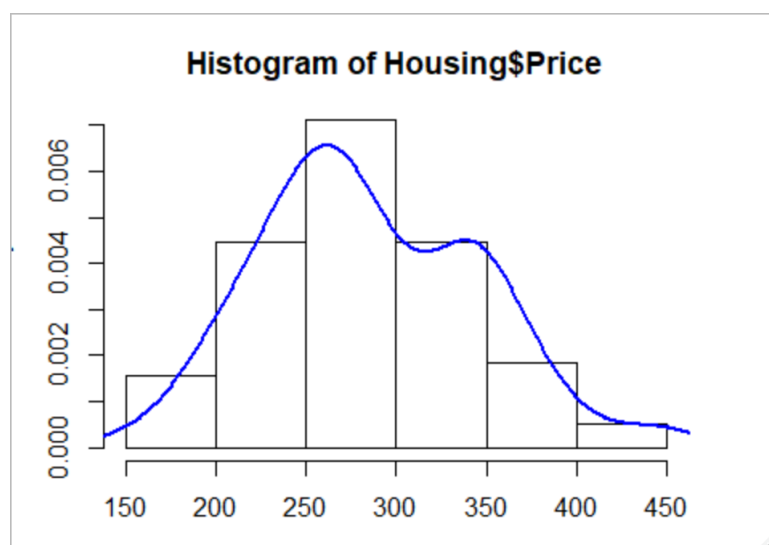
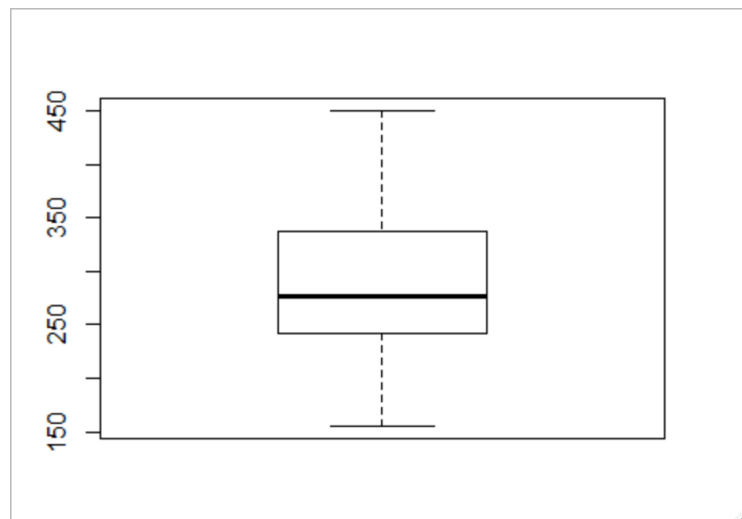

Exploratory Data Analysis:

In order to import correctly I had to use R Studio to remove 'i»ç' from the beginning of the 'Price' column. R Code has also been included at the end as a safety measure

Q1. USING A BOXPLOT, HISTOGRAM AND SUMMARY. DESCRIBE THE DISTRIBUTION OF THE SALES PRICE OF THE HOUSES.

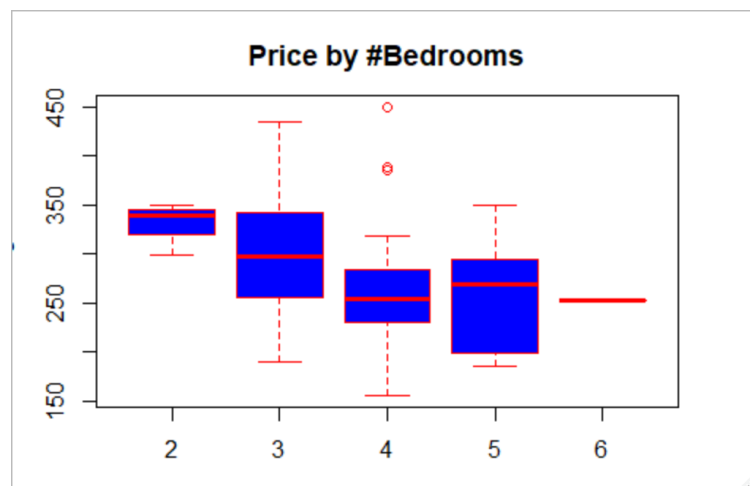


Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
155.5	242.8	276.0	285.8	336.8	450.0

The range of house price is \$155,500 to \$450,000. 50% of house prices fall between \$242,800 and \$336,800. None of the data appears to be above or below the max/min, as in there are no properties above the \$450K mark ($1.5 * \text{IQR}$), i.e no outliers. From the histogram there appears to be a lot of houses just above the \$250K dipping at \$300K and rising slightly again at ~\$350K.

Q2. CONVERT ALL THE CATEGORICAL VARIABLES TO FACTORS. USING THE SUMMARY AND A BOXPLOT DESCRIBE HOW SALES PRICES VARY WITH RESPECT TO THE NUMBER OF BEDROOMS, BATHROOMS, GARAGE SIZE AND SCHOOL.

Since other attributes are already in a numeric representation (i.e. lot size) I have only converted the categorical data School into factors. For comparison I used boxplots grouped by each secondary category.



\$`2`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 299.0 319.4 339.9 329.6 344.9 350.0

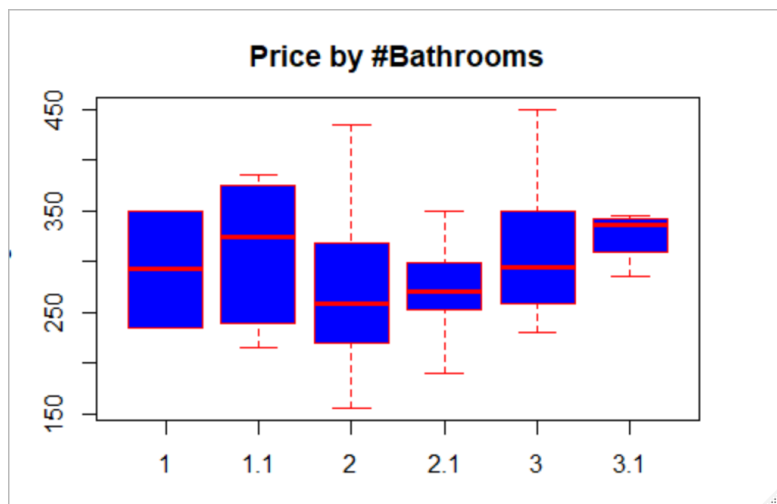
\$`3`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 189.5 256.2 297.0 297.3 342.5 435.0

\$`4`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 155.5 231.5 254.4 266.6 283.5 450.0

\$`5`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 185.0 199.0 269.0 259.5 295.0 349.5

\$`6`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 252.5 252.5 252.5 252.5 252.5 252.5

By number of bedrooms, it would appear that price dips until 4 bedrooms and recovers just as quickly. 2 Bedrooms appears to be the highest on average while 4 bedrooms (one of the cheapest on average) has the highest priced home, an outlier. There only appears to be 1 house with 6 bedrooms, this appears to fall in closely with the median for 4 and 5 bedroom homes.



\$`1`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 235.0 263.8 292.5 292.5 321.2 350.0

\$`1.1`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 215.0 239.5 325.0 307.9 374.5 385.5

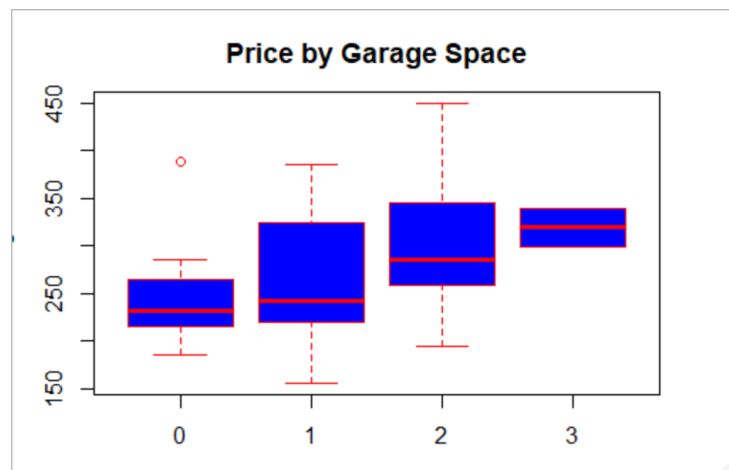
\$`2`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 155.5 220.0 259.0 270.7 319.0 435.0

\$`2.1`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 189.5 254.8 269.9 274.5 297.7 349.5

\$`3`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 230.0 259.0 295.0 307.8 349.5 450.0

\$`3.1`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 285.0 309.4 336.0 324.2 342.5 345.0

Surprisingly homes that had a half-bathroom extra tended to increase the price on average, contradicting the belief that half-bathrooms are not as valued as full bathrooms. The more pricier houses tend to have 1 and a half bathrooms, or 3 bathrooms while 2/2.5 bathroom houses tend to cost less.



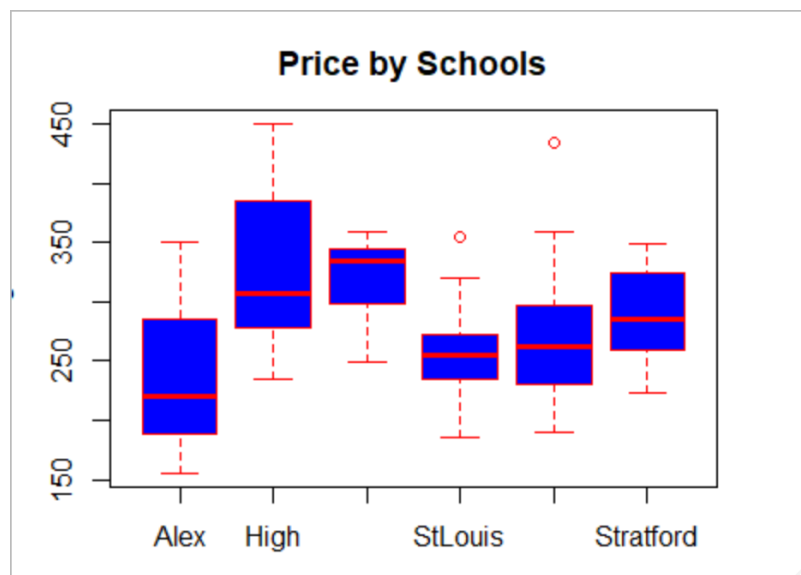
\$`0`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 185.0 216.0 232.0 246.9 264.4 388.0

\$`1`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 155.5 220.0 242.0 260.6 324.5 385.5

\$`2`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 195.0 259.0 285.0 299.6 343.8 450.0

\$`3`
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 299.0 309.2 319.4 319.4 329.7 339.9

Not surprisingly, the number of car spaces increases the house price. Oddly, there is an outlier with 0 car spaces that is unusually high. Having 1-2 spaces seems the most common, with a similar range but a higher price.



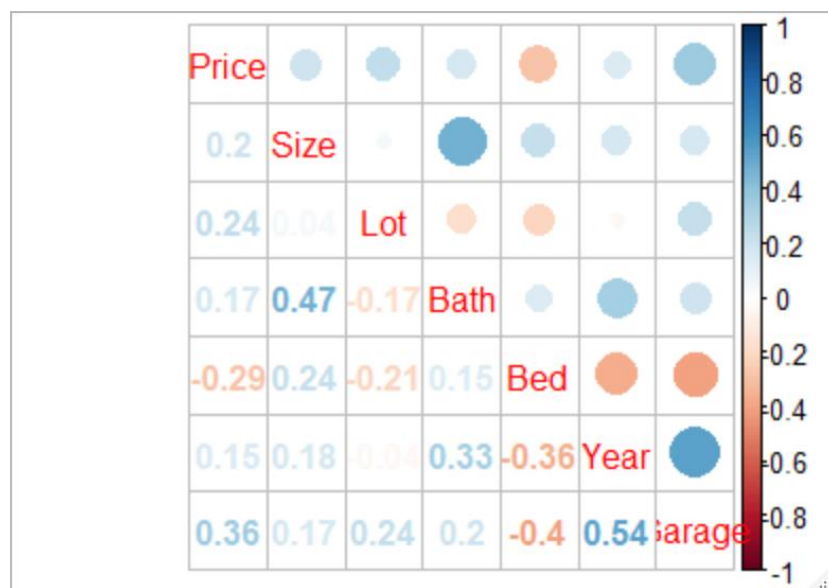
\$Alex
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 155.5 187.8 220.0 241.8 285.0 350.0

\$High
 Min. 1st Qu. Median Mean 3rd Qu. Max.
 235.0 279.2 307.5 327.1 385.6 450.0

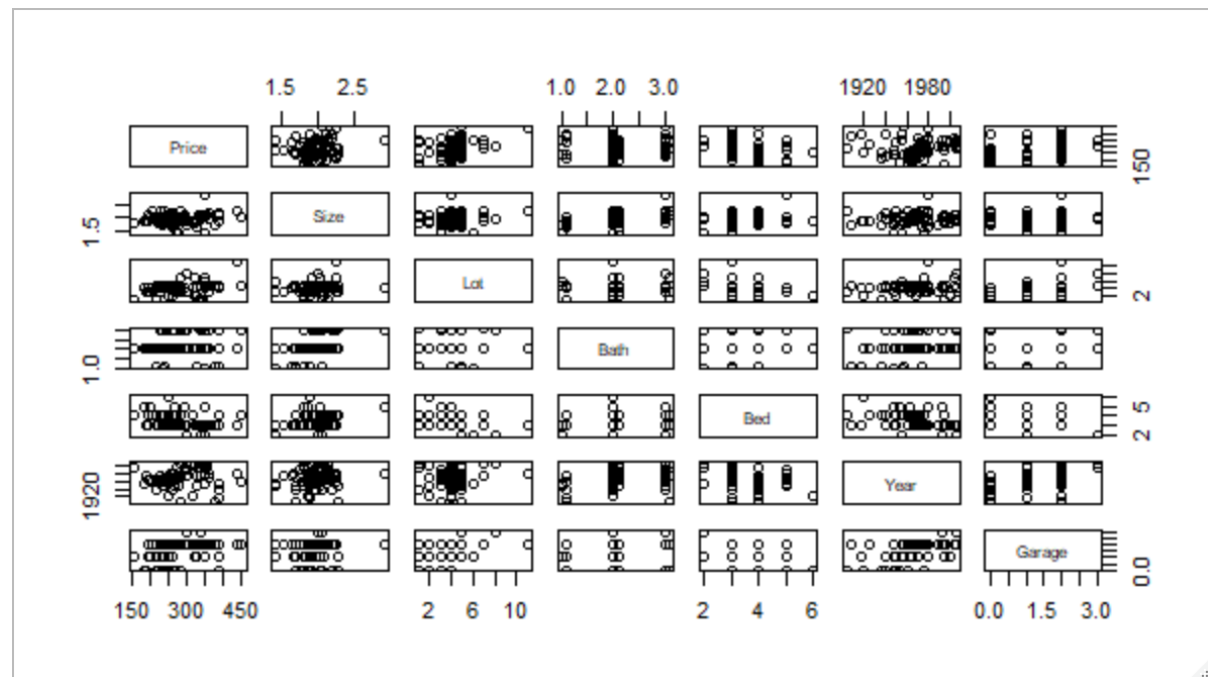
\$NotreDame					
Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
249.9	304.0	334.9	319.1	345.0	359.9
\$StLouis					
Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
185.0	235.4	255.0	257.4	272.4	355.0
\$StMarys					
Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
189.5	231.6	262.0	269.8	296.5	435.0
\$Stratford					
Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
222.5	266.2	285.0	287.8	315.0	349.5

Comparing the price of homes to schools, it would appear that 2 (High) has the largest impact on school price. While 5 (Stratfor) does has a highly priced outlier, as does 4 (St. Marys), their median is not as high. However, 3 (Notre Dame) does have the largest median price increase.

Q3. USING THE SUMMARY, CORRELATION AND THE PAIRS PLOTS DISCUSS THE RELATIONSHIP BETWEEN THE RESPONSE SALES PRICE AND EACH OF THE NUMERIC PREDICTOR VARIABLES.



From the correlation matrix we can see that price and garage are the most positively correlated while bed and price are the most negatively correlated. Meaning that as the number of beds increase, the price is more likely to decrease.



Regression Model

Q1. FIT A MULTIPLE LINEAR REGRESSION MODEL TO THE DATA WITH SALES PRICE AS THE RESPONSE AND SIZE, LOT, BATH, BED, YEAR, GARAGE AND SCHOOL AS THE PREDICTOR VARIABLES. WRITE DOWN THE EQUATION FOR THIS MODEL.

Call:

```
lm(formula = Housing$Price ~ Housing$Size + Housing$Lot.Type.f +
    Housing$Bath.Type.f + Housing$Bed + Housing$Year + Housing$Garage +
    Housing$School.Type.f, data = Housing)
```

Residuals:

Min	1Q	Median	3Q	Max
-83.626	-18.966	1.722	21.676	70.213

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	-855.6150	677.4376	-1.263	0.212111	
Housing\$Size	41.9278	28.3780	1.477	0.145467	
Housing\$Lot.Type.f2	24.8812	37.6375	0.661	0.511428	
Housing\$Lot.Type.f3	5.0597	31.9228	0.158	0.874666	
Housing\$Lot.Type.f4	9.3078	30.9623	0.301	0.764882	
Housing\$Lot.Type.f5	35.4518	32.2761	1.098	0.276999	
Housing\$Lot.Type.f6	279.1165	68.8729	4.053	0.000167	***
Housing\$Lot.Type.f7	38.3898	33.9690	1.130	0.263506	
Housing\$Lot.Type.f8	-27.4835	55.0113	-0.500	0.619426	
Housing\$Lot.Type.f11	174.5166	51.9887	3.357	0.001465	**
Housing\$Bath.Type.f1.1	142.8145	46.8846	3.046	0.003608	**
Housing\$Bath.Type.f2	92.7954	43.6970	2.124	0.038389	*
Housing\$Bath.Type.f2.1	94.9218	44.9945	2.110	0.039630	*
Housing\$Bath.Type.f3	135.6652	46.0876	2.944	0.004806	**
Housing\$Bath.Type.f3.1	107.5644	51.0539	2.107	0.039878	*
Housing\$Bed	-11.3629	8.8708	-1.281	0.205794	
Housing\$Year	0.4473	0.3369	1.328	0.189962	

```

Housing$Garage          13.8429      8.4199      1.644 0.106084
Housing$School.Type.fHigh 132.2855     36.0237     3.672 0.000560 ***
Housing$School.Type.fNotreDame 99.8114     34.2670     2.913 0.005234 **
Housing$School.Type.fStLouis 47.1872     34.3808     1.372 0.175692
Housing$School.Type.fStMarys 50.0332     34.0152     1.471 0.147229
Housing$School.Type.fStratford 72.2833     38.7528     1.865 0.067688 .
---

```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 39.94 on 53 degrees of freedom
 Multiple R-squared: 0.6903, Adjusted R-squared: 0.5617
 F-statistic: 5.369 on 22 and 53 DF, p-value: 2.82e-07

The R code is:

```
lm(Price ~ Size + Lot + Bath + Bed + Year + Garage + School, data=Housing)
```

The maths formula would be:

$$Y_i = \beta_0 + \beta_1 X_{i,1} + \dots + \beta_{p-1} X_{i,p-1}$$

Alternatively:

$$\text{Price} = -855.62 + \beta_1(\text{Size}) + \beta_2(\text{Lot}) + \beta_3(\text{Bath}) + \beta_4(\text{Bed}) + \beta_5(\text{Year}) + \beta_6(\text{Garage}) + \beta_7(\text{School})$$

Q2. INTERPRET THE ESTIMATE OF THE INTERCEPT TERM B0.

The estimated average house price is \$-855,615 when Size, Lot, Bath, Bed, Year, Garage, and school are all at 0. However the P-Value is large (> 0.05) but since it is the intercept it could be ignored. For variables they should be considered if removal is better as it affects the models performance. A house with none of the attributes listed would not get a sale price, you have nothing to sell and therefore no meaning should be attached to this intercept value of the constant.

Q3. INTERPRET THE ESTIMATE OF BSIZE THE PARAMETER ASSOCIATED WITH FLOOR SIZE (SIZE).

A one unit increase to Size has an increase of \$41,927 (since cost is divided by 100K) to the cost of a house. Since the P-Value for Size is $0.178 > 0.05$ we fail to reject the null hypothesis. That is, we do not have enough evidence to support that Size significantly affects the Price of a house.

Q4. INTERPRET THE ESTIMATE OF BBATH1.1 THE PARAMETER ASSOCIATED WITH ONE AND A HALF BATHROOMS.

A one unit increase in Bath1.1 would result in \$142,815 increase to the property price compared to a 1 bathroom house. Since this is a categorical feature, we have to remember that the first entry is taken as reference i.e. Bathroom 1 is the reference and all other bathroom sizes are compared to it, meaning that value added is in comparison to a 1 Bathroom and not 0 bathrooms. As the P value is $0.003605 < 0.05$ it is considered significant. In other words, we have enough evidence to support that Bath1.1 significantly affects the price of a house and can reject the null hypothesis.

Q5. DISCUSS AND INTERPRET THE EFFECT THE PREDICTOR VARIABLE BED ON THE EXPECTED VALUE OF THE HOUSE PRICES.

- From an adjustment made after submission –

It appears that 2 bedroom houses add the most value to the property, for additional beds it appears to reduce the cost, 6 bedrooms has ~double the loss but this may not be accurate market analysis as there is only 1 entry with 6 bedrooms.

Housing\$Bed.Type.f3	-36.7818	44.1971	-0.832	0.409242
Housing\$Bed.Type.f4	-49.5788	45.6583	-1.086	0.282746
Housing\$Bed.Type.f5	-46.7850	49.9368	-0.937	0.353322
Housing\$Bed.Type.f6	-90.6022	68.9303	-1.314	0.194708

Q6. LIST THE PREDICTOR VARIABLES THAT ARE SIGNIFICANTLY CONTRIBUTING TO THE EXPECTED VALUE OF THE HOUSE PRICES

Variable	Estimate	Std. Error	t value	Pr(> t)	
Lot 6	279.1165	68.8729	4.053	0.000167	***
Lot 11	174.5166	51.9887	3.357	0.001465	**
Bath 1.1	142.8145	46.8846	3.046	0.003608	**
Bath 2	92.7954	43.6970	2.124	0.038389	*
Bath 2.1	94.9218	44.9945	2.110	0.039630	*
Bath 3	135.6652	46.0876	2.944	0.004806	**
Bath 3.1	107.5644	51.0539	2.107	0.039878	*
School High	132.2855	36.0237	3.672	0.000560	***
School NotreDame	99.8114	34.2670	2.913	0.005234	**
School Stratford	72.2833	38.7528	1.865	0.067688	.

These variables are the ones that contribute the most to the value of the house prices. Lot 6 contributes the most to expected value of the house prices. Lot 11 also contributes a lot, but considering there is only 1 data entry and it is priced above average its representation could be skewed. Bathroom 1.1 contributes the most from the bathroom category, with bathroom 3 being next. Schools are the next biggest contributors, Stratford is above the 0.05 value, but depending on how strict we need to be with our model I have decided to include it.

Q7. FOR EACH PREDICTOR VARIABLE WHAT IS THE VALUE THAT WILL LEAD TO THE LARGEST EXPECTED VALUE OF THE HOUSE PRICES.

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	-855.6150	677.4376	-1.263	0.212111	
Size	41.9278	28.3780	1.477	0.145467	
Lot 6	279.1165	68.8729	4.053	0.000167	***
Bath 1.1	142.8145	46.8846	3.046	0.003608	**
Bed	-11.3629	8.8708	-1.281	0.205794	
Year	0.4473	0.3369	1.328	0.189962	
Garage	13.8429	8.4199	1.644	0.106084	
School High	132.2855	36.0237	3.672	0.000560	***

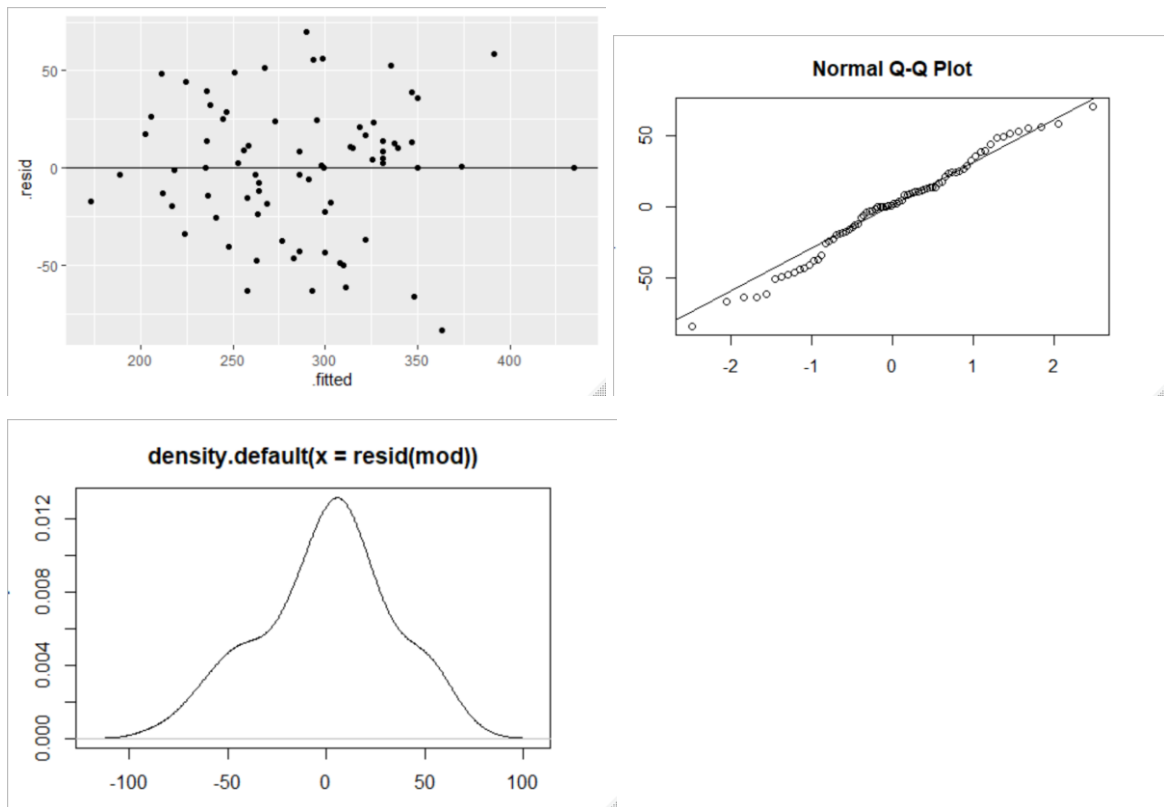
For Size, a one unit increase would increase house price by \$41,928 (in the range of [1.44, 2.896]). Lot of category 6 increases the house price by \$279,117 in comparison to lot 1. Then, a 1.5 Bathrooms would add \$142,815 in price in comparison to having only 1 bathroom. Having a 3 bedroom house would reduce the price of a house by -\$66,276 in comparison with a 2 bedroom house, therefore a 2 bedroom house would be better. For the year, the more modern the better the increase (\$447 per year from 1905 to 2005). Having a garage with 2 car spaces would add \$18,851 to your house in comparison to having no garage/car space. Finally, owning a house near the school the High School would increase its value by \$132,286 compared to living next to Alexandra.

Q8. FOR EACH PREDICTOR VARIABLE WHAT IS THE VALUE THAT WILL LEAD TO THE LOWEST EXPECTED VALUE OF THE HOUSE PRICES.

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-855.6150	677.4376	-1.263	0.212111
Size	41.9278	28.3780	1.477	0.145467
Lot 8	-27.4835	55.0113	-0.500	0.619426
Bath 1				
Bed	-11.3629	8.8708	-1.281	0.205794
Year	0.4473	0.3369	1.328	0.189962
Garage	13.8429	8.4199	1.644	0.106084
School Alexandra				

For the lowest priced house, it would require a relatively small Size, the Lot category should be in 8 as it would reduce the house price by -\$27,484 compared to Lot 1. Bathrooms should only have 1 as it will be the lowest price (any other size increases the house price). The number of bedrooms should be 6 as this reduces the house price the most in comparison to having only 2 beds. The year should be as old as possible, 1905. Any year after that adds a value of \$447 to the house. Having no car space keeps the house price at its lowest, in comparison to having even just 1 car space. And finally, being located near Alexandra school would have the lowest priced houses than any other school.

9. BY LOOKING AT THE INFORMATION ABOUT THE RESIDUALS IN THE SUMMARY AND BY PLOTTING THE RESIDUALS DO YOU THINK THIS IS A GOOD MODEL OF THE EXPECTED VALUE OF THE HOUSE PRICES.



Since the graph of residuals on the left shows homoskedasticity, i.e. the variance from 0 is of ~equal distance in the plus and minus, and the left graph has little variance from the line shows that the model is good. The 3rd graph shows density around the 0 mark which would such a normal distribution.

10. INTERPRET THE ADJUSTED R-SQUARED VALUE

Multiple R-squared: 0.6903, Adjusted R-squared: 0.5617

The R-squared value explains the variance in Y by the model, 0 would mean that none of the variance in Y is explained and 1 would mean all the variance is explained by the model. The adjusted R-square takes into account the number of variables used and is penalised for each one added. While the R-Squared value is at a decent level, 69.03% the gap adjusted on the adjusted r-squared of 56.17% would suggest that our model has some random/unnecessary variables included in it. The Adjusted R-Squared value suggests that 56.17% of variance in Y is explained by the model.

11. INTERPRET THE F-STATISTIC IN THE OUTPUT IN THE SUMMARY OF THE REGRESSION MODEL. HINT: STATE THE HYPOTHESIS BEING TESTED, THE TEST STATISTIC AND P-VALUE AND THE CONCLUSION IN THE CONTEXT OF THE PROBLEM.

F-statistic: 5.369 on 22 and 53 DF, p-value: 2.82e-07

The Hypothesis is: $H_0: \beta_1 = 0$ vs. $H_a: \beta_1 \neq 0$

The test statistic is: 5.39

The P=Value is: 0.000000282 (2.82e-07)

The test statistic is 5.39 and the probability of getting that in an F-distribution consisting of 22 and 53 df (degrees of freedom) is 2.82e-07, since it is less than 0.05 we reject the null hypothesis and conclude that at least one of the variables is non-zero.

ANOVA

Q1 COMPUTE THE TYPE 1 ANOVA TABLE. INTERPRET THE OUTPUT. HINT: STATE THE HYPOTHESIS BEING TESTED, THE TEST STATISTIC AND P-VALUE AND THE CONCLUSION IN THE CONTEXT OF THE PROBLEM.

Analysis of Variance Table
Response: Housing\$Price

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
Housing\$Size	1	11078	11077.7	6.9438	0.0110053	*
Housing\$Lot.Type.f	8	45378	5672.2	3.5555	0.0022973	**
Housing\$Bath.Type.f	5	41999	8399.8	5.2652	0.0005348	***
Housing\$Bed	1	21831	21830.9	13.6842	0.0005150	***
Housing\$Year	1	2517	2516.6	1.5775	0.2146302	
Housing\$Garage	1	6500	6500.2	4.0745	0.0486087	*
Housing\$School.Type.f	5	59147	11829.4	7.4150	2.47e-05	***
Residuals	53	84553	1595.3			

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

From the ANOVA table, we can see the variation in Y explained by the addition of each previous variable. For example, 11,078 of variation in Y is explained by Size when there are no other variables. However, 6,500 of variation in Y is explained by Garage given that Year, Bed, Bath, Lot, and Size are already in the model. Then, 84,553 of the variation in Y is NOT explained by its relationship with all the variables included. The Type I ANOVA table suggests removing 'Year' as it is the only variable that has the least significance.

The Hypothesis is: $H_0: \beta_1 = \beta_2 = \beta_3 = \beta_4 = \beta_5 = \beta_6 = \beta_7 = 0$ vs. H_a : at least one $\beta_k \neq 0$

Given that $F_{0.05,22,53} = 0.5256153$

With the F-test:

$$SSR = 11078 + 45378 + 41999 + 21831 + 2517 + 6500 + 59147 = 188449.41$$

$$MSR = (SSR/7) = 26921.42$$

$$SSE = 84553.06$$

$$MSE = SSE/53 = 1595.34$$

$$F = 16.87498$$

Since $16.875 > 0.525$ we reject H_0 . Which indicates that at least one of the variables are significantly related to Y.

Q2. WHICH PREDICTOR VARIABLE DOES THE TYPE 1 ANOVA TABLE SUGGEST YOU SHOULD REMOVE THE REGRESSION ANALYSIS.

Analysis of Variance Table

Response: Price

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Year	1	2517	2516.6	1.5775	0.2146302

Based on the ANOVA type 1 table, the variable 'Year' should be removed as it has a p-value greater than 0.05 and is not significant.

Q3. COMPUTE A TYPE 2 ANOVA TABLE COMPARING THE FULL MODEL WITH ALL PREDICTOR VARIABLES TO THE REDUCED MODEL WITH THE SUGGESTED PREDICTOR VARIABLE IDENTIFIED IN THE PREVIOUS QUESTION REMOVED. HINT: STATE THE HYPOTHESIS BEING TESTED, THE TEST STATISTIC AND P-VALUE AND THE CONCLUSION IN THE CONTEXT OF THE PROBLEM

Anova Table (Type II tests) **FOR REDUCED MODEL**

Response: Price

	Sum Sq	Df	F value	Pr(>F)
Size	3581	1	2.2131	0.1426584
Lot	54807	8	4.2345	0.0005401 ***
Bath	26817	5	3.3151	0.0110497 *
Bed	4541	1	2.8068	0.0996474 .
Garage	9835	1	6.0787	0.0168910 *
School	56397	5	6.9718	4.35e-05 ***
Residuals	87365	54		

Anova Table (Type II tests) **FOR NORMAL MODEL**

	Sum Sq	Df	F value	Pr(>F)
Size	3483	1	2.1829	0.1454666
Lot	56661	8	4.4395	0.0003676 ***
Bath	29071	5	3.6445	0.0065922 **

Bed	2618	1	1.6408	0.2057936
Year	2812	1	1.7628	0.1899620
Garage	4312	1	2.7029	0.1060842
School	59147	5	7.4150	2.47e-05 ***
Residuals	84553	53		

Hypothesis $H_0: \beta_k = 0$ vs. $H_a: \beta_k \neq 0$

$F^* = 2.655$

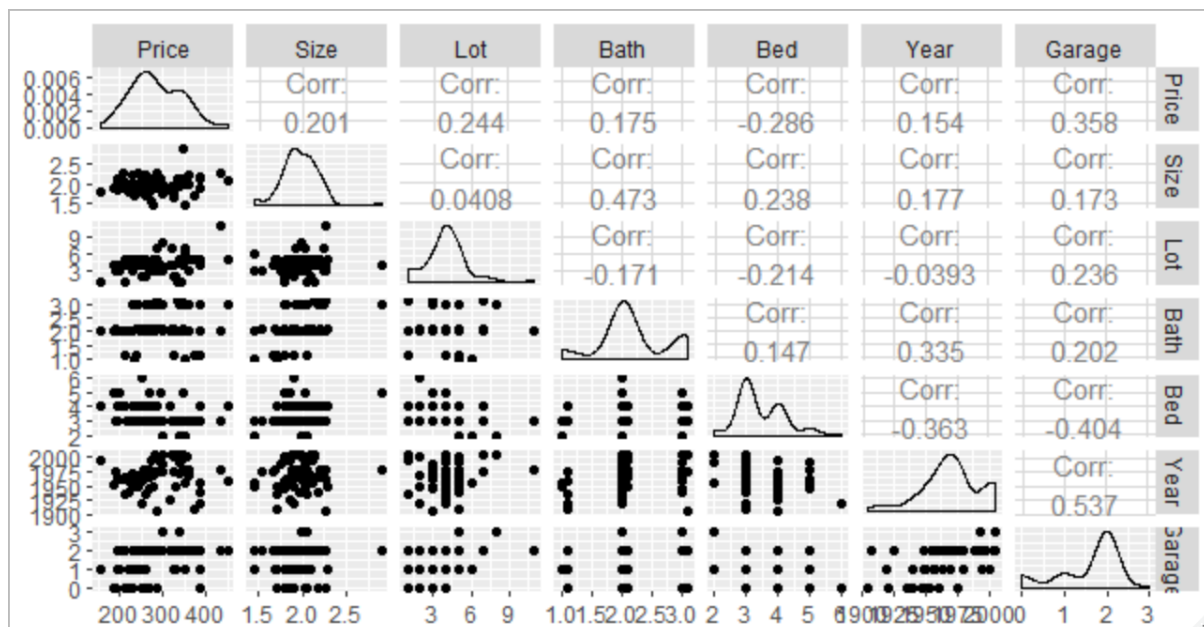
$F_{0.05, 21, 54} = 0.5188$

Since $2.655 > 0.5188$ we can reject H_0 i.e. that keeping the variable 'Year' does not significantly improve the model and can be removed. This can also be achieved using `anova(mod, red_mod)` as it will perform an F test comparison of them both. This gave me a p-value of 0.19 and since $0.19 > 0.05$ it is recommended to reject the second model).

Diagnostics

Since the 'Year' variable was not significant, and by recommendation of the attached video I have removed it and will be using 'red_mod' (model without 'Year') going forward.

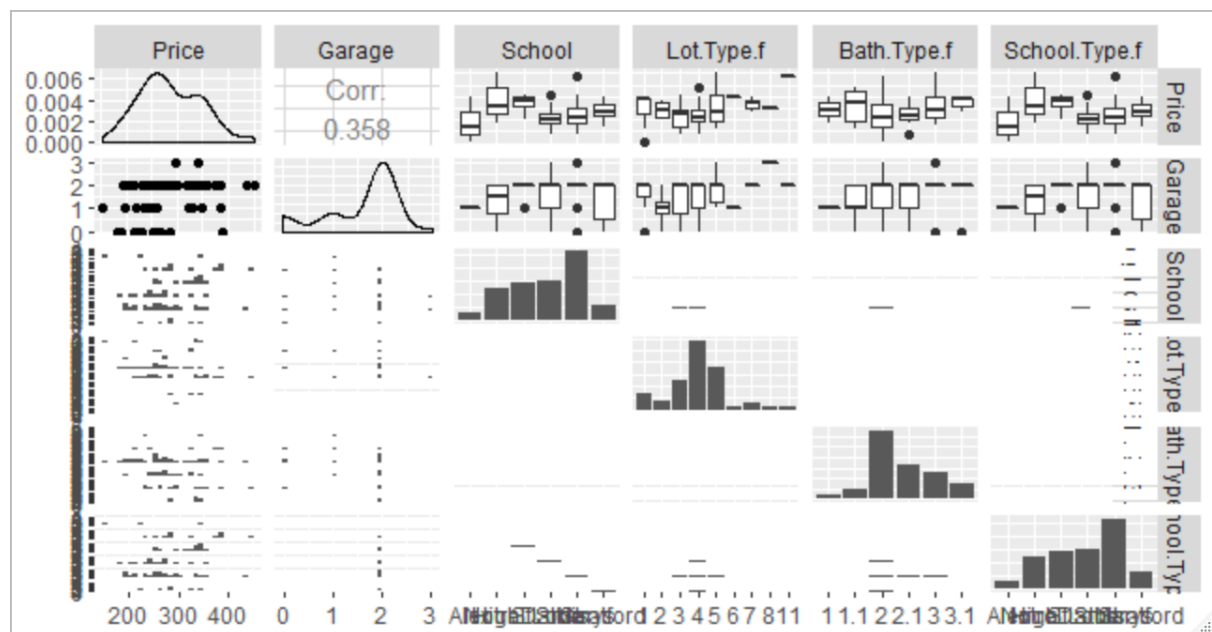
Q1 CHECK THE LINEARITY ASSUMPTION BY INTERPRETING THE ADDED VARIABLE PLOTS AND COMPONENT-PLUS-RESIDUAL PLOTS. WHAT EFFECT WOULD NON-LINEARITY HAVE ON THE REGRESSION MODEL AND HOW MIGHT YOU CORRECT OR IMPROVE THE MODEL IN THE PRESENCE OF NON-LINEARITY?



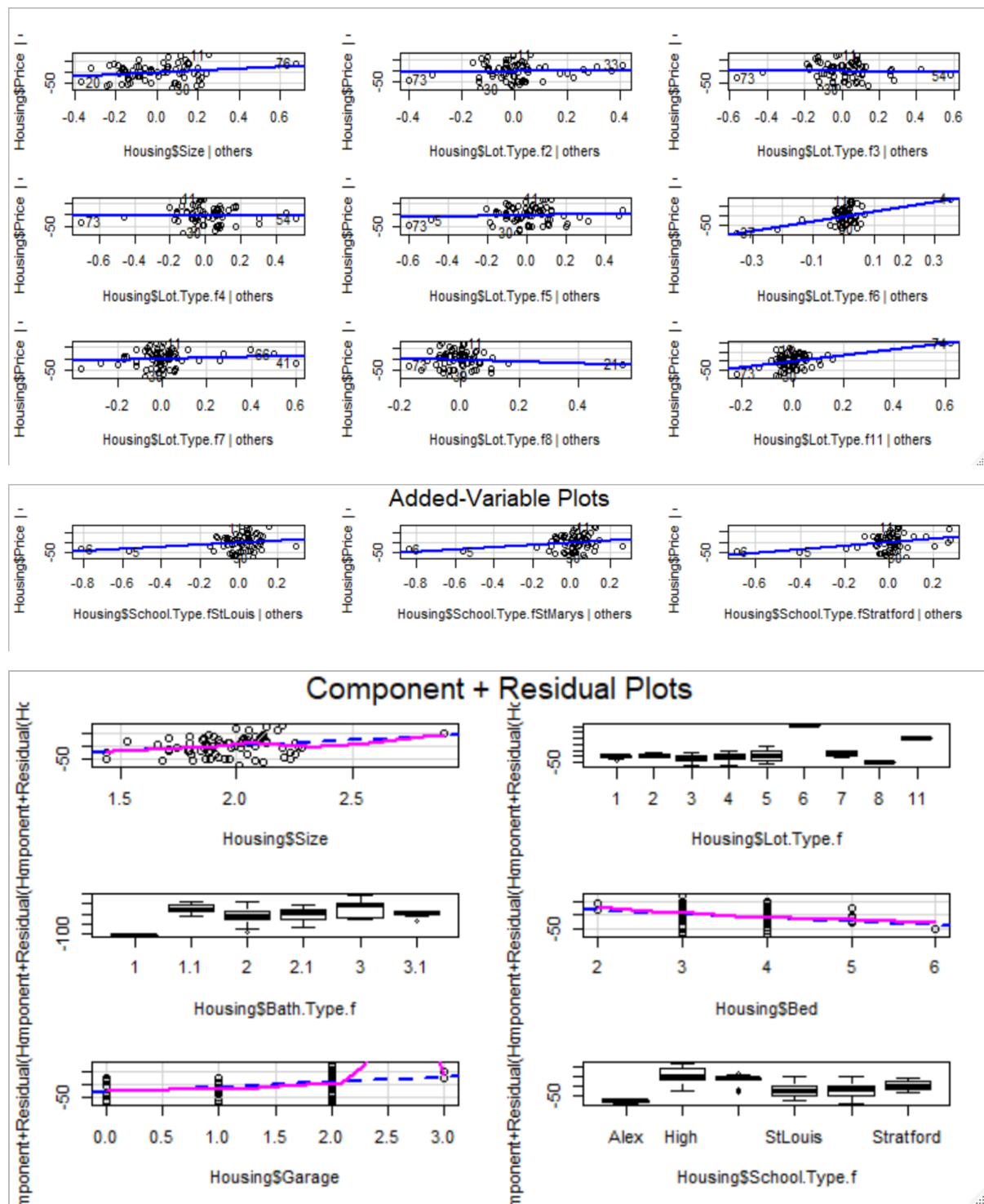
Predictor Variable	Type	Comments
--------------------	------	----------

Size	Linear	Linear, but does not appear to have a positive or negative correlation so the line may just go straight across
Lot	Linear	Perhaps slightly positively correlation
Bath	Linear	Categorical but appears linear
Bed	Linear	Perhaps negatively correlated as there appears to be more located in the lower right and upper left.
Year	Non-Linear	The year appears to rise and fall in an inverted U shape, suggesting possible non-linear relationship. This has been removed in the updated model
Garage		Perhaps positively correlated as there appears to be heavier population in an positive slope moreso than if a negative slope was on it.

The remaining variables I had plotted but they appeared odd and hard to interpret.



From the Added variable plot, it would appear that they all have a positive slope except for beds, when looking at the reduced model ('Year' removed).



From the Component-plus-residual plot graphs, Size (+), Bed (-) have a linear relationship but Garage appears to be non-linear as the purple/pink line deviates from the blue linear line. It can create random outcomes in the model that cannot be explained. To correct this we can use transformations, or polynomials.

Q2. CHECK THE RANDOM/I.I.D. SAMPLE ASSUMPTION BY CAREFULLY READING THE DATA DESCRIPTION AND COMPUTING THE DURBIN WATSON TEST (STATE THE HYPOTHESIS OF THE TEST, THE TEST STATISTIC AND P-VALUE AND THE CONCLUSION IN THE CONTEXT OF THE PROBLEM). WHAT ARE THE TWO COMMON VIOLATIONS OF THE RANDOM/I.I.D. SAMPLE ASSUMPTION? WHAT EFFECT WOULD DEPENDANT SAMPLES HAVE ON THE

REGRESSION MODEL AND HOW MIGHT YOU CORRECT OR IMPROVE THE MODEL IN THE PRESENCE OF DEPENDANT SAMPLES?

lag Autocorrelation D-W Statistic p-value
 1 0.1760459 1.612759 0.038
 Alternative hypothesis: rho != 0

Hypothesis H_0 : There is no correlation among residuals (independent)
 vs.

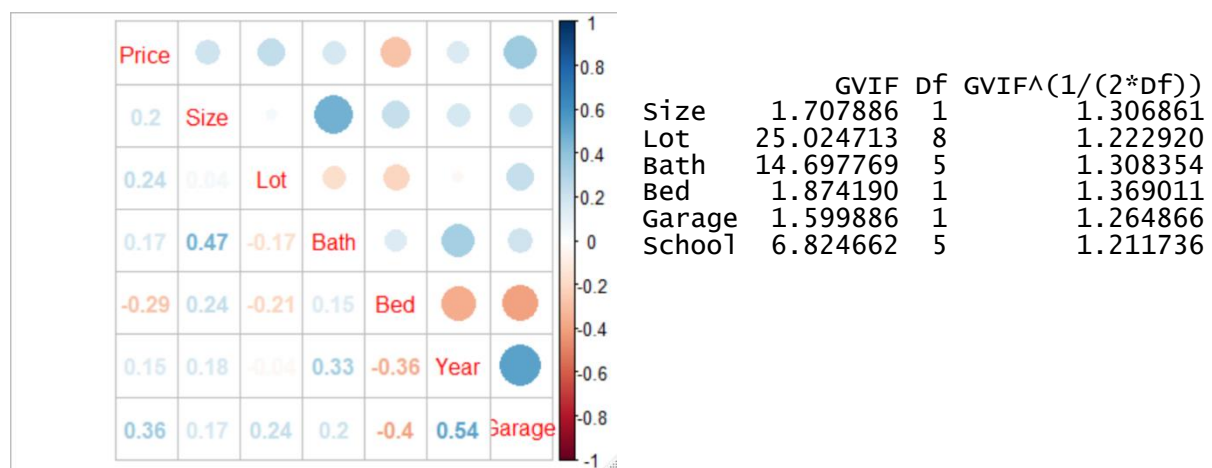
H_a : The residuals are autocorrelated.

Since the test showed a Statistic of 1.6 we can say that it is indicative of positive autocorrelation, but the p-value is < 0.05 so the null hypothesis is rejected, therefore the observations cannot be classed as independent.

The 2 most common violations are repeated measures (the same object is measured on different dates, so medication dose over a month period), and multiple measures (analysing a test when a student performs well on one section, and expect them to perform well on another).

The dependencies would create bias and make the model inefficient as it would become more difficult to interpret. To improve the model, we can use mixed effect models if the data is normally distributed, or log transformations if it isn't, but that data should have repeated measures. If there is dependence then we could look at time series analysis to improve the model.

Q3. CHECK THE COLLINEARITY ASSUMPTION BY INTERPRETING THE CORRELATION AND VARIANCE INFLATION FACTORS. WHAT EFFECT WOULD MULTICOLLINEARITY HAVE ON THE REGRESSION MODEL AND HOW MIGHT YOU CORRECT OR IMPROVE THE MODEL IN THE PRESENCE OF MULTICOLLINEARITY.

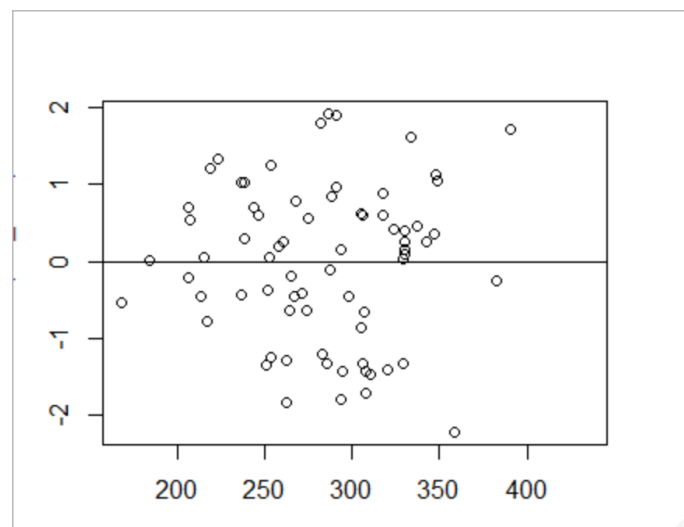


The Variance Inflation Factor using $GVIF^{(1/(2*df))}$ does not appear to show multicollinearity as no value is too high (~4 or higher, using rule of thumb). Since we have many degrees of freedom (df) in some variables the Generalized Variance Inflation Factor formula:

$GVIF^{(1/(2*df))}$ is used instead. A strong correlation between 2 variables X_j and X_k , then $\hat{\beta}$ becomes unstable. The estimate of β_j will then depend heavily on the other variables in the model. Therefore if the variables are correlated we can't interpret the coefficients like before. To improve the model with multicollinearity we could try removing the predictors that are highly correlated, or use Partial Least Squares Regression (PLS), Principal

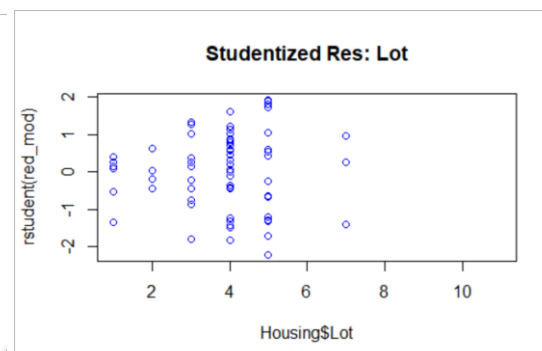
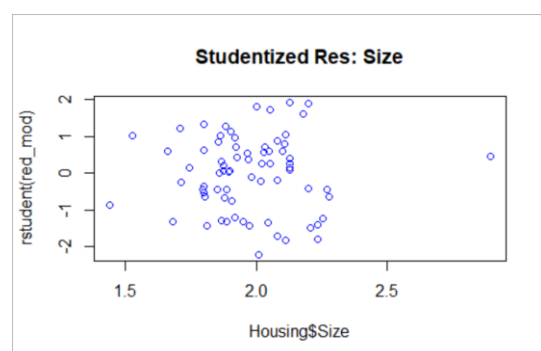
Component Analysis (PCA), or Ridge Regression. Any of these methods achieve improvement by using a subset of variables.

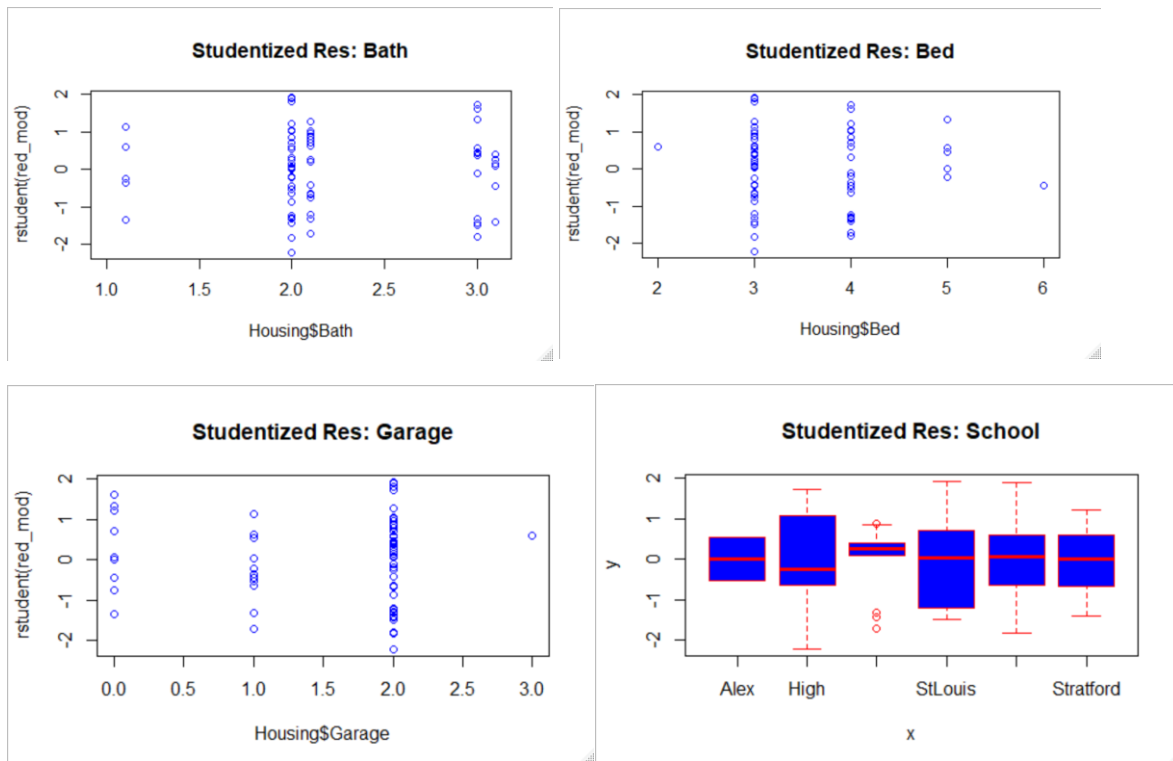
Q4. CHECK THE ZERO CONDITIONAL MEAN AND HOMOSCEDASTICITY ASSUMPTION BY INTERPRETING THE STUDENTIZED RESIDUALS VRS FITTED VALUES PLOTS AND THE STUDENTIZED RESIDUALS VRS PREDICTOR VARIABLE PLOTS. WHAT EFFECT WOULD HETEROSCEDASTICITY HAVE ON THE REGRESSION MODEL AND HOW MIGHT YOU CORRECT OR IMPROVE THE MODEL IN THE PRESENCE OF HETEROSCEDASTICITY.



FITTED VS STUDENTIZED RESIDUALS

From the plot on the left we can see that there is homoscedasticity as there is a symmetric spread from the 0 line in the positive and negative direction. There is a little less spread on the right which could indicate outliers.

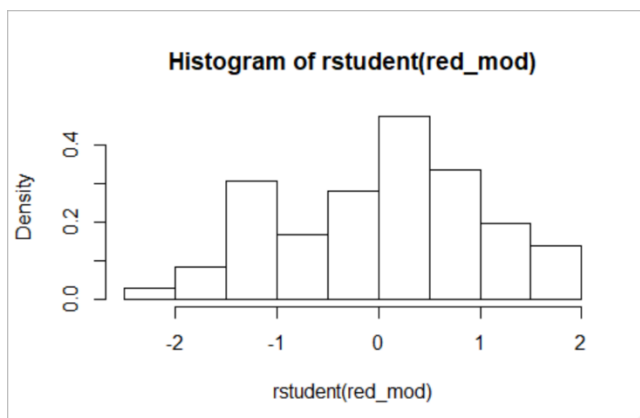




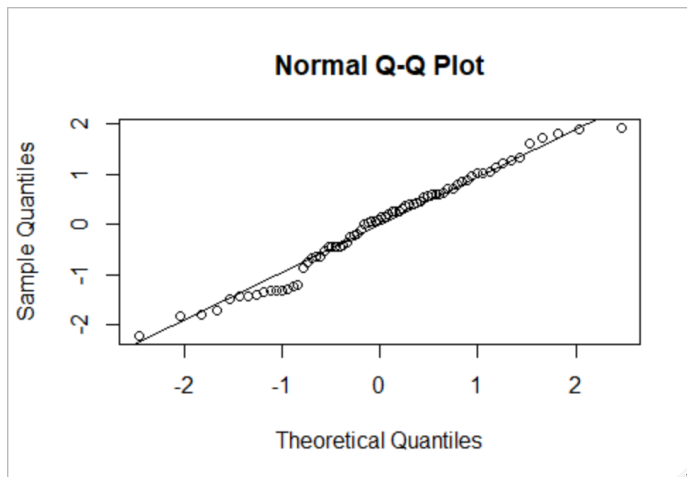
Some of these are categorical and cannot be interpreted correctly, only numeric values should be considered

Heteroskedasticity would cause the standard errors to become biased, which could affect the hypothesis tests. To correct this we can use Weighted Least Squares which adds a weight that are inversely proportionate to the variability of the data, i.e. sparse data is weighted differently to heavily concentrated data. This helps the model become more homoscedastic, but may still have outliers.

Q5. CHECK THE NORMALITY ASSUMPTION BY INTERPRETING THE HISTOGRAM AND QUANTILE-QUANTILE PLOT OF THE STUDENTIZED RESIDUALS. WHAT EFFECT WOULD NON-NORMALITY HAVE ON THE REGRESSION MODEL AND HOW MIGHT YOU CORRECT OR IMPROVE THE MODEL IN THE PRESENCE OF NON-NORMALITY.



The histogram shows relatively normal distribution with a slight increase in -1 to -1.5. This could suggest outliers



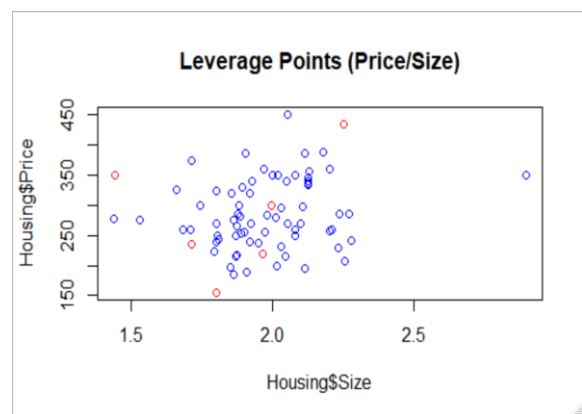
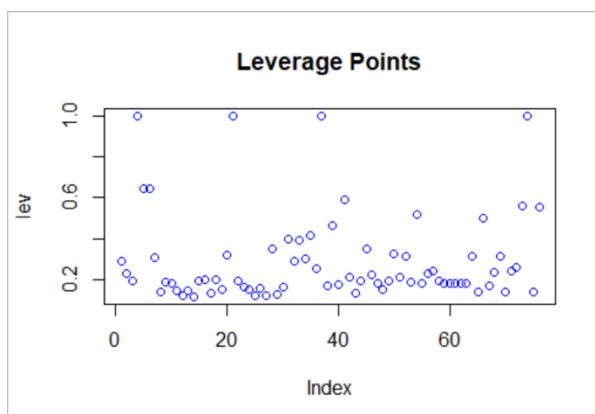
The QQ-Plot shows a very close relationship to the line, with a little deviation around -1, this is expected based from the histogram. However, the points are all close enough to assume normality.

If non-normality was present in our model then it would affect the critical values of the t-test and F-test. To correct this we can perform transformations on response/predictor variables. We can create interaction models, which makes the model more complex. Alternatively, we can use a different model as the current one is not flexible enough for our data.

Leverage, Influence and Outliers

Q1 WHAT IS A LEVERAGE POINT? WHAT EFFECT WOULD A LEVERAGE POINT HAVE ON THE REGRESSION MODEL? USE THE LEVERAGE VALUES AND THE LEVERAGE PLOTS TO SEE IF THERE IS ANY LEVERAGE POINTS.

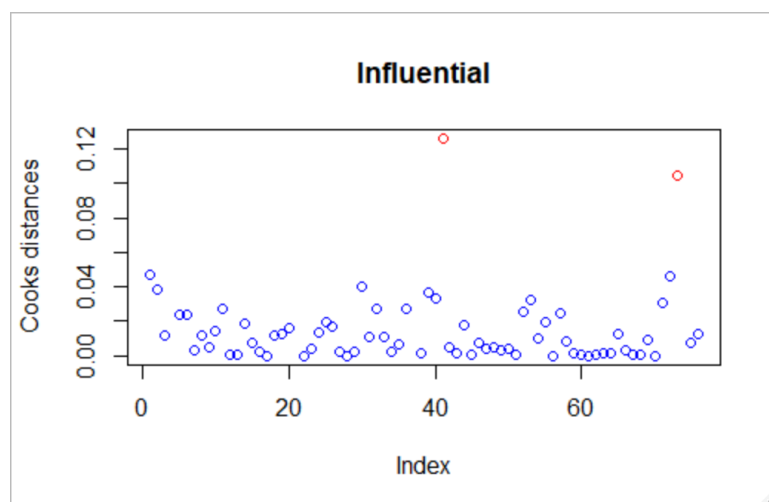
A Leverage point is a data point with an unusual X-value, which can affect the statistics of the model summary (R^2 , SSE, etc) but it will have minimum impact on the estimates of regression coefficients. If they are high, they have the potential to affect the fit of the model.



	Price	Size
4	350.0	1.442
5	155.5	1.800
6	220.0	1.965
21	299.0	1.994

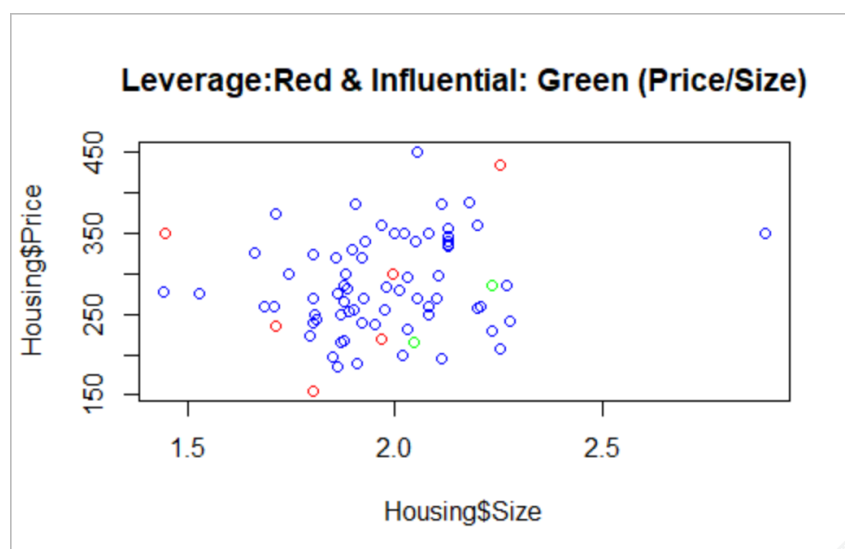
37 235.0 1.712
74 435.0 2.253

Q2. WHAT IS AN INFLUENTIAL POINT? WHAT EFFECT WOULD AN INFLUENTIAL POINT HAVE ON THE REGRESSION MODEL? USE THE INFLUENCE PLOT TO SEE IF THERE IS ANY INFLUENCE POINTS



An Influential Point has an unusual \mathbf{Y} value. It will move the regression model in the direction of its point.

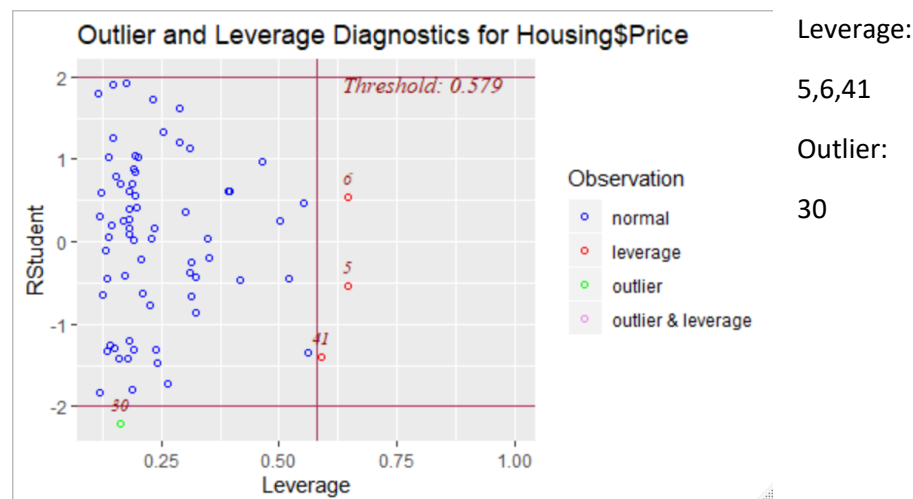
41 and 73 are influential points above 0.08



Q3. WHAT IS AN OUTLIER? WHAT EFFECT WOULD AN OUTLIER HAVE ON THE REGRESSION MODEL? HOW WOULD YOU CORRECT FOR OUTLIERS? USE THE OUTLIER TEST AND OUTLIER AND LEVERAGE DIAGNOSTICS PLOT TO SEE IF THERE IS ANY OUTLIERS. DEAL WITH THE OUTLIERS IF ANY ARE IDENTIFIED.

An outlier is an observation (i) within the data where the response does not relate to the what the model fitted for the majority of the data. The impact of an outlier is that it could affect the estimation of the regression coefficients. If it is not a high leverage/influence point we can remove the outlier from the data.

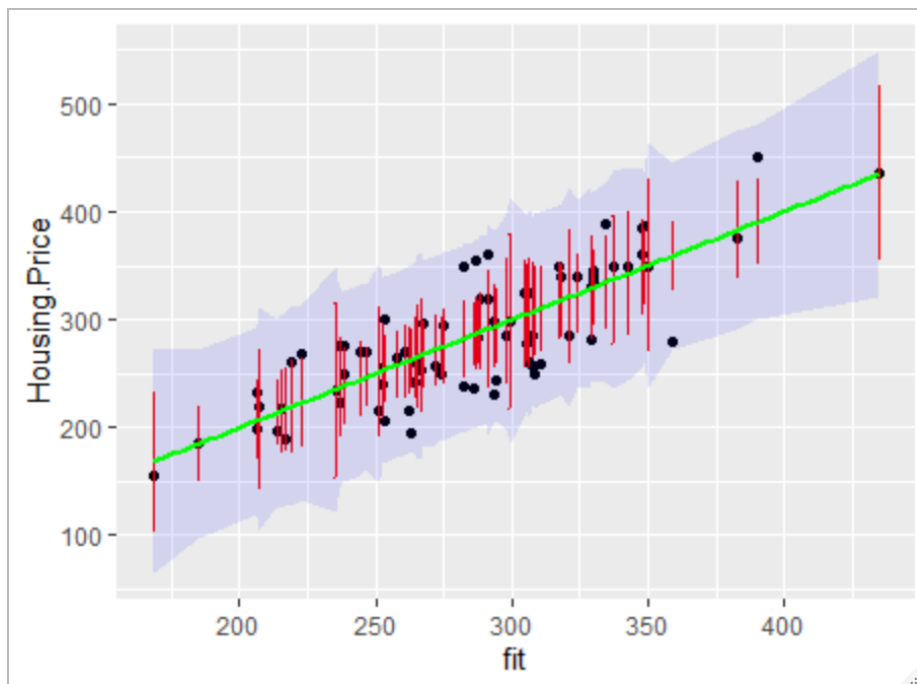
```
No studentized residuals with Bonferroni p < 0.05
Largest |rstudent|:
  rstudent unadjusted p-value Bonferroni p
30 -2.213021          0.031225          NA
```



Since no studentized residuals had a p-value < 0.05. The graphic does highlight entry 30 as an outlier. The other possibility is to compare the model with and without the model and see which model is better.

Expected Value, CI and PI

Q1. PLOT THE OBSERVED HOUSE PRICES, THEIR EXPECTED VALE (FITTED VALUE), CONFIDENCE INTERVALS (IN RED) AND PREDICTION INTERVALS (IN BLUE). LOOKING AT THIS PLOT IS THIS MODEL PROVIDING A GOOD ESTIMATE OF THE HOUSE PRICES.



The model is providing a good measure of predictions, the data falls within the prediction measures and most of the points are within the confidence intervals. The graph shows homoskedasticity

R CODE

Patrick Lowe - 16725829

open the housing CSV

Housing = read.csv(file.choose())

par(mar=c(3,3,3,3)) # this works best for my machine

#####

Q1

#####

Create a boxplot for housing prices

boxplot(Housing\$Price)

Create a Histogram of prices

hist(Housing\$Price, freq=FALSE)

lines(density(Housing\$Price),lwd=2, col="blue")

create a summary

```
summary(Housing)
```

```
#####
```

```
# Q2 #
```

```
#####
```

```
# Convert Categorical to factors
```

```
Housing$Lot.Type.f <- factor(Housing$Lot)
```

```
Housing$Bath.Type.f <- factor(Housing$Bath)
```

```
Housing$Bed.Type.f <- factor(Housing$Bed)
```

```
#Housing$Garage.Type.f <- factor(Housing$Garage) # chosen not to treat as categorical, very well  
could be
```

```
Housing$School.Type.f <- factor(Housing$School)
```

```
#SUMMARY AND A BOXPLOT
```

```
#PRICES: BEDROOMS,BATHROOMS,GARAGE,SCHOOL
```

```
boxplot(Housing$Price~Housing$Bed,data=Housing,main="Price by  
#Bedrooms",col="blue",border="red")
```

```
tapply(Housing$Price, Housing$Bed, summary)
```

```
boxplot(Housing$Price~Housing$Bath,data=Housing,main="Price by  
#Bathrooms",col="blue",border="red")
```

```
tapply(Housing$Price, Housing$Bath, summary)
```

```
boxplot(Housing$Price~Housing$Garage,data=Housing,main="Price by Garage  
Space",col="blue",border="red")
```

```
tapply(Housing$Price, Housing$Garage, summary)
```

```
boxplot(Housing$Price~Housing$School,data=Housing,main="Price by  
Schools",col="blue",border="red")
```

```
tapply(Housing$Price, Housing$School, summary)
```

```
#####
```

```
# Q3 #
```

```
#####
```

```
# Summary, Correlation, Pairs Plots:
```

```
#price and each of the numeric predictorvariables.
```

```
# Summary
```

```
tapply(Housing$Price, Housing$Bed, summary)
```

```
# Correlation
```

```
# step by step: cor(Housing$Price,Housing$Bed)
```

```
# Or graphically:
```

```
library(corrplot)
```

```
M <- cor(Housing[,1:7])
```

```
corrplot.mixed(M)
```

```
# Pairs Plots
```

```
pairs(Housing[,1:7])
```

```
#####
```

```
# Q4 #
```

```
#####
```

```
# MLR Model
```

```
mod = lm(Housing$Price ~ Housing$Size
```

```
  + Housing$Lot.Type.f
```

```
  + Housing$Bath.Type.f
```

```
  + Housing$Bed.Type.f
```

```
  + Housing$Year
```

```
  + Housing$Garage
```

```
  + Housing$School.Type.f,
```

```
  data=Housing)
```

```
summary(mod)
```

```
# Plot Residuals
```

```
resid(mod) #List of residuals
```

```
plot(density(resid(mod)))
qqnorm(resid(mod))
qqline(resid(mod))
library(ggplot2)
residualPlot <- ggplot(aes(x=.fitted, y=.resid),data=mod)+geom_point()+geom_hline(yintercept=0)
residualPlot
```

```
#####
```

```
# Anova #
```

```
#####
```

```
# COMPUTE THE TYPE 1 ANOVA TABLE. INTERPRET THE OUTPUT
```

```
anova(mod)
```

```
SSR <- sum(anova(mod)[1:7,2])
```

```
MSR <- SSR/7
```

```
SSE <- anova(mod)[8,2]
```

```
MSE <- anova(mod)[8,3]
```

```
test_stat <- MSR/MSE
```

```
qf = qf(0.05,22,53)
```

```
#Q3. Compute a type 2 anova table comparing the full model with all predictor
```

```
#variables to the the reduced model with the suggested predictor variable
```

```
#identified in the previous question removed.
```

```
# Reduced Model
```

```
red_mod = lm(Housing$Price ~ Housing$Size
```

```
  + Housing$Lot.Type.f
```

```
  + Housing$Bath.Type.f
```

```
  + Housing$Bed
```

```
  + Housing$Garage
```

```
  + Housing$School.Type.f,
```

```
  data=Housing)
```



```
Anova(red_mod)
SSR2 <- sum(anova(red_mod)[1:6,2])
MSR2 <- SSR/6
SSE2 <- anova(mod)[7,2]
MSE2 <- anova(mod)[7,3]
test_stat2 <- MSR2/MSE2
qf2 = qf(0.05,21,54)

anova(red_mod, mod)

#####
# Diagnostics #
#####
library("GGally")
library(car)
ggpairs(Housing[,1:7])
ggpairs(Housing[,c(1:1, (ncol(Housing) - 4):ncol(Housing))])
avPlots(red_mod)
crPlots(red_mod)

# random/i.i.d sample
dwt(red_mod)

# multicollinearity
gvif(red_mod)
M <- cor(Housing)
corrplot.mixed(M)

# Zero Conditional Mean & Homoskedasticity
plot(fitted(red_mod),rstudent(red_mod))
abline(h=0)
```

```
plot(Housing$Size,rstudent(red_mod),main="Studentized Res: Size",col="blue")
plot(Housing$Lot,rstudent(red_mod),main="Studentized Res: Lot",col="blue")
plot(Housing$Bath,rstudent(red_mod),main="Studentized Res: Bath",col="blue")
plot(Housing$Bed,rstudent(red_mod),main="Studentized Res: Bed",col="blue")
plot(Housing$Garage,rstudent(red_mod),main="Studentized Res: Garage",col="blue")
plot(Housing$School,rstudent(red_mod),main="Studentized Res: School",col="blue",border="red")
```

```
# normality
```

```
hist(rstudent(red_mod),freq=FALSE)
lines(density(rstudent(red_mod)), lwd=2, col="blue")
qqnorm(rstudent(red_mod))
qqline(rstudent(red_mod))
```

```
# Leverage points
```

```
library(olsrr)
lev = hat(model.matrix(red_mod))
plot(lev,main="Leverage Points",col="blue")
# highlighting high leverage points
plot(Housing$Size,Housing$Price,main="Leverage Points (Price/Size)",col="blue")
points(Housing[4,]$Size,Housing[4,]$Price,col="red")
points(Housing[5,]$Size,Housing[5,]$Price,col="red")
points(Housing[6,]$Size,Housing[6,]$Price,col="red")
points(Housing[21,]$Size,Housing[21,]$Price,col="red")
points(Housing[37,]$Size,Housing[37,]$Price,col="red")
points(Housing[74,]$Size,Housing[74,]$Price,col="red")
```

```
# Influential Points
```

```
cook = cooks.distance(red_mod)
plot(cook,ylab="Cooks distances")
which(cook>0.08)
plot(cook,ylab="Cooks distances",main="Influential",col="blue")
```

```
points(41,cook[41],col="red")
points(73,cook[73],col="red")

# plot both points
plot(Housing$Size,Housing$Price,main="Leverage:Red & Influential: Green (Price/Size)",col="blue")
points(Housing[4,]$Size,Housing[4,]$Price,col="red")
points(Housing[5,]$Size,Housing[5,]$Price,col="red")
points(Housing[6,]$Size,Housing[6,]$Price,col="red")
points(Housing[21,]$Size,Housing[21,]$Price,col="red")
points(Housing[37,]$Size,Housing[37,]$Price,col="red")
points(Housing[74,]$Size,Housing[74,]$Price,col="red")
points(Housing[41,]$Size,Housing[41,]$Price,col="green")
points(Housing[73,]$Size,Housing[73,]$Price,col="green")

# Outliers
outlierTest(red_mod)
CD=cooks.distance(red_mod)
CD[as.numeric(which(CD> 0.05))]]

#PLOT PRICES, FITTED VALUE
new.prices <- data.frame(Housing$Price)
CI <- predict(red_mod, newdata = new.prices, interval = "confidence")
PI <- predict(red_mod, newdata = new.prices, interval = "predict")

plotdata <- data.frame(Housing$Price,CI[,1:3],PI[,2:3])
names(plotdata)[names(plotdata) == "lwr"] <- "CI lwr"
names(plotdata)[names(plotdata) == "upr"] <- "CI upr"
names(plotdata)[names(plotdata) == "lwr.1"] <- "PI lwr"
names(plotdata)[names(plotdata) == "upr.1"] <- "PI upr"

plotdata[,1]=round(plotdata[,1],2)
```

```
plotdata[,2]=round(plotdata[,2],2)
plotdata[,3]=round(plotdata[,3],2)
plotdata[,4]=round(plotdata[,4],2)
plotdata[,5]=round(plotdata[,5],2)
plotdata[,6]=round(plotdata[,6],2)
```

```
library(ggplot2)
U <- plotdata$`CI upr`
L <- plotdata$`CI lwr`
U2 <- plotdata$`PI upr`
L2 <- plotdata$`PI lwr`
```

```
ggplot(plotdata, aes(x = fit, y = Housing.Price)) +
  geom_point() +
  geom_errorbar(aes(ymin = L, ymax = U), colour = "red") +
  geom_ribbon(aes(ymin = L2, ymax = U2), fill = "blue", alpha = 0.1) +
  geom_smooth(method = "lm", se = FALSE, colour = "green")
```