

Canst Thou Beggar Thy Neighbour?

Evidence from the 1930s

Paul Bouscasse*

University of Cambridge

First version: November 2021

This version: April 2023

Abstract

Does an exchange rate depreciation depress trading partners' output? I address this question through the lens of a classic episode: the currency devaluations of the 1930s. From 1931 to 1936, many of the biggest economies in the world successively left the gold standard or devalued, leading to a depreciation of their currency by more than 30% against gold. In theory, the effect is ambiguous for countries that did not devalue: expenditure switching can lower their output, but the monetary stimulus to demand might raise it. I use cross-sectional evidence to discipline the strength of these two mechanisms in a multi-country model. This evidence comes in two forms: (i) causal inference of the effect of devaluation on country-level variables, (ii) new product-level data to estimate parameters that are essential to discipline the response of trade — the international elasticity of substitution among foreign varieties, and the pass-through of the exchange rate to international prices. Contrary to the popular narrative in modern policy debates, devaluation did not dramatically lower the output of trading partners in this context. The expenditure switching effect was mostly offset by the monetary stimulus to foreign demand.

Keywords: beggar-thy-neighbour, devaluations, Great Depression, trade elasticities

JEL codes: E5, F3, F4

*Faculty of Economics, University of Cambridge. Email address: paul.bouscasse@econ.cam.ac.uk. I thank Jennifer La'O, Jesse Schreger, Jón Steinsson, and Michael Woodford for invaluable guidance and support, as well as Hassan Afrouzi, Michael Bordo, Mario Crucini, Barry Eichengreen, Luca Fornaro, Joshua Hausman, Olivier Jeanne, Loukas Karabarbounis, Arnaud Mehl, Chris Meissner, Kris Mitchener, Eric Monnet, Emi Nakamura, Mathieu Pedemonte, Sarah Quincy, Alessandro Rebucci, Stephanie Schmitt-Grohé, Anson Soderbery, Martín Uribe, François Velde, Chenzi Xu, as well as seminar participants at various venues for comments and suggestions. I am grateful to Olivier Accominotti and Kris Mitchener for sharing data, and to Sarah Wagner for help with Python. Sabrina Chen, Alejandro Diaz Tur, Nino Dickersin, Mariam Khmaladze, and Maria Thompson provided outstanding research assistance. Part of this paper was written while I was a dissertation fellow at the International Finance division of the Federal Reserve Board. I acknowledge financial support from the Program for Economic Research at Columbia University and the Janeway Institute at the University of Cambridge.

1 Introduction

“Beggar thy neighbour, zero-sum game, currency war”: currency depreciation gets a bad rap. In theory, exchange rate depreciation can cause a contraction in foreign output: if prices or costs are sticky in the producer’s currency, a depreciation cheapens domestic goods on international markets, which shifts demand towards those goods, possibly lowering output in foreign countries. This mechanism is at play in prominent international macroeconomic models since at least Fleming (1962) and Mundell (1963). In policy circles, recriminations fly when a country engages in a monetary expansion that leads to a depreciation of the currency. In 2010, Guido Mantega, then Brazil’s finance minister, famously equated the Federal Reserve’s quantitative easing (QE) to monetary warfare.¹ Such claims resurfaced when the European Central Bank took the QE route in 2015.²

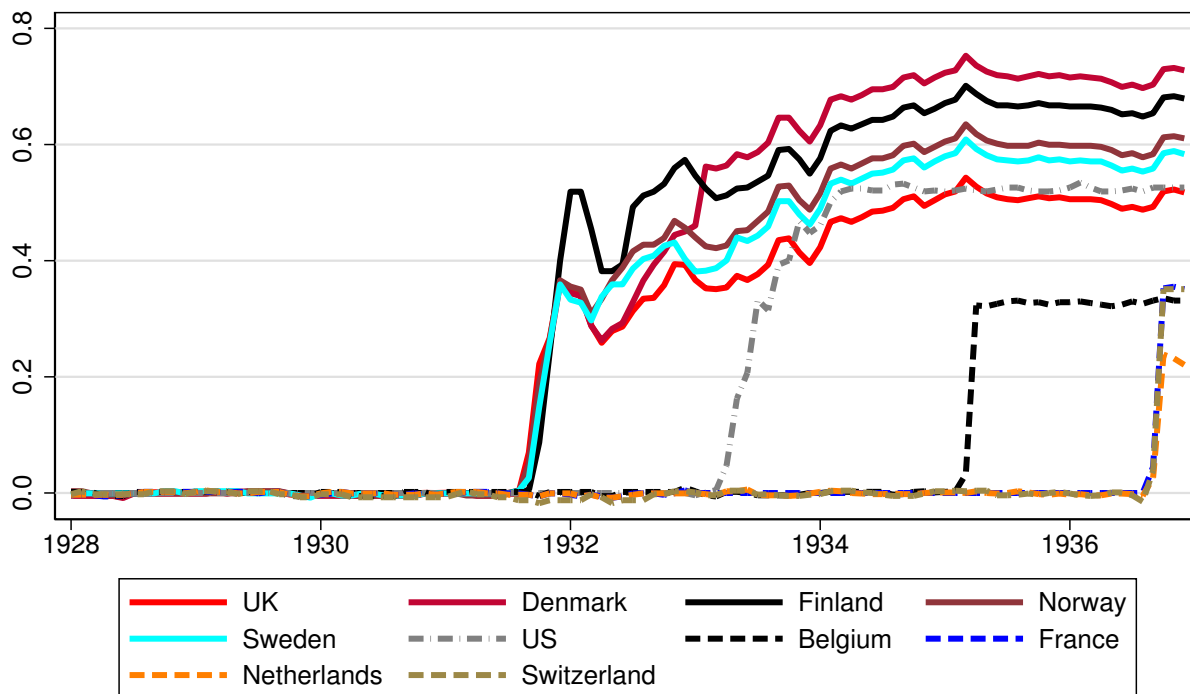
Are these concerns justified? My answer is no: a devaluation has a small effect on foreign output. It is, however, a big stimulus for domestic output. To reach that conclusion, I follow a micro to macro approach: I identify empirical moments in cross-country and new product-level trade data, use these moments to estimate a model, and run counter-factual experiments. To implement this approach, I focus on a specific historical episode: the currency devaluations of the 1930s.

The 1930s currency devaluations are an attractive setting for several reasons. First, they are a classic example of what are generally thought to be competitive devaluations. Second, they form a sequence of large events: many countries, including the biggest economies in the world, devalued their currencies by 30 to 40% against gold. Last but not least, the staggered way in which these countries devalued is ideal for identification. This staggered feature is apparent on figure 1, which shows the nominal price of gold in selected currencies. Until 1931, the countries on the chart were on the gold-exchange standard: the price of gold — hence the exchange rate with other currencies that are on gold — was stable. In September 1931, Britain left the gold standard. It was quickly followed into devaluation by, among others, Scandinavian countries. On the other hand, several countries stayed on the gold standard, sometimes for years: the United States (US) devalued in April 1933; the so-called gold bloc, led by France, only dislocated in October 1936.

¹“We’re in the midst of an international currency war, a general weakening of currency.” See Jonathan Wheatley, “Brazil in ‘currency war’ alert,” *Financial Times*, September 27, 2010, <https://www.ft.com/content/33ff9624-ca48-11df-a860-00144feab49a>.

²David Keohane, “All currency war, all the time”, *Financial Times*, February 5, 2015, <https://www.ft.com/content/c259d418-d0ca-37d5-a3bd-27748fc1024e>.

Figure 1: Gold price in selected currencies



Note: log of the nominal price of gold expressed in country's currency — 1930 is normalized to 0. An increase corresponds to a devaluation.

Such discrepancies in timing create cross-sectional variation that is ripe for identification.

Eichengreen and Sachs (1985) offer the classic treatment of this episode. They show that countries which devalued earlier recovered earlier from the Great Depression. Albeit foundational, this result has several limitations. The first one is that it's not necessarily causal since it is only a regression of the change in industrial production on the change in the exchange rate. If the decision to devalue was endogenous to output, these estimates are biased. Second, even interpreted in a causal way, this evidence only speaks to the relative effect, not the absolute one. Indeed, one does not know if British output was going up compared to French output, because British output was going up compared to a counter-factual scenario where Britain hadn't devalued, or because French output was going down. Finally, as Eichengreen and Sachs (henceforth ES) emphasize, the exchange rate is not the only channel through which devaluations might have affected output. Cutting the gold content of the currency relaxes the gold cover constraint of the central bank. The latter can then expand the money supply or lower its discount rate, thereby stimulating aggregate demand. So the ES regression might capture a closed-economy monetary stimulus, instead of the

effect of a competitive devaluation. Part of my contribution is to overcome these three limitations.

First, I analyze cross-country data to establish a causal relationship between output and devaluation. To do so, I conduct two empirical exercises.

The first one is a difference-in-difference (DD) estimation. Since some countries devalued in 1931 while others stayed on the gold standard, the episode is an ideal setup. Countries which devalued produced more, exported more, and experienced a drop in real wages following the devaluation, relative to countries that did not. They did not import less — in fact, they imported more after a few years — which runs against the beggar-thy-neighbour presumption. Nominal variables such as price and wage indices also increased. On the monetary side, the action did not come from the nominal but from the real interest rate: inflation was higher in devaluing countries, which pushed the real rate down. Finally, devaluing countries do not exhibit preexisting trends in these variables. That they do not rules out some endogeneity concerns — the devaluation decision does not depend on recent economic performance — but it does not rule out all of them. Devaluation may occur in reaction to a contemporaneous shock that would not appear in the preexisting trend but would affect future output. This caveat motivates another identification strategy.

The second empirical exercise relies on high-frequency identification (HFI). I first construct series of policy announcements based on newspaper articles. Most of these events are a devaluation, or an announcement that makes devaluation less likely, such as the fall of an easy money cabinet. I use the change in the forward exchange rate around these announcements as a shock. As long as financial markets anticipate devaluations that systematically correlate with economic conditions, changes in the forward exchange rate around policy announcements are exogenous to those same economic conditions. Armed with these shocks, I can use them as instruments for the spot exchange rate in low-frequency regressions of macroeconomic variables. The identification assumption is conceptually similar to that made by researchers who use changes in federal funds rate futures around Federal Open Market Committee (FOMC) meetings — a well-established identification scheme to infer the causal effect of monetary policy on macroeconomic variables.

These two independent estimation methods show that devaluation stimulated output powerfully. A 30% devaluation increased industrial production by 6% over the following 3 years according to the DD estimates, by 14% according to the HFI ones — compared to a country that did not devalue.

Second, a model is necessary to translate those relative effects into absolute ones. I develop

a multi-country model which features three main ingredients: (i) incomplete pass-through of the exchange rate to international prices (which I show later is an essential feature of the data), (ii) sticky wages to generate monetary non-neutrality, and (iii) a gold standard to engineer the devaluation. The model allows for both expenditure switching and monetary stimulus. In the model, the strength of the expenditure switching channel depends on the elasticity of substitution between domestic and foreign varieties, the elasticity of substitution among imported varieties, and the pass-through of the exchange rate to international prices.

Since those parameters are important, I digitize new product-level data on US imports to estimate the last two — the first one is estimated later out of cross-country data. This data comes from an official publication of the Department of Commerce, the *Foreign Commerce and Navigation of the United States*. Estimating a demand elasticity is always challenging: in general, the price of a good is correlated with demand shocks. An ordinary least square (OLS) regression of quantity on price would suffer from simultaneity bias. The 1930s, however, offer a readily available instrument: the exchange rate. I show that, as long as devaluation is exogenous to demand shocks for different national varieties of the same product, this instrument identifies the within-product elasticity of substitution among foreign goods. I defend this assumption by checking that prices and quantities do not display pre-existing trends. I find a demand elasticity between 2 and 4, and a pass-through of about 0.4. For a one percent decline in their relative price, US imports of British varieties increase by 2–4%, compared to French varieties. For a 1% depreciation of the pound, dollar prices of British products go down by 0.4%.

Third, I estimate the model by matching moments from cross-country and trade data and run counter-factual experiments. I find that the 1930s devaluations had small effects on non-devaluing countries' output. Put another way, British output was going up compared to French output mostly because British output was going up compared to a counter-factual scenario where Britain hadn't devalued — not because French output was going down. An equivalent way to state this conclusion is that the effect estimated in the cross-section is almost entirely attributable to higher output in devaluing countries, not less output in non-devaluing ones: the absolute effect is close to the relative one.

Finally, I clarify how cross-sectional evidence is informative about the absolute effect of a devaluation on devaluing and non-devaluing countries. The absolute effect is the sum of two

terms. One depends on the effect on devaluing countries' output relative to non-devaluing countries, the other one is the absolute effect on world output. The first term is pinned down by cross-sectional evidence, but the second one is absorbed by time fixed effects. Since the model matches cross-sectional evidence, the result depends on the model's predictions for world output. Those predictions in turn rely on certain structural parameters which determine how powerful monetary policy is — expenditure switching is a wash and doesn't matter to world-level variables. The cross-section is informative about those parameters. For instance, the relative path of nominal wages is informative about how sticky they are, which conditions the ability of monetary policy to affect real interest rates. The relative path of output is informative about how it responds to changes in real rates. Overall, parameter estimates point toward a monetary stimulus that is big enough to offset expenditure switching in non-devaluing countries. In devaluing countries, those two forces work together.

Literature review: I contribute to the literature on the Great Depression. Friedman and Schwartz (1963, p. 362) were the first ones to note a cross-country correlation between devaluation and recovery. Building on Eichengreen and Sachs's (1985) contribution, Campa (1990) or Bernanke and Carey (1996) study currency devaluations across countries; Mitchener and Wand-schneider (2015) study capital controls. Eichengreen (1992) analyzes the interwar gold standard. Mathy and Meissner (2011) are interested in how trade and exchange rates affect business cycle co-movement. Accominotti (2009) investigates the credibility of the gold standard in Britain before the devaluation. Cohen-Setton et al. (2017) examine the consequences of supply-side reforms in France — reforms that were contemporaneous to its 1936 devaluation. Hausman et al. (2019) argue that the US devaluation bailed out US farmers through its effect on agricultural prices. de Bromhead et al. (2018) or Albers (2019) delve into tariff policies. Ellison et al. (2021) argue that devaluation affected ex-ante real interest rates. In contemporaneous work, Candia and Pedemonte (2021) are also interested in the spillovers of exchange-rate policy on trading partners. Compared to these papers, I bring causal evidence of the effect of devaluations across countries, translate it an absolute effect, and answer the beggar-thy-neighbour question.

The model builds on the open-economy New Keynesian literature. Svensson and van Wijn-bergen (1989), Obstfeld and Rogoff (1995) are founding papers. Subsequent literature has been

preoccupied with optimal monetary policy in the open economy. Clarida et al. (2002), Benigno and Benigno (2006) showed that under complete markets and local currency pricing, optimal policy consists in stabilizing producer price inflation and output gap, like in the closed economy. With local currency pricing, monetary policy trades off internal objectives with international prices misalignment (Devereux and Engel, 2003, Corsetti and Pesenti, 2005, Engel, 2009). In their handbook chapter, Corsetti et al. (2011) summarize the lessons of the international New Keynesian literature: they argue that concerns for competitive devaluations tend to be overrated since policy responses abroad offset strategic terms-of-trade manipulations. To that tradition, I bring more recent elements from the incomplete pass-through literature (Gopinath and Itskhoki, 2010, Amiti et al., 2014a, Burstein and Gopinath, 2014, Itskhoki and Mukhin, 2019). Compared to these papers, my contribution is to build a model with realistic cross-sectional implications for devaluations, and draw quantitative counter-factual conclusions.

I add to empirical studies of large devaluations. Most of it deals with developing countries in a modern context. For instance, Verhoogen (2008) analyzes firm-level data around the 1994 peso devaluation. Burstein et al. (2005, 2007) argue that the behavior of the real exchange rate after a devaluation is attributable to the slow adjustment of non-tradable goods and services. Mattoo et al. (2017), Rose (2018), Alessandria et al. (2018), Rodnyansky (2019), Kohn et al. (2020) study the behavior of exports following a large depreciation. Compared to these papers, I answer the beggar-thy-neighbour question.

2 Cross-Country Evidence

2.1 Difference-in-Difference Estimation (DD)

From an empirical standpoint, the 1931 devaluations have an appealing feature: before 1931, most countries were on a fixed exchange rate; in 1931, a group of countries devalued their currencies while others stayed on the gold standard for several years, some until 1936. Setting endogeneity aside for now, it is tempting to see the first set of countries as a treated group, and the second one as a control group.

To apply this insight, I run the following specification:

$$\log \left(\frac{Z_t^j}{Z_{1930}^j} \right) = \beta_t \log \left(\frac{XR_{1932}^j}{XR_{1930}^j} \right) \times \mathbb{1}_t + \mu_t + e_t^j \quad (1)$$

where Z_t^j is an aggregate variable of interest (industrial production, prices, etc.) in country j in year t ; XR_t^j the exchange rate, expressed as the local-currency price of gold; $\mathbb{1}_t$ a year dummy; and μ_t a year-fixed effect. The sample is restricted to countries that either devalued in 1931 (e.g. Britain) or did not devalue before 1936 (e.g. France).³ I study years 1928 to 1935 in order to include pre- and post-treatment years.

The change in the exchange rate is instrumented with a devaluation dummy, so that β_t can be interpreted as the performance of devaluing countries, relative to non devaluers, scaled by the average change in the exchange rate. Figure 2 displays those coefficients for 9 macroeconomic variables. In the appendix, I also present results for a specification where I estimate equation (1) by ordinary least squares (figure A.3). These results are denoted DD-OLS while the baseline is DD-IV.

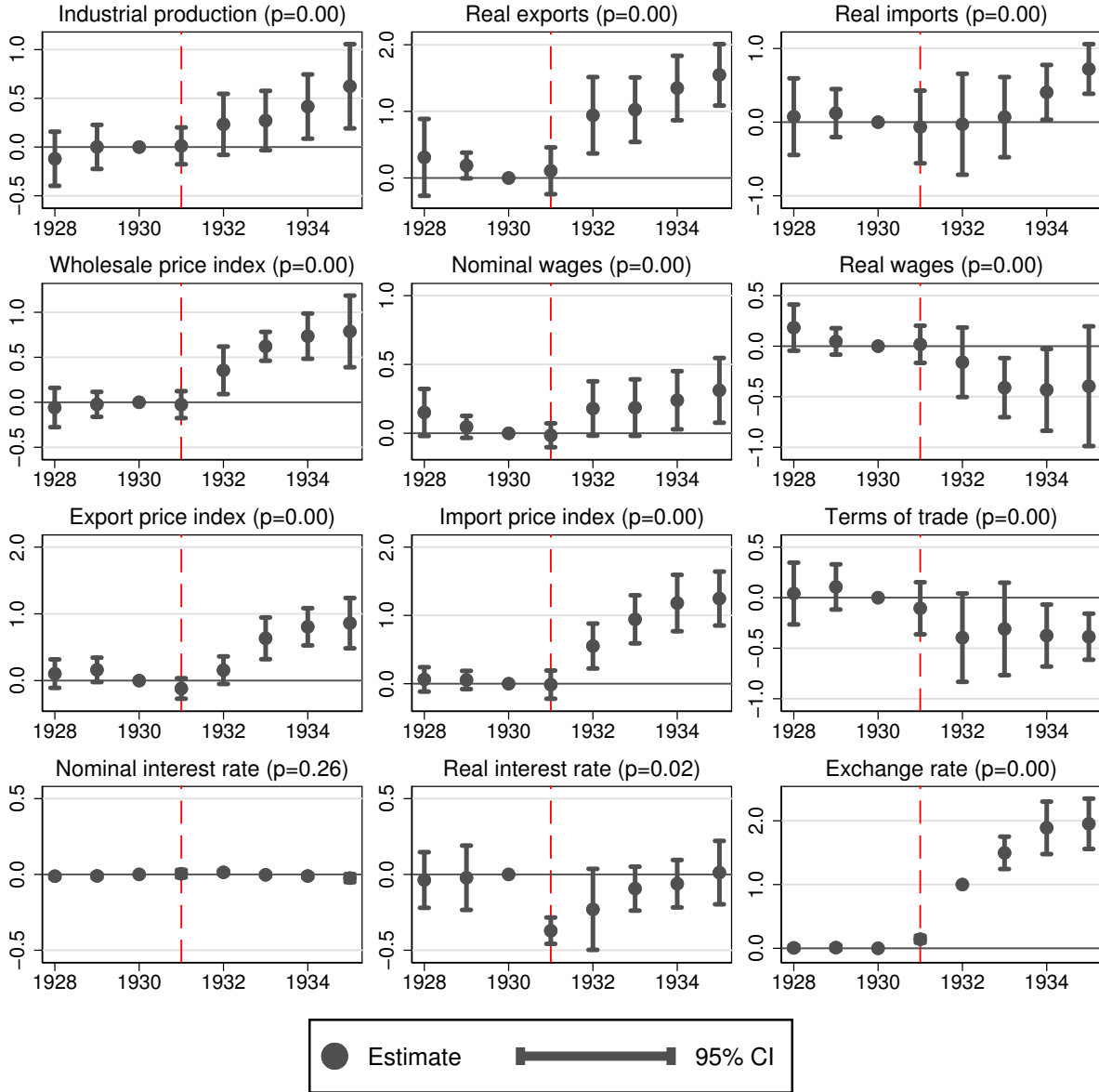
The devaluations were followed by a sharp relative increase in industrial production, exports and, after a few years, imports. The increase in imports may seem surprising since the expenditure switching channel would predict a fall as these become dearer. Nominal quantities also increased, but not one for one. Since the impulse response functions (IRF) are normalized by the change in the exchange rate, full pass-through of the 1930-32 change in the exchange rate to nominal quantities would correspond to a coefficient of 1. For instance, in 1935, wholesale prices had only increased by about 50% in countries that devalued, relative to countries that didn't. The pass-through is even lower for nominal wages, so that real wages went down in countries that devalued.

On the monetary front, the action did not come from the nominal interest rate, but from the real interest rate. The devaluations did not coincide with significant cuts in the discount rate of the central bank. On the other hand, the relative increase in the price level implies that the real interest rate fell sharply.⁴ The 1931 coefficient of -0.37 entails a 11% basis points cut in the real interest rate for a devaluation of 30%. Calling the 1931 depreciation of the pound a devaluation is a slight

³Country samples are summarized in the appendix (table A.1).

⁴Absent direct measures on inflation expectations, I computed the ex-post real interest rate: $r_t \approx i_t - \pi_{t+1}$ where i_t is the discount rate of the central bank and π_{t+1} is the annual inflation rate of the wholesale price index.

Figure 2: Aggregate variables (DD-IV)



Note: response of relevant variable in devaluing countries, relative to non-devaluing ones. Formally, the plots show the estimate of β_t in equation (1). β_{1930} is normalized to 0. Standard errors are bootstrapped with 2,000 replications, and clustered at the country level. The number in parenthesis is the P-value of a test of the joint significance of the 1932 to 1935 coefficients. The y-axis scale of the nominal and real interest rates are identical for visual convenience.

abuse of language. In truth, British authorities suspended the gold convertibility of the pound, let it depreciate for a few months, and actively managed the exchange rate from July 1932 onward. The tie to gold was never formally re-established. Other devaluers similarly allowed their currency to depreciate before pegging to the pound (Eichengreen, 1992, ch. 10). Therefore, I investigate the behavior of the exchange rate after 1932 by estimating equation (1) with the exchange rate on the left-hand side — $\hat{\beta}_{1932} = 1$ by construction. The exchange rate kept depreciating after 1932 (bottom right panel of figure 2).

The behavior of imports is a first dent into the beggar-thy-neighbour story. One side of that story is that, through devaluation, the country makes its domestic goods cheaper, importing less itself, and forcing others to import more. I find no evidence for that mechanism, suggesting either a low substitution between imports and domestic goods, or an offsetting response of domestic demand... or both. In fact, after a few years, those countries import more, which is suggestive of a strong response of domestic demand. These estimates conflict with the findings of a contemporaneous paper by Candia and Pedemonte (2021). Looking at a cross-section of US cities around 1931, they argue that those that were more exposed to sectors which were exporting to devaluing countries experienced a fall in bank debits. (Bank debits are the only city-level measure of economic activity that is available at monthly frequency.) This is puzzling in light of my results: a necessary condition for those cities to be hurt would seem to be that devaluing countries import less. That devaluing countries do not import less will be confirmed by the cleaner high frequency identification.

Did other policies correlate with exchange rate depreciation? The first possibility is changes in tariff. A relative tariff decline in countries that devalued may explain the behavior of imports on figure 1. Figure A.1 in the appendix shows that countries which devalued indeed experienced a relative decline in the ratio of tariff revenues divided by the value of their imports. These changes were driven by tariff hikes in countries that did not devalue, instead of tariff cuts in countries that did (Eichengreen and Irwin, 2010). The second possibility is fiscal policy. After all, the pound devaluation occurred in the face of stalling negotiations around the government budget. It is conceivable that, by relaxing the constraint that it exerted on monetary policy, leaving the gold standard allowed the central bank to print money to finance the fiscal authority. Through higher spending, fiscal policy might have stimulated output. Evidence shown on figure A.1 disproves this

story. If anything, real government spending declined after the devaluations. The test of joint significance of the 1932 to 1935 coefficients returns a p-value that is too high to reject the null that all coefficients are 0. The same is true about revenues.⁵ In any case, these policies do not explain the results. In figure A.2, I add the changes in tariff and fiscal variables as controls, without discernible effects on the point estimates of the main variables. They can lower statistical significance for some variables though (table A.3).

Why did countries devalue, why did they not? According to Eichengreen and Sachs (1985) and Eichengreen (1992), the decision to devalue was shaped by past experience with the gold standard more than contemporaneous economic performance. In Britain, where the pound had been brought back to its pre-WWI parity at the expense of deflation and unemployment, there was little appetite for more austerity. France, on the other hand, had come close to hyperinflation in 1926. The franc convertibility had been restored after a protracted war of attrition between political parties to decide who would bear the costs of adjustment. In the 1930s, even Communists opposed devaluation, which they thought would reduce workers' living standards.

Can figure 2 be interpreted as evidence of a causal link from devaluation to recovery? That none of the variables exhibits a preexisting trend — except nominal wages perhaps — is at least suggestive of a causal relationship. In particular, it rules out the possibility that results are driven by the economy bouncing back from a shock that would have led to devaluation. Nevertheless, devaluing countries may have been reacting to a contemporaneous shock that would have affected future output, such as banking difficulties.⁶ That kind of bias, however, would go in a direction opposite to my results.

Still, it is with these caveats in mind that I turn to a different identification strategy, one which relies on high frequency variation.

⁵Inspecting these results, I found that this (insignificant) fall in real spending and revenues was primarily driven by the increase in the wholesale price index (figure 2), which I use to deflate nominal spending and revenues. Thus, it is plausible that the decline in spending and revenues was due to nominal stickiness in civil servant's salaries or tax brackets.

⁶For instance, Accominotti (2012) argues that London merchant banks were in serious trouble because they were the guarantors of important quantities of Central European debt, whose payments were frozen when Austria, Germany and Hungary implemented capital controls.

2.2 High-Frequency Identification (HFI)

2.2.1 Identification Strategy

The decision to devalue may of course be influenced by fundamentals; but different policymakers may take different decisions in the face of similar circumstances. The 1933 devaluation of the dollar is a good example. Franklin D. Roosevelt, who had just become president, had been unclear about his intentions. The Senate was sympathetic to a *de facto* abandonment of the gold standard in the form of massive purchases of silver. To avoid this radical policy, Roosevelt negotiated in secret an amendment to the Farm Relief Act that gave him authority to cut the gold content of the currency by up to 50%. Edwards (2018, pp. 57–58) tells how he broke the news to his advisers:

On the night of April 18, the president met with his close advisers to discuss issues related to the impending visit of British Prime Minister Ramsay MacDonald. [...] Only [Assistant Secretary of State] Moley knew that Roosevelt had been negotiating a new initiative for “controlled inflation” with a group of key senators, including Elmer Thomas from Oklahoma. When FDR told them, with a chuckle, that the next day he would announce his support for the Thomas Amendment, [adviser] Feis, [Budget Director] Douglas, and [Secretary of State] Warburg became livid; they couldn’t believe what they were hearing and interrupted each other in their efforts to convince the president that this was a mistake of historical proportions. In 1934, Warburg wrote that as late as April 18, those who were in daily contact with FDR had no “idea that he was seriously considering such a move.” [...] After leaving the White House late that night, Lew Douglas told the rest of the group that without a doubt this was “the end of Western civilization.”

That Roosevelt’s closest advisers were stunned by his decision suggests there was nothing ineluctable about it. Had he been president, Budget Director Douglas would have probably chosen a different path for the exchange rate. The silverites of the Senate would have picked yet another path.

That kind of variation is exogenous and can be isolated. To do so, I construct series of policy announcements and use changes in the 3-month forward exchange rate around those announcements as instruments for the exchange rate. Being an asset price, the forward exchange rate should incorporate market expectations about the future path of the exchange rate. So, as long as markets

and policymakers have similar information about current and future economic conditions, movements in the forward rate that are prompted by policy announcements are exogenous, and can be used to infer the effect of exchange rate policy on the economy. Within that framework, a shock can be an unexpected devaluation that happened, or an expected devaluation that did not happen. This strategy is conceptually similar to that of a literature that uses, in a modern context, changes in federal funds rate futures on FOMC days as an instrument for changes in the federal funds rate target in a VAR or a local projection.⁷

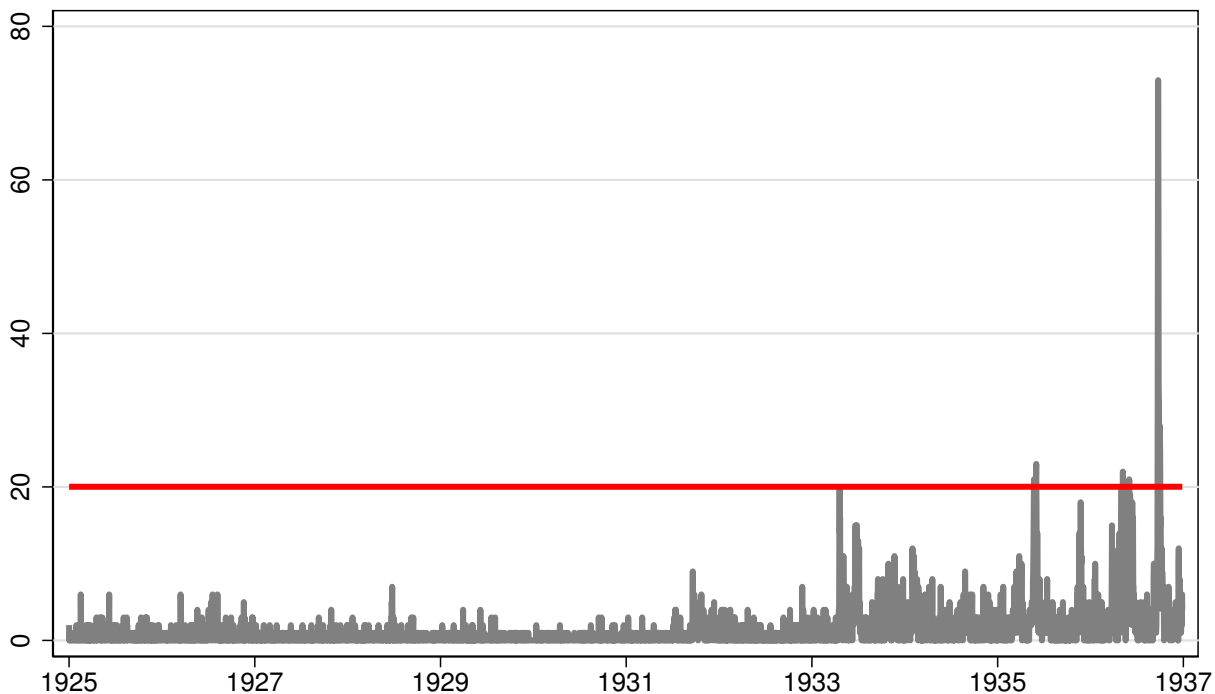
2.2.2 Shock Construction

The main challenge of this identification strategy is to find the dates of the policy announcements. I implement a four-step procedure:

1. Search in ProQuest Historical Newspapers for articles featuring appropriate keywords for each day of the period during which the country is on the gold standard. The keywords are: (i) the name of the country or currency, and (ii) mentions of devaluation, leaving the gold standard, or exchange controls. ProQuest Historical Newspapers is a database of archives of the main US newspapers: New York Times, Wall Street Journal, Chicago Tribune;
2. Retain dates whose number of articles is 6 standard deviations above the mean. With this step, I seek to select days where an important amount of news is released about future exchange rate policy. The threshold of 6 is somewhat arbitrary. I picked it by trading-off quantity and relevance of those dates. A much higher threshold would leave me with too few days, a much lower one would return too many dates where nothing particular is announced;
3. Read those articles to identify the nature and exact timing of the news: retain dates that correspond to a policy announcement (e.g. devaluation), reject them if they're only conveying news about the economic situation (e.g. strike).
4. Use variation in the 3-month forward exchange rate around the announcement as a shock.

⁷See Ramey (2016) for a survey. Bagliano and Favero (1999), Cochrane and Piazzesi (2002), Faust et al. (2004), Barakchian and Crowe (2013), Gertler and Karadi (2015) are early contributors to this identification scheme. Weiss (2020) applies it to silverite agitation during the pre-WWI gold standard era. The use of forward exchange rates to infer devaluation expectations has some precedents in the Great Depression literature (Hsieh and Romer, 2006, Accominotti, 2009). These authors, however, do not study the effect of devaluation shocks on macroeconomic variables.

Figure 3: Daily number of articles for France



Note: number of daily articles returned by ProQuest Historical Newspapers with the keywords described above. The red line is the cutoff defined by the formula: $\text{mean} + 6 \times \text{standard deviations}$.

I illustrate this procedure with the case of France; other countries' details are given in the appendix. France is an ideal example because it waited until 1936 to devalue, and because its decision to finally devalue was preceded by several false alarms — expected devaluations that did not happen. Figure 3 shows the number of daily articles returned by a search of the keywords:

```
ti(France OR French OR franc) AND (devaluation OR ((off OR suspension OR
leave OR quit) AND "gold standard") OR (exchange control))
```

These keywords mean that: (i) “France”, “French” or “franc” must be in the title of the article; and (ii) “devaluation”, or a reference to leaving the gold standard or “exchange control” must be in the title or text of the article. Over the period during which France was on the gold standard (June 1928 to September 1936), the mean of the number of such articles was 1.3 and the standard deviation 3.1. So a day must feature more than 20 articles in order to be retained. The red line on figure 3 is the cutoff.

Table 1 lists the dates that are above that cutoff. First, many dates neighbor each other. For

instance, the September 1936 devaluation of the franc, which was announced on September 26, and officially ratified on October 2, sparked a flurry of articles from September 25 to October 5. Since these articles deal with the same event, I take them as a single shock, which happens on September 26, the day the devaluation was announced. Before devaluation happened, my procedure detects two announcements that correspond to non-devaluations. The first one is the fall of Pierre-Etienne Flandin’s cabinet. While officially opposed to devaluation, Flandin’s government implemented policies that ran against the gold standard constraint: low interest rates, budget deficit, cartelization of the French industry... Flandin’s fall and replacement by Fernand Bouisson, whose government was more committed to deflationary policies, were interpreted as a contractionary shock by the forward exchange rate market. The second non-devaluation is Léon Blum’s speech on May 10, 1936. Blum was a Socialist who led a left-wing coalition, the Popular Front, to victory on May 3, 1936. This victory led to speculation that France would devalue. With that speech, the head of government reassured financial markets about his commitment to the gold standard, thereby generating a contractionary shock. Five months later, the same Blum would take the decision to devalue the franc...

One last difficulty is that, sometimes, financial markets were closed after major announcements. For instance, in 1936, the Paris bourse was closed from September 25 until October 1, when the devaluation law was passed. So there were no forward exchange quotations during that time. As a result, I take the narrowest possible window — 7 days in that case — around the announcement. The French devaluation, however, is an extreme example. In most cases, that narrowest window is 1 or 2 trading days.

2.2.3 Results

I estimate a panel local projection with instrumental variable (Jordà, 2005, Jordà et al., 2015):

$$z_{t+k}^j - z_{t-1}^j = \beta_k(xr_t^j - xr_{t-1}^j) + \gamma_k' X_{t-1}^j + \delta_{t,k} + \zeta_k^j + e_{t,k}^j \quad (2)$$

where z_t^j is the log of the variable of interest in country j in month t , xr_t^j the log of the spot exchange rate, X_t^j is a vector of controls, $\delta_{t,k}$ and ζ_k^j are time and country fixed effects. The spot exchange rate is instrumented with the aforementioned shocks. I estimate equation (2) at quarterly

Table 1: France, shocks

Date	# articles	Event	Shock
27may1935	21	Fall of Flandin's cabinet	-0.046
31may1935	21		
04jun1935	21		
05jun1935	23		
11may1936	22	Blum's devaluation speech	-0.013
05jun1936	21	Strikes	×
25sep1936	30	Devaluation	+0.234
26sep1936	60		
27sep1936	73		
28sep1936	59		
29sep1936	33		
30sep1936	31		
01oct1936	25		
02oct1936	28		
03oct1936	28		
04oct1936	27		
05oct1936	24		

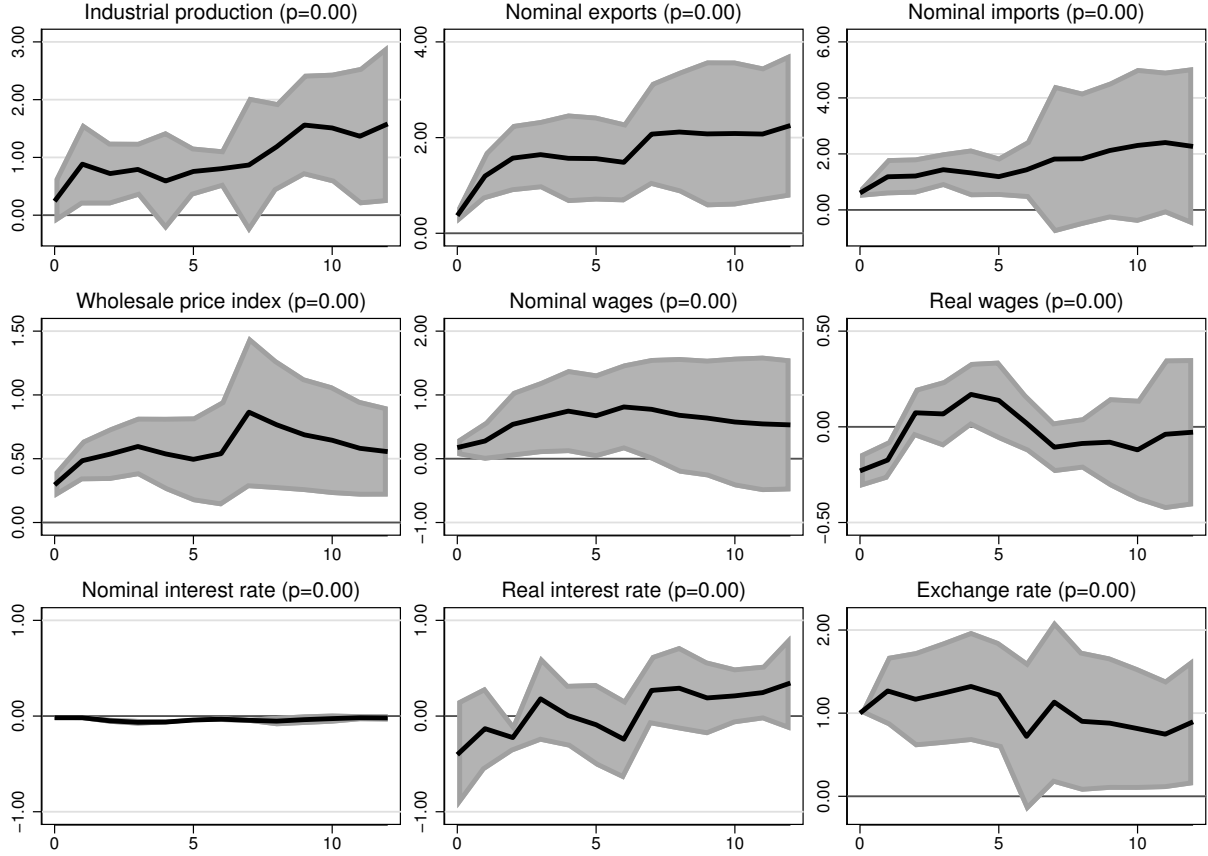
Note: shocks selected by the procedure described in section 2.2. The mean and standard deviations of the daily number of articles are 1.3 and 3.1 respectively. Other countries are shown in table A.4.

frequency. The controls are a year of lagged changes in the outcome variable and the exchange rate.

The countries and outcome variables differ from those of the difference-in-difference exercise. Before, the choice of countries was dictated by the desire to construct a quasi-experimental setup. Now, the sample is determined by data availability: it is made of the eight countries for which forward exchange rate data is continuously available from *The Financial Times* — Germany, Belgium, France, Italy, the Netherlands, Switzerland, the United Kingdom and the US. Moreover, since this exercise requires quarterly data, some of it comes with important limitations. First, exports and imports can only be nominal as quantity and price indices are typically not available above annual frequency. Second, I was not able to find nominal wage series for Belgium, the Netherlands and Switzerland; and that for France is for coal miners only. Third, industrial production is unavailable at quarterly frequency for the Netherlands and Switzerland. For each variable, I include every country for which the data is available.

Figure 4 shows the results for the baseline specification. The results are qualitatively similar

Figure 4: Impulse response functions (HFI)



Note: response of relevant variable to an exogenous 100 log-points devaluation in the exchange rate. Formally, the plots show the estimate of β_k in equation (2). The black line is the point estimate; the gray area is the 95% confidence interval with Driscoll-Kraay standard errors. The number in parenthesis is the p-value of a test of the joint significance of the coefficients. The y-axis scale of the nominal and real interest rates are identical for visual convenience.

to those of section 2.1. The devaluation stimulated industrial production, nominal exports and imports, prices and nominal wages and led to a fall in real wages. Again, the monetary action comes not from the nominal interest rate, but from the real one.

2.2.4 Potential Objections

What about the “information effect” (Nakamura and Steinsson, 2018)? If monetary authorities possess superior information about present and future economic conditions, policy shocks are not exogenous to those same economic conditions. They only reflect superior information. In this context, however, this is less likely to be a problem than with modern data. Governments and central banks had smaller staff than they do today, and they shouldn’t be expected to have had

better information about the state of the economy.

Isn't the timing of the policy announcements endogenous? It certainly is. But this is not a problem as long as that endogeneity is reflected in the forward rate on the eve of the announcement. This would only be a problem if the announcements reflected information about future economic conditions that became known during the day. Since devaluations tended to be protracted affairs whose decisions took months, if not years, to be taken, this is unlikely to be an issue. In the case of the US, the quote above suggests that Roosevelt took his decision at least one day before announcing it.

Don't daily variations in the forward rate partly reflect background noise that correlates with economic fundamentals? It is possible. With modern data, one can keep background noise to a minimum by using 30-minute windows around the announcement. Since only daily data is available in the 1930s, this cannot be done here. Hence, I must assume that there is more good (driven by the announcement) than bad (background noise that correlates with fundamentals) variation on the days of policy announcements. Put another way, the identifying assumption is really that most, if not all, variation in the forward rate, around policy announcements, is exogenous. At any rate, this assumption is much milder than the one required by the DD strategy: that all variation, every day, is exogenous.

Do forward exchange rates capture expectations about future spot rates? Even though, in modern data, the forward rate tends to be a poor proxy for the future spot rate (Fama, 1984), this is less true under fixed exchange rates. For instance, Accominotti et al. (2019) find zero return to the carry trade under fixed exchange rate regimes with the same data, currencies and time period. Their result is in line with other studies from different eras: Flood and Rose (1996) find smaller deviations from uncovered interest rate parity among currencies of the European Monetary System; so do Colacito and Croce (2013) with Bretton-Woods data.

2.3 DD vs. HFI: Comparison

A casual inspection of figures 2 and 4 suggest that, while the two exercises have qualitatively similar conclusions, they differ quantitatively. In the DD case, the industrial production coefficient is mostly below 0.5, meaning that a 30% devaluation stimulates output by less than 15% in relative terms. In the HFI case, that coefficient hovers above 0.5, sometimes even above 1. This kind

of comparison has its limitations for two reasons. First, it does not take standard errors into account — the HFI confidence intervals often include numbers below 0.5. Second, the underlying policy change is not necessarily the same. In the DD case, the exchange rate keeps depreciating, particularly from 1932 to 1933; while the HFI case looks closer to a one-off devaluation.

To refine the comparison, I compute the integral below the coefficients. Formally, for some outcome variable z , I compute:

$$\mathcal{I}^z = \frac{1}{t_2 - t_1 + 1} \sum_{t=t_1}^{t_2} \beta_t^z$$

where β_t^z is the coefficient associated with variable z . To account for the differing exchange rate change, I compute the ratio of the integral for variable z and the integral for the exchange rate: $\mathcal{I}^z / \mathcal{I}^{xr}$. To make the timing of the two exercises as similar as possible, I use the 1932 to 1934 coefficient in the DD case, the horizon-1 to 12 coefficients in the HFI case. Since most of the 1931 devaluations happened in late September or early October, this procedure roughly captures the effect of the devaluation for 3 years, starting one quarter after the devaluation.

This exercise confirms that the HFI strategy delivers larger estimates for industrial production, but the confidence intervals sometimes overlap. In table 2, I present the results for the two versions of the DD strategy — DD-IV where the change in the exchange rate is instrumented with a devaluation dummy, DD-OLS where equation (1) is estimated by OLS — and the HFI strategy. For industrial production, the ratio is more than 3 times higher with the HFI strategy (0.68) than with the DD-IV one (0.20), while DD-OLS lies in between (0.35). The lower bound for the HFI 95% confidence interval is lower than the DD-IV upper bound for the ratio (0.64 against 0.40), but not for the numerator. The differences are smaller and insignificant when it comes to prices.

It is possible that these differences reflect downward bias in the DD estimates. If the 1931 devaluations were driven by negative shocks that affected future output in countries that devalued (such as stress in the financial system), and if these shocks were reflected in the forward rate, then one should expect the HFI strategy to deliver higher estimates. Indeed, the pound was trading at a discount in September 1931, reflecting some anticipations of a currency devaluation. But I caution against reading too much into those differences, for two reasons. First, and once again, the differences are not necessarily statistically significant. Second, the underlying country samples are different. Idiosyncrasies in different countries' experience with the Great Depression or recovery

Table 2: DD vs. HFI: Quantitative Comparison

	DD-IV		DD-OLS		HFI	
	Industrial production					
Numerator	0.39	[0.11,0.66]	0.54	[0.31,0.77]	1.05	[0.47,1.63]
Denominator	1.58	[1.32,1.85]	1.45	[1.22,1.69]	1.03	[0.42,1.64]
Ratio	0.24	[0.06,0.43]	0.37	[0.19,0.56]	1.03	[0.79,1.26]
	Wholesale price index					
Numerator	0.62	[0.38,0.87]	0.56	[0.36,0.76]	0.61	[0.30,0.92]
Denominator	1.44	[1.24,1.64]	1.38	[1.19,1.57]	1.19	[0.52,1.85]
Ratio	0.43	[0.25,0.62]	0.40	[0.25,0.56]	0.51	[0.47,0.56]

Note: point estimates and 95% confidence intervals retrieved from averaging the coefficients whose estimations is detailed in sections 2.1 and 2.2. The numerator is the average coefficient for the relevant variable, the denominator is the average coefficient for the exchange rate. The ratio is the ratio of the two. DD-IV: 1932 to 1934 coefficients of equation (1), instrumenting the change in the exchange rate with a devaluation dummy. DD-OLS: 1932 to 1934 coefficients of equation (1), estimated by OLS. HFI: horizon-1 to 12 coefficients of equation (2).

may also be explanations. For instance, dropping the US from the HFI sample lowers the industrial production ratio from 1.03 to 0.79. This was to be expected as the US experienced a deep depression and a strong recovery.

To summarize section 2, I have presented evidence on the relative effects of devaluations across countries. Once again, a relative statement does not directly answer the question of interest: British output might have been going up compared to French output because it was going up in absolute terms, or because French output was going down. Interpreting these estimates and making a statement about the size of the absolute effect requires a model, to which I turn in the next section.

3 A Multi-Country Model

I consider a world made of a continuum of symmetric countries whose combined size is normalized to 1. In each country, there still are a continuum of firms that each produce a variety of that product. All production is consumed. I relegate detailed proofs to the appendix and focus on the main equations.

3.1 Demand System

In country j , the household minimizes expenditures:

$$\int_f P_t^{jj}(f) C_t^{jj}(f) df + \int_k \int_f P_t^{jk}(f) C_t^{jk}(f) df dk$$

subject to a Kimball (1995) aggregator:

$$C_t^j = (1 - \bar{\Gamma}) C_t^j \times g \left(\int_f g^* \left(\frac{C_t^{jj}(f)}{(1 - \bar{\Gamma}) C_t^j} \right) df \right) + \bar{\Gamma} C_t^j \times g \left(\int_j \int_f g^* \left(\frac{C_t^{jk}(f)}{\bar{\Gamma} C_t^j} \right) df \right) dj \quad (3)$$

$P_t^{jj}(f)$ and $C_t^{jj}(f)$ are the price and consumption, in country j , of the variety produced by firm f of country j . Their product is expenditures on that variety. Summing over all firms f corresponds to expenditures on domestic varieties. $P_t^{jk}(f)$ and $C_t^{jk}(f)$ are the price and consumption, in country j , of the variety produced by firm f of country k . Their product is expenditures on that variety. Summing over all firms f and foreign countries k corresponds to expenditures on foreign varieties. $\bar{\Gamma}$ is the coefficient of openness: the higher $\bar{\Gamma}$, the more the household consumes of the foreign good.

Equation (3) governs aggregation of those varieties through the functions $g(\cdot)$ and $g^*(\cdot)$. I assume particular functional forms for those. $g(\cdot)$, the function that governs aggregation between domestic and foreign varieties, is of the form:

$$g(x) = 1 + \frac{1}{1 - \tilde{\rho}} (x^{1 - \tilde{\rho}} - 1)$$

$g^*(\cdot)$ is the function that governs aggregation among domestic and foreign firms. I borrow the functional form used by Klenow and Willis (2016):

$$h(x) = g^{*-1}(x) = (1 - \bar{m}' \log(x))^{\frac{\theta}{\bar{m}'}} , \quad \theta > 1, \quad \bar{m}' > 0$$

The constant elasticity of substitution (CES) model is nested: this formulation collapses to a CES demand system with elasticity θ if $\tilde{\rho} = 0$ and $\bar{m}' \rightarrow 0$.

Solving the expenditure-minimization problem and log-linearizing around a symmetric steady

state gives the following demand functions:

$$\hat{c}_t^{jk} = -\theta \left(p_t^{jk} - p_t^{j*} \right) + \hat{c}_t^{j*} \quad (4)$$

$$\hat{c}_t^{j*} = -\rho \left(p_t^{j*} - p_t^j \right) + \hat{c}_t^j \quad (5)$$

$$\hat{c}_t^{jj} = -\rho \left(p_t^{jj} - p_t^j \right) + \hat{c}_t^j \quad (6)$$

Equation (4) is the demand for country k 's varieties in country j as a function of their price (p_t^{jk}) relative to the import price index (p_t^{j*}) and of country j 's imports (\hat{c}_t^{j*}). Equations (5) and (6) are the demand for imports (\hat{c}_t^{j*}) and domestic goods (\hat{c}_t^{jj}) as a function of their prices (p_t^{j*} and p_t^{jj}) relative to the consumer price index (p_t^j).

These equations introduce a new parameter, ρ , the elasticity of substitution between domestic and foreign goods, which is given by:

$$\rho \equiv \frac{\theta}{1 + \theta \tilde{\rho}}$$

If $\tilde{\rho} > 0$, $g(\cdot)$ is concave so domestic and foreign varieties become less substitutable ($\rho < \theta$). If $\tilde{\rho} < 0$, $g(\cdot)$ is convex and the opposite happens. In the estimation, I shall argue that $\rho < \theta$ is the relevant case. In theory, however, the relationship is not restricted.

Kimball demand is a standard tool to generate incomplete pass-through of the exchange rate to international prices (Amiti et al., 2014b, Burstein and Gopinath, 2014, Itskhoki and Mukhin, 2019). Compared to these papers, however, I innovate along one dimension: the introduction of an extra aggregator, $g(\cdot)$. This addition allows me to distinguish the elasticity of substitution among imports (θ) from that between imports and domestic varieties (ρ). If $g(\cdot)$ is the identity, then: $\rho = \theta$. Nested CES demand systems within and across sectors can also generate incomplete pass-through if firms are not atomistic within their sectors (Atkeson and Burstein, 2008). To be calibrated, however, this framework requires information on the share of exporting firms and on the concentration of production among producers in a sector, which is not available for this period. Therefore, I opt for the more parsimonious Kimball apparatus.

3.2 Pass-Through

Prices are flexible. Firm f sets its local currency price subject to the demand function that arises from the expenditure minimization of the household. I show in the appendix that this problem implies:

$$p_t^{jk} = (1 - \zeta_1 - \zeta_2)(\Theta + mc_t^k - xr_t^{kj}) + \zeta_1 p_t^{j*} + \zeta_2 p_t^j \quad (7)$$

where:

$$\Theta = \log \frac{\theta}{\theta - 1} \quad \zeta_1 = \frac{\bar{m}(1 - \rho/\theta)}{1 + \bar{m}} \quad \zeta_2 = \frac{\bar{m}\rho/\theta}{1 + \bar{m}} \quad \bar{m} = \frac{\bar{m}'}{\theta - 1}$$

ζ_1 and ζ_2 are the strategic complementarity parameters. Their presence implies that, when they set their price, foreign producers adapt their price to the import (p_t^{j*}) and domestic (p_t^j) price indices of the country of destination. Like the elasticities, the strategic complementarity parameters depend on the aggregator functions $g(\cdot)$ and $g^*(\cdot)$. They lie between 0 and 1. In the CES case, these parameters are equal to 0.

Equation (7) implies that the pass-through of the exchange rate is:

$$1 - \zeta_1 - \zeta_2 = \frac{1}{1 + \bar{m}} < 1 \quad (8)$$

In general this pass-through is below 1. Once again, the CES model is nested. A CES demand implies that the firm charges a constant markup, hence transmits any exchange rate movement one for one — the pass-through is 1. This is apparent in the latter formula: the pass-through converges to 1 if \bar{m} goes to 0. In fact, \bar{m} is the elasticity of the markup. With $\bar{m} \rightarrow 0$, the markup is inelastic, hence constant, which is a feature of the CES case.

3.3 Output

Gathering equations (4–7) implies that output in country j is equal to:

$$\hat{y}_t^j = \underbrace{\psi \left(\int_m (mc_t^m - xr_t^{jm}) dm - mc_t^j \right)}_{\text{expenditure switching}} + \underbrace{(1 - \bar{\Gamma})\hat{c}_t^j}_{\text{domestic consumption}} + \underbrace{\bar{\Gamma} \int_k \hat{c}_t^k dk}_{\text{foreign consumption}} \quad (9)$$

where:

$$\psi \equiv \bar{\Gamma} \left(\frac{(1 - \bar{\Gamma})\rho}{1 + \bar{m}\rho/\theta} + \frac{\theta}{1 + \bar{m}} \right)$$

The last two terms (domestic consumption and foreign consumption) are easy to interpret. If country j consumes more, it demands more of its own goods, hence produces more. Similarly, if its trading partners consume more, they demand more of the country's goods, hence country j produces more.

The first term embodies expenditure switching. It gathers marginal cost in the rest of the world (mc_t^m), adjusted for the bilateral exchange rate (xr_t^{jm}), relative to marginal cost in country j (mc_t^j). If country j 's currency depreciates (xr_t^{mj} increases), it lowers marginal cost expressed in foreign currency. The depreciation finds its way into the price of country j 's varieties through equation (7). Consumers shift expenditures towards those varieties. For given levels of domestic and foreign consumption, country j must produce more.

The beggar-thy-neighbour question amounts to whether the first term dominates the third one — and if so, by how much? Expenditure switching weighs non-devaluing country's output down, but the monetary stimulus to foreign demand pushes it up. In theory, the latter may trump the former. Of course, in equilibrium, whichever dominates will feed into domestic consumption.

The strength of the expenditure switching term is governed by ψ which, besides the coefficient of openness $\bar{\Gamma}$, depends on three parameters: (i) ρ , the elasticity of substitution between foreign and domestic varieties, (ii) θ the elasticity of substitution among foreign varieties and (iii) \bar{m} the elasticity of firms' markup which determines the pass-through.

I now turn to granular trade data to estimate the last two parameters, θ and \bar{m} — the first one, ρ , will be estimated out of the cross-country moments. Doing so, I relax some of the assumptions of this section: single product, absence of demand shocks, symmetric and atomistic countries.

4 Trade Evidence

4.1 Identification Strategy

I assume that demand for product r in country j , C_{rt}^j , is governed by a Kimball (1995) aggregator:

$$C_{rt}^j = (1 - \Gamma_{rt}^j) C_{rt}^j \times g \left(\int_f g^* \left(\frac{C_{rt}^{jj}(f)}{(1 - \Gamma_{rt}^j) C_{rt}^j} \right) df \right) + \Gamma_{rt}^j C_{rt}^j \times g \left(\sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \mathcal{K}_{rt}^{jk} \int_f g^* \left(\frac{C_{rt}^{jk}(f)}{\Gamma_{rt}^j \mathcal{K}_{rt}^{jk} C_{rt}^j} \right) df \right) \quad (10)$$

where $C_{rt}^{jk}(f)$ is demand in country j for product r from firm f of country j .

Equation (10) is a generalization of equation (3). The first generalization is the introduction of many products r : there is now an aggregator within each product category. The second generalization is the introduction of taste shocks for imported (Γ_{rt}^j) and foreign varieties (\mathcal{K}_{rt}^{jk}).

Solving the expenditure minimization problem and log-linearizing, I can write:

$$\Delta c_{rt}^{jk} = -\theta \Delta p_{rt}^{jk} + \theta \Delta p_{rt}^{j*} + \Delta c_{rt}^{j*} + \Delta \kappa_{rt}^{jk} \quad (11)$$

Δ denotes time differentiation, and lower case letters logarithms. c_{rt}^{jk} is country j 's demand for country k 's varieties of product r and p_{rt}^{jk} their price index. c_{rt}^{j*} and p_{rt}^{j*} are country k 's quantity and price of imports of product r . κ_{rt}^{jk} shifts expenditures to country j 's variety away from other foreign varieties.

θ is the price elasticity of substitution among foreign varieties. It is the parameter that I seek to estimate. It is given by: $\theta = -h'(1)$ where $h(\cdot) = (g^{*'})^{-1}(\cdot)$. Importantly, I did not assume a constant elasticity of substitution (CES) demand system to arrive at equation (11). The elasticity is not necessarily constant. On the other hand, θ is the elasticity around a symmetric equilibrium where relative prices are equal to 1.

The empirical analogue of equation (11) is:

$$\Delta c_{rt}^{jk} = -\theta \Delta p_{rt}^{jk} + \mu_{rt}^j + e_{rt}^{jk} \quad (12)$$

The fixed effect μ_{rt}^j soaks up the $\theta\Delta p_{rt}^{j*} + \Delta c_{rt}^{j*}$ term. Yet, simultaneity bias would be looming if I were to estimate equation (12) by OLS. To solve this problem, I instrument the price, p_{rt}^{jk} , with the exchange rate, xr_t^k . Since I use data on US imports ($j = us$), this strategy identifies θ as long as devaluation is uncorrelated with within-product US relative demand shocks, $\Delta\kappa_{rt}^{us,k}$.

Proposition 1 (Elasticity) *Provided devaluation is exogenous to within-product US demand shocks, instrumenting prices with the exchange rate in regression (12) identifies the within-product elasticity of substitution among foreign varieties, θ .*

To understand the strengths and limitations of proposition 1, consider two examples, one where the identification assumption is satisfied and one where it's not. If Britain devalued because it exported a lot of steel and the world demand for steel had collapsed because of the Great Depression, the identification assumption holds. Indeed, such a confounding factor is controlled for by the product-time fixed effects, μ_{rt}^j . If, on the other hand, Britain devalued because the US was suddenly importing less steel from Britain than from France, then the identification assumption fails. Such a story doesn't seem particularly plausible. The identification assumption also allows for any correlation of devaluation with supply-driven shocks. Its failure can only come from a mechanism involving relative demand shocks.

4.2 Data

Probably because tariff revenues were the main source of revenues of the federal government in the first hundred years of its existence, the United States was already collecting detailed trade data in 1790. The first Treasury of the Secretary, Alexander Hamilton, would thus transmit to Congress annual statements on exports and imports. This tradition lasted for the whole 19th century. In the late 1890s, the Bureau of Statistics of the Treasury even started issuing monthly data. In 1904, this responsibility shifted to the newly created Department of Commerce and Labor.

In the 1930s, the annual data would be published in a report of the Department of Commerce called the *Foreign Commerce and Navigation of the United States*. This publication, which lasted until the 1960s, contains the equivalent of the trade data that one can download nowadays from the Census' website. It features yearly information at the product level, broken down by countries. Products are fairly narrow categories. For instance, in 1930, there are 98 products in the metals

Table 3: An example of the data

288

No. 4.—GENERAL IMPORTS, 1931—GROUP 6—METALS

COUNTRY	Zinc—Continued			Antimony						6660. Cobalt ore and metal (free)		6662. Quick-silver or mercury (dut.)		6740. Other ores, metals and alloys (free)
	6558. Blocks, pigs, etc., dross, and old (dut.)	6559. Sheets, dust and manufactures (dut.)	6650. Ore (free)		6651. Needle or liquated, and regulus or metal (dut.)									
			Gross weight	Antimony content										
					Pounds	Dollars	Dollars	Pounds	Dollars	Pounds	Dollars	Pounds	Dollars	
TOTAL	549, 185	14, 829	24, 999	28, 030, 059	9, 726, 015	259, 952	10, 390, 599	445, 369	248, 862	262, 973	27, 054	32, 649	88, 344	
Austria			2, 874											3, 931
Belgium	11, 235	401	4, 480				6, 720	504	99, 198	159, 299				3, 927
France			947											7, 699
Germany	414	157	13, 812				220	27	3, 016	5, 876				303
Italy			261											150
Soviet Russia in Europe														
Spain			7								26, 609	32, 027		
Sweden			255											
Switzerland			212											
United Kingdom	276, 919	8, 934	212				226, 464	20, 624	4, 480	10, 084				9, 998
Canada	567	35	2, 023						118, 872	83, 172				63, 216
Panama														243
Mexico	260, 050	5, 302		26, 917, 869	9, 358, 575	238, 151	486, 169	27, 656			445	622		775
Trinidad and Tobago														107
Cuba														156
Dominican Republic														264
Argentina				1, 112, 200	367, 440	21, 801								80
Chile														161
British India														17
China														10
Hong Kong							9, 671, 026	396, 558						35
Japan			128											9
Palestine														194
Siam														41
Australia														46
Union of South Africa									23, 296	4, 542				582

COUNTRY	6760. Other ores, metals and alloys (dut.)		6790. Manufactures of metals, n. e. s. (dut.)	6810. Gold and silver sweepings (free)	Platinum					6823. Iridium (free)		6824. Osmium and osmium-iridium (free)		
					6820. Ore (platinum content) (free)		6821. Grains, nuggets, sponge, and scrap (free)		6822. Ingots, bars, sheets, and plates (free)					
					Ounces troy	Dollars	Ounces troy	Dollars	Ounces troy	Dollars	Ounces troy	Dollars		
TOTAL	Pounds	Dollars	Dollars	Dollars	Ounces troy	Dollars	Ounces troy	Dollars	Ounces troy	Dollars	Ounces troy	Dollars	Ounces troy	Dollars
Austria			2, 840, 110	167, 149	1, 097	27, 466	67, 078	1, 905, 062	23, 553	657, 413	1, 773	195, 961	13	
Belgium			91, 517											
Czechoslovakia			13, 986											
Denmark			137, 370											
Finland			7, 069											
France	72	20	71											
Germany	32, 753	10, 909	276, 398											
Greece			1, 520, 512						258	7, 941	943	90, 818		
Hungary			687											
Irish Free State			55											
Italy			312											
Latvia			44, 144											
Netherlands			10											
Norway			69, 115											
Poland and Danzig			1, 186											
Portugal			155											
Soviet Russia in Europe			50											
Spain			3, 597											
Sweden			2, 406				1, 412	42, 374	5, 386	173, 523	540	78, 341		
Switzerland			10, 269											
United Kingdom	150	218	19, 495											
Canada			387, 187	249	745	19, 255	25, 470	654, 265	16, 844	439, 717	290	26, 789	5, 833	274, 343
Costa Rica			28, 029	104, 852			138	3, 591					4	616
Nicaragua				800										
Panama				736										
Mexico	205	14												
Newfoundland and Labrador			1, 306											
Bermudas			58											
Jamaica			6											
Cuba			235	9										
Haiti, Republic of			891	28, 167										
Argentina			4											
Bolivia			126	28, 831										
Brazil				34										
Colombia			5	665										
Ecuador			35		340	7, 831	40, 058	1, 204, 832	1, 065	36, 232				
Venezuela	175, 065	158, 883	59											
British India				2, 201	12	380								
British Malaya			2, 234											
Ceylon			61	280										
China			411											
Java and Madura			77, 590	25										
Hong Kong			25											
Iraq			8, 763											
Japan			171											
Palestine	2, 154	2, 036	127, 542											
Philippine Islands			3, 404											
Syria			30	300										
Turkey in Asia and Europe			1, 840											
Australia			211											
British Oceania			81											
Union of South Africa			4											
Nigeria			3											
Egypt			7											
Morocco			1, 275											
			113											

group alone. Among them are steel bars, hand-sewing and darning needles, or manganese ore. For each product, I observe yearly imports both in value and in the relevant quantity. For example, I know that, in 1931, the US imported 1,260,496 golf balls from the United Kingdom, for a value of \$387,513. Table 3 is a snapshot taken from the 1931 issue.

While this data is publicly available, it needs to be digitized.⁸ I used an optical character recognition software, ABBY FineReader. This process entails a lot of errors, and requires many manual adjustments and corrections. Luckily, the tables include subtotals for each product. I can thus check that imports add up at the product level, which greatly limits the possibility of reading mistakes. I digitized the data for years 1929 to 1933. As far as I know, mine is the first paper to use this data at this level of detail.

A notable feature of this data is the way the value of imports is computed. Imported goods were appraised based on the prices prevailing in the country of origin, for domestic or export purposes, inflated with shipping costs. The procedure, as it is described in the Tariff Act of 1930 (p. 132 and following), worked as follows: (i) the importer filed an invoice that included “the purchase price of each item in the currency of the purchase”, (ii) an individual mandated by the Treasury appraised “the merchandise in the unit of quantity in which the merchandise is usually bought and sold by ascertaining or estimating the value thereof by all reasonable ways and means in his power, any statement of cost or cost of production in any invoice, affidavit, declaration, or other document to the contrary notwithstanding”, and (iii) the customs “collector shall give written notice of appraisement to the consignee [...] if the appraised value is higher than the entered value”. This procedure was meant to avoid invoicing fraud. It is unclear how often and how much appraisers deviated from the value contained in the invoice. It is also unclear that it should bias my estimates in a particular way.

4.3 Samples

Since the bulk of the devaluations happened in 1931, I estimate equation (12) with data for two separate years: 1932 and 1933. The year that serves as the basis for comparison is 1930 in both cases. In each case, the identifying variation lies in the change in the exchange rate between 1930

⁸I access it through NewsBank’s “U.S. Congressional Serial Set, 1817-1980” database. The files are high-quality but image-only PDF.

and the relevant year. Moreover, the panel is made exclusively of countries that devalued in 1931, or had not devalued at the end of the relevant year. For instance, Australia, whose currency depreciated as soon as 1930, is dropped from both panels. Estonia, which devalued in 1933, is kept in the first panel (1932), but kicked out of the second one (1933). Among countries that make it in both panels are of course the United Kingdom (it devalued in 1931) and France (it waited until 1936 to devalue). The point of this procedure is to use as large a panel as possible, while avoiding mixing countries that did not devalue at the same time. Table A.1 in the appendix gives the country list of each sample.⁹ Section C in the appendix contains further details on data construction.

4.4 Results: Elasticity

4.4.1 Baseline

I show the baseline results in panel A of table 4. Columns (1), (2) and (3) respectively feature the results for the reduced form, first stage and instrumented equations using the 1930-32 change. The estimated elasticity of quantity with respect to the exchange rate is 0.910 in 1932, while the price response is -0.460. This means that, thanks to its devaluation of 28%, Britain increased the quantity of its exports by 26% compared to France that did not; and decreased their average dollar price by 13% compared to French products. The implied elasticity of substitution among foreign varieties, $\hat{\theta}$, is 2.054. Columns (4), (5) and (6) show the results using the 1930-33 change. The point estimate of the elasticity goes up to 3.084.

Going down to panel B, I run the same regressions while weighting each product-country pair by its 1930 value. For the quantity regression, this specification corresponds to the change in a quantity index, holding prices constant; while for the price regression, it corresponds to the change in a price index, holding quantity constant. The quantity coefficients are mostly unchanged. The price coefficients go down a bit — though not significantly so — which mechanically raises the point estimate for the price elasticity.

⁹I noticed only one case where this selection procedure seems to affect the results: Chile in the 1932 weighted specification. Chile officially devalued in 1932 so I do not include it in my sample. However, it set up an export exchange rate whose devaluation started in 1931. By 1933, this exchange rate was so devalued (-75.4%) that the log-change goes wild and that a few heavily-weighted observations can significantly affect the final result.

Table 4: Elasticity baseline

	1930-32 change			1930-33 change		
	(1) RF	(2) FS	(3) IV	(4) RF	(5) FS	(6) IV
Panel A: Unweighted						
XR	0.910** (0.349)	-0.443*** (0.110)		1.205*** (0.233)	-0.391*** (0.105)	
Elasticity			2.054*** (0.578)			3.084*** (0.413)
Observations	3742	3742	3742	3446	3446	3446
F-statistic			16.142			13.915
Panel B: Weighted						
XR	0.998*** (0.348)	-0.351** (0.136)		0.965*** (0.264)	-0.327*** (0.088)	
Elasticity			2.843** (1.207)			2.951*** (0.853)
Observations	3742	3742	3742	3446	3446	3446
F-statistic			6.643			13.781
Panel C: Unweighted placebo						
XR	0.027 (0.122)	0.049 (0.047)		0.102 (0.069)	0.024 (0.034)	
Elasticity			-0.550 (2.527)			-4.310 (6.439)
Observations	3755	3755	3755	3579	3579	3579
F-statistic			1.097			0.482
Panel D: Weighted placebo						
XR	-0.135 (0.233)	-0.049 (0.054)		-0.181 (0.173)	-0.061 (0.042)	
Elasticity			-2.755 (3.073)			-2.955* (1.565)
Observations	3755	3755	3755	3579	3579	3579
F-statistic			0.824			2.141

Note: reduced form (RF), first stage (FS) and instrumented (IV) estimates of equation (12). Standard errors are in parentheses. They are double clustered at the country and product levels. F-statistic is the Kleibergen-Paap Wald rk F-statistic of the first stage.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

4.4.2 Endogeneity Concerns

Were devaluations exogenous to US demand shocks? To answer this reverse causality concern, I regress pre-devaluation changes in quantity and price on the later change in the exchange rate. To be precise, I estimate equation (12) with the 1929-30 change for Δc_{rt}^{jk} and Δp_{rt}^{jk} , and the 1930-32

and 1930-33 changes for the instrument.

The results of this exercise are presented in panels C and D. None of the quantity coefficients is significantly different from 0. Their sign varies with the specification. The price coefficients are always close to 0. The estimates for the elasticity of substitution are negative. One of the placebo coefficients — column (6) of panel D — is significant at the 10% level. There are 12 regressions so one false positive at that level of confidence is what one should expect when the true coefficients are zero. Moreover, the coefficient is of a different sign than the one for the unweighted placebo, and much smaller than its counterpart of panel B. Most importantly, the point estimates of the elasticity of substitution are all negative while the standard errors explode.

4.4.3 Elasticity: Other Results

In appendix F, I present other results related to the estimation of trade elasticities.

First, I investigate whether the presence of tariffs in the form of fixed nominal duties biases the estimates (section F.1). I find slightly higher elasticities (up to 4) if I restrict the sample to products that are taxed with a proportional duty, suggesting that this feature of the interwar US tariff code may indeed slightly bias the estimates.

Second, I investigate other mechanisms that could explain the findings, such as healthier financial system or differential slack in the exporting economies (section F.2). Since they involve a move of the supply curve, these other mechanisms would not threaten the identification of the elasticity, but they change the interpretation of the reduced-form estimates as being driven by the change in the exchange rate. Based on the results, I conclude that the exchange rate is the most plausible channel.

Third, I allow the elasticity to differ across goods (section F.3). The data is too noisy to draw a sharp conclusion.

Fourth, I discuss how my results shed light on the “international elasticity puzzle” (section F.4). Estimates of international elasticities from the macroeconomics literature are usually low (around 1), while more recent estimates from the trade literature can be much higher (above 2, sometimes up to 10). My estimates (2–4) are on the low side of the trade literature. I argue that they are consistent with explanations that emphasize that these strands of the literature estimate different objects (Feenstra et al., 2018).

4.5 Results: Pass-Through

4.5.1 Empirical Results

The coefficients of columns (2) and (5) of table 4 are not exactly comparable to standard pass-through estimates. Indeed, that literature usually adds a control like the price level to proxy for marginal cost (Burstein and Gopinath, 2014). Thus, to estimate pass-through, I now run:

$$\Delta p_{rt}^{jk} = \alpha \Delta x r_{rt}^{jk} + \beta \Delta w p_t^k + \nu_{rt}^i + e_{rt}^{jk} \quad (13)$$

where $w p_t^k$ the logarithm of the wholesale price index in the country of origin.

I show the results in table 5. The pass-through estimates are around 0.4. This number is similar to those reported by Burstein and Gopinath (2014).

Table 5: Pass-Through

	1930-32 change		1930-33 change	
	(1) Unwght.	(2) Wght.	(3) Unwght.	(4) Unwght.
XR	-0.370** (0.159)	-0.491*** (0.176)	-0.442*** (0.148)	-0.355** (0.144)
WPI	-0.279 (0.286)	0.417 (0.353)	0.159 (0.300)	0.091 (0.357)
Observations	3592	3592	3405	3405

Note: Standard errors are in parentheses. They are double clustered at the country and product levels.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

4.5.2 Identification

Equation (7) makes clear that, if I were equipped with a proper control for marginal cost, I could identify $1 - \zeta_1 - \zeta_2$ by regressing p_t^{jk} on the exchange rate, that control and a time fixed effect with the data of section 4. Since marginal cost is hard to observe, the standard approach — which I followed in section 4.5 — is to use the price level as a proxy. It turns out this control is not entirely valid in this model. As a matter of fact, the producer price index — which I interpret as the model

equivalent of the wholesale price index — is given by:

$$p_t^{kk} = \Theta + (1 - \delta\bar{\Gamma})mc_t^k + \delta\bar{\Gamma} \int_m (mc_t^m - xr_t^{mj}) dm + \delta\bar{\Gamma}xr^{kj} \quad (14)$$

The first term after the constant (Θ) is the desired variation, the second term — which accounts for the marginal cost of foreign producers — does not depend on k and can be absorbed by a time fixed effect. The third term is problematic because it depends on the exchange rate. This means that the movement in the marginal cost will be overestimated, hence that pass-through will be overestimated. This bias is likely to be small since $\bar{\Gamma}$ is the coefficient of openness (0.1 in my calibration) and δ , the strategic complementarity between domestic and foreign producers, is less than 1. Formally, I prove in the appendix that:

Proposition 2 (Pass-through) *The pass-through regression (13) featuring the wholesale price index as control identifies:*

$$\hat{\alpha} = -\frac{1 - \zeta_1 - \zeta_2}{1 - \delta\bar{\Gamma}}$$

Moreover, ζ_1 , ζ_2 and δ are tied together by the elasticity of markups, \bar{m} :

$$\zeta_1 = \frac{\bar{m}(1 - \rho/\theta)}{1 + \bar{m}} \quad \zeta_2 = \frac{\bar{m}\rho/\theta}{1 + \bar{m}} \quad \delta = \frac{\bar{m}\rho}{\theta + \bar{m}\rho}$$

where: $\bar{m} = \bar{m}'/(\theta - 1)$.

To understand proposition 2, it is useful to, once again, go back to the CES benchmark. With CES demand, markups are constant so \bar{m} , the elasticity of markups, is 0. In that case, the pass-through is just 1. As markups become elastic, the true pass-through ($1 - \zeta_1 - \zeta_2$) declines and the bias ($1/(1 - \delta\bar{\Gamma})$) increases. So $\hat{\alpha}$ moves away from -1. Proposition 2 also makes clear that, given ρ , θ and $\bar{\Gamma}$, the pass-through regression pins down \bar{m} , from which one can recover ζ_1 , ζ_2 and δ .

To summarize section 4, I have used granular trade data to estimate an elasticity and a pass-through coefficient. Those will be key in disciplining expenditure switching in the model. I now go back to the model, and close it by presenting its dynamic side.

5 Model: Dynamics

5.1 Monetary Policy

The central bank commits to buy and sell gold at price \mathcal{E}_t^j , and it issues money in proportion of its gold reserves:¹⁰

$$\Lambda M_t^j = \mathcal{E}_t^j G_t^j \quad (15)$$

Λ is the gold cover ratio, which is assumed to be constant for now. Gold is in fixed supply at the world level, and can only be held for monetary purposes. As a result, the world supply of money depends only on the price of gold, up to a first-order approximation:

$$\int_j m_t^j dj = \int_j \epsilon_t^j dj \quad (16)$$

The exchange rate between countries j and k is pinned down by arbitrage at the relative price of gold:

$$XR^{kj} = \frac{\mathcal{E}_t^k}{\mathcal{E}_t^j} \quad (17)$$

Otherwise, households could buy gold where it is cheap, sell it where it is expensive and make a risk-free profit on the exchange rate market. In log-linear terms, this can be written as: $xr^{kj} = \epsilon_t^k - \epsilon_t^j$. Since there is free capital mobility, uncovered interest rate parity holds to a first order: $i_t^j = i_t^k - E_t \Delta xr_{t+1}^{kj}$. As long as the gold standard is credible ($E_t \Delta xr_{t+1}^{kj} = 0$), all central banks have the same interest rate.

To induce stationarity, I follow Schmitt-Grohé and Uribe (2003) and assume that the interest rate earned by households slightly depends on the level of bonds owned by the country:

$$i_t^j = i_t^W - v b_t^j$$

where v is strictly positive but close to 0.

¹⁰The assumptions made in this section are relaxed in section 6.4 and appendix G through a sufficient statistics approach.

5.2 Rest of the Model

In each country, the representative household derives utility from consumption with internal habits, and money holdings. These assumptions yield an Euler equation and a demand for money:¹¹

$$\hat{\mu}_t^j = E_t \hat{\mu}_{t+1}^j + \hat{i}_t^j - E_t \pi_{ct+1}^j \quad (18)$$

$$\hat{i}_t^j = \frac{1-\beta}{\beta^2} \left(-\chi(\widehat{m_t^j - p_t^j}) - \hat{\mu}_t^j \right) \quad (19)$$

where μ_t^j is the marginal utility of consumption, i_t^j is the nominal interest rate, π_t^j inflation, and m_t^j the money stock. The world interest rate can be obtained by integrating over countries in the money demand and substituting out the world money supply thanks to equation (16):

$$\hat{i}_t^W = \frac{1-\beta}{\beta^2} \left(\chi \int_j (\widehat{p_t^j - \epsilon_t^j}) dj - \int_j \hat{\mu}_t^j dj \right) \quad (20)$$

Each firm produces output with only labor as an input:

$$Y_t^j(f) = f(N_t^j(f))$$

Therefore, marginal cost and aggregate output are, in log-linear terms:

$$mc_t^j = w_t^j + \frac{\alpha_1}{1-\alpha_2} \hat{y}_t^j \quad (21)$$

$$\hat{y}_t^j = (1-\alpha_2) \hat{n}_t^j \quad (22)$$

where:

$$\alpha_1 \equiv \frac{\bar{N} f''(\bar{N})}{f'(\bar{N})} \quad \alpha_2 \equiv 1 - \frac{\bar{N} f'(\bar{N})}{f(\bar{N})}$$

α_1 is the curvature of the production function. α_2 would be the profit share under competitive markets. In the Cobb-Douglas case, of course, α_1 and α_2 are equal. Assuming this more general production function will allow me, in the calibration, to match the behavior of producer prices by adjusting α_1 .

¹¹Once again, formal derivations are relegated to the appendix.

Wages are set *à la* Calvo, which yields a New Keynesian wage Phillips curve:

$$\pi_{wt}^j = \beta E_t \pi_{wt+1}^j + \tilde{\kappa} y_t^j - \kappa \hat{\mu}_t^j + \kappa (w_t^j - p_t^j) \quad (23)$$

Now that the static and dynamic sides of the model have been presented, I can estimate it and run counter-factual experiments.

6 Structural Estimation and Counterfactual Analysis

6.1 Structural Estimation

6.1.1 Structural Estimation: Strategy

I simulate an unexpected devaluation in half of the world, and estimate 6 parameters by matching a set of empirical moments. Devaluation is modeled as exogenous. While I took endogeneity seriously in the empirical part, the goal of this theoretical exercise is not to model endogenous devaluation, but to infer the effect of a devaluation, for given economic conditions. Although the results are cast in terms of deviations from steady state, the linearity of the model guarantees that they can be interpreted as deviations from any set of fundamentals.

Following standard practice in the estimation of general equilibrium models, I calibrate some parameters with standard values from external sources, and estimate those that are essential to match the empirical moments (table 6).¹²

I estimate 4 parameters, σ^{-1} , α_1 , ξ and ρ , with the impulse response functions of industrial production, the wholesale price index, nominal wages, imports, and the import price index. σ^{-1} is the inter-temporal elasticity of substitution (IES). In this environment, a devaluation is a proportional increase in the money supply. This stimulates domestic demand by lowering the real interest rate. So it is natural to adjust the IES in order to match the relative response of output. α_1 is the curvature of the production function. Equation (21) shows that it controls the response of marginal cost to output for a given nominal wage. Since marginal costs cannot be observed, their closest equivalent is the producer price index, as equation (14) suggests. So I adjust α_1 to match the response of wholesale prices. ξ is the Calvo parameter for wages. I adjust it to match

¹²See Christiano et al. (2005) or Auclert et al. (2020) for instance.

Table 6: Parameters

Parameter	Value	Concept	Target or source
Calibrated			
β	0.99	Discount rate	Annual interest rate of 4%
ι	0.8	Habit	Eggertsson (2008)
χ	1	Concavity of utility for money	Standard
γ	0.1	Openness	British export-to-GDP ratio
α_2	0.37	Profit share	Labor share of 2/3
η	21	Elasticity of labor demand	Christiano et al. (2005)
ν	0.0007	Stationarity-inducing device	Schmitt-Grohé and Uribe (2003)
g^*	0.14	Gold to output ratio	British gold-reserves-to-GDP ratio
Estimated			
σ^{-1}	0.63 (0.22)	IES	Industrial production
α_1	0.55 (0.13)	Curvature of production function	WPI
ξ	0.91 (0.02)	Calvo wage parameter	Nominal wages
ρ	0.44 (0.25)	Macro trade elasticity	Imports and IPI
θ	3.01 (0.2)	Micro trade elasticity	US imports
\bar{m}	1.3 (0.03)	Markup elasticity	Pass-through

Note: calibrated and estimated parameters. The estimation follows the procedure described in section 6.1. The targets are the DD-IV cross-country moments of section 2.1 and the trade estimates of section 4.

the response of nominal wages. Finally, ρ is the elasticity of substitution between domestic goods and imports. I target the response of imports, and of import prices.

I estimate 2 parameters, θ and \bar{m} , in order to match the estimates of section 4. By proposition 1, the elasticity of substitution among foreign varieties, θ , was estimated in table 4. By proposition 2, the pass-through regression recovers \bar{m} for given ρ , θ and $\bar{\Gamma}$. I use the point estimates and standard deviations of the 1930-33 change in the unweighted case.

In section 3, I introduced the three parameters that determine the strength of expenditure switching: ρ , θ , and \bar{m} . In section 4, I explained how θ and \bar{m} are identified. Let me now focus on ρ . Identification of the other three parameters (σ^{-1} , ξ , α_1) will be explained in section 6.3. Looking back at equation (5), the demand for imports is a function of two terms: the import price index (IPI) relative to that of domestic consumption (CPI), scaled by the elasticity (ρ), and domestic consumption: the cheaper imports are, the more a country imports; the more a country consumes,

the more it imports. Playing with this equation, I can express ρ as:

$$\rho = -\frac{\hat{c}_t^{D*} - \hat{c}_t^{N*} - (c_t^D - c_t^N)}{p_t^{D*} - p_t^{N*} - (p_t^D - p_t^N)} \quad (24)$$

Superscripts D and N respectively denote devaluing and non-devaluing countries. Equation (24) means that what is informative about ρ is the relative response of imports ($\hat{c}_t^{D*} - \hat{c}_t^{N*}$) compared to the relative response of domestic consumption ($c_t^D - c_t^N$), versus the relative response of import prices ($p_t^{D*} - p_t^{N*}$) compared to the relative response of the domestic price index ($p_t^D - p_t^N$). Intuitively, for any given response of consumption and prices, a low response of imports is suggestive of a low elasticity of substitution.

6.1.2 Structural Estimation: Results

I report here the results where I target the DD-IV estimates as empirical targets. Results where I use the HFI ones are in the appendix.¹³ I simulate the model at quarterly frequency with the devaluations happening in the last quarter of 1931. I then aggregate the simulated data over years to obtain regression coefficients that are consistent with the empirical ones.

Formally, I assume that before time 0, the economy is in steady state. At time 0, half of the world experiences a sequence of unexpected cuts in the gold content of their currencies, $\Delta\epsilon_0^j > 0$. In order to reproduce the path of the exchange rate shown in the bottom right panel of figure 2, I assume that some unexpected shocks keep hitting the economy. The value of the shocks is set to minimize the distance with the empirical response of the exchange rate. An alternative route would be to assume that the whole path is expected from time 0 onward. The issue is that, through uncovered interest rate parity, this implies that the nominal interest rate shoots up in devaluing countries, which did not happen.

I estimate the model by the simulated method of moments (SMM). Technical details are relegated to appendix E. The estimation yields values for σ^{-1} and α_1 that are very close to conventional one. σ^{-1} is 0.63, the most commonly used values in calibration exercises being 1/2. $\alpha_1 = 0.55$ implies a slight departure from a Cobb-Douglas production function, but not a very significant one — $\alpha_1 = \alpha_2 = 0.37$ would be the Cobb-Douglas case. Compared to the Cobb-Douglas case,

¹³See figures A.5, A.6 and table A.5.

$\alpha_1 > \alpha_2$ makes marginal cost more elastic to output. $\xi = 0.9$ implies very sticky wages, with a mean duration of 10 quarters. This is the price of matching the data: 4 years after the devaluation, the pass-through of the exchange rate to nominal wages was only 20%. $\rho = 0.51$ is low but not out of line with other estimates, especially when considering the large confidence interval. For instance, using data from that period, Irwin (1998b) estimates an elasticity of 0.8.

Figure 5 shows the fit of the model for real and nominal variables. For the reader’s convenience, targeted IRF are circled in red and denoted “targ.” Untargeted IRF are circled in green and denoted “unt.” Implicitly targeted IRF are circled in blue and denoted “imp.” By implicitly targeted, I mean that the variable is a function of variables that have been targeted. For instance, real wages are nominal wages divided by the price index. Since both are targeted, real wages are implicitly targeted. The model is able to match the cross-country evidence well. The fit is particularly good for wholesale prices and nominal wages. Even moments that I have not targeted, real exports and the export price index, are decently predicted.

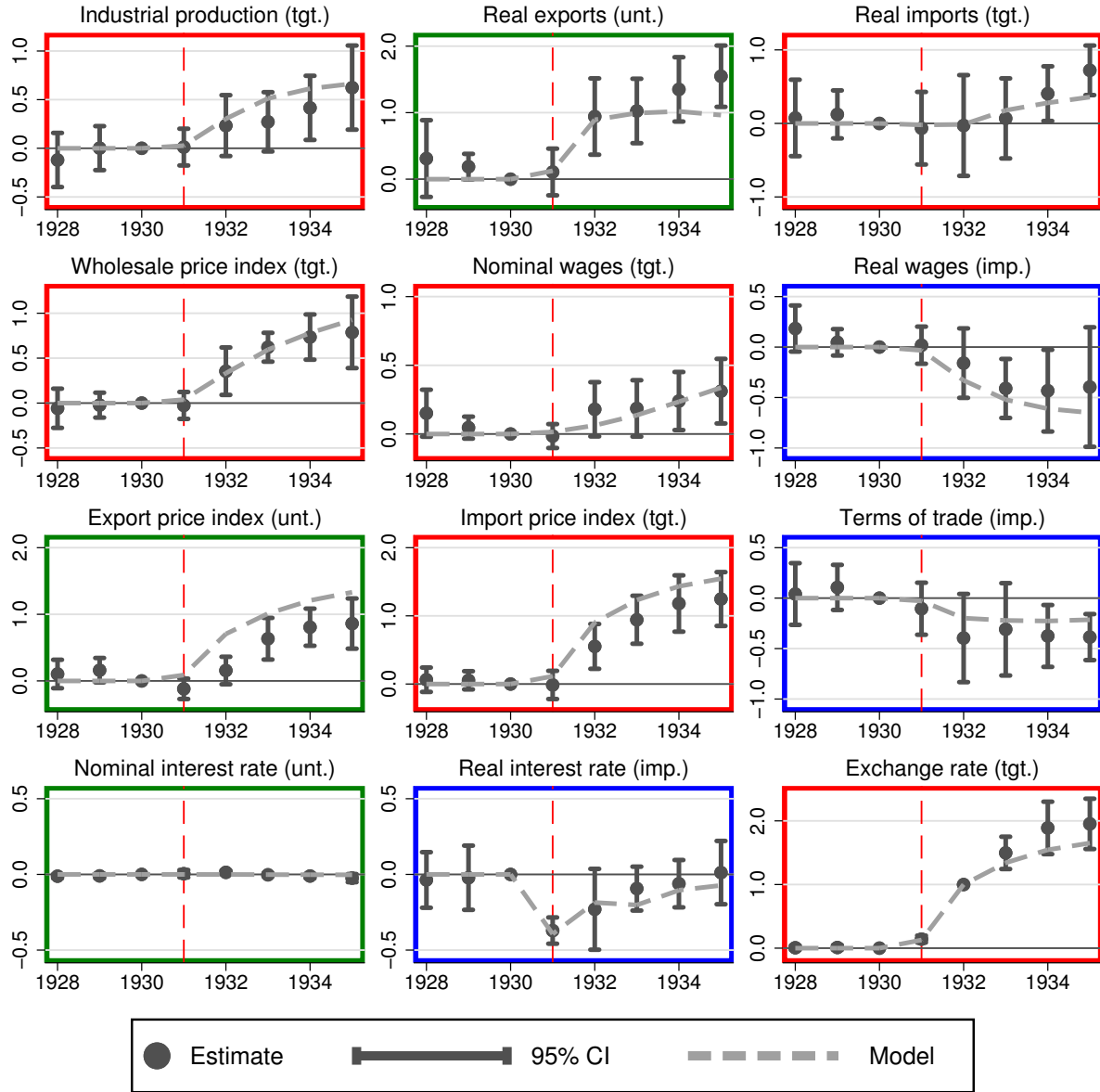
6.2 Counterfactual Experiment

Armed with these parameter values, I can now answer the counterfactual questions of interest: how did devaluation affect countries which devalued, and countries which didn’t? I consider a simple scenario: the world economy is in steady state; at time 0, half of the world devalues by 30% against gold.

Figure 6 illustrates that a foreign devaluation makes little quantitative difference for output (left-hand side panel). It lowers output for a few quarters but this effect is small and quickly dissipates. The explanation lies on the right-hand side panel. The devaluation entails a powerful real interest rate cut in countries that devalue. Through inter-temporal substitution, this cut stimulates consumption hence output. Beggar-thy-neighbour effects are dwarfed by this stimulus: the devaluation makes some countries more competitive, but those same countries consume more. On balance, the two effects compensate each other.

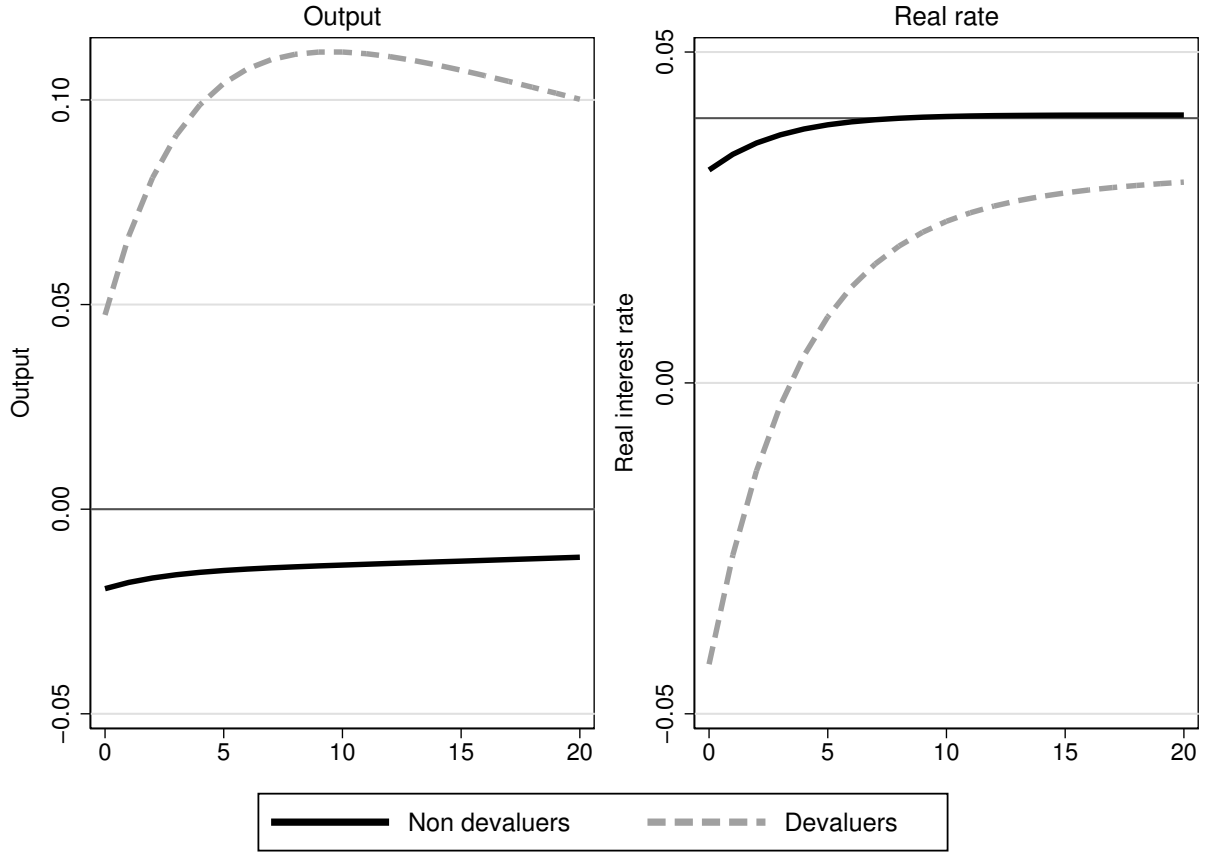
An equivalent way to state the conclusion is that the absolute effect is not very different from the relative effect. Britain was booming compared to France mostly because Britain was booming compared to a counterfactual scenario where it wouldn’t have devalued, not because France was collapsing.

Figure 5: Model fit (DD-IV)



Note: empirical and theoretical responses of relevant variables under estimated parameter values. The response circled in red (tgt.) are explicitly targeted. Those circled in green (untg.) are not. Those circled in blue (imp.) are implicitly targeted since they are functions of targeted moments.

Figure 6: Counterfactual analysis



Note: response of output and the real interest rate to a 30% devaluation by half of the world under parameter values estimated with the DD-IV cross-country moments of section 2.1 and the trade estimates of section 4. Those parameter values are shown in table 6.

The conclusion was not foregone: the model can imply large negative spillovers on foreign output. I show in the appendix that it does if I simplify the Kimball aggregator to a constant elasticity of substitution demand system with a single elasticity of substitution, which I set equal to 3. This model implies full pass-through and as much substitution between imports and domestic goods, as among imports. Other parameters are left unchanged compared to table 6. This calibration implies a sharp fall of non-devaluing countries' output: about 5% on impact with substantial persistence (figure A.7). This calibration, however, has unrealistic implications: by construction, it implies a pass-through of 1, which is inconsistent with the estimates of section 4. Moreover, it fits the cross-country evidence, particularly the response of imports, very poorly (figure A.8). This exercise illustrates that incomplete pass-through is an important ingredient of the result: a theme

that I come back to in the next section.

6.3 Identification

What, in the data, identifies the model, and determines the result? To understand it, I denote devaluing countries, non-devaluing countries, and world-level variables with superscripts D , N , and W . Up to a first-order approximation, the effect of a devaluation by a mass S^D of countries, can be written as

$$\text{devaluing countries:} \quad \hat{y}_t^D = (1 - S^D) (\hat{y}_t^D - \hat{y}_t^N) + \hat{y}_t^W \quad (25)$$

$$\text{non-devaluing countries:} \quad \hat{y}_t^N = -S^D (\hat{y}_t^D - \hat{y}_t^N) + \hat{y}_t^W \quad (26)$$

Those formulas do not depend on the details of the model — they're an immediate implication of symmetry and linearity. They are another way to express the identification problem: the term that features $\hat{y}_t^D - \hat{y}_t^N$ is pinned down by cross-sectional evidence, but \hat{y}_t^W is common across countries and soaked up by the time fixed effect. In so far as the model matches the relative effect, its conclusion is as good as the estimated effect on world-level output.

As table 6 showed, I estimate 6 parameters to match the cross-sectional evidence:

1. σ^{-1} : the inter-temporal elasticity of substitution (IES);
2. ξ : the Calvo parameter for wages;
3. α_1 : the curvature of the production function, which determines how sensitive prices are to changes in output;
4. ρ : the elasticity of substitution between imports and domestic goods;
5. θ : the elasticity of substitution among imports;
6. \bar{m} : the elasticity of the markup, which determines the exchange rate pass-through.

One can show analytically that the last three do not matter to world-level variables. The explanation is simple: those parameters are about substitution and strategic complementarity, hence only affect relative quantities and prices. On the other hand, the first three parameters matter. I focus on the first two here.¹⁴ Intuitively, they correspond to how powerful monetary policy is: how sensitive demand is to real interest rates, how sticky wages are...

¹⁴See figure A.9 for similar pictures about the third one.

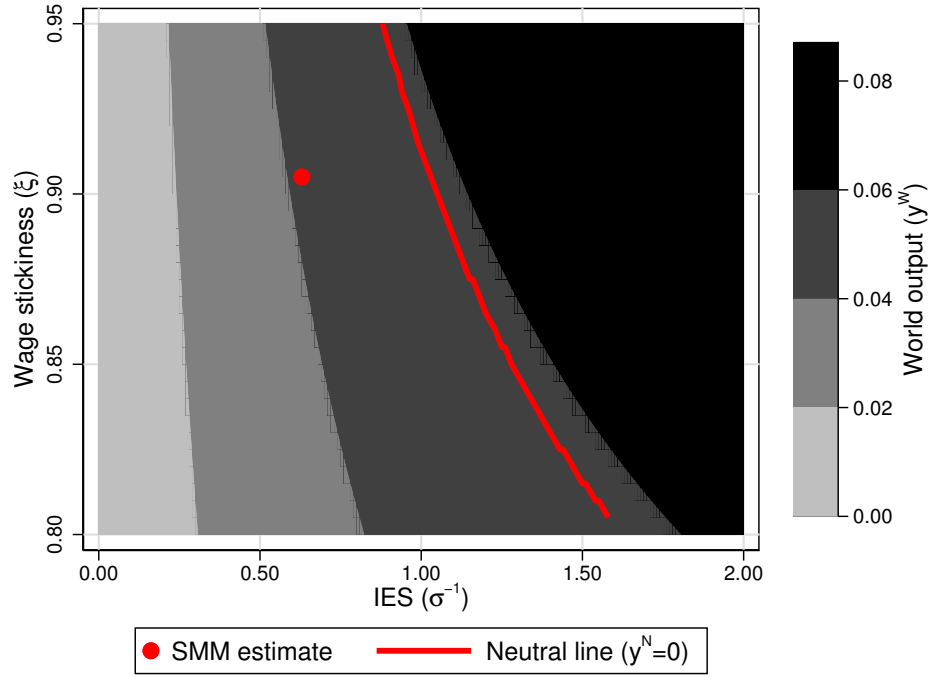
Figure 7.A shows the response of world output as a function of σ^{-1} and ξ . This response is computed as the mean increase in world output over the three years that follow a devaluation by half of the world. Darkest shades correspond to a bigger effect. The higher σ^{-1} and ξ , the bigger the effect. Indeed, a higher σ^{-1} means that consumption is more sensitive to a given cut in the real interest rate; a higher ξ means that wages are more sticky, hence nominal quantities adjust more slowly, and the real rate stays low for longer. The red line corresponds to the (σ^{-1}, ξ) pairs that imply no effect on non-devaluing countries (the neutral line). The red dot is the estimate that serves as the basis for the counter-factual experiments. It lands to the southwest of the red line, which explains why there is a negative effect on non-devaluing countries' output.

What, in the data, nails the red dot? Figure 7.B shows all combinations of σ^{-1} and ξ that imply the observed effect on output in devaluing countries relative to non-devaluing ones (the black line), and nominal wages in devaluing countries relative to non-devaluing ones (the gray line). The output line is downward-sloping because a higher σ^{-1} or a higher ξ act in the same direction: they raise relative output in devaluing countries. The intuition for why they do so is the same as that for world output: the higher these parameters, the less neutral money is. Since monetary stimulus happens in one type of countries but not in the other, a less neutral world means a bigger relative effect. As σ^{-1} and ξ act in the same direction, achieving a *given* relative response of output with a higher σ^{-1} requires a lower ξ . Relative wages effectively pin down ξ : the higher ξ , the faster wages adjust.

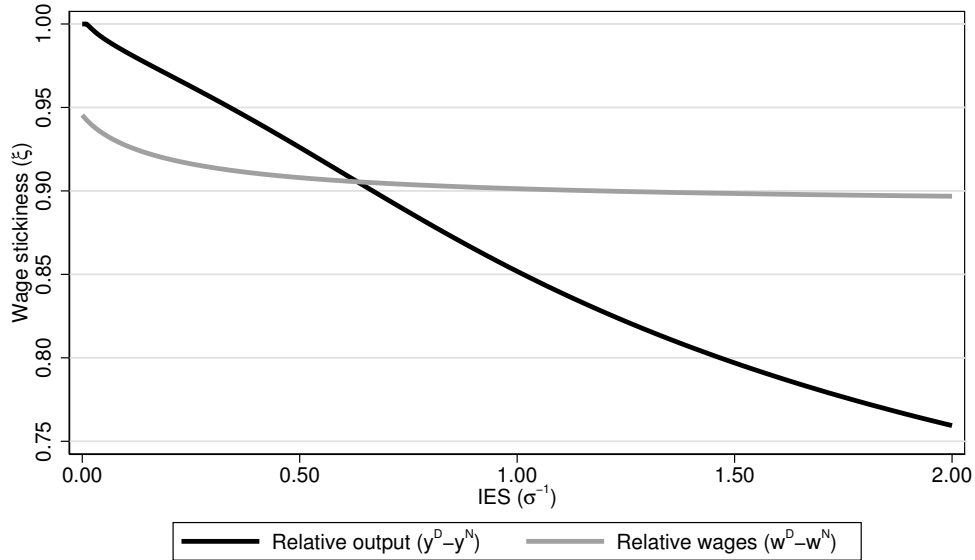
Was the work involving the granular trade data any useful? After all, the arguments invoked on figure 7 do not make any mention of the trade parameters. Perhaps matching relative output is sufficient. No, it's not. The position of the two lines on figure 7.B depends on those parameters. Figure 8.B shows what happens with complete pass-through ($\bar{m} = 0$). The solid black line moves to the left, becoming the dashed black line. Indeed, complete pass-through implies more responsive trade. Therefore, to obtain a given response of output, consumption needs to be less responsive, implying a lower σ^{-1} . Additionally, the solid gray line slightly moves up, becoming the dashed gray line. As a result, the estimated σ^{-1} goes down. The relevant point on figure 8.A becomes the blue triangle. Since σ^{-1} is lower, the stimulus on world output is smaller by more than 2 percentage points. Therefore, the effect on non-devaluing countries is more negative by the same amount (figure A.10).

Figure 7: Identification

Panel A: Response of World Output as a function of σ^{-1} and ξ



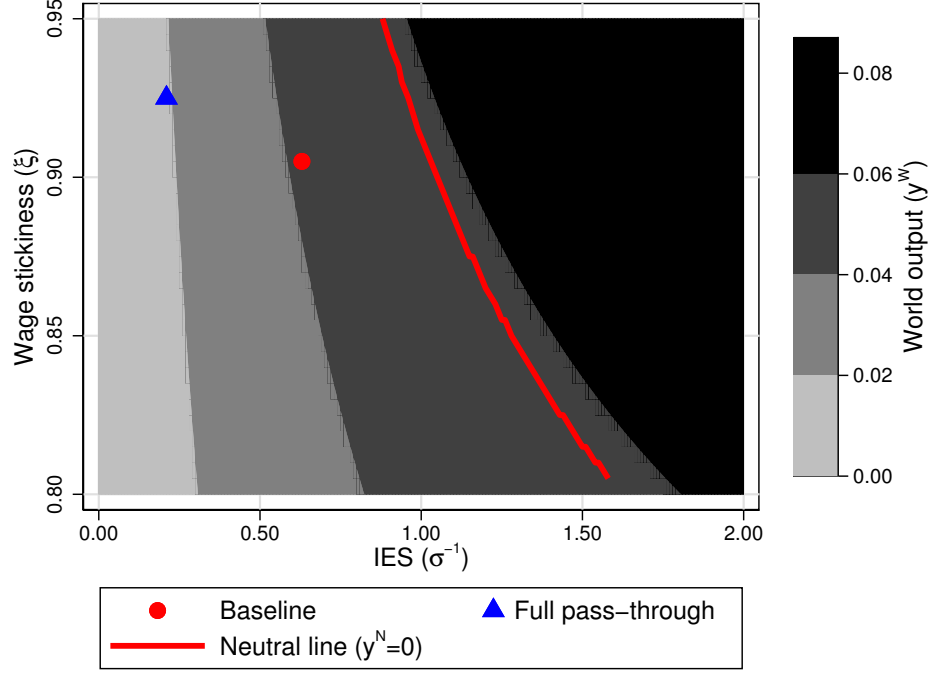
Panel B: Identification of σ^{-1} and ξ



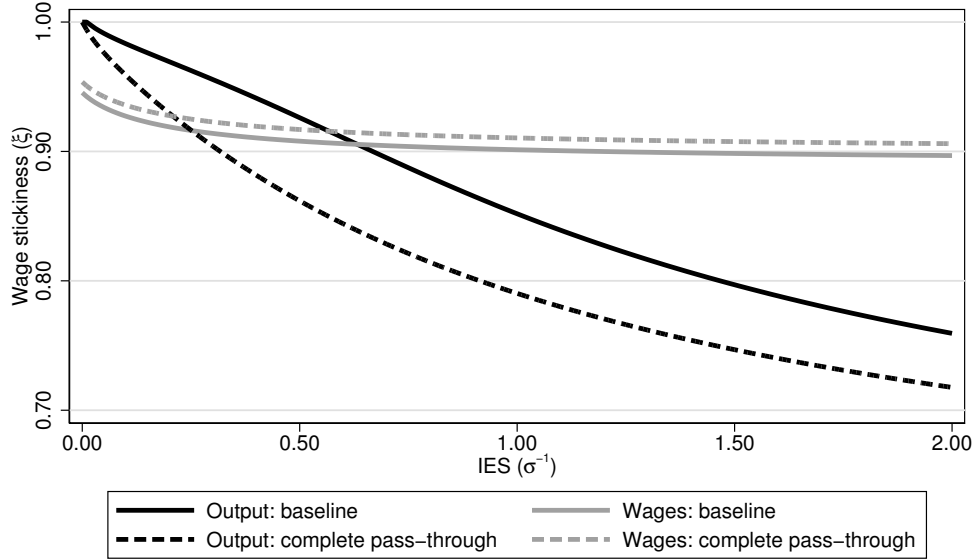
Note: panel A shows the response of world output over three years as a function of the intertemporal elasticity of substitution (σ^{-1}) and the Calvo wage parameter (ξ). The darker the color, the bigger the effect on world output. The red dot corresponds to the (σ^{-1}, ξ) pair returned by the estimation. The red line draws the (σ^{-1}, ξ) pairs which imply no effect on non-devaluing countries' output. Panel B shows how the estimated moments pin down the two parameters: the black line is the (σ^{-1}, ξ) pairs which imply the observed relative effect on output, the gray line the (σ^{-1}, ξ) pairs which imply the observed relative effect on nominal wages. The intersection of the two lines corresponds to the red dot in panel A.

Figure 8: Identification: complete pass-through

Panel A: Response of World Output as a function of σ^{-1} and ξ



Panel B: Identification of σ^{-1} and ξ



Note: see note for figure 7. In panel A, the blue triangle is the (σ^{-1}, ξ) pair that matches the relative effects on output and wages, assuming full pass-through. This pair is pinned down by the intersection of the two dotted lines in panel B.

6.4 Sufficient Statistics Approach

In appendix G, I develop a sufficient statistics approach to relax assumptions made on monetary policy, thus showing that the main conclusion doesn't rely on the specific parametric assumptions of my model. I show analytically that the effect of devaluation in devaluing countries depends on three objects: the relative effect on output, the time path of the world real interest rate, and structural parameters that determine how sensitive output is to the real rate. The first object is pinned down by the cross-country estimation exercises, the third one by the structural estimation of the model. But the time path of the world real rate depends on the specification of monetary policy. To provide empirical discipline on this object, I go back to the data and use the HFI shocks that I constructed in the first part of the paper. Without a time fixed effect, I can estimate the *absolute* effect of devaluation on the real rate, instead of its relative effect. Indeed, those shocks are identified *time series* shocks and I can use them for identification in the time series as well as in the cross-section. The world real interest rate falls after a subset of countries devalues. I then feed the path of the real rate into the model to simulate it. The advantage of this approach is that it does not require specifying world-level monetary policy. The result confirms the main message: devaluation has a small effect in countries that do not devalue, but a large one in those that do.

7 Conclusion

In this paper, I identified key moments from cross-country and granular trade data. Cross-country evidence shows that devaluation stimulated output, inflation, and trade in devaluing countries, compared to countries that did not. The effect on output is large: a 30% devaluation stimulated it by up to 14% over the next three years. Running against the competitive devaluation presumption, real imports did not go down. In fact, they went up after a few years. Trade evidence points toward an elasticity of substitution across foreign varieties between 2 and 4, and a pass-through of the exchange rate of about 0.4. I then used those moments to estimate a New Keynesian model with variable markups. The takeaway of this exercise is that devaluations had quantitatively small effects on the output of countries that did not devalue. Equivalently, the absolute effect of a devaluation is roughly similar to its relative effect. To the question “canst thou beggar thy neighbour,” this episode suggests a negative answer.

References

- Accominotti, O. (2009). The sterling trap: foreign reserves management at the bank of france, 1928–1936. *European Review of Economic History* 13, 349—376.
- Accominotti, O. (2012). London merchant banks, the central european panic, and the sterling crisis of 1931. *Journal of Economic History* 72(1), 1–43.
- Accominotti, O., J. Cen, D. Chambers, and I. W. Marsh (2019). Currency regimes and the carry trade. *Journal of Financial and Quantitative Analysis* 54(5), 2233–2260.
- Acosta, M. and L. Cox (2021). Hysteresis in the U.S. tariff code: Origins and implications. Work in progress.
- Albers, T. (2019). Currency devaluations and beggar-my-neighbour penalties. *Economic History Review* 73(1), 233–257.
- Albers, T. N. H. (2018). The prelude and global impact of the great depression: Evidence from a new macroeconomic dataset. *Explorations in Economic History* 70, 150–163.
- Alessandria, G., S. Pratap, and V. Z. Yue (2018). Export growth and aggregate dynamics in large devaluations. Working paper.
- Amiti, M., O. Itskhoki, and J. Konings (2014a). Importers, exporters, and exchange rate disconnect. *American Economic Review* 104(7), 1942–1978.
- Amiti, M., O. Itskhoki, and J. Konings (2014b). Importers, exporters, and exchange rate disconnect. *American Economic Review* 104(7), 1942–1978.
- Atkeson, A. and A. Burstein (2008). Pricing-to-market, trade costs, and international relative prices. *American Economic Review* 98(5), 1998–2031.
- Auclert, A., M. Rognlie, and L. Straub (2020). Micro jumps, macro humps: Monetary policy and business cycles in an estimated HANK model. Working paper.
- Bagliano, F. C. and C. A. Favero (1999). Information from financial markets and VAR measures of monetary policy. *European Economic Review* 43, 825–837.
- Barakchian, S. M. and C. Crowe (2013). Monetary policy matters: Evidence from new shocks data. *Journal of Monetary Economics* 60, 950–966.
- Benigno, G. and P. Benigno (2006). Designing targeting rules for international monetary policy cooperation. *Journal of Monetary Economics* 53, 473–506.
- Bernanke, B. and H. James (1991). The gold standard, deflation, and financial crisis in the Great Depression: An international comparison. In R. G. Hubbard (Ed.), *Financial Markets and Financial Crises*. Chicago: University of Chicago Press.
- Bernanke, B. S. (1983). Non-monetary effects of the financial crisis in the propagation of the great depression. *American Economic Review* 73(3), 257–276.
- Bernanke, B. S. and K. Carey (1996). Nominal wage stickiness and aggregate supply in the Great Depression. *Quarterly Journal of Economics* 111(3), 853–883.
- Bond, E. W., M. J. Crucini, T. Potter, and J. Rodrigue (2013). Misallocation and productivity effects of the Smoot–Hawley tariff. *Review of Economic Dynamics* 16, 120–134.

- Broda, C. and D. E. Weinstein (2006). Globalization and the gains from variety. *Quarterly Journal of Economics* 121(2), 541–585.
- Burstein, A., M. Eichenbaum, and S. Rebelo (2005). Large devaluations and the real exchange rate. *Journal of Political Economy* 113(4), 742–784.
- Burstein, A., M. Eichenbaum, and S. Rebelo (2007). Modeling exchange rate passthrough after large devaluations. *Journal of Monetary Economics* 54, 346–368.
- Burstein, A. and G. Gopinath (2014). International prices and exchange rates. In G. Gopinath, E. Helpman, and K. Rogoff (Eds.), *Handbook of International Economics*, Chapter 7, pp. 391–451. Amsterdam: Elsevier.
- Caballero, R. J., E. Farhi, and P.-O. Gourinchas (2021). Global imbalances and policy wars at the zero lower bound. *Review of Economic Studies* 88(6), 2570–2621.
- Caliendo, L. and F. Parro (2015). Estimates of the trade and welfare effects of NAFTA. *Review of Economic Studies* 82, 1–44.
- Campa, J. M. (1990). Exchange rates and economic recovery in the 1930s: An extension to Latin America. *Journal of Economic History* 50(3), 677–682.
- Candia, B. and M. Pedemonte (2021). Export-led decay: The trade channel in the gold standard era. Working paper.
- Christiano, L. J., M. Eichenbaum, and C. L. Evans (2005). Nominal rigidities and the dynamic effects of a shock to monetary policy. *Journal of Political Economy* 113(1), 1–45.
- Clarida, R., J. Galí, and M. Gertler (2002). A simple framework for international monetary policy analysis. *Journal of Monetary Economics* 49, 879–904.
- Cochrane, J. H. and M. Piazzesi (2002). The Fed and interest rates: A high-frequency identification. *American Economic Review* 92(2), 90–95.
- Cohen-Setton, J., J. K. Hausman, and J. F. Wieland (2017). Supply-side policies in the Depression: Evidence from France. *Journal of Money, Credit and Banking* 49(2–3), 273–317.
- Colacito, R. and M. M. Croce (2013). International asset pricing with recursive preferences. *Journal of Finance* 68(6), 2651–2686.
- Corsetti, G., L. Dedola, and S. Leduc (2011). Optimal monetary policy in open economies. In B. M. Friedman and M. Woodford (Eds.), *Handbook of Monetary Economics*, Volume 3B, Chapter 16, pp. 861–933. Amsterdam: Elsevier.
- Corsetti, G. and P. Pesenti (2005). International dimensions of optimal monetary policy. *Journal of Monetary Economics* 52, 281–305.
- Crucini, M. J. (1994). Sources of variation in real tariff rates: The United States, 1900–1940. *American Economic Review* 84(3), 732–743.
- David, T. (1995). Un indice de la production industrielle de la Suisse durant l’entre-deux-guerres. *Revue Suisse d’histoire* 45, 109–130.
- de Bromhead, A., A. Fernihough, M. Lampe, and K. H. O’Rourke (2018). When Britain turned inward: The impact of interwar British protection. *American Economic Review* 109(2), 325–352.

- Devereux, M. B. and C. Engel (2003). Monetary policy in the open economy revisited: Price setting and exchange-rate flexibility. *Review of Economic Studies* 70(4), 765–783.
- Economic Commission for Latin America (1951). *Economic Survey of Latin America 1940*. New York: United Nations.
- Edwards, S. (2018). *American Default*. Princeton and Oxford: Princeton University Press.
- Eggertsson, G. B. (2008). Great expectations and the end of the Depression. *American Economic Review* 98(4), 1476–1516.
- Eichengreen, B. (1992). *Golden Fetters: The Gold Standard and the Great Depression, 1919–1939*. Oxford: Oxford University Press.
- Eichengreen, B. and D. A. Irwin (2010). The slide to protectionism in the Great Depression: Who succumbed and why? *Journal of Economic History* 70(4), 871–897.
- Eichengreen, B. and J. Sachs (1985). Exchange rates and economic recovery in the 1930s. *Journal of Economic History* 45(4), 925–946.
- Ellison, M., S. S. Lee, and K. H. O’Rourke (2021). The ends of 30 big depressions. NBER Working Paper No. 27586.
- Engel, C. (2009). Currency misalignments and optimal monetary policy: A reexamination. NBER Working Paper No. 14829.
- Fama, E. F. (1984). Forward and spot exchange rates. *Journal of Monetary Economics* 14, 319–338.
- Faust, J., E. T. Swanson, and J. H. Wright (2004). Identifying VARs based on high frequency futures data. *Journal of Monetary Economics* 51, 1107–1131.
- Feenstra, R. C. (1994). New product varieties and the measurement of international prices. *American Economic Review* 84(1), 157–177.
- Feenstra, R. C., P. Luck, M. Obstfeld, and K. N. Russ (2018). In search of the Armington elasticity. *Review of Economics and Statistics* 100(1), 135–150.
- Fleming, J. M. (1962). Domestic financial policies under fixed and under floating exchange rates. *IMF Staff Papers* 9(3), 369–380.
- Flood, R. P. and A. K. Rose (1996). Fixes: Of the forward discount puzzle. *Review of Economics and Statistics* 78(4), 748–752.
- Fouquin, M. and J. Hugot (2016). Two centuries of bilateral trade and gravity data: 1827–2014. CEPII Working Paper No. 2016-14.
- Friedman, M. and A. J. Schwartz (1963). *A Monetary History of the United States, 1867–1960*. Princeton: Princeton University Press.
- Galí, J. (2015). *Monetary Policy, Inflation, and the Business Cycle: An Introduction to the New Keynesian Framework and Its Applications* (Second ed.). Princeton: Princeton University Press.
- Gertler, M. and P. Karadi (2015). Monetary policy surprises, credit costs, and economic activity. *American Economic Journal: Macroeconomics* 7(1), 44–76.
- Goldstein, M. and M. S. Khan (1985). Income and price effects in foreign trade. In R. W. Jones and P. B. Kenen (Eds.), *Handbook of International Economics*, Volume 2, Chapter 20, pp. 1041–1105. Amsterdam: Elsevier.

- Gopinath, G. and O. Itskhoki (2010). Frequency of price adjustment and pass-through. *Quarterly Journal of Economics* 125(2), 675–727.
- Hausman, J. K., P. W. Rhode, and J. F. Wieland (2019). Recovery from the Great Depression: The farm channel in spring 1933. *American Economic Review* 109(2), 427–472.
- Head, K. and J. Ries (2001). Increasing returns versus national product differentiation as an explanation for the pattern of U.S.-Canada trade. *American Economic Review* 91(4), 858–876.
- Hsieh, C.-T. and C. Romer (2006). Was the federal reserve constrained by the gold standard during the Great Depression? evidence from the 1932 open market purchase program. *Journal of Economic History* 66(1), 140–176.
- Imbs, J. and I. Méjean (2015). Elasticity optimism. *American Economic Journal: Macroeconomics* 7(3), 43–83.
- Irwin, D. A. (1998a). Changes in U.S. tariffs: The role of import prices and commercial policies. *American Economic Review* 88(4), 1015–1026.
- Irwin, D. A. (1998b). The Smoot-Hawley tariff: A quantitative assessment. *Review of Economics and Statistics* 80(2), 326–334.
- Itskhoki, O. and D. Mukhin (2019). Exchange rate disconnect in general equilibrium. Working paper.
- Jordà, O. (2005). Estimation and inference of impulse responses by local projections. *American Economic Review* 95(1), 161–182.
- Jordà, O., M. Schularick, and A. M. Taylor (2015). Betting the house. *Journal of International Economics* 96, S2–S18.
- Kimball, M. S. (1995). The quantitative analytics of the basic neomonetarist model. *Journal of Money, Credit and Banking* 27(4), 1241–1277.
- Klenow, P. J. and J. L. Willis (2016). Real rigidities and nominal price changes. *Economica* 83, 443–472.
- Kohn, D., F. Leibovici, and M. Szkup (2020). Financial frictions and export dynamics in large devaluations. *Journal of International Economics* 122, 1–13.
- League of Nations (1936, July). *Monthly Bulletin of Statistics*. Geneva: League of Nations.
- League of Nations (1938a). *Review of World Trade 1937*. Geneva: League of Nations.
- League of Nations (1938b). *Statistical Year-Book, 1937/38*. Geneva: League of Nations.
- Mathy, G. P. and C. M. Meissner (2011). Business cycle co-movement: Evidence from the Great Depression. *Journal of Monetary Economics* 58, 362–372.
- Mattoo, A., P. Mishra, and A. Subramanian (2017). Beggar-thy-neighbor effects of exchange rates: A study of the renminbi. *American Economic Journal: Economic Policy* 9(4), 333–366.
- Mitchell, B. R. (2003). *International Historical Statistics: Europe, 1750–2000* (Fifth ed.). New York: Palgrave Macmillan.
- Mitchener, K. J. and K. Wandschneider (2015). Capital controls and recovery from the financial crisis of the 1930s. *Journal of International Economics* 95, 188–201.

- Mundell, R. A. (1963). Capital mobility and stabilization policy under fixed and flexible exchange rates. *Canadian Journal of Economics and Political Science* 29(4), 475–485.
- Nakamura, E. and J. Steinsson (2018). High-frequency identification of monetary non-neutrality: The information effect. *Quarterly Journal of Economics* 133(3), 1283–1330.
- Obstfeld, M. and K. Rogoff (1995). Exchange rate dynamics redux. *Journal of Political Economy* 103(3), 624–660.
- Ramey, V. A. (2016). Macroeconomic shocks and their propagation. In J. B. Taylor and H. Uhlig (Eds.), *Handbook of Macroeconomics*, Volume 2A, pp. 71–162. Amsterdam: Elsevier.
- Rodnyansky, A. (2019). (Un)competitive devaluations and firm dynamics. Working paper.
- Romalis, J. (2007). NAFTA’s and CUSFTA’s impact on international trade. *Review of Economics and Statistics* 89(3), 416–435.
- Rose, A. K. (2018). Currency wars? Unconventional monetary policy does not stimulate exports. *Journal of Money, Credit and Banking*. Forthcoming.
- Ruhl, K. J. (2008). The international elasticity puzzle. Unpublished.
- Schmitt-Grohé, S. and M. Uribe (2003). Closing small open economy models. *Journal of International Economics* 61, 163–185.
- Svensson, L. E. O. and S. van Wijnbergen (1989). Excess capacity, monopolistic competition, and international transmission of monetary disturbances. *Economic Journal* 99(397), 785–805.
- Thorp, R. (1984). *From an Export-led to an Import-substituting Economy: Chile*. New York: St. Martin’s Press.
- United States Department of Commerce (1930–1934). *Foreign Commerce and Navigation of the United States*. Washington: Government Printing Office. 1929 to 1933 issues.
- Verhoogen, E. (2008). Trade, quality upgrading, and wage inequality in the mexican manufacturing sector. *Quarterly Journal of Economics* 123(2), 489–530.
- Weiss, C. (2020). Contractionary devaluation risk: Evidence from the free silver movement, 1878–1900. *Review of Economics and Statistics* 102(4), 705–720.

APPENDIX

A Additional Tables

Table A.1: Country Samples

Difference-in-difference (section 2.1)		
Group	Countries	Comment
Treatment	Austria, Canada, Denmark, Finland, Japan, Mexico, Norway, Salvador, Sweden, United Kingdom	Devaluation in 1931
Control	France, Germany, Hungary, Netherlands, Poland, Switzerland	No devaluation before 1936
High frequency identification (section 2.2)		
	Belgium, France, Germany, Italy, Netherlands, United Kingdom, United States, Switzerland	Countries for which forward exchange rate data is available
US imports (section 4)		
Specification	Countries	Comment
1932, 1933	Austria, British India, British Malaya, Canada, Denmark, Egypt, Finland, Japan, Mexico, Norway, Portugal, Salvador, Sweden, United Kingdom	Devaluation in 1931
1932, 1933	Albania, Belgium, Bulgaria, Czechoslovakia, France, Germany, Hungary, Italy, Latvia, Lithuania, Netherlands, Poland and Danzig, Rumania, Switzerland	No devaluation before 1934
1932	Estonia, South Africa	Devaluation in 1933
none	Argentina, Australia, Bolivia, Brazil, New Zealand, Uruguay, Venezuela	Devaluation before 1931
none	Chile, Costa Rica, Colombia, Ecuador, Greece, Peru, Siam, Yugoslavia	Devaluation in 1932
none	China, Hong Kong, Iran	Silver standard
none	Spain, Turkey	Floating currency
none	Cuba, Philippine Islands	Preferential tariff rates
World real rate (appendix G)		
Specification	Countries	Comment
WPI	Austria, Belgium, Czechoslovakia, Denmark, Estonia, Finland, France, Germany, Hungary, Japan, Netherlands, Norway, Poland, South Africa, Sweden, Switzerland, United Kingdom, United States	Countries for which nominal interest rate and price index consistently available
CPI	Same minus Denmark, plus Australia and Greece	from 1925 to 1936

Table A.2: Data Sources

Section	Variable	Source
2, 4	Spot exchange rate	League of Nations (1938b)
2.1	Industrial production	Mitchell (2003) for Europe except Latvia and Switzerland, David (1995) for Switzerland, Thorp (1984) for Latin America and League of Nations (1938b) otherwise
2.1	Exports and imports	League of Nations (1936): special trade, merchandise only, measured in metric tons for Belgium and the Netherlands, Economic Commission for Latin America (1951) for Argentina, Brazil and Mexico, League of Nations (1938a) otherwise
2.1	Wholesale price indices	League of Nations (1936)
2.1	Export and import price indices	League of Nations (1938a)
2.2	Forward exchange rate	<i>Financial Times</i>
2.2	Monthly data	Albers (2018), Mitchener and Wandschneider (2015)
4	US imports	United States Department of Commerce (1934)
G.3	GDP and trade weights	Fouquin and Hugot (2016)

Table A.3: P-values for DD Exercise

	Baseline	Tariffs	Government
IV			
Industrial production	0.00	0.03	0.07
Real exports	0.00	0.19	0.00
Real imports	0.00	0.07	0.00
Wholesale price index	0.00	0.00	0.00
Consumer price index	0.00	0.04	0.04
Export price index	0.00	0.30	0.00
Import price index	0.00	0.00	0.00
Nominal wages	0.00	0.12	0.10
Real wages	0.00	0.73	0.03
Terms of trade	0.00	0.27	0.26
Nominal interest rate	0.26	0.27	0.13
Real interest rate	0.02	0.16	0.03
OLS			
Industrial production	0.00	0.00	0.00
Real exports	0.00	0.00	0.00
Real imports	0.00	0.00	0.02
Wholesale price index	0.00	0.00	0.00
Consumer price index	0.00	0.00	0.00
Export price index	0.00	0.03	0.00
Import price index	0.00	0.00	0.00
Nominal wages	0.00	0.16	0.01
Real wages	0.00	0.00	0.20
Terms of trade	0.00	0.89	0.04
Nominal interest rate	0.25	0.27	0.17
Real interest rate	0.00	0.04	0.00

Note: p-value of a test of the joint significance of the 1932 to 1935 coefficients in equation (1) with no controls (baseline), controlling for the change in tariffs (tariff), or real government revenues and expenditures (government). In the IV rows, the change in the exchange rate is instrumented with a devaluation dummy. Standard errors are clustered at the country level, and bootstrapped with 2,000 replications. Point estimates are shown on figures A.2 and A.3.

Table A.4.1: Shocks

Date	# articles	Event	Shock
Panel A: Belgium (mean=0.2, sd=1.3)			
14nov1934	10	Catholics and liberals against devaluation	-0.007
17mar1935	14	Devaluation	+0.116
18mar1935	12		
19mar1935	22		
20mar1935	23		
21mar1935	8		
23mar1935	10		
25mar1935	11		
26mar1935	10		
27mar1935	12		
28mar1935	12		
29mar1935	25		
30mar1935	45		
31mar1935	24		
01apr1935	24		
02apr1935	14		
08apr1935	8		
Panel B: France (mean=1.3, sd=3.1)			
27may1935	21	Fall of Flandin's cabinet	-0.046
31may1935	21		
04jun1935	21		
05jun1935	23		
11may1936	22	Blum's devaluation speech	-0.013
05jun1936	21	Strikes	×
25sep1936	30	Devaluation	+0.234
26sep1936	60		
27sep1936	73		
28sep1936	59		
29sep1936	33		
30sep1936	31		
01oct1936	25		
02oct1936	28		
03oct1936	28		
04oct1936	27		
05oct1936	24		

Notes: shocks selected by the procedure described in section 2.2. The numbers in parenthesis are the mean and standard deviation of the daily number of articles for the selected country.

Table A.4.2: Shocks

Date	# articles		Event	Shock
Panel C: Germany (mean=0.7, sd=1.3)				
13jul1931	9	}	Exchange controls	-0.067
15jul1931	17			
16jul1931	10			
21jul1931	9			
22jul1931	9			
29sep1931	9		Stock exchange closed indefinitely	n.a.
05dec1931	13	}	Devaluation rumors	×
13jun1934	10			
15jun1934	20			
27sep1936	9	}	Gold bloc demise	×
28sep1936	10			
01oct1936	9			
Panel D: Italy (mean=0.3, sd=1.0)				
10dec1934	9	}	Exchange controls	-0.018
11dec1934	7			
23jul1935	8	}	Gold coverage ratio suspended	+0.042
24jul1935	16			
25jul1935	9			
20nov1935	11		Gold-buying monopoly	-0.002
29nov1935	8		International tensions	×
27sep1936	7	}	Gold bloc demise	×
28sep1936	9			
04oct1936	8			
06oct1936	38	}	Devaluation	n.a.
07oct1936	12			
09oct1936	8			
10nov1936	8		Austro-Italian trade pact	×

Note: shocks selected by the procedure described in section 2.2. The numbers in parenthesis are the mean and standard deviation of the daily number of articles for the selected country.

Table A.4.3: Shocks

Date	# articles	Event	Shock
Panel E: Netherlands (mean=0.2, sd=0.9)			
27jun1933	8	Devaluation rumors	×
06apr1935	8	} Gold bloc reaffirms commitment	-0.023
07apr1935	7		
08apr1935	12		
04jun1935	9	Pro-devaluation minister resigns	-.009
24jul1935	8	} New government	-0.014
25jul1935	13		
26jul1935	8		
27jul1935	13		
28jul1935	11		
29jul1935	15		
17sep1935	7	Bank rate hike	-0.003
18sep1935	8	Deflationary budget	-0.009
26sep1935	8	Devaluation rumors	×
04feb1936	6	Pro-devaluation speech by former minister	×
27sep1936	18	} Devaluation	+0.175
28sep1936	16		
29sep1936	7		
Panel F: Switzerland (mean=0.1, sd=0.5)			
21apr1933	4	US gold embargo	×
23mar1934	5	} Pro-gold minister resigns	+0.003
24mar1934	4		
08apr1935	4	Gold bloc reaffirms commitment	-0.020
03jun1935	8	} Devaluation proposal rejected	-0.011
04jun1935	4		
28oct1935	4	General elections	-0.003
27sep1936	14	} Devaluation	+0.321
28sep1936	14		
30sep1936	6		
01oct1936	4		
09oct1936	4		

Note: shocks selected by the procedure described in section 2.2. The numbers in parenthesis are the mean and standard deviation of the daily number of articles for the selected country.

Table A.4.4: Shocks

Date	# articles	Event	Shock
Panel G: UK (mean=1.2, sd=2.1)			
21sep1931	40	Pound devaluation	+0.175
22sep1931	58		
23sep1931	26		
27sep1931	18		
28sep1931	27		
26nov1932	15	War debt discussions	×
02dec1932	16		
22apr1933	15	Dollar devaluation	×
26apr1933	17	Exchange equalisation fund increased	+0.018
26sep1936	18	Gold bloc demise	×
27sep1936	15		
Panel H: US (mean=18.3, sd=12.7)			
22sep1931	128	Pound devaluation	×
21apr1933	137	Gold embargo	+0.095
23apr1933	96		
16jan1934	156	Dollar devaluation	+0.022
02feb1934	107		
26sep1936	106	Gold bloc demise	×
27sep1936	138		
28sep1936	113		
29sep1936	111		

Note: shocks selected by the procedure described in section 2.2. The numbers in parenthesis are the mean and standard deviation of the daily number of articles for the selected country.

Table A.5: Parameters (HFI)

Parameter	Value	Concept	Target or source
Calibrated			
β	0.99	Discount rate	Annual interest rate of 4%
ι	0.8	Habit	Eggertsson (2008)
χ	1	Concavity of utility for money	Standard
γ	0.1	Openness	British export-to-GDP ratio
α_2	0.37	Profit share	Labor share of 2/3
η	21	Elasticity of labor demand	Christiano et al. (2005)
ν	0.0007	Stationarity-inducing device	Schmitt-Grohé and Uribe (2003)
g^*	0.14	Gold to output ratio	British gold-reserves-to-GDP ratio
Estimated			
σ^{-1}	4.67 (0.86)	IES	Industrial production
α_1	0.21 (0.06)	Curvature of production function	WPI
ξ	0.78 (0.04)	Calvo wage parameter	Nominal wages
ρ	1.2 (0.24)	Macro trade elasticity	Nominal imports
θ	3.08 (0.12)	Micro trade elasticity	US imports
\bar{m}	1.34 (0.02)	Markup elasticity	Pass-through

Note: calibrated and estimated parameters. The estimation follows the procedure described in section 6.1. The targets are the HFI cross-country moments of section 2.2 and the trade estimates of section 4.

Table A.6: Effect on the world real rate (trade-weighted)

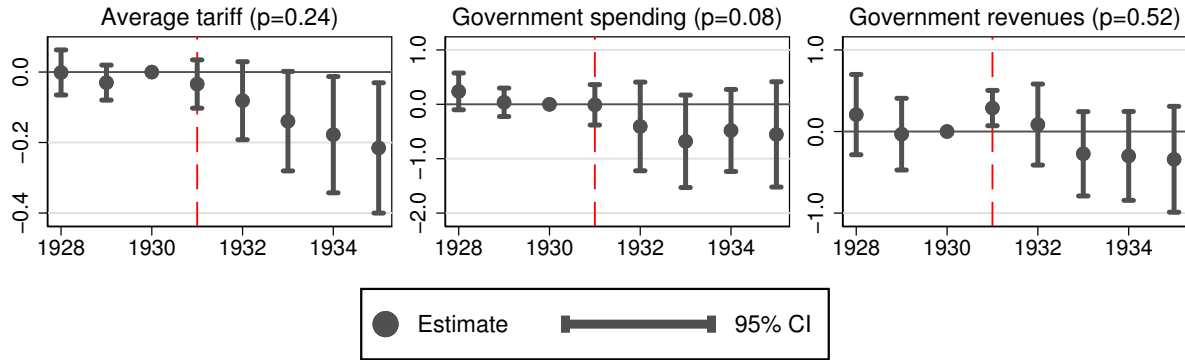
	(1) Numerator	(2) Denominator	(3) Ratio
Real rate (WPI)	−0.412** (0.172)	1.474*** (0.352)	−0.280** (0.133)
Real rate (CPI)	−0.363* (0.191)	1.120*** (0.389)	−0.324* (0.176)

Note: point estimates and standard deviations retrieved from averaging the coefficients whose estimations is detailed in section G.2. The numerator is the average coefficient for the relevant variable, the denominator is the average coefficient for the exchange rate. The ratio is the ratio of the two.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

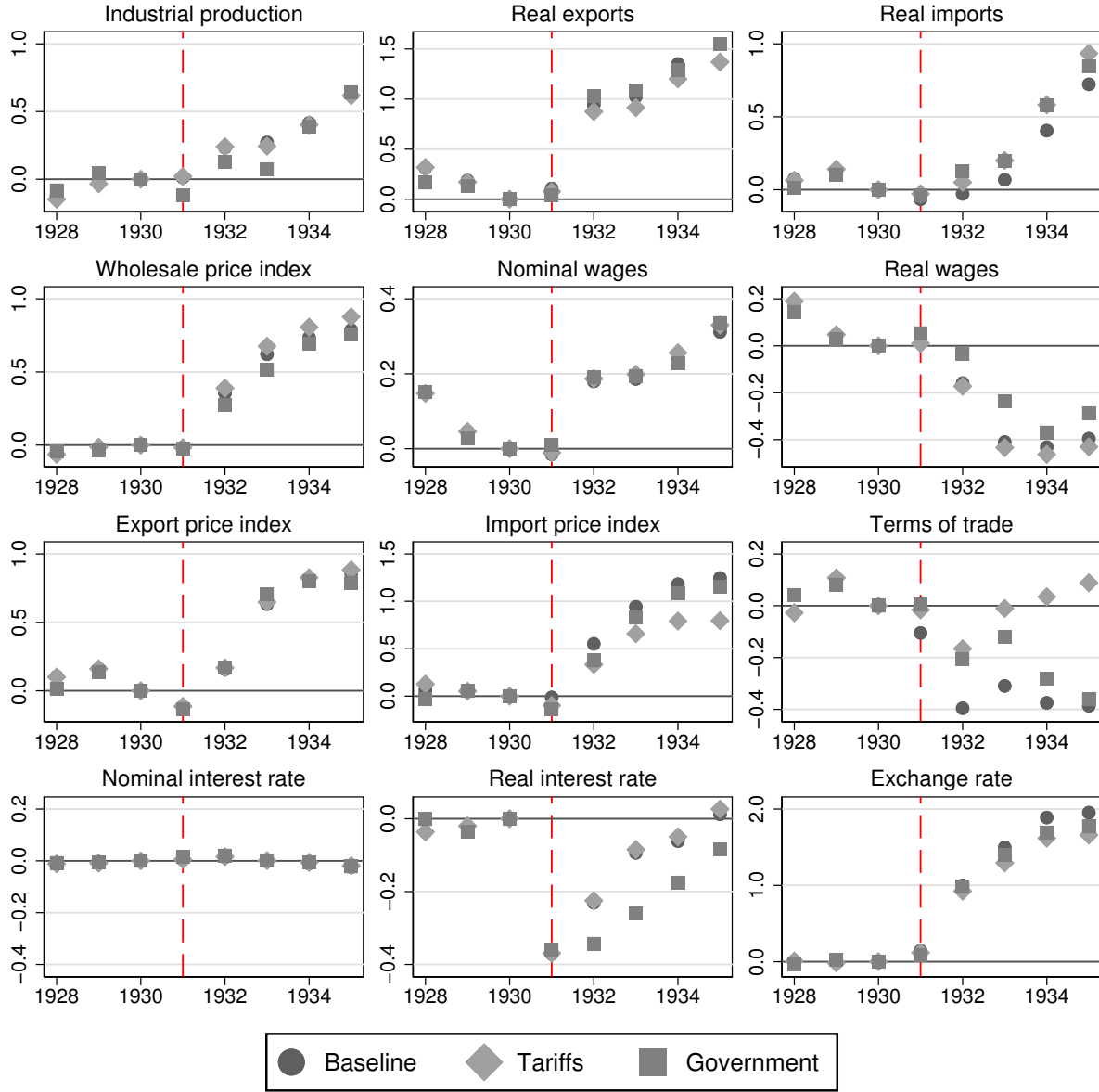
B Additional Figures

Figure A.1: Government variables



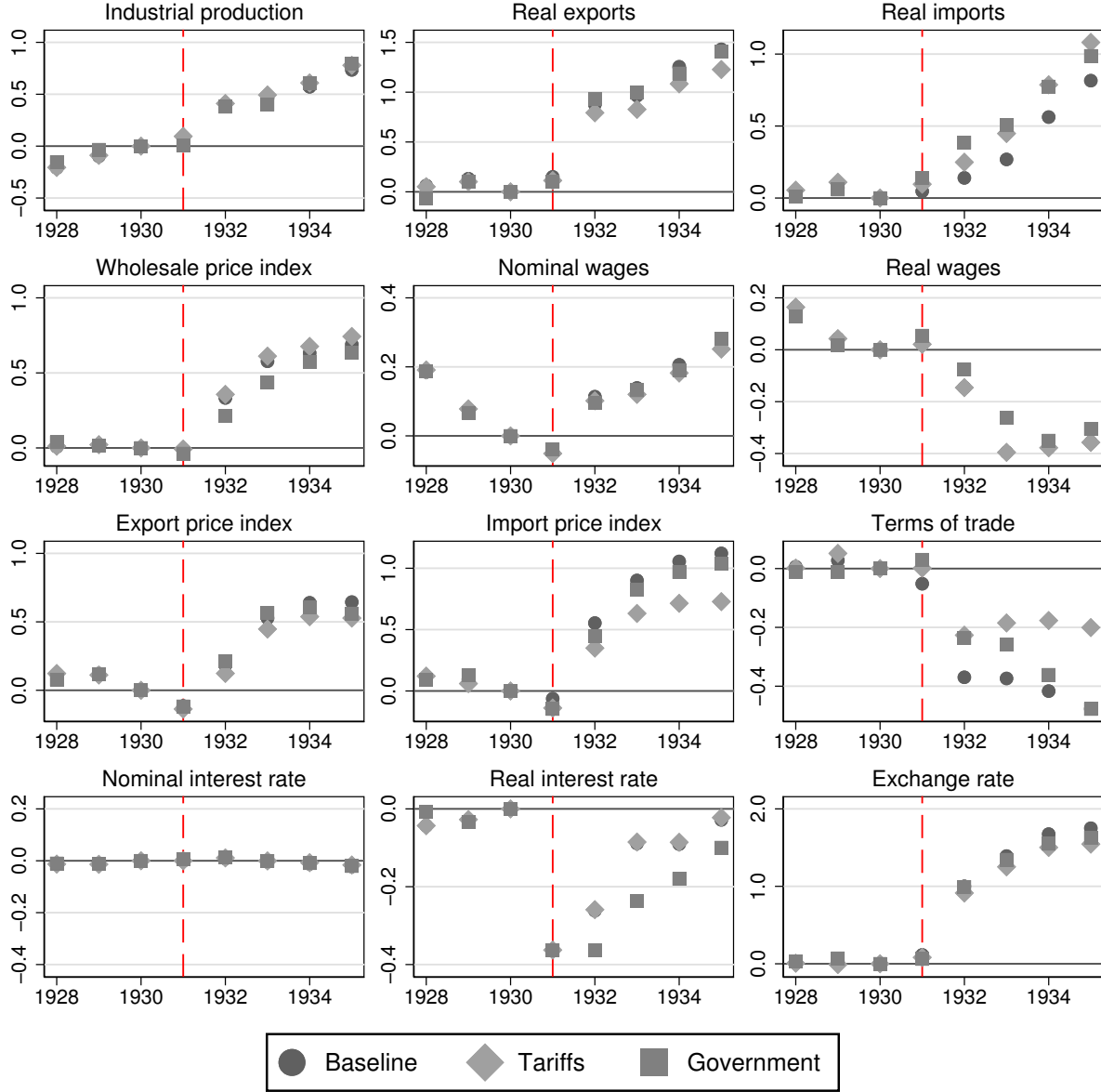
Note: response of relevant variable in devaluing countries, relative to non-devaluing ones. Formally, the plots show the estimate of β_t in equation (1). β_{1930} is normalized to 0. Standard errors are bootstrapped with 2,000 replications, and clustered at the country level. The number in parenthesis is the p-value of a test of the joint significance of the 1932 to 1935 coefficients.

Figure A.2: DD-IV with controls



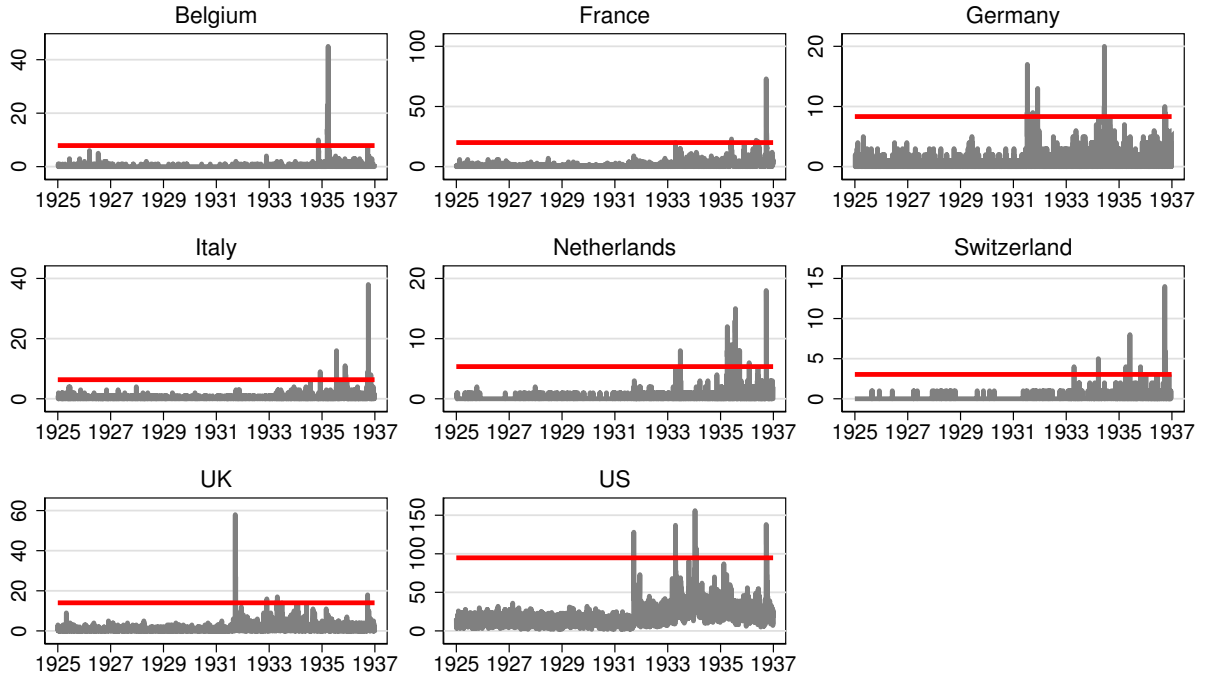
Note: response of relevant variable in devaluing countries, relative to non-devaluing ones. Formally, the plots show the estimate of β_t in equation (1) with no controls (baseline), controlling for the change in tariffs (tariff), or real government revenues and expenditures (government). The change in the exchange rate is instrumented with a devaluation dummy. β_{1930} is normalized to 0. P-values are shown in table A.3.

Figure A.3: DD-OLS with controls



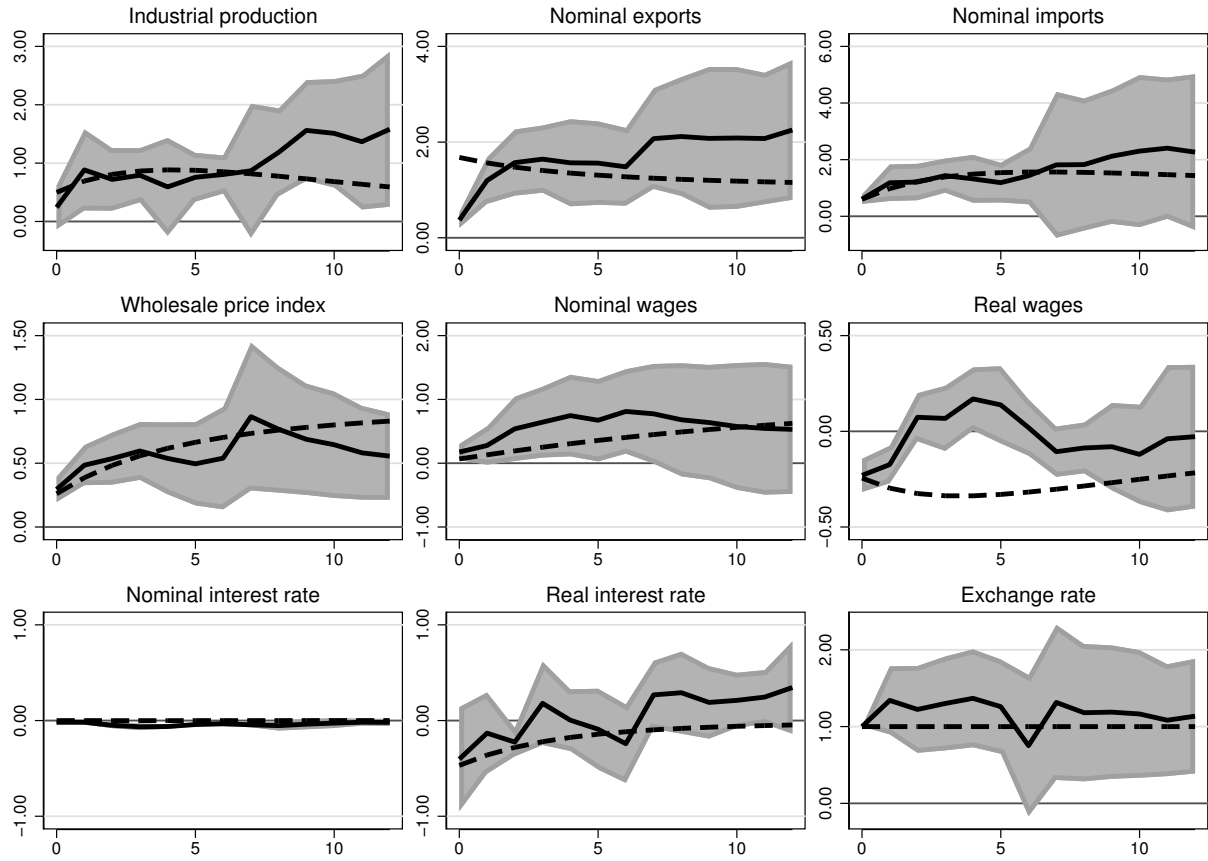
Note: response of relevant variable in devaluing countries, relative to non-devaluing ones. Formally, the plots show the estimate of β_t in equation (1) with no controls (baseline), controlling for the change in tariffs (tariff), or real government revenues and expenditures (government). β_{1930} is normalized to 0. P-values are shown in table A.3.

Figure A.4: Daily number of articles for all countries



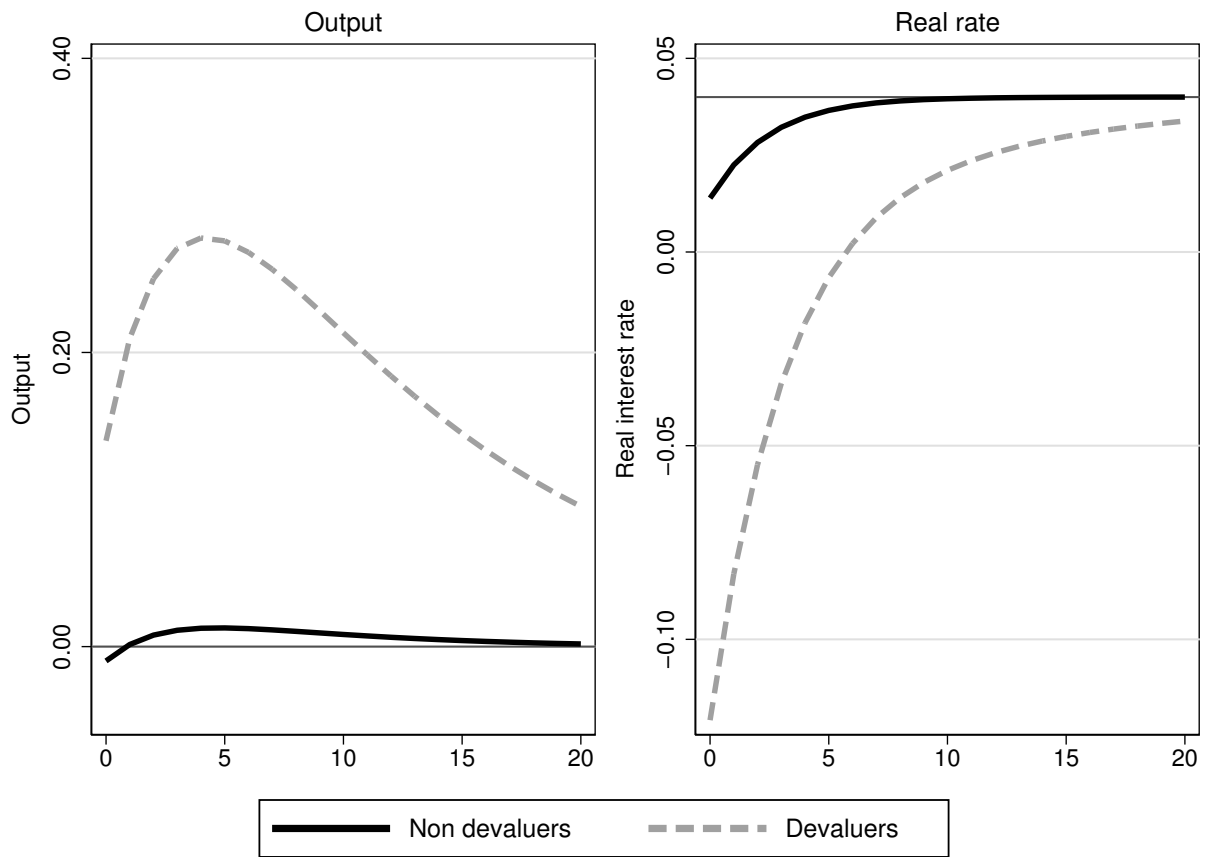
Note: number of daily articles returned by ProQuest Historical Newspapers with the keywords described in section 2.2. The red line is the cutoff defined by the formula: $\text{mean} + 6 \times \text{standard deviations}$.

Figure A.5: Model fit (HFI)



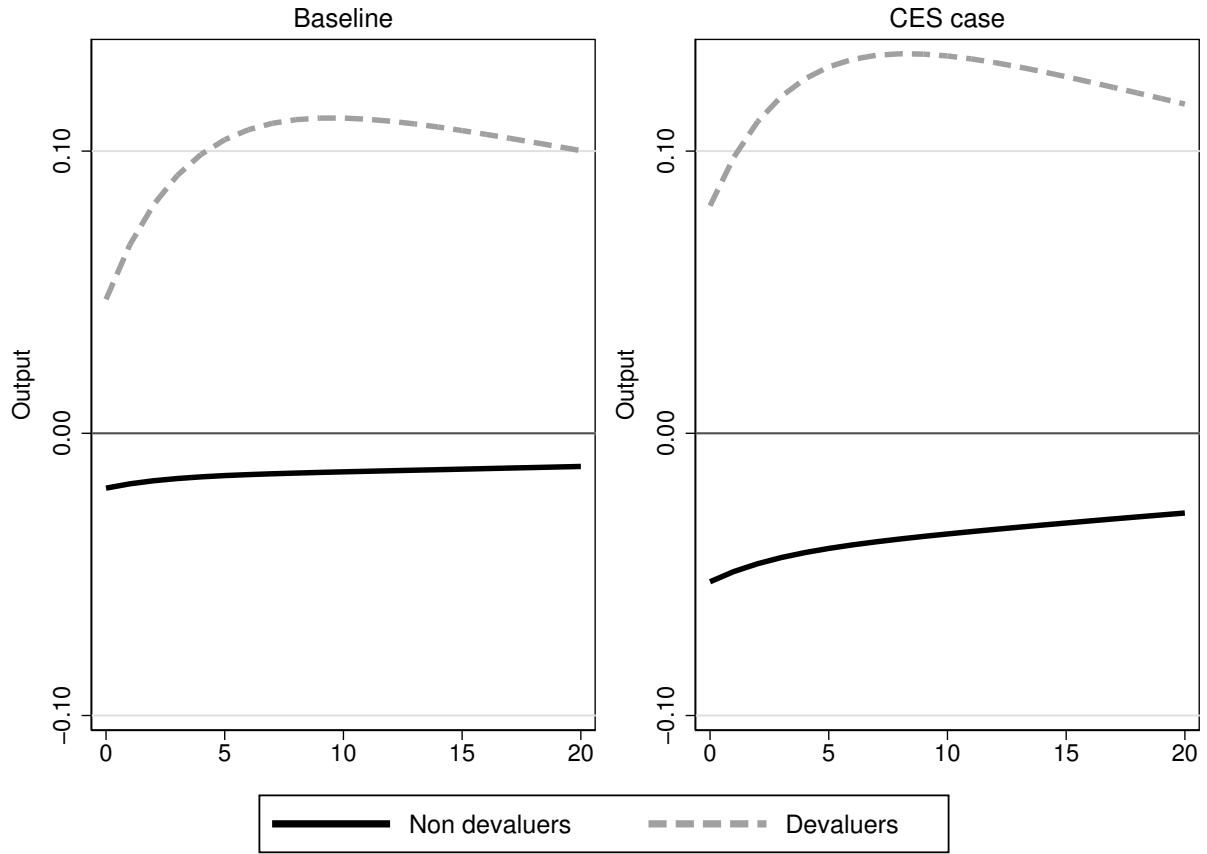
Note: empirical and theoretical responses of relevant variables under estimated parameter values. The black line is the point estimate, the gray area the 95% confidence interval, and the dashed grey line the theoretical response. The response of industrial production, nominal imports, wholesale prices, and nominal wages are explicitly targeted. Those of nominal exports are not. Real wages and the real interest rate are redundant since they respectively are the ratio of nominal wages to the wholesale price index and the difference between the nominal interest rate and wholesale price inflation.

Figure A.6: Counterfactual analysis (HFI)



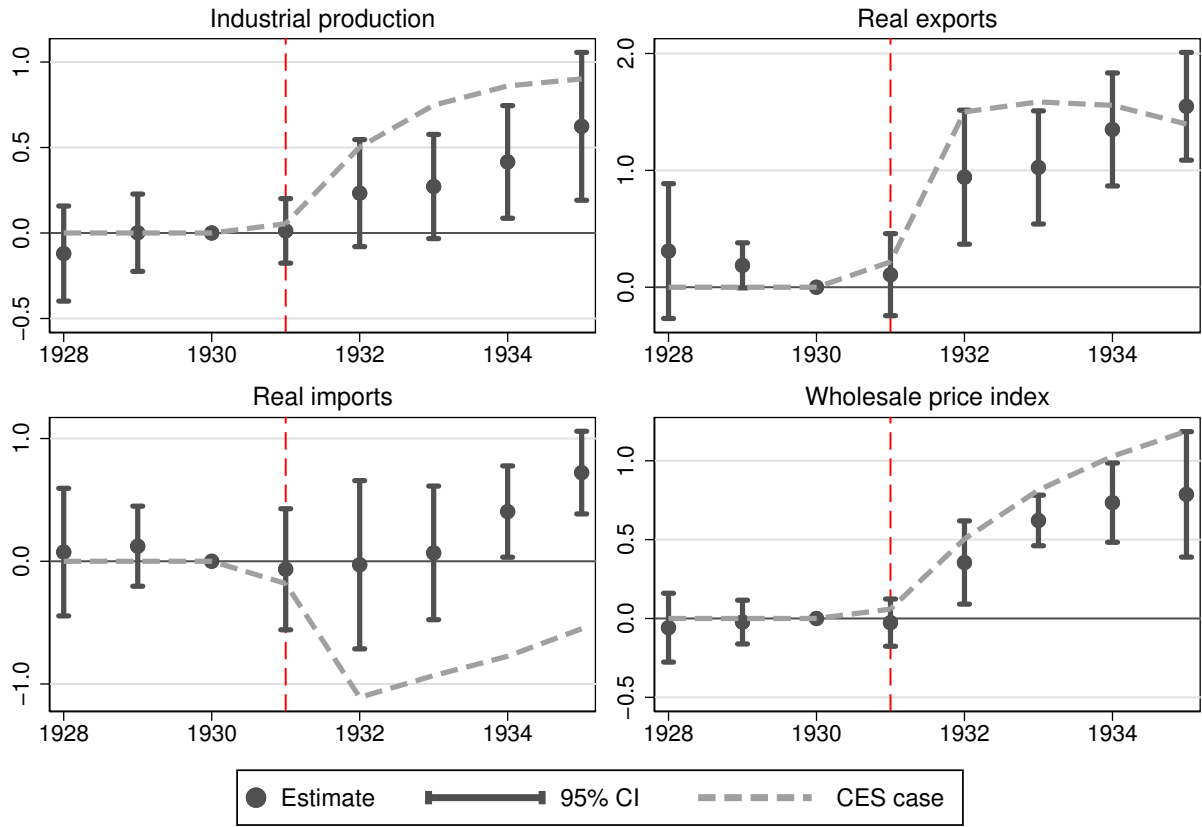
Note: response of output and the real interest rate to a 30% devaluation by half of the world under parameter values estimated with the HFI cross-country moments of section 2.2 and the trade estimates of section 4. Those parameter values are shown in table A.5.

Figure A.7: Counterfactual analysis in a CES world (DD)



Note: response of output to a 30% devaluation in half of the world with estimated parameter values (left-hand side) and with a CES demand for international goods ($\rho = \theta = 3$ and $\bar{m} = 0$). In the latter case, other parameters are set equal to their value in table 6. The left-hand side chart is the same as the left-hand side chart of figure 6.

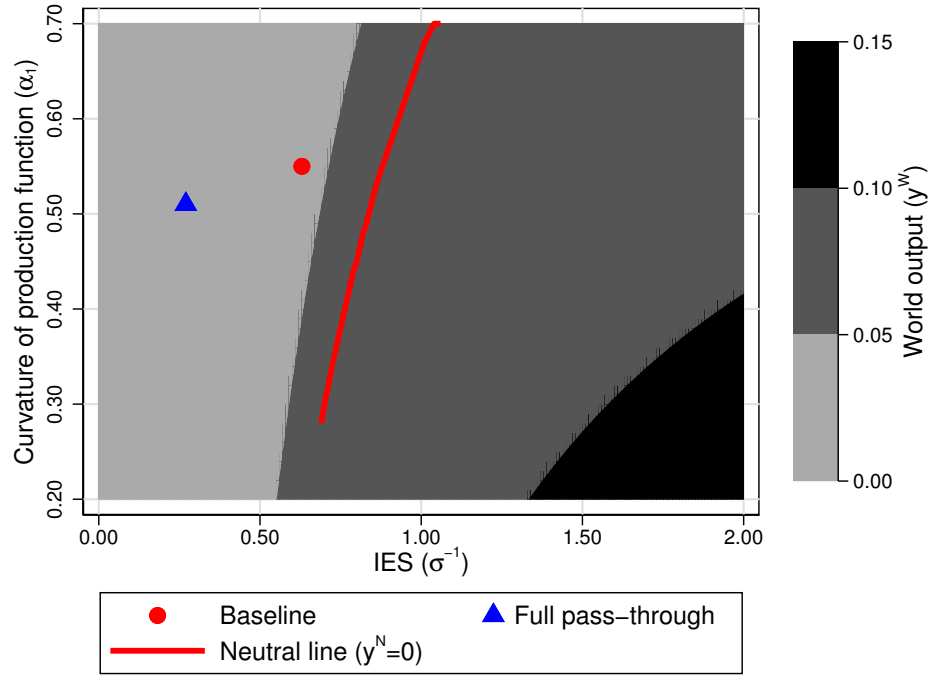
Figure A.8: Poor fit of the CES world (DD)



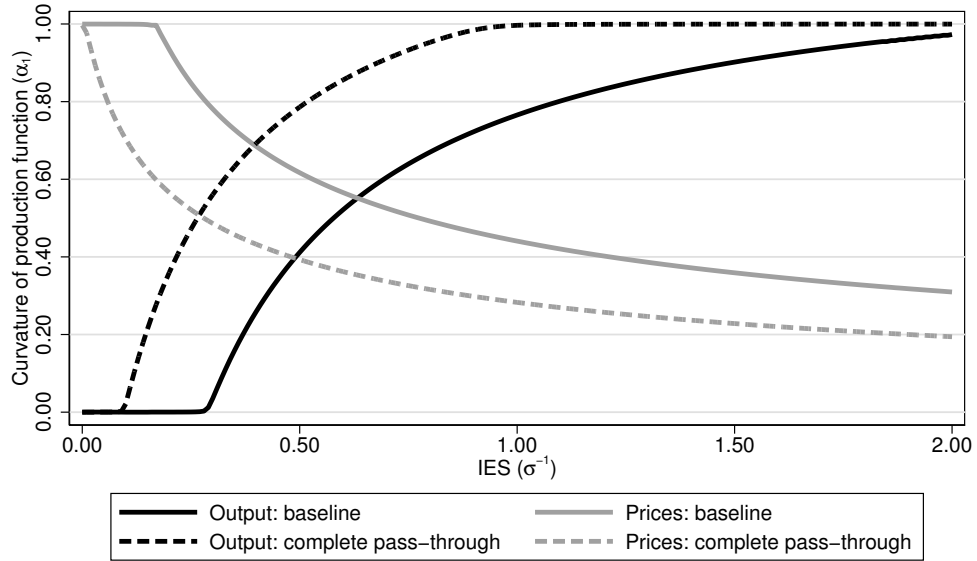
Note: empirical and theoretical responses of relevant variables with a CES demand system for international goods ($\rho = \theta = 3$ and $\bar{m} = 0$)—other parameters are set equal to their value in table 6.

Figure A.9: Identification of α_1

Panel A: Response of World Output as a function of σ^{-1} and α_1

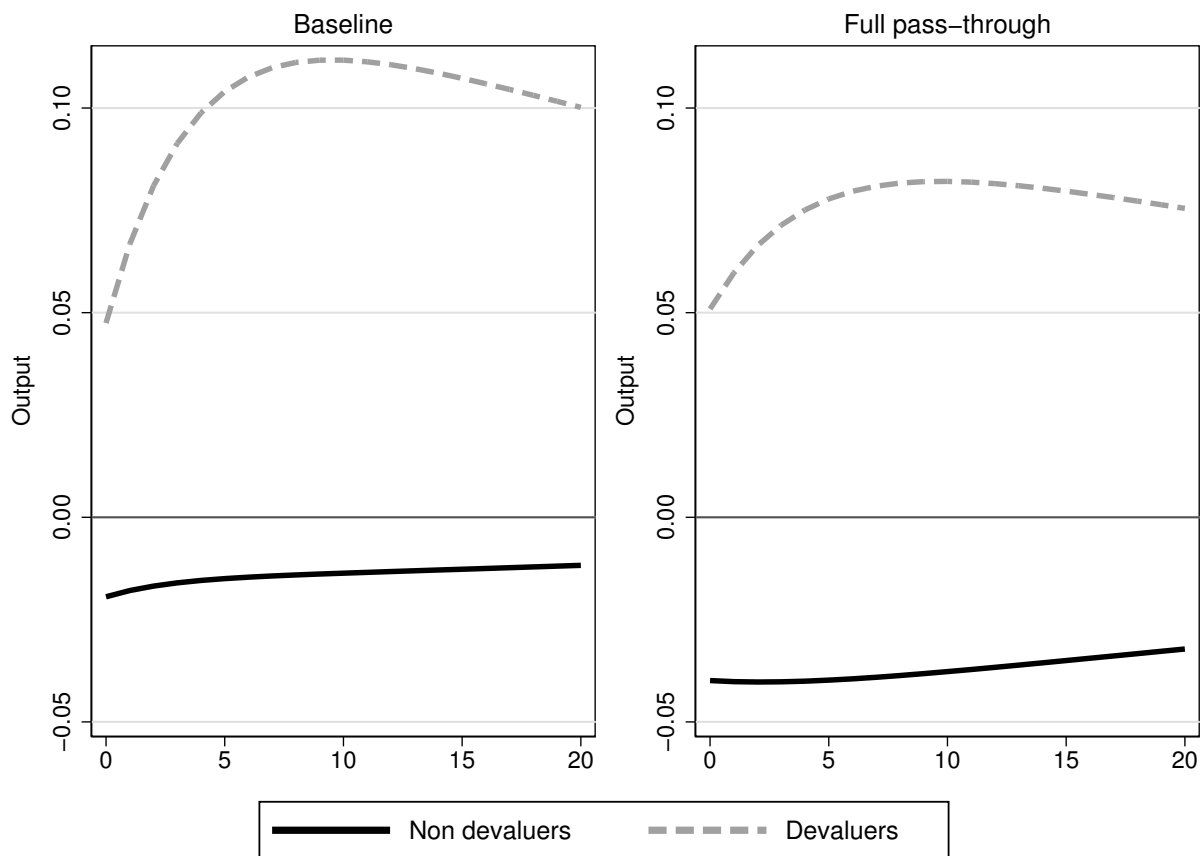


Panel B: Identification of σ^{-1} and ξ



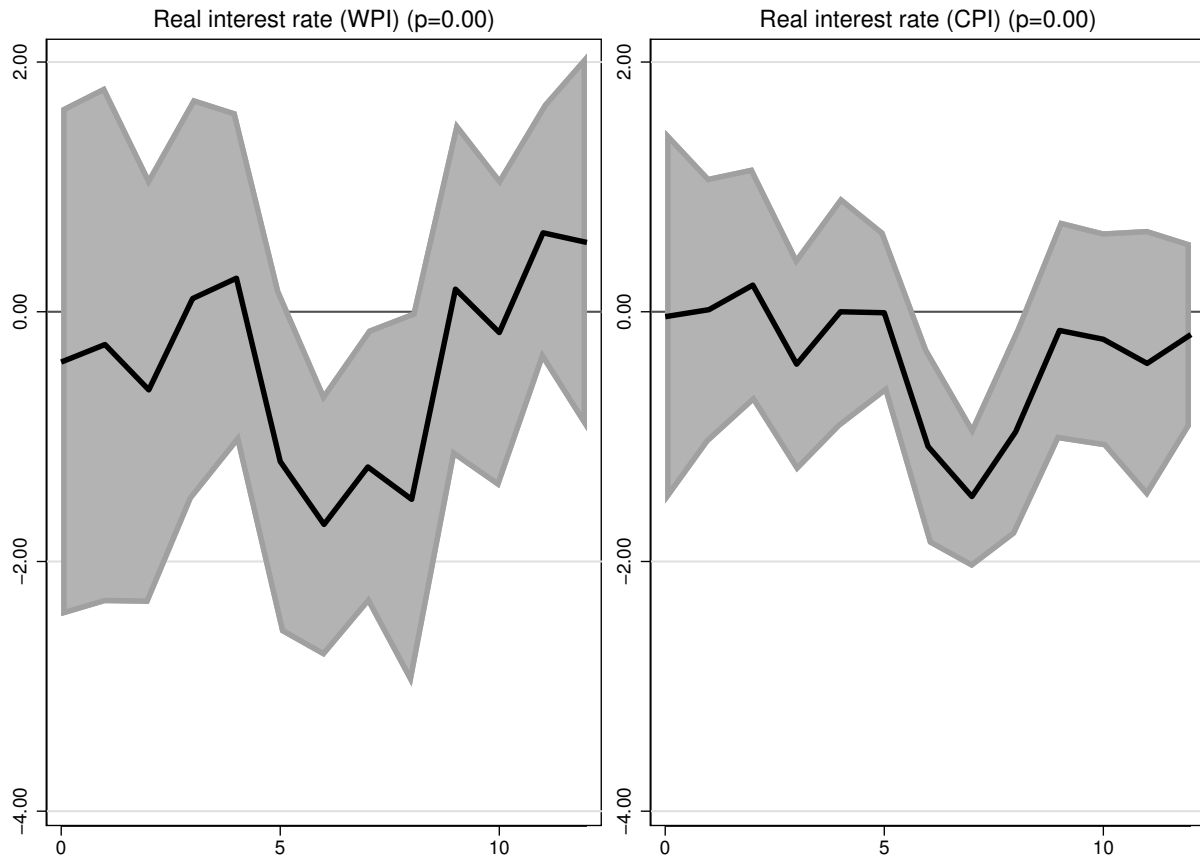
Note: panel A shows the response of world output over three years as a function of the intertemporal elasticity of substitution (σ^{-1}) and the curvature of the production function (α_1). The darker the color, the bigger the effect on world output. The red dot corresponds to the (σ^{-1}, α_1) pair returned by the estimation. The red line draws the (σ^{-1}, α_1) pairs which imply no effect on non-devaluing countries' output. Panel B shows how the estimated moments pin down the two parameters: the black line is the (σ^{-1}, α_1) pairs which imply the observed relative effect on output, the gray line the (σ^{-1}, α_1) pairs which imply the observed relative effect on nominal wages. The intersection of the two lines corresponds to the red dot in panel A.

Figure A.10: Counterfactual analysis with full pass-through and re-estimated σ^{-1} , ξ



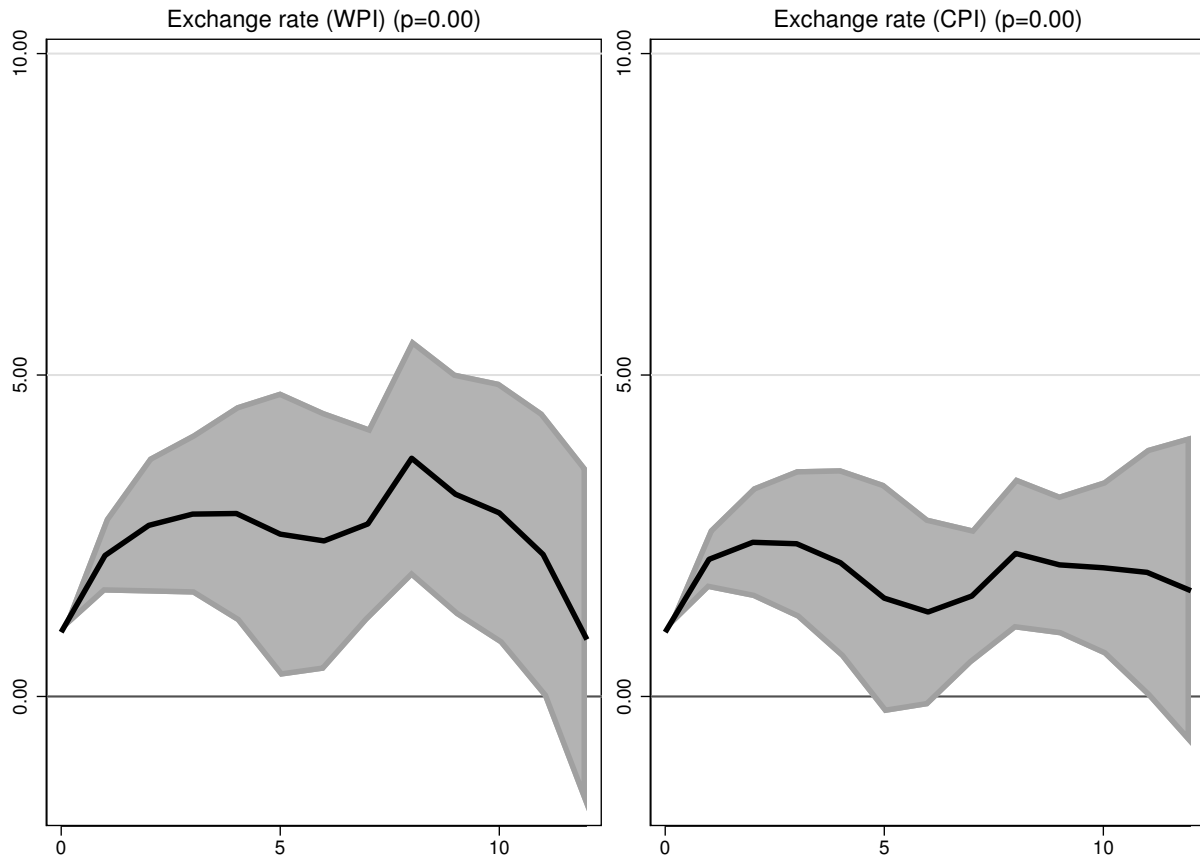
Note: response of output to a 30% devaluation in half of the world with estimated parameter values (left-hand side) and with full pass-through of the exchange rate ($\bar{m} = 0$). In the latter case, σ^{-1} and ξ are re-estimated to match the relative effect on output and wages: their values correspond to the blue triangle in figure 8, panel A; other parameters are set equal to their value in table 6. The left-hand side chart is the same as the left-hand side chart of figure 6.

Figure A.11: Effect on the world real rate (trade-weighted)



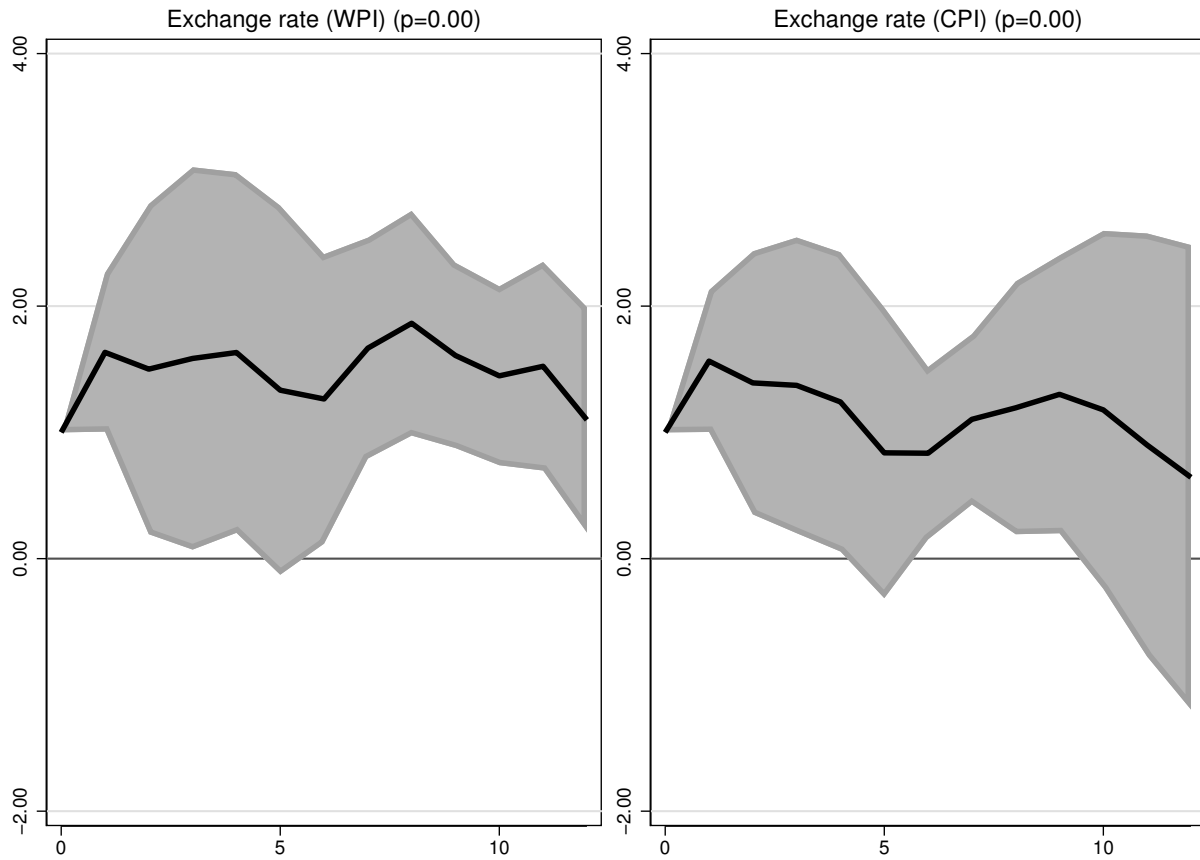
Note: response of relevant variable to an exogenous 100 log-points devaluation in the exchange rate. Formally, the plots show the estimate of β_k in equation (A.54). The black line is the point estimate; the gray area is the 95% confidence interval with Driscoll-Kraay standard errors. The number in parenthesis is the p-value of a test of the joint significance of the coefficients.

Figure A.12: Effect on the world nominal price of gold (GDP-weighted)



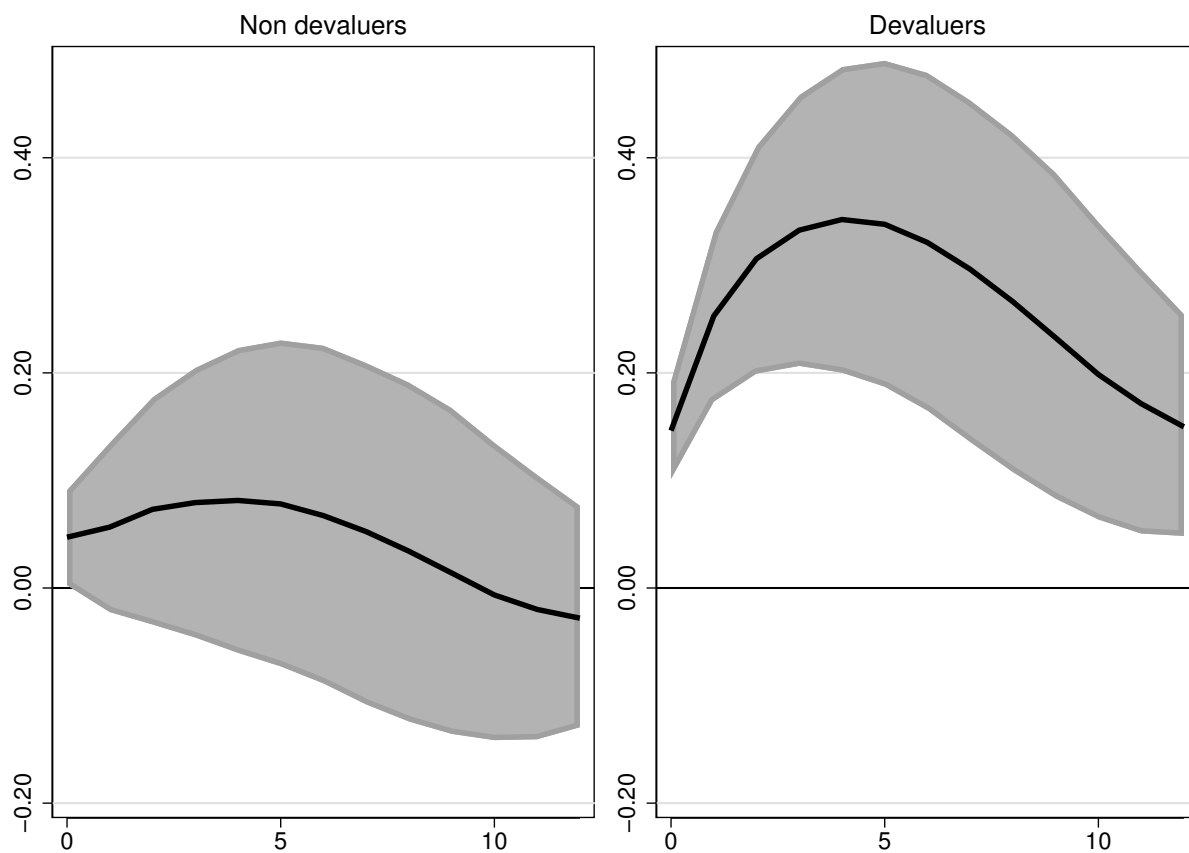
Note: response of relevant variable to an exogenous 100 log-points devaluation in the exchange rate. Formally, the plots show the estimate of β_k in equation (A.54). The black line is the point estimate; the gray area is the 95% confidence interval with Driscoll-Kraay standard errors. The number in parenthesis is the p-value of a test of the joint significance of the coefficients.

Figure A.13: Effect on the world nominal price of gold (trade-weighted)



Note: response of relevant variable to an exogenous 100 log-points devaluation in the exchange rate. Formally, the plots show the estimate of β_k in equation (A.54). The black line is the point estimate; the gray area is the 95% confidence interval with Driscoll-Kraay standard errors. The number in parenthesis is the p-value of a test of the joint significance of the coefficients.

Figure A.14: Effect on output under the sufficient statistics approach



Note: response of output in non-devaluing and devaluing countries constructed following the sufficient statistics approach described in appendix G.3. The black line is the point estimate; the gray area is the 68% confidence interval. The black lines are the same as the two lines shown in the bottom left panel of figure A.16. The confidence interval accounts for uncertainty in the estimate of the effect of devaluation on the path of the world real rate. It is constructed by drawing from a multi-variate normal distribution whose mean and covariance matrix are the point estimates and empirical covariance of these estimates.

C Section 4 Data Construction

I discard imports from countries for which I do not have exchange rate data. After this step, I am left with 89% of the value of US imports from 1930 to 1933. I drop products for which I do not have value and quantity and for which there are two values but only one quantity. When a product is broken down between free and dutiable, I sum quantities and values. When there are two measures of quantity in the same unit, I drop the gross weight. When there are two measures of quantity in different units, I keep the one in pounds. Prices are recovered by dividing value with quantity.

Products are identified by their customs code. In all specifications, a product-country pair is dropped if its value is 0 in one of the relevant years. For instance, when I use 1930 and 1933 data, I only keep product-country pairs for which US imports are strictly positive in both years. This step drops up to 15% of the value of imports remaining from the previous steps. All in all, I am left with 44% of the value of US imports in 1931, 49% in 1932 and 56% in 1933.

The Hawley-Smoot tariff triggered the creation of several product categories in June 1930. As a result, the same product is not always classified consistently before and after the implementation of the tariff. Sometimes, the product category stays the same but switches from free to dutiable. Some other times, a product category is split into other categories. For instance, leather boots and shoes were divided into men's, women's and children's. Whenever possible, I merge all of these product categories to be consistent within and across years.

The customs code were reshuffled between 1929 and 1930, which makes it impossible to use those for the placebo tests. I manually match product categories based on their label. Whenever new product categories were created in 1930, I merge them as described in the previous paragraph.

D Theoretical Derivations

D.1 Demand Functions

The household minimizes expenditures:

$$\sum_{k \in \mathcal{J}} \int_f P_{rt}^{jk}(f) C_{rt}^{jk}(f) df$$

subject to:

$$C_{rt}^j = (1 - \Gamma_{rt}^j) C_{rt}^j \times g \left(\int_f g^* \left(\frac{C_{rt}^{jj}(f)}{(1 - \Gamma_{rt}^j) C_{rt}^j} \right) df \right) + \Gamma_{rt}^j C_{rt}^j \times g \left(\sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \Xi_{rt}^{jk} \int_f g^* \left(\frac{C_{rt}^{jk}(f)}{\Gamma_{rt}^j \Xi_{rt}^{jk} C_{rt}^j} \right) df \right) \quad (\text{A.1})$$

$$g(1) = g^*(1) = g^{*'}(1) = g'(1) = 1 \quad \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \Xi_{rt}^{jk} = 1$$

The measure of firms is normalized to 1.

The first-order conditions are:

$$P_{rt}^{jj}(f) = g^{*'} \left(\frac{C_{rt}^{jj}(f)}{(1 - \Gamma_{rt}^j) C_{rt}^j} \right) \times g' \left(\int_f g^* \left(\frac{C_{rt}^{jj}(f)}{(1 - \Gamma_{rt}^j) C_{rt}^j} \right) df \right) \Lambda_t^j$$

$$P_{rt}^{jk}(f) = g^{*'} \left(\frac{C_{rt}^{jk}(f)}{\Gamma_{rt}^j \Xi_{rt}^{jk} C_{rt}^j} \right) \times g' \left(\sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \Xi_{rt}^{jk} \int_f g^* \left(\frac{C_{rt}^{jk}(f)}{\Gamma_{rt}^j \Xi_{rt}^{jk} C_{rt}^j} \right) df \right) \Lambda_t^j$$

where Λ_t^j is the Lagrange multiplier on the constraint. Rearranging:

$$C_{rt}^{jj}(f) = (1 - \Gamma_{rt}^j) \times h \left(\frac{P_{rt}^{jj}(f)}{\Lambda_t^j \times g' \left(\int_f g^* \left(\frac{C_{rt}^{jj}(f)}{(1 - \Gamma_{rt}^j) C_{rt}^j} \right) df \right)} \right) C_{rt}^j \quad (\text{A.2})$$

$$C_{rt}^{jk}(f) = \Gamma_{rt}^j \Xi_{rt}^{jk} \times h \left(\frac{P_{rt}^{jk}(f)}{\Lambda_t^j g' \left(\sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \Xi_{rt}^{jk} \int_f g^* \left(\frac{C_{rt}^{jk}(f)}{\Gamma_{rt}^j \Xi_{rt}^{jk} C_{rt}^j} \right) df \right)} \right) C_{rt}^j \quad (\text{A.3})$$

where $h(\cdot) = (g^{*'})^{-1}(\cdot)$. Note that $h(1) = 1$ since $g^{*'}(\cdot) = 1$.

Using the first-order conditions, one can check that if relative prices are equal to 1:

$$C_{rt}^{jj}(f) = (1 - \Gamma_{rt}^j) C_{rt}^j$$

$$C_{rt}^{jk}(f) = \Gamma_{rt}^j \Xi_{rt}^{jk} C_{rt}^j$$

Log-linearizing around that symmetric equilibrium when demand shocks are equal to their mean

($\Gamma_{rt}^j = \bar{\Gamma}_{rt}^j$ and $\Xi_{rt}^{jk} = \bar{\Xi}_r^{jk}$):

$$\begin{aligned}\hat{c}_{rt}^j &= (1 - \bar{\Gamma}^j) \hat{c}_{rt}^{jj} + \bar{\Gamma}^j \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}_r^{jk} \hat{c}_{rt}^{jk} \\ \hat{c}_{rt}^{jj} &\equiv \int_f \hat{c}_{rt}^{jj}(f) \quad \hat{c}_{rt}^{jk} \equiv \int_f \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}_r^{jk} \hat{c}_{rt}^{jk}(f) \, df\end{aligned}$$

Similarly defining the price index as:

$$P_{rt}^j C_{rt}^j \equiv \sum_{k \in \mathcal{J}} \int_f P_{rt}^{jk}(f) C_{rt}^{jk}(f) \, df$$

One can show:

$$\begin{aligned}p_{rt}^j &= (1 - \bar{\Gamma}^j) p_{rt}^{jj} + \bar{\Gamma}^j \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}_r^{jk} p_{rt}^{jk} \\ p_{rt}^{jj} &\equiv \int_f p_{rt}^{jj}(f) \quad p_{rt}^{jk} \equiv \int_f \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}_r^{jk} p_{rt}^{jk}(f) \, df\end{aligned}$$

I now log-linearize the demand functions around the symmetric equilibrium:

$$\begin{aligned}\hat{c}_{rt}^{jj}(f) &= \hat{c}_{rt}^j - \theta \left(p_{rt}^{jj}(f) - \lambda_t^j + \tilde{\rho} \left(\hat{c}_{rt}^{jj} - \hat{c}_{rt}^j \right) \right) - \frac{\bar{\Gamma}^j}{1 - \bar{\Gamma}^j} (1 + \theta \tilde{\rho}) \hat{\gamma}_{rt}^j \\ \hat{c}_{rt}^{jk}(f) &= \hat{c}_{rt}^j - \theta \left(p_{rt}^{jk}(f) - \lambda_t^j + \tilde{\rho} \left(\hat{c}_{rt}^{jk} - \hat{c}_{rt}^j \right) \right) + (1 + \theta \tilde{\rho}) \hat{\gamma}_{rt}^j + \hat{\xi}_{rt}^{jk}\end{aligned}$$

where:

$$\begin{aligned}\hat{c}_{rt}^{jj} &\equiv \int_f \hat{c}_{rt}^{jj}(f) \, df & \hat{c}_{rt}^{j*} &\equiv \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}_r^{jk} \int_f \hat{c}_{rt}^{jk}(f) \, df \\ \theta &\equiv -\frac{h'(1)}{h(1)} & \tilde{\rho} &= -\frac{g^{*'}(1)g''(1)}{g'(1)}\end{aligned}$$

I then aggregate over f :

$$\hat{c}_{rt}^{jj} = \hat{c}_{rt}^j - \theta \left(p_{rt}^{jj} - \lambda_t^j + \tilde{\rho} \left(\hat{c}_{rt}^{jj} - \hat{c}_{rt}^j \right) \right) - \frac{\bar{\Gamma}^j}{1 - \bar{\Gamma}^j} (1 + \theta \tilde{\rho}) \hat{\gamma}_{rt}^j \quad (\text{A.4})$$

$$\hat{c}_{rt}^{jk} = \hat{c}_{rt}^j - \theta \left(p_{rt}^{jk} - \lambda_t^j + \tilde{\rho} \left(\hat{c}_{rt}^{j*} - \hat{c}_{rt}^j \right) \right) + (1 + \theta \tilde{\rho}) \hat{\gamma}_{rt}^j + \hat{\xi}_{rt}^{jk} \quad (\text{A.5})$$

The demand for imports is given by averaging over k in equation (A.5):

$$\hat{c}_{rt}^{j*} = \hat{c}_{rt}^j - \theta \left(p_{rt}^{j*} - \lambda_t^j + \tilde{\rho} \left(\hat{c}_{rt}^{j*} - \hat{c}_{rt}^j \right) \right) + (1 + \theta \tilde{\rho}) \hat{\gamma}_{rt}^j \quad (\text{A.6})$$

Subtracting equation (A.6) from equation (A.5):

$$\hat{c}_{rt}^{jk} = \hat{c}_{rt}^{j*} - \theta \left(p_{rt}^{jk} - p_{rt}^{j*} \right) + \hat{\xi}_{rt}^{jk} \quad (\text{A.7})$$

Taking the weighted average of equations (A.4) and (A.6), I can check that: $\lambda_{rt}^j = p_{rt}^j$. Then, equations (A.4) and (A.6) can be rearranged to:

$$\hat{c}_{rt}^{jj} = \hat{c}_{rt}^j - \rho \left(p_{rt}^{jj} - p_{rt}^j \right) - \frac{\bar{\Gamma}^j}{1 - \bar{\Gamma}^j} \hat{\gamma}_{rt}^j \quad (\text{A.8})$$

$$\hat{c}_{rt}^{j*} = \hat{c}_{rt}^j - \rho \left(p_{rt}^{j*} - p_{rt}^j \right) + \hat{\gamma}_{rt}^j \quad (\text{A.9})$$

where:

$$\rho \equiv \frac{\theta}{1 + \theta \tilde{\rho}}$$

ρ is the elasticity of substitution between domestic and foreign varieties, while θ is the elasticity of substitution among foreign varieties.

CES case: Suppose that:

$$g(x) = x \quad g^*(x) = 1 + \frac{\theta}{\theta - 1} \left(x^{\frac{\theta-1}{\theta}} - 1 \right)$$

The Kimball aggregator simplifies to:

$$C_{rt}^j = (1 - \Gamma_{rt}^j) C_{rt}^j \times \int_f \left(\frac{C_{rt}^{jj}(f)}{(1 - \Gamma_{rt}^j) C_{rt}^j} \right)^{\frac{\theta-1}{\theta}} df + \Gamma_{rt}^j C_{rt}^j \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \Xi_{rt}^{jk} \int_f \left(\frac{C_{rt}^{jk}(f)}{\Gamma_{rt}^j \Xi_{rt}^{jk} C_{rt}^j} \right)^{\frac{\theta-1}{\theta}} df$$

Which can be rewritten in a more traditional form:

$$C_{rt}^j = \left[(1 - \Gamma_{rt}^j)^{\frac{1}{\theta}} \int_f \left(C_{rt}^{jj}(f) \right)^{\frac{\theta-1}{\theta}} df + (\Gamma_{rt}^j)^{\frac{1}{\theta}} \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} (\Xi_{rt}^{jk})^{\frac{1}{\theta}} \int_f \left(C_{rt}^{jk}(f) \right)^{\frac{\theta-1}{\theta}} df \right]^{\frac{\theta}{\theta-1}}$$

D.2 Strategic Complementarities

D.2.1 Price indices

A firm maximizes:

$$\left(P_{rt}^{jk}(f) - \frac{MC_t^k(f)}{XR_t^{kj}} \right) C_t^{jk}(f)$$

subject to equations (A.2) or (A.3). The first-order condition for $k \neq j$ is:

$$h(X) + \left(1 - \frac{MC_{rt}^k(f)/XR_t^{kj}}{P_{rt}^{jk}} \right) X h'(X) = 0$$

where:

$$X = \frac{P_{rt}^{jk}(f)}{\Lambda_t^j g' \left(\sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \Xi_{rt}^{jk} \int_f g^* \left(\frac{C_{rt}^{jk}(f)}{\Gamma_{rt}^j \Xi_{rt}^{jk} C_{rt}^j} \right) df \right)}$$

Rearranging:

$$P_{rt}^{jk}(f) = \mathcal{M}(X) \frac{MC_{rt}^k(f)}{XR_t^{kj}}$$

where:

$$\mathcal{M}(X) = \frac{-\frac{Xh'(X)}{h(X)}}{-\frac{Xh'(X)}{h(X)} - 1}$$

Log-linearizing:

$$p_t^{jk}(f) = (-\bar{m} \left(p_t^{jk}(f) - p_t^j + \tilde{\rho}(\hat{c}_{rt}^{j*} - \hat{c}_t^j - \hat{\gamma}_t^j) \right) + (\Theta + mc_t^k(f) - xr_t^{kj}))$$

$$\bar{m} = \frac{d \log \mathcal{M}(X)}{d \log X} \Big|_{X=1} \quad \Theta = \log \frac{\theta}{\theta - 1}$$

Aggregating over f :

$$p_{rt}^{jk} = -\bar{m} \left(p_{rt}^{jk} - p_{rt}^j + \tilde{\rho}(\hat{c}_{rt}^{j*} - \hat{c}_{rt}^j - \hat{\gamma}_{rt}^j) \right) + (\Theta + mc_{rt}^k - xr_t^{kj})$$

Substituting equation (A.9) and rearranging:

$$p_{rt}^{jk} = \frac{\bar{m}\tilde{\rho}}{1 + \bar{m}} p_{rt}^{j*} + \frac{\bar{m}(1 - \tilde{\rho})}{1 + \bar{m}} p_{rt}^j + \frac{1}{1 + \bar{m}} (\Theta + mc_{rt}^k - xr_t^{kj})$$

Aggregating over k :

$$p_{rt}^{j*} = \frac{\bar{m}(1 - \tilde{\rho})}{1 + \bar{m}(1 - \tilde{\rho})} p_{rt}^j + \frac{1}{1 + \bar{m}(1 - \tilde{\rho})} \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}^{jk} (\Theta + mc_{rt}^k - xr_t^{kj})$$

Similarly, for the domestic good:

$$p_{rt}^{jj}(f) = -\bar{m} \left(p_{rt}^{jj}(f) - p_{rt}^j + \tilde{\rho} \left(\hat{c}_{rt}^{jj} - \hat{c}_{rt}^j + \frac{\bar{\Gamma}^j}{1 - \bar{\Gamma}^j} \hat{\gamma}_{rt}^j \right) \right) + \Theta + mc_{rt}^j(f)$$

Aggregating over f :

$$p_{rt}^{jj} = \frac{\bar{m}(1 - \tilde{\rho})}{1 + \bar{m}(1 - \tilde{\rho})} p_{rt}^j + \frac{1}{1 + \bar{m}(1 - \tilde{\rho})} (\Theta + mc_{rt}^j)$$

Therefore:

$$p_{rt}^{jk} = \zeta_1 p_{rt}^{j*} + \zeta_2 p_{rt}^j + (1 - \zeta_1 - \zeta_2)(\Theta + mc_{rt}^k - xr_t^{kj}) \quad (\text{A.10})$$

$$p_{rt}^{j*} = \delta p_{rt}^j + (1 - \delta) \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}^{jk} (\Theta + mc_{rt}^k - xr_t^{kj}) \quad (\text{A.11})$$

$$p_{rt}^{jj} = \delta p_{rt}^j + (1 - \delta) (\Theta + mc_{rt}^j) \quad (\text{A.12})$$

where:

$$\zeta_1 \equiv \frac{\bar{m}\tilde{\rho}\rho}{1 + \bar{m}} \quad \zeta_2 \equiv \frac{\bar{m}(1 - \tilde{\rho}\rho)}{1 + \bar{m}} \quad \delta \equiv \frac{\bar{m}(1 - \tilde{\rho}\rho)}{1 + \bar{m}(1 - \tilde{\rho}\rho)}$$

Substituting in the formula for $\tilde{\rho}$:

$$\zeta_1 = \frac{\bar{m}(1 - \rho/\theta)}{1 + \bar{m}} \quad \zeta_2 = \frac{\bar{m}\rho/\theta}{1 + \bar{m}} \quad \delta = \frac{\bar{m}\rho}{\theta + \bar{m}\rho}$$

Combining equations (A.11) and (A.12):

$$p_{rt}^j = \Theta + (1 - \bar{\Gamma}^j) mc_{rt}^j + \bar{\Gamma}^j \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}^{jk} (mc_{rt}^k - xr_t^{kj}) \quad (\text{A.13})$$

Invoking equation (A.11) one more time:

$$p_{rt}^{j*} = \Theta + \delta(1 - \bar{\Gamma}^j) mc_{rt}^j + (1 - \delta(1 - \bar{\Gamma}^j)) \sum_{\substack{k \in \mathcal{J} \\ k \neq j}} \bar{\Xi}^{jk} (mc_{rt}^k - xr_t^{kj}) \quad (\text{A.14})$$

Plugging the latter two results into equation (7):

$$\begin{aligned} p_{rt}^{jk} = & \Theta + \delta(1 - \bar{\Gamma}^j) mc_{rt}^j + (\zeta_1 + \zeta_2 - \delta(1 - \bar{\Gamma}^j)) \sum_{\substack{m \in \mathcal{J} \\ m \neq j}} \bar{\Xi}^{jm} (mc_{rt}^m - xr_t^{mj}) \\ & + (1 - \zeta_1 - \zeta_2) (mc_{rt}^k - xr_t^{kj}) \end{aligned} \quad (\text{A.15})$$

Klenow and Willis (2016) case: Suppose that: $h(x) = (1 - \bar{m}' \log(x))^{\frac{\theta}{\bar{m}'}}$. Then: $\bar{m} = \bar{m}'/(\theta - 1)$.¹⁵ \bar{m} can take any value from 0 to infinity as long as $\theta > 1$. So $(1 - \zeta_1 - \zeta_2)$ can take any value between 0 and 1.

¹⁵See Itskhoki and Mukhin (2019) for a detailed proof.

D.2.2 Proof of Proposition 2

The producer price index is given by combining equations (A.12) and (A.13):

$$p_t^{kk} = \Theta + (1 - \delta\bar{\Gamma})mc_t^k + \delta\bar{\Gamma} \int_m (mc_t^m - xr_t^{km}) dm$$

Rearranging:

$$\begin{aligned} mc_t^k &= -\frac{\Theta}{1 - \delta\bar{\Gamma}} + \frac{1}{1 - \delta\bar{\Gamma}}p_t^k - \frac{\delta\bar{\Gamma}}{1 - \delta\bar{\Gamma}} \int_m (mc_t^m - xr_t^{km}) dm \\ &= -\frac{\Theta}{1 - \delta\bar{\Gamma}} + \frac{1}{1 - \delta\bar{\Gamma}}p_t^k - \frac{\delta\bar{\Gamma}}{1 - \delta\bar{\Gamma}} \int_m (mc_t^m - xr_t^{mj}) dm - \frac{\delta\bar{\Gamma}}{1 - \delta\bar{\Gamma}}xr_t^{kj} \end{aligned}$$

Plugging into the simplified version of equation (A.10):

$$p_t^{jk} = -\frac{1 - \zeta_1 - \zeta_2}{1 - \delta\bar{\Gamma}}xr_t^{kj} + \frac{1 - \zeta_1 - \zeta_2}{1 - \delta\bar{\Gamma}} \left(p_t^k - \delta\bar{\Gamma} \int_m (mc_t^m - xr_t^{mj}) dm \right) + \zeta_1 p_t^{j*} + \zeta_2 p_t^j$$

D.3 Market Clearing

To a first order, demand for the varieties of country j are:

$$\hat{y}_t^j = (1 - \bar{\Gamma})\hat{c}_t^{jj} + \bar{\Gamma} \int_k \hat{c}_t^{kj} dk$$

Using equations (A.7) and (A.9):

$$\hat{c}_t^{kj} = \hat{c}_t^{k*} - \theta \left(p_t^{kj} - p_t^{k*} \right) = \hat{c}_t^k - \rho \left(p_t^{k*} - p_t^k \right) - \theta \left(p_t^{kj} - p_t^{k*} \right)$$

Now, using equations (A.13) and (A.14):

$$p_t^{k*} - p_t^k = (1 - \delta(1 - \bar{\Gamma})) \left(\int_m (mc_t^m - xr_t^{km}) dm - mc_t^k \right)$$

Moreover, using equations (A.14) and (A.15):

$$p_t^{kj} - p_t^{k*} = (1 - \zeta_1 - \zeta_2) \left(mc_t^j - xr_t^{kj} - \int_m (mc_t^m - xr_t^{km}) dm \right)$$

So:

$$\begin{aligned}\hat{c}_t^{kj} &= \hat{c}_t^k - \rho(1 - \delta(1 - \bar{\Gamma})) \left(\int_m (mc_t^m - xr_t^{km}) dm - mc_t^k \right) \\ &\quad - \theta(1 - \zeta_1 - \zeta_2) \left(mc_t^j - xr_t^{kj} - \int_m (mc_t^m - xr_t^{km}) dm \right)\end{aligned}$$

Note that: $xr^{km} = xr^{mj} - xr^{kj}$. Using this identity and integrating over k :

$$\begin{aligned}\int_k \hat{c}_t^{kj} dk &= \int_k \hat{c}_t^k dk \\ &\quad - \theta(1 - \zeta_1 - \zeta_2) \left(mc_t^j - \int_m (mc_t^m - xr_t^{mj}) dm \right)\end{aligned}$$

Moreover, using equations (A.8) and (A.13):

$$\hat{c}_t^{jj} = \hat{c}_t^j - \rho(p_t^{jj} - p_t^j) = \hat{c}_t^j - \rho(1 - \delta)\bar{\Gamma} \left(mc_t^j - \int_m (mc_t^m - xr_t^{mj}) dm \right)$$

Gathering those results:

$$\begin{aligned}\hat{y}_t^j &= \psi \left(\int_m (mc_t^m - xr_t^{mj}) dm - mc_t^j \right) + (1 - \bar{\Gamma})\hat{c}_t^j + \bar{\Gamma} \int_k \hat{c}_t^k dk \\ \psi &\equiv \bar{\Gamma} [(1 - \delta)(1 - \bar{\Gamma})\rho + (1 - \zeta_1 - \zeta_2)\theta]\end{aligned}\tag{A.16}$$

D.4 Households

The utility of the country's representative household is:

$$E_0 \sum_{t=0}^{\infty} \left(\frac{(C_t^j - \iota C_{t-1}^j)^{1-\sigma} - 1}{1-\sigma} - \int_l \frac{(N_t^j(l))^{1+\phi}}{1+\phi} dl + \frac{(M_t^j/P_t^j)^{1-\chi} - 1}{1-\chi} \right)$$

subject to:

$$\begin{aligned}\frac{B_t^j}{1+i_t^j} - B_{t-1}^j + M_t^j - M_{t-1}^j &= \int_l W_t^j(l) N_t^j(l) dl - P_t^j C_t^j + \text{profits} \\ N_t^j(l) &= \left(\frac{W_t^j(l)}{W_t^j} \right)^{-\eta} N_t^j\end{aligned}$$

The Euler equation is:

$$\begin{aligned}\hat{\mu}_t^j &= E_t \hat{\mu}_{t+1}^j + \hat{i}_t^j - E_t \pi_{ct+1}^j \\ \hat{\mu}_t^j &\equiv \frac{\sigma}{(1-\iota)(1-\beta\iota)} \left(\beta\iota \hat{c}_{t+1}^j - (1+\beta\iota^2) \hat{c}_t^j + \iota \hat{c}_{t-1}^j \right)\end{aligned}$$

The first-order condition for money holdings is:

$$\hat{i}_t^j = \frac{1-\beta}{\beta^2} \left(-\chi(\widehat{m_t^j - p_t^j}) - \hat{\mu}_t^j \right)$$

The Phillips curve can be written as:¹⁶

$$\begin{aligned}\pi_{wt}^j &= \beta E_t \pi_{wt+1}^j + \tilde{\kappa} \hat{y}_t^j - \kappa \hat{\mu}_t^j + \kappa(\widehat{w_t^j - p_t^j}) \\ \kappa &\equiv \frac{(1-\xi)(1-\beta\xi)}{\xi(1+\eta\phi)} \quad \tilde{\kappa} \equiv \kappa \frac{\phi}{1-\alpha_2}\end{aligned}$$

D.5 International Payments

Profits of the central bank are:

$$\Pi_{cbt}^j = M_t^j - M_{t-1}^j - \mathcal{E}_t^j \left(G_t^j - G_{t-1}^j \right)$$

Profits of the firms are:

$$\left(P_t^{jj} - W_t^j \right) C_t^{jj}(f) + \int_k \left(\frac{P_t^{kj}}{X R_t^{kj}} - W_t^j \right) C_t^{kj}(f) dk$$

The balance of payments of the country is given by:

$$\frac{B_t^j}{1+i_t^j} - B_{t-1}^j + \mathcal{E}_t^j \left(G_t^j - G_{t-1}^j \right) = \int_f P_t^{jj} C_t^{jj}(f) df + \int_k \int_f \frac{P_t^{kj}}{X R_t^{kj}} C_t^{kj}(f) dk df - P_t^j C_t^j$$

To a first-order, this is:

$$\beta b_t^j - b_{t-1}^j + \bar{g} \Delta g_t^j = \bar{\Gamma} \int_k \left(p_t^{kj} - x r_t^{kj} - p_t^{jk} \right) dk + \bar{\Gamma} \int_k \left(\hat{c}_t^{kj} dk - \hat{c}_t^{kj} \right) dk$$

¹⁶See Galí (2015, ch. 6) for a derivation.

$$b_t^j = \frac{B_t^j}{P_t^j Y_t^j}$$

$$\bar{g} = \frac{\mathcal{E}_t^j G_t^j}{P_t^j Y_t^j} \text{ expressed in steady state}$$

D.6 Model Summary

For each country, there are 17 variables: output (y), consumption (c), marginal cost (mc), wage (w), price charged in country k (p^{jk}),¹⁷ import price index (IPI, p^{j*}), producer price index (PPI, p^{jj}), consumer price index (CPI, p^j), CPI inflation (π_c^j), wage inflation (π_w^j), marginal utility (μ^j), nominal interest rate (i^j), money supply (m^j), exchange rate (xr^j), nominal gold price (ϵ^j), bonds owned (b^j) and gold stock (g^j). The endogenous state variables are p_{t-1}^j , w_{t-1}^j , b_{t-1}^j and g_{t-1}^j . The exogenous state variable are ϵ_t^j . Other variables are controls.

There are 18 equations:

$$\text{market clearing: } \hat{y}_t^j = \psi \left(\int_m (mc_t^m - xr_t^{mj}) dm - mc_t^j \right) + (1 - \bar{\Gamma}) \hat{c}_t^j + \bar{\Gamma} \int_k \hat{c}_t^k dk \quad (\text{A.17})$$

$$\text{marginal cost: } mc_t^j = w_t^j + \frac{\alpha_1}{1 - \alpha_2} \hat{y}_t^j \quad (\text{A.18})$$

$$\text{CPI: } p_t^j = \Theta + (1 - \bar{\Gamma}) mc_t^j + \bar{\Gamma} \int_m (mc_t^m - xr_t^{mj}) \quad (\text{A.19})$$

$$\text{CPI inflation: } \pi_{ct}^j = p_t^j - p_{t-1}^j \quad (\text{A.20})$$

$$\text{wage inflation: } \pi_{wt}^j = w_t^j - w_{t-1}^j \quad (\text{A.21})$$

$$\text{marginal utility: } \hat{\mu}_t^j = \frac{\sigma}{(1 - \iota)(1 - \beta \iota)} \left(\beta \iota E_t \hat{c}_{t+1}^j - (1 + \beta \iota^2) \hat{c}_t^j + \iota \hat{c}_{t-1}^j \right) \quad (\text{A.22})$$

$$\text{Euler equation: } \hat{\mu}_t^j = E_t \hat{\mu}_{t+1}^j + \hat{v}_t^j - E_t \pi_{c,t+1}^j \quad (\text{A.23})$$

$$\text{money holdings: } \hat{v}_t^j = \frac{1 - \beta}{\beta^2} \left(-\chi (\widehat{m_t^j - p_t^j}) - \hat{\mu}_t^j \right) \quad (\text{A.24})$$

$$\text{Phillips curve: } \pi_{wt}^j = \beta E_t \pi_{wt+1}^j + \tilde{\kappa} \hat{y}_t^j - \kappa \hat{\mu}_t^j + \kappa (\widehat{w_t^j - p_t^j}) \quad (\text{A.25})$$

$$\text{money supply: } \hat{m}_t^j - \hat{\epsilon}_t^j = \hat{g}_t^j \quad (\text{A.26})$$

$$\text{gold peg: } E_t \epsilon_{t+1}^j = \epsilon_t^j \quad (\text{A.27})$$

$$\text{exchange rate: } xr_t^{jk} = \epsilon_t^j - \epsilon_t^k \quad (\text{A.28})$$

$$\text{quasi UIP: } i_t^j - i_t^k - E_t (\Delta \epsilon_t^j - \epsilon_t^k) = v(b_t^k - b_t^j) \quad (\text{A.29})$$

¹⁷To be precise, p^{jk} is not a single variable, but a continuum of variables as there is a continuum of countries k .

$$\text{price setting: } p_t^{jk} = (1 - \zeta_1 - \zeta_2)\Theta + \zeta_1 p_t^{j*} + \zeta_2 p_t^j + (1 - \zeta_1 - \zeta_2)(m c_t^k - x r_t^{kj}) \quad (\text{A.30})$$

$$\text{IPI: } p_t^{j*} = (1 - \delta)\Theta + \delta p_t^j + (1 - \delta) \int_k (m c_t^k - x r_t^{kj}) \quad (\text{A.31})$$

$$\text{PPI: } p_t^{jj} = (1 - \delta)\Theta + \delta p_t^j + (1 - \delta) m c_t^j \quad (\text{A.32})$$

$$\text{bop: } \beta b_t^j - b_{t-1}^j + \bar{g} \Delta g_t^j = \bar{\Gamma} \int_k \left(p_t^{kj} - x r_t^{kj} - p_t^{jk} \right) dk + \bar{\Gamma} \int_k \left(\hat{c}_t^{kj} dk - \hat{c}_t^{kj} \right) dk \quad (\text{A.33})$$

Since equations (A.28) and (A.29) are expressed in relative terms, they are redundant for one country. In which case, they can be replaced with market clearing for gold — the world stock is constant — and bonds at the world level:

$$\int_k \hat{g}_t^k dk = 0 \quad (\text{A.34})$$

$$\int_k b_t^k dk = 0 \quad (\text{A.35})$$

Equation (A.35) is redundant as well. Indeed, it can be deduced from summing over j in equation (A.33). The right-hand side drops out, so:

$$\beta \int_j b_t^j dj = \int_j b_{t-1}^j dj - \bar{g} \Delta \int_j g_t^j dj$$

The last term is 0 for all t by equation (A.34). The initial condition guarantees: $\int_j b_{-1}^j dj = 0$, so that: $\int_j b_0^j dj = 0$. Iterating over t implies equation (A.35). This redundancy verifies Walras law: if the markets for goods and gold clear, so does the market for bonds.

To make the model stationary, I subtract the local gold price from nominal quantities (p^{kj} , p^{j*} , p^j , w^j , $m c^j$), ϵ^j . The exchange rate can be substituted out. I denote those transformed variables with tildes. Random walk variables can be expressed in first difference: $\Delta \epsilon_t^j = \epsilon_t^j - \epsilon_{t-1}^j$.

$$\text{market clearing: } \hat{y}_t^j = \psi \left(\int_m \hat{m} c_t^m dm - \hat{m} c_t^j \right) + (1 - \bar{\Gamma}) \hat{c}_t^j + \bar{\Gamma} \int_k \hat{c}_t^k dk \quad (\text{A.36})$$

$$\text{marginal cost: } \hat{m} c_t^j = \hat{w}_t^j + \frac{\alpha_1}{1 - \alpha_2} \hat{y}_t^j \quad (\text{A.37})$$

$$\text{CPI: } \hat{p}_t^j = (1 - \bar{\Gamma}) \hat{m} c_t^j + \bar{\Gamma} \int_m \hat{m} c_t^m \quad (\text{A.38})$$

$$\text{CPI inflation: } \pi_{ct}^j = \hat{p}_t^j - \hat{p}_{t-1}^j + \Delta \epsilon_t^j \quad (\text{A.39})$$

$$\text{wage inflation: } \pi_{wt}^j = \hat{w}_t^j - \hat{w}_{t-1}^j + \Delta \epsilon_t^j \quad (\text{A.40})$$

$$\text{marginal utility: } \hat{\mu}_t^j = \frac{\sigma}{(1-\iota)(1-\beta\iota)} \left(\beta\iota E_t \hat{c}_{t+1}^j - (1+\beta\iota^2) \hat{c}_t^j + \iota \hat{c}_{t-1}^j \right) \quad (\text{A.41})$$

$$\text{Euler equation: } \hat{\mu}_t^j = E_t \hat{\mu}_{t+1}^j + \hat{i}_t^j - E_t \pi_{c,t+1}^j \quad (\text{A.42})$$

$$\text{money holdings: } \hat{i}_t^j = \frac{1-\beta}{\beta^2} \left(-\chi(\hat{m}_t^j - \hat{p}_t^j) - \hat{\mu}_t^j \right) \quad (\text{A.43})$$

$$\text{Phillips curve: } \pi_{wt}^j = \beta E_t \pi_{wt+1}^j + \tilde{\kappa} \hat{y}_t^j - \kappa \hat{\mu}_t^j + \kappa(\hat{w}_t^j - \hat{p}_t^j) \quad (\text{A.44})$$

$$\text{money supply: } \hat{m}_t^j = \hat{g}_t^j \quad (\text{A.45})$$

$$\text{gold peg: } E_t \Delta \epsilon_{t+1}^j = 0 \quad (\text{A.46})$$

$$\text{quasi UIP: } \hat{i}_t^j - \hat{i}_t^k - E_t \left(\Delta \epsilon_t^j - \Delta \epsilon_t^k \right) = v(b_t^k - b_t^j) \quad (\text{A.47})$$

$$\text{price setting: } \hat{p}_t^{jk} = \zeta_1 \hat{p}_t^{j*} + \zeta_2 \hat{p}_t^j + (1 - \zeta_1 - \zeta_2) \hat{m} c_t^k \quad (\text{A.48})$$

$$\text{IPI: } \hat{p}_t^{j*} = \delta \hat{p}_t^j + (1 - \delta) \int_k \hat{m} c_t^k \quad (\text{A.49})$$

$$\text{PPI: } \hat{p}_t^{jj} = \delta \hat{p}_t^j + (1 - \delta) \hat{m} c_t^j \quad (\text{A.50})$$

$$\begin{aligned} \text{bop: } \beta b_t^j - b_{t-1}^j + \bar{g} \left(\Delta \hat{g}_t^j \right) \\ = \bar{\Gamma} \int_k \left(\hat{p}_t^{kj} - \hat{p}_t^{jk} \right) dk + \bar{\Gamma} \int_k \left(\hat{c}_t^{kj} dk - \hat{c}_t^{jk} \right) dk \end{aligned} \quad (\text{A.51})$$

D.7 Proof of Proposition 3

I plug equation (A.53) into equations (25–26):

$$\begin{aligned} \text{devaluing: } \hat{y}_t^D &= S^D (\hat{y}_t^D - \hat{y}_t^N) - \sigma^{-1} \frac{(1-\iota)(1-\beta\iota)}{(1-\iota L)(1-\beta\iota L^{-1})} \sum_{s=0}^{\infty} E_t \hat{r}_{t+s}^W \\ \text{non-devaluing: } \hat{y}_t^N &= -(1-S^D) (\hat{y}_t^D - \hat{y}_t^N) - \sigma^{-1} \frac{(1-\iota)(1-\beta\iota)}{(1-\iota L)(1-\beta\iota L^{-1})} \sum_{s=0}^{\infty} E_t \hat{r}_{t+s}^W \end{aligned}$$

Knowing σ , ι , $\hat{y}_t^D - \hat{y}_t^N$ and the path of \hat{r}_t^W is sufficient to know \hat{y}_t^D and \hat{y}_t^N .

E SMM Estimation

To estimate the model, I minimize the weighted sum of the empirical moments:

$$(\hat{\mathcal{B}} - \bar{\mathcal{B}})' \times \hat{V}^{-1} \times (\hat{\mathcal{B}} - \bar{\mathcal{B}})$$

where:

$$\mathcal{B} = \begin{pmatrix} \mathcal{B}^{xr} \\ \mathcal{B}^{ip} \\ \mathcal{B}^{wpi} \\ \mathcal{B}^{nw} \\ \mathcal{B}^{imp} \\ \mathcal{B}^{ipi} \\ \mathcal{B}^{tar} \\ \beta^{elas} \\ \beta^{pt} \end{pmatrix}, \quad V = \begin{pmatrix} V^{xr} & (0) & (0) & (0) & (0) & (0) & (0) & (0) & (0) \\ (0) & V^{ip} & (0) & (0) & (0) & (0) & (0) & (0) & (0) \\ (0) & (0) & V^{wpi} & (0) & (0) & (0) & (0) & (0) & (0) \\ (0) & (0) & (0) & V^{nw} & (0) & (0) & (0) & (0) & (0) \\ (0) & (0) & (0) & (0) & V^{imp} & (0) & (0) & (0) & (0) \\ (0) & (0) & (0) & (0) & (0) & V^{ipi} & (0) & (0) & (0) \\ (0) & (0) & (0) & (0) & (0) & (0) & V^{tar} & (0) & (0) \\ (0) & (0) & (0) & (0) & (0) & (0) & (0) & (\sigma^{elas})^2/4 & 0 \\ (0) & (0) & (0) & (0) & (0) & (0) & (0) & 0 & (\sigma^{pt})^2/4 \end{pmatrix}$$

Each \mathcal{B}^y is itself a vector containing the IRF of y from 1932 to 1935:

$$\mathcal{B}^y = \begin{pmatrix} \beta_{1932}^y \\ \vdots \\ \beta_{1935}^y \end{pmatrix}$$

where \mathcal{B} is a vector containing the estimated moments and V the diagonal matrix whose elements are the variance of those moments. A $\hat{\cdot}$ denotes the empirical estimate, and a $\bar{\cdot}$ the model equivalent. Since each IRF is made of four moments, I re-weight the elasticity and pass-through estimates by a factor of 4; so that they receive the same weight than an IRF, for a given variance.

F Robustness of Elasticity Estimates

F.1 Tariff-Induced Bias

A peculiar feature of inter-war US tariffs may bias my elasticities downward. While the US did not raise tariffs selectively against countries that devalued, custom duties on certain products were a fixed levy per physical unit of good. As a consequence, when the dollar price of these products decreased, the effective *ad valorem* tariff rate increased (Crucini, 1994, Irwin, 1998a). Thus the tariff rate could have been correlated with the exchange rate which would compromise the validity of the instrument.

This point can be formalized by going back to equation (11). It is still valid with tariff if p_{rt}^{jk} is interpreted as the after-tariff price. Unfortunately, I do not directly observe after-tariff prices at the product-country level.¹⁸ Fortunately, I do not need the after-tariff price to infer the elasticity. Indeed, if I break down the price into its before-tariff and tariff components, the equation becomes:

$$\Delta c_{rt}^{jk} = -\theta \Delta \tilde{p}_{rt}^{jk} + \underbrace{\theta \Delta \left(\tau_{rt}^{jk*} + \tilde{p}_{rt}^{jk*} \right) + \Delta c_{rt}^{jk*}}_{\text{fixed effect}} + \underbrace{\Delta \left(\kappa_{rt}^{jk} - \theta \tau_{rt}^{jk} \right)}_{\text{error term}} \quad (\text{A.52})$$

where \tilde{p}_{rt}^{jk} is the before-tariff price and τ_{rt}^{jk} is the effective tariff rate paid on varieties imported from country k . The equation again features a term that can be absorbed by a product-time fixed effect. But the error term now includes country-specific tariff.

While the US tariff code did not discriminate across countries until 1934,¹⁹ the effective tariff rate is not necessarily the same across countries because some products were taxed through a fixed nominal duty, not by a fixed proportional rate. As a result, the effective tariff rate becomes inversely proportional to the price:

$$1 + \tau_{rt}^{jk} \left(\tilde{p}_{rt}^{jk} \right) \propto \frac{1}{\tilde{p}_{rt}^{jk}}$$

¹⁸In theory, it should be possible to impute the relevant tariff duty to each product-country pair. In practice, this is impossible to do because the import data is more aggregated than the tariff data, and the tariff data is not broken down by country. As a result, there are several tariff subcategories within a product category as it is defined in the import data, and I do not observe the country information within those subcategories.

¹⁹Cuba and the Philippines benefited from preferential treatment but they are excluded from the sample.

Which implies that the regression actually identifies:

$$\theta \left(1 + \tau_{rt}^{jk'} \left(\tilde{P}_{rt}^{jk} \right) \right) < \theta$$

There is a downward bias in the coefficient.

To quantify the extent of this bias, I re-estimate the elasticity with the same methodology but allowing it to differ for products that are taxed proportionally, and those that are taxed with a fixed nominal duty.²⁰ The results are displayed in table A.7. I now pool 1930-1932 and 1930-33 changes in order to present the results in a more compact way. The products taxed with ad-valorem tariff indeed display a higher elasticity, closer to 4 than 3.

Table A.7: Tariff-induced bias

	Unweighted		Weighted	
	(1)	(2)	(3)	(4)
	All	Ad-valorem	All	Ad-valorem
Panel A: IV				
Elasticity	2.731*** (0.379)	3.784*** (0.763)	2.964*** (0.938)	4.020** (1.689)
Observations	7042	2912	7042	2912
F-statistic	17.170	5.890	13.890	8.970
Panel B: OLS				
Elasticity	1.150*** (0.065)	1.013*** (0.080)	0.878*** (0.174)	0.629* (0.334)
Observations	7042	2912	7042	2912

Note: Observations are unweighted. Standard errors are in parentheses. They are two-way clustered at the country and product levels. F-statistics is the Kleibergen-Paap Wald rk F-statistic of the first stage.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

F.2 Channel

What matters for the identification of θ is that devaluation be uncorrelated with US demand shocks. Could it be, however, that devaluation correlates with supply shocks that drive the results? θ would still be identified but the interpretation of the reduced form results would be different. Exports would not increase because of the change in the exchange rate, but because of other shocks for

²⁰The data on tariff rates was first used by Bond et al. (2013). For the needs of this project, I re-digitized it in collaboration with Acosta and Cox (2021).

which the exchange rate is a proxy. The placebo regressions ruled out shocks that would have occurred before 1931, but what about afterwards?

An obvious suspect is the financing channel (Bernanke, 1983, Bernanke and James, 1991). Devaluations certainly correlated with monetary policy loosening, which could have eased firms' financing conditions and led them to export more. To control for such an effect, I add to the reduced form equation proxies for monetary policy: the change in the discount rate (DR), in money and notes in circulation (M0), and in the deposit to currency ratio (DCR). I show the results in column (2) of table A.8, to be compared to column (1) where there are no controls. Like in table A.7, I pool the 1930-1932 and 1930-33 changes to present the results in a compact way. In column (3), I add the panic variable of Bernanke and James (1991) and the unemployment rate (UR). The results stay the same.

I then go on to various robustness checks. In column (4), I instrument the exchange rate with a devaluation dummy. In column (5), I interact it with a dummy (XC) that takes value 1 if the country implemented exchange controls.

Overall, table A.8 suggests that the estimated substitution occurs because of the change in the exchange rate, not because of other shocks for which the exchange rate is a proxy.

F.3 Heterogeneity

So far, I restricted all goods to have the same elasticity. I relax this assumption here by allowing the elasticities to differ across 10 product categories. These product categories are those that are distinguished by the *Foreign Commerce and Navigation of the United States*: animal, vegetable edible, vegetable inedible, textiles, wood and paper, nonmetallic minerals, metals and manufactures, machinery and vehicles, chemicals and miscellaneous. To gain statistical power, I now group both the 1930-32 and 1930-33 changes.

I show the results in table A.9. Point estimates vary widely, but confidence intervals are large, so that no obvious pattern emerge. Yet, a test of the equality of elasticities across product categories comfortably rejects the null. Anyhow, table A.9 makes clear, that the previous results are not driven by a single class of products. Each category shows the same pattern: a positive coefficient for quantity, and a negative one for prices.

Table A.8: Robustness

	(1)	(2)	(3)	(4)	(5)
XR	1.104*** (0.245)	0.885** (0.390)	1.030** (0.409)	0.704** (0.317)	
DR		-5.870 (4.620)	1.851 (7.165)		
M0		-0.083 (0.430)	-0.425 (0.976)		
DCR		0.274 (0.483)	-0.188 (0.665)		
Panics			-0.002 (0.014)		
UR			-0.012 (0.010)		
XC=0 \times XR					0.791*** (0.264)
XC=1 \times XR					1.362*** (0.168)
Observations	7242	5944	4813	7242	7242

Note: In columns (1–2), I estimate: $\Delta c_{rt}^{us,j} = \beta \Delta x r_t^{us,j} + \zeta' X_{jt}^j + \mu_{rt}^{us} + e_{rt}^{us}$, where X_t^j are the various controls shown in the table. In column (3), I instrument the exchange rate with a dummy for whether the country devalues or not. In column (4), I estimate: $\Delta c_{rt}^{us,j} = \beta \Delta \epsilon_t^{us,j} \times XC_t^j + XC_t^j + \mu_{rt}^{us} + e_{rt}^{us,j}$ where XC_t^j is a dummy which takes value 1 if the country implemented exchange controls. The sample of countries varies with data availability. Observations are unweighted. Standard errors are in parentheses. They are two-way clustered at the country and product levels.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

F.4 An International Elasticity Puzzle?

The macroeconomics and trade literatures have disagreed on the elasticity of substitution between the traded goods of different countries. Old estimations using aggregated data converged toward an elasticity of 1 (Goldstein and Khan, 1985). More recent estimations based on dis-aggregated trade data came up with much higher estimates (Head and Ries, 2001, Feenstra, 1994, Romalis, 2007, Broda and Weinstein, 2006, Caliendo and Parro, 2015).

Various explanations have been proposed. Ruhl (2008), who coined the term “international elasticity puzzle,” emphasized that there is a fixed cost to entering export markets. Hence, trade flows are not very responsive to transitory variations, which are typically used to estimate macro elasticities. On the other hand, they respond a great deal to permanent changes, like tariffs, which are often the source of identifying variation in the trade literature. A second explanation would be that differences reflect the level of aggregation of the data. Broda and Weinstein (2006) show that

Table A.9: Heterogeneity

	RF		FS		IV	
	(1)	(2)	(3)	(4)	(5)	(6)
Animal products	1.165***	(0.327)	−0.192	(0.132)	6.079*	(3.485)
Vegetable edible	1.296**	(0.512)	−0.527***	(0.157)	2.460**	(0.889)
Vegetable inedible	0.509*	(0.273)	−0.343***	(0.062)	1.486*	(0.855)
Textiles	1.294***	(0.364)	−0.505***	(0.097)	2.564***	(0.606)
Wood and paper	0.387	(0.493)	−0.562**	(0.252)	0.688	(0.639)
Nonmetallic minerals	0.850	(0.951)	−0.550	(0.401)	1.545	(0.934)
Metals and manufactures	1.638**	(0.734)	−0.352	(0.234)	4.656**	(1.824)
Machinery and vehicles	2.338**	(0.998)	−0.879	(0.903)	2.659	(3.148)
Chemicals	1.394*	(0.721)	−0.320**	(0.148)	4.358**	(1.916)
Miscellaneous	1.342**	(0.490)	−0.691**	(0.270)	1.941**	(0.912)
Observations	7042		7042		7042	
P-value	0.153		0.218		0.004	

Note: Reduced form (RF), first stage (FS) and instrumented (IV) estimates of equation (12), where coefficients are allowed to differ across categories. Observations are unweighted. Standard errors are in parentheses. They are two-way clustered at the country and product levels. P-value is the p-value for a test of equality of elasticities across product categories.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

using broader categories of goods can substantially lower the elasticity. Imbs and Méjean (2015) argue that the discrepancy can be explained by aggregation biases stemming from heterogeneity in elasticities. Feenstra et al. (2018) propose a third explanation: that these two literatures do not estimate the same parameters. The macro literature estimates the elasticity of substitution between imports and domestic goods, while the trade literature estimates elasticities of substitution between the goods of different foreign countries.

Judging by my estimates, there is no puzzle. Since I have about a thousand product categories, the equivalent level of aggregation in modern data is SITC-5. The medians reported by Broda and Weinstein for this level of aggregation are 2.8 and 2.7, which fall within the 2–4 range that I estimate. Their means are higher but, as they remark, they tend to be overly sensitive to a few extremely elastic goods. The source of variation that I am using is transitory: while the change in the nominal exchange rate may be permanent, prices adjust over the medium run, offsetting gains in competitiveness. Therefore, my siding with the trade literature would go against the fixed cost explanation (Ruhl, 2008).

I cannot discriminate between the second (Broda and Weinstein, 2006, Imbs and Méjean, 2015) and third (Feenstra et al., 2018) explanations with those results alone. The outcome of the struc-

tural estimation, however, favors the latter for two reasons. First, I manage to account for the behavior of aggregate exports with the elasticity estimated in micro data: aggregation does not seem to matter here. Second, that aggregate imports' response is muted while aggregate exports jump up is at least suggestive of more elastic exports. Indeed, I find that a low elasticity between imports and domestic goods is important to account for the behavior of aggregate imports.

G A Sufficient Statistics Approach

In this section, I relax the assumptions made on money demand and supply in section 5. To do so, I develop a sufficient statistics approach and bring additional evidence on the effect of devaluation on the world real rate. This exercise shows that my main conclusion is robust to a more general framework.

G.1 Theoretical Result

Consider again equations (25–26), which I reproduce here for convenience:

$$\text{devaluing countries:} \quad \hat{y}_t^D = (1 - S^D) (\hat{y}_t^D - \hat{y}_t^N) + \hat{y}_t^W \quad (25)$$

$$\text{non-devaluing countries:} \quad \hat{y}_t^N = -S^D (\hat{y}_t^D - \hat{y}_t^N) + \hat{y}_t^W \quad (26)$$

y is output in devaluing countries (D), non-devaluing countries (N), and at the world level (W). S^D is the mass of countries that devalue. In the model, \hat{y}_t^W is pinned down by the path of world real interest rates. This is an implication of iterating forward in the Euler equation. For instance, without habits ($\iota = 0$), world output can be simply written as:

$$\hat{y}_t^W = -\sigma^{-1} \sum_{s=0}^{\infty} E_t \hat{r}_{t+s}^W$$

The presence of habits ($\iota > 0$) slightly complexifies the formula but does not alter its main intuition:

$$\hat{y}_t^W = -\sigma^{-1} \frac{(1 - \iota)(1 - \beta\iota)}{(1 - \iota L)(1 - \beta\iota L^{-1})} \sum_{s=0}^{\infty} E_t \hat{r}_{t+s}^W \quad (\text{A.53})$$

where L is the lag operator.

As I discussed in section 6.3, equations (25–26) show that the effect on devaluing countries' (\hat{y}_t^D) and non-devaluing countries' (\hat{y}_t^N) output can be decomposed into the relative effect ($\hat{y}_t^D - \hat{y}_t^N$) and the effect on world output (\hat{y}_t^W). Since the model matches the relative effect on output, its predictions for \hat{y}_t^D and \hat{y}_t^N are as good as its prediction for \hat{y}_t^W . Equation (A.53) now makes clear that this prediction depends on how sensitive consumption is to real interest rates through σ^{-1} and ι , and the path of the world real rate.

Taken together equations (25) and (A.53) show that knowing σ , ι , $\hat{y}_t^D - \hat{y}_t^N$ and the path of \hat{r}_t^W is enough to answer the question of interest. I formalize this statement in the following proposition:²¹

Proposition 3 (Sufficient statistics) *σ , ι , $\hat{y}_t^D - \hat{y}_t^N$ and the path of \hat{r}_t^W are sufficient statistics for \hat{y}_t^D and \hat{y}_t^N in devaluing and non-devaluing countries.*

G.2 Effect on the World Real Interest Rate

Proposition 3 established that the path of the real rate is key to the result. A model where devaluation is not expansionary at the world level would imply that devaluation is purely a zero-sum game.²² A necessary condition for the main result of the paper to be true is that devaluation be expansionary at the world level.

I provide empirical discipline on devaluations' effect on the world real interest rate. Empirically, constructing a world real interest rate is challenging. To do so, I gather data on the real rate of as many countries as possible and compute a weighted index of their real rate:

$$1 + r_t^W = \prod_{n=1}^N \left(1 + r_t^j\right)^{w_n}$$

The weights w_n are given by their weighted share in world GDP in 1929. Those countries accounted for about 70–75% of world GDP and 60–65% of world trade in 1929.²³

To identify the effect of the devaluations on the world real rate, I use the high-frequency shocks of section 2.2. Indeed, those shocks are identified *time series* shocks. They are valid to infer causal

²¹See the appendix for a complete proof.

²²This is what happens in Caballero et al. (2021). In their environment, the world is at the zero-lower bound and exchange rate depreciation does not lead to any change in the world real interest rate.

²³Those countries are listed in table A.1. In figure A.11 and table A.6, I weight countries by their share in world trade instead of GDP.

effects in the time series. To do so, I estimate:

$$r_{t+k}^W - r_{t-1}^W = \beta_k(\epsilon_t^W - \epsilon_{t-1}^W) + \gamma_k' X_{t-1}^W + \zeta_k + e_{t,k} \quad (\text{A.54})$$

where $\epsilon_t^W - \epsilon_{t-1}^W$ is instrumented with a weighted average of the high-frequency shocks. ϵ_t^W is the weighted average of nominal prices of gold in national currencies, i.e. the world exchange rate against gold. The weights are the same as those used to construct the real rate index.

I show the result in figure A.15. I present results using the wholesale (WPI) and consumer (CPI) price indices. Although the IRF is estimated with noise, it indicates that the real rate falls over the estimable horizon. Since it is the sum of future real rates which appears in equation (A.53), I formally compute the average of the β_k , for $0 \leq k \leq 12$, in table A.10. Like in section 2.2, I compute that average (column 1), the same average for the spot exchange rate (2), and the ratio of the two (3). These objects are all negative and significant at the 1% level. The coefficients are semi-elasticities. The scale is easiest to interpret in column (3). It means that a devaluation leading to an average depreciation of 15% (against gold) implies a cut of more than 500 basis points (0.35×0.15 or 0.41×0.15), on average, over 3 years — a substantial expansion, to say the least.

G.3 Model Estimation and Simulation

The model is the same as before, except that assumptions about monetary policy are relaxed. Devaluation is still an exogenous change in the exchange rate, and uncovered interest rate parity holds. But I do not make any other assumption about monetary policy. In particular, the equations for money demand (19) and supply (15) do not have to be true anymore.²⁴

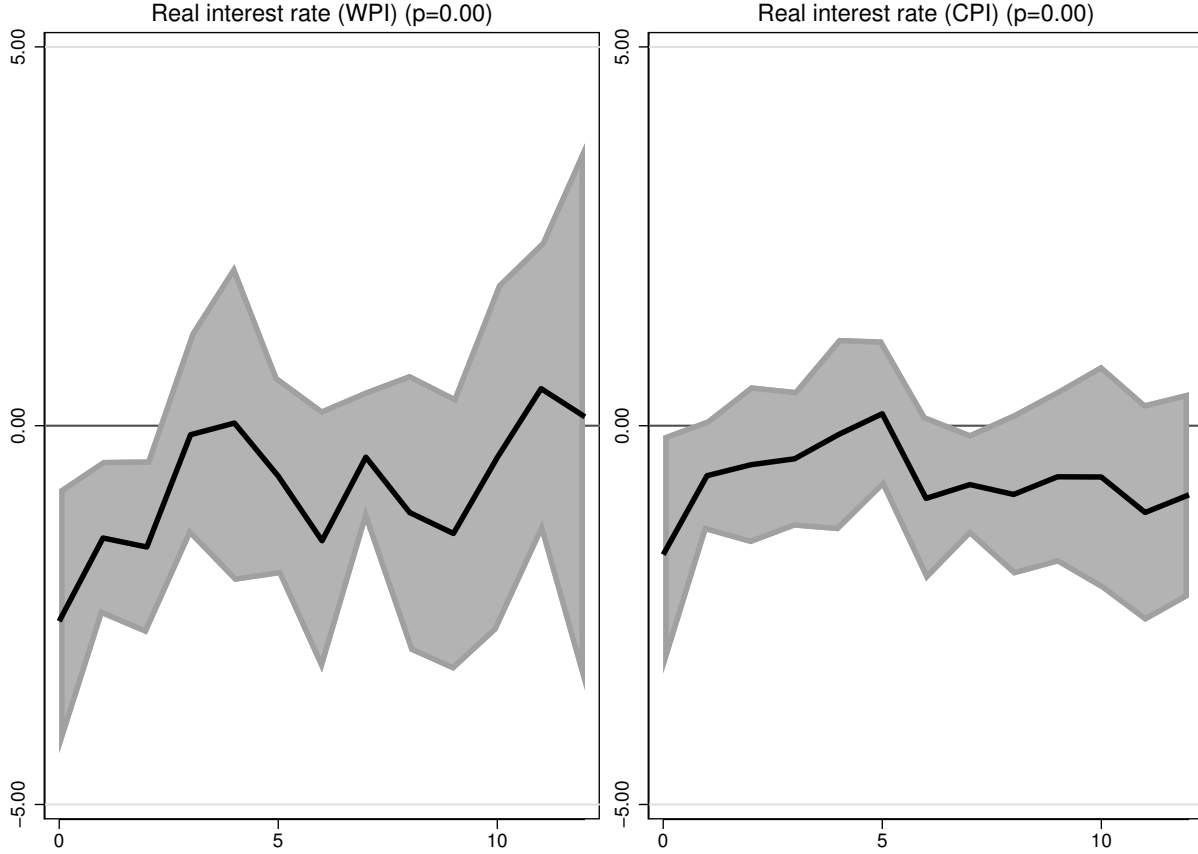
First, I solve the model in relative terms and estimate it out of cross-sectional moments. For any variable x , I denote with superscript R the difference between countries that devalue, and countries that don't:

$$x_t^R \equiv x_t^D - x_t^N$$

I can solve for relative prices and quantities without taking a stance on monetary policy rules and match cross-sectional moments just as in section 6.1. In fact, this new estimation produces the

²⁴To avoid taking a stand on gold flows, I now assume that their value is small compared that of production ($g^* = 0$). Those gold flows are quantitatively irrelevant because they have negligible effects on the current account.

Figure A.15: Effect on the world real rate



Note: response of relevant variable to an exogenous 100 log-points devaluation in the exchange rate. Formally, the plots show the estimate of β_k in equation (A.54). The black line is the point estimate; the gray area is the 95% confidence interval with Driscoll-Kraay standard errors. The number in parenthesis is the p-value of a test of the joint significance of the coefficients.

Table A.10: Effect on the world real rate

	(1) Numerator	(2) Denominator	(3) Ratio
Real rate (WPI)	-0.825*** (0.284)	2.459*** (0.485)	-0.336*** (0.091)
Real rate (CPI)	-0.717*** (0.235)	1.865*** (0.330)	-0.385*** (0.110)

Note: point estimates and standard deviations retrieved from averaging the coefficients whose estimations is detailed in section G.2. The numerator is the average coefficient for the relevant variable, the denominator is the average coefficient for the exchange rate. The ratio is the ratio of the two.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

same parameter estimates as before because it nests the more restrictive model of section 5.1. This allows me to recover the IES (σ^{-1}) which, as equation (A.53) shows, determines the sensitivity of world output to the world interest rate. I then feed the path of the world nominal price of gold shown on figure A.12 into the cross-sectional model to obtain $\hat{y}_t^D - \hat{y}_t^N$.²⁵

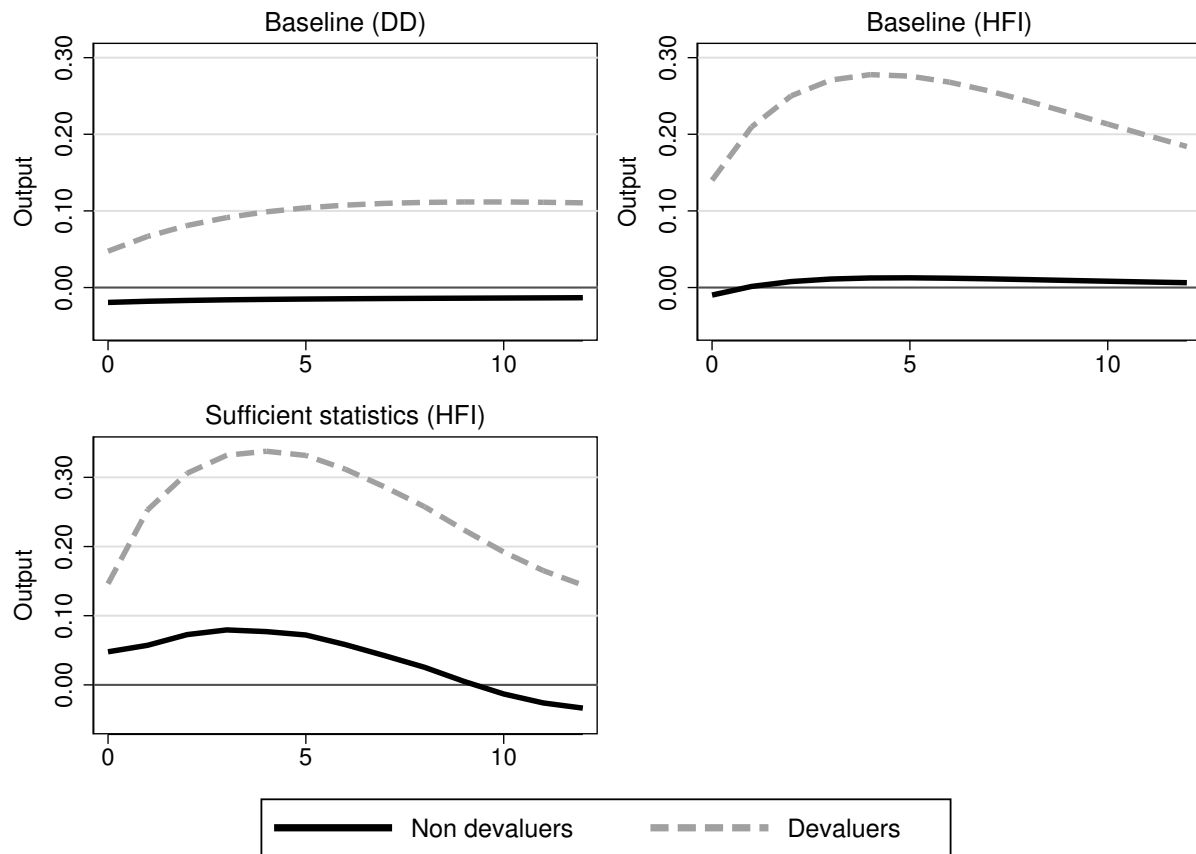
Second, I use the IRF for \hat{r}_t^W to pin down world output, and recover the absolute effect. I feed the path of \hat{r}_t^W shown on figure A.15 into equation (A.53) to obtain that of \hat{y}_t^W . The paths of \hat{y}_t^D and \hat{y}_t^N can then be deduced from equations (25–26).

Again, the approach that I take here nests the model of section 5.1, but it is much more general: it does not assume a specific rule for the money supply or a specific money demand, it can accommodate the zero-lower bound or regime changes, etc. Before, the path of the world real rate relied on a specific model of the gold standard. Now, it is determined by empirical moments.

I show the results in figure A.16. In the first two panels, I use the baseline monetary policy rule, estimating the model out of the cross-sectional IRF implied by the DD (top-left hand side panel) and HFI (top-right hand side) strategies. These pictures are the same as the left-hand side panels of figures 6 and A.6. While those two deliver the same qualitative takeaway, the result is even starker when I use the HFI-identified responses. In the third panel, I use the procedure described in this section. In this scenario, the non devaluers even experience a short-lived stimulus; the effect turns negative after 9 quarters, but that is partly an assumption. Since it is hard to estimate the effect on the real rate beyond 3 years, I assume the latter to be 0 after quarter 12. This implies that the devaluation becomes a zero-sum policy beyond that horizon, hence stacking the odds against my main conclusion.

²⁵Like in section 6, I assume that this depreciation is driven by half of the world, and scale it so that their exchange rate depreciates by 30% on average over three years.

Figure A.16: Sufficient Statistics



Note: response of output in countries that do and do not devalue under the baseline monetary policy rule of section 5.1 (top) and the sufficient statistics procedure described in appendix G.3 (bottom). See figure A.14 for confidence intervals around the IRF implied by the sufficient statistics approach.