dynamics (SGLD) as follows,

$$\mathbf{x}_{j+1} \leftarrow \mathbf{x}_j - \frac{\epsilon^2}{2} \mathcal{E}(\mathbf{x}_j | \mathbf{y}^+) + \epsilon \mathbf{r}_j, \quad j = 1, ..., J$$
 (2)

where $\mathbf{r}_{j} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$ is the stochastic term and the step-size ϵ is typically polynomially decayed [34]. The term $\mathcal{E}(\mathbf{x}_{j}|\mathbf{y}^{+})$ denotes the model energy conditioned on the target class label \mathbf{y}^{+} which we specify as the negative logit corresponding to the target class label \mathbf{y}^{*} . To allow for faster sampling, we follow the common practice of choosing the step-size ϵ and the standard deviation of \mathbf{r}_{j} separately. While \mathbf{x}_{J} is only guaranteed to distribute as $p_{\theta}(\mathbf{x}|\mathbf{y}^{*})$ if $\epsilon \to 0$ and $J \to \infty$, the bias introduced for a small finite ϵ is negligible in practice [22, 8]. Appendix A provides additional implementation details for any tasks related to energy-based modelling.

Generating multiple samples using SGLD thus yields an empirical distribution $\hat{\mathbf{X}}_{\theta,\mathbf{y}^+}$ that approximates what the model has learned about the input data. While in the context of Energy-Based Modelling, this is usually done during training, we propose to repurpose this approach during inference in order to evaluate and generate faithful model explanations.

4.2 Evaluating Plausibility and Faithfulness

The parallels between our definitions of plausibility and faithfulness imply that we can also use similar evaluation metrics in both cases. Since existing work has focused heavily on plausibility, it offers a useful starting point. In particular, Guidotti [10] have proposed an implausibility metric that measures the distance of the counterfactual from its nearest neighbour in the target class. As this distance is reduced, counterfactuals get more plausible under the assumption that the nearest neighbour itself is plausible in the sense of Definition 2.1. In this work, we use the following adapted implausibility metric that relaxes this assumption,

$$impl = \frac{1}{|\mathbf{x} \in \mathbf{X}_{\mathbf{y}^+}|} \sum_{\mathbf{x} \in \mathbf{X}_{\mathbf{y}^+}} dist(\mathbf{x}', \mathbf{x})$$
(3)

where X_{v^+} is a subsample of the training data in the target class y^+ .

This gives rise to a very similar evaluation metric for unfaithfulness. We merely swap out the subsample of individuals in the target class for a subset $\hat{\mathbf{X}}_{\theta,\mathbf{v}^+}^{n_E}$ of the generated conditional samples:

$$\text{unfaith} = \frac{1}{|\mathbf{x} \in \hat{\mathbf{X}}_{\theta, \mathbf{y}^+}^{n_E}|} \sum_{\mathbf{x} \in \hat{\mathbf{X}}_{\theta, \mathbf{y}^+}^{n_E}} \text{dist}(\mathbf{x}', \mathbf{x})$$
(4)

Specifically, we form this subset based on the n_E generated samples with the lowest energy.

5 Energy-Constrained Conformal Counterfactuals (ECCCo)

In this section, we describe our proposed framework for generating Energy-Constrained Conformal Counterfactuals (ECCCos). It is based on the premise that counterfactuals should be faithful, first and foremost. Plausibility, as a secondary concern, is then still attainable but only to the degree that the black-box model itself has learned plausible explanations for the underlying data.

We begin by stating our proposed objective function, which involves tailored loss and penalty functions that we will explain in the following. In particular, we extend Equation 1 as follows:

$$\mathbf{Z}' = \arg\min_{\mathbf{Z}' \in \mathcal{Z}^M} \{ yloss(M_{\theta}(f(\mathbf{Z}')), \mathbf{y}^+) + \lambda_1 dist(f(\mathbf{Z}'), \mathbf{x}) + \lambda_2 dist(f(\mathbf{Z}'), \hat{\mathbf{x}}_{\theta}) + \lambda_3 \Omega(C_{\theta}(f(\mathbf{Z}'); \alpha)) \}$$
(5)

The first penalty term involving λ_1 induces proximity like in Wachter et al. [33]. Our default choice for $\operatorname{dist}(\cdot)$ is the L1 Norm due to its sparsity-inducing properties. The second penalty term involving λ_2 constrains the energy of the generated counterfactual by penalising its distance from the lowest-energy conditional samples as defined in Equation 4. Intuitively, this component induces faithfulness which coincides with plausibility to the extent that the model M_{θ} has learned the true posterior conditional distribution of inputs: $p_{\theta}(\mathbf{X_{v^+}}) \to p(\mathbf{X_{v^+}})$.

The third and final penalty term involving λ_3 introduces a new but familiar concept: it ensures that the generated counterfactual is associated with low predictive uncertainty. As mentioned above, Schut et al. [27] have shown that plausible counterfactuals can be generated implicitly through predictive uncertainty minimization. Unfortunately, this relies on the assumption that the model itself can provide predictive uncertainty estimates, which may be too restrictive in practice.

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To relax this assumption, we leverage recent advances in Conformal Prediction (CP), an approach to predictive uncertainty quantification that has recently gained popularity [3, 19]. Crucially for our intended application, CP is model-agnostic and can be applied during inference without placing any restrictions on model training. Intuitively, CP works under the premise of turning heuristic notions of uncertainty into rigorous uncertainty estimates by repeatedly sifting through the training data or a dedicated calibration dataset. Conformal classifiers produce prediction sets for individual inputs that include all output labels that can be reasonably attributed to the input. These sets tend to be larger for inputs that do not conform with the training data and are therefore characterized by high predictive uncertainty.

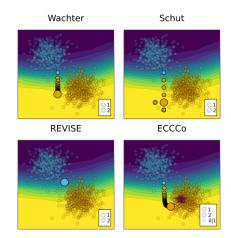
In order to generate counterfactuals that are associated with low predictive uncertainty, we use a smooth set size penalty introduced by Stutz et al. [29] in the context of conformal training:

$$\Omega(C_{\theta}(\mathbf{x}; \alpha)) = \max \left(0, \sum_{\mathbf{y} \in \mathcal{Y}} C_{\theta, \mathbf{y}}(\mathbf{x}_{i}; \alpha) - \kappa\right)$$
(6)

Here, $\kappa \in \{0,1\}$ is a hyper-parameter and $C_{\theta,\mathbf{y}}(\mathbf{x}_i;\alpha)$ can be interpreted as the probability of label y being included in the prediction set.

In order to compute this penalty for any black-box model we merely need to perform a single calibration pass through a holdout set \mathcal{D}_{cal} . Arguably, data is typically abundant and in most applications, practitioners tend to hold out a test data set anyway. Consequently, CP removes the restriction on the family of predictive models, at the small cost of reserving a subset of the available data for calibration. This particular case of conformal prediction is referred to as Split Conformal Prediction (SCP) as it involves splitting the training data into a proper training dataset and a calibration dataset.

In addition to the smooth set size penalty, we have also experimented with the use of a tailored 207 function for yloss(\cdot) that enforces that only the target label y^+ is included in the prediction set Stutz 208 et al. [29]. Further details are described Appendix B.



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Figure 2: An example involving linearly separable synthetic data and illustrating how ECCCo compares to Wachter, Schut and REVISE. The factual class is 2 and the target class is 1. Contours indicate the predicted probability given by a Joint Energy Model that the counterfactual belongs to the target class.

Algorithm 1: Generating ECCCos (For more details, see Appendix C)

$$\mathbf{x}, \mathbf{y}^+, M_{\theta}, f, \Lambda, \alpha, \mathcal{D}, T, \eta, n_{\mathcal{B}}, N_{\mathcal{B}}$$

where $M_{\theta}(\mathbf{x}) \neq \mathbf{y}^+$

Output: x'

- 1: Initialize $\mathbf{z}' \leftarrow f^{-1}(\mathbf{x})$
- 2: Generate buffer \mathcal{B} of $N_{\mathcal{B}}$ conditional samples $\hat{\mathbf{x}}_{\theta}|\mathbf{y}^{+}$ using SGLD (Equation 2)
- 3: Run *SCP* for M_{θ} using \mathcal{D}
- 4: Initialize $t \leftarrow 0$
- 5: **while** not converged or t < T **do**

6:
$$\hat{\mathbf{x}}_{\theta,t} \leftarrow \operatorname{rand}(\mathcal{B}, n_{\mathcal{B}})$$

7:
$$\mathbf{z}' \leftarrow \mathbf{z}' \quad \mathbf{z}' \quad \eta \nabla_{\mathbf{z}'} \mathcal{L}(\mathbf{z}', \mathbf{y}^+, \hat{\mathbf{x}}_{\theta, t}; \Lambda, \alpha)$$

- 8: $t \leftarrow t + 1$ 9: end while
- 10: $\mathbf{x}' \leftarrow f(\mathbf{z}')$

The entire procedure for generating ECCCos is described in Algorithm 1. For the sake of simplicity and without loss of generality, we limit our attention to generating a single counterfactual $\mathbf{x}' = f(\mathbf{z}')$ where in contrast to Equation 5 \mathbf{z}' denotes a 1-dimensional array containing a single counterfactual state. That state is initialized by passing the factual x through the encoder f^{-1} which in our case corresponds to a simple feature transformer, rather than the encoder part of VAE as in REVISE [11]. Next, we generate a buffer of N_B conditional samples $\hat{\mathbf{x}}_{\theta}|\mathbf{y}^+$ using SGLD (Equation 2) and conformalise the model M_{θ} through Split Conformal Prediction on training data \mathcal{D} .

Finally, we search counterfactuals through gradient descent. Let $\mathcal{L}(\mathbf{z}', \mathbf{y}^+, \hat{\mathbf{x}}_{\theta,t}; \Lambda, \alpha)$ denote our loss 218 function defined in Equation 5. Then in each iteration, we first randomly draw n_B samples from 219 the buffer \mathcal{B} before updating the counterfactual state \mathbf{z}' by moving in the negative direction of that 220 loss function. The search terminates once the convergence criterium is met or the maximum number of iterations T has been exhausted. Note that the choice of convergence criterium has important 222 implications on the final counterfactual (for more detail on this see Appendix C). 223

Figure 2 illustrates how ECCCos compare to counterfactuals generated using Wachter, Schut and REVISE. The example involves synthetically generated linearly separable data that belong to one of two classes. Contours indicate the predicted probabilities of a Joint Energy Model that has been jointly trained to predict the output class and generate inputs Grathwohl et al. [8]. We have drawn a random sample from the factual class 1 and used each generator to produce a counterfactual in the target class 2. Both Wachter and Schut yield valid counterfactuals but fail to achieve plausibility in the sense that the generated counterfactuals are far away from the densely populated region in the target class. Conversely, ECCCo yields a faithful and plausible counterfactual in the neighbourhood of the generated conditional samples. REVISE fails to yield a valid counterfactual because the underlying surrogate has failed to learn the DGP.

Empirical Analysis

Our goal in this section is to shed light on the following research questions: 235

Research Question 6.1 (Faithfulness). Are ECCCos more faithful than counterfactuals produced by our benchmark generators?

Research Question 6.2 (Plausibility). How do ECCCos compare to state-of-the-art generators with respect to plausibility?

We first briefly describe our experimental setup, before presenting our main results.

6.1 Experimental Setup

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To assess and benchmark the performance of ECCCo against the state of the art, we generate multiple 242 counterfactuals for different black-box models and datasets. In particular, we compare ECCCo to the following counterfactual generators that were introduced above: firstly; Schut [27], which works under the premise of minimizing predictive uncertainty; secondly, REVISE [11], which is 245 state-of-the-art with respect to plausibility; and, finally, Wachter [33], which serves as our baseline. 246 We also consider two variations of ECCCo: 'ECCCo (no CP)' involves no set size penalty ($\lambda_3 = 0$ in Equation 5), while 'ECCCo (no EBM)' does not penalise the distance to samples generated through 248 SGLD ($\lambda_2 = 0$ in Equation 5). These have been added to gain some sense of the degree to which 249 the two components underlying ECCCo-namely energy-based modelling (EBM) and conformal 250 prediction (CP)—drive the results.

We use both synthetic and real-world datasets from different domains, all of which are publically available and commonly used to train and benchmark classification algorithms. We synthetically generate a dataset containing two **Linearly Separable** Gaussian clusters (n=1000), as well as the well-known **Circles** (n=1000) and **Moons** (n=2500) data. Since these data are generated by distributions of varying degrees of complexity, they allow us to gauge how complexity affects the different generators.

As for real-world data, we follow Schut et al. [27] and use the MNIST [15] dataset containing images of handwritten digits such as the examples shown above. From the social sciences domain, we include Give Me Some Credit (GMSC) [12]: a tabular dataset that has been studied extensively in the literature on Algorithmic Recourse [23]. It consists of 11 numeric features that can be used to predict the binary outcome variable indicating whether or not retail borrowers experience financial distress.

For the predictive modelling tasks, we use simple neural networks (MLP) and Joint Energy Models 263 (JEM). For the more complex real-world datasets we also use ensembling in each case. Both joint-energy modelling and ensembling are associated with generative properties and adversarial 265 robustness, so we expect this to be positively correlated with the plausibility of ECCCos. To account 266 for stochasticity, we generate multiple counterfactuals for each possible target class, generator, model 267 and dataset. Specifically, we randomly sample n^- times from the subset of individuals for which 268 the given model predicts the non-target class y^- given the current target. We set $n^- = 25$ for all 269 of our synthetic datasets, $n^- = 10$ for GMSC and $n^- = 5$ for MNIST. Full details concerning our 270 parameter choices, training procedures and model performance can be found in Appendix D. 271

6.2 Results for Synthetic Data

Table 1 shows the key results for the synthetic datasets separated by model (first columns) and generator (second column). The numerical columns show the average values of our key evaluation metrics computed across all counterfactuals. Standard deviations are shown in parentheses. In bold we have highlighted the best outcome for each model and metric. To provide some sense of effect sizes, we have added asterisks to indicate that a given value is at least one (*) or two (**) standard deviations lower than the baseline (Wachter).

Starting with the high-level results for our Linearly Separable data, we find that ECCCo produces 279 the most faithful counterfactuals for both black-box models. This is consistent with our design since 280 281 ECCCo directly enforces faithfulness through regularization. Crucially though, ECCCo also produces the most plausible counterfactuals for both models. This dataset is so simple that even the MLP has learned plausible explanations of the input data. Zooming in on the granular details for the Linearly 283 Separable data, the results for 'ECCCo (no CP)' and 'ECCCo (no EBM)' indicate that the positive 284 results are dominated by the effect of quantifying and leveraging the model's generative property 285 (EBM). Conformal Prediction alone only leads to marginally improved faithfulness and plausibility 286 relative to the benchmark generators. 287

The findings for the Moons dataset are broadly in line with the findings so far: for the JEM, ECCCo yields significantly more faithful and plausible counterfactuals than all other generators. For the MLP,

Table 1: Results for synthetic datasets. Standard deviations across samples are shown in parentheses. Best outcomes are highlighted in bold. Asterisks indicate that the given value is more than one (*) or two (**) standard deviations away from the baseline (Wachter).

		Linearly Separable		Moons		Circles	
Model	Generator	Unfaithfulness ↓	Implausibility ↓	Unfaithfulness ↓	Implausibility ↓	Unfaithfulness ↓	Implausibility \downarrow
JEM	ECCCo	0.03 (0.06)**	0.20 (0.08)**	0.31 (0.30)*	1.20 (0.15)**	0.52 (0.36)	1.22 (0.46)
	ECCCo (no CP)	0.03 (0.06)**	0.20 (0.08)**	0.37 (0.30)*	1.21 (0.17)**	0.54(0.39)	1.21 (0.46)
	ECCCo (no EBM)	0.16 (0.11)	0.34 (0.19)	0.91 (0.32)	1.71 (0.25)	0.70 (0.33)	1.30 (0.37)
	REVISE	0.19 (0.03)	0.41 (0.01)**	0.78 (0.23)	1.57 (0.26)	0.48 (0.16)*	0.95 (0.32)*
	Schut	0.39 (0.07)	0.73 (0.17)	0.67 (0.27)	1.50 (0.22)*	0.54 (0.43)	1.28 (0.53)
	Wachter	0.18 (0.10)	0.44 (0.17)	0.80 (0.27)	1.78 (0.24)	0.68 (0.34)	1.33 (0.32)
MLP	ECCCo	0.29 (0.05)**	0.23 (0.06)**	0.80 (0.62)	1.69 (0.40)	0.65 (0.53)	1.17 (0.41)
	ECCCo (no CP)	0.29 (0.05)**	0.23 (0.07)**	0.79 (0.62)	1.68 (0.42)	0.49 (0.35)	1.19 (0.44)
	ECCCo (no EBM)	0.46 (0.05)	0.28 (0.04)**	1.34 (0.47)	1.68 (0.47)	0.84 (0.51)	1.23 (0.31)
	REVISE	0.56 (0.05)	0.41 (0.01)	1.45 (0.44)	1.64 (0.31)	0.58 (0.52)	0.95 (0.32)
	Schut	0.43 (0.06)*	0.47 (0.36)	1.45 (0.55)	1.73 (0.48)	0.58 (0.37)	1.23 (0.43)
	Wachter	0.51 (0.04)	0.40 (0.08)	1.32 (0.41)	1.69 (0.32)	0.83 (0.50)	1.24 (0.29)

faithfulness is maintained but counterfactuals are not plausible. This high-level pattern is broadly consistent other more complex datasets and supportive of our narrative, so it is worth highlighting: ECCCos consistently achieve high faithfulness, which—subject to the quality of the model itself—coincides with high plausibility. By comparison, REVISE yields the most plausible counterfactuals for the MLP, but it does so at the cost of faithfulness. We also observe that the best results for ECCCo are achieved when using both penalties. Once again though, the generative component (EBM) has a stronger impact on the positive results for the JEM.

For the Circles data, it appears that REVISE performs well but we note that it generates valid counterfactuals only half of the time (see Appendix E for a complete overview of all evaluation metrics). It turns out that in this case, the underlying VAE with default parameters has not adequately learned the data-generating process. Of course, it is possible to achieve better generative performance through hyperparameter tuning but this example serves to illustrate that REVISE depends strongly on the quality of the surrogate model. Independent of the outcome for REVISE, however, the results do not seem to indicate that ECCCo significantly improves faithfulness and plausibility for the Circles data. We think this points to a limitation of our evaluation metrics rather than ECCCo itself: computing average distances fails to account for the 'wraparound' effect associated with circular data [7].

6.3 Results for Real-World Data

The results for our real-world datasets are shown in Table 2. Once again the findings indicate that the plausibility of ECCCos is positively correlated with the capacity of the black-box model to distinguish plausible from implausible inputs. The case is very clear for MNIST: ECCCos are consistently more faithful than the corresponding counterfactuals produced by any of the benchmark generators and their plausibility gradually improves through ensembling and joint-energy modelling. For the JEM Ensemble, ECCCo is essentially on par with REVISE and does significantly better than the baseline generator. We also note that ECCCo is the only generator that consistently achieves full validity for all models (Appendix E). Interestingly, ECCCo also yields lower-cost outcomes than the baseline generator for the JEMs.

For the tabular credit dataset (GMSC) it is inherently challenging to use deep neural networks in order to achieve good discriminative performance [5, 9] and discriminative performance [16], respectively. In order to achieve high plausibility, ECCCo effectively requires classifiers to achieve good performance for both tasks. Since this is a challenging task even for Joint Energy Models, it is not surprising to find that even though ECCCo once again achieves state-of-the-art faithfulness, it is outperformed by REVISE and Schut with respect to plausibility.

6.4 Key Takeways

To conclude this section, we summarize our findings with reference to the opening questions. The results have clearly demonstrated that ECCCo consistently achieves state-of-the-art faithfulness, as it was designed to do (Research Question 6.1). A related important finding is that ECCCo yields highly

Table 2: Results for real-world datasets. Standard deviations across samples are shown in parentheses. Best outcomes are highlighted in bold. Asterisks indicate that the given value is more than one (*) or two (**) standard deviations away from the baseline (Wachter).

		MN	IST	GMSC			
Model	Generator	Unfaithfulness↓	Implausibility ↓	Unfaithfulness ↓	Implausibility \downarrow		
	ECCCo	81.78 (17.49)**	299.40 (29.48)**	199.40 (38.02)	17.26 (5.64)**		
JEM	REVISE	190.01 (28.89)**	263.01 (46.46)**	206.57 (41.88)	4.86 (0.90)**		
JEM	Schut	210.14 (27.35)**	286.50 (40.67)**	197.85 (37.95)	6.46 (2.11)**		
	Wachter	280.70 (26.17)	499.25 (38.25)	195.02 (32.35)	68.48 (60.80)		
	ECCCo	72.46 (11.11)**	276.48 (26.75)**	182.04 (26.62)	16.85 (4.49)**		
JEM Ensemble	REVISE	173.81 (22.22)**	248.50 (41.54)**	206.02 (41.79)	4.76 (0.63)**		
JEM Elisellible	Schut	202.69 (22.90)**	282.77 (39.72)**	204.53 (24.20)	6.53 (1.55)**		
	Wachter	272.32 (23.03)	494.77 (37.74)	185.59 (33.79)	59.26 (49.56)		
	ECCCo	155.25 (22.13)**	519.58 (33.92)	177.98 (39.03)	19.09 (5.00)**		
MLP	REVISE	367.93 (14.90)**	256.16 (44.23)**	201.61 (30.74)	5.33 (1.74)**		
MILP	Schut	382.40 (16.67)*	286.14 (41.39)**	199.35 (32.06)	6.84 (1.96)**		
	Wachter	406.24 (17.34)	488.30 (39.64)	195.51 (23.99)	81.62 (54.15)		
	ECCCo	144.74 (20.08)**	484.56 (31.26)	196.45 (34.80)*	20.18 (5.20)**		
MLP Ensemble	REVISE	340.33 (13.32)**	251.30 (42.13)**	202.67 (27.80)*	4.82 (0.40)**		
WILF Elisemble	Schut	358.83 (13.17)*	283.12 (43.27)**	199.64 (42.29)*	6.35 (1.66)**		
	Wachter	375.22 (18.91)	456.68 (47.21)	244.65 (44.55)	63.00 (53.77)		

plausible explanations provided that they faithfully describe model behaviour (Research Question 6.2).
Our findings here also indicate that ECCCo achieves this result primarily by leveraging the model's generative property.

7 Limitations

Even though we have taken considerable measures to study our proposed methodology carefully, this work is limited in scope, which caveats our findings. In particular, we have found that the performance of ECCCo is sensitive to hyperparameter choices. In order to achieve faithfulness, we generally had to penalise the distance from generated samples slightly more than the distance from factual values. This choice is associated with relatively higher costs to individuals since the proposed recourses typically involve more substantial feature changes than for our benchmark generators.

Conversely, we have not found that penalising prediction set sizes disproportionately strongly had any discernable effect on our results. Our results indicate that Conformal Prediction alone is often not sufficient to achieve faithfulness and plausibility, although we acknowlege that this needs to be investigated more thoroughly through future work.

Furthermore, while our approach is readily applicable to models with gradient access like deep neural networks, more work is needed to generalise our methodology to other popular machine learning models such as gradient-boosted trees. Relatedly, common challenges associated with energy-based modelling during our experiments including sensitivity to scale, training instabilities and sensitivity to hyperparameters also apply to ECCCo.

8 Conclusion

This work leverages recent advances in energy-based modelling and conformal prediction in the context of Explainable Artificial Intelligence. We have proposed a new way to generate Counterfactual Explanations that are maximally faithful to the black-model they aim to explain. Our proposed counterfactual generator, ECCCo, produces plausible counterfactual if and only if the black-model itself has learned realistic representations of the data, which we demonstrate through rigorous empirical analysis. This should enable researchers and practitioners to use counterfactuals in order to discern trustworthy models from unreliable ones. While the scope of this work limits its generalizability, we believe that ECCCo offers a solid baseline for future work on faithful Counterfactual Explanations.

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446 Appendices

447 **A JEM**

While \mathbf{x}_J is only guaranteed to distribute as $p_{\theta}(\mathbf{x}|\mathbf{y}^+)$ if $\epsilon \to 0$ and $J \to \infty$, the bias introduced for a small finite ϵ is negligible in practice [22, 8]. While Grathwohl et al. [8] use Equation 2 during training, we are interested in applying the conditional sampling procedure in a post-hoc fashion to any standard discriminative model.

452 B Conformal Prediction

The fact that conformal classifiers produce set-valued predictions introduces a challenge: it is not 453 immediately obvious how to use such classifiers in the context of gradient-based counterfactual search. Put differently, it is not clear how to use prediction sets in Equation 1. Fortunately, Stutz et al. 455 [29] have recently proposed a framework for Conformal Training that also hinges on differentiability. 456 Specifically, they show how Stochastic Gradient Descent can be used to train classifiers not only 457 for the discriminative task but also for additional objectives related to Conformal Prediction. One 458 such objective is efficiency: for a given target error rate α , the efficiency of a conformal classifier 459 improves as its average prediction set size decreases. To this end, the authors introduce a smooth set 460 size penalty defined in Equation 6 in the body of this paper 461

Formally, it is defined as $C_{\theta,\mathbf{y}}(\mathbf{x}_i;\alpha) := \sigma\left((s(\mathbf{x}_i,\mathbf{y}) - \alpha)T^{-1}\right)$ for $\mathbf{y} \in \mathcal{Y}$, where σ is the sigmoid function and T is a hyper-parameter used for temperature scaling [29].

Intuitively, CP works under the premise of turning heuristic notions of uncertainty into rigorous uncertainty estimates by repeatedly sifting through the data. It can be used to generate prediction intervals for regression models and prediction sets for classification models [1]. Since the literature on CE and AR is typically concerned with classification problems, we focus on the latter. A particular variant of CP called Split Conformal Prediction (SCP) is well-suited for our purposes, because it imposes only minimal restrictions on model training.

Specifically, SCP involves splitting the data $\mathcal{D}_n = \{(\mathbf{x}_i, \mathbf{y}_i)\}_{i=1,...,n}$ into a proper training set $\mathcal{D}_{\text{train}}$ and a calibration set \mathcal{D}_{cal} . The former is used to train the classifier in any conventional fashion. The latter is then used to compute so-called nonconformity scores: $\mathcal{S} = \{s(\mathbf{x}_i, \mathbf{y}_i)\}_{i \in \mathcal{D}_{\text{cal}}}$ where $s: (\mathcal{X}, \mathcal{Y}) \mapsto \mathbb{R}$ is referred to as *score function*. In the context of classification, a common choice for the score function is just $s_i = 1 - M_{\theta}(\mathbf{x}_i)[\mathbf{y}_i]$, that is one minus the softmax output corresponding to the observed label \mathbf{y}_i [3].

Finally, classification sets are formed as follows,

$$C_{\theta}(\mathbf{x}_i; \alpha) = \{ \mathbf{y} : s(\mathbf{x}_i, \mathbf{y}) \le \hat{q} \}$$
(7)

calibration set increases, the probability that the classification set $C(\mathbf{x}_{test})$ for a newly arrived sample \mathbf{x}_{test} does not cover the true test label \mathbf{y}_{test} approaches α [3].

Observe from Equation 7 that Conformal Prediction works on an instance-level basis, much like Counterfactual Explanations are local. The prediction set for an individual instance \mathbf{x}_i depends only on the characteristics of that sample and the specified error rate. Intuitively, the set is more likely to include multiple labels for samples that are difficult to classify, so the set size is indicative of predictive uncertainty. To see why this effect is exacerbated by small choices for α consider the case of $\alpha = 0$, which requires that the true label is covered by the prediction set with probability equal to

where \hat{q} denotes the $(1-\alpha)$ -quantile of S and α is a predetermined error rate. As the size of the

487 C Conformal Prediction

488 D Experimental Setup

489 E Results

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Table 3: All results for all datasets. Standard deviations across samples are shown in parentheses. Best outcomes are highlighted in bold. Asterisks indicate that the given value is more than one (*) or two (**) standard deviations away from the baseline (Wachter).

Model	Data	Generator ECCCo	Cost ↓ 39.14 (3.71)	Unfaithfulness \(\) 236.79 (51.16)	Implausibility ↓ 39.78 (3.18)	Redundancy ↑ 0.00 (0.00)	Uncertainty ↓ 2.00 (0.00)	Validity ↑ 1.00 (0.00)
	JEM	REVISE	4.39 (2.08)	284.51 (52.74)	5.58 (0.81)**	0.01 (0.03)	1.85 (0.32)	1.00 (0.00)
	JEWI	Schut	4.17 (1.84)	263.55 (60.56)	8.00 (2.03)	0.25 (0.24)*	1.88 (0.31)	1.00 (0.00)
		Wachter ECCCo	2.03 (1.01) 34.85 (4.67)	274.55 (51.17) 249.44 (58.53)	7.32 (1.80) 35.09 (5.56)	0.00 (0.00)	1.90 (0.31) 2.00 (0.00)	1.00 (0.00) 1.00 (0.00)
	JEM Ensemble	REVISE	4.53 (1.97)	268.45 (66.87)	5.44 (0.74)**	0.00 (0.00)	1.95 (0.21)	1.00 (0.00)
	JEWI EIISEIIIDIE	Schut	0.98 (0.38)**	279.38 (63.23)	7.64 (1.47)	0.84 (0.06)**	2.00 (0.00)	1.00 (0.00)
California Housing		Wachter ECCCo	2.00 (0.59) 37.47 (4.59)	268.59 (68.66) 230.92 (48.86)	7.16 (1.46) 37.53 (5.40)	0.00 (0.00)	1.90 (0.31) 1.00 (0.00)**	1.00 (0.00) 1.00 (0.00)
	MD	REVISE	3.38 (2.06)	281.10 (53.01)	5.34 (0.67)**	0.00 (0.00)	1.10 (0.31)	1.00 (0.00)
	MLP	Schut	0.88 (0.51)**	285.12 (56.00)	6.48 (1.18)**	0.72 (0.22)**	1.00 (0.00)**	1.00 (0.00)
		Wachter ECCCo	5.35 (10.88) 38.33 (4.99)	262.50 (56.87) 212.47 (59.27)*	9.21 (10.41)	0.00 (0.00)	1.05 (0.22)	1.00 (0.00)
		REVISE	3.41 (1.79)	284.65 (49.52)	38.17 (6.18) 5.64 (1.13)*	0.00 (0.00)	1.00 (0.00)** 1.05 (0.22)	1.00 (0.00) 1.00 (0.00)
	MLP Ensemble	Schut	0.84 (0.56)**	269.19 (46.08)	7.30 (1.94)	0.81 (0.11)**	1.00 (0.00)**	1.00 (0.00)
		Wachter	2.00 (1.39)	278.09 (73.65)	7.32 (1.75)	0.00 (0.00)	1.07 (0.23)	1.00 (0.00)
		ECCCo ECCCo (no CP)	1.34 (1.48) 1.33 (1.49)	0.63 (1.58) 0.64 (1.61)	1.44 (1.37) 1.45 (1.38)	0.00 (0.00) 0.00 (0.00)	0.98 (0.14) 0.98 (0.14)	0.98 (0.14) 0.98 (0.14)
	JEM	ECCCo (no EBM)	0.85 (1.49)	1.41 (1.51)	1.50 (1.38)	0.00 (0.00)	1.04 (0.28)	0.98 (0.14)
		REVISE	0.99 (0.35)	0.96 (0.32)*	0.95 (0.32)*	0.00 (0.00)	0.50 (0.51)	0.50 (0.51)
		Schut Wachter	1.00 (0.43) 0.74 (1.50)	0.99 (0.80) 1.41 (1.50)	1.28 (0.53) 1.51 (1.35)	0.25 (0.25) 0.00 (0.00)	1.11 (0.38) 0.98 (0.14)	1.00 (0.00)** 0.98 (0.14)
Circles		ECCCo	1.39 (0.23)	0.37 (0.65)**	1.30 (0.68)	0.00 (0.00)	1.00 (0.00)**	1.00 (0.00)
		ECCCo (no CP)	1.33 (0.28)	0.50 (0.85)*	1.28 (0.66)	0.00 (0.00)	1.04 (0.20)*	1.00 (0.00)
	MLP	ECCCo (no EBM)	1.15 (0.69)	2.00 (1.46)	1.83 (1.00)	0.00 (0.00)	0.97 (0.10)**	1.00 (0.00)
		REVISE Schut	0.98 (0.36) 0.61 (0.11)	1.16 (1.05) 1.60 (1.15)	0.95 (0.32)* 1.24 (0.44)	0.00 (0.00) 0.34 (0.24)*	0.50 (0.51)* 1.00 (0.00)**	0.50 (0.51) 1.00 (0.00)
		Wachter	0.53 (0.15)	1.67 (1.05)	1.31 (0.43)	0.00 (0.00)	1.28 (0.46)	1.00 (0.00)
		ECCCo	859.68 (91.05)	40.65 (5.67)**	605.67 (19.56)	0.00 (0.00)	3.00 (0.00)**	1.00 (0.00)
	JEM	REVISE Schut	500.28 (86.07) 10.00 (0.00)**	693.81 (118.47)* 871.82 (64.75)	467.88 (132.24) 561.81 (94.76)	0.00 (0.00) 0.99 (0.00)**	3.20 (2.28)** 0.00 (0.00)**	0.80 (0.45) 0.00 (0.00)
		Wachter	100.86 (13.85)	902.84 (88.79)	586.49 (97.17)	0.00 (0.00)	10.00 (0.00)	1.00 (0.00)
		ECCCo	679.19 (66.95)	59.61 (32.93)**	500.50 (27.51)	0.00 (0.00)	4.00 (0.00)**	1.00 (0.00)
	JEM Ensemble	REVISE	476.47 (147.09)	533.64 (102.81)*	356.60 (79.57)*	0.00 (0.00)	4.80 (1.30)**	1.00 (0.00)
		Schut Wachter	10.00 (0.00)** 92.50 (9.31)	688.61 (86.83) 714.63 (54.58)	445.55 (99.03) 470.54 (96.18)	0.99 (0.00)** 0.00 (0.00)	0.00 (0.00)** 10.00 (0.00)	0.00 (0.00) 1.00 (0.00)
FashionMNIST		ECCCo	885.97 (29.70)	65.36 (20.64)**	791.07 (14.51)	0.00 (0.00)	2.00 (0.00)**	1.00 (0.00)**
	MLP	REVISE	323.10 (102.63)	856.08 (73.66)	394.73 (252.67)	0.00 (0.00)	1.00 (1.00)**	0.60 (0.55)
		Schut Wachter	10.00 (0.00)** 94.57 (10.26)	928.77 (42.27) 916.45 (50.09)	518.98 (143.30) 546.35 (145.24)	0.99 (0.00)** 0.00 (0.00)	0.00 (0.00)** 3.61 (4.01)	0.00 (0.00) 0.80 (0.45)
	MLP Ensemble	ECCCo	869.65 (67.92)	47.37 (7.72)**	751.83 (11.87)	0.00 (0.00)	1.00 (0.00)**	1.00 (0.00)
		REVISE	267.88 (69.67)	822.34 (57.55)	307.50 (105.09)*	0.00 (0.00)	3.00 (4.00)	0.80 (0.45)
		Schut Wachter	10.00 (0.00)** 91.50 (16.35)	891.57 (70.10) 874.21 (59.36)	449.79 (149.32) 476.59 (150.76)	0.99 (0.00)** 0.00 (0.00)	0.00 (0.00)** 4.60 (4.93)	0.00 (0.00) 1.00 (0.00)
	JEM	ECCCo	40.78 (8.79)**	41.65 (17.24)**	40.57 (8.74)**	0.00 (0.00)	1.50 (0.51)	1.00 (0.00)**
		REVISE	5.10 (6.48)**	74.89 (15.82)**	6.01 (5.75)**	0.00 (0.00)	1.81 (0.40)	1.00 (0.00)**
		Schut Wachter	1.10 (0.39)** 127.26 (75.11)	76.23 (15.54)** 146.02 (64.48)	6.02 (0.72)** 128.93 (74.00)	0.77 (0.09)** 0.00 (0.00)	1.55 (0.51) 1.00 (1.03)	1.00 (0.00)** 0.50 (0.51)
		ECCCo	33.87 (8.25)**	26.55 (12.94)**	33.65 (8.33)**	0.00 (0.00)	2.00 (0.00)	1.00 (0.00)**
	JEM Ensemble	REVISE	6.00 (4.92)**	52.47 (14.12)**	6.69 (3.37)**	0.00 (0.00)	1.80 (0.52)	0.95 (0.22)**
	JEW Ensemble	Schut Wachter	1.29 (0.92)**	56.34 (15.00)** 125.72 (70.80)	6.27 (1.06)**	0.74 (0.16)** 0.00 (0.00)	1.62 (0.52) 1.00 (1.03)	1.00 (0.00)** 0.50 (0.51)
GMSC		ECCCo	124.35 (95.08) 38.91 (7.68)**	46.90 (15.80)**	126.55 (93.75) 37.78 (8.40)**	0.00 (0.00)	1.00 (0.00)	1.00 (0.00)
	MLP	REVISE	4.16 (2.35)**	81.08 (19.53)**	4.60 (0.72)**	0.00 (0.00)	1.23 (0.40)	1.00 (0.00)
	WILI	Schut	0.72 (0.32)**	90.67 (20.80)**	5.56 (0.81)**	0.87 (0.06)** 0.00 (0.00)	1.00 (0.00) 1.00 (0.00)	1.00 (0.00)
		Wachter ECCCo	199.28 (14.78) 72.42 (145.72)	191.68 (30.86) 74.65 (144.69)*	200.23 (15.05) 71.87 (145.19)	0.00 (0.00)	1.00 (0.00)	1.00 (0.00)
	MLP Ensemble	REVISE	4.75 (2.94)**	80.90 (14.59)**	5.20 (1.52)**	0.00 (0.00)	1.07 (0.12)	1.00 (0.00)
	WILL Elisemble	Schut	0.65 (0.24)**	85.63 (19.15)**	6.00 (0.99)**	0.88 (0.04)**	1.00 (0.00)**	1.00 (0.00)
		Wachter ECCCo	202.64 (14.71) 0.91 (0.14)	220.05 (17.41) 0.10 (0.06)**	203.65 (14.77) 0.19 (0.03)**	0.00 (0.00)	1.00 (0.00) 0.97 (0.03)**	1.00 (0.00) 1.00 (0.00)
		ECCCo (no CP)	0.91 (0.14)	0.10 (0.07)**	0.19 (0.03)**	0.00 (0.00)	0.98 (0.03)**	1.00 (0.00)
	JEM	ECCCo (no EBM)	0.90 (0.17)	0.37 (0.28)	0.38 (0.26)	0.00 (0.00)	1.23 (0.49)	1.00 (0.00)
		REVISE Schut	0.42 (0.14)* 1.14 (0.27)	0.41 (0.02)** 0.66 (0.23)	0.41 (0.01)** 0.66 (0.22)	0.00 (0.00) 0.21 (0.25)	0.81 (0.82) 1.74 (0.43)	0.50 (0.51) 1.00 (0.00)
Linearly Cananal-1-		Wachter	0.61 (0.12)	0.44 (0.16)	0.44 (0.15)	0.00 (0.00)	1.50 (0.50)	1.00 (0.00)
Linearly Separable		ECCCo	1.52 (0.16)	0.03 (0.02)**	0.69 (0.10)	0.00 (0.00)	1.00 (0.00)**	1.00 (0.00)
		ECCCo (no CP) ECCCo (no EBM)	1.52 (0.16) 2.66 (1.10)	0.03 (0.02)** 1.25 (0.87)	0.68 (0.10) 1.84 (1.10)	0.00 (0.00) 0.00 (0.00)	1.00 (0.00)** 1.00 (0.00)**	1.00 (0.00) 1.00 (0.00)
	MLP	REVISE	0.44 (0.13)*	1.10 (0.10)	0.40 (0.01)**	0.00 (0.00)	1.64 (0.78)	0.82 (0.39)
		Schut	0.76 (0.14)	0.81 (0.10)*	0.47 (0.24)	0.26 (0.25)*	1.00 (0.00)**	1.00 (0.00)
		Wachter ECCCo	0.60 (0.14) 269.99 (57.02)**	0.94 (0.11) 116.09 (30.70)**	0.44 (0.15) 281.33 (41.51)**	0.00 (0.00)	1.54 (0.50) NA	1.00 (0.00) 1.00 (0.00)**
	JEM	REVISE	143.79 (43.43)**	348.74 (65.65)**	246.69 (36.69)**	0.00 (0.01)	NA	0.80 (0.40)
	JEWI	Schut	9.90 (0.55)**	355.58 (64.84)**	270.06 (40.41)**	0.99 (0.00)**	NA	0.15 (0.36)
		Wachter ECCCo	453.86 (16.96) 260.94 (52.14)**	694.08 (50.86) 89.89 (27.26)**	630.99 (33.01) 240.59 (37.41)**	0.00 (0.00)	NA NA	0.90 (0.30) 1.00 (0.00)**
	IEME ::	REVISE	138.82 (33.99)**	292.52 (53.13)**	240.50 (35.73)**	0.00 (0.00)	NA NA	0.81 (0.39)
	JEM Ensemble	Schut	9.97 (0.28)**	319.45 (59.02)**	266.80 (40.46)**	0.99 (0.00)**	NA	0.05 (0.22)
MNIST		Wachter	365.46 (35.14)	582.52 (58.46)	543.90 (44.24) 649.63 (58.80)	0.00 (0.00)	NA NA	0.96 (0.20) 1.00 (0.00)
	MLP	ECCCo REVISE	658.48 (65.03) 150.41 (51.81)**	212.45 (36.70)** 839.79 (77.14)*	649.63 (58.80) 244.33 (38.69)**	0.00 (0.00)	NA NA	0.95 (0.22)
		Schut	9.95 (0.41)**	842.80 (82.01)*	264.94 (42.18)**	0.99 (0.00)**	NA	0.06 (0.25)
	MLP Ensemble	Wachter	400.08 (34.33)	982.32 (61.81)	561.23 (45.08)	0.00 (0.00)	NA NA	1.00 (0.00)
		ECCCo REVISE	616.12 (102.01) 149.48 (47.90)**	741.30 (125.98)*	587.65 (95.01) 242.76 (41.16)**	0.00 (0.00) 0.00 (0.01)	NA NA	1.00 (0.00)** 0.92 (0.27)
		Schut	9.98 (0.23)**	754.35 (132.26)	266.94 (42.55)**	0.99 (0.00)**	NA	0.03 (0.18)
		Wachter	374.37 (41.37)	871.09 (92.36)	536.24 (48.73)	0.00 (0.00)	NA 0.00 (0.18)**	1.00 (0.05)
	JEM	ECCCo ECCCo (no CP)	1.87 (0.79) 1.83 (0.80)	0.57 (0.58)** 0.63 (0.64)*	1.29 (0.21)* 1.30 (0.21)*	0.00 (0.00)	0.99 (0.18)** 1.13 (0.35)	1.00 (0.00) 1.00 (0.00)
		ECCCo (no EBM)	1.30 (1.72)	1.73 (1.34)	1.73 (1.42)	0.00 (0.00)	0.94 (0.27)*	1.00 (0.00)
		REVISE	1.07 (0.26)	1.59 (0.55)	1.55 (0.20)	0.00 (0.00)	1.30 (0.40)	1.00 (0.00)
		Schut Wachter	1.36 (0.35) 0.89 (0.21)	1.55 (0.61) 1.77 (0.48)	1.42 (0.16)* 1.67 (0.15)	0.03 (0.12) 0.00 (0.00)	1.11 (0.30)* 1.45 (0.47)	1.00 (0.00) 1.00 (0.00)
Moons	MLP	ECCCo	2.53 (1.24)	1.68 (1.74)	2.02 (0.86)	0.00 (0.00)	1.45 (0.47)	1.00 (0.00)
		ECCCo (no CP)	2.45 (1.36)	1.34 (1.66)	2.11 (0.88)	0.00 (0.00)	1.24 (0.41)	1.00 (0.00)
		ECCCo (no EBM) REVISE	2.53 (2.03) 0.98 (0.33)*	2.98 (1.89) 2.46 (1.05)	2.29 (1.75) 1.54 (0.27)*	0.00 (0.00) 0.00 (0.00)	0.99 (0.07)** 1.40 (0.49)	1.00 (0.00) 1.00 (0.00)
		Schut	0.75 (0.23)**	2.71 (1.15)	1.62 (0.42)	0.31 (0.27)*	0.94 (0.24)*	0.94 (0.24)
		Wachter	1.49 (1.76)	2.95 (1.42)	1.84 (1.33)	0.00 (0.00)	1.33 (0.48)	1.00 (0.00)

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