

1 The economy consists of a representative agent with preferences:

$$U(c) = E_0 \sum_{t=0}^{\infty} \beta^t u(c_t), \quad u(c) = \frac{c^{1-\alpha}}{1-\alpha}$$

There is no production. Aggregate endowment each period follows the stochastic process

$$y_{t+1} = \lambda_{t+1} y_t$$

where the growth rate  $\lambda$  takes on one of two value,  $\lambda_1$  or  $\lambda_2$ , with probabilities given by the first order Markovian transition matrix:

$$\Pi = \begin{bmatrix} (1+\rho)/2 & (1-\rho)/2 \\ (1-\rho)/2 & (1+\rho)/2 \end{bmatrix} = \begin{bmatrix} \phi & 1-\phi \\ 1-\phi & \phi \end{bmatrix}$$

Let  $\lambda_1 = \mu + \sigma$ , and  $\lambda_2 = \mu - \sigma$ . In equilibrium  $y_t = c_t$ .

Download Non-durable and service consumption data from WRDS. Generate real-per capita consumption growth for two sample (a) one that starts at 1929, (b) one that starts at 1950.

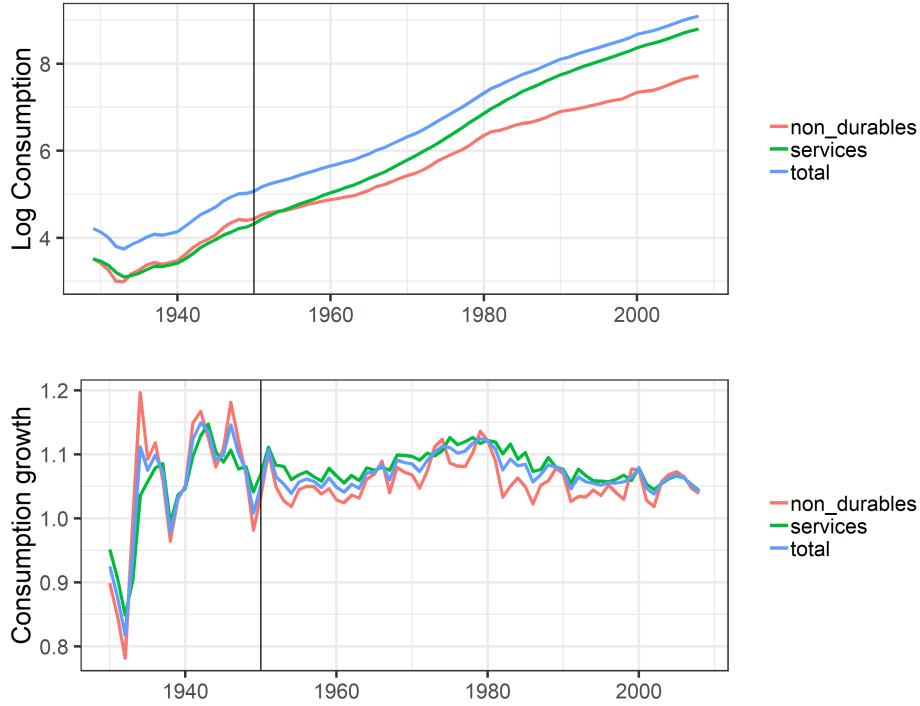


Figure 1: Log Consumption and Consumption Growth (Gross). Vertical line represents starting point of sample (b)

Mehra and Prescott find 0.018, 0.036, and  $-0.14$ , are the mean, standard deviation and autocorrelation of continuous mean consumption growth in their sample, respectively. We can then calculate the parameters of the Markov process as follows:

- **Persistence:** To compute the persistence of the process, consider the indicator function, denoted as  $\xi_t$ , which is equal to 1 in the high growth state and zero otherwise. Given our transition matrix  $\Pi$ , we then

have

$$\begin{bmatrix} \xi_t \\ 1 - \xi_t \end{bmatrix} = \begin{bmatrix} \phi & 1 - \phi \\ 1 - \phi & \phi \end{bmatrix} \begin{bmatrix} \xi_{t-1} \\ 1 - \xi_{t-1} \end{bmatrix} + \begin{bmatrix} v_{1,t} \\ v_{2,t} \end{bmatrix}$$

which implies  $\xi_{t+1} = (1 - \phi) + \xi_t(\phi - 1 + \phi) + v_{1,t+1}$ . Taking the unconditional expectation implies the persistence in the process is given by  $2\phi - 1 = 2(\frac{1}{2}(1 + \rho)) - 1 = \rho$ . Thus,  $2\phi - 1 = -0.014$ , and  $\phi = 0.43$ . Further note that  $E[\xi_{t+1}] = \frac{1-\phi}{2(1-\phi)}$ . Thus, the unconditional probability  $\pi_1 = \pi_2 = 0.5$ .

- **Mean:** Given that the mean gross growth rates in the high and low growth regimes are given by  $\mu + \sigma$  and  $\mu - \sigma$ , we know that the unconditional mean is given by  $\mu = 1.018$ .
- **Standard Deviation:** Once again using the unconditional probability derived above, we have that the unconditional variance of the process is given by  $\sigma^2 = 0.5\sigma^2 + 0.5\sigma^2$ , which implies  $\sigma = 0.036$ .

We can follow the same the process to calculate the parameters for sample (a) and sample (b).

#### a) Markov Chains

- Compute the conditional moments of the Markov chain which describes the evolution of the  $\lambda$  process.

Letting  $s_t = h/l$  denote whether we are currently in the high or low state, the conditional means are given by

$$\begin{aligned} E[\lambda_{t+1}|s_t = h] &= \pi_{hh}\lambda_h + \pi_{hl}\lambda_l \\ &= \phi(\mu + \sigma) + (1 - \phi)(\mu - \sigma) \\ &= \mu + \sigma(2\phi - 1) \\ E[\lambda_{t+1}|s_t = l] &= \pi_{lh}\lambda_h + \pi_{ll}\lambda_l \\ &= (1 - \phi)(\mu + \sigma) + \phi(\mu - \sigma) \\ &= \mu + \sigma(1 - 2\phi) \end{aligned}$$

The conditional variance is generally given by

$$\begin{aligned} Var(\lambda_{t+1}|s_t = i) &= E[\lambda_{t+1}^2|s_t = i] - E[\lambda_{t+1}|s_t = i]^2 \\ &= \sum_j \pi_{ij}\lambda_j^2 - \left( \sum_j \pi_{ij}\lambda_j \right)^2 \end{aligned}$$

Thus, we have

$$\begin{aligned} Var(\lambda_{t+1}|s_t = h) &= \pi_{hh}\lambda_h^2 + \pi_{hl}\lambda_l^2 - (\pi_{hh}\lambda_h + \pi_{hl}\lambda_l)^2 \\ &= \phi(\mu + \sigma)^2 + (1 - \phi)(\mu - \sigma)^2 - (\phi(\mu + \sigma) + (1 - \phi)(\mu - \sigma))^2 \\ &= 4\sigma^2\phi(1 - \phi) \\ Var(\lambda_{t+1}|s_t = l) &= \pi_{lh}\lambda_h^2 + \pi_{ll}\lambda_l^2 - (\pi_{lh}\lambda_h + \pi_{ll}\lambda_l)^2 \\ &= 4\sigma^2\phi(1 - \phi) \end{aligned}$$

Hence, we see that the conditional variances are equal to each other.

- Compute the stationary distribution  $\Pi^*$ , which satisfies  $\Pi^* = \Pi \times \Pi^*$ .

We know that  $\Pi^*\iota = 1$ , where  $\iota$  denotes the vector of ones. Using these two restrictions, we can solve for  $\Pi^*$ . Define the matrix  $A$  as

$$A = \begin{bmatrix} I_2 - \Pi \\ \iota' \end{bmatrix}$$

Then we have  $A\Pi^* = [0 \ 0 \ 1]'$ . The solution for  $\Pi^*$  is

$$\begin{aligned} A\Pi^* &= [0 \ 0 \ 1]' \\ A'A\Pi^* &= A'[0 \ 0 \ 1]' \\ \Pi^* &= (A'A)^{-1}A'[0 \ 0 \ 1]' \end{aligned}$$

Plugging in our parameters for  $\phi$ , we get

$$\Pi^* = \begin{bmatrix} 0.5 \\ 0.5 \end{bmatrix}$$

III. Confirm that the unconditional mean, standard deviation, and first order autocorrelation coefficient for the  $\lambda$  process are  $\mu$ ,  $\sigma$ , and  $\rho$  respectively.

We can calculate the unconditional mean and standard deviation using  $\Pi^*$

$$\begin{aligned} E[\lambda] &= \sum_i \pi_i^* \lambda_i \\ &= 0.5(\mu + \sigma) + 0.5(\mu - \sigma) \\ &= \mu \\ \text{Var}(\lambda) &= \sum_i \pi_i^* \lambda_i^2 - \left( \sum_i \pi_i^* \lambda_i \right)^2 \\ &= 0.5(\mu + \sigma)^2 + 0.5(\mu - \sigma)^2 - (0.5(\mu + \sigma) + 0.5(\mu - \sigma))^2 \\ &= 0.5((\mu + \sigma)^2 + (\mu - \sigma)^2) - \mu^2 \\ &= 0.5(2(\mu^2 + \sigma^2)) - \mu^2 \\ &= \sigma^2 \end{aligned}$$

Thus, we have verified that the unconditional mean and standard deviation are  $\mu$  and  $\sigma$ , respectively.