

# COM 5120 Communication Theory

## Midterm Exam Answer

November 10, 2022  
15:30 ~ 17:20

1. (13%) The answer is **(c)**, recall the definition of **sufficient statistic**, option (c) only contains  $r_1$  and  $r_2$  clearly not satisfied the definition. ■

2. (15%)

$$P_r\{x[n], n = 1, 2, \dots, N|p\} = p^{\sum x[n]}(1-p)^{N-\sum x[n]}$$

Denote  $\bar{X} = \frac{1}{N} \sum_{n=1}^N x[n]$ ,

$$P_r\{x[n], n = 1, 2, \dots, N|p\} = p^{N\bar{X}}(1-p)^{N-N\bar{X}}$$

To find

$$\begin{aligned} & \max_p P_r\{x[n], n = 1, 2, \dots, N|p\} \\ & \frac{d \ln P_r\{x[n], n = 1, 2, \dots, N|p\}}{dp} = \dots \\ & = \frac{N\bar{X}}{p} + \frac{N - N\bar{X}}{1-p}(-1) = 0 \end{aligned}$$

obtain  $\hat{p} = \bar{X} = \frac{1}{N} \sum_{n=1}^N x[n]$ . ■

3. (15%)

(a)  $S_{\hat{X}}(f) = |-j \operatorname{sgn}(f)|^2 S_X(f) = S_X(f)$ , hence  $R_{\hat{X}}(\tau) = R_X(\tau)$

(b)  $S_{X\hat{X}}(f) = S_X(f)(-j \operatorname{sgn}(f))^* = j \operatorname{sgn}(f) S_X(f)$ , therefore,  $R_{X\hat{X}}(\tau) = -R_X(\tau)$

(c)  $R_Z(\tau) = E \left[ \left( X(t+\tau) + j\hat{X}(t+\tau) \right) \left( X(t) - j\hat{X}(t) \right) \right]$ , expanding we have

$$R_Z(\tau) = R_X(\tau) + R_{\hat{X}}(\tau) - j [R_{X\hat{X}}(\tau) - R_{\hat{X}X}(\tau)]$$

Using  $R_{\hat{X}}(\tau) = R_X(\tau)$ , and the fact that  $R_{X\hat{X}}(\tau) = -\hat{R}_X(\tau)$  is an odd function (since it is the HT of an even signal) we have  $R_{\hat{X}X}(\tau) = R_{X\hat{X}}(-\tau) = -R_{X\hat{X}}(\tau)$ , we have

$$R_Z(\tau) = 2R_X(\tau) - jR_{X\hat{X}}(\tau) = 2R_X(\tau) + j2\hat{R}_X(\tau)$$

Taking FT of both sides we have

$$S_Z(f) = 2S_X(f) + j2(-j \operatorname{sgn}(f) S_X(f)) = 2(1 + \operatorname{sgn}(f)) S_X(f) = 4S_X(f) u_{-1}(f)$$

4. (14%)

(a) To show that the waveforms  $f_n(t)$ ,  $n = 1, 2, 3$  are orthogonal we have to prove that:

$$\int_{-\infty}^{\infty} f_m(t)f_n(t) dt = 0, \quad m \neq n$$

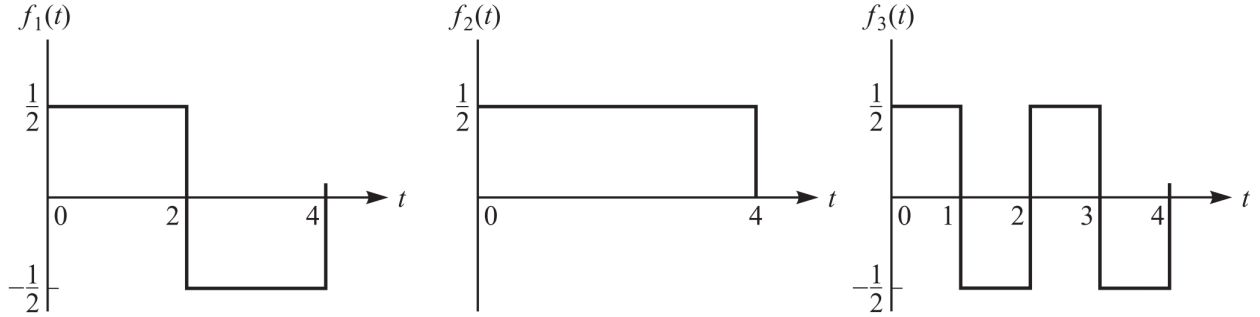


Figure 1: three waveforms  $f_1(t)$ ,  $f_2(t)$ ,  $f_3(t)$

Clearly:

$$\begin{aligned} c_{12} &= \int_{-\infty}^{\infty} f_1(t)f_2(t) dt = \int_0^4 f_1(t)f_2(t) dt = \int_0^2 f_1(t)f_2(t) dt + \int_2^4 f_1(t)f_2(t) dt \\ &= \int_0^2 \frac{1}{2} \cdot \frac{1}{2} dt + \int_2^4 \frac{-1}{2} \cdot \frac{1}{2} dt = \frac{1}{4} \int_0^2 1 dt - \frac{1}{4} \int_2^4 1 dt \\ &= \frac{1}{4} \cdot 2 - \frac{1}{4} \cdot (4 - 2) = 0 \end{aligned}$$

Similarly:

$$\begin{aligned} c_{13} &= \int_{-\infty}^{\infty} f_1(t)f_3(t) dt = \int_0^4 f_1(t)f_3(t) dt \\ &= \int_0^1 f_1(t)f_3(t) dt + \int_1^2 f_1(t)f_3(t) dt + \int_2^3 f_1(t)f_3(t) dt + \int_3^4 f_1(t)f_3(t) dt \\ &= \int_0^1 \frac{1}{2} \cdot \frac{1}{2} dt + \int_1^2 \frac{1}{2} \cdot \frac{-1}{2} dt + \int_2^3 \frac{-1}{2} \cdot \frac{1}{2} dt + \int_3^4 \frac{-1}{2} \cdot \frac{-1}{2} dt \\ &= \frac{1}{4} \int_0^1 1 dt - \frac{1}{4} \int_1^2 1 dt - \frac{1}{4} \int_2^3 1 dt + \frac{1}{4} \int_3^4 1 dt = 0 \end{aligned}$$

and:

$$\begin{aligned} c_{23} &= \int_{-\infty}^{\infty} f_2(t)f_3(t) dt = \int_0^4 f_2(t)f_3(t) dt \\ &= \int_0^1 f_2(t)f_3(t) dt + \int_1^2 f_2(t)f_3(t) dt + \int_2^3 f_2(t)f_3(t) dt + \int_3^4 f_2(t)f_3(t) dt \\ &= \int_0^1 \frac{1}{2} \cdot \frac{1}{2} dt + \int_1^2 \frac{1}{2} \cdot \frac{-1}{2} dt + \int_2^3 \frac{1}{2} \cdot \frac{1}{2} dt + \int_3^4 \frac{1}{2} \cdot \frac{-1}{2} dt \\ &= \frac{1}{4} \int_0^1 1 dt - \frac{1}{4} \int_1^2 1 dt + \frac{1}{4} \int_2^3 1 dt - \frac{1}{4} \int_3^4 1 dt = 0 \end{aligned}$$

Thus, the signals  $f_n(t)$  are orthogonal. It is also straightforward to prove that the signals have unit energy:

$$\int_{-\infty}^{\infty} |f_i(t)|^2 dt = 1, \quad i = 1, 2, 3.$$

Hence, they are orthonormal.

(b) We first determine the weighting coefficients

$$x_n = \int_{-\infty}^{\infty} x(t) f_n(t) dt, \quad n = 1, 2, 3, \quad \text{where } x = \begin{cases} -2, & 0 \leq t < 1 \\ 6, & 1 \leq t < 3 \\ 4, & 3 \leq t < 4 \end{cases}$$

$$\begin{aligned} x_1 &= \int_0^4 x(t) f_1(t) dt \\ &= \int_0^1 x(t) f_1(t) dt + \int_1^2 x(t) f_1(t) dt + \int_2^3 x(t) f_1(t) dt + \int_3^4 x(t) f_1(t) dt \\ &= \int_0^1 -2 \cdot \frac{1}{2} dt + \int_1^2 6 \cdot \frac{1}{2} dt + \int_2^3 6 \cdot \frac{-1}{2} dt + \int_3^4 4 \cdot \frac{-1}{2} dt \\ &= -1 \int_0^1 1 dt + 3 \int_1^2 1 dt - 3 \int_2^3 1 dt - 2 \int_3^4 1 dt = -3 \\ x_2 &= \int_0^4 x(t) f_2(t) dt \\ &= \int_0^1 x(t) f_2(t) dt + \int_1^2 x(t) f_2(t) dt + \int_2^3 x(t) f_2(t) dt + \int_3^4 x(t) f_2(t) dt \\ &= \int_0^1 -2 \cdot \frac{1}{2} dt + \int_1^2 6 \cdot \frac{1}{2} dt + \int_2^3 6 \cdot \frac{1}{2} dt + \int_3^4 4 \cdot \frac{1}{2} dt \\ &= -1 + 3 + 3 + 2 = 7 \\ x_3 &= \int_0^4 x(t) f_3(t) dt \\ &= \int_0^1 x(t) f_3(t) dt + \int_1^2 x(t) f_3(t) dt + \int_2^3 x(t) f_3(t) dt + \int_3^4 x(t) f_3(t) dt \\ &= \int_0^1 -2 \cdot \frac{1}{2} dt + \int_1^2 6 \cdot \frac{-1}{2} dt + \int_2^3 6 \cdot \frac{1}{2} dt + \int_3^4 4 \cdot \frac{-1}{2} dt \\ &= -3 \end{aligned}$$

As it is observed,  $x(t) \neq -3f_1(t) + 7f_2(t) - 3f_3(t)$  and thus it **can not** be represented as a linear combination of these functions. ■

5. (14%)

(a) Consider the QAM constellation of Figure 2. Using the Pythagorean theorem we can find the radius of the inner circle as:

$$a^2 + a^2 = A^2 \implies a = \frac{1}{\sqrt{2}} A$$

The radius of the outer circle can be found using the cosine rule. Since  $b$  is the third side of a triangle with  $a$  and  $A$  the two other sides and angle between them equal to  $\theta = 105^\circ$ , we obtain:

$$b^2 = a^2 + A^2 - 2aA \cos 105^\circ \implies b = \frac{1 + \sqrt{3}}{2}A$$

(b) If we denote by  $r$  the radius of the circle, then using the cosine theorem we obtain:

$$A^2 = r^2 + r^2 - 2r \cos 45^\circ \implies r = \frac{A}{\sqrt{2} - \sqrt{2}}$$

■

6. (14%)

(a) The correlation type demodulator employs a filter:

$$f(t) = \begin{cases} \frac{1}{\sqrt{T}}, & 0 \leq t < T \\ 0, & \text{otherwise} \end{cases}$$

Hence, the sampled outputs of the crosscorrelators are:

$$r = s_m + n, \quad m = 0, 1$$

where  $s_0 = 0$ ,  $s_1 = A\sqrt{T}$  and the noise term  $n$  is a zero-mean Gaussian random variable with variance:

$$\sigma_n^2 = \frac{N_0}{2}$$

The probability density function for the sampled output is:

$$p(r|s_0) = \frac{1}{\sqrt{\pi N_0}} e^{-\frac{r^2}{N_0}}$$

$$p(r|s_1) = \frac{1}{\sqrt{\pi N_0}} e^{-\frac{(r - A\sqrt{T})^2}{N_0}}$$

Since the signals are equally probable, the optimal detector decides in favor of  $s_0$  if

$$PM(\mathbf{r}|\mathbf{s}_0) = p(r|s_0) > p(r|s_1) = PM(\mathbf{r}|\mathbf{s}_1)$$

otherwise it decides in favor of  $s_1$ . The decision rule may be expressed as:

$$\frac{PM(\mathbf{r}, \mathbf{s}_0)}{PM(\mathbf{r}, \mathbf{s}_1)} = e^{\frac{(r - A\sqrt{T})^2 - r^2}{N_0}} = e^{-\frac{(2r - A\sqrt{T})A\sqrt{T}}{N_0}} \begin{matrix} s_0 \\ > \\ < \\ s_1 \end{matrix} 1$$

or equivalently:

$$\begin{array}{c} s_1 \\ r > \frac{1}{2}A\sqrt{T} \\ r < \frac{1}{2}A\sqrt{T} \\ s_0 \end{array}$$

The optimum threshold is  $\frac{1}{2}A\sqrt{T}$ .

(b) The average probability of error is:

$$\begin{aligned} P(e) &= \frac{1}{2}P(e|s_0) + \frac{1}{2}P(e|s_1) \\ &= \frac{1}{2} \int_{\frac{1}{2}A\sqrt{T}}^{\infty} p(r|s_0)dr + \frac{1}{2} \int_{-\infty}^{\frac{1}{2}A\sqrt{T}} p(r|s_0)dr \\ &= \frac{1}{2} \int_{\frac{1}{2}A\sqrt{T}}^{\infty} \frac{1}{\sqrt{\pi N_0}} e^{-\frac{r^2}{N_0}} dr + \frac{1}{2} \int_{-\infty}^{\frac{1}{2}A\sqrt{T}} \frac{1}{\sqrt{\pi N_0}} e^{-\frac{(r-A\sqrt{T})^2}{N_0}} dr \\ &= \frac{1}{2} \int_{\frac{1}{2}\sqrt{\frac{2}{N_0}}A\sqrt{T}}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx + \frac{1}{2} \int_{-\infty}^{-\frac{1}{2}\sqrt{\frac{2}{N_0}}A\sqrt{T}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= Q\left[\frac{1}{2}\sqrt{\frac{2}{N_0}}A\sqrt{T}\right] = Q[\sqrt{\text{SNR}}], \text{ where } \text{SNR} = \frac{\frac{1}{2}A^2T}{N_0} \end{aligned}$$

Thus, the on-off signaling requires a factor of two more energy to achieve the same probability of error as the antipodal signaling. ■

7. (15%) The PDF of the noise  $n$  is:

$$p(n) = \frac{1}{\sqrt{2\sigma}} e^{-|n|\sqrt{\frac{2}{\sigma}}} = \frac{\lambda}{2} e^{-\lambda|n|}, \text{ where } \lambda = \sqrt{\frac{2}{\sigma}}$$

The optimal receiver uses the criterion:

$$\frac{p(r|A)}{p(r|-A)} = e^{-\lambda[|r-A|-|r+A|]} \begin{array}{ccc} A & & A \\ > & 1 \implies r & > \\ < & & < \\ -A & & -A \end{array} 0$$

The average probability of error is:

$$\begin{aligned} P(e) &= \frac{1}{2}P(e|A) + \frac{1}{2}P(e|-A) \\ &= \frac{1}{2} \int_{-\infty}^0 f(r|A)dr + \frac{1}{2} \int_0^{\infty} f(r|-A)dr \\ &= \frac{1}{2} \int_{-\infty}^0 \frac{\lambda}{2} e^{-\lambda|r-A|}dr + \frac{1}{2} \int_0^{\infty} \frac{\lambda}{2} e^{-\lambda|r+A|}dr \\ &= \frac{\lambda}{4} \int_{-\infty}^A e^{-\lambda|x|}dx + \frac{\lambda}{4} \int_A^{\infty} e^{-\lambda|x|}dx \\ &= 2 \cdot \frac{\lambda}{4} \int_A^{\infty} e^{-\lambda|x|}dx \\ &= \frac{\lambda}{2} \cdot \frac{-1}{\lambda} \cdot (e^{-\lambda\infty} - e^{-\lambda A}) = \frac{-1}{2} \cdot (0 - e^{-\lambda A}) \\ &= \frac{1}{2}e^{-\lambda A} = \frac{1}{2}e^{-\sqrt{\frac{2}{\sigma}}A} \end{aligned}$$

■