1 Factor Models

a) Use monthly data on the univariate size-sorted portfolios and book-to-market sorted decile portfolios to etimate alphas and betas, as well as compute the GRS test statistics to test the CAPM and Fama-French 3 factor model.

Solution: Fama and French consider regressions of the form

$$R_t^{ei} = \alpha_i + \beta_i f_t + \epsilon_{it}, \quad t = 1, 2, \dots T \forall i$$
 (1)

Where R^e denotes an excess return and f is a factor portfolio, which is an excess return. These time series regressions directly imply

$$E(R^{ei}) = \alpha_i + \beta_i E(f)$$

by taking averages. CAPM uses the excess return as the only factor. We can then implement Equation 1 for each portfolio. We get the estimates shown in Table 1, which show that β 's and returns go in the correct direction. However, the test of whether or not CAPM works is a test of whether or not the intercepts are jointly equal to zero.

Table 1: CAPM Test for Size and Book-to-Market Portfolios (Full Sample)

	Size estimates			BE/ME estimates			
	α	β	E(R)	α	β	E(R)	
d1	0.451	1.414	1.384	0.207	1.009	0.873	
d2	0.342	1.381	1.254	0.347	0.945	0.971	
d3	0.370	1.325	1.245	0.334	0.966	0.972	
d4	0.382	1.254	1.209	0.244	1.043	0.933	
d5	0.351	1.226	1.160	0.340	1.000	1.000	
d6	0.380	1.201	1.173	0.385	1.027	1.063	
d7	0.343	1.149	1.102	0.260	1.096	0.984	
d8	0.341	1.111	1.074	0.433	1.136	1.183	
d9	0.304	1.060	1.004	0.478	1.275	1.320	
d10	0.272	0.928	0.885	0.355	1.459	1.318	

The GRS test statistic allows us to test the hypothesis $\alpha_i = 0 \quad \forall i$. The GRS test is given by

$$\left(\frac{T}{N}\right) \left(\frac{T-N-L}{T-L-1}\right) \left[\frac{\hat{\alpha}'\hat{\Sigma}^{-1}\hat{\alpha}}{1+\bar{f}'\hat{\Omega}^{-1}\bar{f}}\right] \sim F(N,T-N-L)$$

where $\hat{\alpha}$ is a $N \times 1$ vector of estimated intercepts, $\hat{\Sigma}$ is an unbiased estimate of the residual covariance matrix, \bar{f} is a $L \times 1$ vector of the factor portfolios' covariance matrix. Hence in our case, the GRS statistic will be distributed as F(10, T - 10 - 1). The critical value for %95 confidence for this

distribution in 1.84. We get GRS test statistics of 35.77 and 72.36 for book to market and size, respectively. Hence, we can reject the null hypothesis that the α 's are jointly zero are the %95 confidence level.

Next we consider the Fama-French 3 factor model

$$R_t^{ei} = \alpha_i + b_i(R_{Mt} - R_{ft}) + s_i(SMB_t) + h_i(HML_t) + \epsilon_{it}$$

Table 2

Size estimates			BE/ME estimates			
α	β	E(R)	α	β	E(R)	
0.301	1.074	0.873	0.118	0.997	1.384	
0.397	0.975	0.971	0.112	1.065	1.254	
0.350	0.982	0.972	0.189	1.076	1.245	
0.206	1.027	0.933	0.228	1.038	1.209	
0.275	0.975	1.000	0.241	1.059	1.160	
0.277	0.973	1.063	0.281	1.071	1.173	
0.114	1.008	0.984	0.274	1.050	1.102	
0.247	1.012	1.183	0.291	1.048	1.074	
0.237	1.104	1.320	0.268	1.030	1.004	
0.024	1.190	1.318	0.298	0.974	0.885	
	0.301 0.397 0.350 0.206 0.275 0.277 0.114 0.247 0.237	0.301 1.074 0.397 0.975 0.350 0.982 0.206 1.027 0.275 0.975 0.277 0.973 0.114 1.008 0.247 1.012 0.237 1.104	0.301 1.074 0.873 0.397 0.975 0.971 0.350 0.982 0.972 0.206 1.027 0.933 0.275 0.975 1.000 0.277 0.973 1.063 0.114 1.008 0.984 0.247 1.012 1.183 0.237 1.104 1.320	0.301 1.074 0.873 0.118 0.397 0.975 0.971 0.112 0.350 0.982 0.972 0.189 0.206 1.027 0.933 0.228 0.275 0.975 1.000 0.241 0.277 0.973 1.063 0.281 0.114 1.008 0.984 0.274 0.247 1.012 1.183 0.291 0.237 1.104 1.320 0.268	0.301 1.074 0.873 0.118 0.997 0.397 0.975 0.971 0.112 1.065 0.350 0.982 0.972 0.189 1.076 0.206 1.027 0.933 0.228 1.038 0.275 0.975 1.000 0.241 1.059 0.277 0.973 1.063 0.281 1.071 0.114 1.008 0.984 0.274 1.050 0.247 1.012 1.183 0.291 1.048 0.237 1.104 1.320 0.268 1.030	

We can repeat this analysis for the past 10 years and past 25 years. The GRS test statistics and associated critical F-values are shown in 3 Hence, we strongly reject the model in all samples.

Table 3: GRS Statistics

	BE/ME	Size	Critical Value (%95)
CAPM (Full Sample)	35.771	72.363	1.839
FF3 (Full Sample)	35.792	85.591	1.839
CAPM (Last 10 Years)	2.747	2.176	1.839
FF3 (Last 10 Years)	280.745	209.916	1.839
CAPM (Last 25 Years)	8.520	599.408	1.839
FF3 (Last 25 Years)	138.727	292.573	1.839

b) Run Fama-MacBeth cross-sectional regressions of excess returns on the Fama-French factors as well as the two characteristics: Size and average Book/Market. Do the latter have significant slope coefficients? DO the results change if you omit the SMB and HML factors from the regression. Solution: We use the following procedure

1. Run the time series regressions: $R_t^{ei} = a_i + \beta_i' f_t + \epsilon_{it}$. In this case $f = [R^e, \text{SMB}, \text{HML}, \text{Size}, \text{BE/ME}]$.

- 2. Then we run the cross-section as a separate regression for each t: $R_t^{ei} = \beta_i' \lambda_t + \alpha_{it}$, where β from step 1 is the right hand side variable and λ (the price of risk) is the coefficient to be estimated. We get a different price of risk and error term for each t.
- 3. Estimate the price of risk as $\hat{\lambda}_t = E_T(\hat{\lambda}_t)$ and $\hat{\alpha} = E_T(\hat{\alpha}_T)$
- 4. Estimate the standard errors as $\sigma(\hat{\lambda}) = \frac{\sigma(\hat{\lambda}_t)}{\sqrt{T}}$.

Table 4: Fama-MacBeth Slope Coefficients

	RX	SMB	HML	ln(Size)	$\ln(\mathrm{BE/ME})$
lambda.size.	0.136	-0.113	-0.002	3.654	-0.364
t.lambda.	4.596	-1.730	-0.070	3.123	-2.571
lambda.BE.ME.	0.083	0.019	0.051	5.055	-0.369
t.lambda1	3.380	0.340	2.447	3.363	-2.025

Hence, we see that including that size and book to market ratios are significant. We can reestimate the parameters with a specification that leaves these out and get the following. Results generally do

Table 5: Fama-MacBeth Slope Coefficients

	RX	SMB	HML
lambda.size.	0.157	-0.131	-0.006
t.lambda.	4.885	-3.233	-0.304
lambda.BE.ME.	0.114	-0.019	0.047
t.lambda1	5.619	-0.448	2.346

not change when we omit the size and book to market factors.

c) Download (real) aggregate nondurable and services consumption from BEA or FRED, at monthly and quarterly frequencies. Test the standard consumption-based CAPM, using both Fama-MacBeth cross-sectional regressions and SDF-GMM methods, report estimates of cross-sectional factor risk premia (λ 's) and SDF loadings (b's) and their standard errors, as well as the relevant cross-sectional tests. Is the model rejected? What is the price of consumption risk that is implied by this cross-section of asset returns? Plot the actual mean excess returns against the predicted excess returns.

Solution: The CCAPM states

$$E\left[R^{ei}\right] = \gamma cov(\Delta c_{t+1}, R_{t+1}^{ei}) = \frac{cov(\Delta c_{t+1}, R_{t+1}^{ei})}{var(\Delta c_{t+1})} \left[\gamma var(\Delta)c_{t+1}\right] = \beta_{i,M} \lambda_{\Delta c}$$

It states that the expected return of asset i is proportional to beta times the market price of consumption risk. This representation implies we can run the following test of the model

$$R_{t+1}^{ei} = a_i + b_{i,\Delta c} \Delta c_{t+1} + \epsilon_{t+1}^i \quad t = 1, ..., T$$
$$E(R^{ei}) = b_{i,\Delta c} \lambda_{\Delta c} + \alpha_i$$

In the first stage the point is to estimate the tendency for the asset to fall when consumption falls (i.e. estimate b). The second stage is the test of the statement of the model: α 's should be zero and we should see expected return to be high where β 's are high.

Using GMM, we use the following procedure to compute the SDF loadings

$$0 = E(MR^e)$$

$$M = 1 - b'(f - Ef) = 1 - b'\tilde{f}$$

$$g_T = E_T \left[R^e (1 - b\tilde{f}) \right]$$

$$d = \frac{\partial g_t}{\partial b} - E(R^e \tilde{f}') = -cov(R^e, f') = -c$$

$$\min g'_t W g_t \implies \text{FOC: } d'W g_t = 0$$

$$0 = d'W E_T (R^e (1 - \tilde{f}'b)) \implies 0 = -c'W E_T (R^e (1 - \tilde{f}'b))$$

$$c'W E_T (R^e) = c'W cb$$

$$b = \left[c'W c \right]^{-1} c'W E_T (R^e)$$

In the case of CCAPM the factor, f, is consumption growth, Δc_{t+1} , so we see that GMM is just a cross sectional regression of expected returns on covariances. That is, GMM is equivalent to considering the cross-sectional model $E(R^e) = cov(R^e, \Delta c_{t+1})b$, where we can just run a regression to estimate b.