

## 8.4 Moving average models

Rather than using past values of the forecast variable in a regression, a moving average model uses past forecast errors in a regression-like model.

$$y_t = c + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \cdots + \theta_q \varepsilon_{t-q},$$

where  $\varepsilon_t$  is white noise. We refer to this as an **MA( $q$ ) model**, a moving average model of order  $q$ . Of course, we do not *observe* the values of  $\varepsilon_t$ , so it is not really a regression in the usual sense.

Notice that each value of  $y_t$  can be thought of as a weighted moving average of the past few forecast errors. However, moving average *models* should not be confused with the moving average *smoothing* we discussed in Chapter 6. A moving average model is used for forecasting future values, while moving average smoothing is used for estimating the trend-cycle of past values.

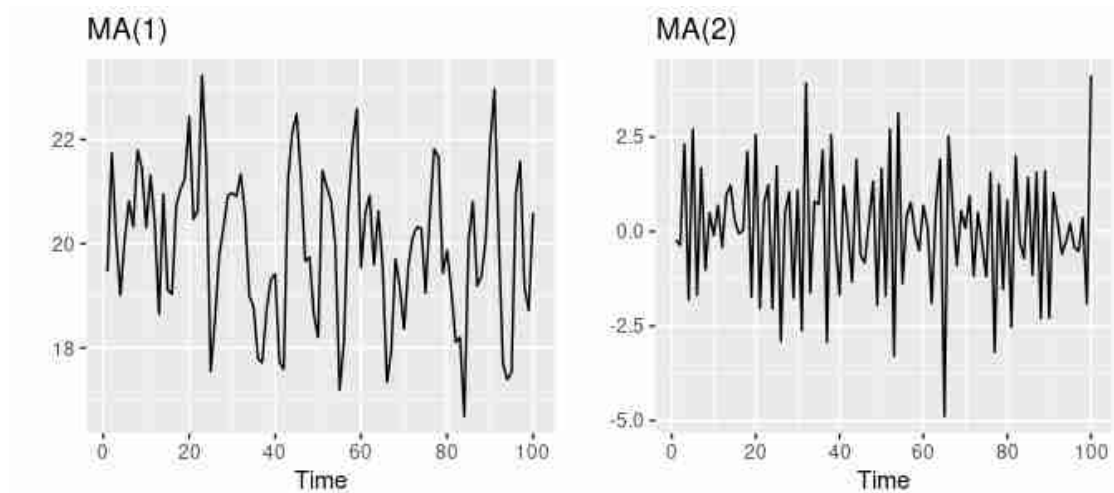


Figure 8.6: Two examples of data from moving average models with different parameters. Left: MA(1) with  $y_t = 20 + \varepsilon_t + 0.8\varepsilon_{t-1}$ . Right: MA(2) with  $y_t = \varepsilon_t - \varepsilon_{t-1} + 0.8\varepsilon_{t-2}$ . In both cases,  $\varepsilon_t$  is normally distributed white noise with mean zero and variance one.

Figure 8.6 shows some data from an MA(1) model and an MA(2) model. Changing the parameters  $\theta_1, \dots, \theta_q$  results in different time series patterns. As with autoregressive models, the variance of the error term  $\varepsilon_t$  will only change the scale of the series, not the patterns.

It is possible to write any stationary AR( $p$ ) model as an MA( $\infty$ ) model. For example, using repeated substitution, we can demonstrate this for an AR(1) model:

$$\begin{aligned} y_t &= \phi_1 y_{t-1} + \varepsilon_t \\ &= \phi_1 (\phi_1 y_{t-2} + \varepsilon_{t-1}) + \varepsilon_t \\ &= \phi_1^2 y_{t-2} + \phi_1 \varepsilon_{t-1} + \varepsilon_t \\ &= \phi_1^3 y_{t-3} + \phi_1^2 \varepsilon_{t-2} + \phi_1 \varepsilon_{t-1} + \varepsilon_t \\ &\text{etc.} \end{aligned}$$

Provided  $-1 < \phi_1 < 1$ , the value of  $\phi_1^k$  will get smaller as  $k$  gets larger. So eventually we obtain

$$y_t = \varepsilon_t + \phi_1 \varepsilon_{t-1} + \phi_1^2 \varepsilon_{t-2} + \phi_1^3 \varepsilon_{t-3} + \cdots,$$

an MA( $\infty$ ) process.

The reverse result holds if we impose some constraints on the MA parameters. Then the MA model is called **invertible**. That is, we can write any invertible MA( $q$ ) process as an AR( $\infty$ ) process. Invertible models are not simply introduced to enable us to convert from MA models to AR models. They also have some desirable mathematical properties.

For example, consider the MA(1) process,  $y_t = \varepsilon_t + \theta_1 \varepsilon_{t-1}$ . In its AR( $\infty$ ) representation, the most recent error can be written as a linear function of current and past observations:

$$\varepsilon_t = \sum_{j=0}^{\infty} (-\theta)^j y_{t-j}.$$

When  $|\theta| > 1$ , the weights increase as lags increase, so the more distant the observations the greater their influence on the current error. When  $|\theta| = 1$ , the weights are constant in size, and the distant observations have the same influence as the recent observations. As neither of these situations make much sense, we require  $|\theta| < 1$ , so the most recent observations have higher weight than observations from the more distant past. Thus, the process is invertible when  $|\theta| < 1$ .

The invertibility constraints for other models are similar to the stationarity constraints.

- For an MA(1) model:  $-1 < \theta_1 < 1$ .
- For an MA(2) model:  $-1 < \theta_2 < 1$ ,  $\theta_2 + \theta_1 > -1$ ,  $\theta_1 - \theta_2 < 1$ .

More complicated conditions hold for  $q \geq 3$ . Again, R will take care of these constraints when estimating the models.