Investigate washout

Some studies have washout, which can mean different things for different studies.

These are all the studies that are labeled as having washout:

```
kable(df %>% filter(WashOut == 'Y') %>%
select(JournalVolumePage, Author, Year, subtype, TypesofMIGSifany))
```

JournalVolumePage	Author	Year	subtype	TypesofMIGSifany
JCRS, 36(407-412)	Fea et. al	2010	OAG	NA
JCRS, 36(407-412)	Fea et. al	2010	OAG	iStent
AAO, 118(459-467)	Samuelson et al.	2011	OAG	NA
AAO, 118(459-467)	Samuelson et al.	2011	OAG	iStent
AAO, 122(1283-1293)	Pfeiffer et al.	2015	OAG	NA
AAO, 122(1283-1293)	Pfeiffer et al.	2015	OAG	Hydrus Microstent
Lancet, $388(1389-1397)$	Azuara-Blanco et al.	2016	ACG	NA
O, 123(2103-2112)	Vold et al.	2016	OAG	NA
O, 123(2103-2112)	Vold et al.	2016	OAG	Cypass

This study has pre- and post-washout and no measurements without washout. This is equivalent to setting the pre and post op meds to 0:

```
kable(df %>% filter(Washoutbaseline == PreOpIOPMean, !is.na(WashoutIOP)) %>%
    select(JournalVolumePage, Author, Year, TypesofMIGSifany))
```

JournalVolumePage	Author	Year	TypesofMIGSifany
AAO, 122(1283-1293)	Pfeiffer et al.	2015	NA
AAO, 122(1283-1293)	Pfeiffer et al.	2015	Hydrus Microstent
O, 123(2103-2112)	Vold et al.	2016	NA
O, 123(2103-2112)	Vold et al.	2016	Cypass

This study has preop washout and no washout in the post period. Virgin patients were enrolled, and had 0 pre-op meds; only after were they put on meds:

```
kable(df %>% filter(Washoutbaseline == PreOpIOPMean, is.na(WashoutIOP)) %>%
select(JournalVolumePage, Author, Year, TypesofMIGSifany))
```

JournalVolumePage	Author	Year	TypesofMIGSifany
Lancet, 388(1389-1397)	Azuara-Blanco et al.	2016	NA

These studies had both washout and regular measurements in the pre-period. That tells us about the relationship between meds and IOP:

rel.p, rel.p.drop.per.med), digits = 2)

Author	Year	PreOpIOPMean	RxPreOpMean	mm.Hg.per.med	rel.p	rel.p.drop.per.med
Samuelson et al.	2011	18.0	1.5	5.00	70.59	20.72
Samuelson et al.	2011	18.7	1.5	4.33	74.21	18.04
Neiweem et al.	2016	19.9	2.4	1.88	81.56	8.14

Most commonly, 1 med corresponds to about a 15-20% drop in IOP; or a 4-5 mmHg drop.

Note: Pfeiffer has washout in the pre-periods and the post-periods. It also has a baseline measurement (with meds).

This study has washout only in the last period; we can also use this to estimate the drop in IOP per med:

Author	Year	${\bf Last Period IOP Mean}$	${\bf RxPostOpMean}$	mm.Hg.per.med	rel.p	rel.p.drop.per.med
Fea et. al	2010	15.7	1.3	2.69	81.77	14.34
Fea et. al	2010	14.8	0.4	4.50	89.16	24.94

This study is also in line with the estimate of 4-5mmHg, 15-20% drop in IOP per med.

Study classifications

I went back and classified the studies depending on the washout type:

```
kable(df %>% filter(washout.type != 'None') %>%
select(JournalVolumePage, Author, Year, subtype, TypesofMIGSifany, washout.type))
```

JournalVolumePage	Author	Year	subtype	TypesofMIGSifany	washout.type
JCRS, 36(407-412)	Fea et. al	2010	OAG	NA	Partial
JCRS, 36(407-412)	Fea et. al	2010	OAG	iStent	Partial
AAO, 118(459-467)	Samuelson et al.	2011	OAG	NA	Partial
AAO, 118(459-467)	Samuelson et al.	2011	OAG	iStent	Partial
AAO, 122(1283-1293)	Pfeiffer et al.	2015	OAG	NA	Both
AAO, 122(1283-1293)	Pfeiffer et al.	2015	OAG	Hydrus Microstent	Both
Lancet, 388(1389-1397)	Azuara-Blanco et al.	2016	ACG	NA	Pre
PLoS ONE, 11(4)e0153585	Neiweem et al.	2016	OAG	Ab Interno Trabeculectomy	Partial
O, 123(2103-2112)	Vold et al.	2016	OAG	NA	Both
O, 123(2103-2112)	Vold et al.	2016	OAG	Cypass	Both

Net effect - including IOP and meds

Let's see what happens when we add the IOP drop effect to the Rx drop effects. We try different values of the translation value between meds and mmHg drop: 0, 2, 3, 4, or 5 mm Hg per Rx.

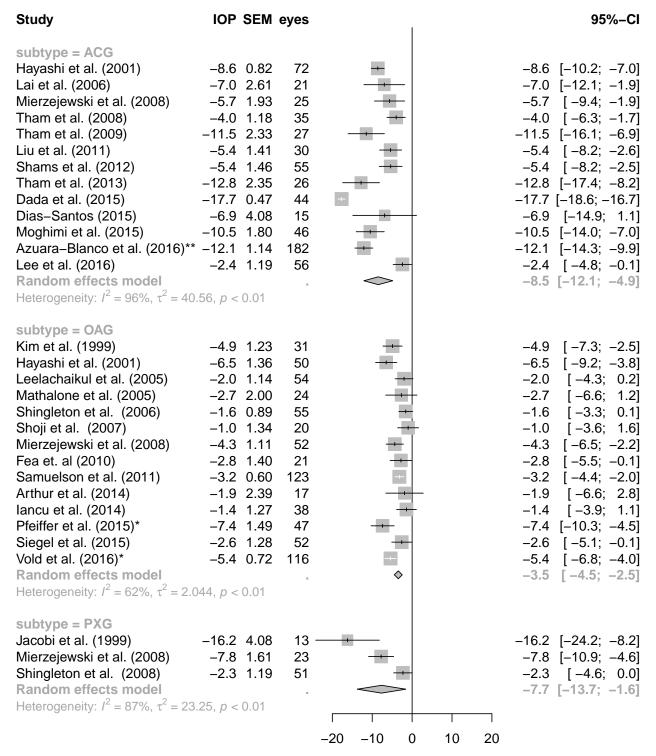
Last period

```
mmhg.per.meds \leftarrow c(0, 2, 3, 4, 5)
for(mmhg.per.med in mmhg.per.meds) {
  df_ <- df %>%
    filter(!is.na(RxChangeMean),
           df$subtype != "acute",
           MIGsYorN == 'N',
           !is.na(LastPeriodAbsIOPChangeStdDev) | !is.na(OneYAbsIOPChangeStdDev)) %>%
    mutate(subtype=factor(subtype),
           net.effect=ifelse(is.na(LastPeriodAbsIOPChangeMean), OneYAbsIOPChangeMean,
                                                                                          LastPeriodAbsIOP
             mmhg.per.med * RxChangeMean * (washout.type %in% c('None', 'Partial')) +
             mmhg.per.med * RxPostOpMean * (washout.type %in% c('Pre')),
           net.sem=2 * ifelse(is.na(LastPeriodAbsIOPChangeStdDev),
                               OneYAbsIOPChangeStdDev,
                               LastPeriodAbsIOPChangeStdDev) / sqrt(LastPeriodEyes))
  m <- metagen(net.effect,</pre>
               net.sem,
               study.name,
               data=df_,
               byvar=subtype,
               n.e=LastPeriodEyes)
  forest(m,
         comb.fixed=FALSE,
         digits=1,
         digits.se = 2,
         overall=FALSE,
         leftcols=c("studlab", "TE", "seTE", "n.e"),
         leftlabs=c("Study", "IOP", "SEM", "eyes"))
  grid.text(
    pasteO("Simulated net change in IOP in last period, ", mmhg.per.med, " mm Hg per med"), .5, .97, gp
 print(" ")
```

Simulated net change in IOP in last period, 0 mm Hg per med

Study	IOP SEM	eyes	1	95%-CI
subtype = ACG Hayashi et al. (2001) Lai et al. (2006) Mierzejewski et al. (2008) Tham et al. (2009) Liu et al. (2011) Shams et al. (2012) Tham et al. (2013) Dada et al. (2015) Dias–Santos (2015) Moghimi et al. (2015) Azuara–Blanco et al. (2016)** Lee et al. (2016) Random effects model Heterogeneity: I² = 95%, τ² = 27	-1.7 1.19	72 21 25 35 27 30 55 26 44 15 46 182 56		-7.2 [-8.8; -5.6] -4.2 [-9.3; 0.9] -5.1 [-8.9; -1.3] -1.8 [-4.1; 0.5] -8.3 [-12.9; -3.7] -3.8 [-6.6; -1.0] -4.5 [-7.4; -1.6] -8.4 [-13.0; -3.8] -13.9 [-14.8; -13.0] -5.4 [-13.4; 2.6] -8.3 [-11.8; -4.8] -12.9 [-15.1; -10.7] -1.7 [-4.0; 0.6] -6.6 [-9.7; -3.6]
subtype = OAG Kim et al. (1999) Hayashi et al. (2001) Leelachaikul et al. (2005) Mathalone et al. (2006) Shingleton et al. (2006) Shoji et al. (2007) Mierzejewski et al. (2008) Fea et. al (2010) Samuelson et al. (2011) Arthur et al. (2014) Iancu et al. (2014) Pfeiffer et al. (2015)* Siegel et al. (2015) Vold et al. (2016)* Random effects model Heterogeneity: I² = 70%, τ² = 2.	-2.9 1.23 -5.3 1.36 -1.6 1.14 -1.9 2.00 -1.4 0.89 -1.0 1.34 -4.0 1.11 -1.6 1.40 -1.0 0.60 -2.1 2.39 -1.9 1.27 -7.4 1.49 -2.2 1.28 -5.4 0.72	31 50 54 24 55 20 52 21 123 17 38 47 52 116		-2.9 [-5.3; -0.5] -5.3 [-8.0; -2.6] -1.6 [-3.8; 0.6] -1.9 [-5.8; 2.0] -1.4 [-3.1; 0.3] -1.0 [-3.6; 1.6] -4.0 [-6.2; -1.8] -1.6 [-4.3; 1.1] -1.0 [-2.2; 0.2] -2.1 [-6.8; 2.6] -1.9 [-4.4; 0.6] -7.4 [-10.3; -4.5] -2.2 [-4.7; 0.3] -5.4 [-6.8; -4.0] -2.8 [-3.9; -1.7]
subtype = PXG Jacobi et al. (1999) Mierzejewski et al. (2008) Shingleton et al. (2008) Random effects model Heterogeneity: $I^2 = 86\%$, $\tau^2 = 21$	-14.0 4.08 -6.6 1.61 -1.1 1.19	13 - 23 51	20 -10 0	-14.0 [-22.0; -6.0] -6.6 [-9.8; -3.5] -1.1 [-3.4; 1.2] -6.2 [-12.0; -0.3]

Simulated net change in IOP in last period, 2 mm Hg per med



Simulated net change in IOP in last period, 3 mm Hg per med

Study	IOP SEM	eyes	1	95%-CI
subtype = ACG Hayashi et al. (2001) Lai et al. (2006) Mierzejewski et al. (2008) Tham et al. (2009) Liu et al. (2011) Shams et al. (2012) Tham et al. (2013) Dada et al. (2015) Dias–Santos (2015) Moghimi et al. (2015) Azuara–Blanco et al. (2016)** Lee et al. (2016) Random effects model Heterogeneity: I² = 97%, τ² = 48	-2.8 1.19	72 21 25 35 27 30 55 26 44 15 46 182 56		-9.3 [-10.9; -7.7] -8.4 [-13.5; -3.3] -5.9 [-9.7; -2.1] -5.1 [-7.4; -2.8] -13.1 [-17.7; -8.5] -6.2 [-9.0; -3.4] -5.8 [-8.7; -3.0] -15.0 [-19.6; -10.4] -19.6 [-20.5; -18.6] -7.6 [-15.6; 0.4] -11.6 [-15.1; -8.1] -11.7 [-13.9; -9.5] -2.8 [-5.1; -0.4] -9.4 [-13.4; -5.5]
subtype = OAG Kim et al. (1999) Hayashi et al. (2001) Leelachaikul et al. (2005) Mathalone et al. (2006) Shingleton et al. (2006) Shoji et al. (2007) Mierzejewski et al. (2008) Fea et. al (2010) Samuelson et al. (2011) Arthur et al. (2014) lancu et al. (2014) Pfeiffer et al. (2015)* Siegel et al. (2015) Vold et al. (2016)* Random effects model Heterogeneity: $I^2 = 66\%$, $\tau^2 = 2$.	-5.9 1.23 -7.1 1.36 -2.2 1.14 -3.1 2.00 -1.7 0.89 -1.0 1.34 -4.5 1.11 -3.4 1.40 -4.3 0.60 -1.8 2.39 -1.2 1.27 -7.4 1.49 -2.8 1.28 -5.4 0.72	31 50 54 24 55 20 52 21 123 17 38 47 52 116		-5.9 [-8.3; -3.5] -7.1 [-9.8; -4.4] -2.2 [-4.5; 0.0] -3.1 [-7.0; 0.8] -1.7 [-3.4; 0.0] -1.0 [-3.6; 1.6] -4.5 [-6.7; -2.3] -3.4 [-6.1; -0.7] -4.3 [-5.5; -3.1] -1.8 [-6.5; 2.9] -1.2 [-3.7; 1.3] -7.4 [-10.3; -4.5] -2.8 [-5.3; -0.3] -5.4 [-6.8; -4.0] -3.8 [-4.8; -2.7]
subtype = PXG Jacobi et al. (1999) Mierzejewski et al. (2008) Shingleton et al. (2008) Random effects model Heterogeneity: $I^2 = 88\%$, $\tau^2 = 24$	-17.3 4.08 -8.4 1.61 -2.9 1.19	13 · 23 51	-20 -10 0 10 20	-17.3 [-25.3; -9.3] -8.4 [-11.5; -5.2] -2.9 [-5.2; -0.6] -8.4 [-14.6; -2.2]

Simulated net change in IOP in last period, 4 mm Hg per med

Study	IOP SEM	eyes	1		95%-CI
subtype = ACG Hayashi et al. (2001) Lai et al. (2006) Mierzejewski et al. (2008) Tham et al. (2008) Tham et al. (2009) Liu et al. (2011) Shams et al. (2012) Tham et al. (2013) Dada et al. (2015) Dias–Santos (2015) Moghimi et al. (2015) Azuara–Blanco et al. (2016)* Lee et al. (2016) Random effects model Heterogeneity: $I^2 = 98\%$, $\tau^2 = 58$	-3.1 1.19	72 21 25 35 27 30 55 26 44 15 46 182 56	++++++		-10.0 [-11.6; -8.4] -9.8 [-14.9; -4.6] -6.2 [-10.0; -2.4] -6.2 [-8.5; -3.9] -14.7 [-19.3; -10.1] -7.0 [-9.8; -4.2] -6.3 [-9.1; -3.4] -17.2 [-21.8; -12.6] -21.4 [-22.4; -20.5] -8.4 [-16.4; -0.4] -12.7 [-16.2; -9.2] -11.3 [-13.5; -9.1] -3.1 [-5.5; -0.8] -10.3 [-14.7; -6.0]
subtype = OAG Kim et al. (1999) Hayashi et al. (2001) Leelachaikul et al. (2005) Mathalone et al. (2005) Shingleton et al. (2006) Shoji et al. (2007) Mierzejewski et al. (2008) Fea et. al (2010) Samuelson et al. (2011) Arthur et al. (2014) Iancu et al. (2014) Pfeiffer et al. (2015)* Siegel et al. (2015) Vold et al. (2016)* Random effects model Heterogeneity: $I^2 = 72\%$, $\tau^2 = 3$	-6.9 1.23 -7.7 1.36 -2.4 1.14 -3.5 2.00 -1.8 0.89 -1.0 1.34 -4.7 1.11 -4.0 1.40 -5.4 0.60 -1.7 2.39 -0.9 1.27 -7.4 1.49 -3.0 1.28 -5.4 0.72	31 50 54 24 55 20 52 21 123 17 38 47 52 116			-6.9 [-9.3; -4.5] -7.7 [-10.4; -5.0] -2.4 [-4.7; -0.2] -3.5 [-7.4; 0.4] -1.8 [-3.5; -0.1] -1.0 [-3.6; 1.6] -4.7 [-6.9; -2.5] -4.0 [-6.7; -1.3] -5.4 [-6.6; -4.2] -1.7 [-6.4; 3.0] -0.9 [-3.4; 1.5] -7.4 [-10.3; -4.5] -3.0 [-5.5; -0.5] -5.4 [-6.8; -4.0] -4.1 [-5.2; -2.9]
subtype = PXG Jacobi et al. (1999) Mierzejewski et al. (2008) Shingleton et al. (2008) Random effects model Heterogeneity: $I^2 = 88\%$, $\tau^2 = 28\%$	-18.4 4.08 -8.9 1.61 -3.5 1.19 5.53, <i>p</i> < 0.01	13 · 23 51	-20 -10 0	10 20	-18.4 [-26.4; -10.4] -8.9 [-12.1; -5.8] -3.5 [-5.8; -1.2] -9.2 [-15.5; -2.9]

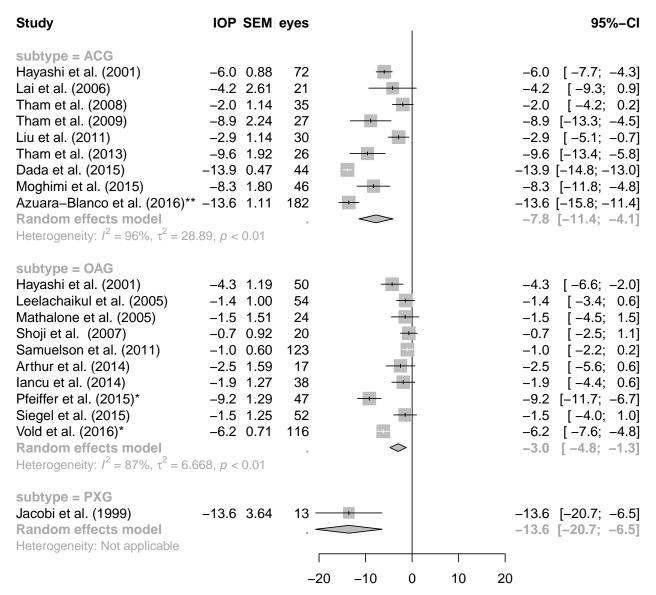
Simulated net change in IOP in last period, 5 mm Hg per med

Study	IOP SEM	eyes	I	95%-CI
subtype = ACG Hayashi et al. (2001) Lai et al. (2006) Mierzejewski et al. (2008) Tham et al. (2008) Tham et al. (2009) Liu et al. (2011) Shams et al. (2012) Tham et al. (2013) Dada et al. (2015) Dias—Santos (2015) Moghimi et al. (2015) Azuara—Blanco et al. (2016)* Lee et al. (2016) Random effects model Heterogeneity: $I^2 = 98\%$, $\tau^2 = 7$	-3.5 1.19	72 21 25 35 27 30 55 26 44 15 46 182 56	++++++	-10.7 [-12.3; -9.1] -11.1 [-16.3; -6.0] -6.5 [-10.3; -2.7] -7.3 [-9.6; -5.0] -16.3 [-20.9; -11.7] -7.8 [-10.6; -5.0] -6.7 [-9.6; -3.8] -19.4 [-24.0; -14.8] -23.3 [-24.2; -22.4] -9.1 [-17.1; -1.1] -13.8 [-17.3; -10.3] -10.9 [-13.1; -8.7] -3.5 [-5.8; -1.2] -11.3 [-16.0; -6.6]
subtype = OAG Kim et al. (1999) Hayashi et al. (2001) Leelachaikul et al. (2005) Mathalone et al. (2005) Shingleton et al. (2006) Shoji et al. (2007) Mierzejewski et al. (2008) Fea et. al (2010) Samuelson et al. (2011) Arthur et al. (2014) Iancu et al. (2014) Pfeiffer et al. (2015)* Siegel et al. (2015) Vold et al. (2016)* Random effects model Heterogeneity: I² = 79%, τ² = 4	-7.9 1.23 -8.3 1.36 -2.6 1.14 -3.9 2.00 -1.9 0.89 -1.0 1.34 -4.9 1.11 -4.6 1.40 -6.5 0.60 -1.6 2.39 -0.7 1.27 -7.4 1.49 -3.2 1.28 -5.4 0.72	31 50 54 24 55 20 52 21 123 17 38 47 52 116		-7.9 [-10.3; -5.5] -8.3 [-11.0; -5.6] -2.6 [-4.9; -0.4] -3.9 [-7.8; 0.0] -1.9 [-3.6; -0.2] -1.0 [-3.6; 1.6] -4.9 [-7.0; -2.7] -4.6 [-7.3; -1.9] -6.5 [-7.7; -5.3] -1.6 [-6.3; 3.1] -0.7 [-3.2; 1.8] -7.4 [-10.3; -4.5] -3.2 [-5.7; -0.7] -5.4 [-6.8; -4.0] -4.4 [-5.7; -3.1]
subtype = PXG Jacobi et al. (1999) Mierzejewski et al. (2008) Shingleton et al. (2008) Random effects model Heterogeneity: $I^2 = 89\%$, $\tau^2 = 20$	-19.5 4.08 -9.5 1.61 -4.1 1.19 6.74, $p < 0.01$	13 23 51	*	-19.5 [-27.5; -11.5] -9.5 [-12.7; -6.4] -4.1 [-6.4; -1.8] -9.9 [-16.4; -3.5]

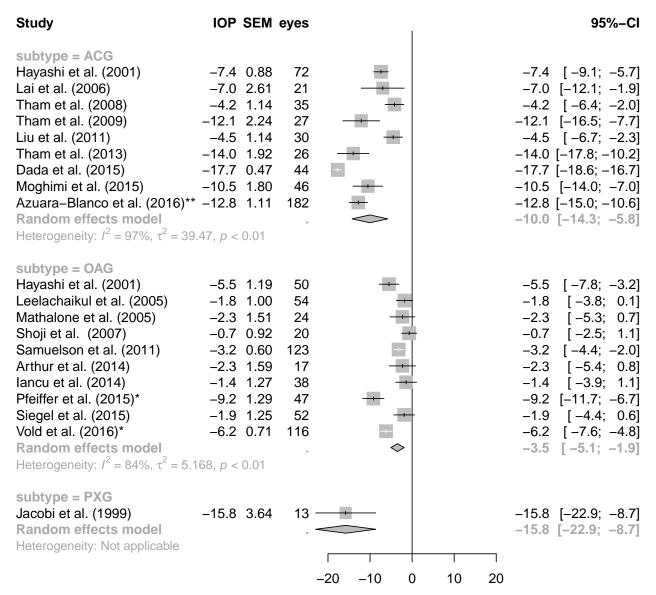
One year follow-up

```
for(mmhg.per.med in mmhg.per.meds) {
 df_ <- df %>%
   filter(!is.na(RxChangeMean),
           df$subtype != "acute",
           MIGsYorN == 'N',
           !is.na(OneYAbsIOPChangeStdDev)) %>%
   mutate(subtype=factor(subtype),
           net.effect=OneYAbsIOPChangeMean +
           mmhg.per.med * RxChangeMean * (washout.type %in% c('None', 'Partial')) +
           mmhg.per.med * RxPostOpMean * (washout.type %in% c('Pre')),
           net.sem=2 * OneYAbsIOPChangeStdDev / sqrt(OneYEyes))
 m <- metagen(net.effect,</pre>
               net.sem,
               study.name,
               data=df_,
               byvar=subtype,
               n.e=LastPeriodEyes)
 forest(m,
         comb.fixed=FALSE,
         digits=1,
        digits.se = 2,
         overall=FALSE,
         leftcols=c("studlab", "TE", "seTE", "n.e"),
        leftlabs=c("Study", "IOP", "SEM", "eyes"))
  grid.text(paste0("Simulated net change in IOP at one year, ", mmhg.per.med, " mm Hg per med"), .5, .9
 print(" ")
```

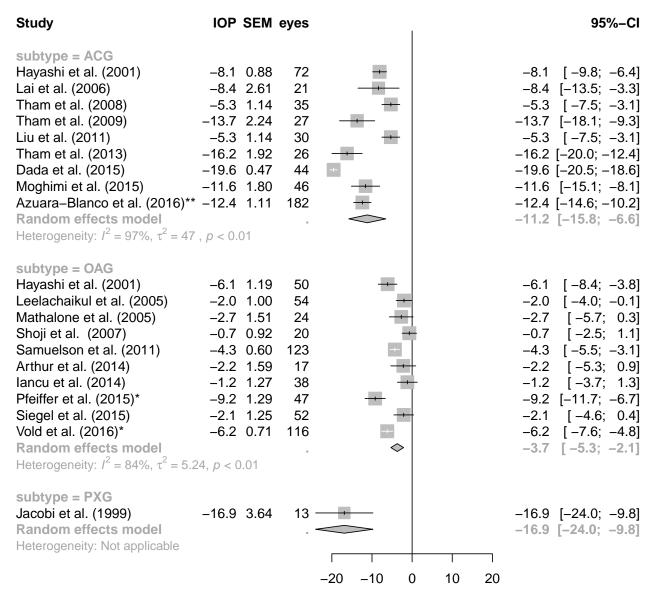
Simulated net change in IOP at one year, 0 mm Hg per med



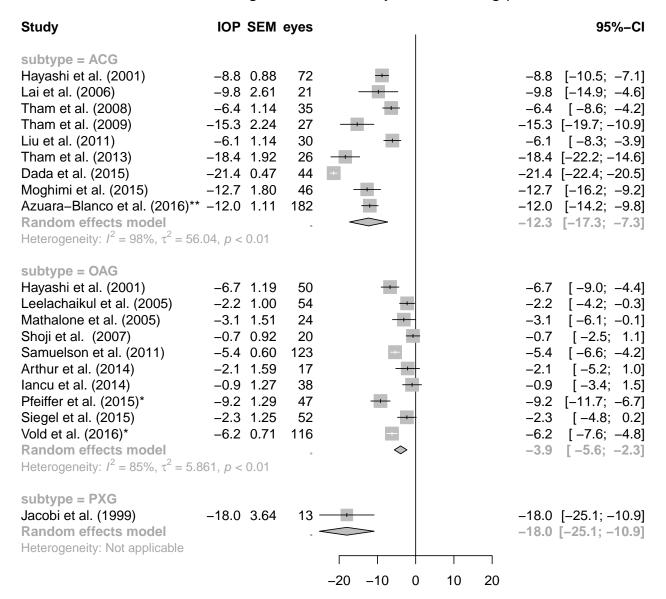
Simulated net change in IOP at one year, 2 mm Hg per med



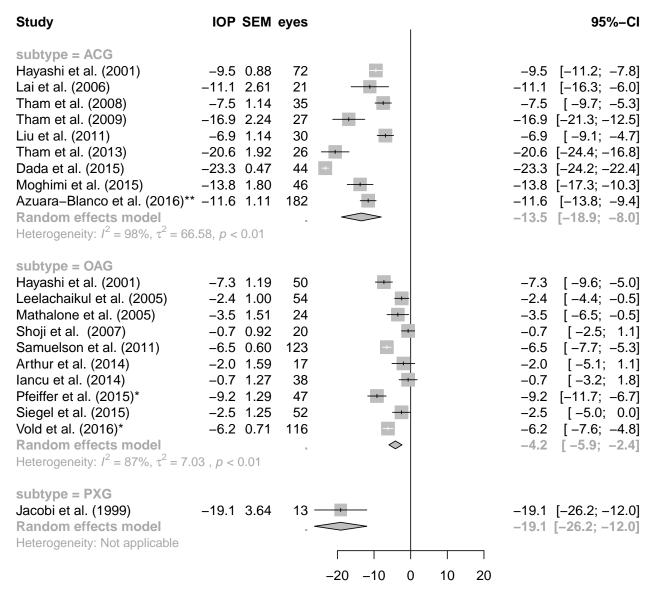
Simulated net change in IOP at one year, 3 mm Hg per med



Simulated net change in IOP at one year, 4 mm Hg per med



Simulated net change in IOP at one year, 5 mm Hg per med



[1] " "

Conclusion

If you take into account the drop in number of meds in each study, you get an additional ~1mmHg drop in the OAG group. That's a net effect around 4 mmHg drop at 12 months and in the last period: quite a bit larger than the uncorrected estimates we had previously.

There's another source of bias we can't correct for here: in some studies, less potent medicines with better side effect profiles might have replaced more potent medicines in the post period, in response to better control of IOP after surgery. That would be measured as a net change of 0 between RxPreOp and RxPostOp, and it would result in an apparent increase in IOP, even though that's a net positive for the patient. I'll leave it to the clinician to determine how likely this is.