

Multi-state Model for the Estimation of Overdiagnosis in the Breast Cancer Screening Program

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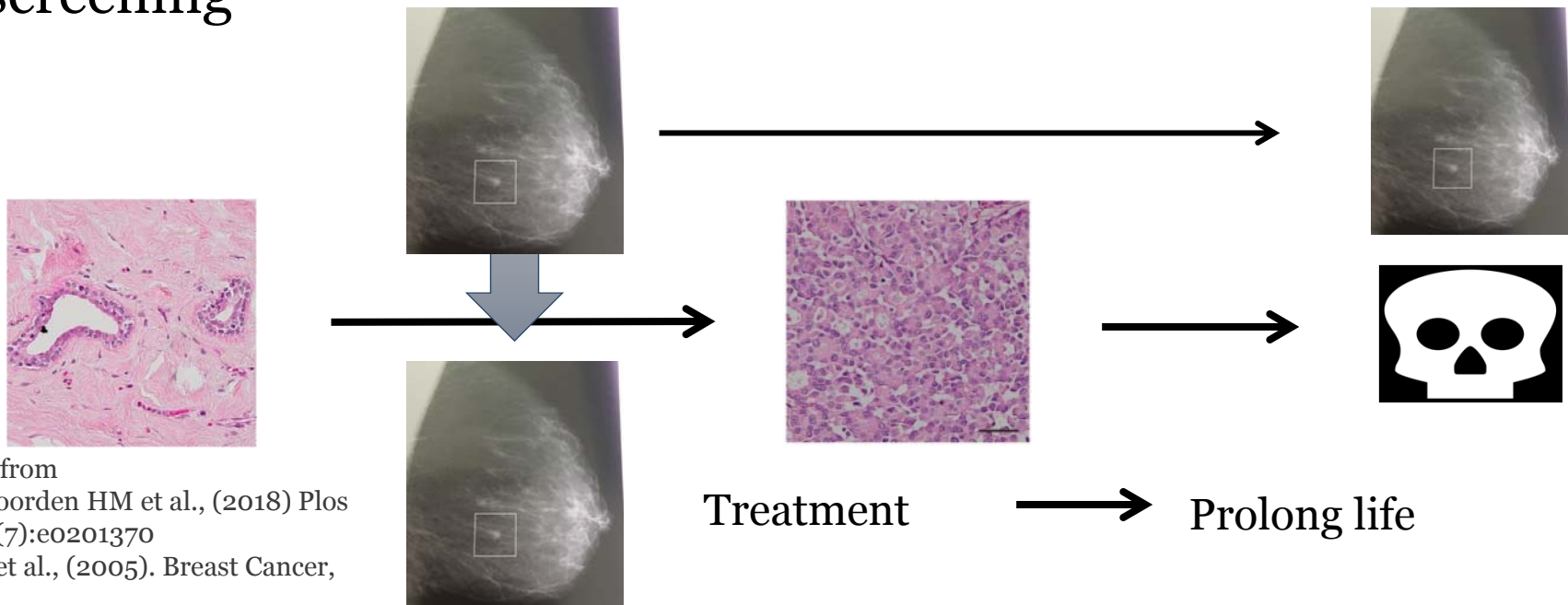
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Overdiagnosis

Screening detects a breast cancer that would not have presented clinically in the woman's lifetime in the absence of screening



Pictures from
Duivenvoorden HM et al., (2018) Plos
ONE, 13(7):e0201370
Tabar L et al., (2005). Breast Cancer,
Thieme

Quantification is difficult!

Estimates varies from 0-56% (EUROSCREEN)

**Control
group?**

**Evaluation
method?**

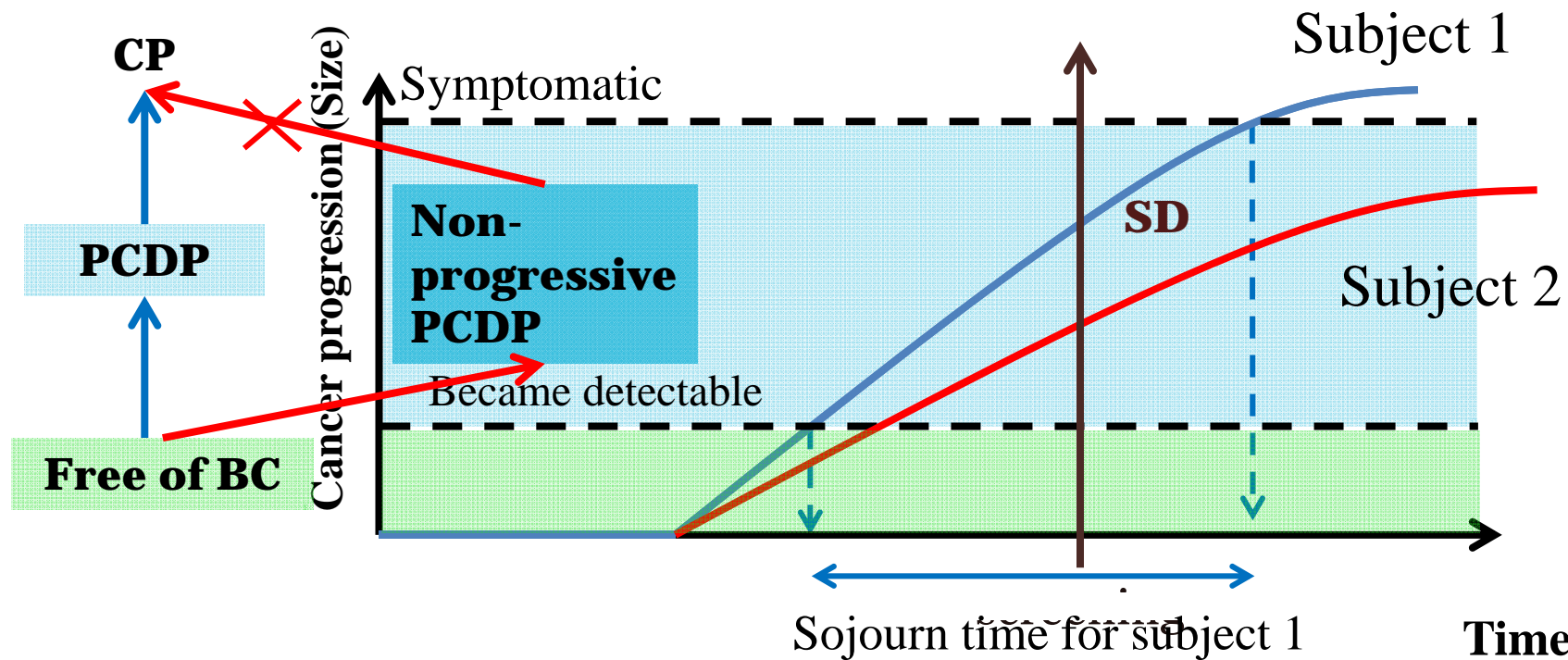
**Useful method for
service-screening?**

Definition?

**Lead time
adjustment?**

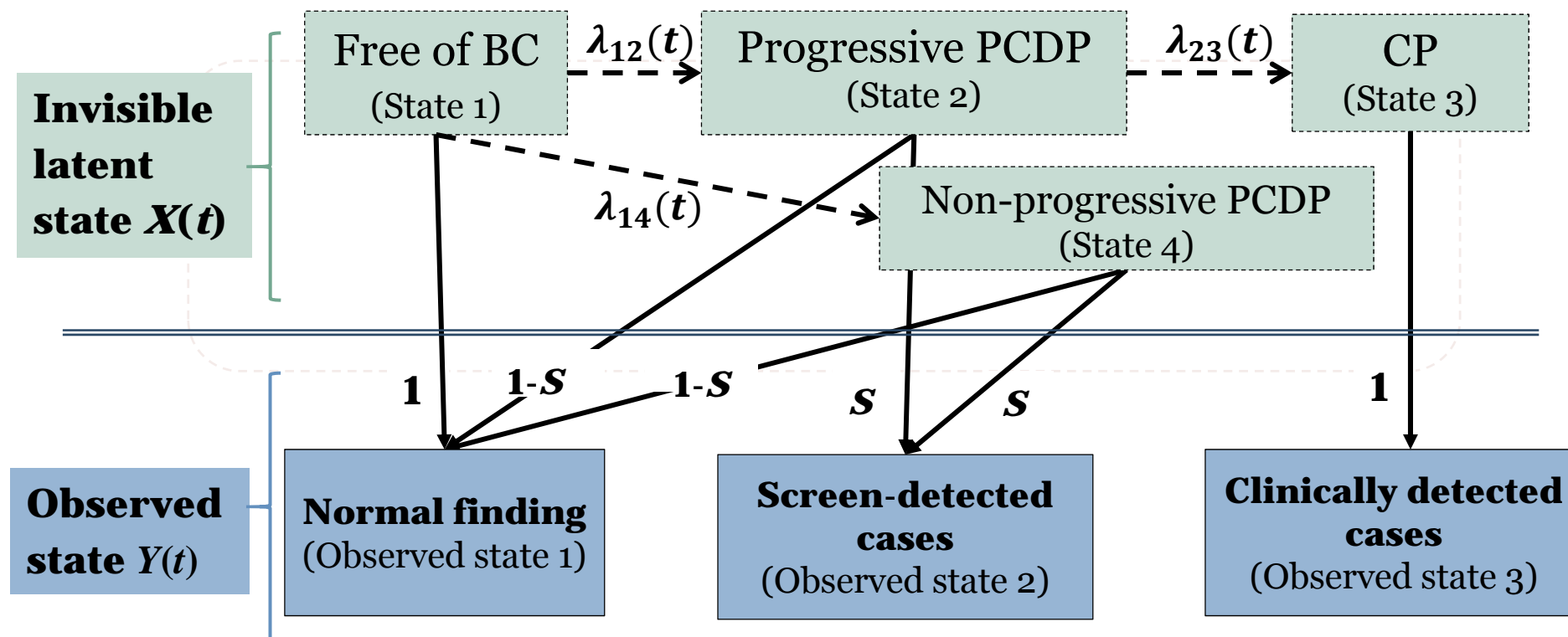
Denominators?

Illustration



BC: Breast cancer, PCDP: Preclinical screen-detected phase, CP: Clinical phase, SD: Screen-detected case

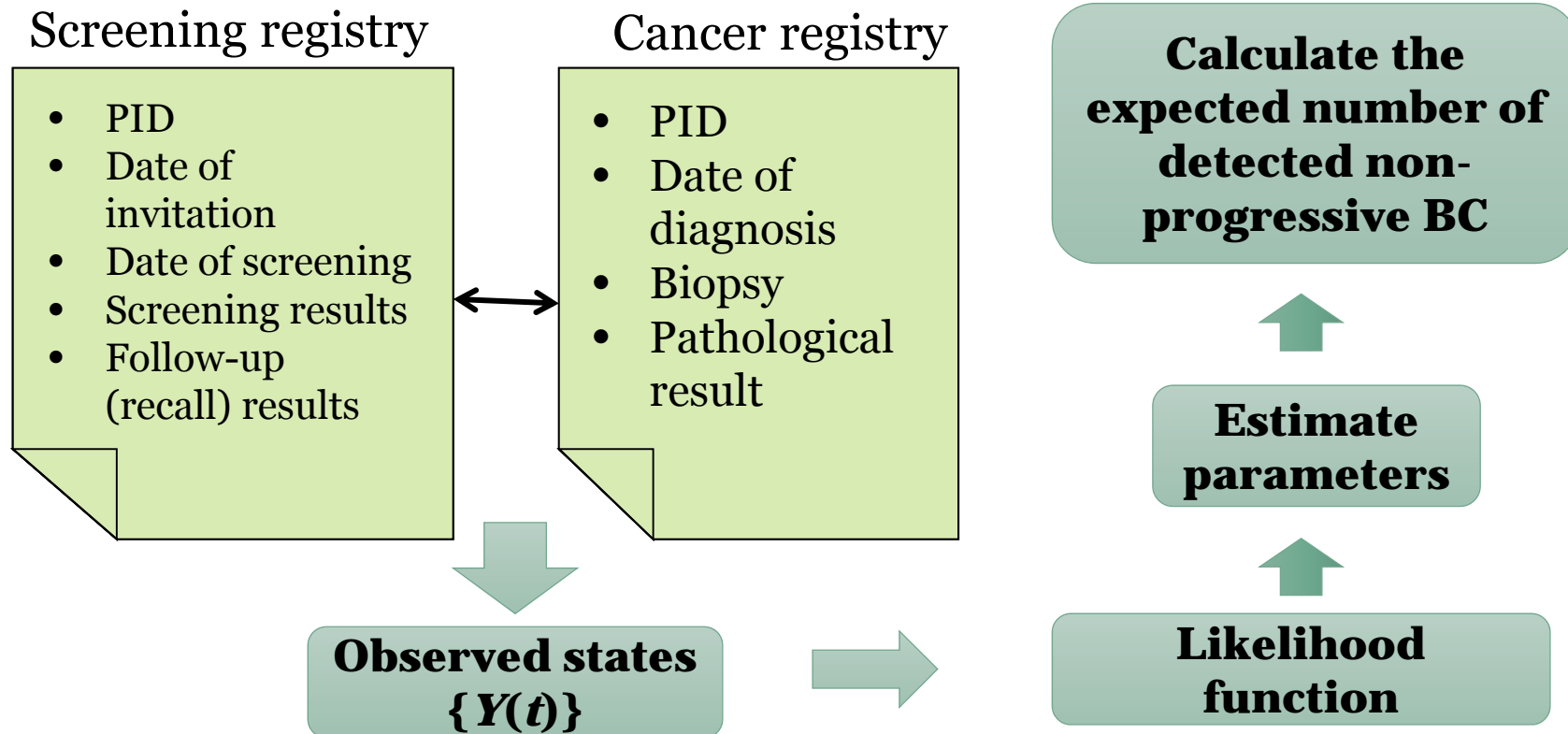
Multi-state model



$\lambda_{ij}(t)$: transition rate from state i to state j at time t , S : test sensitivity

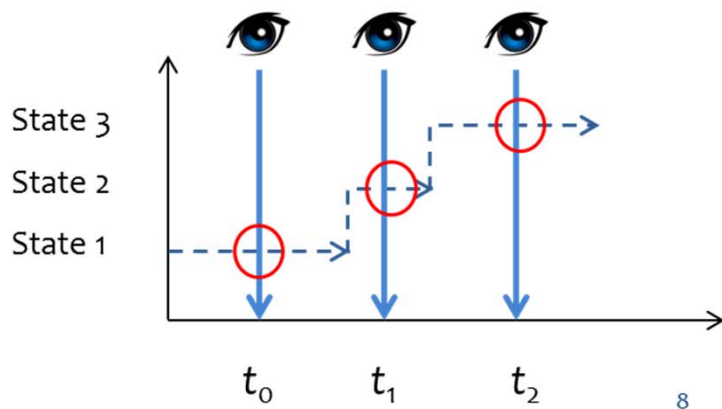
BC: breast cancer, PCDP: preclinical screen-detectable phase, CP: clinical phase

Data sources and work process



Screening Data

- Panel data (longitudinal follow-up data)
 - State only observed at a finite number of times (interval censoring)
 - Don't know the state between times



Let $Y(t)$ denote the state occupied at time t by a random chosen individual

$$\{Y(t_0) = 1, Y(t_1) = 2, Y(t_2) = 3\}$$

Screening Data

- Left truncation
 - Only the asymptomatic subjects will be invited
- Time-inhomogeneous model
 - The incidence rate depends on the age
- Measurement error
 - Sensitivity \neq 100%
- Informative censoring ??
 - The subjects with shorter time to clinical phase tend to be clinically detected if the participation rate is low

Conditional
probability

Piece-wise rate/
parametric

Hidden Markov
model

EM algorithm??

Supplementary Table 2. The Observed Data and Corresponding Likelihood in The Screened and Control Groups by Modes of Detection

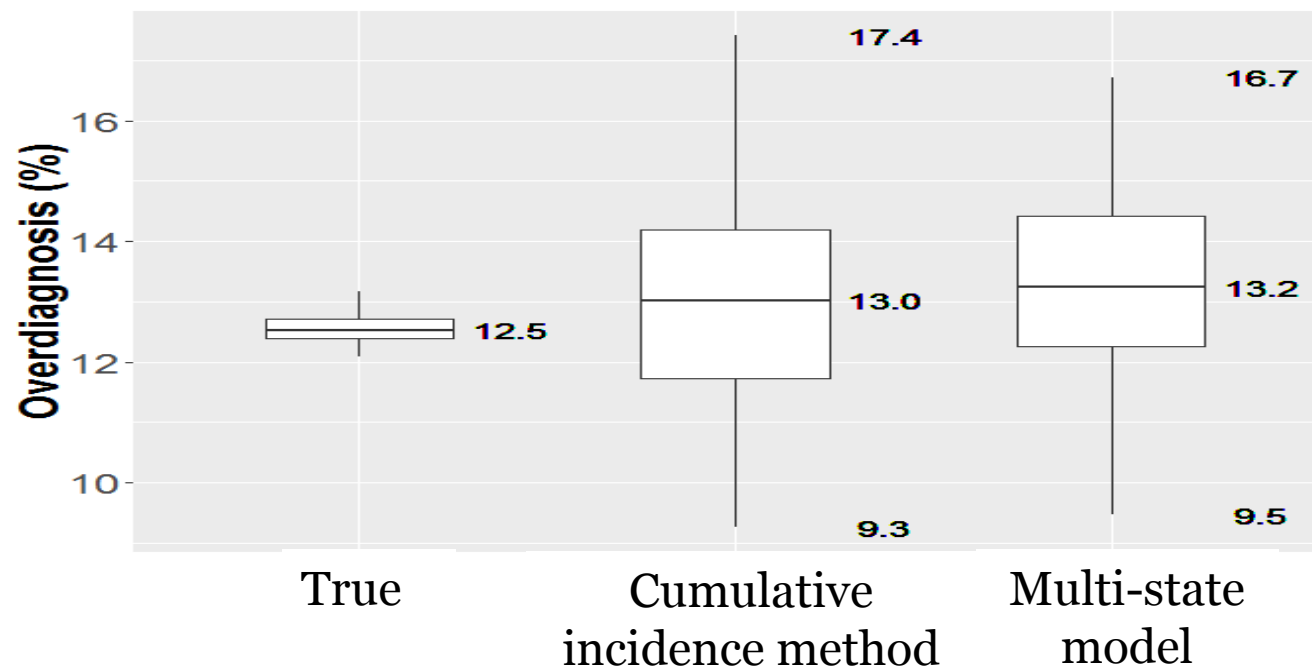
Type	Observed states	Likelihood ^a
Screened group		
SD cases at k^{th} round	$\{Y(t_1) = 2 \mid X(t_1) \neq 3\}$ for $k=1$, and $\{Y(t_k) = 2, Y(t_{k-1}) = 1, \dots, Y(t_1) = 1 \mid X(t_1) \neq 3\}$ for $1 < k \leq m$	$\frac{\left(\prod_{\ell=0}^{k-2} P_{11}(t_{\ell-1}, t_{\ell}) \right)}{1 - P_{13}(t_0, t_1)} \times \left\{ P_{11}(t_{k-2}, t_{k-1}) \times (P_{12}(t_{k-1}, t_k) + P_{14}(t_{k-1}, t_k)) \times S \right.$ $\left. + (P_{12}(t_{k-2}, t_{k-1}) \times P_{22}(t_{k-1}, t_k) + P_{14}(t_{k-2}, t_{k-1})) \times (1 - S) \right\}$
CD cases between k^{th} and $k+1^{\text{th}}$ round	$\{Y(t) = 3, Y(t_k) = 1, \dots, Y(t_1) = 1 \mid X(t_1) \neq 3\}$ for $1 \leq k < m$, and $t_k < t < t_{k+1}$.	$\frac{\left(\prod_{\ell=0}^{k-1} P_{11}(t_{\ell-1}, t_{\ell}) \right)}{1 - P_{13}(t_0, t_1)} \times \left\{ P_{11}(t_{k-1}, t_k) \times P_{12}(t_k, t) \times \lambda_{23}(t) \right.$ $\left. + P_{12}(t_{k-1}, t_k) \times (1 - S) \times P_{22}(t_k, t) \times \lambda_{23}(t) \right\}$
NF women in the last round of screening	$\{Y(t_m) = 1, Y(t_{m-1}) = 1, \dots, Y(t_1) = 1 \mid X(t_1) \neq 3\}$	$\frac{\left(\prod_{\ell=0}^{m-1} P_{11}(t_{\ell-1}, t_{\ell}) \right)}{1 - P_{13}(t_0, t_1)} \times \left\{ P_{12}(t_{m-1}, t_m) \times (1 - S) + P_{11}(t_{m-1}, t_m) \right\}$
Control group		
CD	$\{Y(t) = 3 \mid X(t_1) \neq 3\}$ for $t_1 < t \leq t_m$	$\frac{1}{1 - P_{13}(t_0, t_1)} \times \left\{ P_{11}(t_0, t_1) \times P_{12}(t_1, t) \times \lambda_{23}(t) + P_{12}(t_0, t_1) \times P_{22}(t_1, t) \times \lambda_{23}(t) \right\}$
Non-BC cases	$\{Y(t_m) \neq 3 \mid X(t_1) \neq 3\}$	$\frac{1}{1 - P_{13}(t_0, t_1)} \times \left\{ P_{11}(t_0, t_1) \times (1 - P_{13}(t_1, t_m)) + P_{12}(t_0, t_1) \times P_{22}(t_1, t_m) + P_{14}(t_0, t_1) \right\}$

^a for $\ell \leq 0$, $P_{1j}(t_{\ell-1}, t_{\ell}) = 1$ if $j = 1$, otherwise $P_{1j}(t_{\ell-1}, t_{\ell}) = 0$

BC: breast cancer; CD: clinically detected; m : the total number of screening; NF: negative finding; SD: screen-detected cases; t_k : the time at k^{th} round of screening; $X(t)$: invisible latent state; $Y(t)$: observed state

Estimation of overdiagnosis in breast cancer screening using a non-homogeneous multi-state model: A simulation study

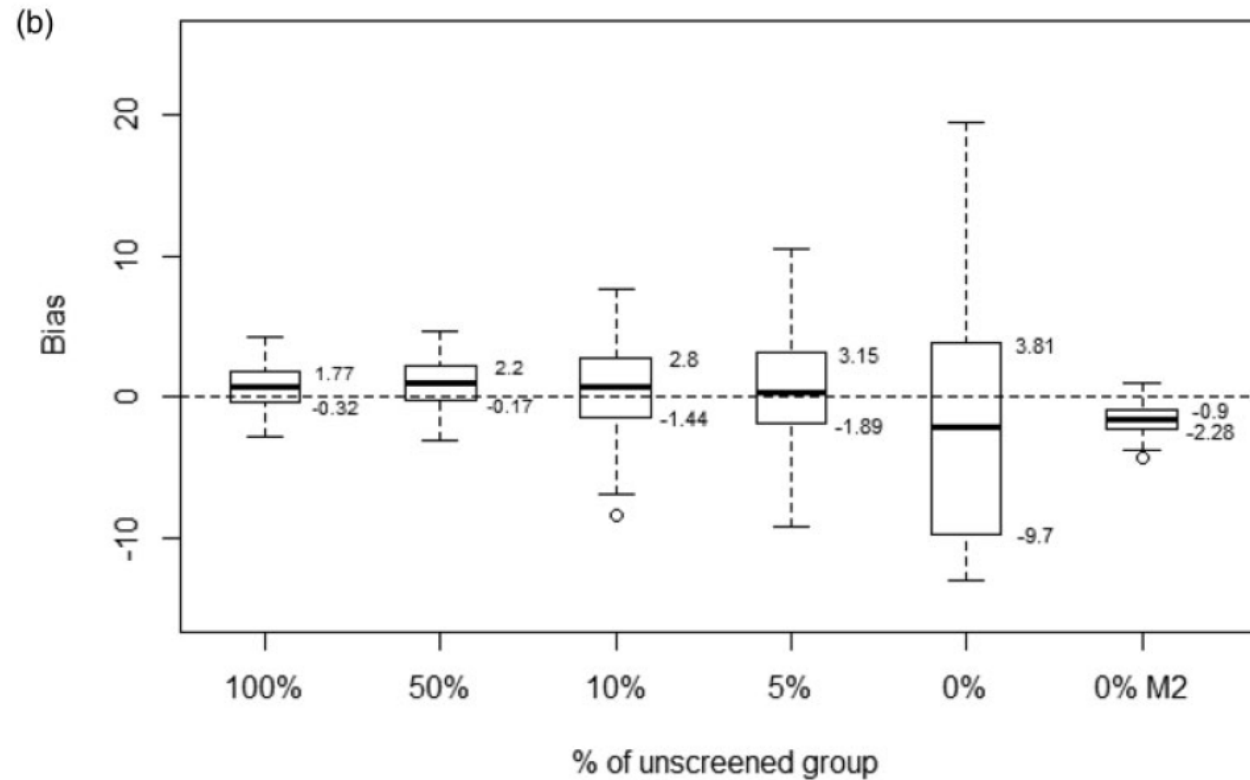
J Med Screen
0(0) 1–8
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DOI: 10.1177/0969141317733294
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SAGE



Randomized
controlled
trial

1: 1 ratio screened
and control,
100% participation
rate

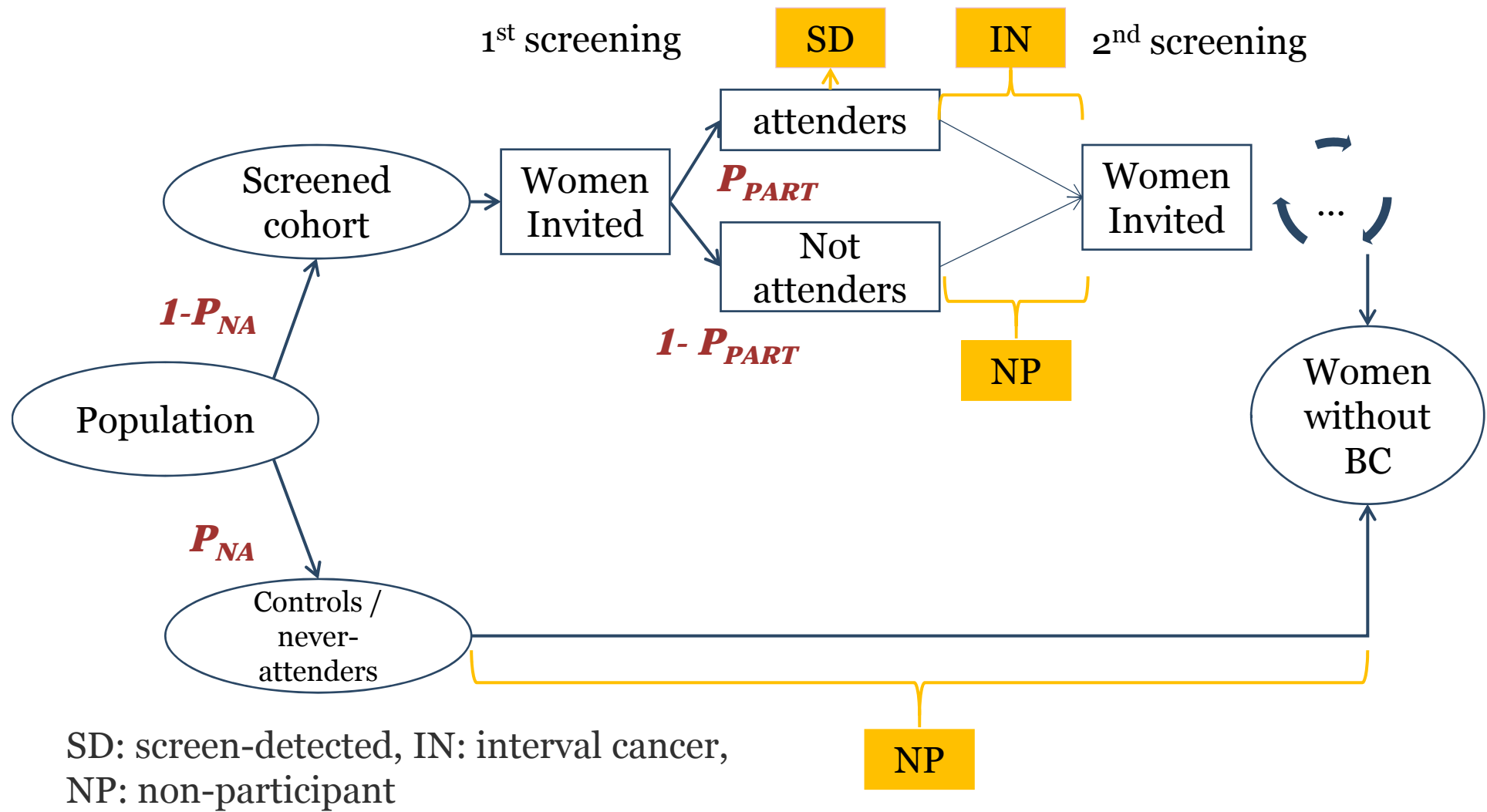
Conclusion: Both methods can provide a proper estimate



**Control group provided information to stabilize the model.
(help to deal with the identifiability problem)**

Control group in the service screening program?

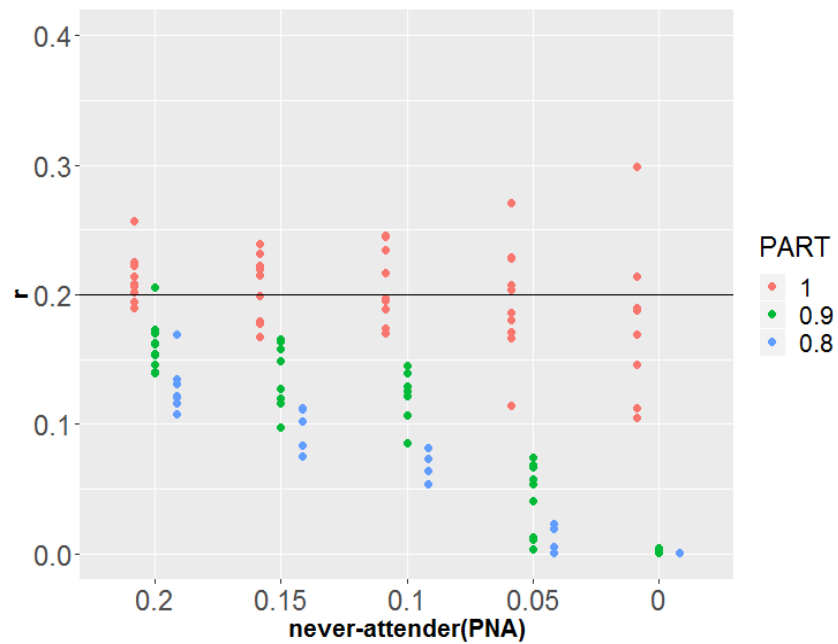
- **Never-attenders** (who were invited but never participated in the screening program) can provide the similar information under certain conditions, i.e. similar incidence rate
 - Group A: those who will never attend any round of screening
 - Group B: those who might attend the screening but does not participate in the current/previous rounds
 - Shorter time to CP → few rounds of invitation → more clinical cases (**observation process and disease process are not independent= informative censoring??**)



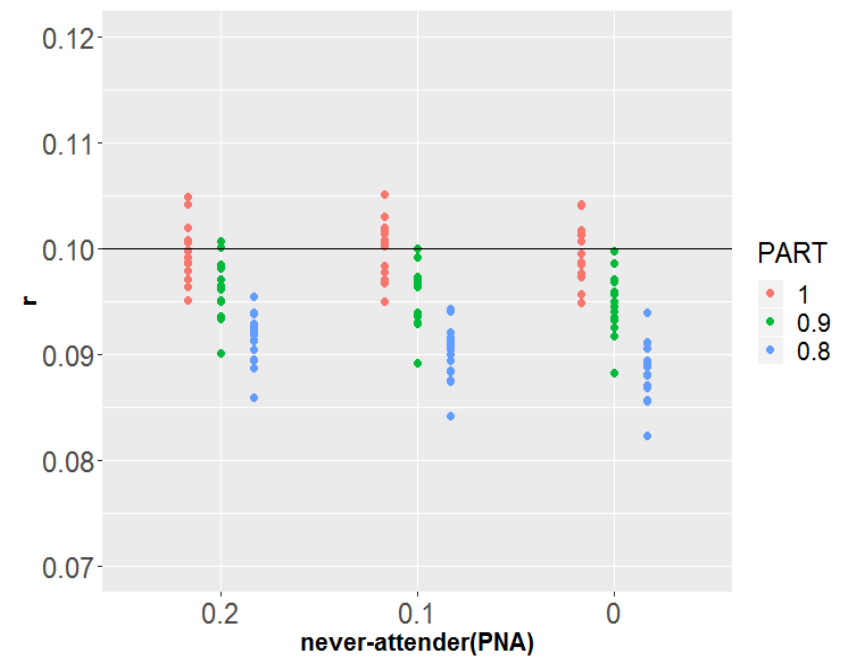
$$r = \frac{\lambda_{14}(t)}{\lambda_{12}(t)}$$

Simulation results

Non-homogeneous model

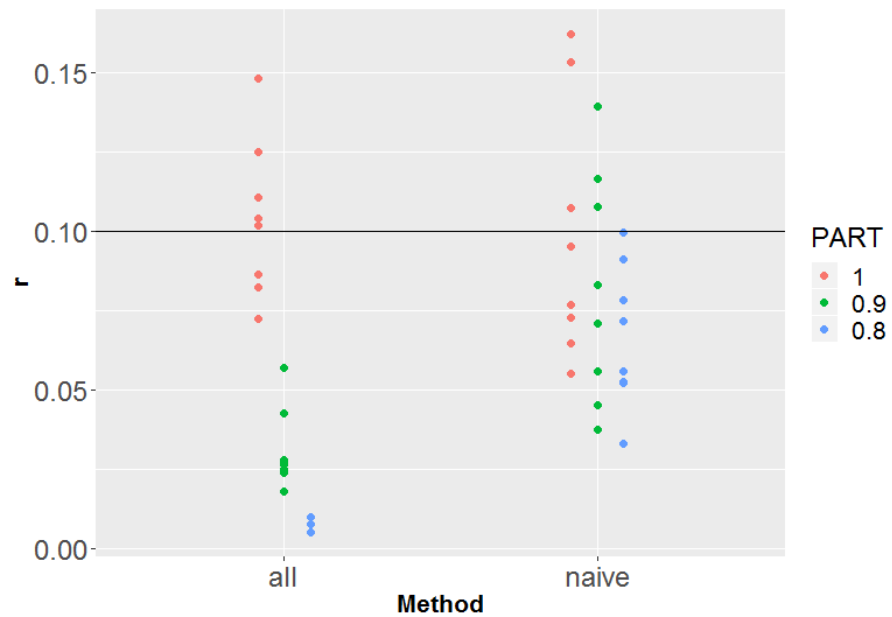


Homogeneous model

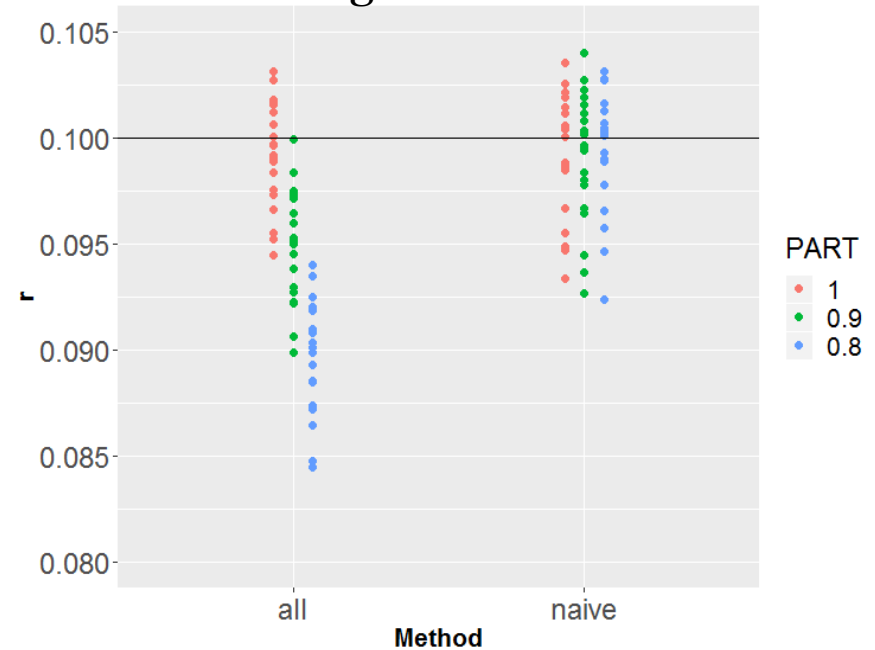


A naive method (removing all the never-attenders)

Non-homogeneous model



Homogeneous model



Given on $P_{NA}=0.1$

Future studies

Method

- Develop a flexible multi-state model
- Age, period or cohort effect
- Informative censoring

Sweden

- Hormone replacement therapy
- Digital mammography
- Different evaluation methods

Collaboration

- Norway
- Finland
- ...

THANKS FOR YOUR ATTENTION

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