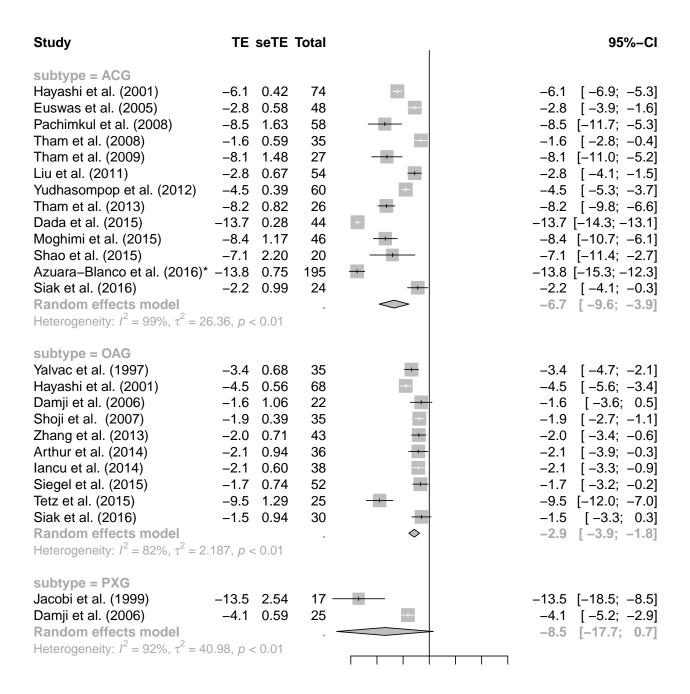
Phaco meta analysis

Patrick Mineault
14/02/2017

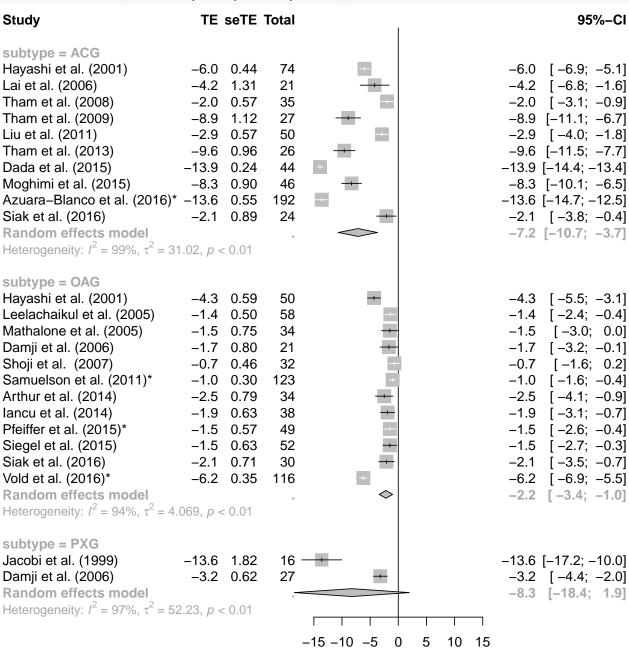
Analysis of full dataset

6 month follow-up



12-month follow up

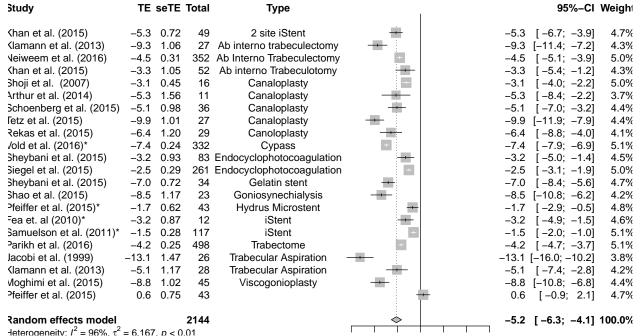
```
digits.se = 2,
overall=FALSE,
leftcols=c("studlab", "TE", "seTE", "n.e"))
```



Last period

Study	TE seTE	Total	1	95%-CI
subtype = ACG Hayashi et al. (2001) Mierzejewski et al. (2008) Tham et al. (2008) Tham et al. (2009) Liu et al. (2011) Shams et al. (2012) Tham et al. (2013) Dias-Santos (2015) Azuara-Blanco et al. (2016)* Lee et al. (2016) Random effects model Heterogeneity: I ² = 97%, τ ² = 18	-1.7 0.60	72 25 35 27 30 55 26 15 182 56		-7.2 [-8.0; -6.4] -5.1 [-7.0; -3.2] -1.8 [-3.0; -0.6] -8.3 [-10.6; -6.0] -3.8 [-5.2; -2.4] -4.5 [-5.9; -3.1] -8.4 [-10.7; -6.1] -5.4 [-9.4; -1.4] -12.9 [-14.0; -11.8] -1.7 [-2.9; -0.5] -5.9 [-8.4; -3.4]
subtype = OAG Kim et al. (1999) Hayashi et al. (2001) Leelachaikul et al. (2005) Mathalone et al. (2005) Damji et al. (2006) Shingleton et al. (2006) Shoji et al. (2007) Mierzejewski et al. (2008) Fea et. al (2010)* Samuelson et al. (2011)* Arthur et al. (2014) Pfeiffer et al. (2015) Siegel et al. (2015) Vold et al. (2016)* Random effects model Heterogeneity: $l^2 = 91\%$, $\tau^2 = 3.5$	-2.9 0.62 -5.3 0.68 -1.6 0.57 -1.9 1.00 -1.5 0.86 -1.4 0.44 -1.0 0.67 -4.0 0.56 -1.6 0.70 -1.0 0.30 -2.1 1.20 0.3 0.69 -2.2 0.64 -5.4 0.36	31 50 54 24 25 20 52 21 123 17 47 52 116	* * * * * * * * * * * * * * * * * * * *	-2.9 [-4.1; -1.7] -5.3 [-6.6; -4.0] -1.6 [-2.7; -0.5] -1.9 [-3.9; 0.1] -1.5 [-3.2; 0.1] -1.4 [-2.3; -0.5] -1.0 [-2.3; 0.3] -4.0 [-5.1; -2.9] -1.6 [-3.0; -0.2] -1.0 [-1.6; -0.4] -2.1 [-4.4; 0.2] 0.3 [-1.0; 1.6] -2.2 [-3.5; -0.9] -5.4 [-6.1; -4.7] -2.3 [-3.3; -1.3]
subtype = PXG Jacobi et al. (1999) Damji et al. (2006) Mierzejewski et al. (2008) Shingleton et al. (2008) Random effects model Heterogeneity: $I^2 = 95\%$, $\tau^2 = 13$	-14.0 2.04 -3.1 0.78 -6.6 0.81 -1.1 0.59 3.25, <i>p</i> < 0.01	13 · 27 23 51	-15 -10 -5 0 5 10 15	-14.0 [-18.0; -10.0] -3.1 [-4.7; -1.6] -6.6 [-8.2; -5.0] -1.1 [-2.3; 0.1] -5.8 [-9.5; -2.0]

MIGS

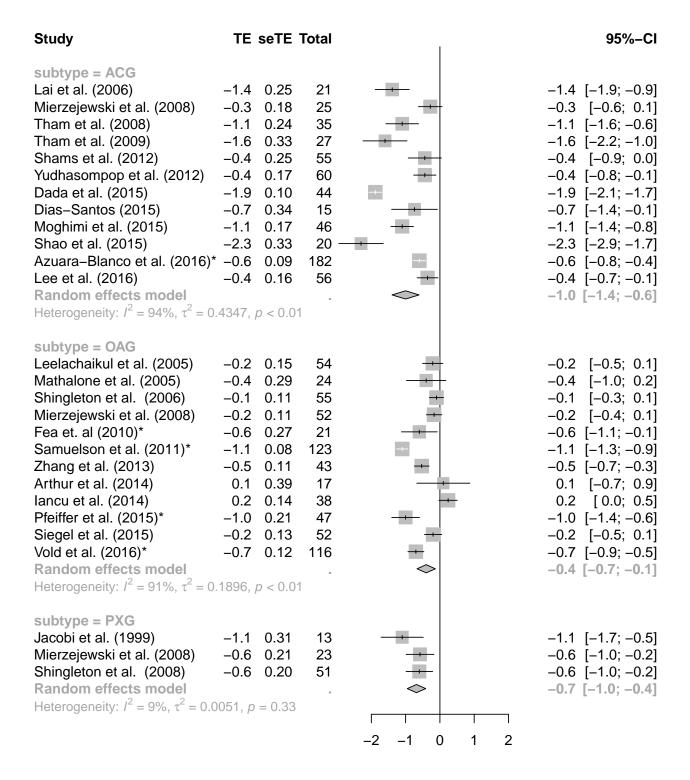


Acute

```
n.e=SixMoEyes, comb.fixed = FALSE)
print(m)
cat("One year: \n")
cat("=====\n")
df_ <- df %>%
 filter(!is.na(df$OneYAbsIOPChangeMean), df$subtype == "acute", MIGsYorN == 'N') %>%
 mutate(subtype=factor(subtype)) %>% dplyr::arrange(Year)
m <- metagen(OneYAbsIOPChangeMean,</pre>
         OneYAbsIOPChangeStdDev / sqrt(OneYEyes),
         study.name,
         data=df_,
         n.e=OneYEyes, comb.fixed = FALSE)
print(m)
cat("-----\n")
cat("Last period: \n")
cat("-----\n")
df <- df %>%
 filter(!is.na(df$LastPeriodAbsIOPChangeMean), df$subtype == "acute", MIGsYorN == 'N') %>%
 mutate(subtype=factor(subtype)) %>% dplyr::arrange(Year)
m <- metagen(LastPeriodAbsIOPChangeMean,
         LastPeriodAbsIOPChangeStdDev / sqrt(LastPeriodEyes),
         study.name,
         data=df ,
         n.e=LastPeriodEyes, comb.fixed = FALSE)
print(m)
## Six months:
                                 95%-CI %W(random)
## Lam et al. (2008) -47.0000 [-51.0521; -42.9479]
                                         50.2
## Hou et al. (2015) -38.2000 [-42.4159; -33.9841]
                                           49.8
## Number of studies combined: k = 2
##
##
                                  95%-CI
                                          z p-value
## Random effects model -42.62 [-51.2438; -33.9963] -9.69 < 0.0001
## Quantifying heterogeneity:
## tau^2 = 34.2694; H = 2.95; I^2 = 88.5%;
## Rb = 88.5\%
##
## Test of heterogeneity:
     Q d.f. p-value
          0.0032
## 8.70
       1
##
## Details on meta-analytical method:
## - Inverse variance method
## - DerSimonian-Laird estimator for tau^2
## One year:
```

```
##
                                       95%-CI %W(random)
## Lam et al. (2008) -47.6000 [-50.4731; -44.7269]
## Hou et al. (2015) -35.9600 [-38.9540; -32.9660]
                                                   49.9
## Number of studies combined: k = 2
##
##
                                          95%-CI
                                                    z p-value
## Random effects model -41.7879 [-53.1949; -30.3809] -7.18 < 0.0001
##
## Quantifying heterogeneity:
## tau^2 = 65.5036; H = 5.50; I^2 = 96.7\%;
## Rb = 96.7\%
##
## Test of heterogeneity:
##
      Q d.f. p-value
## 30.23 1 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - DerSimonian-Laird estimator for tau^2
## Last period:
##
                                          95%-CI %W(random)
## Jacobi et al. (2002) -22.7000 [-23.6565; -21.7435]
## Lam et al. (2008)
                     -47.1000 [-50.0449; -44.1551]
                                                      25.3
                     -35.8000 [-39.5586; -32.0414]
## Lee et al. (2010)
                                                      25.1
## Husain et al. (2012) -44.5000 [-51.8668; -37.1332]
                                                      24.0
## Number of studies combined: k = 4
##
##
                                          95%-CI
                                                    z p-value
## Random effects model -37.3974 [-51.7129; -23.0820] -5.12 < 0.0001
## Quantifying heterogeneity:
## tau^2 = 208.4467; H = 9.89 [8.06; 12.13]; I^2 = 99.0% [98.5%; 99.3%];
## Rb = 97.7\% [93.1%; 100.0%]
##
## Test of heterogeneity:
       Q d.f. p-value
            3 < 0.0001
## 293.22
## Details on meta-analytical method:
## - Inverse variance method
## - DerSimonian-Laird estimator for tau^2
```

Meds



Correlation between meds and drop in IOP

How is IOP drop related to change in meds? Two hypotheses:

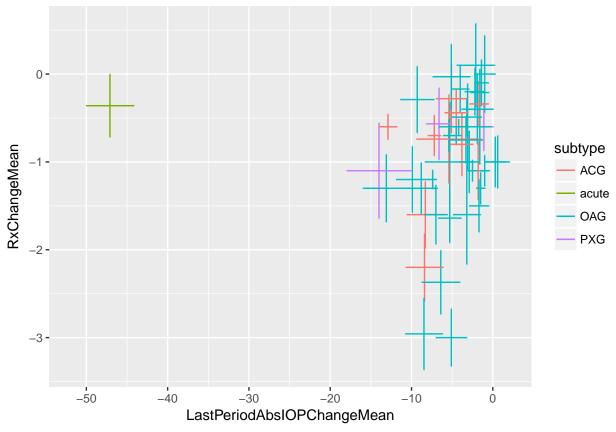
• Those studies that see the largest IOP drops also have drop in meds, as doctors see that can use the newfound slack to decrease the number of meds people take

• The studies that see the largest IOP drops are those that don't change meds, because dropping meds would also increase IOP

So which is it?

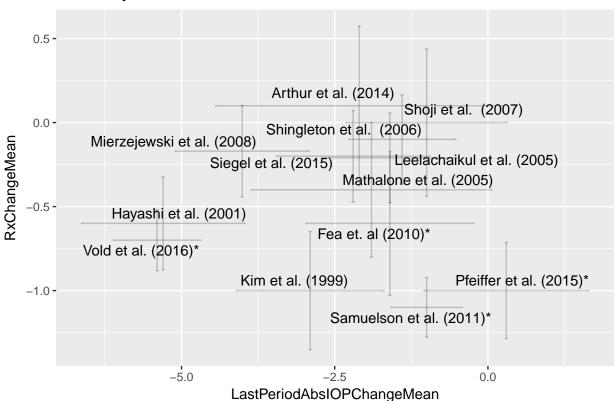
Warning: Removed 18 rows containing missing values (geom_errorbar).

Warning: Removed 18 rows containing missing values (geom_errorbarh).





OAG only



There is an apparent positive correlation between the two effect sizes: studies with larger drops in IOP also tend to see larger drops in Rx. This is clearer when we reject the studies with washout.

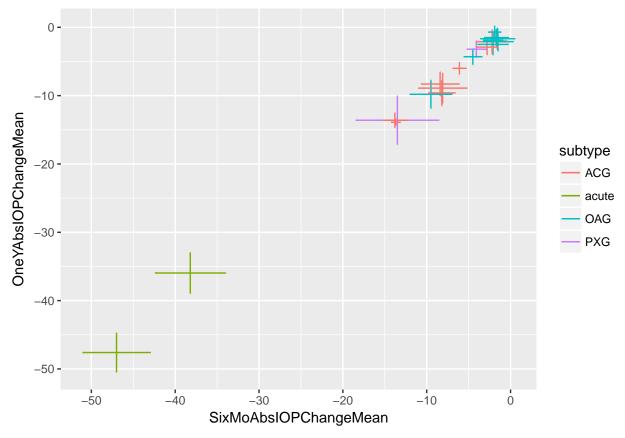
```
draw.corr <- function(x.mean, x.sem, y.mean, y.sem) {</pre>
  d_ <- data.frame(x = rnorm(n = length(x.mean), mean=x.mean, sd = x.sem),</pre>
                  y = rnorm(n = length(x.mean), mean=y.mean, sd = y.sem))
  with(d_ %>% filter(!is.na(x) & !is.na(y)), cor(x, y))
}
cat("Mean +- SE correlation, all studies\n")
## Mean +- SE correlation, all studies
df_ <- df %>% filter(subtype=='OAG', MIGsYorN == 'N')
drawn.corrs <- with(df_, replicate(n = 100,</pre>
                                    draw.corr(LastPeriodAbsIOPChangeMean,
                                               LastPeriodAbsIOPChangeStdDev / sqrt(LastPeriodEyes),
                                               RxPostOpMean - RxPreOpMean,
                                               1 / sqrt(LastPeriodEyes))))
mean(drawn.corrs)
## [1] -0.001909355
sd(drawn.corrs)
## [1] 0.1418261
```

Joint inferences

Measure the correlation between different outcomes (IOP at 6 months vs. 12 months).

```
## Warning: Removed 46 rows containing missing values (geom_errorbar).
```

Warning: Removed 46 rows containing missing values (geom_errorbarh).



It's very clear that six months and 12 months IOP are highly correlated.

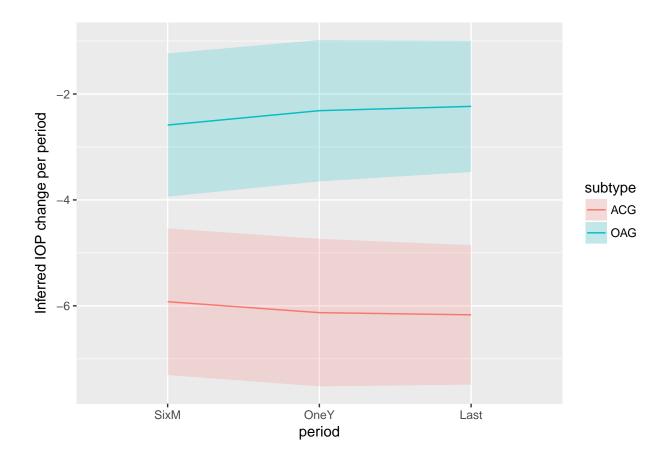
Mean +- SE correlation, All subtypes

```
df_ <- df %>% filter(subtype=='OAG')
drawn.corrs <- with(df_, replicate(n = 100,</pre>
                                    draw.corr(SixMoAbsIOPChangeMean,
                                              SixMoAbsIOPChangeStdDev / sqrt(SixMoEyes),
                                              OneYAbsIOPChangeMean,
                                              OneYAbsIOPChangeStdDev / sqrt(OneYEyes))))
cat("Mean +- SE correlation, OAG only\n")
## Mean +- SE correlation, OAG only
print(mean(drawn.corrs))
## [1] 0.9298547
print(sd(drawn.corrs))
## [1] 0.04253995
df <- df
drawn.corrs <- with(df_, replicate(n = 100,</pre>
                                    draw.corr(SixMoAbsIOPChangeMean,
                                              SixMoAbsIOPChangeStdDev / sqrt(SixMoEyes),
                                              OneYAbsIOPChangeMean,
                                              OneYAbsIOPChangeStdDev / sqrt(OneYEyes))))
cat("Mean +- SE correlation, All subtypes\n")
```

```
print(mean(drawn.corrs))
## [1] 0.9917813
print(sd(drawn.corrs))
## [1] 0.003656665
That's really high. Let's use mymeta to infer the effect size for all periods together.
library(mvmeta)
## This is mvmeta 0.4.7. For an overview type: help('mvmeta-package').
fill.na <- function(x, y, z) {
  return(ifelse(!is.na(x),
                   ifelse(is.na(y),
                           z,
                           ifelse(is.na(z),
                             sqrt((y**2 + z**2) / 2 )))))
}
get.correlation.matrices.tri <- function(x, y, z, assumed.rho) {</pre>
  S <- list()
  for(i in 1:length(x)) {
    xx <- fill.na(x[i], y[i], z[i])</pre>
    yy <- fill.na(y[i], x[i], z[i])</pre>
    zz <- fill.na(z[i], x[i], y[i])</pre>
    S[[i]] \leftarrow matrix(c(xx ** 2, xx * yy * assumed.rho, xx * zz * assumed.rho ** 2,
                        xx * yy * assumed.rho, yy ** 2, zz * yy* assumed.rho,
                        xx * zz * assumed.rho ** 2, zz * yy * assumed.rho, zz * zz), ncol=3)
 }
 S
}
df_ <- df %>% filter(!is.na(LastPeriodAbsIOPChangeStdDev) |
                      !is.na(SixMoAbsIOPChangeStdDev) |
                      !is.na(OneYAbsIOPChangeStdDev), subtype %in% c('OAG', 'ACG'), MIGsYorN == 'N')
thefit <- mvmeta(cbind(SixMoAbsIOPChangeMean, OneYAbsIOPChangeMean, LastPeriodAbsIOPChangeMean) ~ subty
       S=get.correlation.matrices.tri(df_$SixMoAbsIOPChangeStdDev / sqrt(df_$SixMoEyes),
                                       df_$OneYAbsIOPChangeStdDev / sqrt(df_$OneYEyes),
                                       df_$LastPeriodAbsIOPChangeStdDev / sqrt(df_$LastPeriodEyes), .7),
       data=df_,
       method="reml")
summary(thefit)
## Call: mvmeta(formula = cbind(SixMoAbsIOPChangeMean, OneYAbsIOPChangeMean,
       LastPeriodAbsIOPChangeMean) ~ subtype, S = get.correlation.matrices.tri(df_$SixMoAbsIOPChangeStd
##
##
       df_$OneYAbsIOPChangeStdDev/sqrt(df_$OneYEyes), df_$LastPeriodAbsIOPChangeStdDev/sqrt(df_$LastPer
       0.7), data = df , method = "reml")
##
## Multivariate random-effects meta-regression
```

```
## Dimension: 3
## Estimation method: REML
##
## Fixed-effects coefficients
## SixMoAbsIOPChangeMean :
                Estimate Std. Error
                                             z Pr(>|z|) 95%ci.lb
                                                                     95%ci.ub
##
                 -5.9217
                               0.7061 -8.3865
                                                   0.0000
                                                            -7.3057
                                                                       -4.5378
## (Intercept)
                               0.9870
                                                                       5.2699
## subtypeOAG
                  3.3355
                                        3.3795
                                                   0.0007
                                                             1.4011
##
## (Intercept)
## subtypeOAG
## OneYAbsIOPChangeMean :
                                             z Pr(>|z|)
                Estimate
                          Std. Error
                                                           95%ci.lb
                                                                     95%ci.ub
                                                                       -4.7365
## (Intercept)
                 -6.1285
                               0.7102 - 8.6292
                                                   0.0000
                                                            -7.5205
## subtypeOAG
                  3.8138
                               0.9841
                                        3.8754
                                                   0.0001
                                                             1.8850
                                                                       5.7425
##
## (Intercept)
## subtypeOAG
## LastPeriodAbsIOPChangeMean :
                Estimate Std. Error
                                             z Pr(>|z|)
                                                           95%ci.lb
                                                                     95%ci.ub
## (Intercept)
                 -6.1695
                              0.6717 -9.1853
                                                   0.0000
                                                            -7.4860
                                                                       -4.8531
## subtypeOAG
                  3.9360
                               0.9221
                                        4.2683
                                                   0.0000
                                                             2.1286
                                                                       5.7433
##
## (Intercept)
## subtypeOAG
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Between-study random-effects (co)variance components
    Structure: General positive-definite
##
                                Std. Dev
                                                            Corr
## SixMoAbsIOPChangeMean
                                  2.7657
                                          SixMoAbsIOPChangeMean
## OneYAbsIOPChangeMean
                                  2.7806
                                                          0.9963
## LastPeriodAbsIOPChangeMean
                                  2.5526
                                                          0.9813
                                OneYAbsIOPChangeMean
## SixMoAbsIOPChangeMean
## OneYAbsIOPChangeMean
## LastPeriodAbsIOPChangeMean
                                              0.9677
##
## Multivariate Cochran Q-test for residual heterogeneity:
## Q = 1537.5043 (df = 62), p-value = 0.0000
## I-square statistic = 96.0%
## 36 studies, 68 observations, 6 fixed and 6 random-effects parameters
      logLik
                    AIC
                                BIC
## -118.6183
               261.2367
                           286.7623
newdata <- data.frame(subtype=c('OAG', 'ACG'))</pre>
pred <- predict(thefit, newdata, se=TRUE)</pre>
newdata$SixMoAbsIOPChangeMean <- pred$fit[,1]</pre>
newdata$OneYAbsIOPChangeMean <- pred$fit[,2]</pre>
newdata$LastPeriodAbsIOPChangeMean <- pred$fit[,3]</pre>
newdata$SixMoAbsIOPChangeSEM <- pred$se[,1]</pre>
newdata$OneYAbsIOPChangeSEM <- pred$se[,2]</pre>
```

```
newdata$LastPeriodAbsIOPChangeSEM <- pred$se[,3]</pre>
library(reshape2)
##
## Attaching package: 'reshape2'
## The following object is masked from 'package:tidyr':
##
##
       smiths
nd <- melt(newdata)</pre>
## Using subtype as id variables
nd$period <- substr(nd$variable, 0, 4)</pre>
nd$metric <- substr(nd$variable, nchar(as.character(nd$variable)) - 3, nchar(as.character(nd$variable))</pre>
##
      subtype
                                 variable
                                               value period metric
## 1
          OAG
                   SixMoAbsIOPChangeMean -2.5862365
                                                       SixM
                                                               Mean
## 2
          ACG
                   SixMoAbsIOPChangeMean -5.9217293
                                                       SixM
                                                               Mean
## 3
          OAG
                    OneYAbsIOPChangeMean -2.3147326
                                                       OneY
                                                               Mean
## 4
          ACG
                    OneYAbsIOPChangeMean -6.1284889
                                                       OneY
                                                               Mean
## 5
          OAG LastPeriodAbsIOPChangeMean -2.2335730
                                                       Last
                                                               Mean
## 6
          ACG LastPeriodAbsIOPChangeMean -6.1695441
                                                       Last
                                                               Mean
## 7
          OAG
                    SixMoAbsIOPChangeSEM 0.6895898
                                                       SixM
                                                               eSEM
## 8
          ACG
                    SixMoAbsIOPChangeSEM 0.7061031
                                                       SixM
                                                               eSEM
## 9
          OAG
                     OneYAbsIOPChangeSEM 0.6811988
                                                       OneY
                                                               eSEM
## 10
          ACG
                     OneYAbsIOPChangeSEM 0.7102074
                                                       OneY
                                                               eSEM
## 11
          OAG LastPeriodAbsIOPChangeSEM 0.6318083
                                                       Last
                                                               eSEM
## 12
          ACG LastPeriodAbsIOPChangeSEM 0.6716733
                                                       Last
                                                               eSEM
df_ <- dcast(nd, formula = subtype + period ~ metric)</pre>
df_$period <- factor(df_$period, c('SixM', 'OneY', 'Last'))</pre>
ggplot(df_, aes(x=period,
                ymin=Mean - 1.96*eSEM,
                ymax=Mean + 1.96*eSEM,
                group=subtype,
                fill=subtype)) + geom_ribbon(alpha=.2) + geom_line(aes(color=subtype)) + ylab("Inferred
```



Meta-regression

Consider relationships between different covariates and outcomes. Focus on the IOP drop at one year and its correlation with different factors.

```
df_ <- df %>% filter(!is.na(df$OneYAbsIOPChangeMean), df$subtype != "acute", MIGsYorN == 'N') %>%
  mutate(subtype=relevel(factor(subtype), ref="OAG"))
m <- metagen(OneYAbsIOPChangeMean,
             OneYAbsIOPChangeStdDev / sqrt(OneYEyes),
             study.name,
             data=df_,
             byvar=subtype,
             n.e=OneYEyes)
print(metareg(~ OneYEyes, x=m))
##
## Mixed-Effects Model (k = 24; tau^2 estimator: DL)
##
## tau^2 (estimated amount of residual heterogeneity):
                                                            30.8712 (SE = 12.8681)
## tau (square root of estimated tau^2 value):
                                                            5.5562
## I^2 (residual heterogeneity / unaccounted variability): 99.02%
## H^2 (unaccounted variability / sampling variability):
                                                            102.25
## R^2 (amount of heterogeneity accounted for):
                                                            0.00%
##
## Test for Residual Heterogeneity:
```

```
## QE(df = 22) = 2249.5296, p-val < .0001
##
## Test of Moderators (coefficient(s) 2):
## QM(df = 1) = 0.9026, p-val = 0.3421
## Model Results:
##
##
            estimate
                          se
                                  zval
                                         pval
                                                 ci.lb
                                                         ci.ub
## intrcpt
             -3.3325 1.8721 -1.7801 0.0751
                                               -7.0019
                                                        0.3368
## OneYEyes
             -0.0275 0.0290 -0.9500 0.3421
                                               -0.0843 0.0293
##
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
print(metareg(~ OneYEyes * subtype, x=m))
##
## Mixed-Effects Model (k = 24; tau^2 estimator: DL)
## tau^2 (estimated amount of residual heterogeneity):
                                                          17.0380 (SE = 7.3921)
## tau (square root of estimated tau^2 value):
                                                          4.1277
## I^2 (residual heterogeneity / unaccounted variability): 98.16%
## H^2 (unaccounted variability / sampling variability):
                                                          54.33
## R^2 (amount of heterogeneity accounted for):
                                                          39.83%
## Test for Residual Heterogeneity:
## QE(df = 18) = 977.8565, p-val < .0001
## Test of Moderators (coefficient(s) 2,3,4,5,6):
## QM(df = 5) = 14.6923, p-val = 0.0118
##
## Model Results:
##
                                                              ci.lb
                                                                       ci.ub
                       estimate
                                             zval
                                                     pval
                                      se
## intrcpt
                        -1.1839
                                  2.3639
                                          -0.5008 0.6165
                                                            -5.8170
                                                                      3.4491
                                                   0.6202
                                                            -0.0936
                                                                      0.0558
## OneYEyes
                        -0.0189
                                  0.0381
                                          -0.4956
## subtypeACG
                                  3.0929
                                          -1.1585
                                                   0.2467
                                                            -9.6452
                                                                       2.4789
                        -3.5831
## subtypePXG
                        -27.5433 12.8476
                                          -2.1438
                                                   0.0320
                                                           -52.7243 -2.3624
                        -0.0253
                                         -0.5397
                                                   0.5894
## OneYEyes:subtypeACG
                                 0.0469
                                                            -0.1172
                                                                      0.0666
                         0.9644
## OneYEyes:subtypePXG
                                  0.5600
                                          1.7220 0.0851
                                                            -0.1333
                                                                      2.0620
##
## intrcpt
## OneYEyes
## subtypeACG
## subtypePXG
## OneYEyes:subtypeACG
## OneYEyes:subtypePXG
##
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
print(metareg(~ Year, x=m))
## Mixed-Effects Model (k = 24; tau^2 estimator: DL)
```

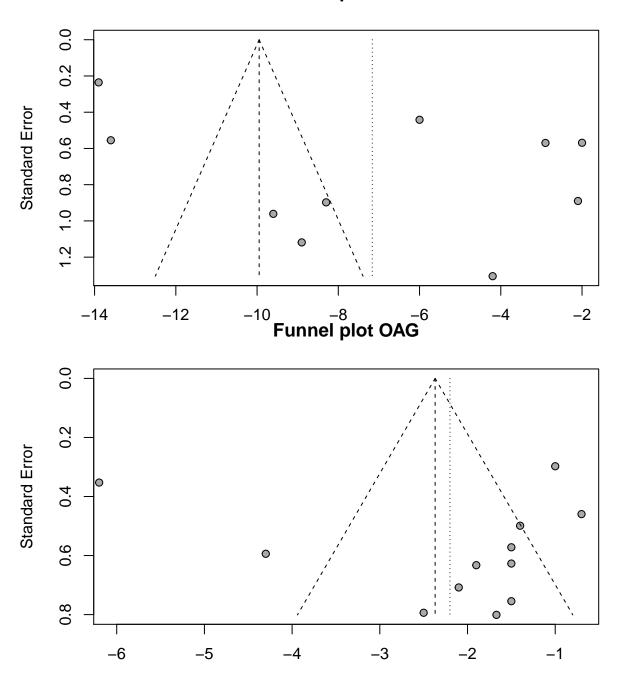
```
##
## tau^2 (estimated amount of residual heterogeneity):
                                                           25.2200 \text{ (SE = } 10.5747)
## tau (square root of estimated tau^2 value):
## I^2 (residual heterogeneity / unaccounted variability): 98.82%
## H^2 (unaccounted variability / sampling variability):
                                                           84.52
## R^2 (amount of heterogeneity accounted for):
                                                           10.94%
## Test for Residual Heterogeneity:
## QE(df = 22) = 1859.3429, p-val < .0001
## Test of Moderators (coefficient(s) 2):
## QM(df = 1) = 0.0302, p-val = 0.8621
## Model Results:
##
##
            estimate
                                   zval
                                                     ci.lb
                                                               ci.ub
                            se
                                           pval
                                                 -707.3556
                                                            834,4924
## intrcpt
             63.5684
                     393.3358
                                 0.1616 0.8716
             -0.0340
                        0.1957 -0.1736 0.8621
                                                   -0.4175
                                                              0.3496
##
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
print(metareg(~ Year * subtype, x=m))
## Mixed-Effects Model (k = 24; tau^2 estimator: DL)
## tau^2 (estimated amount of residual heterogeneity):
                                                           11.2606 (SE = 4.8038)
## tau (square root of estimated tau^2 value):
                                                           3.3557
## I^2 (residual heterogeneity / unaccounted variability): 97.24%
## H^2 (unaccounted variability / sampling variability):
## R^2 (amount of heterogeneity accounted for):
                                                           60.23%
##
## Test for Residual Heterogeneity:
## QE(df = 18) = 651.7540, p-val < .0001
## Test of Moderators (coefficient(s) 2,3,4,5,6):
## QM(df = 5) = 19.9554, p-val = 0.0013
##
## Model Results:
##
                                                      pval
##
                      estimate
                                                                 ci.lb
                                              zval
                                       se
## intrcpt
                       68.7364
                                 390.2739
                                            0.1761 0.8602
                                                             -696.1864
## Year
                      -0.0353
                                   0.1941 -0.1817 0.8558
                                                               -0.4158
## subtypeACG
                      668.7765
                                 603.7615
                                            1.1077 0.2680
                                                             -514.5744
## subtypePXG
                    -3052.2792 1516.1880 -2.0131 0.0441
                                                            -6023.9531
## Year:subtypeACG
                       -0.3350
                                 0.3003 -1.1158 0.2645
                                                               -0.9235
## Year:subtypePXG
                        1.5210
                                   0.7568
                                          2.0097 0.0445
                                                                0.0377
                        ci.ub
## intrcpt
                     833.6592
## Year
                       0.3452
## subtypeACG
                    1852.1273
## subtypePXG
                     -80.6053 *
## Year:subtypeACG
                       0.2535
## Year:subtypePXG
                       3.0043 *
```

```
##
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
print(metareg(~ PreOpIOPMean, x=m))
## Mixed-Effects Model (k = 24; tau^2 estimator: DL)
## tau^2 (estimated amount of residual heterogeneity):
                                                          4.4783 \text{ (SE = 1.8384)}
## tau (square root of estimated tau^2 value):
                                                          2.1162
## I^2 (residual heterogeneity / unaccounted variability): 93.40%
## H^2 (unaccounted variability / sampling variability):
## R^2 (amount of heterogeneity accounted for):
                                                          84.19%
##
## Test for Residual Heterogeneity:
## QE(df = 22) = 333.0958, p-val < .0001
## Test of Moderators (coefficient(s) 2):
## QM(df = 1) = 65.1891, p-val < .0001
## Model Results:
##
##
                estimate
                                     zval
                                             pval
                                                     ci.lb
                                                              ci.ub
## intrcpt
                13.0744 2.2388
                                   5.8399 <.0001
                                                    8.6864 17.4624
                 -0.8673 0.1074 -8.0740 <.0001 -1.0779 -0.6568
## PreOpIOPMean
##
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
print(metareg(~ PreOpIOPMean * subtype, x=m))
## Mixed-Effects Model (k = 24; tau^2 estimator: DL)
##
## tau^2 (estimated amount of residual heterogeneity):
                                                          1.2655 (SE = 0.6125)
## tau (square root of estimated tau^2 value):
                                                          1.1249
## I^2 (residual heterogeneity / unaccounted variability): 79.37%
## H^2 (unaccounted variability / sampling variability):
                                                          4.85
## R^2 (amount of heterogeneity accounted for):
                                                          95.53%
## Test for Residual Heterogeneity:
## QE(df = 18) = 87.2305, p-val < .0001
##
## Test of Moderators (coefficient(s) 2,3,4,5,6):
## QM(df = 5) = 218.5070, p-val < .0001
##
## Model Results:
##
##
                           estimate
                                         se
                                                zval
                                                        pval
                                                                ci.lb
                            5.2227 2.4821
## intrcpt
                                              2.1041 0.0354
                                                               0.3578
## PreOpIOPMean
                            -0.3961 0.1307 -3.0318 0.0024 -0.6522
## subtypeACG
                            8.2042 3.2360
                                              2.5353 0.0112
                                                               1.8617
## subtypePXG
                            8.4784 5.4425
                                              1.5578 0.1193 -2.1888
## PreOpIOPMean:subtypeACG -0.5532 0.1601 -3.4547 0.0006 -0.8671
```

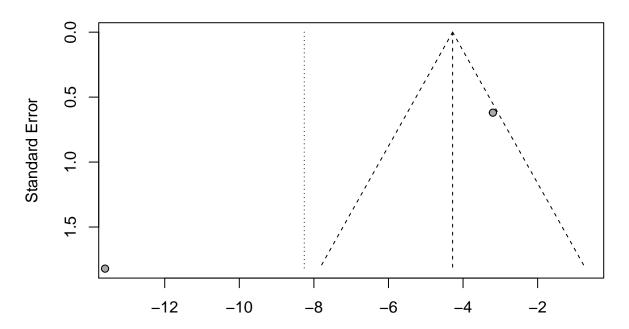
```
## PreOpIOPMean:subtypePXG
                             -0.4570 0.2429 -1.8817 0.0599 -0.9330
##
                              ci.ub
## intrcpt
                            10.0876
## PreOpIOPMean
                            -0.1400
## subtypeACG
                            14.5467
## subtypePXG
                            19.1455
## PreOpIOPMean:subtypeACG
                            -0.2394
## PreOpIOPMean:subtypePXG
                             0.0190
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
# Same, restricted to OAG only
df_ <- df %>% filter(!is.na(df$OneYAbsIOPChangeMean), df$subtype == "OAG", MIGsYorN == 'N') %>%
  mutate(subtype=relevel(factor(subtype), ref="OAG"))
m <- metagen(OneYAbsIOPChangeMean,
             OneYAbsIOPChangeStdDev / sqrt(OneYEyes),
             study.name,
             data=df_,
             byvar=subtype,
             n.e=OneYEyes)
bubble(metareg(~ PreOpIOPMean, x=m))
      T
      7
Treatment effect ()
      က
      -5
      9
                                                             22
           16
                           18
                                            20
                                                                             24
                                   Covariate PreOpIOPMean
```

Small study bias

Funnel plot ACG



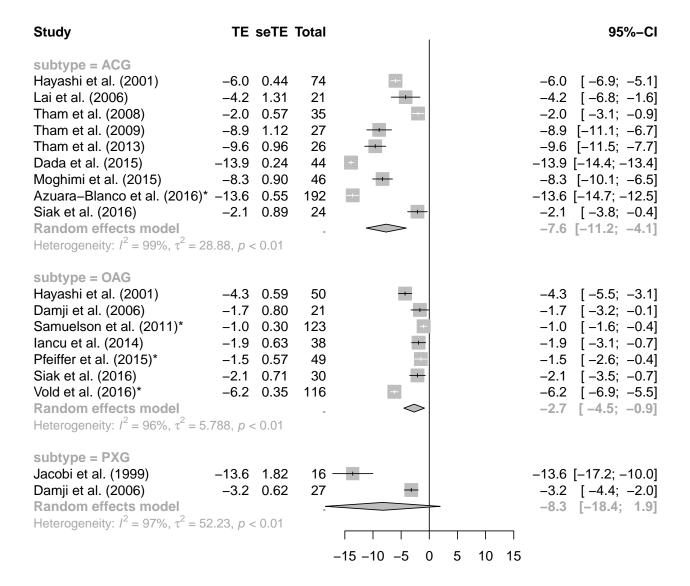
Funnel plot PXG



Alternative filterings of the data

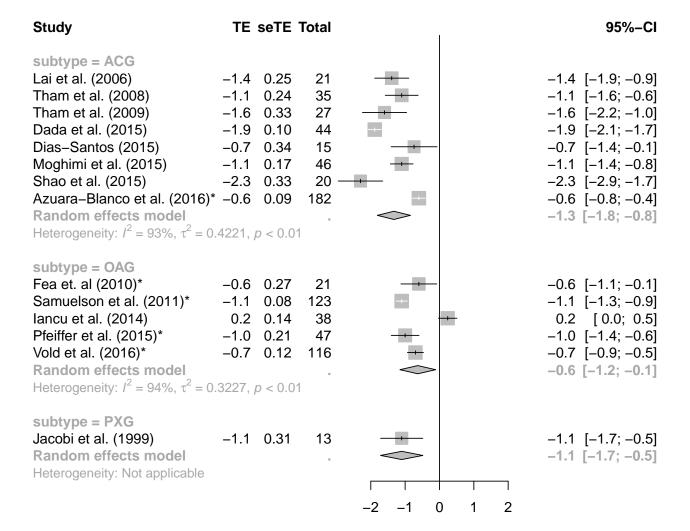
Prospective studies only

One year



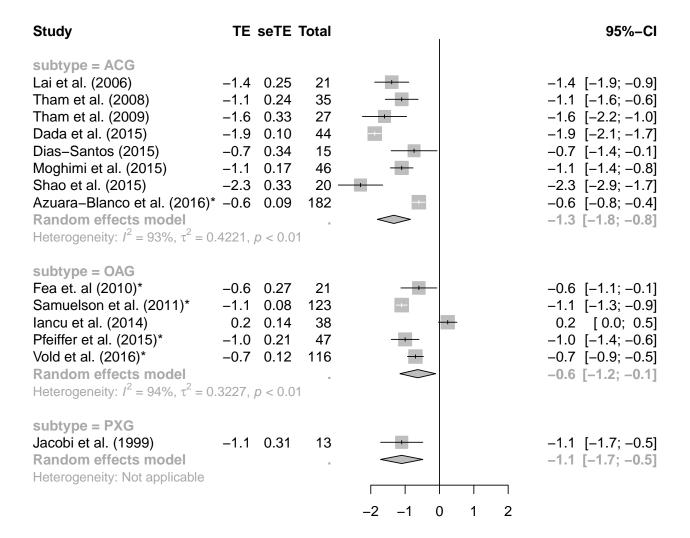
Last period

```
Study
                              TE seTE Total
                                                                                         95%-CI
subtype = ACG
Hayashi et al. (2001)
                             -7.2 0.41
                                          72
                                                                              -7.2 [-8.0; -6.4]
Tham et al. (2008)
                             -1.8 \quad 0.59
                                           35
                                                                              -1.8 [-3.0; -0.6]
                                                                              -8.3 [-10.6; -6.0]
Tham et al. (2009)
                             -8.3 1.16
                                          27
Tham et al. (2013)
                             -8.4 1.18
                                           26
                                                                              -8.4 [-10.7; -6.1]
Dias-Santos (2015)
                             -5.4 2.04
                                           15
                                                                              -5.4 [-9.4; -1.4]
Azuara-Blanco et al. (2016)* -12.9 0.57
                                         182
                                                                             –12.9 [–14.0; –11.8]
Random effects model
                                                                              -7.4 [-11.0; -3.8]
Heterogeneity: I^2 = 97\%, \tau^2 = 18.87, p < 0.01
subtype = OAG
Hayashi et al. (2001)
                             -5.3 0.68
                                          50
                                                                              -5.3 [ -6.6; -4.0]
Damji et al. (2006)
                                           24
                                                                              -1.5 [-3.2; 0.1]
                             -1.5 0.86
Fea et. al (2010)*
                             -1.6 0.70
                                          21
                                                                              -1.6 [-3.0; -0.2]
Samuelson et al. (2011)*
                             -1.0 0.30
                                                                              -1.0 [ -1.6; -0.4]
                                         123
Pfeiffer et al. (2015)*
                              0.3 0.69
                                          47
                                                                               0.3 [-1.0; 1.6]
Vold et al. (2016)*
                             -5.4 0.36
                                         116
                                                                              -5.4 [-6.1; -4.7]
Random effects model
                                                                              -2.4 [ -4.5; -0.4]
Heterogeneity: I^2 = 96\%, \tau^2 = 6.343, p < 0.01
subtype = PXG
Jacobi et al. (1999)
                            -14.0 2.04
                                                                             -14.0 [-18.0; -10.0]
                                           13 -
Damji et al. (2006)
                             -3.1 0.78
                                           27
                                                                              -3.1 [-4.7; -1.6]
Random effects model
                                                                              -8.4 [-19.0; 2.2]
Heterogeneity: I^2 = 96\%, \tau^2 = 56.48, p < 0.01
                                               -15 -10 -5
                                                           0
                                                                5
                                                                    10 15
df_ <- df %>%
  filter(!is.na(RxChangeMean), !is.na(RxChangeStdDev),
         df$subtype != "acute", MIGsYorN == 'N') %>%
  mutate(subtype=factor(subtype))
m <- metagen(RxChangeMean,
             sqrt(RxPostOpStdDev** 2 + RxPreOpStdDev ** 2) / sqrt(LastPeriodEyes),
             study.name,
             data=df ,
             byvar=subtype,
             n.e=LastPeriodEyes)
forest(m,
       comb.fixed=FALSE,
       digits=1,
       digits.se = 2,
       overall=FALSE,
       leftcols=c("studlab", "TE", "seTE", "n.e"))
```



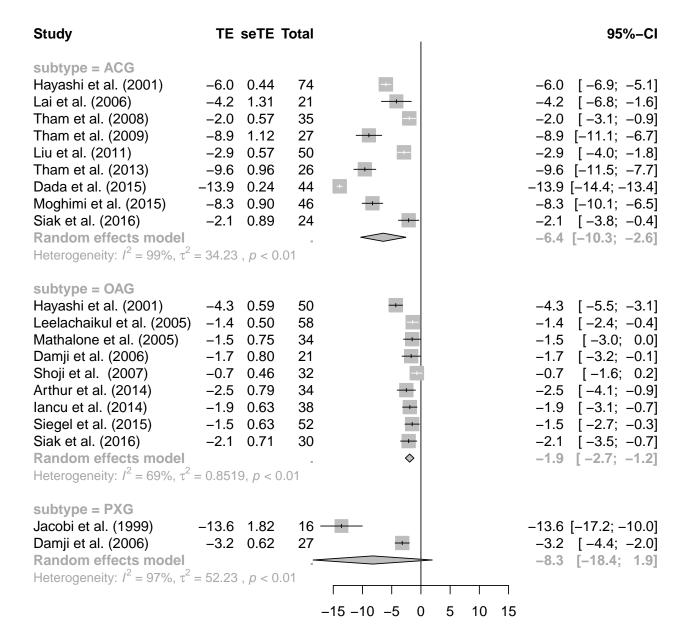
Meds

```
df_ <- df %>%
  filter(!is.na(RxChangeMean), !is.na(RxChangeStdDev),
         df$subtype != "acute", MIGsYorN == 'N') %>%
  mutate(subtype=factor(subtype))
m <- metagen(RxChangeMean,
             sqrt(RxPostOpStdDev** 2 + RxPreOpStdDev ** 2) / sqrt(LastPeriodEyes),
             study.name,
             data=df_,
             byvar=subtype,
             n.e=LastPeriodEyes)
forest(m,
       comb.fixed=FALSE,
       digits=1,
       digits.se = 2,
       overall=FALSE,
       leftcols=c("studlab", "TE", "seTE", "n.e"))
```



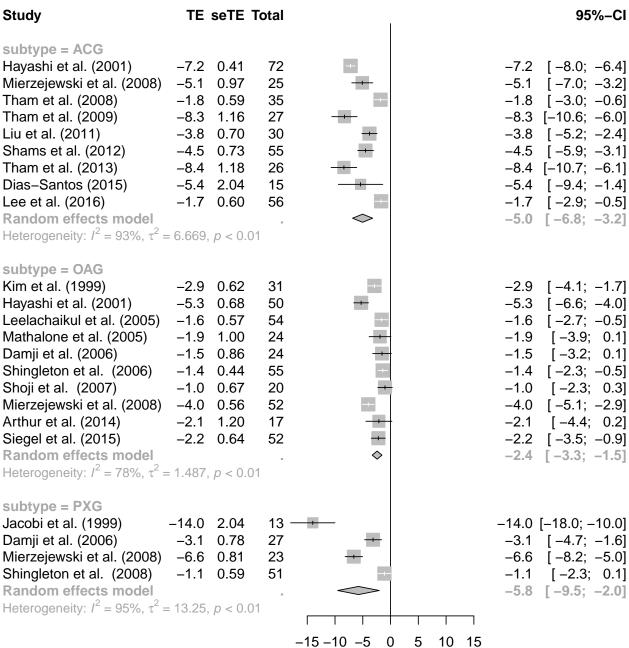
Excluding washout studies

One year



Last period

```
digits.se = 2,
overall=FALSE,
leftcols=c("studlab", "TE", "seTE", "n.e"))
```



Meds

